CSSS 510: Lab 3

Logistic Regression

2017-10-13

0. Agenda

1. Deriving a likelihood function for the logistic regression model
2. Fitting a logit model using optim() and glm()
3. Simulating predicted values and confidence intervals
4. Simulating first differences

1. Deriving a likelihood function for the logistic regression model

Recall from lecture the logit model:

$$y_i \sim \text{Bern}(y_i|\pi_i)$$

$$\pi_i = \text{logit}^{-1}(\boldsymbol{x}_i\beta)$$

$$\pi_i = \frac{\exp(\boldsymbol{x}_i\beta)}{1+\exp(\boldsymbol{x}_i\beta)} = \frac{1}{1+\exp(-\boldsymbol{x}_i\beta)}$$

1. Deriving a likelihood function for the logistic regression model

In the simple case, this stems from the latent variable model:

$$y^* = \beta_0 + \beta_1 x + \epsilon$$

where the relationship between latent variable y^* and the explanatory variable x is modeled using simple linear regression, and the binary outcome y is a function of the sign of y^*:

$$y = \begin{cases} 1, & \text{if } y^* > 0 \\ 0, & \text{if } y^* \leq 0 \end{cases}$$

1. Deriving a likelihood function for the logistic regression model

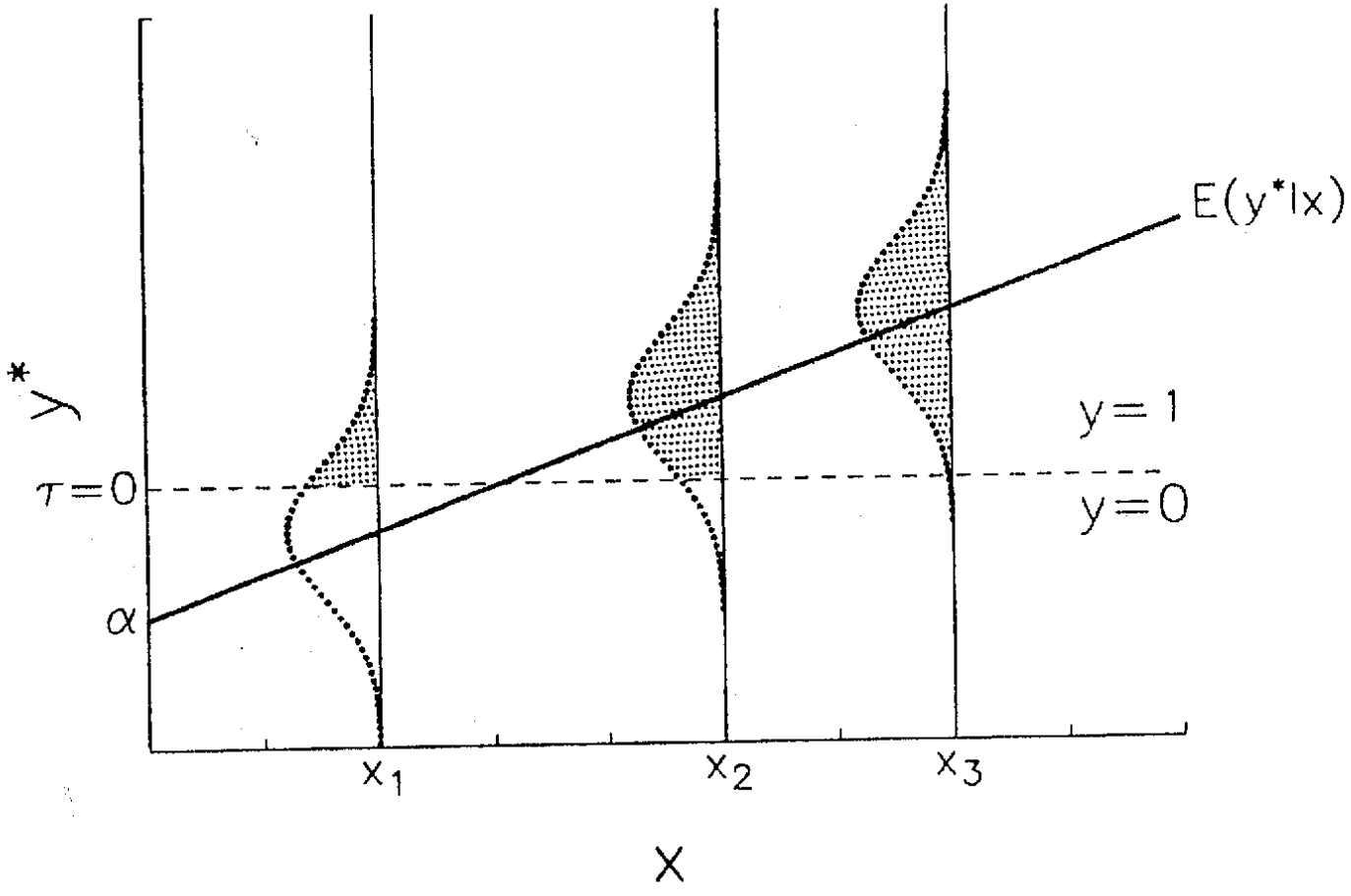


Figure 3.2. The Distribution of y^* Given x in the Binary Response Model

1. Deriving a likelihood function for the logistic regression model

The logistic regression model is obtained if we assume the errors of this latent variable model follow a standard logistic distribution.

