

Representing and comparing probabilities with kernels: Part 2

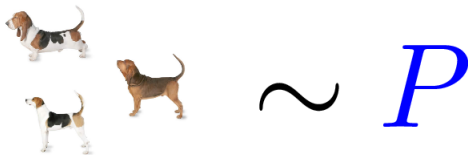
Arthur Gretton

Gatsby Computational Neuroscience Unit,
University College London

MLSS Madrid, 2018

Comparing two samples

- **Given:** Samples from unknown distributions P and Q .
- **Goal:** do P and Q differ?



Outline

Two sample testing

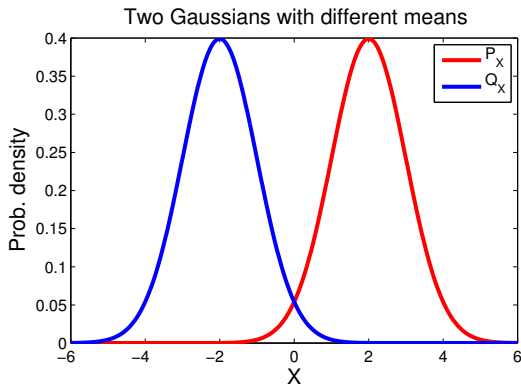
- Test statistic: Maximum Mean Discrepancy (MMD)...
 - ...as a difference in feature means
 - ...as an integral probability metric (not just a technicality!)
- Statistical testing with the MMD
- “How to choose the best kernel”

Training GANs with MMD

Maximum Mean Discrepancy

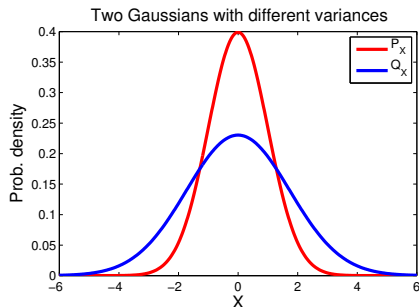
Feature mean difference

- Simple example: 2 Gaussians with different means
- Answer: t-test



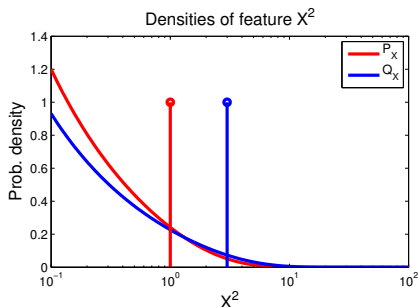
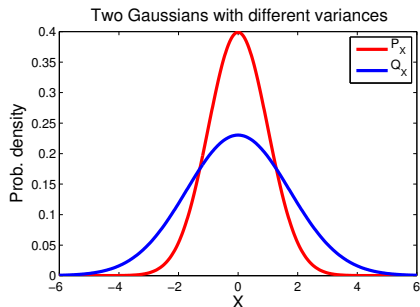
Feature mean difference

- Two Gaussians with same means, different variance
- Idea: look at difference in **means of features** of the RVs
- In Gaussian case: second order features of form $\varphi(x) = x^2$



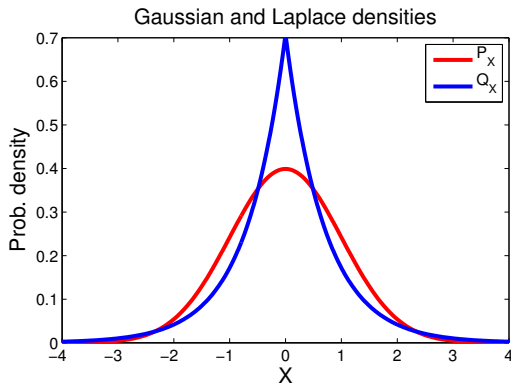
Feature mean difference

- Two Gaussians with same means, different variance
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Feature mean difference

- Gaussian and Laplace distributions
- Same mean *and* same variance
- Difference in means using **higher order features**...RKHS



Infinitely many features using kernels

**Kernels: dot products
of features**

Feature map $\varphi(x) \in \mathcal{F}$,

$$\varphi(x) = [\dots \varphi_i(x) \dots] \in \ell_2$$

For **positive definite** k ,

$$k(x, x') = \langle \varphi(x), \varphi(x') \rangle_{\mathcal{F}}$$

Infinitely many features
 $\varphi(x)$, dot product in
closed form!

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 $\varphi(x)$, dot product in
closed form!

Exponentiated quadratic kernel

$$k(x, x') = \exp \left(-\gamma \|x - x'\|^2 \right)$$

$$\varphi(x) = \begin{bmatrix} \varphi_1(x) \\ \varphi_2(x) \\ \varphi_3(x) \\ \vdots \end{bmatrix}$$

Infinitely many features of *distributions*

Given P a Borel **probability measure** on \mathcal{X} , define **feature map of probability P** ,

$$\mu_P = [\dots \mathbf{E}_P [\varphi_i(X)] \dots]$$

For **positive definite** $k(x, x')$,

$$\langle \mu_P, \mu_Q \rangle_{\mathcal{F}} = \mathbf{E}_{P, Q} k(x, y)$$

for $x \sim P$ and $y \sim Q$.

Fine print: feature map $\varphi(x)$ must be Bochner integrable for all probability measures considered.
Always true if kernel bounded.

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Fine print: feature map $\varphi(\mathbf{x})$ must be Bochner integrable for all probability measures considered.
Always true if kernel bounded.

The maximum mean discrepancy

The maximum mean discrepancy is the distance between **feature means**:

$$\begin{aligned}MMD^2(P, Q) &= \|\mu_P - \mu_Q\|_{\mathcal{F}}^2 \\&= \langle \mu_P, \mu_P \rangle_{\mathcal{F}} + \langle \mu_Q, \mu_Q \rangle_{\mathcal{F}} - 2 \langle \mu_P, \mu_Q \rangle_{\mathcal{F}} \\&= \underbrace{\mathbf{E}_P k(X, X')}_{(a)} + \underbrace{\mathbf{E}_Q k(Y, Y')}_{(a)} - 2 \underbrace{\mathbf{E}_{P, Q} k(X, Y)}_{(b)}\end{aligned}$$

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(a)= within distrib. similarity, (b)= cross-distrib. similarity.

Illustration of MMD

- Dogs ($= P$) and fish ($= Q$) example revisited
- Each entry is one of $k(\text{dog}_i, \text{dog}_j)$, $k(\text{dog}_i, \text{fish}_j)$, or $k(\text{fish}_i, \text{fish}_j)$

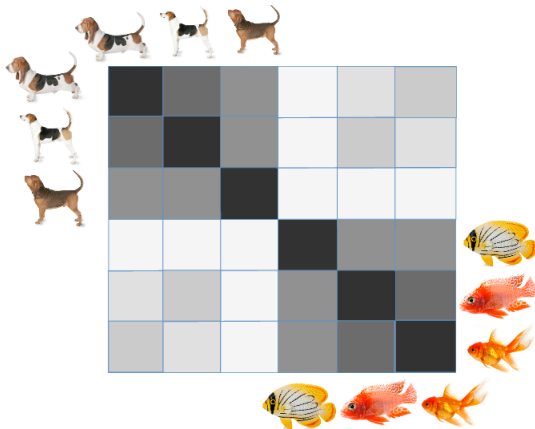
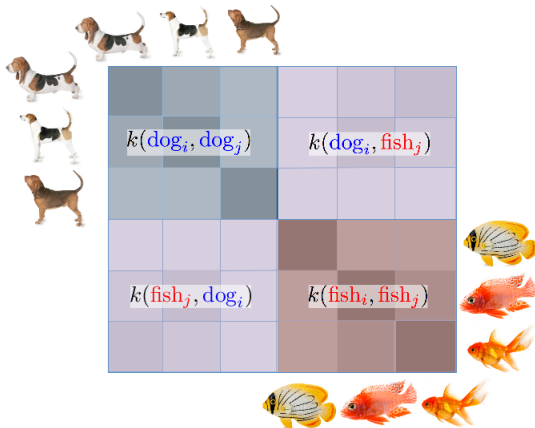


Illustration of MMD

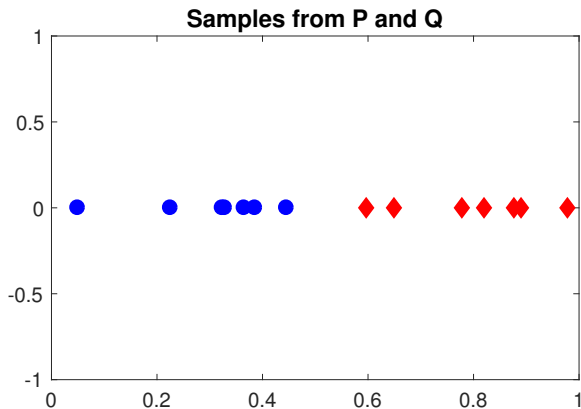
The maximum mean discrepancy:

$$\widehat{MMD}^2 = \frac{1}{n(n-1)} \sum_{i \neq j} k(\text{dog}_i, \text{dog}_j) + \frac{1}{n(n-1)} \sum_{i \neq j} k(\text{fish}_i, \text{fish}_j) - \frac{2}{n^2} \sum_{i,j} k(\text{dog}_i, \text{fish}_j)$$



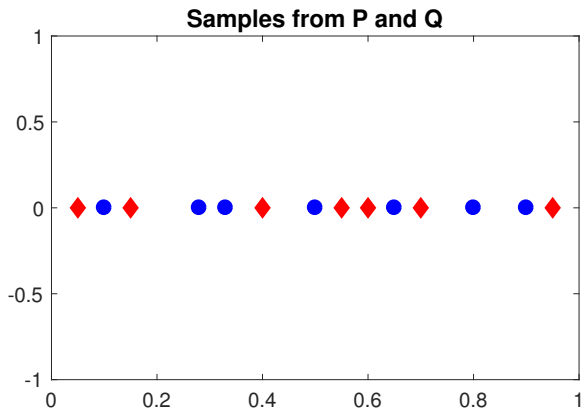
MMD as an integral probability metric

Are P and Q different?



