We toss a coin 10 times and get 7 tails. Is this sufficient evidence to conclude that the coin is biased?

We toss a coin 10 times and get 7 tails. Is this sufficient evidence to conclude that the coin is biased?

The **null hypothesis**, H_0 , states that "nothing extraordinary is going on". So in this case

$$H_0$$
: $P(T) = \frac{1}{2}$

We toss a coin 10 times and get 7 tails. Is this sufficient evidence to conclude that the coin is biased?

The **null hypothesis**, H_0 , states that "nothing extraordinary is going on". So in this case

$$H_0$$
: $P(T) = \frac{1}{2}$

The alternative hypothesis, H_A , states that there is a different chance process that generates the data. Here we can take

$$H_A$$
: $P(T) \neq \frac{1}{2}$

We toss a coin 10 times and get 7 tails. Is this sufficient evidence to conclude that the coin is biased?

The **null hypothesis**, H_0 , states that "nothing extraordinary is going on". So in this case

$$H_0$$
: $P(T) = \frac{1}{2}$

The alternative hypothesis, H_A , states that there is a different chance process that generates the data. Here we can take

$$H_A$$
: $P(T) \neq \frac{1}{2}$

Hypothesis testing proceeds by collecting data and evaluating whether the data are compatible with H_0 or not (in which case one rejects H_0).

A different example: A company develops a new drug to lower blood pressure. It tests it with an experiment involving 1,000 patients.

A different example: A company develops a new drug to lower blood pressure. It tests it with an experiment involving 1,000 patients.

In this case "nothing extraordinary going on" means that the drug has no effect.

A different example: A company develops a new drug to lower blood pressure. It tests it with an experiment involving 1,000 patients.

In this case "nothing extraordinary going on" means that the drug has no effect. So

 H_0 : no change in blood pressure H_A : blood pressure drops

A different example: A company develops a new drug to lower blood pressure. It tests it with an experiment involving 1,000 patients.

In this case "nothing extraordinary going on" means that the drug has no effect. So

 H_0 : no change in blood pressure H_A : blood pressure drops

Note that in this case the company would like to reject $H_0!$

A different example: A company develops a new drug to lower blood pressure. It tests it with an experiment involving 1,000 patients.

In this case "nothing extraordinary going on" means that the drug has no effect. So

 H_0 : no change in blood pressure H_A : blood pressure drops

Note that in this case the company would like to reject $H_0!$

So the logic of testing is typically indirect: One assumes that nothing extraordinary is happening and then hopes to reject this assumption H_0 .

A test statistic measures how far away the data are from what we would expect if H_0 were true.

A **test statistic** measures how far away the data are from what we would expect if H_0 were true.

The most common test statistic is the **z-statistic**:

$$z = \frac{\text{observed} - \text{expected}}{\text{SE}}$$

A test statistic measures how far away the data are from what we would expect if H_0 were true.

The most common test statistic is the **z-statistic**:

$$z = \frac{\text{observed} - \text{expected}}{\text{SE}}$$

'Observed' is a statistic that is appropriate for assessing H_0 . In the example of the 10 coin tosses, appropriate statistics would be the number of tails or the percent of tails.

A test statistic measures how far away the data are from what we would expect if H_0 were true.

The most common test statistic is the **z-statistic**:

$$z = \frac{\mathsf{observed} - \mathsf{expected}}{\mathsf{SE}}$$

'Observed' is a statistic that is appropriate for assessing H_0 . In the example of the 10 coin tosses, appropriate statistics would be the number of tails or the percent of tails.

'Expected' and SE are the expected value and the SE of this statistic, computed under the assumption that H_0 is true.

A **test statistic** measures how far away the data are from what we would expect if H_0 were true.

The most common test statistic is the **z-statistic**:

$$z = \frac{\mathsf{observed} - \mathsf{expected}}{\mathsf{SE}}$$

'Observed' is a statistic that is appropriate for assessing H_0 . In the example of the 10 coin tosses, appropriate statistics would be the number of tails or the percent of tails.

'Expected' and SE are the expected value and the SE of this statistic, computed under the assumption that H_0 is true.