Recall that the pdf and cdf of the standard logisitic distribution are as follows:

$$f(t) = \frac{\exp(t)}{(1+\exp(t))^2}$$

$$F(t) = \frac{\exp(t)}{1+\exp(t)}$$

1. Deriving a likelihood function for the logistic regression model

We therefore have the following:

$$
\begin{aligned}
\Pr(y = 1|x) &= \Pr(y^* > 0|x) \\
&= \Pr(\beta_0 + \beta_1 x + \epsilon > 0|x) \\
&= \Pr(\epsilon > -(\beta_0 + \beta_1 x)) \\
&= \Pr(\epsilon < \beta_0 + \beta_1 x) \\
&= F(\beta_0^L + \beta_1^L x)
\end{aligned}
$$

1. Deriving a likelihood function for the logistic regression model

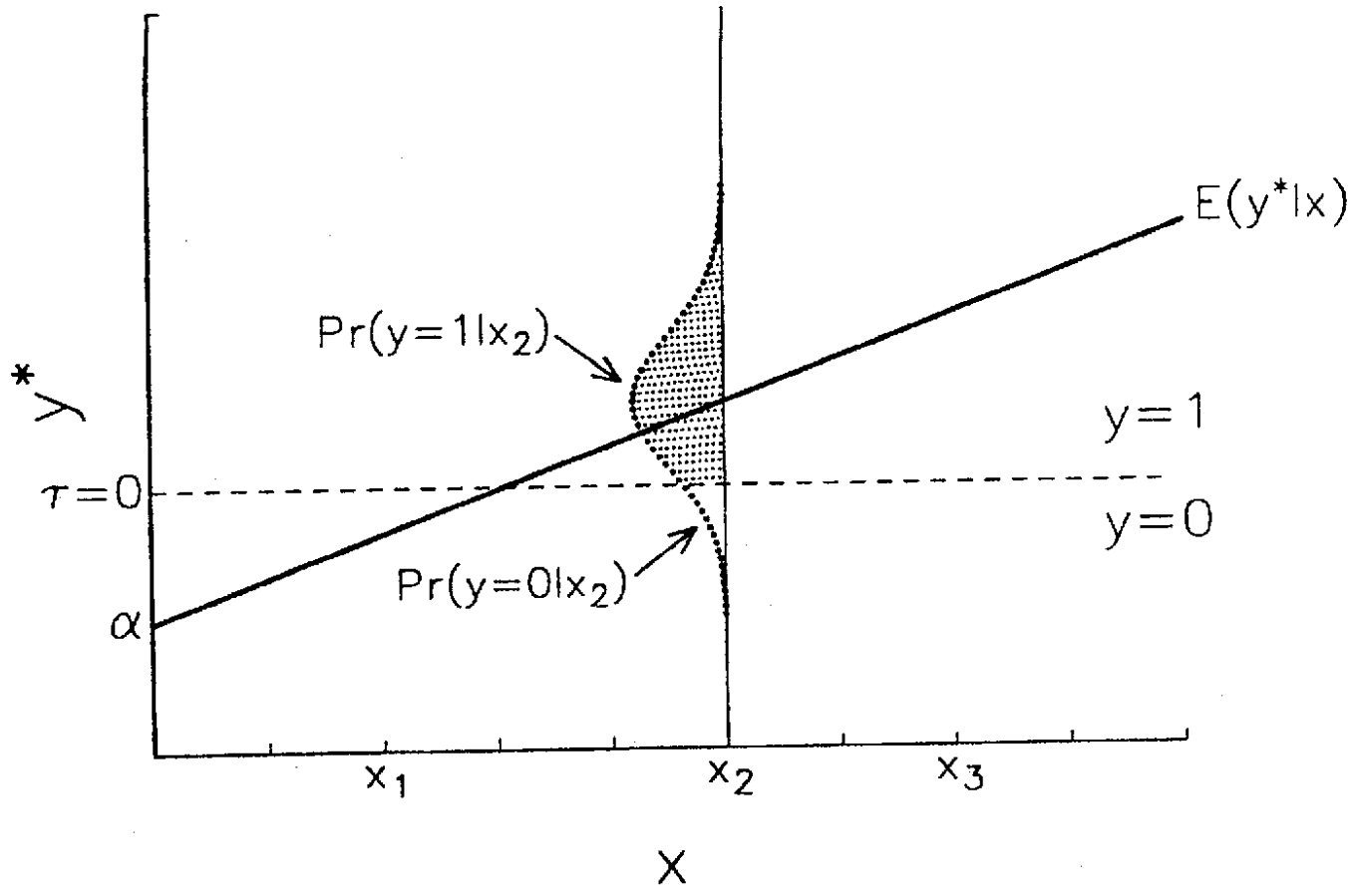


Figure 3.4. Probability of Observed Values in the Binary Response Model

1. Deriving a likelihood function for the logistic regression model

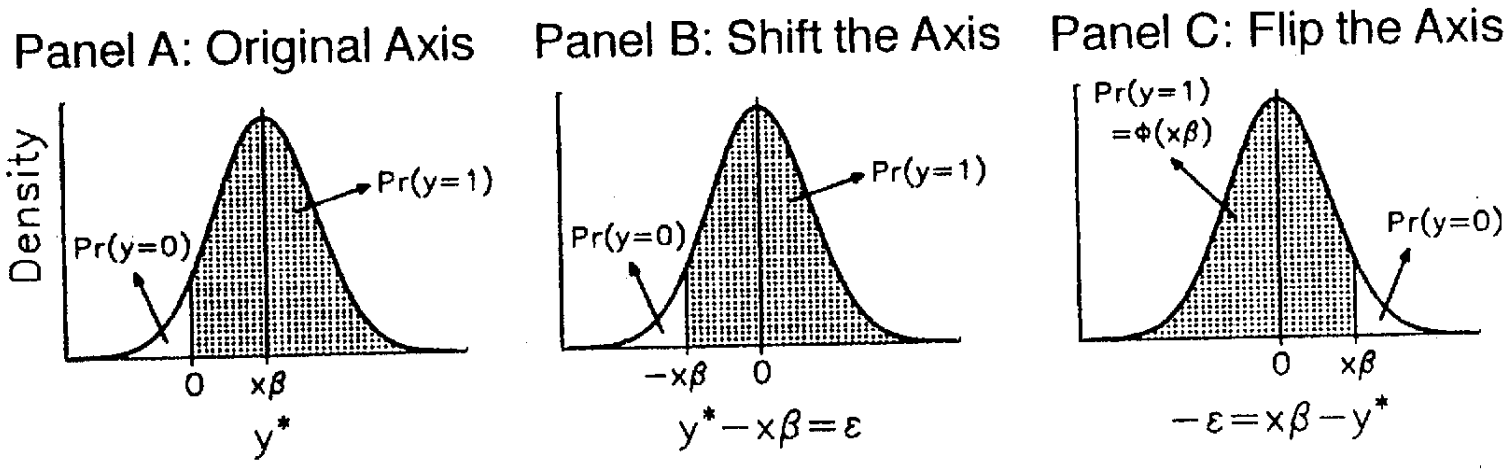


Figure 3.5: Computing $\Pr(y = 1 \mid x)$ in the Binary Response Model

1. Deriving a likelihood function for the logistic regression model

Since we assume the errors follow a standard logistic distribution, we have

$$\Pr(y = 1|x) = F(\beta_0^L + \beta_1^L x)$$

$$= \frac{\exp(\beta_0^L + \beta_1^L x)}{1 + \exp(\beta_0^L + \beta_1^L x)}$$

$\mathrm{E}(\epsilon)$=0 and $\mathrm{Var}(\epsilon) = \frac{\pi^2}{3}$.

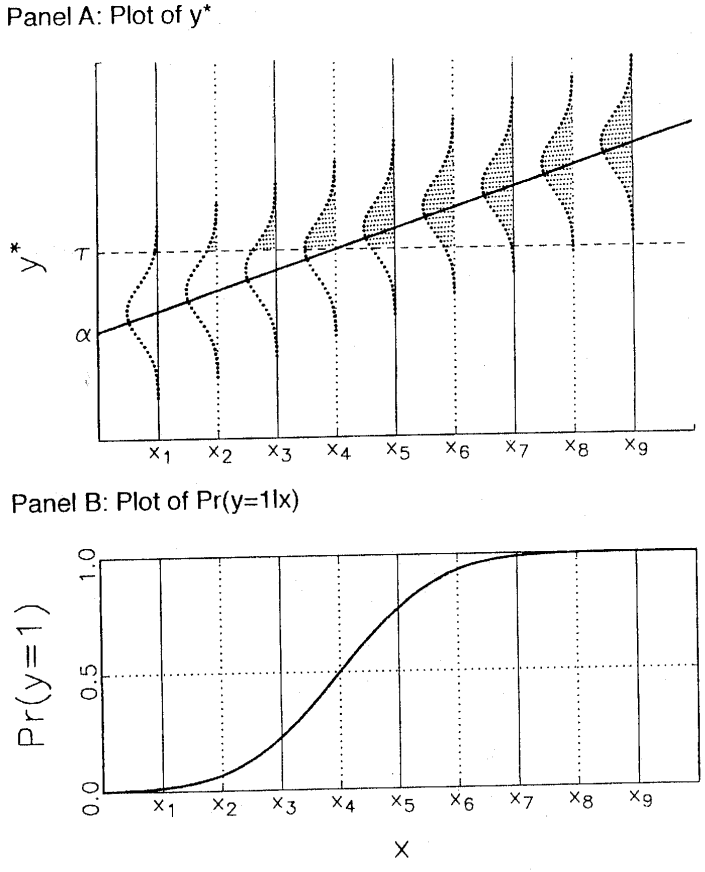


Figure 3.6. Plot of y^* and $\Pr(y = 1 \mid \mathbf{x})$ in the Binary Response Model

1. Deriving a likelihood function for the logistic regression model

The logit function is the inverse of the logistic function:

$$\text{logit}(p) = \log \frac{p}{1-p}$$

or

$$\text{logit}^{-1}(p) = \frac{\exp(x)}{1+\exp(x)}$$

We therefore have the following

