

Supplementary Materials for “The Effect of Legislature Size on Public Spending: A Meta-Analysis”

Huzeyfe Alptekin* Danilo Freire[†] Umberto Mignozzetti[‡] Catarina Roman[§]

21 June 2020

Contents

A Search Criteria	3
B Article Selection	3
B.1 Exclusion Analysis	4
B.2 Flow Chart	4
B.3 Study Sample	5
C Meta-Analysis Dataset	8
D Descriptive Statistics	8
D.1 Study Year	9
D.2 Frequency of Published Papers	9
D.3 Electoral System	10
D.4 Dependent Variables	11
D.5 Independent Variables	12
D.6 Histogram of the Coefficients and the Standard Errors	13
D.7 Sign Coefficients	15
E Descriptive Statistics of Moderators	15

*Research Associate, Contemporary Brazilian History Research and Documentation Center, School of Social Sciences, Getulio Vargas Foundation, Brazil, huzeyfealptekin@gmail.com.

[†]Postdoctoral Research Associate, The Political Theory Project, Brown University, Providence, RI 02912, USA, danilofreire@brown.edu, <http://danilofreire.github.io>.

[‡]School of International Relations, Fundação Getulio Vargas, São Paulo, SP, Brazil and Wilf Family Department of Politics, NYU, NY, USA, umberto.mig@nyu.edu, <http://umbertomig.com>.

[§]Research Associate, School of International Relations, Fundação Getulio Vargas, São Paulo, SP, Brazil, catarinamroman@gmail.com.

F	Binomial Tests for Coefficient Signs	17
G	Meta-Analysis	19
G.1	Estimation Method	19
G.2	Lower House Size and Expenditure per Capita	20
G.3	Log Lower House Size and Expenditure per Capita	26
G.4	Upper House Size and Expenditure per Capita	26
G.5	Lower House Size and Log Expenditure Per Capita	29
G.6	Log of Lower House Size and Log of Expenditure Per Capita	31
G.7	Log of Upper House Size and Log of Expenditure Per Capita	34
G.8	Lower House Size and Expenditure as Percentage of GDP	34
G.9	Log Lower House Size and Expenditure as Percentage of GDP	36
G.10	Upper House Size and Expenditure as Percentage of GDP	38
H	Meta-Analysis (All Coefficients)	41
H.1	Lower House Size and Expenditure Per Capita	41
H.2	Log of Lower House Size and Expenditure Per Capita	50
H.3	Upper House Size and Expenditure Per Capita	50
H.4	Lower House Size and Log of Expenditure Per Capita	54
H.5	Log of Lower House Size and Log of Expenditure Per Capita	58
H.6	Upper House Size and Log of Expenditure Per Capita	61
H.7	Lower House Size and Expenditure as Percentage of GDP	61
H.8	Log of Lower House Size and Expenditure as Percentage of GDP	65
H.9	Upper House Size and Expenditure as Percentage of GDP	68
I	Meta-Regressions	71
I.1	Meta-Regressions for Expenditure as a Percentage of the GDP	71
I.2	Meta-Regressions for Expenditure Per Capita	79
I.3	Meta-Regressions for the Log of Expenditure Per Capita	87
J	Robustness: Full Model Meta-Regressions Combined	96
K	Auxiliary Functions	103
K.1	Function to Generate Meta-Analytic Figures	103
K.2	Webscraping Code	105

A Search Criteria

The first step in our systematic review consisted in gathering a study sample. We started our data collection with a manual search based on a set of keywords we scouted from the distributive politics literature. This search produced a database with many entries that were unrelated to our subject of investigation. To reduce the number of false positives in our sample, we restricted our search to studies that cited Weingast, Shepsle and Johnsen’s 1981 paper “The Political Economy of Benefits and Costs: A Neoclassical Approach to Distributive Politics”, which is a seminal contribution in the field. Although [Google Scholar](#) reports the article amounts to 2,180 citations¹, our search resulted in 2,664 records as of the 21st of November 2019.

We webscraped three large academic databases: [Google Scholar](#) (n = 1001); [Microsoft Academic](#) (n = 927); and [Scopus](#) (n = 736). The R script we wrote extracted the article title, abstract, authors, year, journal of publication, and database from which the record originated. Our code is available in section [K.2](#) below. We screened these results with an English language and article restriction, that is, we excluded all records written in other languages and all that were not academic papers, such as book chapters or doctoral theses. To reduce the risk of publication bias, we set no restriction to unpublished articles.

B Article Selection

The selection process was conducted by two authors in three stages. In the first round, we excluded all titles that were obviously unrelated to our topic. For instance, we curiously found articles about automobile motors amidst our sample. We consider this a preliminary step, since we were not able to eliminate a large number of entries. Then, we read all abstracts. We chose to maintain those which indicated that either government expenditure or legislative structures were the main subject of the paper. For instance, if the paper sought to identify variables that increased government size, it was maintained. Abstracts that indicated the paper discussed or estimated the impacts of representative institutions, elections, or chamber dynamics were also included. This allowed us to significantly reduce our sample to 376 records.

In the second phase, we assessed full texts. To remain in our sample, the paper should (i) conduct a quantitative analysis, (ii) report data on the number of legislators, and (iii) also on public expenditure. If the record was marked as positive for all three, it was maintained. Disagreements in this phase were discussed among the authors, and a third investigator was consulted when needed.

¹As of May 11th, 2020.

The third phase consisted of filling out tables for each of the remaining 50 articles to systematically evaluate their eligibility. Since authors use different measures for government spending and the number of lower/upper house members, we extracted all coefficients that provided this information. We decided which variables to keep by following the current practices of the literature. In this phase, we also collected information on whether or not the paper had been published, and if it explicitly discussed the *law of 1/n*. Upon choosing the variables, we excluded the non-conforming studies, arriving at our final sample of 26 articles.

B.1 Exclusion Analysis

We selected the final pool of articles based on two criteria regarding their reported coefficients:

1. Matched treatment variable:

- N : Number Legislators in the Lower House
- $\log N$: Log Number Legislators in the Lower House
- K : Number Legislators in the Upper House

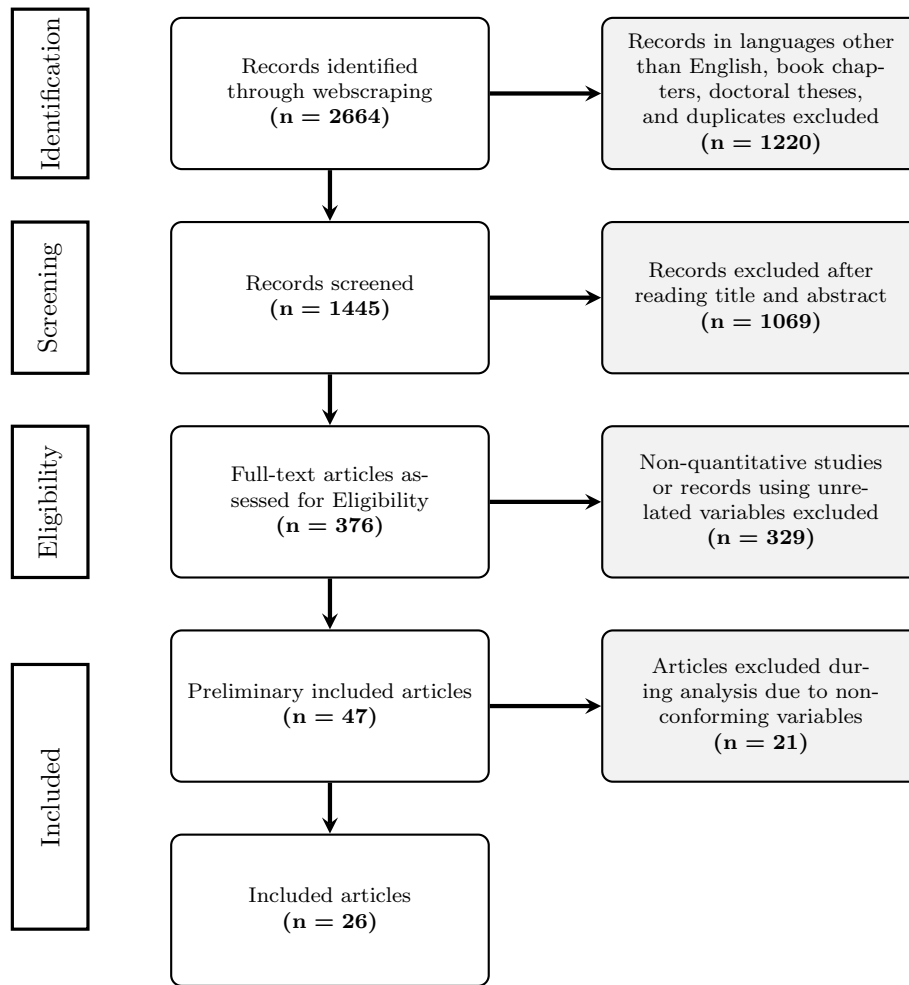
2. Matched outcome variable:

- $ExpPC$: Expenditure Per Capita
- $\log ExpPC$: Log Expenditure Per Capita
- $PCTGDP$: Percent GDP Public Expenditure

B.2 Flow Chart

The diagram below shows each step of our article selection process. We followed the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement to conduct our study². The column to the right depicts the amount of articles excluded in each phase, and the one to the left shows the number of records evaluated.

²More information about the PRISMA statement is available at <http://www.prisma-statement.org>.



B.3 Study Sample

A table with the 26 selected articles is available below. It includes information on the article author, the journal in which the paper was published, the publication year, the country of analysis, the level of data aggregation, and whether the electoral system is majoritarian or not. For unpublished papers, we included the year the paper was made available online.

Table 1: Papers included in the meta-analysis, listed alphabetically by last name

Author(s)	Title	Journal	Year	Country	Level	Electoral System
R. Baqir	Districts, Spillovers, and Government Overspending	Unpub.	1999	USA	Municipal	M
R. Baqir	Districting and Government Overspending	JPE	2002	USA	Municipal	M
T. Baskaran	Coalition Governments, Cabinet Size, and the Common Pool Problem: Evidence from the German States	EJPE	2013	DEU	State	NM

Table 1: Papers included in the meta-analysis, listed alphabetically by last name

Author(s)	Title	Journal	Year	Country	Level	Electoral System
T. Bjedov, S. Lapointe & T. Madiès	The Impact of Within-Party and Between-Party Ideological Dispersion on Fiscal Outcomes: Evidence from Swiss Cantonal Parliaments	PC	2014	CHE	Canton	NM
J. Bradbury & M. Crain	Legislative Organization and Government Spending: Cross-Country Evidence	JPubE	2001	Multiple (37)	Country	NM
J. Bradbury & E. F. Stephenson	Spatially Targeted Government Spending and Heterogeneous Constituent Cost Shares	JPriE	2009	USA	State	M
J. Chen & N. Malhotra	The Law of 1/K: The Effect of Chamber Size on Government Spending in Bicameral Legislatures	APSR	2007	USA	State	M
G. Crowley	The Law of 1/N Revisited: Distributive Politics, Legislature Size, and the Costs of Collective Action	SEJ	2019	USA	State	M
J. Drew & B. Dollery	The Price of Democracy? Political Representation Structure and Per Capita Expenditure in Victorian Local Government	UAR	2017	AUS	State	NM
H. A. Erler	Legislative Term Limits and State Spending	PC	2007	USA	State	M
N. Fiorino & R. Ricciuti	Legislature Size and Government Spending in Italian Regions: Forecasting the Effects of a Reform	PC	2007	ITA	Region	NM
D. Höhmann	The Effect of Legislature Size on Public Spending: Evidence from a Regression Discontinuity Design	PC	2017	DEU	Municipal	NM
A. Kessler	Communication in Federal Politics: Universalism, Policy Uniformity, and the Optimal Allocation of Fiscal Authority	JPE	2014	USA	Municipal	M

Table 1: Papers included in the meta-analysis, listed alphabetically by last name

Author(s)	Title	Journal	Year	Country	Level	Electoral System
D. Lee	Supermajority Rule and Bicameral Bargaining	PC	2016	USA	State	M
D. Lee & S. Park	Court-Ordered Redistricting and the Law of 1/N	PC	2018	USA	District	M
B. Lewis	Legislature Size, Local Government Expenditure and Taxation, and Public Service Access in Indonesia	SCID	2019	IDN	District	NM
V. Lledo	Electoral Systems, Legislative Fragmentation and Public Spending: A Comparative Analysis of Brazilian States	Unpub.	2003	BRA	State	NM
L. MacDonald	The Impact of Government Structure on Local Public Expenditures	PC	2008	USA	Municipal	M
B. Maldonado	Legislature, Leaders, and Leviathans: How Constitutional Institutions Affect the Size of Government Spending	SSQ	2013	Multiple (92)	Country	NM
J. Matsusaka	The Endogeneity of the Initiative: A Comment on Marschall and Ruhil	SPPQ	2005	USA	State	M
B. Mukherjee	Political Parties and the Size of Government in Multiparty Legislatures: Examining Cross-Country and Panel Data Evidence	CPS	2003	Multiple (110)	Country	NM
P. Pettersson-Lidbom	Does the Size of the Legislature Affect the Size of Government? Evidence from Two Natural Experiments	JPubE	2012	FIN & SWE	Municipal	NM
D. Primo	Stop Us before We Spend Again: Institutional Constraints on Government Spending	E&P	2006	USA	State	M
R. Ricciuti	Legislature Size, Bicameralism and Government Spending: Evidence from Democratic Countries	Unpub.	2004	Multiple (75)	Country	NM

Table 1: Papers included in the meta-analysis, listed alphabetically by last name

Author(s)	Title	Journal	Year	Country	Level	Electoral System
C. Schaltegger & L. Feld	Do Large Cabinets Favor Large Governments? Evidence on the Fiscal Commons Problems for Swiss Cantons	JPubE	2009	CHE	Canton	NM
E. Stein, E. Talvi & A. Grisanti	Institutional Arrangements and Fiscal Performance: The Latin American Experience	Unpub.	1998	Multiple (26)	Country	NM

Note: Unpub=Unpublished, JPE=Journal of Political Economy, EJPE=European Journal of Political Economy, PC=Public Choice, JPubE=Journal of Public Economics, JPriE=Journal of Private Enterprise, APSR=American Political Science Review, SEJ=Southern Economic Journal, UAR=Urban Affairs Review, SCID=Studies in Comparative International Development, SSQ=Social Science Quarterly, SPPQ=State Politics and Policy Quarterly, CPS=Comparative Political Studies, E&P=Economics and Politics, M=Majoritarian, NM=Non-majoritarian (mixed or proportional representation). Country codes follow the ISO 3166-1 alpha-3 international standard.

C Meta-Analysis Dataset

Our meta-analytic data is comprised of two datasets. The first dataset has the main coefficients reported in the studies. These data include only the most rigorous model from each paper, that is, those estimated with the largest n , most control variables, and fixed effects if the authors added them. If the article employed a regression discontinuity design, we chose the coefficient from the optimal bandwidth or from the intermediate one. This sample encompasses 36 estimates, as 10 articles analysed two dependent or independent variables of interest. Our second sample, in contrast, contains all the 126 effect sizes reported in the 26 papers.

In the main text, we focus on the results for our restricted sample as we consider them more robust, but the findings are very similar when we use the extended set. Below is the data extraction process for all relevant coefficients in the selected articles. Here we present the results of all tests performed in both reduced and full samples.

D Descriptive Statistics

In this section, we show the descriptive statistics for our study sample. We focus on the following paper characteristics: study year, whether the paper has been published or not, the electoral system of the country discussed in the original study, the data aggregation level, as well as the distribution of the dependent and independent variables of interest. We also add a descriptive statistics table similar to the one in the main paper.

