CSSS 510: Lab 3

Logistic Regression

0. Agenda

- 1. Deriving a likelihood function for the logistic regression model
- 2. Fitting a logit model using optim() and glm()
- 3. Simulating predicted values and confidence intervals
- 4. Simulating first differences

Recall from lecture the logit model:

$$y_i \sim \mathsf{Bern}(\pi_i)$$
 $\pi_i = \mathsf{logit}^{-1}(m{x}_im{eta})$ $\pi_i = \frac{\mathsf{exp}(m{x}_im{eta})}{1 + \mathsf{exp}(m{x}_im{eta})} = rac{1}{1 + \mathsf{exp}(-m{x}_im{eta})}$

In the simple case, this stems from the latent variable model:

$$y^* = \beta_0 + \beta_1 x + \epsilon$$

where the relationship between latent variable y^* and the explanatory variable x is modeled using simple linear regression, and the binary outcome y is a function of the sign of y^* :

$$y = \begin{cases} 1, & \text{if } y^* > 0 \\ 0, & \text{if } y^* \le 0 \end{cases}$$

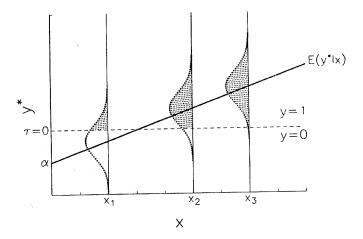


Figure 3.2. The Distribution of y^* Given x in the Binary Response Model

The logistic regression model is obtained if we assume the errors of this latent variable model follow a standard logistic distribution.

Recall that the pdf and cdf of the standard logisitic distribution are as follows:

$$f(t) = \frac{\exp(t)}{(1 + \exp(t))^2}$$

$$F(t) = \frac{\exp(t)}{1 + \exp(t)}$$

We therefore have the following:

$$\begin{aligned} \Pr(y = 1|x) &= \Pr(y^* > 0|x) \\ &= \Pr(\beta_0 + \beta_1 x + \epsilon > 0|x) \\ &= \Pr(\epsilon > -(\beta_0 + \beta_1 x)) \\ &= \Pr(\epsilon < \beta_0 + \beta_1 x) \\ &= F(\beta_0^L + \beta_1^L x) \end{aligned}$$

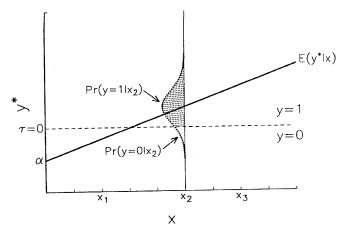


Figure 3.4. Probability of Observed Values in the Binary Response Model

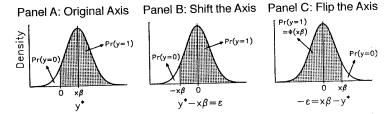
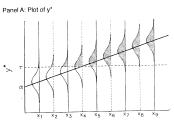


Figure 3.5. Computing Pr(y = 1 | x) in the Binary Response Model

Since we assume the errors follow a standard logistic distribution, we have

$$Pr(y = 1|x) = F(\beta_0^L + \beta_1^L x)$$
$$= \frac{\exp(\beta_0^L + \beta_1^L x)}{1 + \exp(\beta_0^L + \beta_1^L x)}$$

$$E(\epsilon)=0$$
 and $Var(\epsilon)=\frac{\pi^2}{3}$.