$$\Pr(y = 1|x) = \text{logit}^{-1}(\beta_1^L + \beta_1^L x)$$

or

$$\text{logit}(\Pr(y = 1|x)) = \beta_1^L + \beta_1^L x$$

or

$$\log \frac{\Pr(y = 1|x)}{\Pr(y = 0|x)} = \beta_0^L + \beta_1^L x.$$

1. Deriving a likelihood function for the logistic regression model

Recall from lecture that the Bernoulli distribution has the following pdf:

$$\Pr(y_i = 1|\pi_i) = \pi_i^{y_i}(1-\pi_i)^{1-y_i}$$

And the likelihood function can be derived from the joint probability:

$$\mathcal{L}(\boldsymbol{\pi}|\boldsymbol{y}) \propto \prod_{i=1}^{n} \pi_i^{y_i}(1-\pi_i)^{1-y_i}$$

$$\mathcal{L}(\beta|\boldsymbol{y}) \propto \prod_{i=1}^{n} \left(\frac{1}{1+\exp(-\boldsymbol{x}_i\beta)}\right)^{y_i} \left(1-\frac{1}{1+\exp(-\boldsymbol{x}_i\beta)}\right)^{1-y_i}$$

$$\mathcal{L}(\beta|\boldsymbol{y}) \propto \prod_{i=1}^{n} (1+\exp(-\boldsymbol{x}_i\beta))^{-y_i}(1+\exp(-\boldsymbol{x}_i\beta))^{-(1-y_i)}$$

$$\log\mathcal{L}(\beta|\boldsymbol{y}) \propto \sum_{i=1}^{n} -y_i\log(1+\exp(-\boldsymbol{x}_i\beta)) - (1-y_i)\log(1+\exp(\boldsymbol{x}_i\beta))$$

2. Fitting a logit model using optim() and glm()

```
rm(list = ls()) # clear up the memory

#install and load the packages needed
#from CRAN: install.packages("MASS", dependencies = TRUE)
library(MASS)
library(RColorBrewer)

# download simcf and tile packages
# from- http://faculty.washington.edu/cadolph/software
# don't unzip the archive (tar) file
library(simcf)
library(tile)
```

```
## Loading required package: grid
```

```
# Load data
file <- "nes00a.csv"
data <- read.csv(file, header=TRUE)
# attach(data)
```