D.1 Study Year

For study year, we have an average of 2009.33, with standard deviation of 6.5. The oldest study included in the paper is dated from 1998, while the most recent paper is dated from 2019. Therefore, we cover 21 years of tests of the *law of 1/n*.

```
dat %>%  
  select(id, year) %>%  
  unique() %>%  
  ggplot(aes(x = as.factor(year))) +  
    geom_bar(color = "black") +  
  labs(x = "",  
        y = "") +  
  theme_bw()
```

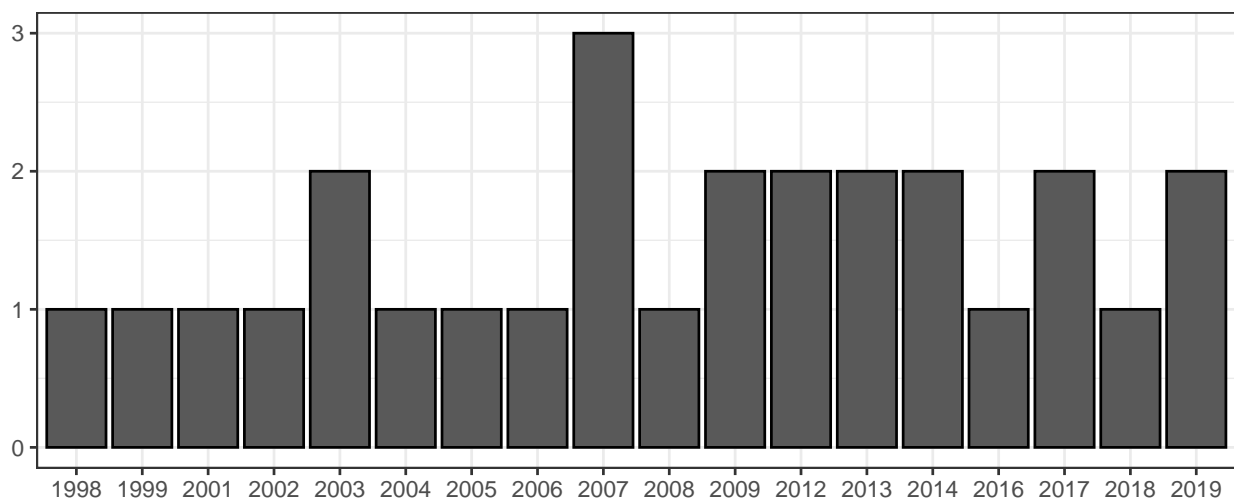


Figure 1: Study Year Frequencies

D.2 Frequency of Published Papers

Studies were included in our sample regardless of their publication status. From the 26 papers in the sample, 22 were published while 4 were not published.

```
dat %>%  
  select(id, published) %>%  
  unique() %>%  
  ggplot(aes(x = as.factor(published))) +
```

```
geom_bar(color = "black") +
labs(x = "Published Study?",
     y = "") +
theme_bw()
```

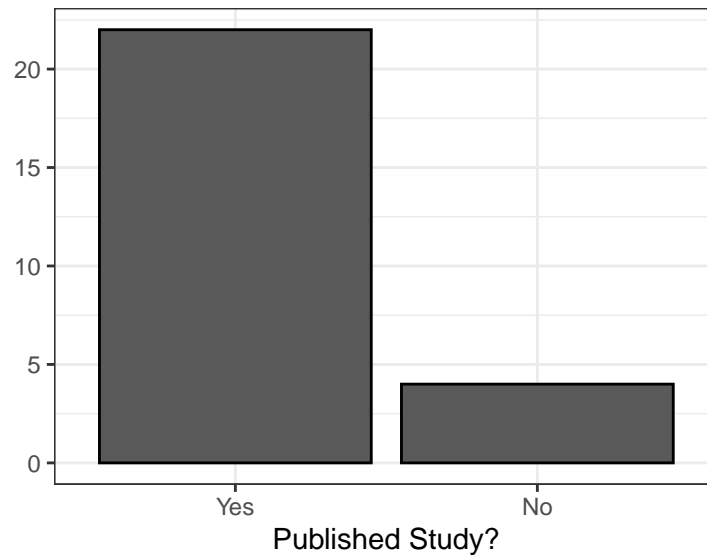


Figure 2: Was the study published?

D.3 Electoral System

Our study selection differs considerably in regards to research design. One remarkable difference is that several authors apply the logics of the *law of 1/n*, which was built with majoritarian systems in mind, to non-majoritarian democracies. In our sample, 12 of the papers study *Majoritarian* systems while 14 study *Non-Majoritarian* electoral systems.³

```
dat %>%
  select(id, elecsys2) %>%
  unique() %>%
  ggplot(aes(x=as.factor(elecsys2))) +
    geom_bar(color = "black") +
  labs(x = "Electoral Systems",
       y = "") +
  theme_bw()
```

³Note that the argument for working in a non-majoritarian system, we need to assume that despite the fact that politicians are able to campaign in every place in the district, the votes are geographically concentrated. The concentration facilitates politicians to use pork-barrel projects to captivate their electoral supporters.

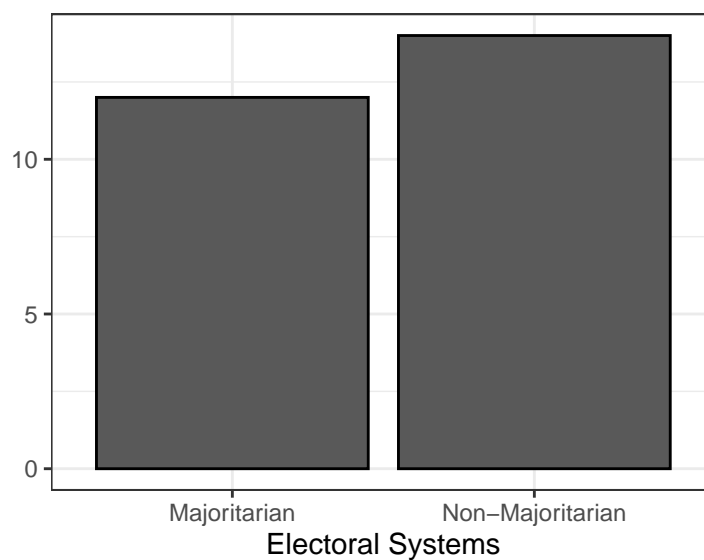


Figure 3: Electoral Systems

D.4 Dependent Variables

The outcome variables included in the paper are:

- 13 Per Capita Expenditure papers
- 7 Natural Log of Per Capita Expenditure papers
- 8 Expenditure as a Percentage of the GDP papers

```
dat %>%
  select(id, depvar2) %>%
  unique() %>%
  mutate(depvar2 = factor(depvar2,
    labels = c("Per Capita Expenditure",
               "Percentage GDP Government Expenditure",
               "Log Per Capita Expenditure"))) %>%

  ggplot(aes(x = depvar2)) +
    geom_bar(color = "black") +
  labs(x = "Dependent Variables",
       y = "") +
  coord_flip() +
  theme_bw()
```



Figure 4: Dependent variables across the law of 1/n studies

D.5 Independent Variables

Most of papers in our sample analyse the number of legislators in the lower house (20). The second most frequent independent variable is the number of legislators in the upper house (9). Finally, the minority of papers use the natural log of the number of legislators in the lower house as an independent variable (5). As we noted above, some papers had multiple coefficients, and thus the total number of coefficients is 36, while the number of papers is only 26.

```
dat %>%
  select(id, indepvar2) %>%
  unique() %>%
  mutate(indepvar2 = factor(indepvar2, labels = c("K (upper house legislators)",
                                                  "N (lower house legislators)",
                                                  "Log N"))) %>%

  ggplot(aes(x = indepvar2)) +
    geom_bar(color = "black") +
    labs(x = "Independent Variables",
         y = "") +
    coord_flip() +
    theme_bw()
```

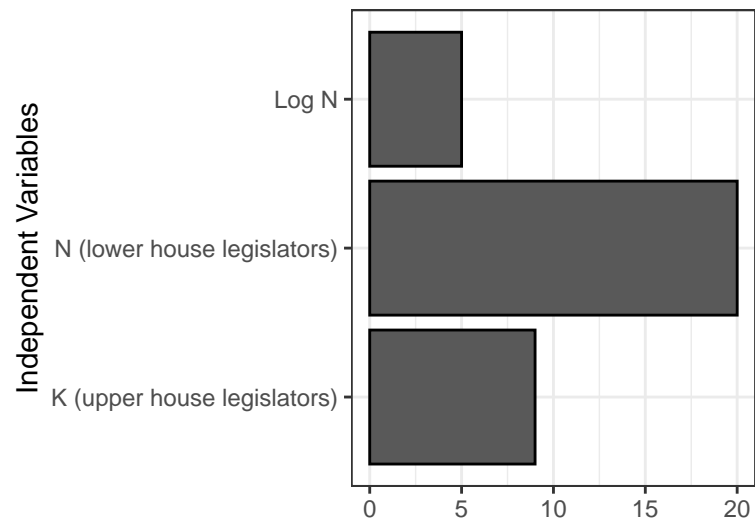


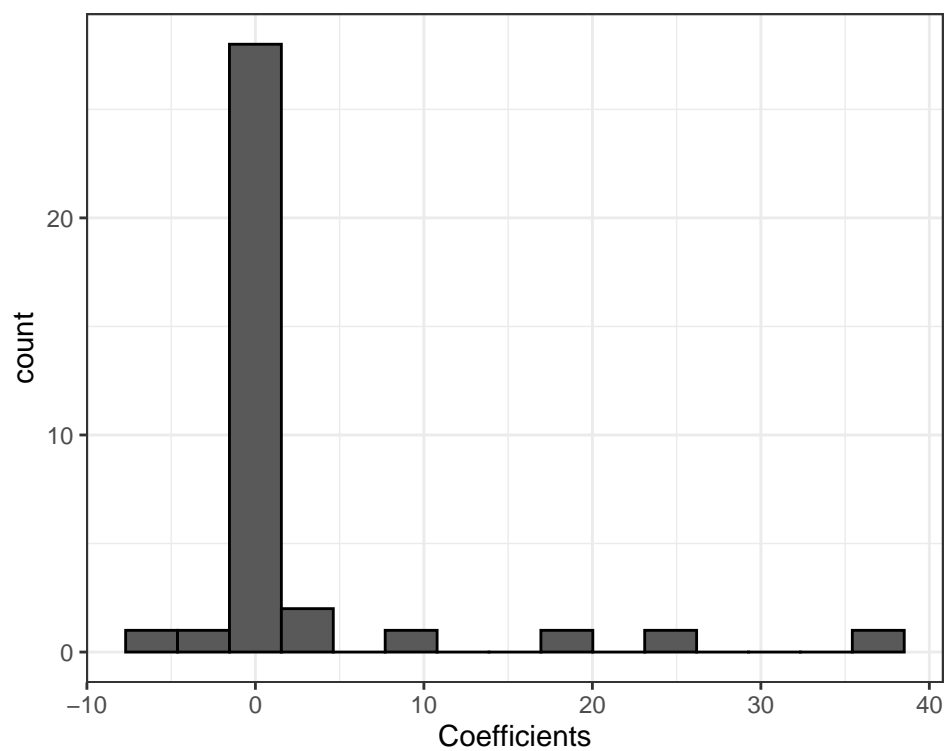
Figure 5: Independent variables across the law of 1/n studies

D.6 Histogram of the Coefficients and the Standard Errors

The coefficients in the papers present a striking variability. In this section, we plot a histogram of the coefficients for all measurements included in the meta-analytic dataset.

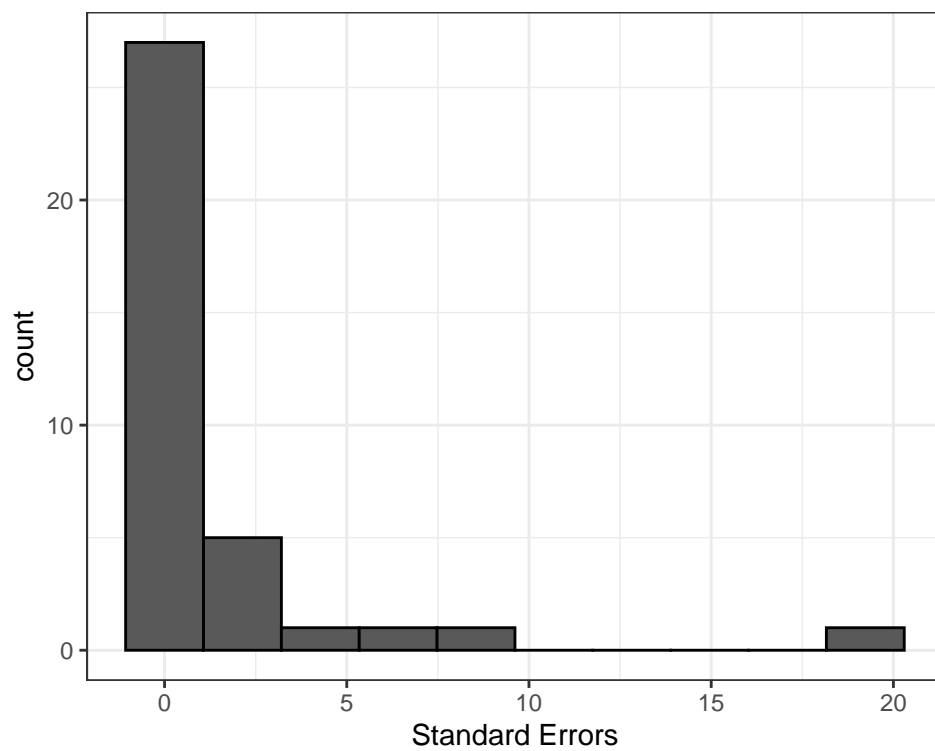
Coefficients:

```
dat %>%
  ggplot(aes(x = coef)) +
  geom_histogram(bins = 15, color = "black") +
  labs(x = "Coefficients") +
  theme_bw()
```



Standard errors:

```
dat %>%  
  ggplot(aes(x = SE)) +  
  geom_histogram(bins = 10, color = "black") +  
  labs(x = "Standard Errors") +  
  theme_bw()
```



D.7 Sign Coefficients

One simple statistic that we can compute to assess the validity of the *law of 1/n* is the frequency of positive and negative estimates in the study sample. Below we plot the frequency for all the papers included in the meta-analytic dataset.

```
dat %>%  
  ggplot(aes(x=as.factor(scoef))) +  
  geom_bar(color = "black") +  
  labs(x = "Coefficient Sign",  
       y = "") +  
  theme_bw()
```

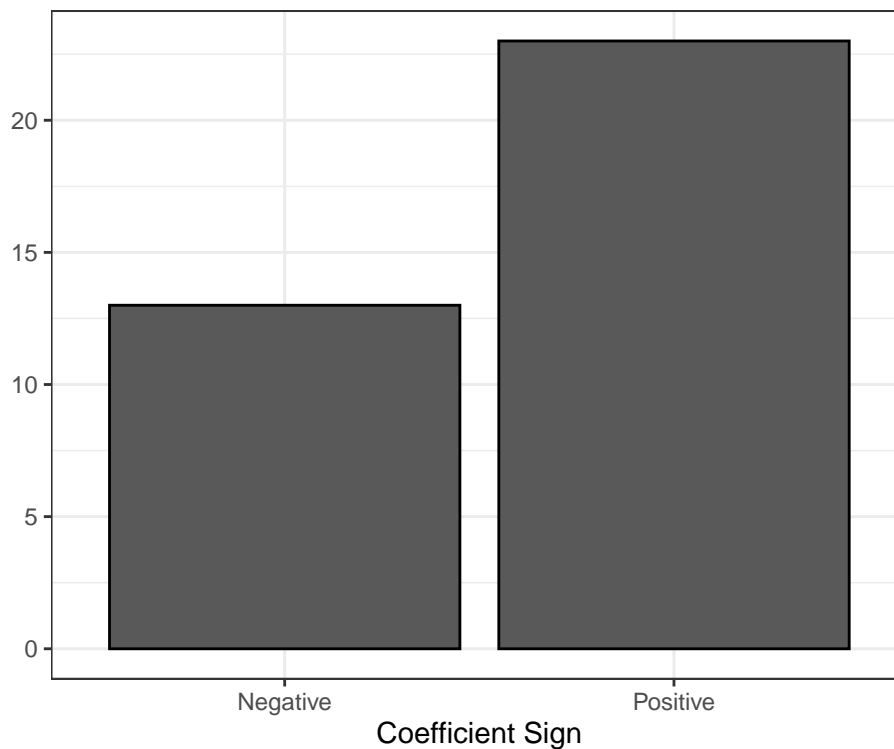


Figure 6: Coefficient Sign

E Descriptive Statistics of Moderators

We chose a set of moderators that frequently appear in the literature and may help us interpret our results. We included them in our meta-regressions alongside an indicator for the type of independent variable used in the original study (n , $\log(n)$, or k). The additional moderators are: 1) electoral system; 2) data aggregation level; 3) estimation method; 4) publication year; 5) paper publication in an academic journal. The table below presents descriptive statistics for these moderators in our selection of articles.

```

fulldat$usemeta2 <- factor(fulldat$usemeta)

levels(fulldat$usemeta2) <- c("Other Coefficients", "Main Coefficients")

aux <- select(fulldat, usemeta2, indepvar2, year, published,
             elecsys2, method) %>%

  rename(`Independent Variables` = indepvar2,
         `Year` = year,
         `Published work` = published,
         `Electoral system` = elecsys2,
         `Estimation method` = method)

descrTable(~.-usemeta2, aux, y = aux$usemeta2,
           show.p.overall = F, show.all = T)

```

```

##
## -----Summary descriptives table by 'usemeta2'-----
##
## -----
##           [ALL]   Other Coefficients Main Coefficients
##           N=128      N=92           N=36
## -----
## Independent Variables:
##      K           38 (29.7%)    29 (31.5%)    9 (25.0%)
##      N           72 (56.2%)    51 (55.4%)    21 (58.3%)
##      logN        18 (14.1%)    12 (13.0%)    6 (16.7%)
##      Year        2008 (5.98)   2008 (5.75)   2009 (6.50)
## Published work:
##      Yes         104 (81.2%)   74 (80.4%)   30 (83.3%)
##      No          24 (18.8%)    18 (19.6%)    6 (16.7%)
## Electoral system:
##      Majoritarian 58 (45.3%)    40 (43.5%)    18 (50.0%)
##      Non-Majoritarian 70 (54.7%)  52 (56.5%)    18 (50.0%)
## Estimation method:
##      OLS         56 (43.8%)    43 (46.7%)    13 (36.1%)
##      PANEL       57 (44.5%)    40 (43.5%)    17 (47.2%)

```


##	IV	6 (4.69%)	3 (3.26%)	3 (8.33%)
##	RDD	9 (7.03%)	6 (6.52%)	3 (8.33%)
##	-----			

F Binomial Tests for Coefficient Signs

The *law of 1/n* posits that we should expect a positive influence of legislature size on public expenditures. A general test of the theory could investigate whether the papers tend to find a higher frequency of positive coefficients in their estimations. In statistical terms, consider a random variable representing the coefficient sign for the papers. As each sign of the paper is a Bernoulli trial, the aggregate result for all papers follows a Binomial distribution with parameters n equals the number of papers, and p the chance of a positive sign. The *law of 1/n* can be reformulated as the chance of $p > 0.5$, which facilitates the testing of the theory. The null hypothesis for such a test is that:

- H_0 : the proportion of positive and negative signs are indistinguishable ($p = 0.5$).