MMD as an integral probability metric

Are P and Q different?

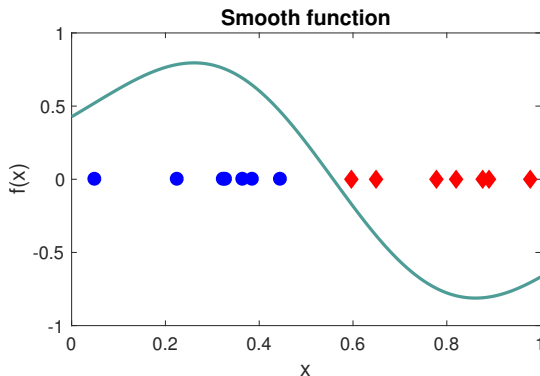


MMD as an integral probability metric

Integral probability metric:

Find a "well behaved function" $f(x)$ to maximize

$$\mathbf{E}_P f(X) - \mathbf{E}_Q f(Y)$$

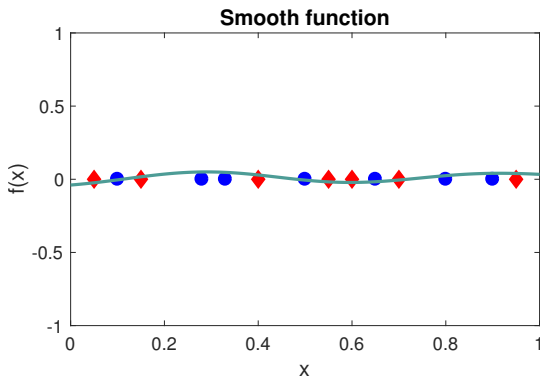


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MMD as an integral probability metric

Maximum mean discrepancy: smooth function for P vs Q

$$MMD(P, Q; \mathcal{F}) := \sup_{\|f\| \leq 1} [\mathbf{E}_P f(X) - \mathbf{E}_Q f(Y)]$$

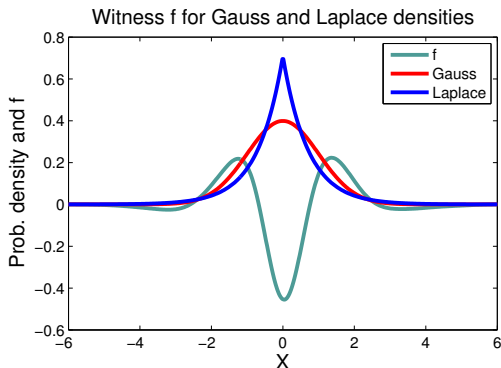
(\mathcal{F} = unit ball in RKHS \mathcal{F})

MMD as an integral probability metric

Maximum mean discrepancy: smooth function for P vs Q

$$MMD(P, Q; F) := \sup_{\|f\| \leq 1} [\mathbf{E}_P f(X) - \mathbf{E}_Q f(Y)]$$

(F = unit ball in RKHS \mathcal{F})



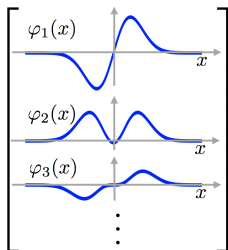
MMD as an integral probability metric

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(F = unit ball in RKHS \mathcal{F})

Functions are linear combinations of features:

$$f(x) = \langle f, \varphi(x) \rangle_{\mathcal{F}} = \sum_{\ell=1}^{\infty} f_{\ell} \varphi_{\ell}(x) = \begin{bmatrix} f_1 \\ f_2 \\ f_3 \\ \vdots \end{bmatrix}^{\top} \begin{bmatrix} \varphi_1(x) \\ \varphi_2(x) \\ \varphi_3(x) \\ \vdots \end{bmatrix}$$


MMD as an integral probability metric

Maximum mean discrepancy: smooth function for P vs Q

$$MMD(P, Q; F) := \sup_{\|f\| \leq 1} [\mathbf{E}_P f(X) - \mathbf{E}_Q f(Y)]$$

(F = unit ball in RKHS \mathcal{F})

Expectations of functions are linear combinations of expected features

$$\mathbf{E}_P(f(X)) = \langle f, \mathbf{E}_P \varphi(X) \rangle_{\mathcal{F}} = \langle f, \mu_P \rangle_{\mathcal{F}}$$

(always true if kernel is bounded)

MMD as an integral probability metric

Maximum mean discrepancy: smooth function for P vs Q

$$MMD(P, Q; \mathcal{F}) := \sup_{\|f\| \leq 1} [\mathbf{E}_P f(X) - \mathbf{E}_Q f(Y)]$$

(\mathcal{F} = unit ball in RKHS \mathcal{F})

For characteristic RKHS \mathcal{F} , $MMD(P, Q; \mathcal{F}) = 0$ iff $P = Q$

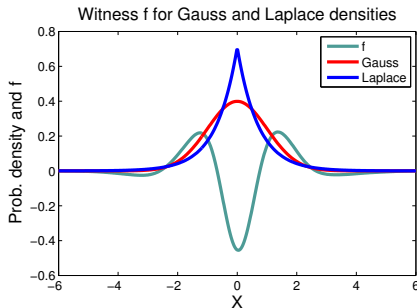
Other choices for witness function class:

- Bounded continuous [Dudley, 2002]
- Bounded variation 1 (Kolmogorov metric) [Müller, 1997]
- Bounded Lipschitz (Wasserstein distances) [Dudley, 2002]

Integral prob. metric vs feature difference

The MMD:

$$\begin{aligned} &MMD(P, Q; F) \\ &= \sup_{f \in F} [\mathbf{E}_P f(X) - \mathbf{E}_Q f(Y)] \end{aligned}$$



Integral prob. metric vs feature difference

The MMD:

use

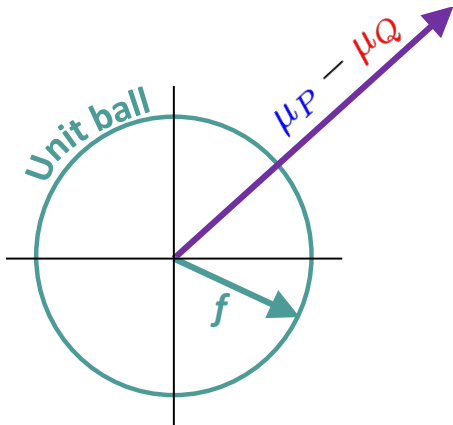
$$\begin{aligned} \text{MMD}(P, Q; F) \\ &= \sup_{f \in F} [\mathbf{E}_P f(X) - \mathbf{E}_Q f(Y)] \\ &= \sup_{f \in F} \langle f, \mu_P - \mu_Q \rangle_{\mathcal{F}} \end{aligned}$$

$$\mathbf{E}_P f(X) = \langle \mu_P, f \rangle_{\mathcal{F}}$$

Integral prob. metric vs feature difference

The MMD:

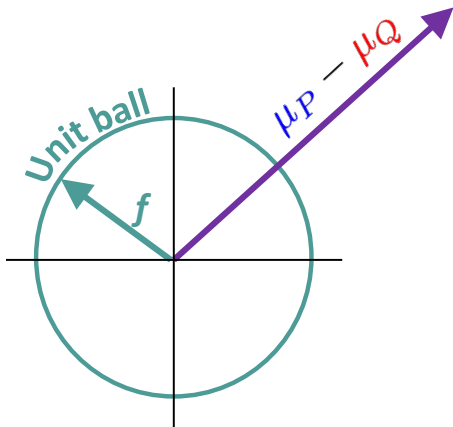
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Integral prob. metric vs feature difference

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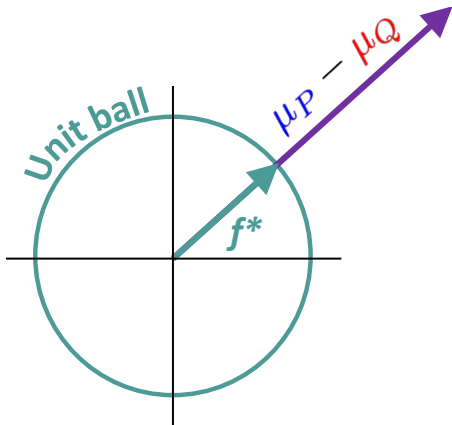
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$$f^* = \frac{\mu_P - \mu_Q}{\|\mu_P - \mu_Q\|}$$

Integral prob. metric vs feature difference

The MMD:

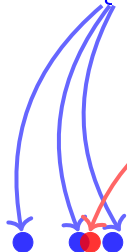
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Function view and feature view equivalent

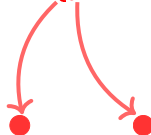
Construction of MMD witness

Construction of empirical **witness function** (proof: next slide!)