2. Fitting a logit model using optim() and glm()

```
# Estimate logit model using optim()
# Construct variables and model objects
y <- data$vote00
x <- cbind(data$age,data$hsdeg,data$coldeg)

# Likelihood function for logit
llk.logit <- function(param,y,x) {
  os <- rep(1,length(x[,1])) # constant
  x <- cbind(os,x) # constant+covariates
  b <- param[ 1 : ncol(x) ]
  # number of parameters to be estimated equals number of columns in x
  # (i.e, one for constant and one for each covariates : total 4)
  xb <- x%*%b
  sum( y*log(1+exp(-xb)) + (1-y)*log(1+exp(xb)))
  # log-likelihood function for logit model
  # (based on our choice of standard logistic cdf as the systematic component)
                # optim is a minimizer, so use -lnL here
}
```

2. Fitting a logit model using optim() and glm()

```
# Fit logit model using optim
ls.result <- lm(y~x)  # use ls estimates as starting values (for convenience)
stval <- ls.result$coefficients  # initial guesses
logit.result.opt <- optim(stval,llk.logit,method="BFGS",hessian=T,y=y,x=x)
# call minimizer procedure or max by adding control=list(fnscale=-1)
pe.opt <- logit.result.opt$par   # point estimates
vc.opt <- solve(logit.result.opt$hessian)  # var-cov matrix
se.opt <- sqrt(diag(vc.opt))    # standard errors
ll.opt <- -logit.result.opt$value  # likelihood at maximum

logit.optim<-data.frame(cbind(round(pe.opt,3), round(se.opt,3)))
rownames(logit.optim)<-c("intercept", "age", "highschool" , "college")
colnames(logit.optim)<-c("pe", "std.err")
logit.optim
```

```
##                pe std.err
## intercept  -2.149   0.257
## age         0.031   0.003
## highschool  1.213   0.179
## college     1.102   0.130
```

2. Fitting a logit model using optim() and glm()

```
#p-value based on t-statistics
2*pt(abs(logit.optim$pe/logit.optim$std.err),
     df=length(y)-length(pe.opt) , lower.tail = FALSE)
```

```
## [1] 1.229175e-16 2.393414e-24 1.668155e-11 4.780627e-17
```

```
# Estimate logit model using glm()

# Run logit & extract results using glm.
# GLM solves the likelihood equations with a common numeric algorithm
# called iteratively re-weighted least squares (IRWLS).

logit.result <- glm(vote00~age +hsdeg+coldeg, family=binomial, data=data)
# family "binomial" calls logit transformation
# (log(pi/1-pi)) as a "link function"
# corresponding to logistic distribution.
# Link function transforms pi so that it follows a linear model.
# (So although pi itself is dependent on covariates in a non-linear way,
# logit transformed pi is dependent on covariates in a linear way.)
```

2. Fitting a logit model using optim() and glm()

```
summary(logit.result)
```

```
##
## Call:
## glm(formula = vote00 ~ age + hsdeg + coldeg, family = binomial,
##     data = data)
##
## Deviance Residuals:
##     Min       1Q   Median       3Q      Max
## -2.4487  -1.1347   0.6360   0.8973   1.8854
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -2.147997   0.256608  -8.371  < 2e-16 ***
## age          0.030885   0.003386   9.121  < 2e-16 ***
## hsdeg        1.212882   0.179447   6.759 1.39e-11 ***
## coldeg       1.102465   0.130426   8.453  < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##     Null deviance: 2293.5  on 1782  degrees of freedom
## Residual deviance: 2076.0  on 1779  degrees of freedom
## AIC: 2084
##
## Number of Fisher Scoring iterations: 4
```

2. Fitting a logit model using optim() and glm()

```
# now a new model adding age^2
model <- vote00 ~ age + I(age^2) + hsdeg + coldeg
mdata <- extractdata(model, data, na.rm=TRUE) # needs library(simcf)

logit.result <- glm(model, family=binomial, data=mdata)

pe <- logit.result$coefficients  # point estimates
vc <- vcov(logit.result)         # var-cov matrix

pe
```

```
##   (Intercept)           age      I(age^2)         hsdeg        coldeg
## -3.0193890720  0.0747251922 -0.0004427014  1.1243907749  1.0795702357
```

```
sqrt(diag(vc))
```

```
##  (Intercept)           age      I(age^2)         hsdeg        coldeg
## 0.4181899367 0.0168440337 0.0001655466 0.1800068893 0.1312113265
```

3. Simulating predicted values and confidence intervals

```
# Simulate parameter distributions
sims <- 10000
simbetas <- mvrnorm(sims, pe, vc) #needs library(MASS)
# draw 10000 sets of simulated
# parameter (beta) estimates from a multivariate normal distribution with
# mean pe and variance-covariance vc

# Now let's plan counterfactuals: We will have three
# sets of counterfactuals based on education lavel
# (less than hs edu, hs edu, college or higher edu), and for each set
# we will make age varies between 18 years old and 97 years old.

# Set up counterfactuals:  all ages, each of three educations
xhyp <- seq(18,97,1)  # create age vector
nscen <- length(xhyp)  # we will have total 80 different age scenarios
#for each education level

nohsScen <- hsScen <- collScen <- cfMake(model, mdata, nscen)
#this is just to initialize 80 scenarios for each education level.
#As default, all covariate values are set at the mean.
```