In the example: Using the formulas for the sum of 0/1 labels we get 'expected' = $10 \times \frac{1}{2} = 5$ and SE = $\sqrt{10} \sqrt{\frac{1}{2} \times \frac{1}{2}} = 1.58$. So

$$z = \frac{7-5}{1.58} = 1.27$$

p-values measure the evidence agains H_0

Large values of |z| are evidence against H_0 : The larger |z| is, the stronger the evidence. The strength of the evidence is measured by the **p-value** (or: **observed significance level**):

p-values measure the evidence agains H₀

Large values of |z| are evidence against H_0 : The larger |z| is, the stronger the evidence. The strength of the evidence is measured by the **p-value** (or: **observed significance level**):

The p-value is the probability of getting a value of z as extreme or more extreme than the observed z, assuming H_0 is true.

p-values measure the evidence agains H₀

Large values of |z| are evidence against H_0 : The larger |z| is, the stronger the evidence. The strength of the evidence is measured by the **p-value** (or: **observed significance level**):

The p-value is the probability of getting a value of z as extreme or more extreme than the observed z, assuming H_0 is true.

But if H_0 is true, then z follows that standard normal curve, according to the central limit theorem, so the p-value can be computed with normal approximation:

p-values measure the evidence agains H₀

Large values of |z| are evidence against H_0 : The larger |z| is, the stronger the evidence. The strength of the evidence is measured by the **p-value** (or: **observed significance level**):

The p-value is the probability of getting a value of z as extreme or more extreme than the observed z, assuming H_0 is true.

But if H_0 is true, then z follows that standard normal curve, according to the central limit theorem, so the p-value can be computed with normal approximation:

The smaller the p-value, the stronger the evidence against H_0 . Often the criterion for rejecting H_0 is a p-value smaller than 5%. Then the result is called **statistically significant**.

p-values measure the evidence against H₀

In the example:

p-values measure the evidence against H₀

In the example:

Note that the p-value **does not** give the probability that H_0 is true, as H_0 is either true or not - there are no chances involved. Rather, it gives the probability of seeing a statistic as extreme, or more extreme, that the observed one, assuming H_0 is true.

It has been said that it is difficult to distinguish Coke and Pepsi by taste alone, without the visual cue of the bottle or can.

It has been said that it is difficult to distinguish Coke and Pepsi by taste alone, without the visual cue of the bottle or can.

In an experiment that I did in a class at Stanford, 10 cups were filled at random with either Coke or Pepsi. A student volunteer tasted each of the 10 cups and correctly named the contents of seven. Is this sufficient evidence to conclude that the student can tell apart Coke and Pepsi?

It has been said that it is difficult to distinguish Coke and Pepsi by taste alone, without the visual cue of the bottle or can.

In an experiment that I did in a class at Stanford, 10 cups were filled at random with either Coke or Pepsi. A student volunteer tasted each of the 10 cups and correctly named the contents of seven. Is this sufficient evidence to conclude that the student can tell apart Coke and Pepsi?

"Nothing extraordinary is going on" means that the student does not have any special ability to tell them apart and is just guessing.

It has been said that it is difficult to distinguish Coke and Pepsi by taste alone, without the visual cue of the bottle or can.

In an experiment that I did in a class at Stanford, 10 cups were filled at random with either Coke or Pepsi. A student volunteer tasted each of the 10 cups and correctly named the contents of seven. Is this sufficient evidence to conclude that the student can tell apart Coke and Pepsi?

"Nothing extraordinary is going on" means that the student does not have any special ability to tell them apart and is just guessing.

To write this down formally we introduce 0/1 labels since we are counting correct answers: 1 = correct answer, 0 = wrong answer

It has been said that it is difficult to distinguish Coke and Pepsi by taste alone, without the visual cue of the bottle or can.

In an experiment that I did in a class at Stanford, 10 cups were filled at random with either Coke or Pepsi. A student volunteer tasted each of the 10 cups and correctly named the contents of seven. Is this sufficient evidence to conclude that the student can tell apart Coke and Pepsi?

"Nothing extraordinary is going on" means that the student does not have any special ability to tell them apart and is just guessing.

