Homework 2: Linear Regression

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SOLUTION

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In class we studied the linear regression model $\mathbf{y} = \mathbf{X}\boldsymbol{\theta}^* + \boldsymbol{\epsilon}$, under the homoskedastic assumption $\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I})$. In this homework you will derive the same results for the slightly more general heteroskedastic model where $\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}^*)$. Each subproblem is worth 10 points.

Problem 2.1. Derive an expression for the coefficient vector $\boldsymbol{\theta}$ that minimizes the mean squared error, i.e.,

$$\underset{\boldsymbol{\theta}}{\operatorname{arg\,min}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\theta}\|_{2}^{2}.$$

Solution. This is the same as in the homoskedastic case, $\hat{\boldsymbol{\theta}} = (\mathbf{X}^\mathsf{T} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \mathbf{y}$.

Problem 2.2. Derive an expression for the maximum likelihood estimator (MLE) of θ^* , i.e.,

$$\mathop{\arg\max}_{\boldsymbol{\theta}} \; \mathbb{P}(\mathbf{y}, \mathbf{X} | \boldsymbol{\theta}, \boldsymbol{\Sigma}^{\star})$$

Solution. For this, write:

$$\begin{split} \hat{\boldsymbol{\theta}} &= \underset{\boldsymbol{\theta}}{\operatorname{arg\,max}} \ \mathbb{P}(\mathbf{y}, \mathbf{X} | \boldsymbol{\theta}, \boldsymbol{\Sigma}^{\star}) \\ &= \underset{\boldsymbol{\theta}}{\operatorname{arg\,max}} \ \frac{1}{\sqrt{(2\pi)^{\mathrm{N}} |\boldsymbol{\Sigma}^{\star}|}} e^{-\frac{1}{2}(\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^{\mathsf{T}} \boldsymbol{\Sigma}^{\star - 1} (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})} \\ &= \underset{\boldsymbol{\theta}}{\operatorname{arg\,max}} \ - (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^{\mathsf{T}} \boldsymbol{\Sigma}^{\star - 1} (\mathbf{y} - \mathbf{X}\boldsymbol{\theta}) \\ &= \underset{\boldsymbol{\theta}}{\operatorname{arg\,min}} \ \mathbf{y}^{\mathsf{T}} \boldsymbol{\Sigma}^{\star - 1} \mathbf{y} - \mathbf{y}^{\mathsf{T}} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X}\boldsymbol{\theta} - \boldsymbol{\theta}^{\mathsf{T}} \mathbf{X}^{\mathsf{T}} \boldsymbol{\Sigma}^{\star - 1} \mathbf{y} + \boldsymbol{\theta}^{\mathsf{T}} \mathbf{X}^{\mathsf{T}} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X}\boldsymbol{\theta} \\ &= \underset{\boldsymbol{\theta}}{\operatorname{arg\,min}} \ -2\boldsymbol{\theta}^{\mathsf{T}} \mathbf{X}^{\mathsf{T}} \boldsymbol{\Sigma}^{\star - 1} \mathbf{y} + \boldsymbol{\theta}^{\mathsf{T}} \mathbf{X}^{\mathsf{T}} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X}\boldsymbol{\theta}. \end{split}$$

Taking derivative with respect to θ and setting to zero, we have:

$$-2\mathbf{X}^\mathsf{T}\mathbf{\Sigma}^{\star-1}\mathbf{y} + 2\mathbf{X}^\mathsf{T}\mathbf{\Sigma}^{\star-1}\mathbf{X}\boldsymbol{\theta} \ = \ \mathbf{0},$$

and solving we obtain:

$$\hat{\boldsymbol{\theta}} = (\mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{y}.$$

Problem 2.3. What is the distribution of the MLE of θ^* ?

Solution. Since $\mathbf{y} \sim \mathcal{N}(\mathbf{X}\boldsymbol{\theta}^{\star}, \boldsymbol{\Sigma}^{\star})$,

$$\begin{split} \hat{\boldsymbol{\theta}} &\sim \mathcal{N} \Big((\mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X} \boldsymbol{\theta}^{\star}, (\mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \boldsymbol{\Sigma}^{\star} ((\mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1})^\mathsf{T} \Big) \\ &= \mathcal{N} \Big(\boldsymbol{\theta}^{\star}, (\mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X} (\mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X})^{-1} \Big) \\ &= \mathcal{N} \Big(\boldsymbol{\theta}^{\star}, (\mathbf{X}^\mathsf{T} \boldsymbol{\Sigma}^{\star - 1} \mathbf{X})^{-1} \Big). \end{split}$$

Problem 2.4. Given a new sample with feature vector \mathbf{x} , what is the MLE of the response, \hat{y} ?