3. Simulating predicted values and confidence intervals

```
# Create three sets of education counterfactuals

for (i in 1:nscen) {
  # No High school scenarios (loop over each age, total 80 scenarios)
  nohsScen <- cfChange(nohsScen, "age", x = xhyp[i], scen = i)
  nohsScen <- cfChange(nohsScen, "hsdeg", x = 0, scen = i)#no hs degree
  nohsScen <- cfChange(nohsScen, "coldeg", x = 0, scen = i)#no college degree

  # HS grad scenarios (loop over each age, total 80 scenarios)
  hsScen <- cfChange(hsScen, "age", x = xhyp[i], scen = i)
  hsScen <- cfChange(hsScen, "hsdeg", x = 1, scen = i)#has hs degree
  hsScen <- cfChange(hsScen, "coldeg", x = 0, scen = i)#no college degree

  # College grad scenarios (loop over each age, total 80 scenarios)
  collScen <- cfChange(collScen, "age", x = xhyp[i], scen = i)
  collScen <- cfChange(collScen, "hsdeg", x = 1, scen = i)#has hs degree
  collScen <- cfChange(collScen, "coldeg", x = 1, scen = i)#has college degree
}
```

3. Simulating predicted values and confidence intervals

```
# Now given the counterfactual covariates (nohsScen/hsScen/collScen)
# and simulated parameters (simbetas), we can calculate expected value
# of the response. In this case, expected probability of voting!

head(nohsScen$x) #we will fit the counterfactual data
```

```
##      vote00 age hsdeg coldeg
## 1 0.6567583  18     0      0
## 2 0.6567583  19     0      0
## 3 0.6567583  20     0      0
## 4 0.6567583  21     0      0
## 5 0.6567583  22     0      0
## 6 0.6567583  23     0      0
```

```
nohsScen$model # in the model specification
```

```
## vote00 ~ age + I(age^2) + hsdeg + coldeg
```

3. Simulating predicted values and confidence intervals

```
nohsSims <- logitsimev(nohsScen, simbetas, ci=0.95) # using simulated
# betas to get expected values.
# Built-in function "logitsimev" calculates the expected value
# for every individual scenario you created.
# reports lower and upper confidence intervals
# as well as expected probabilities

# same thing for two other sets of sceneraios
hsSims <- logitsimev(hsScen, simbetas, ci=0.95)
collSims <- logitsimev(collScen, simbetas, ci=0.95)
```

3. Simulating predicted values and confidence intervals

```
# Get 3 nice colors for traces
col <- brewer.pal(3,"Dark2")

# Set up lineplot traces of expected probabilities

#Traces are elements of tile : lines, labels, legends...

# no hs
nohsTrace <- lineplot(x=xhyp, # age on x-axis
  y=nohsSims$pe, #expected probability on y-axis
  lower=nohsSims$lower, # lower confidence interval
  upper=nohsSims$upper, #upper confidence interval
  col=col[1], #color choice
  extrapolate=list(data=mdata[,2:ncol(mdata)],
  #actual covariates (i.e., values in your data)
  cfact=nohsScen$x[,2:ncol(hsScen$x)],#counterfactual covariates
  omit.extrapolated=FALSE), #don't show extrapolated values
  plot=1)
```

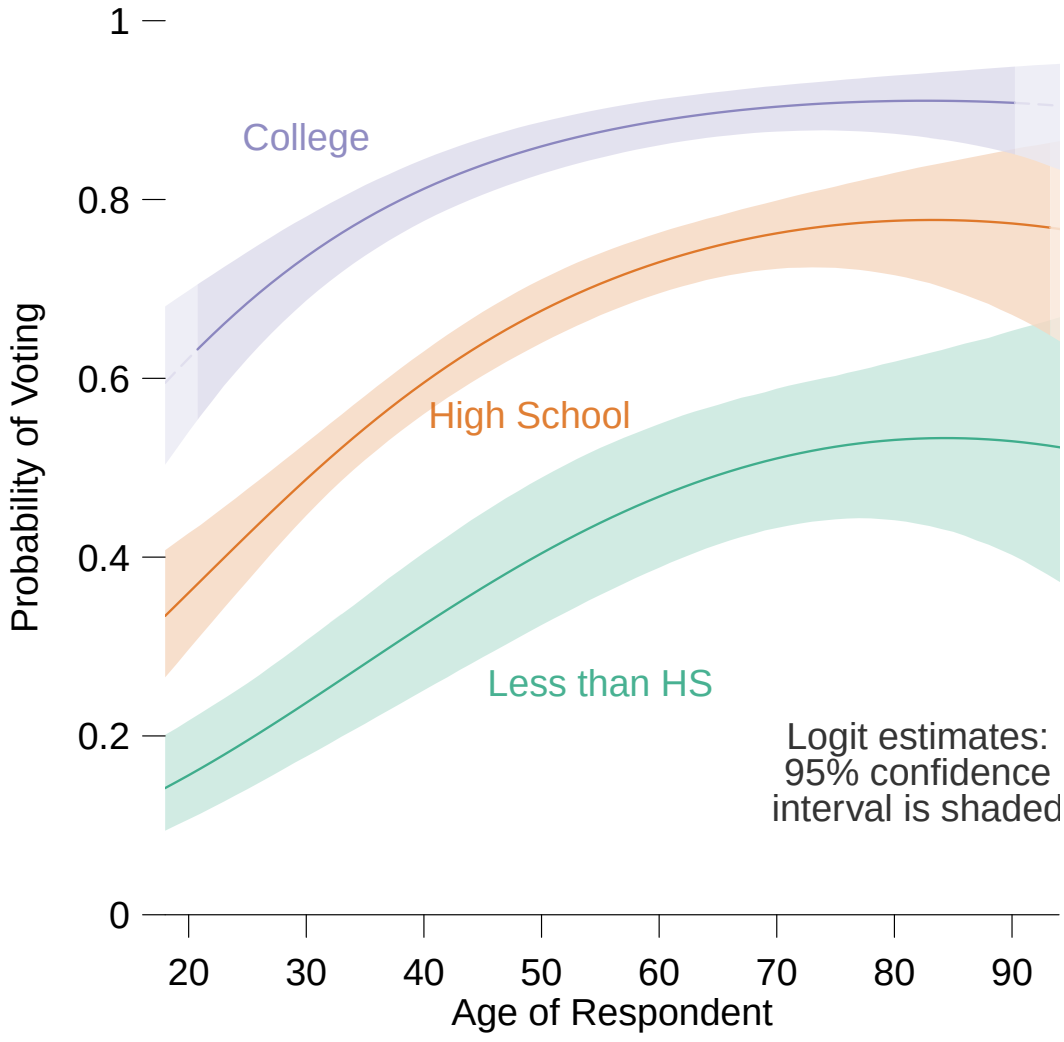
3. Simulating predicted values and confidence intervals

```
# hs but no college
hsTrace <- lineplot(x=xhyp,
  y=hsSims$pe,
  lower=hsSims$lower,
  upper=hsSims$upper,
  col=col[2],
  extrapolate=list(data=mdata[,2:ncol(mdata)],
  cfact=hsScen$x[,2:ncol(hsScen$x)],
  omit.extrapolated=FALSE),
  plot=1)

#college
collTrace <- lineplot(x=xhyp,
  y=collSims$pe,
  lower=collSims$lower,
  upper=collSims$upper,
  col=col[3],
  extrapolate=list(data=mdata[,2:ncol(mdata)],
  cfact=collScen$x[,2:ncol(hsScen$x)],
  omit.extrapolated=FALSE),
  plot=1)
```