To write this down formally we introduce 0/1 labels since we are counting correct answers: 1 = correct answer, 0 = wrong answer

$$H_0: P(0) = P(1) = \frac{1}{2}$$
 $H_A: P(1) > \frac{1}{2}$

It has been said that it is difficult to distinguish Coke and Pepsi by taste alone, without the visual cue of the bottle or can.

In an experiment that I did in a class at Stanford, 10 cups were filled at random with either Coke or Pepsi. A student volunteer tasted each of the 10 cups and correctly named the contents of seven. Is this sufficient evidence to conclude that the student can tell apart Coke and Pepsi?

"Nothing extraordinary is going on" means that the student does not have any special ability to tell them apart and is just guessing.

To write this down formally we introduce 0/1 labels since we are counting correct answers: 1 = correct answer, 0 = wrong answer

$$H_0: P(0) = P(1) = \frac{1}{2}$$
 $H_A: P(1) > \frac{1}{2}$

This is a **one-sided test**: the alternative hypothesis for for P(1) we are interested in is on one side of $\frac{1}{2}$.

Since we are looking at the sum of ten 0/1 labels, the z-statistic is the same that we had for coin-tossing:

$$z = rac{ ext{observed sum} - ext{expected sum}}{ ext{SE of sum}} = rac{7-5}{1.58} = 1.27$$

Since we are looking at the sum of ten 0/1 labels, the z-statistic is the same that we had for coin-tossing:

$$z = \frac{\text{observed sum} - \text{expected sum}}{\text{SE of sum}} = \frac{7-5}{1.58} = 1.27$$

But since we do a one-sided test instead of a two-sided test, the p-value is only half as large:

Since we are looking at the sum of ten 0/1 labels, the z-statistic is the same that we had for coin-tossing:

$$z \ = \ \frac{\text{observed sum} - \text{expected sum}}{\text{SE of sum}} \ = \ \frac{7-5}{1.58} \ = \ 1.27$$

But since we do a one-sided test instead of a two-sided test, the p-value is only half as large:

Since 10.2% is not smaller than 5%, we don't reject H_0 : We are not convinced that the student can distinguish Coke and Pepsi.

Distinguishing Coke and Pepsi

A two-sided alternative might also be appropriate:

$$H_A: P(1) \neq \frac{1}{2}$$

 H_A corresponds to a student who is more likely than not to distinguish Coke and Pepsi, but who may confuse them. Such a student might get one correct answer (say).

Distinguishing Coke and Pepsi

A two-sided alternative might also be appropriate:

$$H_A: P(1) \neq \frac{1}{2}$$

 H_A corresponds to a student who is more likely than not to distinguish Coke and Pepsi, but who may confuse them. Such a student might get one correct answer (say).

One has to carefully consider whether the alternative should be one-sided or two-sided, as the p-value gets doubled in the latter case.

Distinguishing Coke and Pepsi

A two-sided alternative might also be appropriate:

$$H_A$$
: $P(1) \neq \frac{1}{2}$

 H_A corresponds to a student who is more likely than not to distinguish Coke and Pepsi, but who may confuse them. Such a student might get one correct answer (say).

One has to carefully consider whether the alternative should be one-sided or two-sided, as the p-value gets doubled in the latter case.

It is not ok to change the alternative afterwards in order to get the p-value below 5%.

The health guideline for lead in drinking water is a concentration of not more than 15 parts per billion (ppb).

The health guideline for lead in drinking water is a concentration of not more than 15 parts per billion (ppb).

Five independent samples from a reservoir average 15.6 ppb. Is this sufficient evidence to conclude that the concentration μ in the reservoir is above the standard of 15 ppb?

The health guideline for lead in drinking water is a concentration of not more than 15 parts per billion (ppb).

Five independent samples from a reservoir average 15.6 ppb. Is this sufficient evidence to conclude that the concentration μ in the reservoir is above the standard of 15 ppb?

Recall our model for measurements:

measurement = μ + measurement error

The health guideline for lead in drinking water is a concentration of not more than 15 parts per billion (ppb).

Five independent samples from a reservoir average 15.6 ppb. Is this sufficient evidence to conclude that the concentration μ in the reservoir is above the standard of 15 ppb?