Observe $X = \{x_1, \dots, x_n\} \sim P$

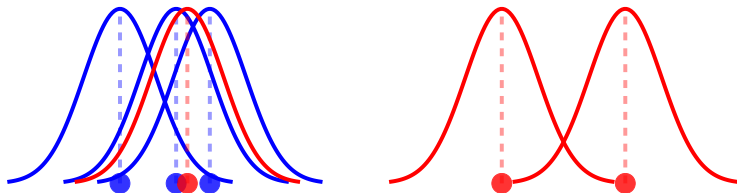


Observe $Y = \{y_1, \dots, y_n\} \sim Q$



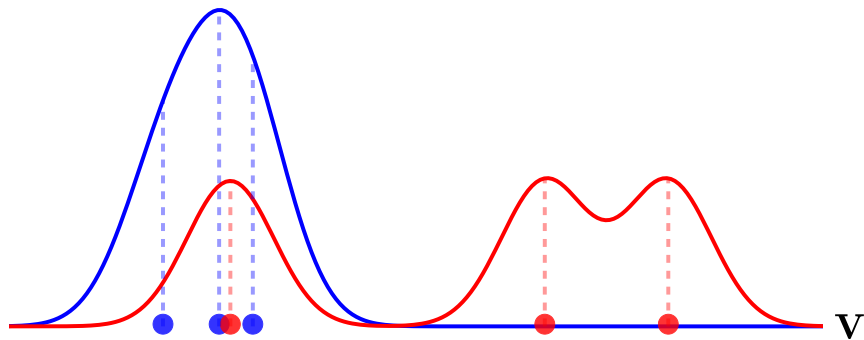
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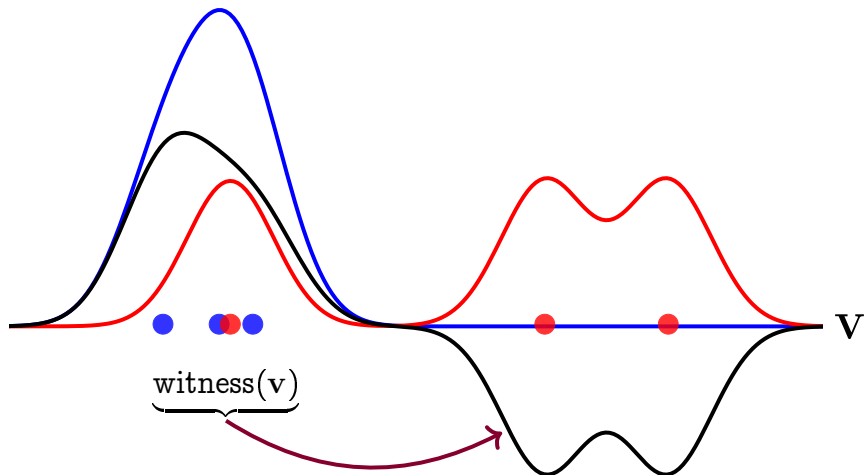
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Derivation of empirical witness function

Recall the witness function expression

$$f^* \propto \mu_P - \mu_Q$$

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The empirical feature mean for P

$$\hat{\mu}_P := \frac{1}{n} \sum_{i=1}^n \varphi(x_i)$$

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The empirical witness function at v

$$f^*(v) = \langle f^*, \varphi(v) \rangle_{\mathcal{F}}$$

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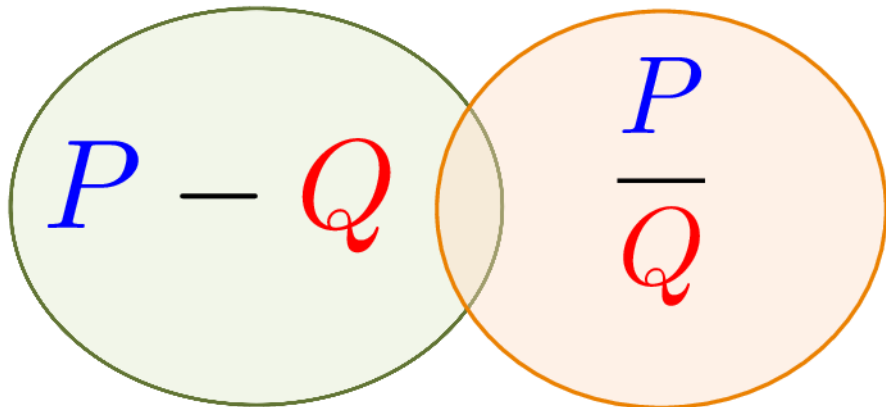
The empirical witness function at v

$$\begin{aligned} f^*(v) &= \langle f^*, \varphi(v) \rangle_{\mathcal{F}} \\ &\propto \langle \hat{\mu}_P - \hat{\mu}_Q, \varphi(v) \rangle_{\mathcal{F}} \\ &= \frac{1}{n} \sum_{i=1}^n k(\mathbf{x}_i, v) - \frac{1}{n} \sum_{i=1}^n k(\mathbf{y}_i, v) \end{aligned}$$

Don't need explicit feature coefficients $f^* := \begin{bmatrix} f_1^* & f_2^* & \dots \end{bmatrix}$

Interlude: divergence measures

Divergences



Divergences

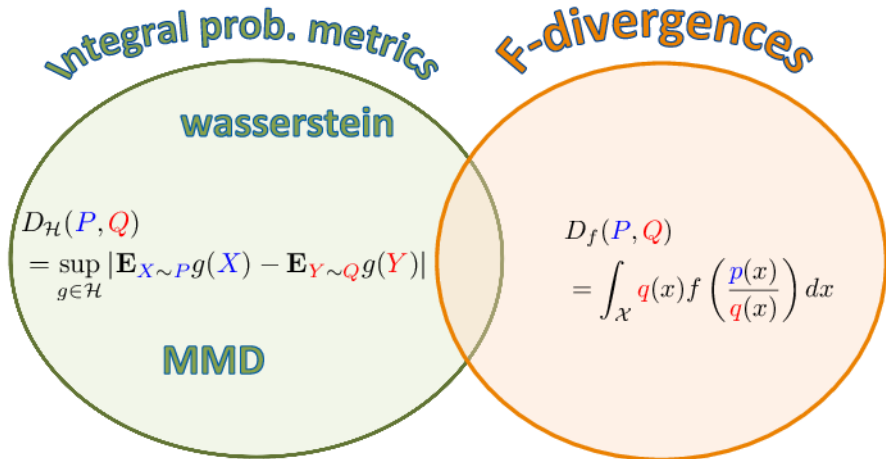
Integral prob. metrics

$$D_{\mathcal{H}}(P, Q) \\ = \sup_{g \in \mathcal{H}} |\mathbf{E}_{X \sim P} g(X) - \mathbf{E}_{Y \sim Q} g(Y)|$$

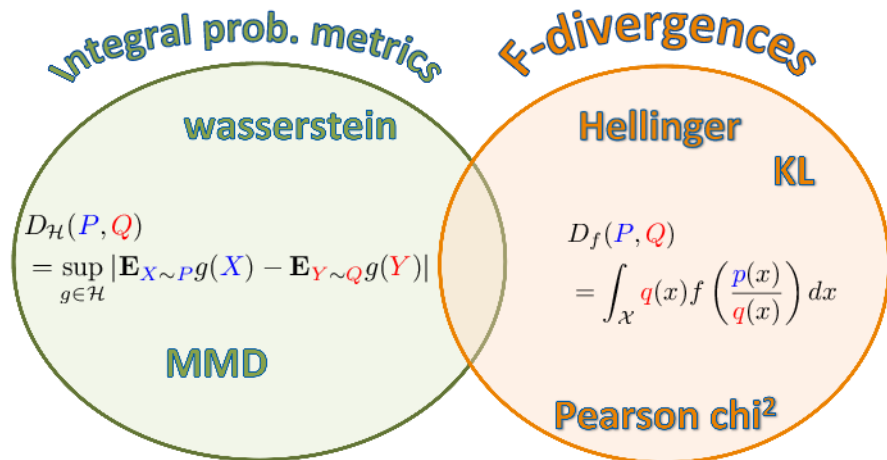
f-divergences

$$D_f(P, Q) \\ = \int_{\mathcal{X}} q(x) f\left(\frac{p(x)}{q(x)}\right) dx$$

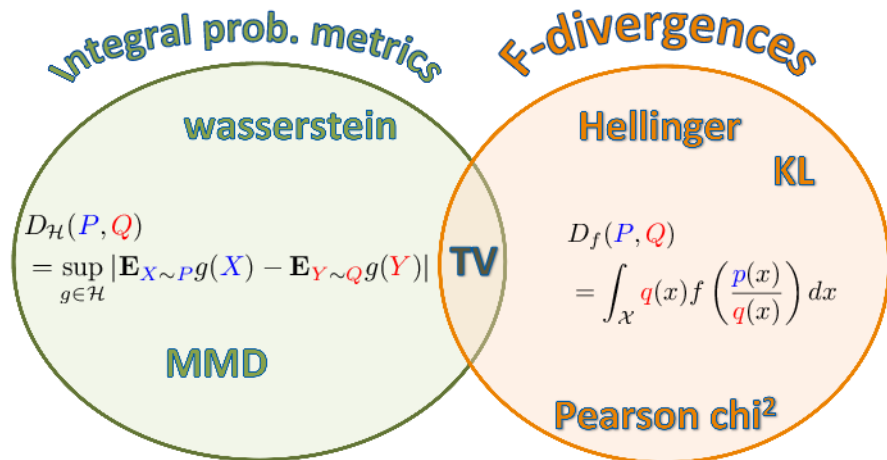
Divergences



Divergences



Divergences



Sriperumbudur, Fukumizu, G, Schoelkopf, Lanckriet (2012)

Two-Sample Testing with MMD

A statistical test using MMD

The empirical MMD:

$$\widehat{MMD}^2 = \frac{1}{n(n-1)} \sum_{i \neq j} k(\mathbf{x}_i, \mathbf{x}_j) + \frac{1}{n(n-1)} \sum_{i \neq j} k(\mathbf{y}_i, \mathbf{y}_j) - \frac{2}{n^2} \sum_{i,j} k(\mathbf{x}_i, \mathbf{y}_j)$$

How does this help decide whether $P = Q$?