As we are taking an agnostic approach, we acknowledge that either the *law of 1/n* ($p > 0.5$), or the *reverse law of 1/n* ($p < 0.5$) could be true. In this case, the alternative hypothesis is $p \neq 0.5$. To perform this test, we run binomial tests in R, using the function `binom.test()`.

This test has two advantages. First, it is robust to the design of the paper. This is an important feature as papers analyse different countries, samples, and have distinct characteristics, such as whether they were published or not. All these factors increase the levels of study heterogeneity. The binomial test ignores the design discrepancies and focuses on the overall reported effect. Second, this test has the advantage of being straightforward and easy to interpret. It requires very few assumption and has a direct statistical formulation. The disadvantage is that we can extract more information from the articles with meta-regressions, as we see in the next sections.

For the number of legislators in the lower house (N), the results follow below.

```
aux <- filter(dat, indepvar2 == "N")
aux2 <- binom.test(table(aux$scoef)[2], sum(table(aux$scoef)), p = 0.5)
aux2

##
## Exact binomial test
##
## data:  table(aux$scoef)[2] and sum(table(aux$scoef))
```

```
## number of successes = 10, number of trials =
## 21, p-value = 1
## alternative hypothesis: true probability of success is not equal to 0.5
## 95 percent confidence interval:
## 0.2571306 0.7021932
## sample estimates:
## probability of success
## 0.4761905
```

Under the null hypothesis of $p = 0.5$, we find that 10 studies, out of 21, had positive sign. The chance of a distribution with $p = 0.5$ generate this sample is equal to p-value = 1. Therefore, we reject the hypothesis that $p \neq 0.5$.

For the log of the number of legislators in the lower house ($\log(N)$), the results follow below.

```
aux <- filter(dat, indepvar2 == "logN")
aux2 <- binom.test(table(aux$scoef)[2], sum(table(aux$scoef)), p = 0.5)
aux2
```

```
##
## Exact binomial test
##
## data: table(aux$scoef)[2] and sum(table(aux$scoef))
## number of successes = 5, number of trials =
## 6, p-value = 0.2188
## alternative hypothesis: true probability of success is not equal to 0.5
## 95 percent confidence interval:
## 0.3587654 0.9957893
## sample estimates:
## probability of success
## 0.8333333
```

Under the null hypothesis of $p = 0.5$, we find that 5 studies, out of 6, had positive sign. The chance of a distribution with $p = 0.5$ generate this sample is equal to p-value = 0.219. Therefore, we reject the hypothesis that $p \neq 0.5$.

Finally, for the number of legislators in the upper house (K), the results follow below.

```

aux <- filter(dat, indepvar2=='K')

aux2 <- binom.test(table(aux$scoef)[2], sum(table(aux$scoef)), p=0.5)

aux2

##

## Exact binomial test

##

## data: table(aux$scoef)[2] and sum(table(aux$scoef))

## number of successes = 8, number of trials =

## 9, p-value = 0.03906

## alternative hypothesis: true probability of success is not equal to 0.5

## 95 percent confidence interval:

## 0.5175035 0.9971909

## sample estimates:

## probability of success

## 0.8888889

```

Under the null hypothesis of $p = 0.5$, we find that 8 studies, out of 9, had positive sign. The chance of a distribution with $p = 0.5$ generate this sample is equal to $p\text{-value} = 0.039$. Therefore, we accept the hypothesis that $p \neq 0.5$. This is the only test that presents evidence of an association between the legislature size and expenditure.

G Meta-Analysis

G.1 Estimation Method

In general terms, there are two main ways to conduct a meta-analysis, either using fixed effects or random effects models. The fixed effects model assumes that there is one true effect in reality, and that all estimates are an attempt to uncover this true, unknown effect. The random effects model, on the other hand, assumes that there is a distribution of true effects, one which varies based on sample and tests characteristics.

In this paper, we use the random effects model. We do so because the empirical papers testing the *law of 1/n* are very diverse. We tried to capture some of this diversity by considering the main dependent and independent variables separately, but they have at least three other important sources of dispersion:

1. **Subjects:** Counties, Municipalities, States, Provinces, Countries.
2. **Electoral systems:** Majoritarian, PR, Mixed.

3. **Modelling strategies:** Panel data, Standard OLS, IV, RDD.

These sources of heterogeneity have two implications. First, they make our estimates very disperse. All but one of our heterogeneity tests are statistically significant. When the sample sizes were large enough, we removed more heterogeneous studies, but we still had considerable dispersion in our estimates. Second, the amount of heterogeneity makes fixed effects estimates unrealistic and biased. Thus, we opt for random effects model.

Let each study having an effect of T_i . In a random effects model, we can decompose this effect in two components, the true effect that the study with the same specifications as i comes from, θ_i , and a within-study error ε_i :

$$T_i = \theta_i + \varepsilon_i$$

And the random effects model assumes that the θ_i varies from study to study, having a true parameter μ , plus a between-study error, ξ_i :

$$T_i = \mu + \xi_i + \varepsilon_i$$

And the random effects model estimates the parameter μ , under the challenge of estimating both the within-and-between-study sampling errors.

In all empirical estimates, we use the package `meta`, and the package `dmeter`, described in [Doing Meta-Analysis with R](#). To empirically implement the random effects model, we need to choose a method to estimate the true effect size variance, τ^2 , which in our formulation, represents the variance of ξ_i . We selected the **Restricted Maximum Likelihood Estimator**, as the literature regards it the most precise when analysing continuous measures, such as we have in our data.

We combined the three independent variables (N , $\log(N)$, and K) with our dependent variables of interest (Expenditure Per Capita, Log of Expenditure Per Capita, Expenditure as a Percentage of the GDP). This formed a 3×3 table, and in the following pages we present the results for each of these combinations.

G.2 Lower House Size and Expenditure per Capita

```
# Pooling effects analysis -- ExpPC x N  
  
aux <- dat %>%  
  filter(indepvar2 == 'N',  
         depvar2 == 'ExpPC')
```

```

mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
               comb.fixed = FALSE,
               comb.random = TRUE,
               method.tau = "REML",
               hakn = TRUE,
               prediction=TRUE,
               sm="SMD")

```

```
mod
```

##	SMD
## Crowley (2019)	-0.3510
## Lee and Park (2018)	-0.8510
## Lee (2016)	0.0164
## Kessler (2014)	0.1740
## Bjedov et al. (2014)	-0.0030
## Baskaran (2013)	0.9740
## Erler (2007)	3.9300
## Chen and Malhotra (2007)	-2.0400
## Fiorino and Ricciuti (2007)	0.2130
## Primo (2006)	-0.8200
## Matsusaka (2005)	-0.9600
## Schaltegger and Feld (2009)	0.0010
##	95%-CI
## Crowley (2019)	[-1.8112; 1.1092]
## Lee and Park (2018)	[-3.5851; 1.8831]
## Lee (2016)	[-2.5570; 2.5898]
## Kessler (2014)	[0.0074; 0.3406]
## Bjedov et al. (2014)	[-0.0226; 0.0166]
## Baskaran (2013)	[-0.1212; 2.0692]
## Erler (2007)	[1.6172; 6.2428]
## Chen and Malhotra (2007)	[-4.6468; 0.5668]
## Fiorino and Ricciuti (2007)	[0.1777; 0.2483]

```

## Primo (2006) [-1.1924; -0.4476]
## Matsusaka (2005) [-1.3128; -0.6072]
## Schaltegger and Feld (2009) [-0.0010; 0.0030]
## %W(random)
## Crowley (2019) 5.3
## Lee and Park (2018) 2.1
## Lee (2016) 2.4
## Kessler (2014) 13.1
## Bjedov et al. (2014) 13.4
## Baskaran (2013) 7.3
## Erler (2007) 2.8
## Chen and Malhotra (2007) 2.3
## Fiorino and Ricciuti (2007) 13.4
## Primo (2006) 12.2
## Matsusaka (2005) 12.3
## Schaltegger and Feld (2009) 13.4
##
## Number of studies combined: k = 12
##
## SMD 95%-CI
## Random effects model -0.0699 [-0.6712; 0.5314]
## Prediction interval [-1.5540; 1.4142]
## t p-value
## Random effects model -0.26 0.8028
## Prediction interval
##
## Quantifying heterogeneity:
## tau^2 = 0.3690 [0.1794; 4.7570]; tau = 0.6075 [0.4236; 2.1810];
## I^2 = 94.7% [92.3%; 96.3%]; H = 4.34 [3.61; 5.21]
##
## Test of heterogeneity:
## Q d.f. p-value
## 206.92 11 < 0.0001

```

```
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model
```

And the forest plot:

```
build_forest(mod, NULL)
```

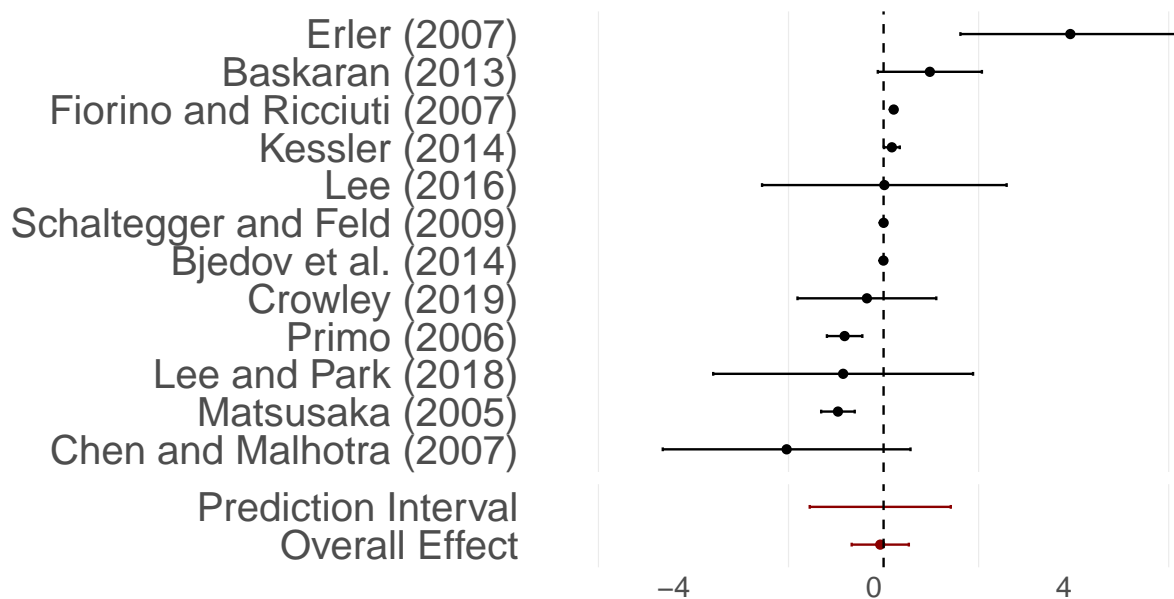


Figure 7: Effect of lower houses size (N) on Per Capita Expenditure (ExpPC)

Highlights:

1. The results are highly heterogeneous: $I^2 = 94.68$.
2. The estimated SMD in the random effects model is $g = -0.07$ ($SE = 0.273$).
3. The prediction interval ranges from -1.55 to 1.41. Therefore, it encompasses zero.

G.2.1 Electoral System Subgroup Analysis

The *law of 1/n* was created for majoritarian systems. In the theoretical section below, we explain why the argument have potential issues when applied to non-majoritarian electoral systems. We estimated a subgroup analysis using a binary electoral system.

```

mod <- update(mod, byvar = aux$selecsys2, print.byvar = F)

mod2 <- tibble(
  TE = mod$TE,
  seTE = mod$seTE,
  studlab = mod$studlab,
  lower = mod$lower,
  upper = mod$upper,
  group = "A") %>%
  bind_rows(.,
    aux = tibble(
      TE = c(mod$TE.random, NA),
      seTE = c(mod$seTE.random, NA),
      studlab = c("Overall Effect", "Prediction Interval"),
      lower = c(mod$lower.random, mod$lower.predict),
      upper = c(mod$upper.random, mod$upper.predict),
      group = "B")) %>%
  group_by(studlab) %>%
  mutate(studlab2 = paste0(studlab, "_", 1:n())) %>%
  ungroup()

f1b <- mod2 %>%
  ggplot(aes(y = reorder(studlab2, TE), x = TE, xmin = lower, xmax = upper)) +
  # Studies coefs
  geom_point(aes(color = group)) +
  # Error Bars
  geom_errorbarh(aes(color = group), height = 0.1) +
  # Colors
  scale_color_manual(values = c("#000000", "#8b0000")) +
  # X-axis limit
  scale_x_continuous(limits=c(1.2*(min(mod2$lower)), 1.2*(max(mod2$upper)))) +
  # Y-axis names

```



```

scale_y_discrete(labels = function(x) str_replace(x, "_[0-9]*$", "")) +

# Vertical dashed line

geom_vline(xintercept=0, color="#000000", linetype="dashed") +

# Labels

labs(x = "",

      y = "") +

# Facet - Separating Studies from Overall Effect

facet_grid(group~., scales = "free", space = "free") +

# Theme

theme_minimal() %+replace%

theme(strip.text.y = element_blank(),

      legend.position = "none",

      axis.text.y = element_text(size = 15, hjust = 1.1),

      axis.text.x = element_text(size = 15, hjust = 1.1))

```

f1b

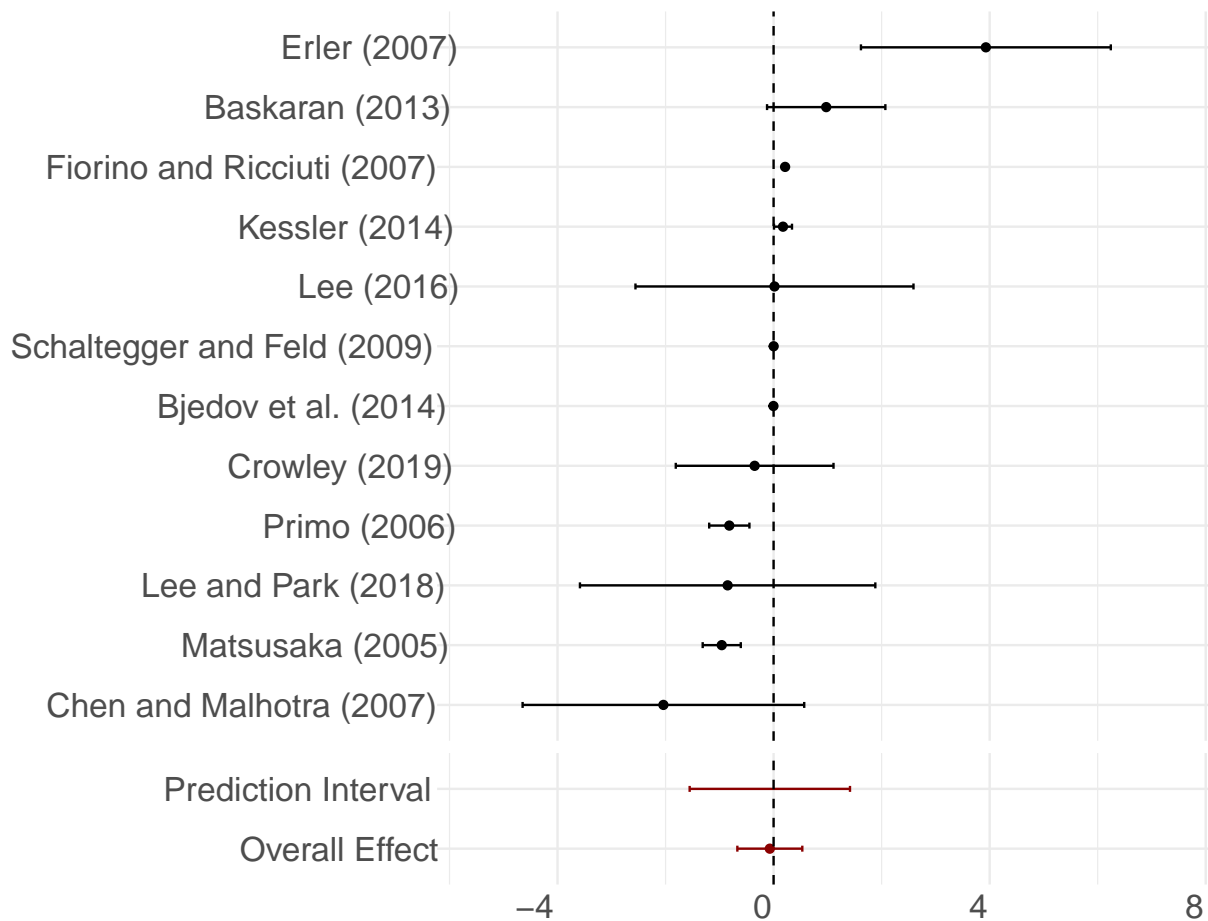


Figure 8: Subgroup Analysis of (N) x (ExpPC), controlling by electoral system

Therefore, we can see that the hypothesis that majoritarian systems produce systematic positive effects was disproved. The majoritarian systems in the sample had a random effects model estimate of -0.25, while the random effects model in the non-majoritarian subgroup fitted a value of 0.08. Both are non-significant, but they reassure us that the absence of effect is not caused by pooling multiple types of electoral systems.

