STAT 433: HOMEWORK 1

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Problem 1. A poker hand consists of five cards drawn from a standard 52-card deck. Find the expected number of aces in a poker hand given that the first card drawn is an ace.

Solution. Let X be a random variable denoting the number of aces in a poker hand, and let 1_A be an indicator function that equals 1 if the first card drawn is an ace and 0 otherwise. Then,

$$E(X \mid 1_A) = \frac{E(X1_A)}{P(A)}$$

$$= \frac{1}{P(A)} \sum_{x=1}^{4} x P(\{X = x\} \cap A)$$

$$= \sum_{x=1}^{4} x P(X = x \mid A)$$

$$= \frac{\binom{3}{x-1} \binom{48}{5-x}}{\binom{51}{4}} = \frac{21}{17}.$$

Problem 2. An urn has n balls. Balls are drawn one at a time and then put back in the urn. Let X denote the number of drawings required until some ball is drawn more than once. Find the probability distribution of X.

Solution. On the first draw, a ball cannot be drawn more than once, so

$$P(X = 1) = 0.$$

On the second draw, the probability of drawing the same ball, i.e., the ball from the first trial, is

$$P(X=2) = \frac{1}{n}.$$

To deduce the probability that X = 3, note that with probability 1 the same ball is not drawn on the first trial (i.e., 1 - P(X = 1)), on the second draw, the same ball is not drawn with probability 1 - 1/n (i.e., 1 - P(X = 2)), and on the third drawn, there are 2 balls in the urn containing n ball which can be drawn again. Hence

$$P(X=3) = \left(1 - \frac{1}{n}\right) \frac{2}{n}.$$

Continuing in this fashion, we obtain a probability distribution given by

$$P(X = k) = \prod_{i=0}^{k} \left(1 - \frac{i}{n}\right) \frac{k-1}{n}$$

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Problem 3. A man with n keys wants to open his door. He tries the keys in a random manner. Let X be the number of trials required to open the door.

(i) Find E(X) and var(X) if unsuccessful keys are eliminated from further selection. Solution. Note that, since the number of keys n decreases by 1 each trial, we have

$$P(X = 1) = \frac{1}{n}$$

$$P(X = 2) = \left(1 - \frac{1}{n}\right) \frac{1}{n - 1}$$

$$\vdots$$

$$P(X = k) = \left(1 - \frac{1}{n}\right) \left(1 - \frac{1}{n - 1}\right) \cdots \left(1 - \frac{1}{n - (k - 2)}\right) \frac{1}{n - (k - 1)},$$

which can be expressed more compactly as

$$P(X = k) = \prod_{i=0}^{k-1} \left(1 - \frac{1}{n-i}\right).$$

Then, compute

$$E(X) = \frac{1}{n} + \frac{2}{n} + \dots + \frac{n}{n} = \frac{n(n+1)}{2n} = \frac{n+1}{2}$$

and

$$E(X^2) = \frac{1}{n} + \frac{4}{n} + \dots + \frac{n^2}{n} = \frac{n(n+1)(2n+1)}{6n} = \frac{(n+1)(2n+1)}{6}.$$

So, using the fact that $var(X) = EX^2 - (EX)^2$, we get

$$var(X) = \frac{(n+1)(2n+1)}{6} - \left(\frac{n+1}{2}\right)^2 = \frac{n^2 - 1}{12}.$$

(ii) Find E(X) and var(X) if unsuccessful keys are not eliminated from further selection.

Solution. If unsuccessful keys not are eliminated from further selection, then $X \sim \text{Geom}(1/n)$, and hence

$$E(X) = \frac{1}{1/n} = n$$
$$var(X) = \frac{1 - 1/n}{1/n^2} = n(n - 1).$$

Problem 4. Let X_1, X_2, \ldots, X_n be i.i.d. random variables with $X_i \sim \mathcal{U}([0, t])$, and let $Y_n = \min\{X_1, \ldots, X_n\}$.

(i) Find $P(Y_n > x)$ for $x \in [0, t]$.

Solution. Since Y_n is greater than the minimum of X_1, X_2, \dots, X_n , we know Y_n must also be greater than every X_1, X_2, \dots, X_n . So,

$$P(Y_n > x) = P(X_1 > x, X_2 > x, \dots, X_n > x)$$

$$= \prod_{i=1}^{n} \left(1 - \frac{x}{t}\right)$$

$$= \left(1 - \frac{x}{t}\right)^n$$

(ii) Let $t \text{ map } n \mapsto t(n)$ such that $\lim_{n \to \infty} n/t(n) = \lambda$. Show that

$$\lim_{n \to \infty} P(Y_n > x) = e^{-\lambda x}.$$

Solution. Evaluating the limit gives

$$\lim_{n \to \infty} P(Y_n > x) = \lim_{n \to \infty} \left(1 - \frac{x}{t(n)} \right)^n$$

$$= \lim_{n \to \infty} \left(1 - \frac{n}{n} \frac{x}{t(n)} \right)^n$$

$$= \lim_{n \to \infty} \left(1 - \frac{(n/t(n))x}{n} \right)^n$$

$$= \lim_{n \to \infty} \left(1 - \frac{\lambda x}{n} \right)^n$$

$$= e^{-\lambda x}.$$

since by definition t satisfies $\lim_{n\to\infty} n/t(n) = \lambda$.