3. Simulating predicted values and confidence intervals

```
# Set up traces with labels
labelTrace <- textTile(labels=c("Less than HS", "High School", "College"),
    x=c( 55,    49,     30),
    y=c( 0.26,  0.56,   0.87),
    col=col,
    plot=1)


# For legend
legendTrace <-
  textTile(labels=c("Logit estimates:", "95% confidence", "interval is shaded"),
  x=c(82, 82, 82),
  y=c(0.2, 0.16, 0.12),
  plot=1)


#options(device="quartz")
# Plot traces using tile
voting<-tile(nohsTrace,
     hsTrace,
     collTrace,
     labelTrace,
     legendTrace,
     limits=c(18,94,0,1),
     xaxis=list(at=c(20,30,40,50,60,70,80,90)),
     yaxis=list(label.loc=-0.5, major=FALSE),
     xaxistitle=list(labels="Age of Respondent"),
     yaxistitle=list(labels="Probability of Voting"),
     width=list(null=5,yaxistitle=4,yaxis.labelspace=-0.5)
     ,output=list(file="educationEV",width=5.5)
)
```

3. Simulating predicted values and confidence intervals



4. Simulating first differences

```
#################################################################
#
# Now consider a new specification adding the variable
# "ever married", or marriedo
# We will estimate this new model with glm(), then
# simulate new scenarios for marrieds and non-marrieds

# Estimate logit model using glm()
# Set up a new model formula and model specific data frame

model2 <- vote00 ~ age + I(age^2) + hsdeg + coldeg + marriedo
mdata2 <- extractdata(model2, data, na.rm=TRUE)

# Run logit & extract results
logit.m2 <- glm(model2, family=binomial, data=mdata2)
pe.m2 <- logit.m2$coefficients  # point estimates
vc.m2 <- vcov(logit.m2)         # var-cov matrix

# Simulate parameter distributions
sims <- 10000
simbetas.m2 <- mvrnorm(sims, pe.m2, vc.m2)
```

4. Simulating first differences

```
# Set up counterfactuals:  all ages
xhyp <- seq(18,97,1)
nscen <- length(xhyp)
marriedScen <- notmarrScen <- cfMake(model2, mdata2, nscen)
for (i in 1:nscen) {


  #  - we will use the marriedScen counterfactuals in FDs and RRs as well as EVs
  # Note below the careful use of before scenarios (xpre) and after scenarios (x)
  # :i.e., use of the same age range (18-97) for both x and xpre, only marriedo values differ.

  # Married (loop over each age)
  marriedScen <- cfChange(marriedScen, "age", x = xhyp[i], xpre= xhyp[i], scen = i)
  marriedScen <- cfChange(marriedScen, "marriedo", x = 1, xpre= 0, scen = i)

  # Not Married (loop over each age)
  notmarrScen <- cfChange(notmarrScen, "age", x = xhyp[i], scen = i)
  notmarrScen <- cfChange(notmarrScen, "marriedo", x = 0, scen = i)
}

# Simulate expected probabilities for all age scenarios for married and not married respectively
marriedSims <- logitsimev(marriedScen, simbetas.m2, ci=0.95)
notmarrSims <- logitsimev(notmarrScen, simbetas.m2, ci=0.95)

# Simulate first difference of voting wrt marriage: E(y|married)-E(y|notmarried)
marriedFD <- logitsimfd(marriedScen, simbetas.m2, ci=0.95)

# Simulate relative risk of voting wrt marriage: E(y|married)/E(y|notmarried)
marriedRR <- logitsimrr(marriedScen, simbetas.m2, ci=0.95)
```