Recall our model for measurements:

 $measurement = \mu + measurement error$

So it may be that the concentration μ is below 15 ppb, but measurement error results in an average of 15.6 ppb.

The health guideline for lead in drinking water is a concentration of not more than 15 parts per billion (ppb).

Five independent samples from a reservoir average 15.6 ppb. Is this sufficient evidence to conclude that the concentration μ in the reservoir is above the standard of 15 ppb?

Recall our model for measurements:

$$measurement = \mu + measurement error$$

So it may be that the concentration μ is below 15 ppb, but measurement error results in an average of 15.6 ppb.

$$H_0$$
: $\mu = 15 \text{ ppb}$ H_A : $\mu > 15 \text{ ppb}$

The health guideline for lead in drinking water is a concentration of not more than 15 parts per billion (ppb).

Five independent samples from a reservoir average 15.6 ppb. Is this sufficient evidence to conclude that the concentration μ in the reservoir is above the standard of 15 ppb?

Recall our model for measurements:

$$measurement = \mu + measurement error$$

So it may be that the concentration μ is below 15 ppb, but measurement error results in an average of 15.6 ppb.

$$H_0$$
: $\mu = 15 \text{ ppb}$ H_A : $\mu > 15 \text{ ppb}$

We can try a z-test for the average of the measurements:

$$z \ = \ \frac{\text{observed average} - \text{expected average}}{\text{SE of average}} \ = \ \frac{15.6 \text{ ppb} - 15 \text{ ppb}}{\text{SE of average}}$$

since the measurement error has expected value zero.

SE of average $=\frac{\sigma}{\sqrt{n}}$, but the standard deviation σ of the measurement error is unknown.

SE of average = $\frac{\sigma}{\sqrt{n}}$, but the standard deviation σ of the measurement error is unknown.

We can estimate σ by s, the sample standard deviation of the measurements. However:

SE of average $=\frac{\sigma}{\sqrt{n}}$, but the standard deviation σ of the measurement error is unknown.

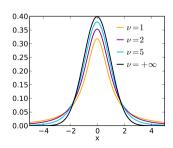
We can estimate σ by s, the sample standard deviation of the measurements. However:

If we estimate σ and n is small ($n \leq 20$), then the normal curve is not a good enough approximation to the distribution of the z-statistic. Rather, an appropriate approximation is **Student's t-distribution with** n-1 **degrees of freedom**:

SE of average $=\frac{\sigma}{\sqrt{n}}$, but the standard deviation σ of the measurement error is unknown.

We can estimate σ by s, the sample standard deviation of the measurements. However:

If we estimate σ and n is small ($n \leq 20$), then the normal curve is not a good enough approximation to the distribution of the z-statistic. Rather, an appropriate approximation is **Student's t-distribution with** n-1 **degrees of freedom**:



The fatter tails account for the additional uncertainty introduced by estimating σ by $s = \sqrt{\frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})^2}$.

The fatter tails account for the additional uncertainty introduced by estimating σ by $s = \sqrt{\frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})^2}$.

Using the **t-test** in place of the z-test is only necessary for small samples: $n \leq 20$ (say).

The fatter tails account for the additional uncertainty introduced by estimating σ by $s = \sqrt{\frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})^2}$.

Using the **t-test** in place of the z-test is only necessary for small samples: $n \leq 20$ (say).

In that case it is also better to replace the confidence interval $\bar{x}\pm z$ SE by

$$\bar{x} \pm t_{n-1} \mathsf{SE}$$

▶ Statistically significant does not mean that the effect size is important:

► Statistically significant does not mean that the effect size is important:

Suppose the sample average shows a lead concentration that is only slightly above the health standard of 15 ppb: say the sample average is 15.05 ppb.

► Statistically significant does not mean that the effect size is important: Suppose the sample average shows a lead concentration that is only slightly above the health standard of 15 ppb: say the sample average is 15.05 ppb.

That may not be of practical concern, even though the test may be highly significant: Statistical significance convinces us that there is an effect, but it doesn't say how big the effect is.

► Statistically significant does not mean that the effect size is important: Suppose the sample average shows a lead concentration that is only slightly above the health standard of 15 ppb: say the sample average is 15.05 ppb.