G.3 Log Lower House Size and Expenditure per Capita

There were no studies that had per capita expenditure in the dependent variable and log of lower house size in the treatment variable.

G.4 Upper House Size and Expenditure per Capita

Now, we look into the upper house size (K). In this model, we investigate the effect of upper house size on expenditure per capita (ExpPC).

```
# Pooling effects analysis -- ExpPC x K
```

```
aux <- dat %>%
```

```
  filter(indepvar2 == 'K',
         depvar2 == 'ExpPC')
```

```
mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
               comb.fixed = FALSE,
               comb.random = TRUE,
               method.tau = "REML",
               hakn = TRUE,
               prediction=TRUE,
               sm="SMD")

mod
```

##	SMD
## Crowley (2019)	8.2100
## Lee and Park (2018)	19.7400
## Lee (2016)	38.4400
## Bradbury and Stephenson (2009)	0.6240
## Chen and Malhotra (2007)	26.0900
## Primo (2006)	0.9700
##	95%-CI
## Crowley (2019)	[0.2702; 16.1498]
## Lee and Park (2018)	[3.2645; 36.2155]
## Lee (2016)	[0.7499; 76.1301]
## Bradbury and Stephenson (2009)	[0.2295; 1.0185]
## Chen and Malhotra (2007)	[11.4883; 40.6917]
## Primo (2006)	[-0.4804; 2.4204]
##	%W(random)
## Crowley (2019)	20.0
## Lee and Park (2018)	13.8
## Lee (2016)	5.1
## Bradbury and Stephenson (2009)	23.1

```

## Chen and Malhotra (2007)          15.1
## Primo (2006)                     23.0
##
## Number of studies combined: k = 6
##
##                               SMD          95%-CI
## Random effects model 10.6134 [ -2.6210; 23.8479]
## Prediction interval          [-21.1303; 42.3571]
##                               t p-value
## Random effects model 2.06  0.0943
## Prediction interval
##
## Quantifying heterogeneity:
## tau^2 = 104.2124 [20.3551; >1042.1236]; tau = 10.2084 [4.5117; >32.2819];
## I^2 = 79.4% [55.1%; 90.6%]; H = 2.20 [1.49; 3.26]
##
## Test of heterogeneity:
##      Q d.f. p-value
## 24.31    5 0.0002
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model

```

And the forest plot:

```
build_forest(mod, NULL)
```

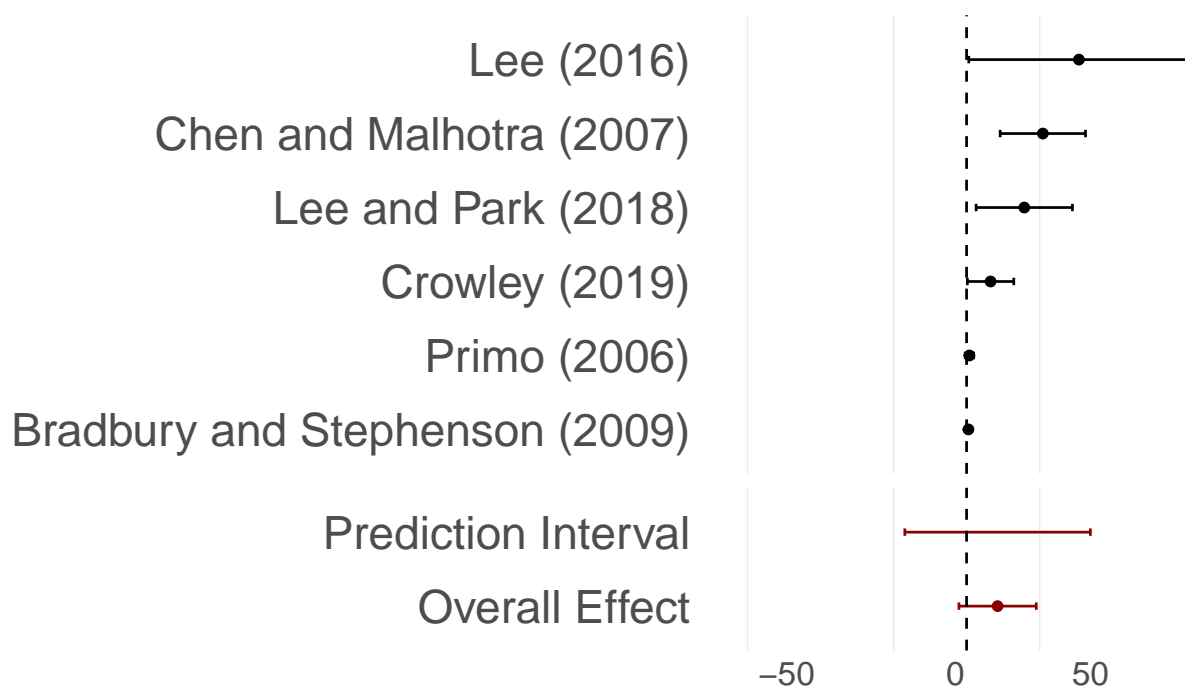


Figure 9: Effect of upper house size (K) on the per capita government expenditure (ExpPC)

Highlights:

1. The results are highly heterogeneous: $I^2 = 79.43$.
2. The estimated SMD in the random effects model is $g = 10.61$ ($SE = 5.148$).
3. The prediction interval ranges from -21.13 to 42.36. Therefore, it encompasses zero.

G.5 Lower House Size and Log Expenditure Per Capita

This model estimates the Log of Per Capita Expenditure as the dependent variable, and the number of lower house legislators as the treatment variable.

```
# Pooling effects analysis -- logExpPC x N
aux <- dat %>%
  filter(indepvar2 == 'N',
         depvar2 == 'logExpPC')

mod <- metagen(
  coef, SE, data=aux,
  studlab=paste(authoryear),
  comb.fixed = FALSE,
  comb.random = TRUE,
```

```

method.tau = "REML",
hakn = TRUE,
prediction = TRUE,
sm="SMD"
)

```

```
mod
```

```

##                               SMD
## Lewis (2019)                  -0.1740
## H hmann (2017)                -0.0300
## Drew and Dollery (2017)      0.0770
## Pettersson-Lidbom (2012)    -0.1590
##                               95%-CI
## Lewis (2019)                  [-0.2450; -0.1030]
## H hmann (2017)                [-0.0496; -0.0104]
## Drew and Dollery (2017)      [ 0.0221;  0.1319]
## Pettersson-Lidbom (2012)    [-0.2394; -0.0786]
##                               %W(random)
## Lewis (2019)                  24.3
## H hmann (2017)                26.6
## Drew and Dollery (2017)      25.3
## Pettersson-Lidbom (2012)     23.7
##
## Number of studies combined: k = 4
##
##                               SMD          95%-CI
## Random effects model -0.0686 [-0.2560; 0.1188]
## Prediction interval          [-0.6179; 0.4807]
##                               t p-value
## Random effects model -1.17  0.3282
## Prediction interval
##
## Quantifying heterogeneity:

```

```
## tau^2 = 0.0128 [0.0034; 0.1933]; tau = 0.1133 [0.0584; 0.4396];
## I^2 = 92.5% [84.1%; 96.5%]; H = 3.66 [2.51; 5.34]
##
## Test of heterogeneity:
##      Q d.f.  p-value
## 40.11    3 < 0.0001
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model
```

And the forest plot:

```
build_forest(mod, NULL)
```

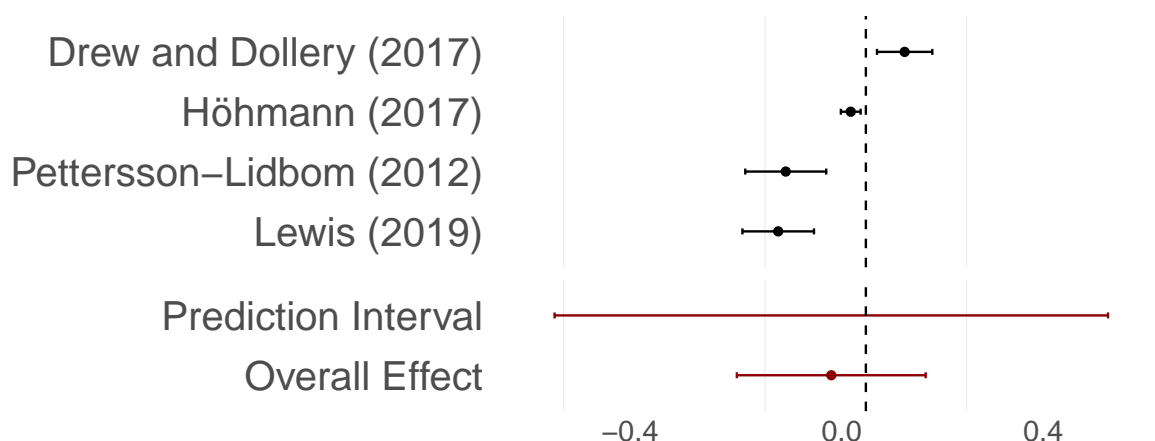


Figure 10: Effect of lower houses size (N) on log of per capita expenditure (logExpPC)

Highlights:

1. The results are highly heterogeneous: $I^2 = 92.52$.
2. The estimated SMD in the random effects model is $g = -0.07$ ($SE = 0.059$).
3. The prediction interval ranges from -0.62 to 0.48. Therefore, it encompasses zero.

G.6 Log of Lower House Size and Log of Expenditure Per Capita

In this specification, we study the log of per capita expenditure (logExpPC) as a function of the log of lower house size (logN).

```
# Pooling effects analysis -- logExpPC x logN
```

```
aux <- dat %>%
```

```
  filter(indepvar2 == 'logN',
         depvar2 == 'logExpPC')
```

```
mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
               comb.fixed = FALSE,
               comb.random = TRUE,
               method.tau = "REML",
               hakn = TRUE,
               prediction=TRUE,
               sm="SMD")

mod
```

```
##              SMD              95%-CI
## MacDonald (2008) 0.1360 [0.0447; 0.2273]
## Baqir (2002)     0.1127 [0.0396; 0.1858]
## Baqir (1999)     0.3020 [0.2269; 0.3771]
##              %W(random)
## MacDonald (2008)      31.9
## Baqir (2002)          34.2
## Baqir (1999)          33.9
##
## Number of studies combined: k = 3
##
##              SMD              95%-CI      t
## Random effects model 0.1844 [-0.0738; 0.4425] 3.07
## Prediction interval      [-1.2580; 1.6267]
##              p-value
## Random effects model  0.0916
## Prediction interval
##
## Quantifying heterogeneity:
```



```
## tau^2 = 0.0093 [0.0014; 0.4193]; tau = 0.0964 [0.0372; 0.6476];
## I^2 = 85.9% [59.0%; 95.2%]; H = 2.66 [1.56; 4.54]
##
## Test of heterogeneity:
##      Q d.f. p-value
## 14.18    2  0.0008
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model
```

The forest plot is available below:

```
build_forest(mod, NULL)
```

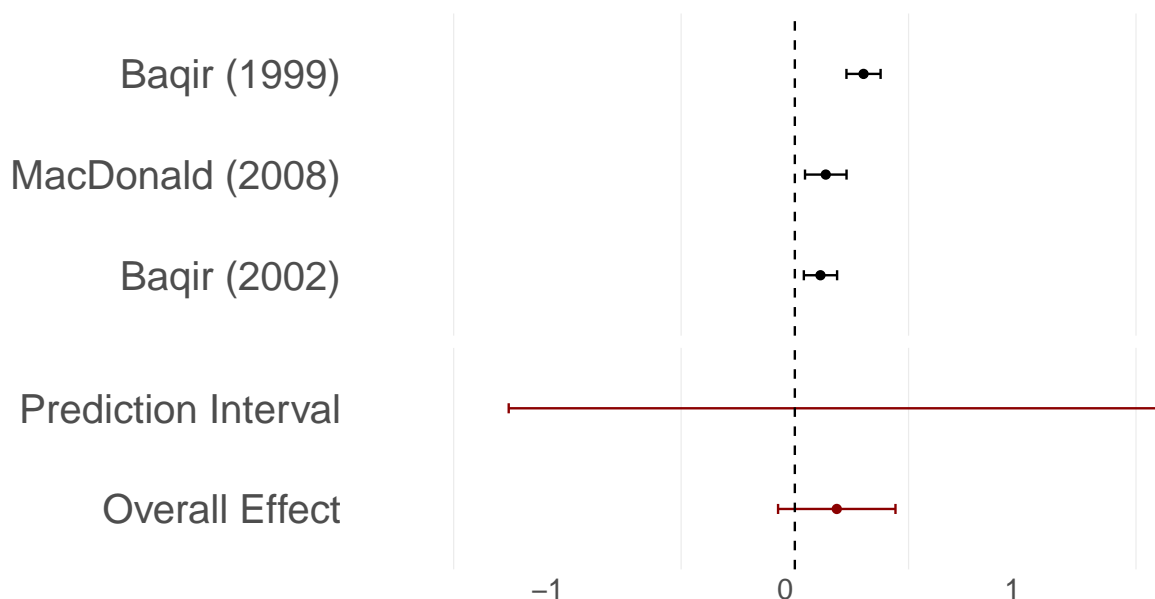


Figure 11: Effect of log lower houses size (logN) on the log of per capita government expenditure (logExpPC)

Highlights:

1. The results are highly heterogeneous: $I^2 = 85.9$.
2. The estimated SMD in the random effects model is $g = 0.18$ ($SE = 0.06$). **This model is significant at the 10% confidence level.**
3. The prediction interval ranges from -1.26 to 1.63. Therefore, it encompasses zero.

G.7 Log of Upper House Size and Log of Expenditure Per Capita

No studies related the log of per capita expenditure with the size of upper house (K).