A statistical test using MMD

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Perspective from statistical hypothesis testing:

- Null hypothesis \mathcal{H}_0 when $P = Q$
 - should see \widehat{MMD}^2 “close to zero”.
- Alternative hypothesis \mathcal{H}_1 when $P \neq Q$
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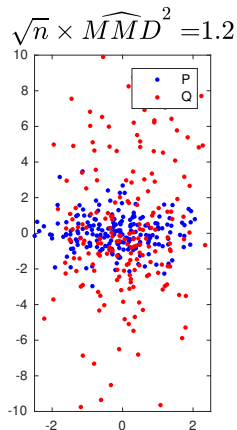
Want Threshold c_α for \widehat{MMD}^2 to get false positive rate α

Behaviour of \widehat{MMD}^2 when $P \neq Q$

Draw $n = 200$ i.i.d samples from P and Q

■ Laplace with different y-variance.

■ $\sqrt{n} \times \widehat{MMD}^2 = 1.2$

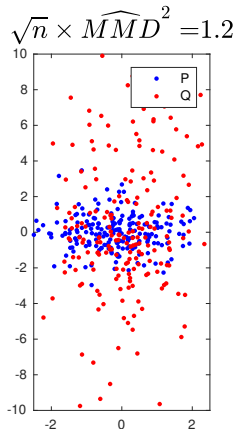
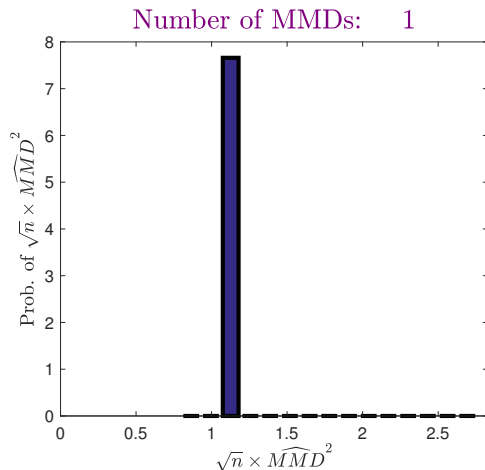


Behaviour of \widehat{MMD}^2 when $P \neq Q$

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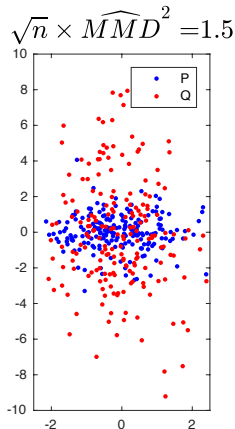
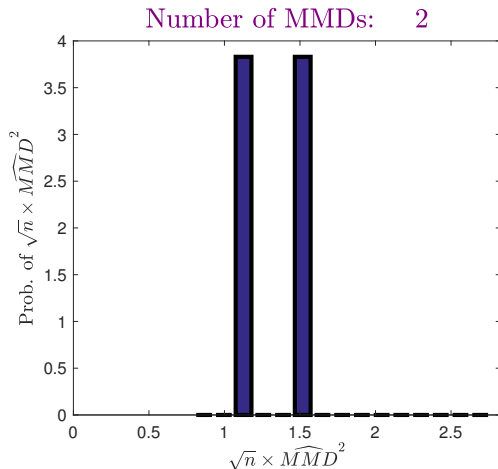


Behaviour of \widehat{MMD}^2 when $P \neq Q$

Draw $n = 200$ **new** samples from P and Q

■ Laplace with different y-variance.

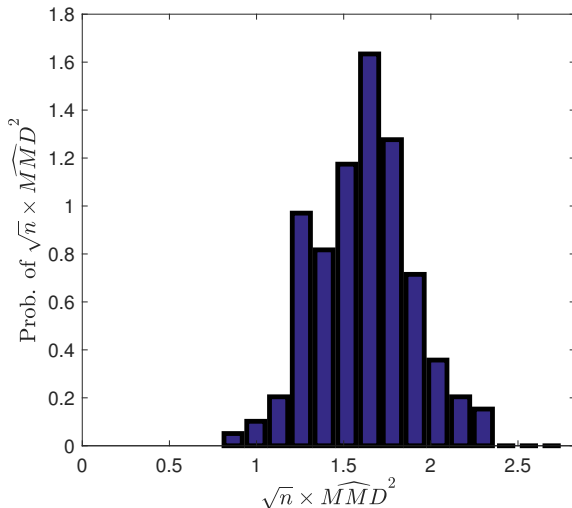
■ $\sqrt{n} \times \widehat{MMD}^2 = 1.5$



Behaviour of \widehat{MMD}^2 when $P \neq Q$

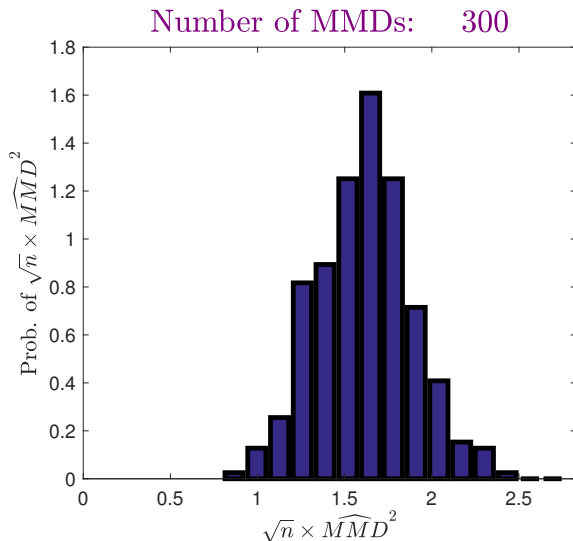
Repeat this 150 times ...

Number of MMDs: 150



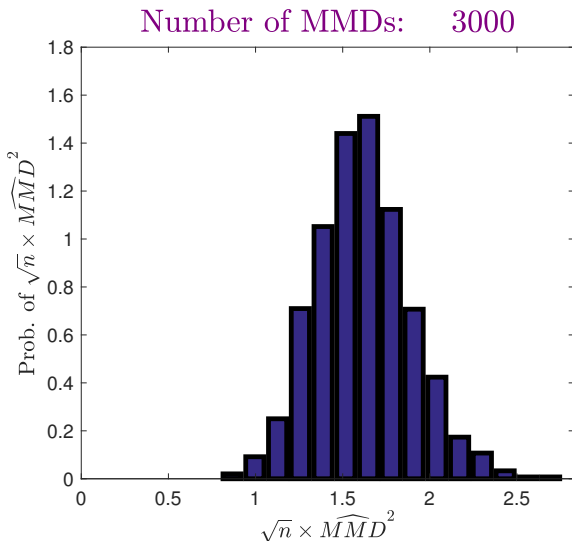
Behaviour of \widehat{MMD}^2 when $P \neq Q$

Repeat this 300 times ...



Behaviour of \widehat{MMD}^2 when $P \neq Q$

Repeat this 3000 times ...



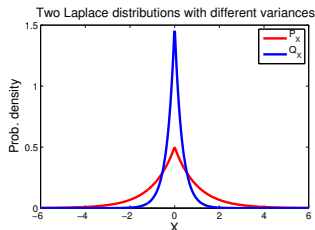
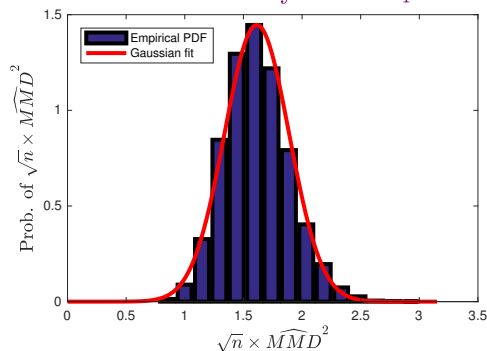
Asymptotics of \widehat{MMD}^2 when $P \neq Q$

When $P \neq Q$, statistic is asymptotically normal,

$$\frac{\widehat{MMD}^2 - \text{MMD}(P, Q)}{\sqrt{V_n(P, Q)}} \xrightarrow{D} \mathcal{N}(0, 1),$$

where variance $V_n(P, Q) = O(n^{-1})$.

MMD density under \mathcal{H}_1

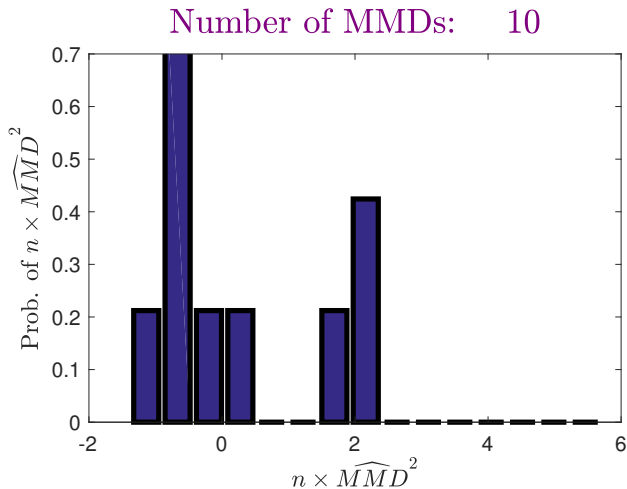


Behaviour of \widehat{MMD}^2 when $P = Q$

What happens when P and Q are the same?