That may not be of practical concern, even though the test may be highly significant: Statistical significance convinces us that there is an effect, but it doesn't say how big the effect is.

Reason: A large sample size n makes $SE = \frac{\sigma}{\sqrt{n}}$ small, so even a small exceedance over the limit by (say) 0.05 ppb may give a statistically significant result.

► Statistically significant does not mean that the effect size is important: Suppose the sample average shows a lead concentration that is only slightly above the health standard of 15 ppb: say the sample average is 15.05 ppb.

That may not be of practical concern, even though the test may be highly significant: Statistical significance convinces us that there is an effect, but it doesn't say how big the effect is.

Reason: A large sample size n makes $SE = \frac{\sigma}{\sqrt{n}}$ small, so even a small exceedance over the limit by (say) 0.05 ppb may give a statistically significant result.

Therefore it is helpful to complement a test with a confidence interval: In the above case a 95% confidence interval for μ might be [15.02 ppb, 15.08 ppb].

▶ There is a general connection between confidence intervals and tests:

▶ There is a general connection between confidence intervals and tests:

A 95% confidence interval contains all values for the null hypothesis that will not be rejected by a two-sided test at a 5% significance level.

(A 5% significance level means that the threshold for the p-value is 5%).

▶ There is a general connection between confidence intervals and tests:

A 95% confidence interval contains all values for the null hypothesis that will not be rejected by a two-sided test at a 5% significance level.

(A 5% significance level means that the threshold for the p-value is 5%).

▶ There are two ways that a test can result in a wrong decision:

▶ There is a general connection between confidence intervals and tests:

A 95% confidence interval contains all values for the null hypothesis that will not be rejected by a two-sided test at a 5% significance level.

(A 5% significance level means that the threshold for the p-value is 5%).

There are two ways that a test can result in a wrong decision:

 H_0 is true, but was erroneously rejected $\to \mathsf{Type}\;\mathsf{I}\;\mathsf{error}\;\mathsf{(`false\;positive')}$

▶ There is a general connection between confidence intervals and tests:

A 95% confidence interval contains all values for the null hypothesis that will not be rejected by a two-sided test at a 5% significance level. (A 5% significance level means that the threshold for the p-value is 5%).

► There are two ways that a test can result in a wrong decision:

 H_0 is true, but was erroneously rejected o Type I error ('false positive')

 H_0 is false, but we fail to reject it \rightarrow Type II error

▶ There is a general connection between confidence intervals and tests:

A 95% confidence interval contains all values for the null hypothesis that will not be rejected by a two-sided test at a 5% significance level.

(A 5% significance level means that the threshold for the p-value is 5%).

▶ There are two ways that a test can result in a wrong decision:

 H_0 is true, but was erroneously rejected o Type I error ('false positive')

 H_0 is false, but we fail to reject it \rightarrow Type II error

Rejecting H $_0$ if the p-value is smaller than 5% means P(type I error) $\leq 5\%$

Last month, the President's approval rating in a sample of 1,000 likely voters was 55%. This month, a poll of 1,500 likely voters resulted in a rating of 58%. Is this sufficient evidence to conclude that the rating has changed?

Last month, the President's approval rating in a sample of 1,000 likely voters was 55%. This month, a poll of 1,500 likely voters resulted in a rating of 58%. Is this sufficient evidence to conclude that the rating has changed?

We want to assess whether

 $p_1 =$ proportion of all likely voters approving last month is equal to

 p_2 =proportion of all likely voters approving this month

Last month, the President's approval rating in a sample of 1,000 likely voters was 55%. This month, a poll of 1,500 likely voters resulted in a rating of 58%. Is this sufficient evidence to conclude that the rating has changed?

We want to assess whether

 $p_1 =$ proportion of all likely voters approving last month is equal to

 p_2 =proportion of all likely voters approving this month

"nothing unusual is going on" means $p_1 = p_2$.

Last month, the President's approval rating in a sample of 1,000 likely voters was 55%. This month, a poll of 1,500 likely voters resulted in a rating of 58%. Is this sufficient evidence to conclude that the rating has changed?