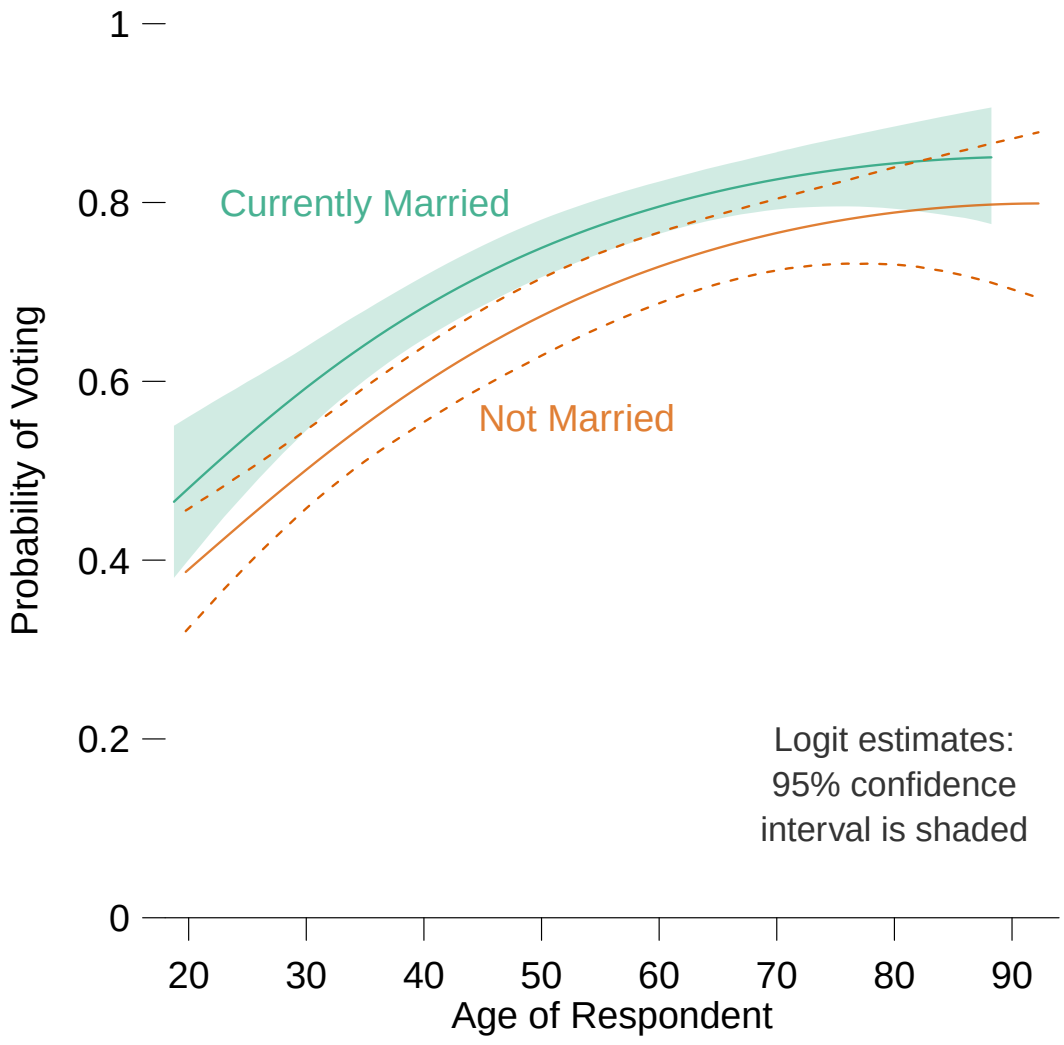
4. Simulating first differences

```
## Make plots using tile

# Get 3 nice colors for traces
col <- brewer.pal(3,"Dark2")

# Set up lineplot traces of expected probabilities
marriedTrace <- lineplot(x=xhyp,
                         y=marriedSims$pe,
                         lower=marriedSims$lower,
                         upper=marriedSims$upper,
                         col=col[1],
                         extrapolate=list(data=mdata2[,2:ncol(mdata2)],
                                          cfact=marriedScen$x[,2:ncol(marriedScen$x)],
                                          omit.extrapolated=TRUE),
                         plot=1)

notmarrTrace <- lineplot(x=xhyp,
                         y=notmarrSims$pe,
                         lower=notmarrSims$lower,
                         upper=notmarrSims$upper,
                         col=col[2],
                         ci = list(mark="dashed"),
                         extrapolate=list(data=mdata2[,2:ncol(mdata2)],
                                          cfact=notmarrScen$x[,2:ncol(notmarrScen$x)],
                                          omit.extrapolated=TRUE),
                         plot=1)
```

4. Simulating first differences

```
# Set up traces with labels and legend
labelTrace <- textTile(labels=c("Currently Married", "Not Married"),
                       x=c( 35,    53),
                       y=c( 0.8,  0.56),
                       col=col,
                       plot=1)



legendTrace <- textTile(labels=c("Logit estimates:", "95% confidence", "interval is shaded"),
                        x=c(80, 80, 80),
                        y=c(0.2, 0.15, 0.10),
                        cex=0.9,
                        plot=1)



# Plot traces using tile
tile(marriedTrace,
     notmarrTrace,
     labelTrace,
     legendTrace,
     limits=c(18,94,0,1),
     xaxis=list(at=c(20,30,40,50,60,70,80,90)),
     yaxis=list(label.loc=-0.5, major=FALSE),
     xaxistitle=list(labels="Age of Respondent"),
     yaxistitle=list(labels="Probability of Voting"),
     width=list(null=5,yaxistitle=4,yaxis.labelspace=-0.5)
     #,output=list(file="marriedEV",width=5.5)
)
```