Behaviour of \widehat{MMD}^2 when $P = Q$

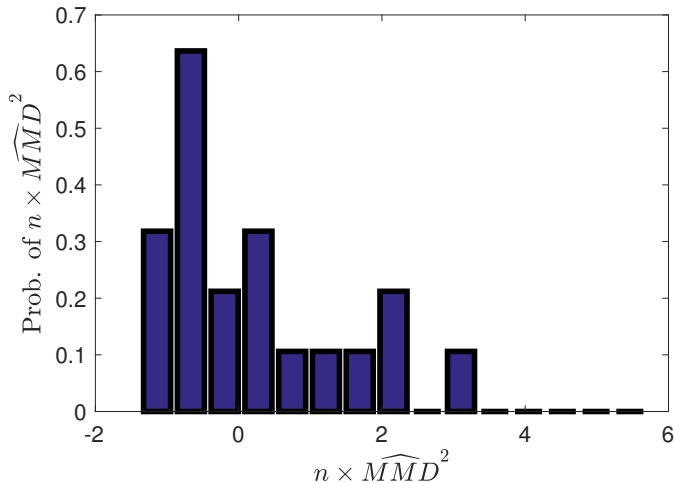
■ Case of $P = Q = \mathcal{N}(0, 1)$



Behaviour of \widehat{MMD}^2 when $P = Q$

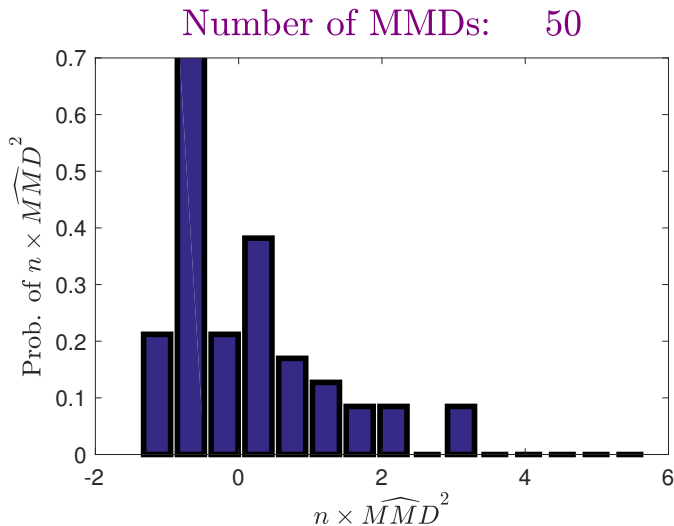
■ Case of $P = Q = \mathcal{N}(0, 1)$

Number of MMDs: 20



Behaviour of \widehat{MMD}^2 when $P = Q$

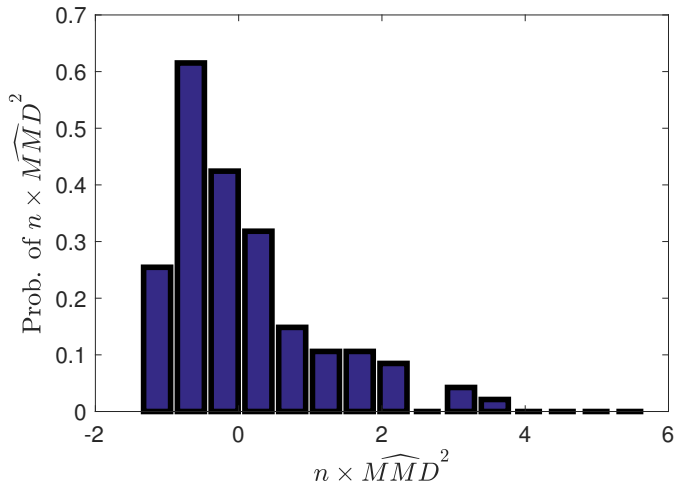
■ Case of $P = Q = \mathcal{N}(0, 1)$



Behaviour of \widehat{MMD}^2 when $P = Q$

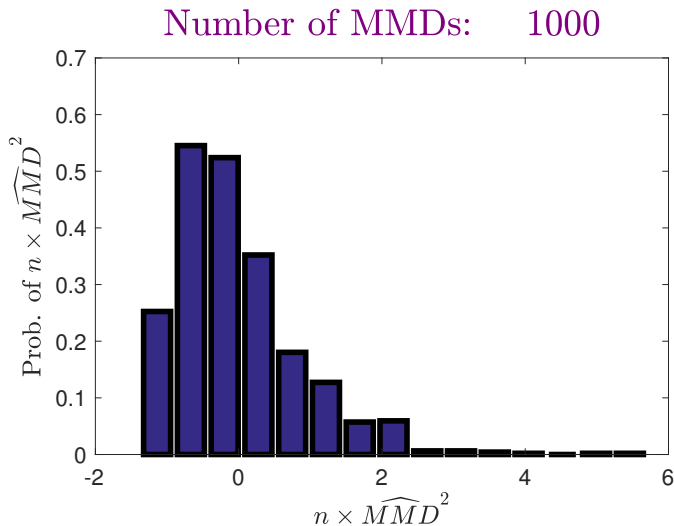
■ Case of $P = Q = \mathcal{N}(0, 1)$

Number of MMDs: 100



Behaviour of \widehat{MMD}^2 when $P = Q$

■ Case of $P = Q = \mathcal{N}(0, 1)$



Asymptotics of \widehat{MMD}^2 when $P = Q$

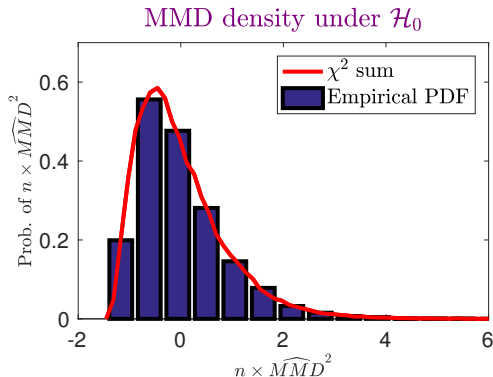
Where $P = Q$, statistic has asymptotic distribution

$$n\widehat{MMD}^2 \sim \sum_{l=1}^{\infty} \lambda_l [z_l^2 - 2]$$

where

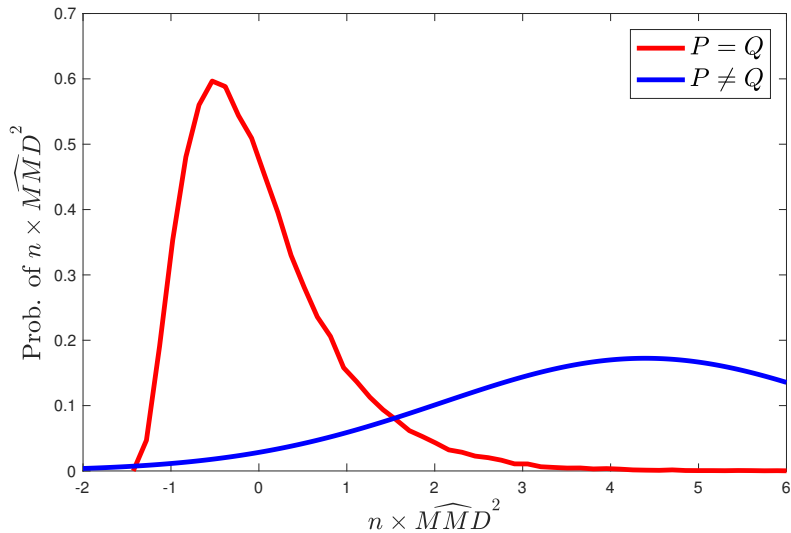
$$\lambda_i \psi_i(x') = \int_{\mathcal{X}} \underbrace{\tilde{k}(x, x')}_{\text{centred}} \psi_i(x) dP(x)$$

$$z_l \sim \mathcal{N}(0, 2) \quad \text{i.i.d.}$$



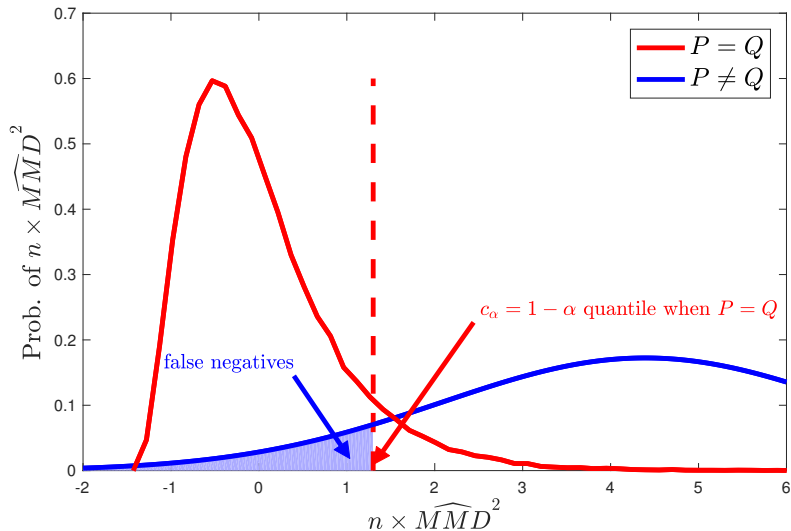
A statistical test

A summary of the asymptotics:



A statistical test

Test construction: (G., Borgwardt, Rasch, Schoelkopf, and Smola, JMLR 2012)



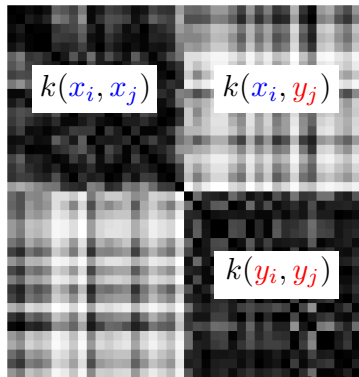
How do we get test threshold c_α ?