We want to assess whether

 $p_1 =$ proportion of all likely voters approving last month is equal to

 p_2 =proportion of all likely voters approving this month

"nothing unusual is going on" means $p_1=p_2$. It's common to look at the difference p_2-p_1 instead:

$$\mathsf{H}_0:\ p_2-p_1=0 \qquad \qquad \mathsf{H}_1:\ p_2-p_1\neq 0$$

Last month, the President's approval rating in a sample of 1,000 likely voters was 55%. This month, a poll of 1,500 likely voters resulted in a rating of 58%. Is this sufficient evidence to conclude that the rating has changed?

We want to assess whether

 $p_1 =$ proportion of all likely voters approving last month is equal to

 $p_2 =$ proportion of all likely voters approving this month

"nothing unusual is going on" means $p_1=p_2$. It's common to look at the difference p_2-p_1 instead:

$$\mathsf{H}_0:\ p_2-p_1=0$$
 $\mathsf{H}_1:\ p_2-p_1\neq 0$

 p_1 is estimated by $\hat{p}_1=55\%$, p_2 by $\hat{p}_2=58\%$.

Last month, the President's approval rating in a sample of 1,000 likely voters was 55%. This month, a poll of 1,500 likely voters resulted in a rating of 58%. Is this sufficient evidence to conclude that the rating has changed?

We want to assess whether

 $p_1 =$ proportion of all likely voters approving last month is equal to

 p_2 =proportion of all likely voters approving this month

"nothing unusual is going on" means $p_1=p_2$. It's common to look at the difference p_2-p_1 instead:

$$H_0: p_2 - p_1 = 0$$
 $H_1: p_2 - p_1 \neq 0$

 p_1 is estimated by $\hat{p}_1 = 55\%$, p_2 by $\hat{p}_2 = 58\%$. The central limit theorem applies to the difference $\hat{p}_2 - \hat{p}_1$ just as it does to \hat{p}_1 and \hat{p}_2 . So we can use a z-test:

We can use a z-test for the difference $\hat{p}_2 - \hat{p}_1$:

We can use a z-test for the difference $\hat{p}_2 - \hat{p}_1$:

$$z=rac{ ext{observed difference}- ext{expected difference}}{ ext{SE of difference}}=rac{(\hat{p}_2-\hat{p}_1)-(p_2-p_1)}{ ext{SE of difference}}$$

We can use a z-test for the difference $\hat{p}_2 - \hat{p}_1$:

$$z \ = \ \frac{\text{observed difference} - \text{expected difference}}{\text{SE of difference}} \ = \ \frac{(\hat{p}_2 - \hat{p}_1) - (p_2 - p_1)}{\text{SE of difference}}$$

An important fact is that if \hat{p}_1 and \hat{p}_2 are independent, then

$$SE(\hat{p}_2 - \hat{p}_1) = \sqrt{(SE(\hat{p}_1))^2 + (SE(\hat{p}_2))^2}.$$

We can use a z-test for the difference $\hat{p}_2 - \hat{p}_1$:

$$z \ = \ \frac{\text{observed difference} - \text{expected difference}}{\text{SE of difference}} \ = \ \frac{(\hat{p}_2 - \hat{p}_1) - (p_2 - p_1)}{\text{SE of difference}}$$

An important fact is that if \hat{p}_1 and \hat{p}_2 are independent, then

$$SE(\hat{p}_2 - \hat{p}_1) = \sqrt{(SE(\hat{p}_1))^2 + (SE(\hat{p}_2))^2}.$$

$$z = \frac{(\hat{p}_2 - \hat{p}_1) - 0}{\sqrt{\sqrt{\frac{p_1(1 - p_1)}{1000}^2} + \sqrt{\frac{p_2(1 - p_2)}{1500}^2}}} =$$

We can use a z-test for the difference $\hat{p}_2 - \hat{p}_1$:

$$z=rac{ ext{observed difference}- ext{expected difference}}{ ext{SE of difference}}=rac{(\hat{p}_2-\hat{p}_1)-(p_2-p_1)}{ ext{SE of difference}}$$

An important fact is that if \hat{p}_1 and \hat{p}_2 are independent, then

$$SE(\hat{p}_2 - \hat{p}_1) = \sqrt{(SE(\hat{p}_1))^2 + (SE(\hat{p}_2))^2}. So$$

$$z = \frac{(\hat{p}_2 - \hat{p}_1) - 0}{\sqrt{\sqrt{\frac{p_1(1-p_1)}{1000}^2} + \sqrt{\frac{p_2(1-p_2)}{1500}^2}}} = \frac{0.03}{0.0202} = 1.48$$

The two-sample z-test is applicable in the same way to the difference of two sample means in order to test for equality of two population means.