Panel B: Plot of Pr(y=1lx)

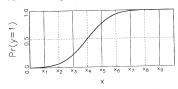


Figure 3.6. Plot of y^* and Pr(y = 1 | x) in the Binary Response Model

The logit function is the inverse of the logistic function:

$$\mathsf{logit}(p) = \mathsf{log}\frac{p}{1-p}$$

or

$$\operatorname{logit}^{-1}(p) = \frac{\operatorname{exp}(x)}{1 + \operatorname{exp}(x)}$$

We therefore have the following

$$\Pr(y = 1|x) = \operatorname{logit}^{-1}(\beta_1^L + \beta_1^L x)$$

or

$$logit(Pr(y = 1|x)) = \beta_1^L + \beta_1^L x$$

or

$$\log \frac{\Pr(y = 1|x)}{\Pr(y = 0|x)} = \beta_0^L + \beta_1^L x.$$

Recall from lecture that the Bernoulli distribution has the following pdf:

$$\Pr(y_i = 1 | \pi_i) = \pi_i^{y_i} (1 - \pi_i)^{1 - y_i}$$

And the likelihood function can be derived from the joint probability:

$$egin{aligned} \mathcal{L}(oldsymbol{\pi}|oldsymbol{y}) &\propto \prod_{i=1}^n \pi_i^{y_i} (1-\pi_i)^{1-y_i} \ \mathcal{L}(oldsymbol{eta}|oldsymbol{y}) &\propto \prod_{i=1}^n \left(rac{1}{1+\exp(-oldsymbol{x}_ioldsymbol{eta})}
ight)^{y_i} \left(1-rac{1}{1+\exp(-oldsymbol{x}_ioldsymbol{eta})}
ight)^{1-y_i} \ \mathcal{L}(oldsymbol{eta}|oldsymbol{y}) &\propto \prod_{i=1}^n (1+\exp(-oldsymbol{x}_ioldsymbol{eta}))^{-y_i} (1+\exp(-oldsymbol{x}_ioldsymbol{eta}))^{-(1-y_i)} \ \log \mathcal{L}(oldsymbol{eta}|oldsymbol{y}) &\propto \sum_{i=1}^n -y_i \log(1+\exp(-oldsymbol{x}_ioldsymbol{eta})) - (1-y_i) \log(1+\exp(oldsymbol{x}_ioldsymbol{eta})) \end{aligned}$$

```
rm(list = ls()) # clear up the memory
#install and load the packages needed
#from CRAN: install.packages("MASS", dependencies = TRUE)
library(MASS)
library(RColorBrewer)

# download simcf and tile packages
# from- http://faculty.washington.edu/cadolph/software
# don't unzip the archive (tar) file
library(simcf)
library(tile)
```

Loading required package: grid

```
# Load data
file <- "nes00a.csv"
data <- read.csv(file, header=TRUE)
# attach(data)</pre>
```

```
# Estimate logit model using optim()
# Construct variables and model objects
v <- data$vote00
x <- cbind(data$age,data$hsdeg,data$coldeg)
# Likelihood function for logit
llk.logit <- function(param, v, x) {</pre>
  os <- rep(1,length(x[,1])) # constant
  x <- cbind(os,x) # constant+covariates
  b <- param[ 1 : ncol(x) ]
  # number of parameters to be estimated equals number of columns in x
  # (i.e, one for constant and one for each covariates : total 4)
  xb <- x%*%b
  sum(y*log(1+exp(-xb)) + (1-y)*log(1+exp(xb)))
  # log-likelihood function for logit model
  # (based on our choice of standard logistic cdf as the systematic component)
               # optim is a minimizer, so use -lnL here
```

```
# Fit logit model using optim
ls.result \leftarrow lm(y\sim x) # use is estimates as starting values (for convenience)
stval <- ls.result$coefficients # initial guesses
logit.result.opt <- optim(stval,llk.logit,method="BFGS",hessian=T,y=y,x=x)</pre>
# call minimizer procedure or max by adding control=list(fnscale=-1)
pe.opt <- logit.result.opt$par # point estimates</pre>
vc.opt <- solve(logit.result.opt$hessian) # var-cov matrix</pre>
se.opt <- sqrt(diag(vc.opt)) # standard errors</pre>
11.opt <- -logit.result.opt$value # likelihood at maximum
logit.optim<-data.frame(cbind(round(pe.opt,3), round(se.opt,3)))</pre>
rownames(logit.optim)<-c("intercept", "age", "highschool" , "college")</pre>
colnames(logit.optim)<-c("pe", "std.err")</pre>
logit.optim
```

```
## pe std.err
## intercept -2.149 0.257
## age 0.031 0.003
## highschool 1.213 0.179
## college 1.102 0.130
```

```
#p-value based on t-statistics
2*pt(abs(logit.optim$pe/logit.optim$std.err),
     df=length(y)-length(pe.opt) , lower.tail = FALSE)
## [1] 1.229175e-16 2.393414e-24 1.668155e-11 4.780627e-17
# Estimate logit model using qlm()
# Run logit & extract results using qlm.
# GLM solves the likelihood equations with a common numeric algorithm
# called iteratively re-weighted least squares (IRWLS).
logit.result <- glm(vote00~age +hsdeg+coldeg, family=binomial, data=data)</pre>
# family "binomial" calls logit transformation
# (log(pi/1-pi)) as a "link function"
# corresponding to logistic distribution.
# Link function transforms pi so that it follows a linear model.
# (So although pi itself is dependent on covariates in a non-linear way,
# logit transformed pi is dependent on covariates in a linear way.)
```