G.8 Lower House Size and Expenditure as Percentage of GDP

This model fits the random effects for the percentage of GDP as public expenditure as the main outcome, and the size of lower house as the main treatment variable.

```
# Pooling effects analysis -- PCTGDP x N
```

```
aux <- dat %>%
```

```
  filter(indepvar2 == 'N',
         depvar2 == 'PCTGDP')
```

```
mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
               comb.fixed = FALSE,
               comb.random = TRUE,
               method.tau = "REML",
               hakn = TRUE,
               prediction=TRUE,
               sm="SMD")
```

```
mod
```

##	SMD
## Bjedov et al. (2014)	-0.0040
## Maldonado (2013)	-0.0609
## Mukherjee (2003)	0.0030
## Bradbury and Crain (2001)	0.0036
## Ricciuti (2004)	0.0140
##	95%-CI
## Bjedov et al. (2014)	[-0.0432; 0.0352]
## Maldonado (2013)	[-0.0838; -0.0380]
## Mukherjee (2003)	[0.0010; 0.0050]
## Bradbury and Crain (2001)	[0.0008; 0.0065]
## Ricciuti (2004)	[-0.0095; 0.0375]

```

##                                %W(random)

## Bjedov et al. (2014)           15.1
## Maldonado (2013)              19.5
## Mukherjee (2003)              23.0
## Bradbury and Crain (2001)     23.0
## Ricciuti (2004)              19.4

##

## Number of studies combined: k = 5

##

##                                SMD            95%-CI
## Random effects model -0.0083 [-0.0450; 0.0285]
## Prediction interval      [-0.1054; 0.0889]

##                                t p-value
## Random effects model -0.62  0.5667
## Prediction interval

##

## Quantifying heterogeneity:
## tau^2 = 0.0008 [0.0002; 0.0072]; tau = 0.0275 [0.0129; 0.0849];
## I^2 = 87.1% [72.2%; 94.0%]; H = 2.78 [1.90; 4.08]

##

## Test of heterogeneity:
##      Q d.f.  p-value
## 30.97    4 < 0.0001

##

## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model

```

Below, you may find the forest plot:

```
build_forest(mod, NULL)
```

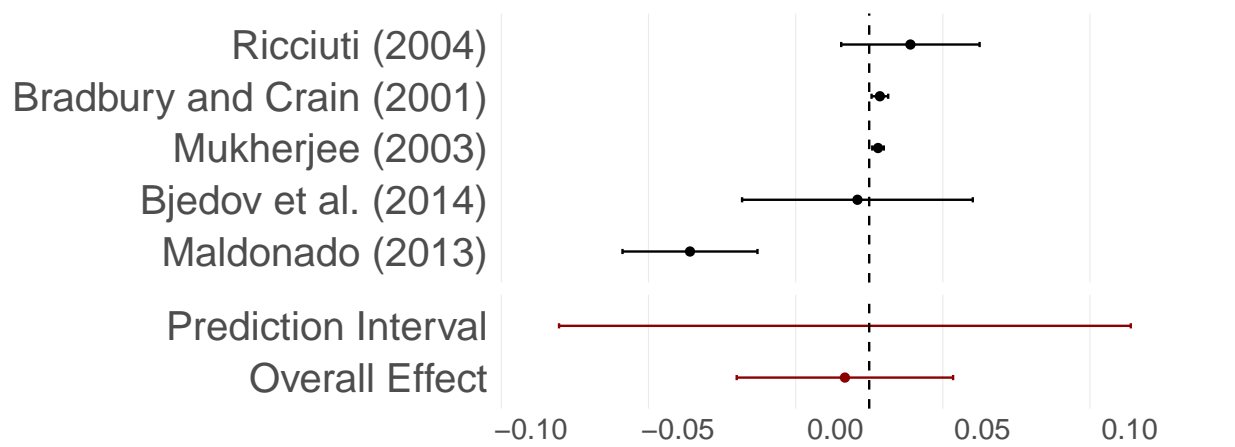


Figure 12: Effect of lower houses size (N) on percentage of public expenditure GDP (PCTGDP)

Highlights:

1. The results are highly heterogeneous: $I^2 = 87.08$.
2. The estimated SMD in the random effects model is $g = -0.01$ ($SE = 0.013$).
3. The prediction interval ranges from -0.11 to 0.09. Therefore, it encompasses zero.

G.9 Log Lower House Size and Expenditure as Percentage of GDP

This model investigates the percentage of GDP as public expenditure as the dependent variable and the log lower house size (logN) as the treatment variable.

```
# Pooling effects analysis -- PCTGDP x logN

aux <- dat %>%

  filter(indepvar2 == 'logN',
         depvar2 == 'PCTGDP')

mod <- metagen(
  coef, SE, data=aux,
  studlab=paste(authoryear),
  comb.fixed = FALSE,
  comb.random = TRUE,
  method.tau = "REML",
  hakn = TRUE,
  prediction=TRUE,
  sm="SMD"
```

)

mod

```
##              SMD              95%-CI
## Baqir (1999)      2.0660 [ 1.4887; 2.6433]
## Lledo (2003)     -4.6900 [-9.9427; 0.5627]
## Stein et al. (1998) 0.0109 [-0.0171; 0.0389]
##              %W(random)
## Baqir (1999)      40.8
## Lledo (2003)      17.7
## Stein et al. (1998) 41.5
##
## Number of studies combined: k = 3
##
##              SMD              95%-CI
## Random effects model 0.0203 [ -7.1961; 7.2367]
## Prediction interval      [-36.2058; 36.2465]
##              t p-value
## Random effects model 0.01  0.9914
## Prediction interval
##
## Quantifying heterogeneity:
## tau^2 = 5.3156 [0.5756; >100.0000]; tau = 2.3056 [0.7587; >10.0000];
## I^2 = 96.1% [91.8%; 98.2%]; H = 5.08 [3.48; 7.42]
##
## Test of heterogeneity:
##      Q d.f.  p-value
## 51.65    2 < 0.0001
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
```

```
## - Hartung-Knapp adjustment for random effects model
```

The forest plot follows below:

```
build_forest(mod, NULL)
```

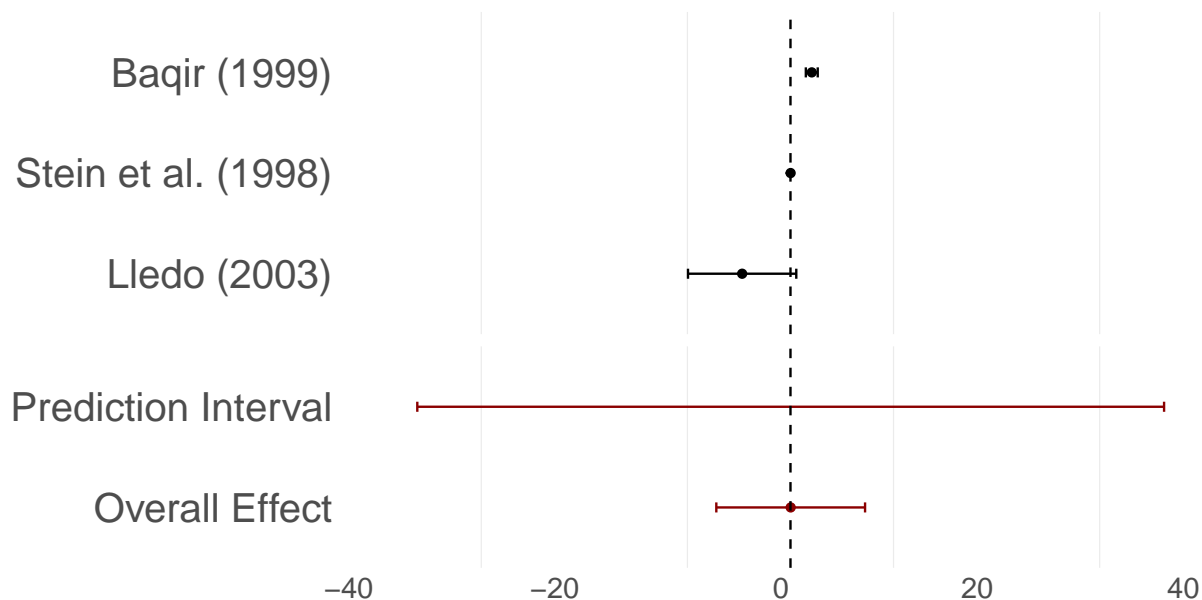


Figure 13: Effect of log lower houses size (logN) on the GDP share of public expenditure (PCTGDP)

Highlights:

1. The results are highly heterogeneous: $I^2 = 96.13$.
2. The estimated SMD in the random effects model is $g = 0.02$ ($SE = 1.677$).
3. The prediction interval ranges from -36.21 to 36.25. Therefore, it encompasses zero.

G.10 Upper House Size and Expenditure as Percentage of GDP

This model looks into the effect of upper house size (K) on the public expenditure share of the GDP (PCTGDP).

```
# Pooling effects analysis -- PCTGDP x K

aux <- dat %>%
  filter(indepvar2 == 'K',
         depvar2 == 'PCTGDP')

mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
```

```

comb.fixed = FALSE,
comb.random = TRUE,
method.tau = "REML",
hakn = TRUE,
prediction=TRUE,
sm="SMD")

```

mod

```

##                               SMD
## Maldonado (2012)             -0.0400
## Bradbury and Crain (2001)    0.0126
## Ricciuti (2004)             0.0160
##                               95%-CI
## Maldonado (2012)            [-0.0659; -0.0141]
## Bradbury and Crain (2001) [ 0.0010;  0.0243]
## Ricciuti (2004)            [-0.0075;  0.0395]
##                               %W(random)
## Maldonado (2012)            31.3
## Bradbury and Crain (2001)   36.4
## Ricciuti (2004)            32.3
##
## Number of studies combined: k = 3
##
##                               SMD          95%-CI
## Random effects model -0.0027 [-0.0793; 0.0738]
## Prediction interval          [-0.4284; 0.4229]
##                               t p-value
## Random effects model -0.15  0.8915
## Prediction interval
##
## Quantifying heterogeneity:
## tau^2 = 0.0008 [0.0001; 0.0388]; tau = 0.0284 [0.0101; 0.1970];
## I^2 = 85.8% [58.6%; 95.1%]; H = 2.65 [1.55; 4.53]
##

```

```
## Test of heterogeneity:

##      Q d.f. p-value
## 14.07   2  0.0009

##

## Details on meta-analytical method:

## - Inverse variance method

## - Restricted maximum-likelihood estimator for tau^2

## - Q-profile method for confidence interval of tau^2 and tau

## - Hartung-Knapp adjustment for random effects model
```

The forest plot follows below:

```
build_forest(mod, NULL)
```

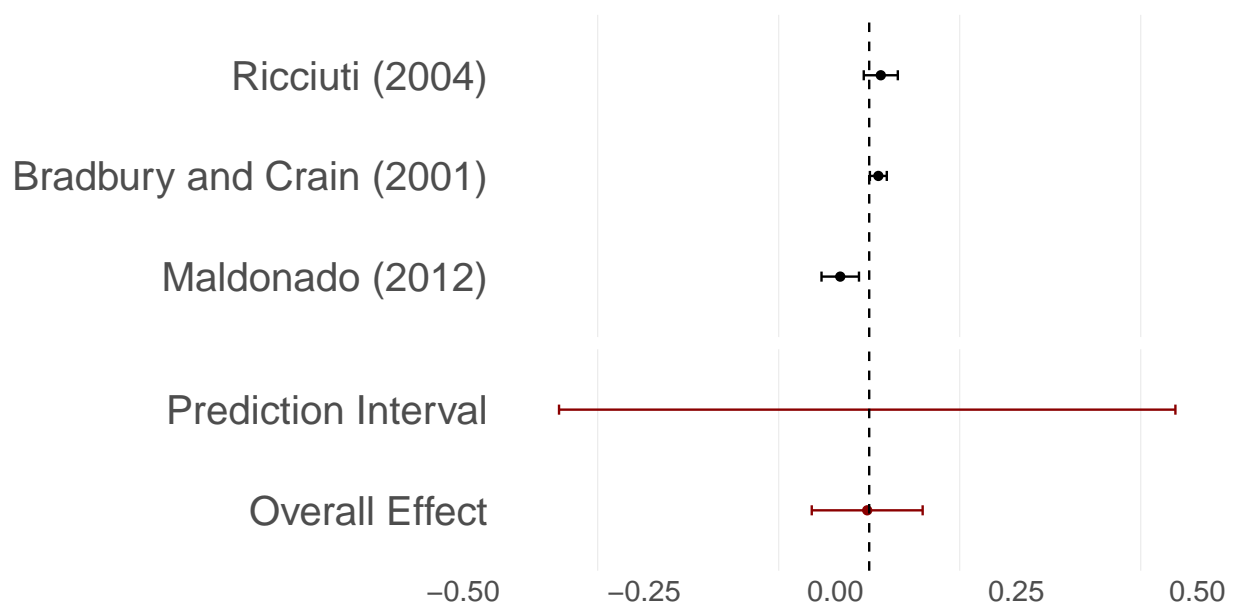


Figure 14: Effect of upper house size (K) on the public expenditure share of the GDP (PCTGDP)

Highlights:

1. The results are highly heterogeneous: $I^2 = 85.79$.
2. The estimated SMD in the random effects model is $g = 0$ ($SE = 0.018$).
3. The prediction interval ranges from -0.43 to 0.42. Therefore, it encompasses zero.

H Meta-Analysis (All Coefficients)

H.1 Lower House Size and Expenditure Per Capita

```
# Pooling effects analysis -- ExpPC x N
```

```
aux <- fulldat %>%
```

```
  filter(indepvar2 == 'N',  
         depvar2 == 'ExpPC')
```

```
mod <- metagen(coef, SE, data=aux,  
               studlab=paste(authoryear),  
               comb.fixed = FALSE,  
               comb.random = TRUE,  
               method.tau = "REML",  
               hakn = TRUE,  
               prediction=TRUE,  
               sm="SMD")
```

```
mod
```

##	SMD
## Crowley (2019)	-0.3510
## Crowley (2019)	5.9750
## Crowley (2019)	7.6580
## Lee and Park (2018)	-0.8510
## Lee and Park (2018)	-1.6890
## Lee and Park (2018)	7.6320
## Lee (2016)	0.0164
## Kessler (2014)	0.1740
## Kessler (2014)	0.2230
## Kessler (2014)	0.2150
## Kessler (2014)	0.1580
## Bjedov et al. (2014)	-0.0030
## Bjedov et al. (2014)	-0.0060
## Baskaran (2013)	0.9740
## Erler (2007)	3.9300

## Chen and Malhotra (2007)	-2.0400
## Chen and Malhotra (2007)	-1.4000
## Fiorino and Ricciuti (2007)	0.2130
## Fiorino and Ricciuti (2007)	0.2290
## Fiorino and Ricciuti (2007)	0.4550
## Fiorino and Ricciuti (2007)	0.4110
## Fiorino and Ricciuti (2007)	0.2260
## Fiorino and Ricciuti (2007)	0.2130
## Fiorino and Ricciuti (2007)	0.1850
## Fiorino and Ricciuti (2007)	0.2350
## Fiorino and Ricciuti (2007)	0.3740
## Fiorino and Ricciuti (2007)	0.8110
## Fiorino and Ricciuti (2007)	0.7950
## Fiorino and Ricciuti (2007)	0.8490
## Primo (2006)	-0.8200
## Primo (2006)	-1.7000
## Primo (2006)	-2.3700
## Primo (2006)	-2.0300
## Matsusaka (2005)	-0.9600
## Schaltegger and Feld (2009)	0.0010
## Schaltegger and Feld (2009)	-0.0010
##	95%-CI
## Crowley (2019)	[-1.8112; 1.1092]
## Crowley (2019)	[0.7889; 11.1611]
## Crowley (2019)	[-0.0290; 15.3450]
## Lee and Park (2018)	[-3.5851; 1.8831]
## Lee and Park (2018)	[-3.0551; -0.3229]
## Lee and Park (2018)	[3.1064; 12.1576]
## Lee (2016)	[-2.5570; 2.5898]
## Kessler (2014)	[0.0074; 0.3406]
## Kessler (2014)	[0.1211; 0.3249]
## Kessler (2014)	[0.0954; 0.3346]
## Kessler (2014)	[0.0522; 0.2638]

## Bjedov et al. (2014)	[-0.0226; 0.0166]
## Bjedov et al. (2014)	[-0.0256; 0.0136]
## Baskaran (2013)	[-0.1212; 2.0692]
## Erler (2007)	[1.6172; 6.2428]
## Chen and Malhotra (2007)	[-4.6468; 0.5668]
## Chen and Malhotra (2007)	[-2.6544; -0.1456]
## Fiorino and Ricciuti (2007)	[0.1777; 0.2483]
## Fiorino and Ricciuti (2007)	[0.1565; 0.3015]
## Fiorino and Ricciuti (2007)	[0.3805; 0.5295]
## Fiorino and Ricciuti (2007)	[0.3150; 0.5070]
## Fiorino and Ricciuti (2007)	[0.1221; 0.3299]
## Fiorino and Ricciuti (2007)	[-0.4083; 0.8343]
## Fiorino and Ricciuti (2007)	[-0.4128; 0.7828]
## Fiorino and Ricciuti (2007)	[-0.4235; 0.8935]
## Fiorino and Ricciuti (2007)	[0.2486; 0.4994]
## Fiorino and Ricciuti (2007)	[0.4562; 1.1658]
## Fiorino and Ricciuti (2007)	[0.4500; 1.1400]
## Fiorino and Ricciuti (2007)	[0.3825; 1.3155]
## Primo (2006)	[-1.1924; -0.4476]
## Primo (2006)	[-2.3076; -1.0924]
## Primo (2006)	[-3.0952; -1.6448]
## Primo (2006)	[-2.7552; -1.3048]
## Matsusaka (2005)	[-1.3128; -0.6072]
## Schaltegger and Feld (2009)	[-0.0010; 0.0030]
## Schaltegger and Feld (2009)	[-0.0030; 0.0010]
##	%W(random)
## Crowley (2019)	2.0
## Crowley (2019)	0.3
## Crowley (2019)	0.2
## Lee and Park (2018)	0.9
## Lee and Park (2018)	2.1
## Lee and Park (2018)	0.4
## Lee (2016)	1.0