The two-sample z-test is applicable in the same way to the difference of two sample means in order to test for equality of two population means.

If the two samples are independent, then again

$$SE(\bar{x}_2 - \bar{x}_1) = \sqrt{(SE(\bar{x}_1))^2 + (SE(\bar{x}_2))^2}$$

The two-sample z-test is applicable in the same way to the difference of two sample means in order to test for equality of two population means.

If the two samples are independent, then again

$$SE(\bar{x}_2 - \bar{x}_1) = \sqrt{(SE(\bar{x}_1))^2 + (SE(\bar{x}_2))^2}$$

and
$$SE(\bar{x}_1) = \frac{\sigma_1}{\sqrt{n_1}}$$
 is estimated by $\frac{s_1}{\sqrt{n_1}}$.

The two-sample z-test is applicable in the same way to the difference of two sample means in order to test for equality of two population means.

If the two samples are independent, then again

$$SE(\bar{x}_2 - \bar{x}_1) = \sqrt{(SE(\bar{x}_1))^2 + (SE(\bar{x}_2))^2}$$

and
$${\sf SE}(\bar{x}_1) = \frac{\sigma_1}{\sqrt{n_1}}$$
 is estimated by $\frac{s_1}{\sqrt{n_1}}.$

All of the above two-sample tests require that the two samples are independent. They are also applicable in special situations where the samples are dependent, e.g. to compare the treatment effect when subjects are randomized into treatment and control groups.

The paired-difference test

Do husbands tend to be older than their wives?

Do husbands tend to be older than their wives?

The ages of five couples:

Husband's age	Wife's age	age difference
43	41	2
71	70	1
32	31	1
68	66	2
27	26	1

Do husbands tend to be older than their wives?

The ages of five couples:

Husband's age	Wife's age	age difference
43	41	2
71	70	1
32	31	1
68	66	2
27	26	1

The two-sample t-test is not applicable since the two samples are not independent.

Do husbands tend to be older than their wives?

The ages of five couples:

Husband's age	Wife's age	age difference
43	41	2
71	70	1
32	31	1
68	66	2
27	26	1

The two-sample t-test is not applicable since the two samples are not independent. Even if they were independent, the small differences in ages would not be significant since the standard deviations are large for husbands and also for the wives.

Since we have paired data, we can simply analyze the differences obtained from each pair with a regular t-test, which in this context of **matched pairs** is called **paired t-test**:

Since we have paired data, we can simply analyze the differences obtained from each pair with a regular t-test, which in this context of **matched pairs** is called **paired t-test**:

H₀: population difference has mean zero

Since we have paired data, we can simply analyze the differences obtained from each pair with a regular t-test, which in this context of **matched pairs** is called **paired t-test**:

H₀: population difference has mean zero

 $t=rac{ar{d}-0}{\mathsf{SE}(ar{d})}$, where d_i is the age difference of the ith couple.

Since we have paired data, we can simply analyze the differences obtained from each pair with a regular t-test, which in this context of **matched pairs** is called **paired t-test**:

H₀: population difference has mean zero

 $t=rac{ar{d}-0}{\mathsf{SE}(ar{d})}$, where d_i is the age difference of the ith couple.

 $SE(\bar{d}) = \frac{\sigma_d}{\sqrt{n}}$. Estimate σ_d by $s_d = 0.55$.