Solution. By the invariance property of the MLE, $\hat{y} = \mathbf{x}^{\mathsf{T}} \hat{\boldsymbol{\theta}}$.

Problem 2.5. Given a new sample with feature vector \mathbf{x} , what is the distribution of the MLE \hat{y} ?

Solution. Since
$$\hat{\boldsymbol{\theta}} \sim \mathcal{N}(\boldsymbol{\theta}^{\star}, (\mathbf{X}^{\mathsf{T}}\boldsymbol{\Sigma}^{\star-1}\mathbf{X})^{-1})$$
, we have that $\hat{y} \sim \mathcal{N}(\mathbf{x}\boldsymbol{\theta}^{\star}, \mathbf{x}(\mathbf{X}^{\mathsf{T}}\boldsymbol{\Sigma}^{\star-1}\mathbf{X})^{-1}\mathbf{x}^{\mathsf{T}})$.

Problem 2.6. Derive an expression for the MLE of Σ^* , i.e.,

$$\argmax_{\boldsymbol{\Sigma}} \ \mathbb{P}(\mathbf{y}, \mathbf{X} | \boldsymbol{\theta}^{\star}, \boldsymbol{\Sigma})$$

Solution. For this, write:

$$\hat{\boldsymbol{\theta}} = \underset{\boldsymbol{\Sigma}}{\operatorname{arg \, max}} \ \mathbb{P}(\mathbf{y}, \mathbf{X} | \boldsymbol{\theta}, \boldsymbol{\Sigma})$$

$$= \underset{\boldsymbol{\Sigma}}{\operatorname{arg \, max}} \ \frac{1}{\sqrt{(2\pi)^N |\boldsymbol{\Sigma}|}} e^{-\frac{1}{2}(\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^\mathsf{T} \boldsymbol{\Sigma}^{-1}(\mathbf{y} - \mathbf{X}\boldsymbol{\theta})}$$

$$= \underset{\boldsymbol{\Sigma}}{\operatorname{arg \, max}} \ -\frac{N}{2} \log(2\pi) - \frac{1}{2} \log |\boldsymbol{\Sigma}| - (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^\mathsf{T} \boldsymbol{\Sigma}^{-1}(\mathbf{y} - \mathbf{X}\boldsymbol{\theta})$$

$$= \underset{\boldsymbol{\Sigma}}{\operatorname{arg \, min}} \ \frac{1}{2} \log |\boldsymbol{\Sigma}| + (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^\mathsf{T} \boldsymbol{\Sigma}^{-1}(\mathbf{y} - \mathbf{X}\boldsymbol{\theta})$$

Taking derivative with respect to Σ and setting to zero, we have:

$$\frac{1}{|\Sigma|} |\Sigma| \Sigma^{-1} - \Sigma^{-1} (\mathbf{y} - \mathbf{X}\boldsymbol{\theta}) (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^{\mathsf{T}} \Sigma^{-1} = \mathbf{0}$$
$$\mathbf{I} - \Sigma^{-1} (\mathbf{y} - \mathbf{X}\boldsymbol{\theta}) (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^{\mathsf{T}} = \mathbf{0},$$

and solving we obtain:

$$\Sigma = (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})(\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^{\mathsf{T}}.$$

Problem 2.7. Consider the following vector \mathbf{y} , containing information about glucose level of four individuals, and the following data matrix \mathbf{X} containing information about height and weight of the corresponding individuals:

$$\mathbf{y} = \begin{bmatrix} 110 \\ 140 \\ 180 \\ 190 \end{bmatrix}, \qquad \mathbf{X} = \begin{bmatrix} 180 & 150 \\ 150 & 175 \\ 170 & 165 \\ 185 & 210 \end{bmatrix}.$$

Given these data

- (a) What are your maximum likelihood estimates of Σ^* and θ^* ?
- (b) Given a new sample with feature vector $\mathbf{x} = [175 \ 170]^{\mathsf{T}}$, what is the maximum likelihood estimate of its response $\hat{\mathbf{y}}$?
- (c) Derive a 95% confidence interval for \hat{y} .
- (d) Would you conclude that height is a significant feature for this model? Why?
- (e) Would you conclude that weight is a significant feature for this model? Why?