## Kessler (2014)	3.6
## Kessler (2014)	3.6
## Kessler (2014)	3.6
## Kessler (2014)	3.6
## Bjedov et al. (2014)	3.6
## Bjedov et al. (2014)	3.6
## Baskaran (2013)	2.5
## Erler (2007)	1.2
## Chen and Malhotra (2007)	1.0
## Chen and Malhotra (2007)	2.3
## Fiorino and Ricciuti (2007)	3.6
## Fiorino and Ricciuti (2007)	3.6
## Fiorino and Ricciuti (2007)	3.6
## Fiorino and Ricciuti (2007)	3.6
## Fiorino and Ricciuti (2007)	3.6
## Fiorino and Ricciuti (2007)	3.1
## Fiorino and Ricciuti (2007)	3.2
## Fiorino and Ricciuti (2007)	3.1
## Fiorino and Ricciuti (2007)	3.6
## Fiorino and Ricciuti (2007)	3.4
## Fiorino and Ricciuti (2007)	3.5
## Fiorino and Ricciuti (2007)	3.3
## Primo (2006)	3.4
## Primo (2006)	3.2
## Primo (2006)	3.0
## Primo (2006)	3.0
## Matsusaka (2005)	3.4
## Schaltegger and Feld (2009)	3.6
## Schaltegger and Feld (2009)	3.6

##

Number of studies combined: k = 36

##

##	SMD	95%-CI
----	-----	--------

```

## Random effects model -0.0169 [-0.4166; 0.3829]
## Prediction interval          [-1.7588; 1.7250]
##                               t p-value
## Random effects model -0.09  0.9322
## Prediction interval
##
## Quantifying heterogeneity:
## tau^2 = 0.6959 [0.7202; 4.3553]; tau = 0.8342 [0.8486; 2.0869];
## I^2 = 95.3% [94.2%; 96.1%]; H = 4.60 [4.16; 5.08]
##
## Test of heterogeneity:
##      Q d.f.  p-value
## 739.53   35 < 0.0001
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model

```

The forest plot:

```
build_forest(mod, NULL)
```

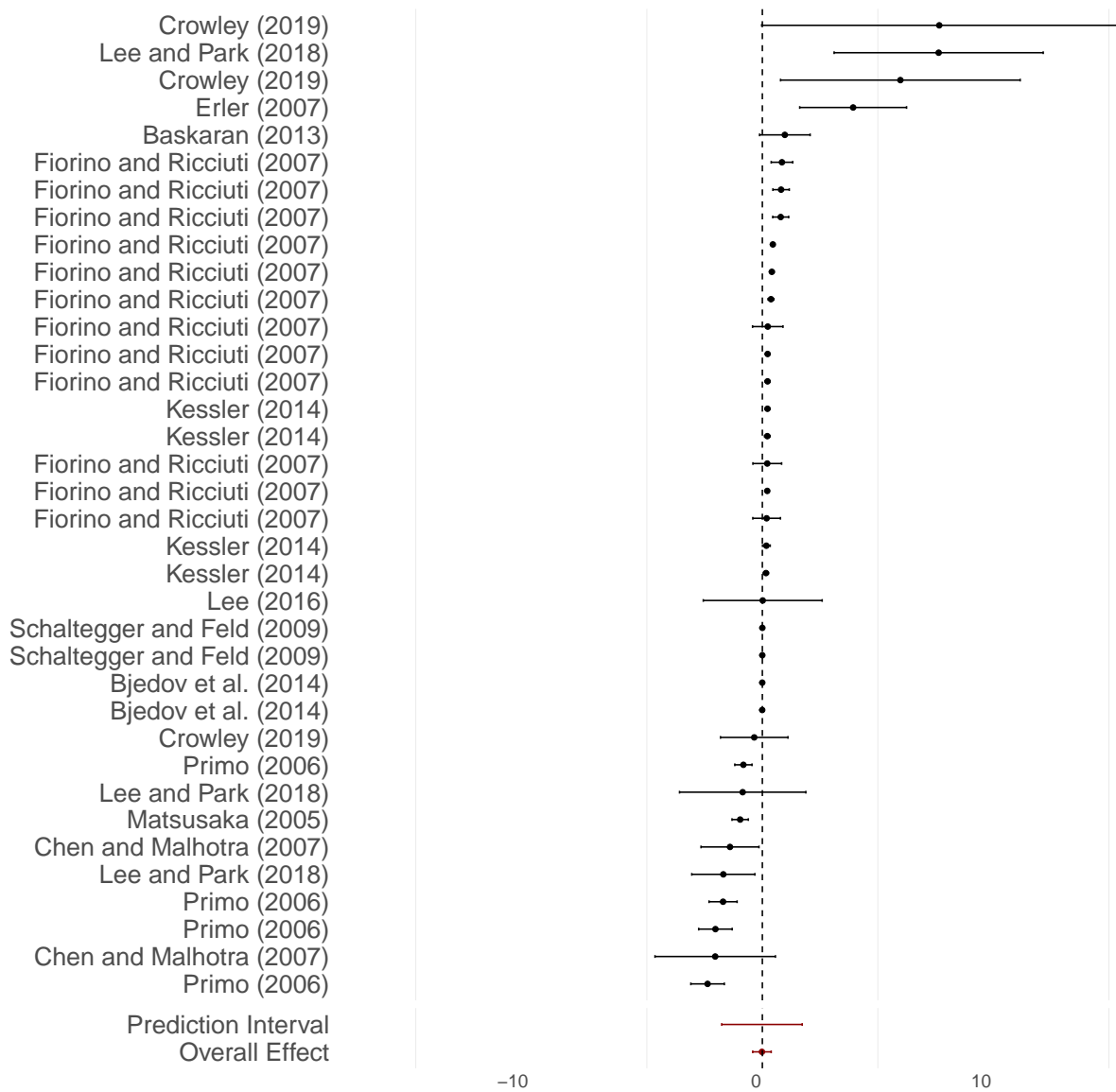


Figure 15: Effect of Lower House Size (N) on Per Capita Expenditure (ExpPC)

Highlights:

1. The results are highly heterogeneous: $I^2 = 95.27$.
2. The estimated SMD in the random effects model is $g = -0.02$ ($SE = 0.197$).
3. The prediction interval ranges from -1.76 to 1.73. Therefore, it encompasses zero.

H.1.1 Electoral System Subgroup Analysis

The *law of 1/n* was created for majoritarian systems. In the theoretical section below, we explain why the argument have potential issues when applied to non-majoritarian electoral systems. We estimated a subgroup analysis using a binary electoral system.

```

mod2 <- tibble(
  TE = mod$TE,
  seTE = mod$seTE,
  studlab = mod$studlab,
  lower = mod$lower,
  upper = mod$upper,
  group = "A") %>%
  bind_rows(.,
    aux = tibble(
      TE = c(mod$TE.random, NA),
      seTE = c(mod$seTE.random, NA),
      studlab = c("Overall Effect", "Prediction Interval"),
      lower = c(mod$lower.random, mod$lower.predict),
      upper = c(mod$upper.random, mod$upper.predict),
      group = "B")) %>%
  group_by(studlab) %>%
  mutate(studlab2 = paste0(studlab, "_", 1:n())) %>%
  ungroup()

f8b <- mod2 %>%

  ggplot(aes(y = reorder(studlab2, TE), x = TE, xmin = lower, xmax = upper)) +
  # Studies coeffs
  geom_point(aes(color = group)) +
  # Error Bars
  geom_errorbarh(aes(color = group), height = 0.1) +
  # Colors
  scale_color_manual(values = c("#000000", "#8b0000")) +
  # X-axis limit
  scale_x_continuous(limits=c(1.2*(min(mod2$lower)), 1.2*(max(mod2$upper)))) +
  # Y-axis names
  scale_y_discrete(labels = function(x) str_replace(x, "_[0-9]*$", "")) +
  # Vertical dashed line
  geom_vline(xintercept=0, color="#000000", linetype="dashed") +

```

```

# Labels

labs(x = "",
     y = "") +

# Facet - Separating Studies from Overall Effect

facet_grid(group~., scales = "free", space = "free") +

# Theme

theme_minimal() %+replace%

theme(strip.text.y = element_blank(),
      legend.position = "none",
      axis.text.y = element_text(size = 13, hjust = 1.1),
      axis.text.x = element_text(size = 15, hjust = 1.1))

```

f8b

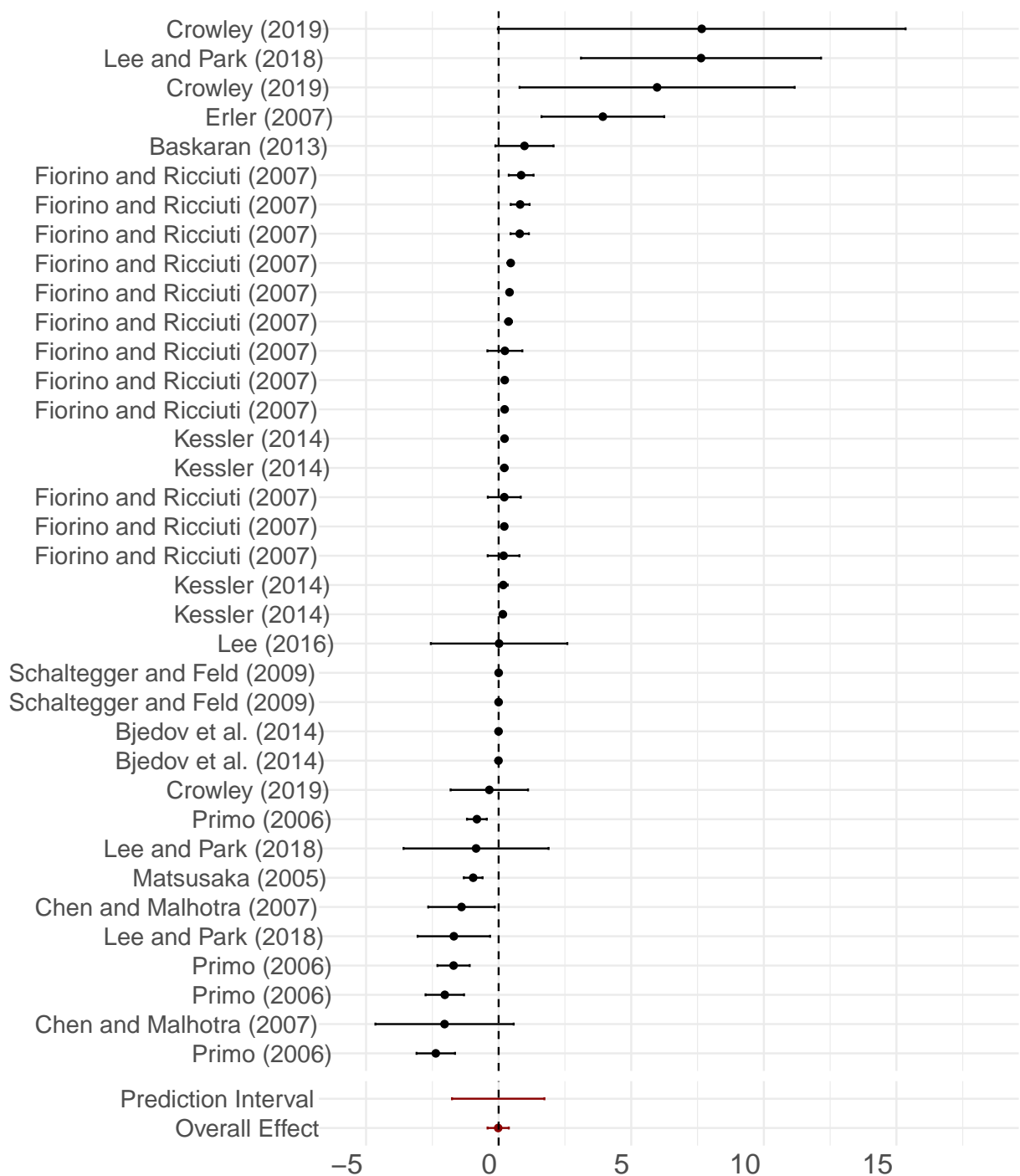


Figure 16: Subgroup Analysis of (N) x (ExpPC), controlling by electoral system

Therefore, we can see that the hypothesis that majoritarian systems produce systematic positive effects was disproved. The majoritarian systems in the sample had a random effects model estimate of -0.25, while the random effects model in the non-majoritarian subgroup fitted a value of 0.08. Both are non-significant, but they reassure us that the absence of effect is not caused by pooling multiple types of electoral systems.

H.2 Log of Lower House Size and Expenditure Per Capita

There were no studies that had per capita expenditure in the dependent variable and log of lower house size in the treatment variable.

H.3 Upper House Size and Expenditure Per Capita

Now we investigate the effect of the upper house size (K) on government spending. In the model below, we evaluate the relationship between upper house size and expenditure per capita (ExpPC).

```
# Pooling effects analysis -- ExpPC x K
```

```
aux <- fulldat %>%
```

```
  filter(indepvar2 == 'K',
         depvar2 == 'ExpPC')
```

```
mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
               comb.fixed = FALSE,
               comb.random = TRUE,
               method.tau = "REML",
               hakn = TRUE,
               prediction=TRUE,
               sm="SMD")
```

```
mod
```

##	SMD
## Crowley (2019)	8.2100
## Crowley (2019)	8.4230
## Crowley (2019)	9.5940
## Lee and Park (2018)	19.7400
## Lee and Park (2018)	10.0600
## Lee and Park (2018)	9.0620
## Lee (2016)	38.4400
## Lee (2016)	37.8500
## Lee (2016)	25.6100
## Lee (2016)	5.9960

## Lee (2016)	25.5600
## Lee (2016)	4.6930
## Bradbury and Stephenson (2009)	0.6240
## Chen and Malhotra (2007)	26.0900
## Chen and Malhotra (2007)	8.3000
## Chen and Malhotra (2007)	5.1400
## Chen and Malhotra (2007)	4.7800
## Chen and Malhotra (2007)	20.3800
## Chen and Malhotra (2007)	4.8700
## Chen and Malhotra (2007)	26.7500
## Primo (2006)	0.9700
## Primo (2006)	5.9000
## Primo (2006)	5.7500
## Primo (2006)	6.9600
##	95%-CI
## Crowley (2019)	[0.2702; 16.1498]
## Crowley (2019)	[-27.1895; 44.0355]
## Crowley (2019)	[2.1383; 17.0497]
## Lee and Park (2018)	[3.2645; 36.2155]
## Lee and Park (2018)	[2.2887; 17.8313]
## Lee and Park (2018)	[-30.8821; 49.0061]
## Lee (2016)	[0.7499; 76.1301]
## Lee (2016)	[3.0214; 72.6786]
## Lee (2016)	[-0.8103; 52.0303]
## Lee (2016)	[-19.6011; 31.5931]
## Lee (2016)	[-0.8799; 51.9999]
## Lee (2016)	[-19.5126; 28.8986]
## Bradbury and Stephenson (2009)	[0.2295; 1.0185]
## Chen and Malhotra (2007)	[11.4883; 40.6917]
## Chen and Malhotra (2007)	[3.6941; 12.9059]
## Chen and Malhotra (2007)	[0.1813; 10.0987]
## Chen and Malhotra (2007)	[-0.9039; 10.4639]
## Chen and Malhotra (2007)	[7.6990; 33.0610]

```

## Chen and Malhotra (2007)      [ 1.2833; 8.4567]
## Chen and Malhotra (2007)      [ 0.8589; 52.6411]
## Primo (2006)                  [ -0.4804; 2.4204]
## Primo (2006)                  [ 2.6857; 9.1143]
## Primo (2006)                  [ 2.3593; 9.1407]
## Primo (2006)                  [ 2.6089; 11.3111]
##                                %W(random)
## Crowley (2019)                 4.8
## Crowley (2019)                 0.4
## Crowley (2019)                 5.1
## Lee and Park (2018)            1.7
## Lee and Park (2018)            4.9
## Lee and Park (2018)            0.3
## Lee (2016)                    0.4
## Lee (2016)                    0.4
## Lee (2016)                    0.8
## Lee (2016)                    0.8
## Lee (2016)                    0.8
## Lee (2016)                    0.9
## Bradbury and Stephenson (2009) 10.0
## Chen and Malhotra (2007)       2.1
## Chen and Malhotra (2007)       7.3
## Chen and Malhotra (2007)       7.0
## Chen and Malhotra (2007)       6.4
## Chen and Malhotra (2007)       2.6
## Chen and Malhotra (2007)       8.2
## Chen and Malhotra (2007)       0.8
## Primo (2006)                   9.7
## Primo (2006)                   8.5
## Primo (2006)                   8.4
## Primo (2006)                   7.6
##
## Number of studies combined: k = 24

```

```
##
##              SMD              95%-CI
## Random effects model 7.2162 [ 4.4400;  9.9925]
## Prediction interval      [-1.2217; 15.6542]
##              t   p-value
## Random effects model 5.38 < 0.0001
## Prediction interval
##
## Quantifying heterogeneity:
##  tau^2 = 14.7532 [5.4141; 111.2304]; tau = 3.8410 [2.3268; 10.5466];
##  I^2 = 77.7% [67.3%; 84.8%]; H = 2.12 [1.75; 2.57]
##
## Test of heterogeneity:
##      Q d.f.  p-value
##  103.34   23 < 0.0001
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model
```