summary(logit.result)

```
##
## Call:
## glm(formula = vote00 ~ age + hsdeg + coldeg, family = binomial,
      data = data)
##
## Deviance Residuals:
##
      Min
               10 Median
                                30
                                        Max
## -2.4487 -1.1347 0.6360 0.8973 1.8854
##
## Coefficients:
##
              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -2.147997  0.256608 -8.371 < 2e-16 ***
## age
              0.030885 0.003386 9.121 < 2e-16 ***
## hsdeg 1.212882 0.179447 6.759 1.39e-11 ***
## coldeg 1.102465 0.130426 8.453 < 2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
      Null deviance: 2293.5 on 1782 degrees of freedom
## Residual deviance: 2076.0 on 1779 degrees of freedom
## ATC: 2084
##
## Number of Fisher Scoring iterations: 4
```

```
# now a new model adding age 2
model <- vote00 ~ age + I(age^2) + hsdeg + coldeg
mdata <- extractdata(model, data, na.rm=TRUE) # needs library(simcf)</pre>
logit.result <- glm(model, family=binomial, data=mdata)</pre>
pe <- logit.result$coefficients # point estimates</pre>
vc <- vcov(logit.result) # var-cov matrix</pre>
ре
    (Intercept) age I(age^2) hsdeg
##
                                                                coldeg
## -3.0193890720 0.0747251922 -0.0004427014 1.1243907749 1.0795702357
sqrt(diag(vc))
```

```
## (Intercept) age I(age^2) hsdeg coldeg
## 0.4181899367 0.0168440337 0.0001655466 0.1800068893 0.1312113265
```

```
# Simulate parameter distributions
sims < -10000
simbetas <- mvrnorm(sims, pe, vc) #needs library(MASS)</pre>
# draw 10000 sets of simulated
# parameter (beta) estimates from a multivariate normal distribution with
# mean pe and variance-covariance vc
# Now let's plan counterfactuals: We will have three
# sets of counterfactuals based on education lavel
# (less than hs edu, hs edu, college or higher edu), and for each set
# we will make age varies between 18 years old and 97 years old.
# Set up counterfactuals: all ages, each of three educations
xhyp <- seq(18,97,1) # create age vector</pre>
nscen <- length(xhyp) # we will have total 80 different age scenarios
#for each education level
nohsScen <- hsScen <- collScen <- cfMake(model, mdata, nscen)
#this is just to initialize 80 scenarios for each education level.
#As default, all covariate values are set at the mean.
```

```
# Create three sets of education counterfactuals
for (i in 1:nscen) {
  # No High school scenarios (loop over each age, total 80 scenarios)
  nohsScen <- cfChange(nohsScen, "age", x = xhyp[i], scen = i)</pre>
  nohsScen <- cfChange(nohsScen, "hsdeg", x = 0, scen = i) #no hs degree
  nohsScen <- cfChange(nohsScen, "coldeg", x = 0, scen = i) #no college degree
  # HS grad scenarios (loop over each age, total 80 scenarios)
  hsScen <- cfChange(hsScen, "age", x = xhyp[i], scen = i)
  hsScen <- cfChange(hsScen, "hsdeg", x = 1, scen = i) #has hs degree
  hsScen <- cfChange(hsScen, "coldeg", x = 0, scen = i) #no college degree
  # College grad scenarios (loop over each age, total 80 scenarios)
  collScen <- cfChange(collScen, "age", x = xhyp[i], scen = i)</pre>
  collScen <- cfChange(collScen, "hsdeg", x = 1, scen = i) #has hs degree
  collScen <- cfChange(collScen, "coldeg", x = 1, scen = i) #has college degree
```

```
# Now given the counterfactual covariates (nohsScen/hsScen/collScen)
# and simulated parameters (simbetas), we can calculate expected value
# of the response. In this case, expected probability of voting!

head(nohsScen$x) #we will fit the counterfactual data
```

```
## vote00 age hsdeg coldeg
## 1 0.6567583 18 0 0
## 2 0.6567583 19 0 0
## 3 0.6567583 20 0 0
## 4 0.6567583 21 0 0
## 5 0.6567583 22 0 0
## 6 0.6567583 23 0
```

```
nohsScen$model # in the model specification
```

```
## vote00 ~ age + I(age^2) + hsdeg + coldeg
```