4. Simulating first differences



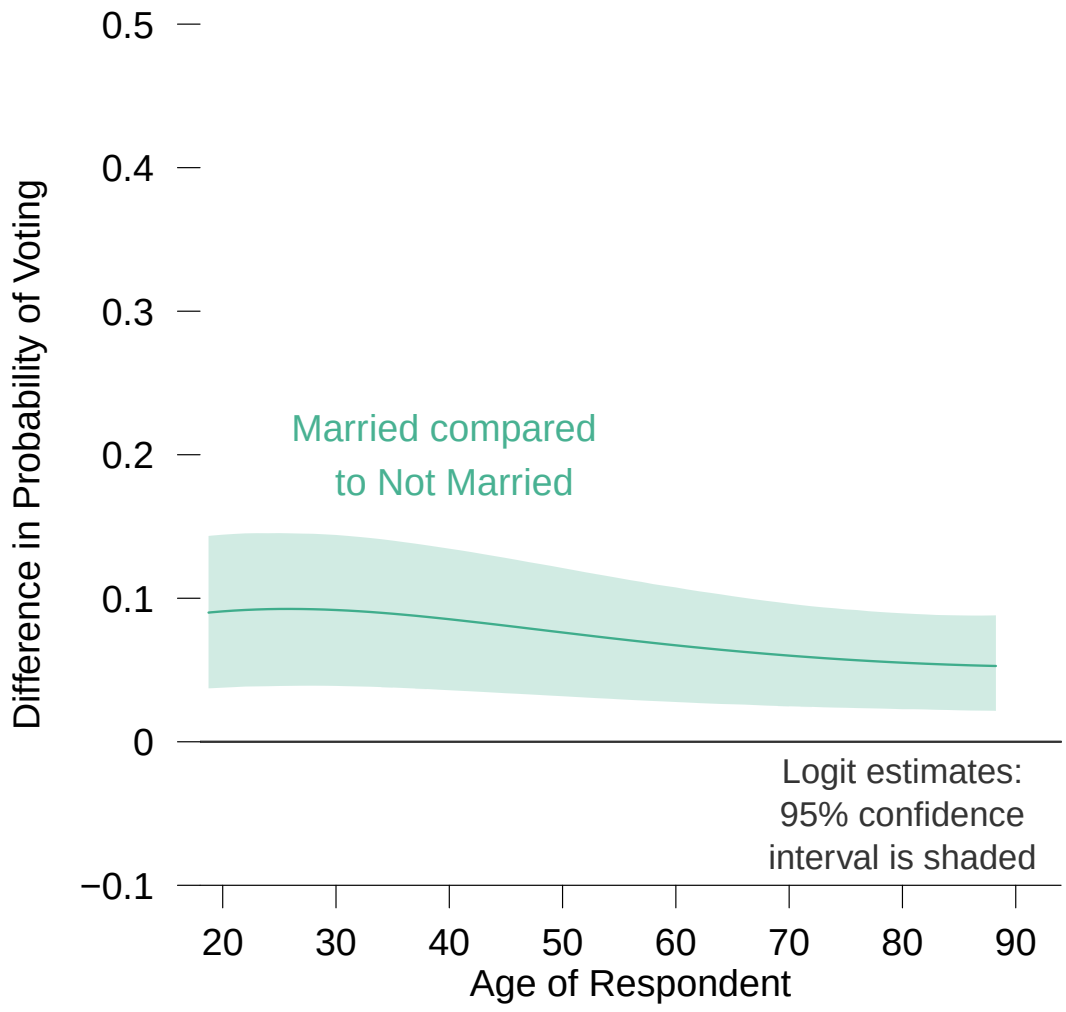
4. Simulating first differences

```
# Plot First Difference
# Set up lineplot trace of first difference

marriedFDTrace <- lineplot(x=xhyp,
                           y=marriedFD$pe,
                           lower=marriedFD$lower,
                           upper=marriedFD$upper,
                           col=col[1],
                           extrapolate=list(data=mdata2[,2:ncol(mdata2)],
                                            cfact=marriedScen$x[,2:ncol(marriedScen$x)],
                                            omit.extrapolated=TRUE),
                           plot=1)


# Set up baseline: for first difference, this is 0
baseline <- linesTile(x=c(18,94),
                      y=c(0,0),
                      plot=1)
```

4. Simulating first differences

```
# Set up traces with labels and legend
labelFDTrace <- textTile(labels=c("Married compared \n to Not Married"),
                         x=c( 40),
                         y=c( 0.20),
                         col=col[1],
                         plot=1)



legendFDTrace <- textTile(labels=c("Logit estimates:", "95% confidence", "interval is shaded"),
                          x=c(80, 80, 80),
                          y=c(-0.02, -0.05, -0.08),
                          cex=0.9,
                          plot=1)



# Plot traces using tile
tile(marriedFDTrace,
     labelFDTrace,
     legendFDTrace,
     baseline,
     limits=c(18,94,-0.1,0.5),
     xaxis=list(at=c(20,30,40,50,60,70,80,90)),
     yaxis=list(label.loc=-0.5, major=FALSE),
     xaxistitle=list(labels="Age of Respondent"),
     yaxistitle=list(labels="Difference in Probability of Voting"),
     width=list(null=5,yaxistitle=4,yaxis.labelspace=-0.5)
     #,output=list(file="marriedFD",width=5.5)
)
```

4. Simulating first differences



4. Simulating first differences

```
# Plot Relative Risk

# Set up lineplot trace of relative risk
marriedRRTrace <- lineplot(x=xhyp,
                           y=marriedRR$pe,
                           lower=marriedRR$lower,
                           upper=marriedRR$upper,
                           col=col[1],
                           extrapolate=list(data=mdata2[,2:ncol(mdata2)],
                                            cfact=marriedScen$x[,2:ncol(marriedScen$x)],
                                            omit.extrapolated=TRUE),
                           plot=1)


# Set up baseline: for relative risk, this is 1

baseline <- linesTile(x=c(18,94),
                      y=c(1,1),
                      plot=1)
```

4. Simulating first differences

```
# Set up traces with labels and legend
labelRRTrace <- textTile(labels=c("Married compared \n to Not Married"),
                         x=c( 55),
                         y=c( 1.25),
                         col=col[1],
                         plot=1)



legendRRTrace <- textTile(labels=c("Logit estimates:", "95% confidence", "interval is shaded"),
                          x=c(80, 80, 80),
                          y=c(0.98, 0.95, 0.92),
                          cex=0.9,
                          plot=1)



# Plot traces using tile
tile(marriedRRTrace,
     labelRRTrace,
     legendRRTrace,
     baseline,
     limits=c(18,94,0.9,1.5),
     xaxis=list(at=c(20,30,40,50,60,70,80,90)),
     yaxis=list(label.loc=-0.5, major=FALSE),
     xaxistitle=list(labels="Age of Respondent"),
     yaxistitle=list(labels="Relative Risk of Voting"),
     width=list(null=5,yaxistitle=4,yaxis.labelspace=-0.5)
     #,output=list(file="marriedRR",width=5.5)
)
```

4. Simulating first differences