## 1 E-M for latent variable models

The lecture notes and slides over-complicate E-M for mixture models. The derivation is quite simple.

Let  $f(x;\theta)$  be the full distribution of a latent variable model. The density for mixture models can be decomposed as follows (assuming independence):

$$f(x;\theta) = f(x|z;\theta_x)f(z;\theta_z)$$

Where Z follows a discrete distribution and  $\theta_x$  and  $\theta_z$  are parameters specifically for X and Z respectively. Since Z is discrete, this function can be written as

$$f(x; \theta) = \sum_{j} p_{j} f(x|Z = p_{j}; \theta_{x,j}), \sum_{j} p_{j} = 1$$

For notational simplicity, I will drop the notation  $\theta_{x,j}$  and simply revert to  $\theta_j$ .

Intuitively, we should maximize  $f(x;\theta)$  over the observed data to estimate  $\theta$ . We proceed down the path, though we should feel some discomfort since we don't observe z and so the problem may be tricky.

## 2 Maximize the likelihood

Given data  $x_i, i \in [1, 2, ..., n]$ , we can try to maximize the log-likelihood:

$$\max_{\theta, p} \sum_{i} \log \left( \sum_{j} p_{j} f(x_{i}|Z = p_{j}; \theta_{j}) \right)$$

We typically do this by taking derivatives with respect to the parameters and setting the derivatives equal to zero. However, we have the additional constraint that the probabilities for Z must sum to one, which requires us to create the Lagrangian:

$$\ell(\theta, p, q) = \sum_{i} \log \left( \sum_{j} p_{j} f(x_{i}|Z = p_{j}; \theta_{j}) \right) + q \left( \sum_{j} p_{j} - 1 \right)$$

Now we can start taking derivatives with respect to the Lagrangian:

$$\begin{split} \frac{\partial \ell}{\partial \theta_j} &= \sum_i \left( p_j \frac{\partial f(x_i|Z=p_j;\theta_j)}{\partial \theta_j} \right) \frac{1}{\sum_j p_j f(x_i|Z=p_j;\theta_j)} \\ &= \sum_i \frac{\partial \log \left( f(x_i|Z=p_j;\theta_j) \right)}{\partial \theta_j} \frac{p_j f(x_i|Z=p_j;\theta_j)}{\sum_j p_j f(x_i|Z=p_j;\theta_j)} \end{split}$$

$$\begin{split} &= \sum_{i} h_{i,j} \frac{\partial \log \left( f(x_i | Z = p_j; \theta_j) \right)}{\partial \theta_j} \\ \text{Where } h_{i,j} &= \frac{p_j f(x_i | Z = p_j; \theta_j)}{\sum_{j} p_j f(x_i | Z = p_j; \theta_j)} \\ &\qquad \qquad \frac{\partial \ell}{\partial q} = \sum_{j} p_j - 1 \\ &\qquad \qquad \frac{\partial \ell}{\partial p_j} = \sum_{i} \frac{f(x_i | Z = p_j; \theta_j)}{\sum_{j} p_j f(x_i | Z = p_j; \theta_j)} + q \\ &\qquad \qquad = \sum_{i} \frac{h_{i,j}}{p_j} + q \end{split}$$

Setting each of these to zero (and treating  $h_{i,j}$  like a constant!) yields

$$q = -\sum_{j} \sum_{i} h_{i,j}$$
$$p_{j} = \frac{\sum_{i} h_{i,j}}{\sum_{i} \sum_{i} h_{i,j}}$$

The E-M algorithm now becomes obvious: simply maximize the likelihood (for a given set of initial p), then set  $h_{i,j}$  with the updated  $p, \theta$ , and repeat until convergence.

## 3 Example

Assume an mixture of two exponential distributions. The maximization step (fixing  $h_{i,j}$ ), is

$$\frac{\partial \ell}{\partial \theta_j} = \sum_{i} h_{i,j} \left( \frac{1}{\theta_j} - x_i \right)$$

Setting this equal to zero yields

$$\theta_j = \frac{n}{\sum_i x_i}$$

Updating the  $p_j$  are trivial.

Then the next iteration of  $h_{i,j}$  can be found using the (new)  $\theta_j$ ,  $p_j$ .