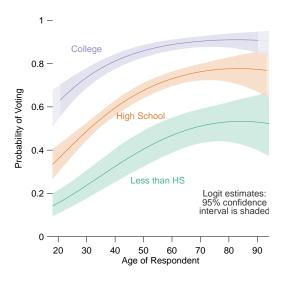
```
nohsSims <- logitsimev(nohsScen, simbetas, ci=0.95) # using simulated
# betas to get expected values.
# Built-in function "logitsimev" calculates the expected value
# for every individual scenario you created.
# reports lower and upper confidence intervals
# as well as expected probabilities

# same thing for two other sets of sceneraios
hsSims <- logitsimev(hsScen, simbetas, ci=0.95)
collSims <- logitsimev(collScen, simbetas, ci=0.95)
```

```
# Get 3 nice colors for traces
col <- brewer.pal(3,"Dark2")</pre>
# Set up lineplot traces of expected probabilities
#Traces are elements of tile: lines, labels, legends...
# no hs
nohsTrace <- lineplot(x=xhyp, # age on x-axis
  y=nohsSims$pe, #expected probability on y-axis
  lower=nohsSims$lower, # lower confidence interval
  upper=nohsSims$upper, #upper confidence interval
  col=col[1], #color choice
  extrapolate=list(data=mdata[,2:ncol(mdata)],
  #actual covariates (i.e., values in your data)
  cfact=nohsScen$x[,2:ncol(hsScen$x)],#counterfactual covariates
  omit.extrapolated=FALSE), #don't show extrapolated values
  plot=1)
```

```
# hs but no college
hsTrace <- lineplot(x=xhyp,
  y=hsSims$pe,
  lower=hsSims$lower.
  upper=hsSims$upper,
  col=col[2].
  extrapolate=list(data=mdata[,2:ncol(mdata)],
  cfact=hsScen$x[,2:ncol(hsScen$x)],
  omit.extrapolated=FALSE),
  plot=1)
#college
collTrace <- lineplot(x=xhyp,</pre>
  v=collSims$pe,
  lower=collSims$lower,
  upper=collSims$upper,
  col=col[3],
  extrapolate=list(data=mdata[,2:ncol(mdata)],
  cfact=collScen$x[,2:ncol(hsScen$x)],
  omit.extrapolated=FALSE),
  plot=1)
```

```
# Set up traces with labels
labelTrace <- textTile(labels=c("Less than HS", "High School", "College"),
    x=c(55, 49,
                       30).
    y=c( 0.26, 0.56, 0.87).
    col=col,
    plot=1)
# For legend
legendTrace <-
 textTile(labels=c("Logit estimates:", "95% confidence", "interval is shaded"),
 x=c(82, 82, 82),
 v=c(0.2, 0.16, 0.12).
 plot=1)
#options(device="quartz")
# Plot traces using tile
voting<-tile(nohsTrace.
    hsTrace.
    collTrace.
    labelTrace.
    legendTrace.
    limits=c(18,94,0,1),
    xaxis=list(at=c(20,30,40,50,60,70,80,90)),
    vaxis=list(label.loc=-0.5, major=FALSE).
    xaxistitle=list(labels="Age of Respondent"),
    vaxistitle=list(labels="Probability of Voting"),
    width=list(null=5.vaxistitle=4.vaxis.labelspace=-0.5)
     .output=list(file="educationEV".width=5.5)
```