Since we have paired data, we can simply analyze the differences obtained from each pair with a regular t-test, which in this context of **matched pairs** is called **paired t-test**:

H₀: population difference has mean zero

 $t = \frac{\bar{d}-0}{\mathsf{SF}(\bar{d})}$, where d_i is the age difference of the ith couple.

$$SE(\bar{d}) = \frac{\sigma_d}{\sqrt{n}}$$
. Estimate σ_d by $s_d = 0.55$. Then $t = \frac{1.4 - 0}{0.55/\sqrt{5}} = 5.69$

Since we have paired data, we can simply analyze the differences obtained from each pair with a regular t-test, which in this context of **matched pairs** is called **paired t-test**:

 H_0 : population difference has mean zero

$$t=\frac{\bar{d}-0}{\mathsf{SF}(\bar{d})}$$
, where d_i is the age difference of the i th couple.

$$SE(\bar{d}) = \frac{\sigma_d}{\sqrt{n}}$$
. Estimate σ_d by $s_d = 0.55$. Then $t = \frac{1.4 - 0}{0.55/\sqrt{5}} = 5.69$

The independence assumption is in the sampling of the couples.

What if didn't know the age difference d_i but only if the husband was older or not?

What if didn't know the age difference d_i but only if the husband was older or not?

We can test

H₀: half the husbands in the population are older than their wives

What if didn't know the age difference d_i but only if the husband was older or not?

We can test

 H_0 : half the husbands in the population are older than their wives using 0/1 labels and a z-test, just as we tested whether a coin is fair:

What if didn't know the age difference d_i but only if the husband was older or not?

We can test

 H_0 : half the husbands in the population are older than their wives using 0/1 labels and a z-test, just as we tested whether a coin is fair:

$$z = \frac{{
m sum \ of \ 1s} - \frac{n}{2}}{{
m SE \ of \ sum}} = \frac{5 - \frac{5}{2}}{\sqrt{5} \frac{1}{2}} = \ 2.24 \quad {
m since} \ \sigma = \frac{1}{2} \ {
m on} \ {
m H}_0.$$

What if didn't know the age difference d_i but only if the husband was older or not?

We can test

 H_0 : half the husbands in the population are older than their wives using 0/1 labels and a z-test, just as we tested whether a coin is fair:

$$z = \frac{\text{sum of } 1\text{s} - \frac{n}{2}}{\text{SE of sum}} = \frac{5 - \frac{5}{2}}{\sqrt{5}\frac{1}{2}} = 2.24$$
 since $\sigma = \frac{1}{2}$ on H_0 .

The p-value of this **sign-test** is less significant than that of the paired t-test. This is because the latter uses more information, namely the size of the differences.

What if didn't know the age difference d_i but only if the husband was older or not?

We can test

 H_0 : half the husbands in the population are older than their wives using 0/1 labels and a z-test, just as we tested whether a coin is fair:

$$z = \frac{\text{sum of 1s} - \frac{n}{2}}{\text{SE of sum}} = \frac{5 - \frac{5}{2}}{\sqrt{5} \frac{1}{2}} = 2.24$$
 since $\sigma = \frac{1}{2}$ on H_0 .

The p-value of this **sign-test** is less significant than that of the paired t-test. This is because the latter uses more information, namely the size of the differences. On the other hand, the sign test has the virtue of easy interpretation due to the analogy to coin tossing.

Mini quiz

1. True or false:

- a. The p-value depends on the data.
- b. If the p-value is smaller than 5%, then there is less than a 5% chance that the null hypothesis is true.
- c. If the null hypothesis is true, then there is less than a 5% chance to get a p-value that is smaller than 5%.
- d. If a data scientist does many tests, then even if the all the null hypotheses are true, a certain proportion will be rejected in error.

- 2. For each of the following situations, indicate which test is appropriate to address the respective question: z-test, t-test, two-sample z-test, sign test, or paired-difference test.
 - a. You want to test whether plain M&Ms really contain 24% blue M&Ms as claimed on the manufacturer's web site. You sample 500 plain M&Ms at random and count the fraction of blue M&Ms
 - b. A high school principal wants to find out whether the average SAT score of this year's graduating class is higher than last year's. She samples 13 students from this year's graduating class at random and wants to compare their average SAT score to the average SAT score from last year's graduating class.
 - c. To investigate whether there are difference in scholastic abilities between first-borns and second-born siblings, 600 families that have at least two children were randomly selected. The scholastic abilities of the first-born and the second-born siblings were assessed with a test and are to be compared.