Original empirical MMD for dogs and fish:

$$X = \left[\text{dog1} \quad \text{dog2} \quad \text{dog3} \quad \dots \right]$$

$$Y = \left[\text{fish1} \quad \text{fish2} \quad \text{fish3} \quad \dots \right]$$

$$\begin{aligned} \widehat{MMD}^2 &= \frac{1}{n(n-1)} \sum_{i \neq j} k(\mathbf{x}_i, \mathbf{x}_j) \\ &\quad + \frac{1}{n(n-1)} \sum_{i \neq j} k(\mathbf{y}_i, \mathbf{y}_j) \\ &\quad - \frac{2}{n^2} \sum_{i,j} k(\mathbf{x}_i, \mathbf{y}_j) \end{aligned}$$



How do we get test threshold c_α ?

Permuted dog and fish samples (**merdogs**):

$$\tilde{X} = \left[\begin{array}{c} \text{fish} \quad \text{dog} \quad \text{fish} \quad \dots \end{array} \right]$$

$$\tilde{Y} = \left[\begin{array}{c} \text{dog} \quad \text{fish} \quad \text{dog} \quad \dots \end{array} \right]$$

How do we get test threshold c_α ?

Permuted dog and fish samples (**merdogs**):

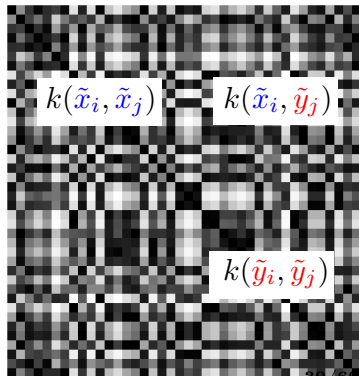
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Permutation simulates

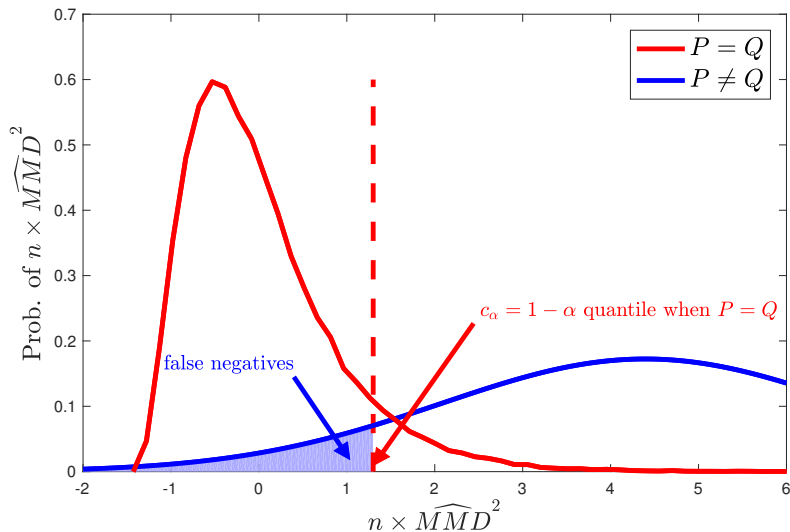
$$P = Q$$



How to choose the best kernel (1)
optimising the kernel parameters

Graphical illustration

- Maximising test power same as minimizing false negatives



Optimizing kernel for test power

The power of our test (\Pr_1 denotes probability under $P \neq Q$):

$$\Pr_1 \left(\widehat{n\text{MMD}}^2 > \hat{c}_\alpha \right)$$

Optimizing kernel for test power

The power of our test (\Pr_1 denotes probability under $P \neq Q$):

$$\begin{aligned} & \Pr_1 \left(\widehat{n\text{MMD}}^2 > \hat{c}_\alpha \right) \\ & \rightarrow \Phi \left(\frac{n\text{MMD}^2(P, Q)}{\sqrt{V_n(P, Q)}} - \frac{c_\alpha}{\sqrt{V_n(P, Q)}} \right) \end{aligned}$$

where

- Φ is the CDF of the standard normal distribution.
- \hat{c}_α is an estimate of c_α test threshold.

Optimizing kernel for test power

The power of our test (\Pr_1 denotes probability under $P \neq Q$):

$$\begin{aligned} & \Pr_1 \left(\widehat{n\text{MMD}}^2 > \hat{c}_\alpha \right) \\ & \rightarrow \Phi \left(\underbrace{\frac{\text{MMD}^2(P, Q)}{\sqrt{V_n(P, Q)}}}_{O(n^{1/2})} - \underbrace{\frac{c_\alpha}{n\sqrt{V_n(P, Q)}}}_{O(n^{-1/2})} \right) \end{aligned}$$

Variance under \mathcal{H}_1 decreases as $\sqrt{V_n(P, Q)} \sim O(n^{-1/2})$

For large n , second term negligible!

Optimizing kernel for test power

The power of our test (\Pr_1 denotes probability under $P \neq Q$):

$$\begin{aligned} & \Pr_1 \left(n \widehat{\text{MMD}}^2 > \hat{c}_\alpha \right) \\ & \rightarrow \Phi \left(\frac{\text{MMD}^2(P, Q)}{\sqrt{V_n(P, Q)}} - \frac{c_\alpha}{n \sqrt{V_n(P, Q)}} \right) \end{aligned}$$

To maximize test power, maximize

$$\frac{\text{MMD}^2(P, Q)}{\sqrt{V_n(P, Q)}}$$

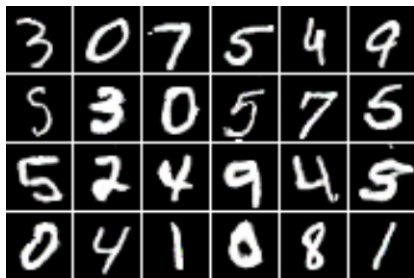
(Sutherland, Tung, Strathmann, De, Ramdas, Smola, G., ICLR 2017)

Code: github.com/dougalsutherland/opt-mmd

Troubleshooting for generative adversarial networks



MNIST samples

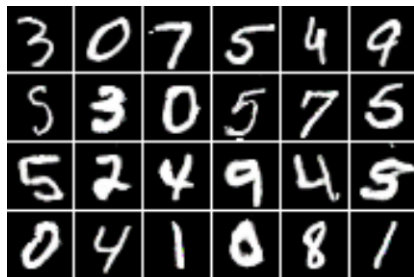


Samples from a GAN

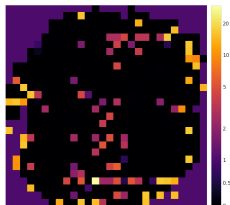
Troubleshooting for generative adversarial networks



MNIST samples



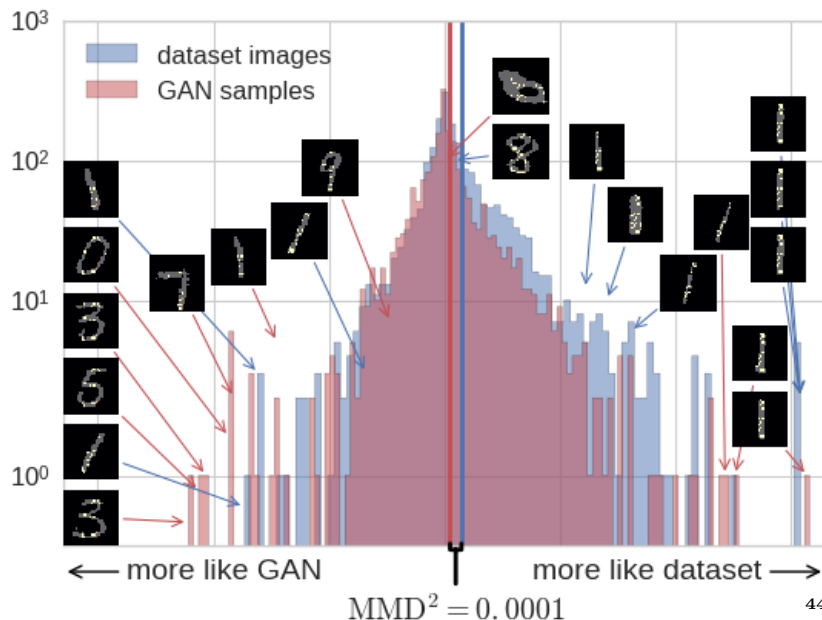
Samples from a GAN



ARD map

- Power for **optimized ARD kernel**: 1.00 at $\alpha = 0.01$
- Power for optimized RBF kernel: 0.57 at $\alpha = 0.01$

Troubleshooting generative adversarial networks



How to choose the best kernel (2)

characteristic kernels

Characteristic kernels

Characteristic: MMD a metric $MMD = 0$ iff $P = Q$)

[NIPS07b, JMLR10]

In the next slides:

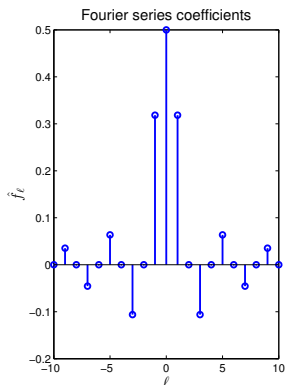
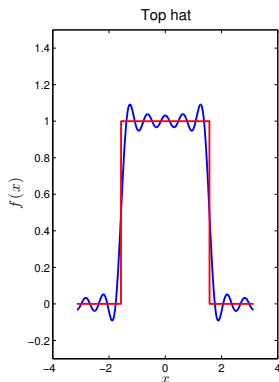
- Characteristic property on $[-\pi, \pi]$ with periodic boundary
- Characteristic property on \mathbb{R}^d

Characteristic kernels on $[-\pi, \pi]$

Reminder: **Fourier series**

Function on $[-\pi, \pi]$ with periodic boundary.

$$f(x) = \sum_{\ell=-\infty}^{\infty} \hat{f}_{\ell} \exp(\imath \ell x) = \sum_{\ell=-\infty}^{\infty} \hat{f}_{\ell} (\cos(\ell x) + \imath \sin(\ell x)).$$