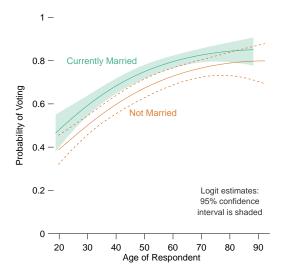


```
# Now consider a new specification adding the variable
# "ever married", or marriedo
# We will estimate this new model with qlm(), then
# simulate new scenarios for marrieds and non-marrieds
# Estimate logit model using qlm()
# Set up a new model formula and model specific data frame
model2 <- vote00 ~ age + I(age^2) + hsdeg + coldeg + marriedo
mdata2 <- extractdata(model2, data, na.rm=TRUE)</pre>
# Run logit & extract results
logit.m2 <- glm(model2, family=binomial, data=mdata2)</pre>
pe.m2 <- logit.m2$coefficients # point estimates</pre>
vc.m2 <- vcov(logit.m2) # var-cov matrix
# Simulate parameter distributions
sims <- 10000
simbetas.m2 <- mvrnorm(sims, pe.m2, vc.m2)</pre>
```

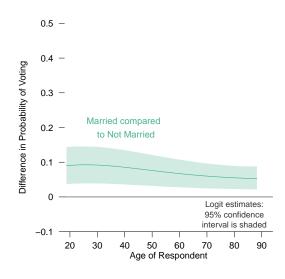
```
# Set up counterfactuals: all ages
xhvp < -seg(18.97.1)
nscen <- length(xhvp)
marriedScen <- notmarrScen <- cfMake(model2, mdata2, nscen)
for (i in 1:nscen) {
 # - we will use the marriedScen counterfactuals in FDs and RRs as well as EVs
  # Note below the careful use of before scenarios (xpre) and after scenarios (x)
  #:i.e., use of the same age range (18-97) for both x and xpre, only marriedo values differ.
  # Married (loop over each age)
 marriedScen <- cfChange(marriedScen, "age", x = xhyp[i], xpre= xhyp[i], scen = i)
 marriedScen <- cfChange(marriedScen, "marriedo", x = 1, xpre= 0, scen = i)
  # Not Married (loop over each age)
 notmarrScen <- cfChange(notmarrScen, "age", x = xhyp[i], scen = i)
 notmarrScen <- cfChange(notmarrScen, "marriedo", x = 0, scen = i)
# Simulate expected probabilities for all age scenarios for married and not married respectively
marriedSims <- logitsimev(marriedScen, simbetas.m2, ci=0.95)
notmarrSims <- logitsimev(notmarrScen, simbetas.m2, ci=0.95)
# Simulate first difference of voting wrt marriage: E(y|married)-E(y|notmarried)
marriedFD <- logitsimfd(marriedScen, simbetas.m2, ci=0.95)
# Simulate relative risk of voting wrt marriage: E(y/married)/E(y/notmarried)
marriedRR <- logitsimrr(marriedScen, simbetas.m2, ci=0.95)
```

```
## Make plots using tile
# Get 3 nice colors for traces
col <- brewer.pal(3,"Dark2")</pre>
# Set up lineplot traces of expected probabilities
marriedTrace <- lineplot(x=xhyp,
                         y=marriedSims$pe,
                         lower=marriedSims$lower.
                         upper=marriedSims$upper,
                         col=col[1],
                         extrapolate=list(data=mdata2[,2:ncol(mdata2)],
                                           cfact=marriedScen$x[.2:ncol(marriedScen$x)].
                                           omit.extrapolated=TRUE),
                         plot=1)
notmarrTrace <- lineplot(x=xhyp,
                         y=notmarrSims$pe,
                         lower=notmarrSims$lower.
                         upper=notmarrSims$upper,
                         col=col[2],
                         ci = list(mark="dashed").
                         extrapolate=list(data=mdata2[,2:ncol(mdata2)],