And the forest plot:

```
build_forest(mod, NULL)
```

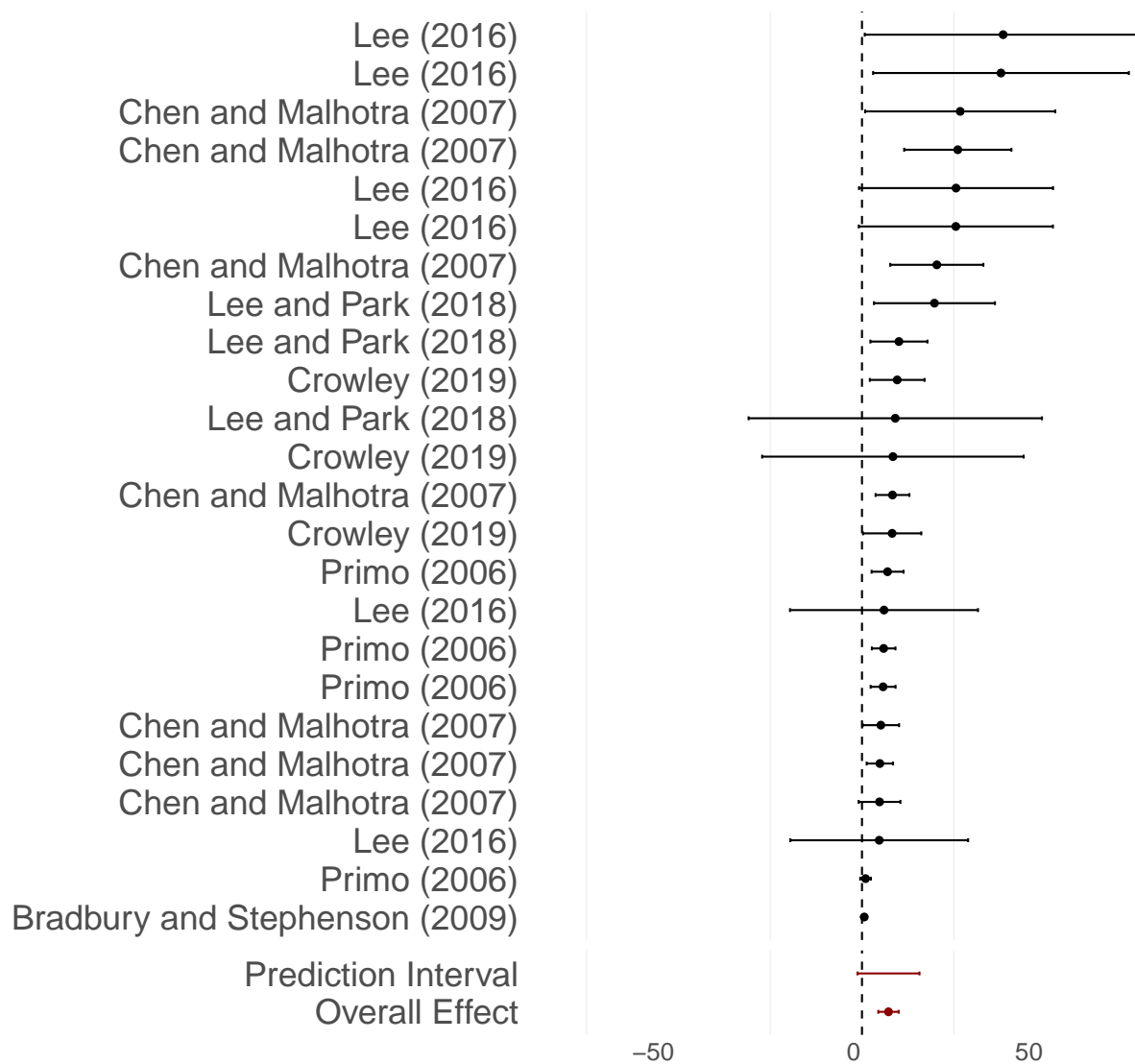


Figure 17: Effect of upper house size (K) on the per capita government expenditure (ExpPC)

Highlights:

1. The results are highly heterogeneous: $I^2 = 77.74$.
2. The estimated SMD in the random effects model is $g = 7.22$ ($SE = 1.342$).
3. The prediction interval ranges from -1.22 to 15.65. Therefore, it encompasses zero.

H.4 Lower House Size and Log of Expenditure Per Capita

This model estimates the Log of Per Capita Expenditure as the dependent variable, and the number of lower house legislators as the treatment variable.

```
# Pooling effects analysis -- logExpPC x N
aux <- fulldat %>%
  filter(indepvar2 == 'N',
```

```

depvar2 == 'logExpPC')

mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
               comb.fixed = FALSE,
               comb.random = TRUE,
               method.tau = "REML",
               hakn = TRUE,
               prediction=TRUE,
               sm="SMD")

mod

```

##	SMD
## Lewis (2019)	-0.1740
## Höhmann (2017)	-0.0300
## Höhmann (2017)	-0.0300
## Höhmann (2017)	-0.0400
## Drew and Dollery (2017)	0.0770
## Drew and Dollery (2017)	0.0310
## Pettersson-Lidbom (2012)	-0.1590
## Pettersson-Lidbom (2012)	-0.1470
## Pettersson-Lidbom (2012)	-0.0900
## Pettersson-Lidbom (2012)	-0.0810
## Pettersson-Lidbom (2012)	-0.0880
## Pettersson-Lidbom (2012)	0.2100
## Pettersson-Lidbom (2012)	0.1570
## Pettersson-Lidbom (2012)	-0.1990
## Pettersson-Lidbom (2012)	-0.1690
##	95%-CI
## Lewis (2019)	[-0.2450; -0.1030]
## Höhmann (2017)	[-0.0496; -0.0104]
## Höhmann (2017)	[-0.0496; -0.0104]
## Höhmann (2017)	[-0.0792; -0.0008]
## Drew and Dollery (2017)	[0.0221; 0.1319]

```

## Drew and Dollery (2017) [-0.0121; 0.0741]
## Pettersson-Lidbom (2012) [-0.2394; -0.0786]
## Pettersson-Lidbom (2012) [-0.2274; -0.0666]
## Pettersson-Lidbom (2012) [-0.1645; -0.0155]
## Pettersson-Lidbom (2012) [-0.1574; -0.0046]
## Pettersson-Lidbom (2012) [-0.1625; -0.0135]
## Pettersson-Lidbom (2012) [ 0.1649; 0.2551]
## Pettersson-Lidbom (2012) [ 0.0845; 0.2295]
## Pettersson-Lidbom (2012) [-0.2774; -0.1206]
## Pettersson-Lidbom (2012) [-0.2494; -0.0886]
##
##                               %W(random)
## Lewis (2019)                  6.6
## Höhmann (2017)                7.1
## Höhmann (2017)                7.1
## Höhmann (2017)                7.0
## Drew and Dollery (2017)       6.8
## Drew and Dollery (2017)       6.9
## Pettersson-Lidbom (2012)      6.4
## Pettersson-Lidbom (2012)      6.4
## Pettersson-Lidbom (2012)      6.5
## Pettersson-Lidbom (2012)      6.5
## Pettersson-Lidbom (2012)      6.5
## Pettersson-Lidbom (2012)      6.9
## Pettersson-Lidbom (2012)      6.5
## Pettersson-Lidbom (2012)      6.4
## Pettersson-Lidbom (2012)      6.4
##
## Number of studies combined: k = 15
##
##                               SMD          95%-CI
## Random effects model -0.0463 [-0.1142; 0.0216]
## Prediction interval      [-0.3105; 0.2178]
##
##                               t p-value

```



```
## Random effects model -1.46  0.1655

## Prediction interval

##

## Quantifying heterogeneity:

##  tau^2 = 0.0139 [0.0070; 0.0364]; tau = 0.1181 [0.0836; 0.1908];

##  I^2 = 93.8% [91.2%; 95.6%]; H = 4.00 [3.38; 4.75]

##

## Test of heterogeneity:

##      Q d.f.  p-value

## 224.56   14 < 0.0001

##

## Details on meta-analytical method:

## - Inverse variance method

## - Restricted maximum-likelihood estimator for tau^2

## - Q-profile method for confidence interval of tau^2 and tau

## - Hartung-Knapp adjustment for random effects model
```

The forest plot is shown below:

```
build_forest(mod, NULL)
```

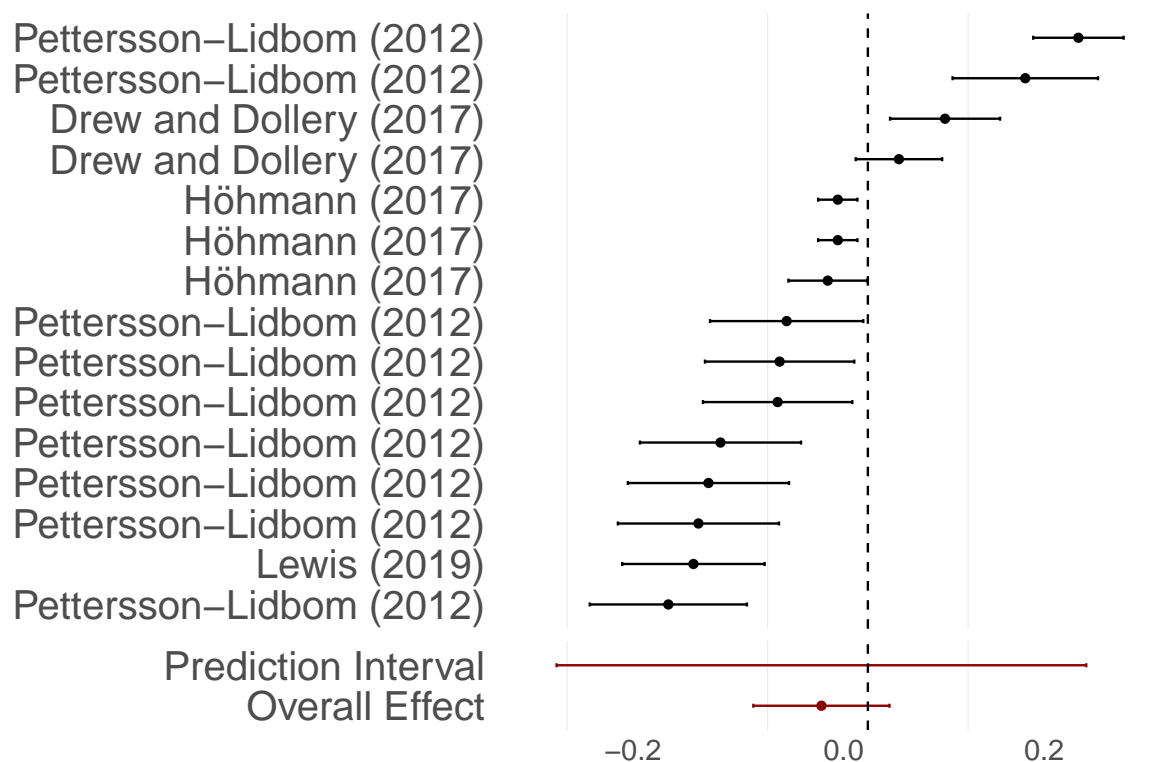


Figure 18: Effect of lower houses size (N) on log of per capita expenditure (logExpPC)

Highlights:

1. The results are highly heterogeneous: $I^2 = 93.77$.
2. The estimated SMD in the random effects model is $g = -0.05$ ($SE = 0.032$).
3. The prediction interval ranges from -0.31 to 0.22. Therefore, it encompasses zero.

H.5 Log of Lower House Size and Log of Expenditure Per Capita

In this specification, we study the log of per capita expenditure (logExpPC) as a function of the log of lower house size (logN).

```
# Pooling effects analysis -- logExpPC x logN

aux <- fulldat %>%
  filter(indepvar2 == 'logN',
         depvar2 == 'logExpPC')

mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
               comb.fixed = FALSE,
```

```

comb.random = TRUE,

method.tau = "REML",

hakn = TRUE,

prediction=TRUE,

sm="SMD")

```

mod

```

##                SMD                95%-CI
## MacDonald (2008) 0.1360 [0.0447; 0.2273]
## MacDonald (2008) 0.2319 [0.1322; 0.3316]
## MacDonald (2008) 0.1443 [0.0471; 0.2415]
## MacDonald (2008) 0.1594 [0.0667; 0.2521]
## MacDonald (2008) 0.2259 [0.1163; 0.3355]
## Baqir (2002)     0.1127 [0.0396; 0.1858]
## Baqir (2002)     0.2760 [0.2007; 0.3513]
## Baqir (2002)     0.3021 [0.2270; 0.3772]
## Baqir (2002)     0.3203 [0.2450; 0.3956]
## Baqir (1999)     0.3020 [0.2269; 0.3771]
## Baqir (1999)     0.2760 [0.2007; 0.3513]
## Baqir (1999)     0.2950 [0.2165; 0.3735]
##
##                %W(random)
## MacDonald (2008)      7.9
## MacDonald (2008)      7.4
## MacDonald (2008)      7.6
## MacDonald (2008)      7.8
## MacDonald (2008)      6.9
## Baqir (2002)          9.1
## Baqir (2002)          8.9
## Baqir (2002)          8.9
## Baqir (2002)          8.9
## Baqir (1999)          8.9
## Baqir (1999)          8.9
## Baqir (1999)          8.7
##

```

```

## Number of studies combined: k = 12

##
##              SMD              95%-CI      t
## Random effects model 0.2346 [0.1864; 0.2828] 10.71
## Prediction interval      [0.0848; 0.3844]
##              p-value
## Random effects model < 0.0001
## Prediction interval
##
## Quantifying heterogeneity:
## tau^2 = 0.0040 [0.0011; 0.0145]; tau = 0.0636 [0.0335; 0.1203];
## I^2 = 70.0% [45.6%; 83.4%]; H = 1.82 [1.36; 2.45]
##
## Test of heterogeneity:
##      Q d.f. p-value
## 36.62  11  0.0001
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model

```

The forest plot:

```
build_forest(mod, NULL)
```

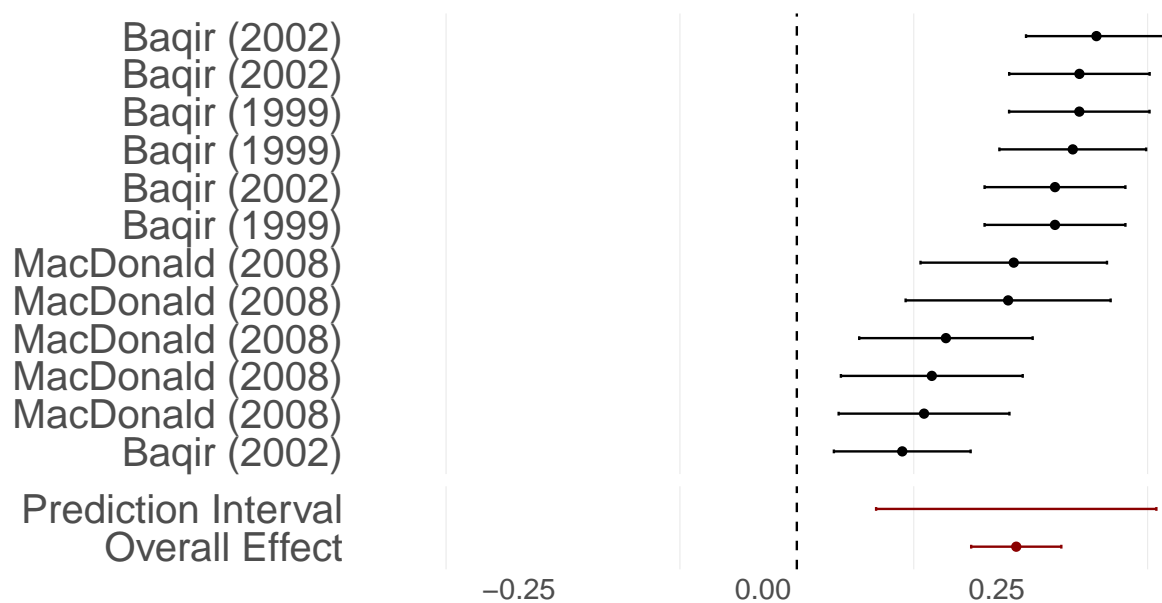


Figure 19: Effect of log lower houses size (logN) on the log of per capita government expenditure (logExpPC)

Highlights:

1. The results are highly heterogeneous: $I^2 = 69.96$.
2. The estimated SMD in the random effects model is $g = 0.23$ ($SE = 0.022$). **This model is significant at the 10% confidence level.**
3. The prediction interval ranges from 0.08 to 0.38. Therefore, it does not encompass zero.

H.6 Upper House Size and Log of Expenditure Per Capita

No studies related the log of per capita expenditure with the size of upper house (K).