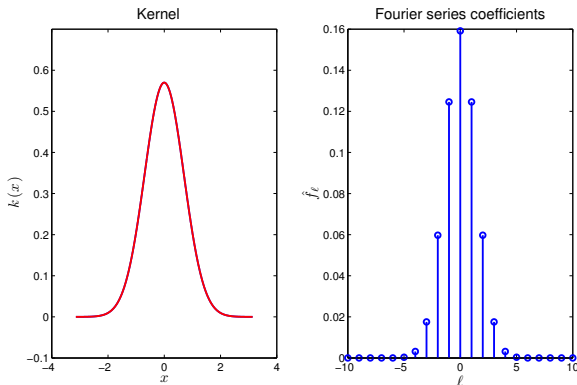


Characteristic kernels on $[-\pi, \pi]$

Jacobi theta kernel (close to exponentiated quadratic):

$$k(x - y) = \frac{1}{2\pi} \vartheta \left(\frac{x - y}{2\pi}, \frac{i\sigma^2}{2\pi} \right), \quad \hat{k}_\ell = \frac{1}{2\pi} \exp \left(-\frac{\sigma^2 \ell^2}{2} \right).$$

ϑ is the Jacobi theta function, close to Gaussian when σ^2 small



The MMD in a Fourier representation

Maximum mean embedding via Fourier series:

- Fourier series for P is characteristic function $\varphi_{P,\ell}$
- Fourier series for mean embedding is product of fourier series!
(convolution theorem)

$$\begin{aligned}\mu_P(x) &= \langle \mu_P, k(\cdot, x) \rangle_{\mathcal{F}} \\ &= E_{X \sim P} k(X - x) \\ &= \int_{-\pi}^{\pi} k(x - t) dP(t) \quad \hat{\mu}_{P,\ell} = \hat{k}_{\ell} \times \bar{\varphi}_{P,\ell}\end{aligned}$$

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MMD can be written in terms of Fourier series:

$$\begin{aligned}MMD(P, Q; F) &= \|\mu_P - \mu_Q\|_{\mathcal{F}} \\ &= \left\| \sum_{\ell=-\infty}^{\infty} [(\bar{\varphi}_{P,\ell} - \bar{\varphi}_{Q,\ell}) \hat{k}_\ell] \exp(i\ell x) \right\|_{\mathcal{F}}\end{aligned}$$

A simpler Fourier representation for MMD

From previous slide,

$$MMD(P, Q; F) = \left\| \sum_{\ell=-\infty}^{\infty} [(\bar{\varphi}_{P,\ell} - \bar{\varphi}_{Q,\ell}) \hat{k}_{\ell}] \exp(i\ell x) \right\|_{\mathcal{F}}$$

Reminder: the squared norm of a function f in \mathcal{F} is:

$$\|f\|_{\mathcal{F}}^2 = \sum_{\ell=-\infty}^{\infty} \frac{|\hat{f}_{\ell}|^2}{\hat{k}_{\ell}}.$$

Simple, interpretable expression for squared MMD:

$$MMD^2(P, Q; F) = \sum_{\ell=-\infty}^{\infty} \frac{[|\varphi_{P,\ell} - \varphi_{Q,\ell}|^2 \hat{k}_{\ell}]^2}{\hat{k}_{\ell}} = \sum_{\ell=-\infty}^{\infty} |\varphi_{P,\ell} - \varphi_{Q,\ell}|^2 \hat{k}_{\ell}$$

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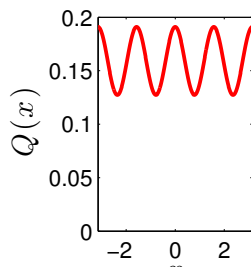
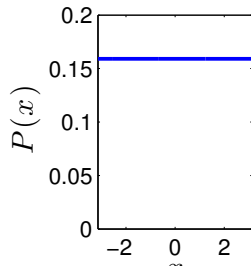
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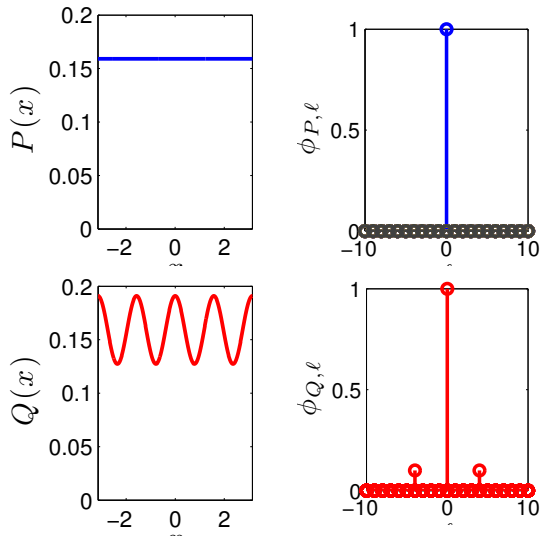
Characteristic kernels on $[-\pi, \pi]$

Example: P differs from Q at one frequency:



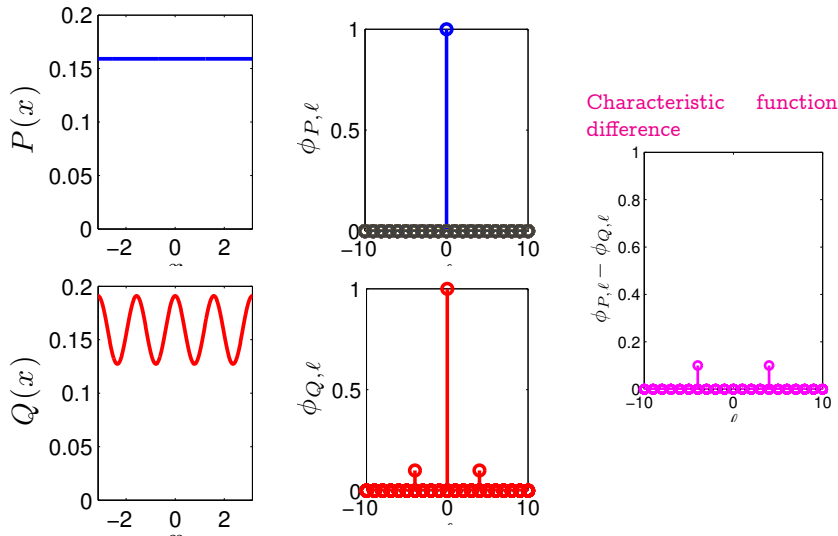
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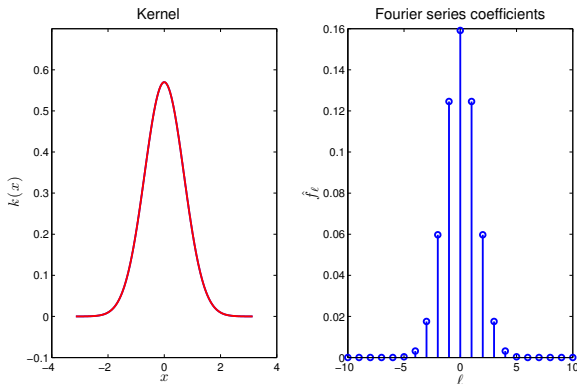
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Characteristic kernels on $[-\pi, \pi]$

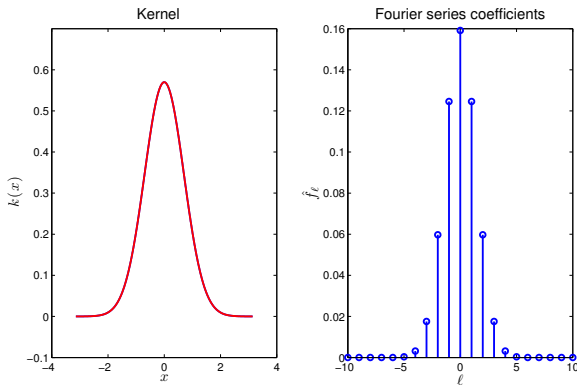
Is the Gaussian spectrum kernel characteristic?



$$MMD^2(P, Q; F) = \sum_{\ell=-\infty}^{\infty} |\varphi_{P,\ell} - \varphi_{Q,\ell}|^2 \hat{k}_\ell$$

Characteristic kernels on $[-\pi, \pi]$

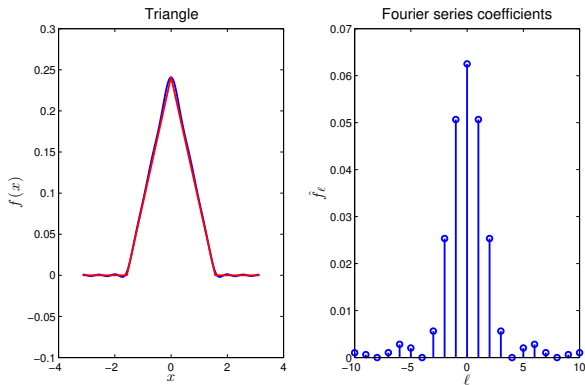
Is the Gaussian spectrum kernel characteristic? **YES**



$$MMD^2(P, Q; F) = \sum_{l=-\infty}^{\infty} |\varphi_{P,l} - \varphi_{Q,l}|^2 \hat{k}_\ell$$

Characteristic kernels on $[-\pi, \pi]$

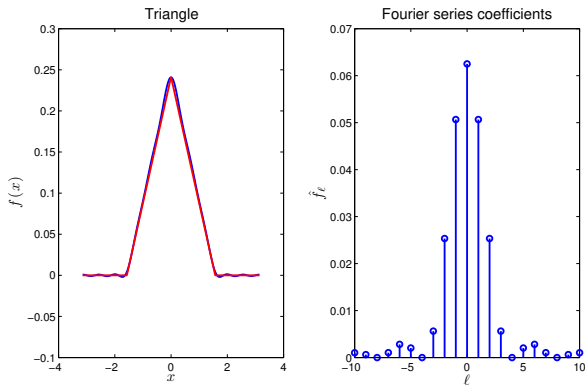
Is the **triangle kernel** characteristic?



$$MMD^2(P, Q; F) = \sum_{\ell=-\infty}^{\infty} |\varphi_{P,\ell} - \varphi_{Q,\ell}|^2 \hat{k}_\ell$$

Characteristic kernels on $[-\pi, \pi]$

Is the triangle kernel characteristic? **NO**



$$MMD^2(P, Q; F) = \sum_{\ell=-\infty}^{\infty} |\varphi_{P,\ell} - \varphi_{Q,\ell}|^2 \hat{k}_\ell$$

Characteristic kernels on \mathbb{R}^d

Can we prove **characteristic on \mathbb{R}^d** ?