                                           cfact=notmarrScen$x[,2:ncol(notmarrScen$x)],
                                           omit.extrapolated=TRUE).
                         plot=1)
```

```
# Set up traces with labels and legend
labelTrace <- textTile(labels=c("Currently Married", "Not Married"),</pre>
                       x=c(35.53).
                       y=c(0.8, 0.56),
                       col=col,
                       plot=1)
legendTrace <- textTile(labels=c("Logit estimates:", "95% confidence", "interval is shaded"),</pre>
                        x=c(80, 80, 80),
                        v=c(0.2, 0.15, 0.10).
                        cex=0.9,
                        plot=1)
# Plot traces using tile
tile(marriedTrace.
     notmarrTrace,
    labelTrace.
    legendTrace.
     limits=c(18,94,0,1),
     xaxis=list(at=c(20,30,40,50,60,70,80,90)),
     vaxis=list(label.loc=-0.5, major=FALSE).
     xaxistitle=list(labels="Age of Respondent"),
     vaxistitle=list(labels="Probability of Voting"),
     width=list(null=5,yaxistitle=4,yaxis.labelspace=-0.5)
     #.output=list(file="marriedEV",width=5.5)
```



```
# Set up traces with labels and legend
labelFDTrace <- textTile(labels=c("Married compared \n to Not Married"),
                         x=c(40).
                         y=c(0.20),
                         col=col[1],
                         plot=1)
legendFDTrace <- textTile(labels=c("Logit estimates:", "95% confidence", "interval is shaded"),</pre>
                          x=c(80, 80, 80),
                          v=c(-0.02, -0.05, -0.08).
                          cex=0.9,
                          plot=1)
# Plot traces using tile
tile(marriedFDTrace,
     labelFDTrace,
     legendFDTrace.
     baseline.
     limits=c(18,94,-0.1,0.5),
     xaxis=list(at=c(20,30,40,50,60,70,80,90)),
     vaxis=list(label.loc=-0.5, major=FALSE).
     xaxistitle=list(labels="Age of Respondent"),
     vaxistitle=list(labels="Difference in Probability of Voting"),
     width=list(null=5,yaxistitle=4,yaxis.labelspace=-0.5)
     #.output=list(file="marriedFD",width=5.5)
```



```
# Set up traces with labels and legend
labelRRTrace <- textTile(labels=c("Married compared \n to Not Married"),
                         x=c(55).
                         y=c(1.25),
                         col=col[1],
                         plot=1)
legendRRTrace <- textTile(labels=c("Logit estimates:", "95% confidence", "interval is shaded"),</pre>
                          x=c(80, 80, 80),
                          v=c(0.98, 0.95, 0.92).
                          cex=0.9,
                          plot=1)
# Plot traces using tile
tile(marriedRRTrace,
     labelRRTrace,
     legendRRTrace.
     baseline.
     limits=c(18,94,0.9,1.5),
     xaxis=list(at=c(20,30,40,50,60,70,80,90)),
     vaxis=list(label.loc=-0.5, major=FALSE).
     xaxistitle=list(labels="Age of Respondent"),
     vaxistitle=list(labels="Relative Risk of Voting"),
     width=list(null=5,yaxistitle=4,yaxis.labelspace=-0.5)
     #.output=list(file="marriedRR",width=5.5)
```