Characteristic function of P via **Fourier transform**

$$\varphi_P(\omega) = \int_{\mathbb{R}^d} e^{ix^\top \omega} dP(x)$$

For translation invariant kernels: $k(x, y) = k(x - y)$, **Bochner's theorem**:

$$k(x - y) = \int_{\mathbb{R}^d} e^{-i(x-y)^\top \omega} d\Lambda(\omega)$$

$\Lambda(\omega)$ finite non-negative Borel measure.

Characteristic kernels on \mathbb{R}^d

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Characteristic kernels on \mathbb{R}^d

Fourier representation of MMD on \mathbb{R}^d :

$$\text{MMD}^2(P, Q; F) = \int |\varphi_P(\omega) - \varphi_Q(\omega)|^2 d\Lambda(\omega)$$

Proof: an exercise! But recall the Fourier series case for $[-\pi, \pi]$:

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Characteristic kernels on \mathbb{R}^d

Fourier representation of MMD on \mathbb{R}^d :

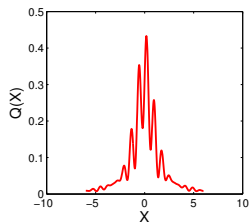
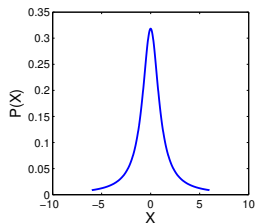
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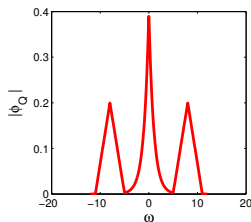
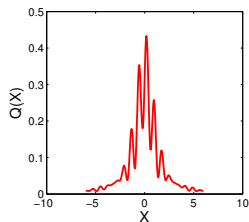
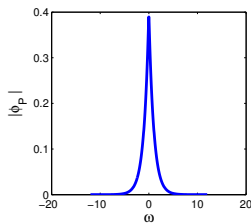
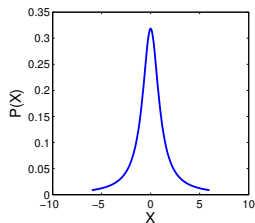
Characteristic kernels on \mathbb{R}^d

Example: P differs from Q at roughly one frequency:



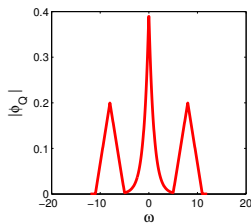
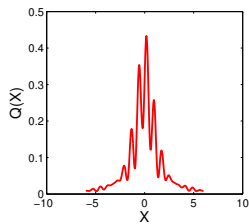
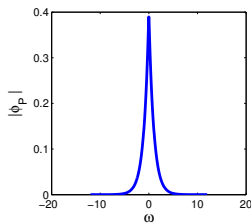
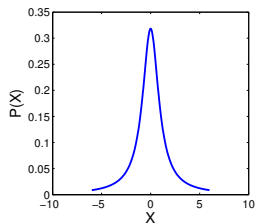
Characteristic kernels on \mathbb{R}^d

Example: P differs from Q at **roughly** one frequency:

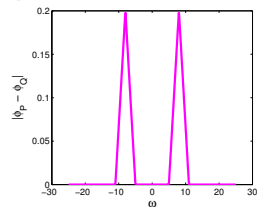


Characteristic kernels on \mathbb{R}^d

Example: P differs from Q at **roughly** one frequency:



Characteristic function difference

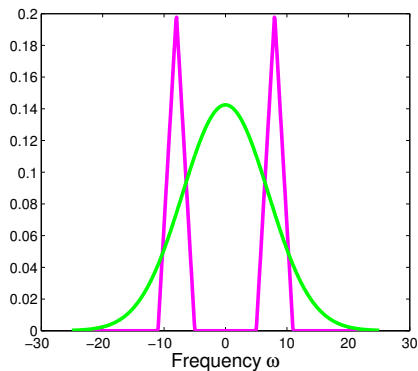


Characteristic kernels on \mathbb{R}^d

Example: P differs from Q at (roughly) one frequency:

Exponentiated quadratic kernel spectrum $\Lambda(\omega)$

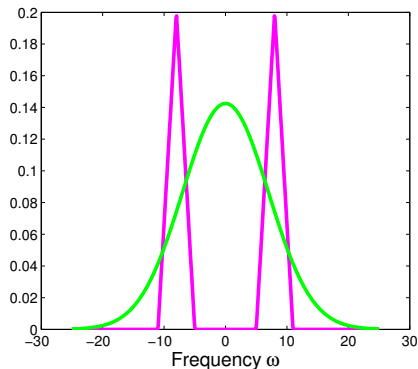
Difference $|\varphi_P - \varphi_Q|$



Characteristic kernels on \mathbb{R}^d

Example: P differs from Q at (roughly) one frequency:

Characteristic

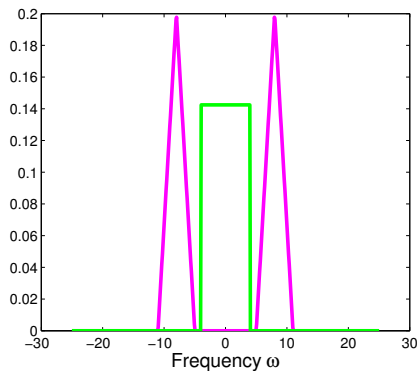


Characteristic kernels on \mathbb{R}^d

Example: P differs from Q at (roughly) one frequency:

Sinc kernel spectrum $\Lambda(\omega)$

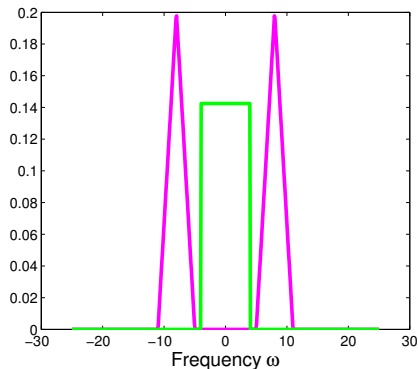
Difference $|\varphi_P - \varphi_Q|$



Characteristic kernels on \mathbb{R}^d

Example: P differs from Q at (roughly) one frequency:

Not characteristic

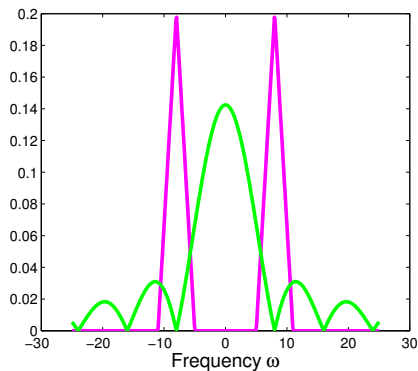


Characteristic kernels on \mathbb{R}^d

Example: P differs from Q at (roughly) one frequency:

Triangle (B-spline) kernel spectrum $\Lambda(\omega)$

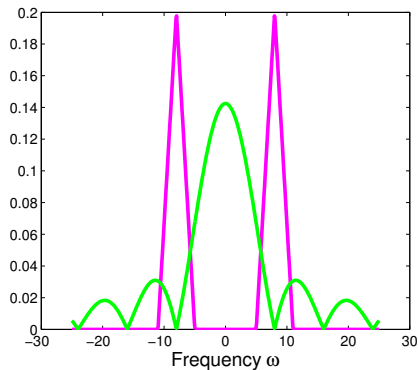
Difference $|\phi_P - \phi_Q|$



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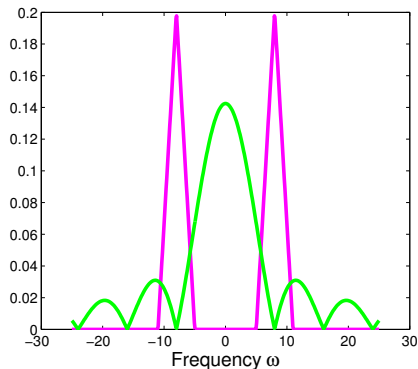
???



Characteristic kernels on \mathbb{R}^d

Example: P differs from Q at (roughly) one frequency:

Characteristic



Summary: characteristic kernels on \mathbb{R}^d

Characteristic kernel: $MMD = 0$ iff $P = Q$ Fukumizu et al. [NIPS07b], Sriperumbudur et al. [COLT08]

Main theorem: A translation invariant k is **characteristic** for prob. measures on \mathbb{R}^d if and only if

$$\text{supp}(\Lambda) = \mathbb{R}^d$$

(i.e. support zero on at most a countable set) Sriperumbudur et al. [COLT08, JMLR10]

Corollary: any continuous, compactly supported k characteristic (since Fourier spectrum $\Lambda(\omega)$ cannot be zero on an interval).

1-D proof sketch from [Mallat, 99, Theorem 2.6], proof on \mathbb{R}^d via distribution theory in Sriperumbudur et al. [JMLR10, Corollary 10 p. 1535]

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