H.7 Lower House Size and Expenditure as Percentage of GDP

This model fits the random effects for the percentage of GDP as public expenditure as the main outcome, and the size of lower house as the main treatment variable.

```
# Pooling effects analysis -- PCTGDP x N

aux <- fulldat %>%
  filter(indepvar2 == 'N',
         depvar2 == 'PCTGDP')

mod <- metagen(coef, SE, data = aux,
               studlab = paste(authoryear),
```

```

comb.fixed = FALSE,
comb.random = TRUE,
method.tau = "REML",
hakn = TRUE,
prediction=TRUE,
sm = "SMD")

```

mod

##	SMD
## Bjedov et al. (2014)	-0.0040
## Bjedov et al. (2014)	-0.0080
## Maldonado (2013)	-0.0609
## Mukherjee (2003)	0.0030
## Mukherjee (2003)	0.0090
## Mukherjee (2003)	0.0110
## Mukherjee (2003)	0.0050
## Mukherjee (2003)	0.0400
## Mukherjee (2003)	0.0300
## Mukherjee (2003)	0.0100
## Mukherjee (2003)	0.0200
## Bradbury and Crain (2001)	0.0036
## Bradbury and Crain (2001)	0.0005
## Bradbury and Crain (2001)	0.0169
## Bradbury and Crain (2001)	0.0123
## Ricciuti (2004)	0.0140
## Ricciuti (2004)	-0.0110
## Ricciuti (2004)	0.0070
## Ricciuti (2004)	0.0050
## Ricciuti (2004)	0.0050
## Ricciuti (2004)	0.0120
##	95%-CI
## Bjedov et al. (2014)	[-0.0432; 0.0352]
## Bjedov et al. (2014)	[-0.0472; 0.0312]
## Maldonado (2013)	[-0.0838; -0.0380]

## Mukherjee (2003)	[0.0010; 0.0050]
## Mukherjee (2003)	[0.0051; 0.0129]
## Mukherjee (2003)	[0.0051; 0.0169]
## Mukherjee (2003)	[-0.0009; 0.0109]
## Mukherjee (2003)	[0.0380; 0.0420]
## Mukherjee (2003)	[0.0280; 0.0320]
## Mukherjee (2003)	[0.0061; 0.0139]
## Mukherjee (2003)	[0.0122; 0.0278]
## Bradbury and Crain (2001)	[0.0008; 0.0065]
## Bradbury and Crain (2001)	[-0.0016; 0.0027]
## Bradbury and Crain (2001)	[0.0131; 0.0208]
## Bradbury and Crain (2001)	[0.0087; 0.0160]
## Ricciuti (2004)	[-0.0095; 0.0375]
## Ricciuti (2004)	[-0.0286; 0.0066]
## Ricciuti (2004)	[-0.0067; 0.0207]
## Ricciuti (2004)	[-0.0126; 0.0226]
## Ricciuti (2004)	[-0.0126; 0.0226]
## Ricciuti (2004)	[-0.0017; 0.0257]
##	%W(random)
## Bjedov et al. (2014)	2.1
## Bjedov et al. (2014)	2.1
## Maldonado (2013)	3.6
## Mukherjee (2003)	5.6
## Mukherjee (2003)	5.5
## Mukherjee (2003)	5.4
## Mukherjee (2003)	5.4
## Mukherjee (2003)	5.6
## Mukherjee (2003)	5.6
## Mukherjee (2003)	5.5
## Mukherjee (2003)	5.3
## Bradbury and Crain (2001)	5.6
## Bradbury and Crain (2001)	5.6
## Bradbury and Crain (2001)	5.6

```

## Bradbury and Crain (2001)          5.6
## Ricciuti (2004)                   3.5
## Ricciuti (2004)                   4.2
## Ricciuti (2004)                   4.7
## Ricciuti (2004)                   4.2
## Ricciuti (2004)                   4.2
## Ricciuti (2004)                   4.7
##
## Number of studies combined: k = 21
##
##                               SMD          95%-CI      t
## Random effects model 0.0078 [-0.0003; 0.0160] 2.01
## Prediction interval          [-0.0259; 0.0416]
##                               p-value
## Random effects model 0.0579
## Prediction interval
##
## Quantifying heterogeneity:
## tau^2 = 0.0002 [0.0002; 0.0007]; tau = 0.0156 [0.0136; 0.0261];
## I^2 = 98.5% [98.2%; 98.7%]; H = 8.11 [7.40; 8.88]
##
## Test of heterogeneity:
##      Q d.f.  p-value
## 1314.54  20 < 0.0001
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model

```

Here is the forest plot:


```
build_forest(mod, NULL)
```

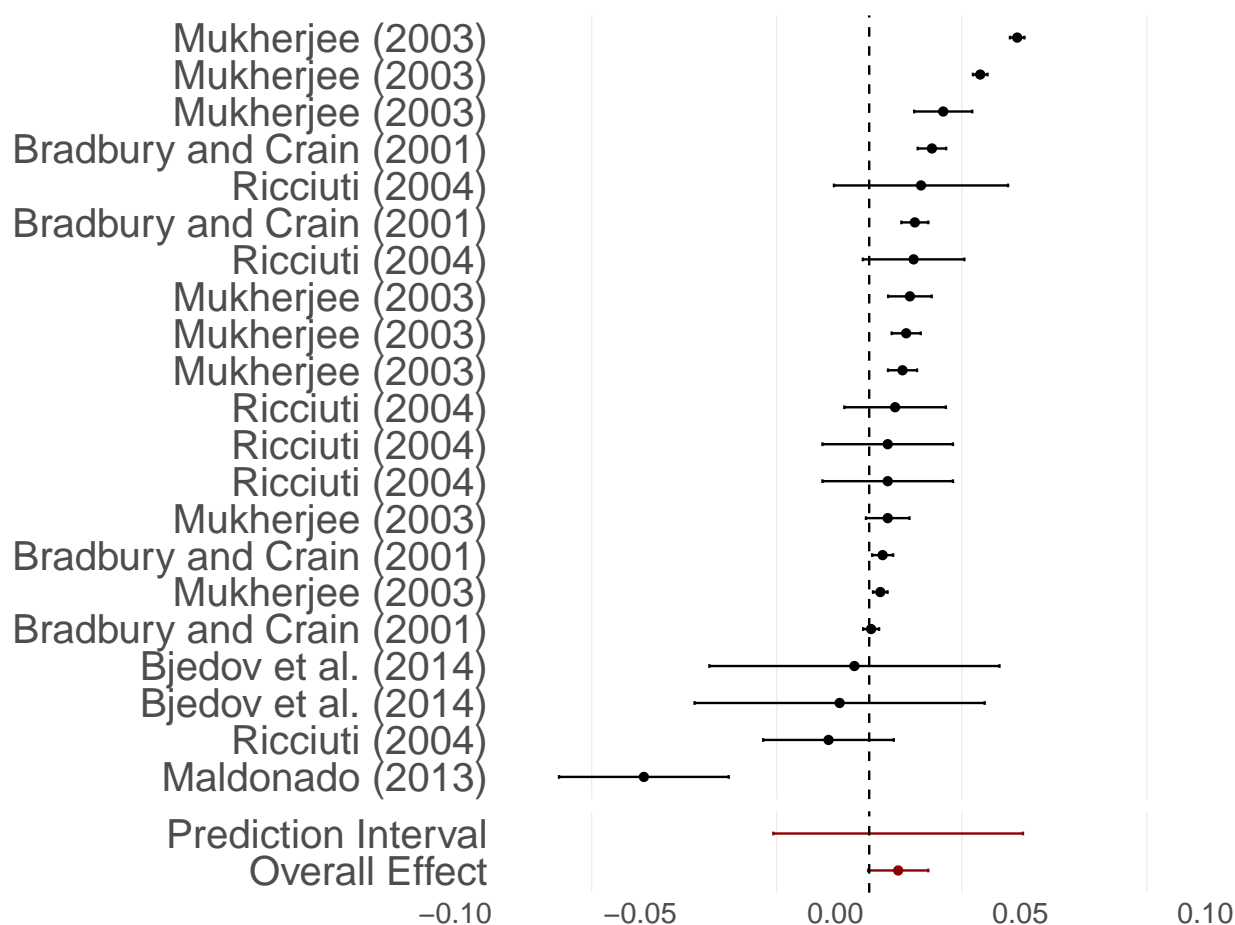


Figure 20: Effect of lower houses size (N) on percentage of public expenditure GDP (PCTGDP)

Highlights:

1. The results are highly heterogeneous: $I^2 = 98.48$.
2. The estimated SMD in the random effects model is $g = 0.01$ ($SE = 0.004$).
3. The prediction interval ranges from -0.03 to 0.04. Therefore, it encompasses zero.

H.8 Log of Lower House Size and Expenditure as Percentage of GDP

This meta-regression investigates the percentage of GDP as public expenditure as the dependent variable and the natural logarithm of lower house size ($\log(N)$) as the treatment variable.

```
# Pooling effects analysis -- PCTGDP x logN

aux <- fulldat %>%
  filter(indepvar2 == 'logN',
         depvar2 == 'PCTGDP')
```

```
mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
               comb.fixed = FALSE,
               comb.random = TRUE,
               method.tau = "REML",
               hakn = TRUE,
               prediction=TRUE,
               sm="SMD")
```

```
mod
```

```
##                SMD                95%-CI
## Baqir (1999)      2.0660 [ 1.4887; 2.6433]
## Baqir (1999)      2.0120 [ 1.4235; 2.6005]
## Baqir (1999)      2.4680 [ 1.8817; 3.0543]
## Lledo (2003)      -4.6900 [-9.9427; 0.5627]
## Stein et al. (1998) 0.0109 [-0.0171; 0.0389]
## Stein et al. (1998) 0.0135 [-0.0102; 0.0372]
##                %W(random)
## Baqir (1999)      18.9
## Baqir (1999)      18.8
## Baqir (1999)      18.8
## Lledo (2003)      3.8
## Stein et al. (1998) 19.8
## Stein et al. (1998) 19.8
##
## Number of studies combined: k = 6
##
##                SMD                95%-CI    t
## Random effects model 1.0619 [-0.7256; 2.8493] 1.53
## Prediction interval      [-3.0267; 5.1504]
##                p-value
## Random effects model  0.1873
## Prediction interval
```

```
##
## Quantifying heterogeneity:
## tau^2 = 1.6850 [0.6497; 38.1618]; tau = 1.2981 [0.8060; 6.1775];
## I^2 = 96.9% [95.2%; 98.1%]; H = 5.71 [4.55; 7.16]
##
## Test of heterogeneity:
##      Q d.f.  p-value
## 163.00    5 < 0.0001
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model
```

The forest plot:

```
build_forest(mod, NULL)
```

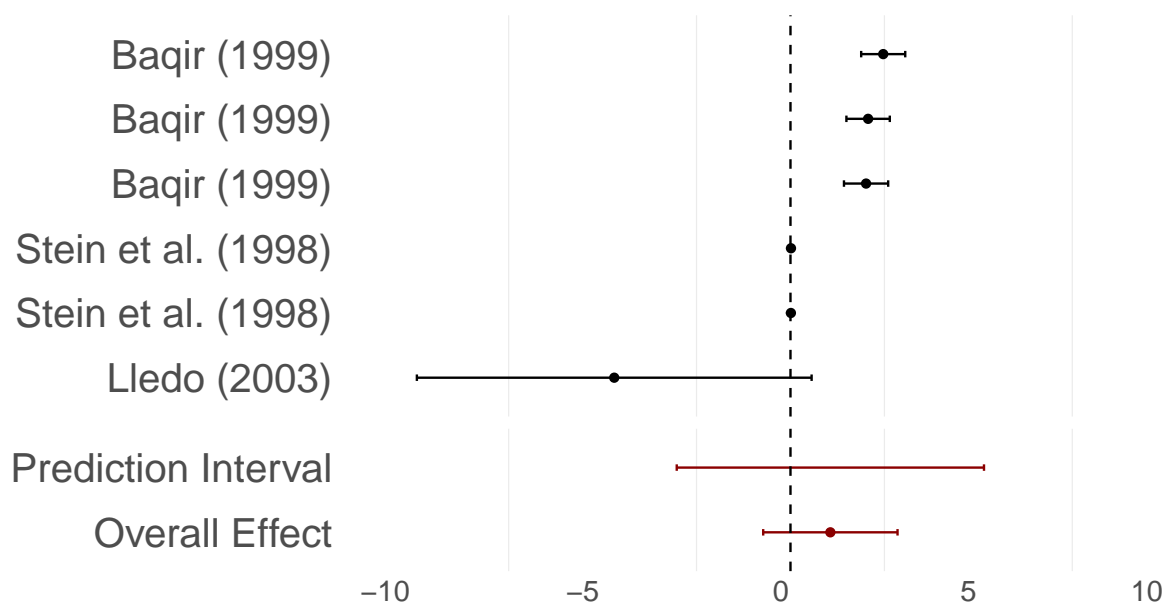


Figure 21: Effect of log lower houses size (logN) on the GDP share of public expenditure (PCTGDP)

Highlights:

1. The results are highly heterogeneous: $I^2 = 96.93$.

2. The estimated SMD in the random effects model is $g = 1.06$ ($SE = 0.695$).
3. The prediction interval ranges from -3.03 to 5.15. Therefore, it encompasses zero.

H.9 Upper House Size and Expenditure as Percentage of GDP

This model looks into the effect of upper house size (K) on the public expenditure share of the GDP (PCTGDP).

```
# Pooling effects analysis -- PCTGDP x K
```

```
aux <- fulldat %>%
```

```
  filter(indepvar2 == 'K',
         depvar2 == 'PCTGDP')
```

```
mod <- metagen(coef, SE, data=aux,
               studlab=paste(authoryear),
               comb.fixed = FALSE,
               comb.random = TRUE,
               method.tau = "REML",
               hakn = TRUE,
               prediction=TRUE,
               sm="SMD")
```

```
mod
```

##	SMD
## Maldonado (2012)	-0.0400
## Bradbury and Crain (2001)	0.0126
## Bradbury and Crain (2001)	0.0050
## Bradbury and Crain (2001)	-0.0113
## Bradbury and Crain (2001)	-0.0056
## Ricciuti (2004)	0.0160
## Ricciuti (2004)	0.0210
## Ricciuti (2004)	0.0140
## Ricciuti (2004)	0.0030
## Ricciuti (2004)	0.0300
## Ricciuti (2004)	0.0300
## Ricciuti (2004)	0.0390

## Ricciuti (2004)	0.0127
## Ricciuti (2004)	0.0160
##	95%-CI
## Maldonado (2012)	[-0.0659; -0.0141]
## Bradbury and Crain (2001)	[0.0010; 0.0243]
## Bradbury and Crain (2001)	[0.0016; 0.0083]
## Bradbury and Crain (2001)	[-0.0163; -0.0064]
## Bradbury and Crain (2001)	[-0.0102; -0.0010]
## Ricciuti (2004)	[-0.0075; 0.0395]
## Ricciuti (2004)	[-0.0006; 0.0426]
## Ricciuti (2004)	[-0.0036; 0.0316]
## Ricciuti (2004)	[-0.0088; 0.0148]
## Ricciuti (2004)	[-0.0210; 0.0810]
## Ricciuti (2004)	[-0.0210; 0.0810]
## Ricciuti (2004)	[-0.0022; 0.0802]
## Ricciuti (2004)	[-0.0147; 0.0401]
## Ricciuti (2004)	[-0.0075; 0.0395]
##	%W(random)
## Maldonado (2012)	5.7
## Bradbury and Crain (2001)	9.8
## Bradbury and Crain (2001)	11.8
## Bradbury and Crain (2001)	11.5
## Bradbury and Crain (2001)	11.6
## Ricciuti (2004)	6.2
## Ricciuti (2004)	6.7
## Ricciuti (2004)	7.9
## Ricciuti (2004)	9.7
## Ricciuti (2004)	2.2
## Ricciuti (2004)	2.2
## Ricciuti (2004)	3.1
## Ricciuti (2004)	5.3
## Ricciuti (2004)	6.2
##	

```

## Number of studies combined: k = 14

##
##              SMD              95%-CI      t
## Random effects model 0.0056 [-0.0042; 0.0155] 1.24
## Prediction interval          [-0.0233; 0.0346]
##
##              p-value
## Random effects model 0.2376
## Prediction interval
##
## Quantifying heterogeneity:
## tau^2 = 0.0002 [0.0001; 0.0008]; tau = 0.0125 [0.0109; 0.0279];
## I^2 = 80.0% [67.3%; 87.8%]; H = 2.24 [1.75; 2.86]
##
## Test of heterogeneity:
##      Q d.f.  p-value
## 65.02   13 < 0.0001
##
## Details on meta-analytical method:
## - Inverse variance method
## - Restricted maximum-likelihood estimator for tau^2
## - Q-profile method for confidence interval of tau^2 and tau
## - Hartung-Knapp adjustment for random effects model

```

And the forest plot:

```
build_forest(mod, NULL)
```

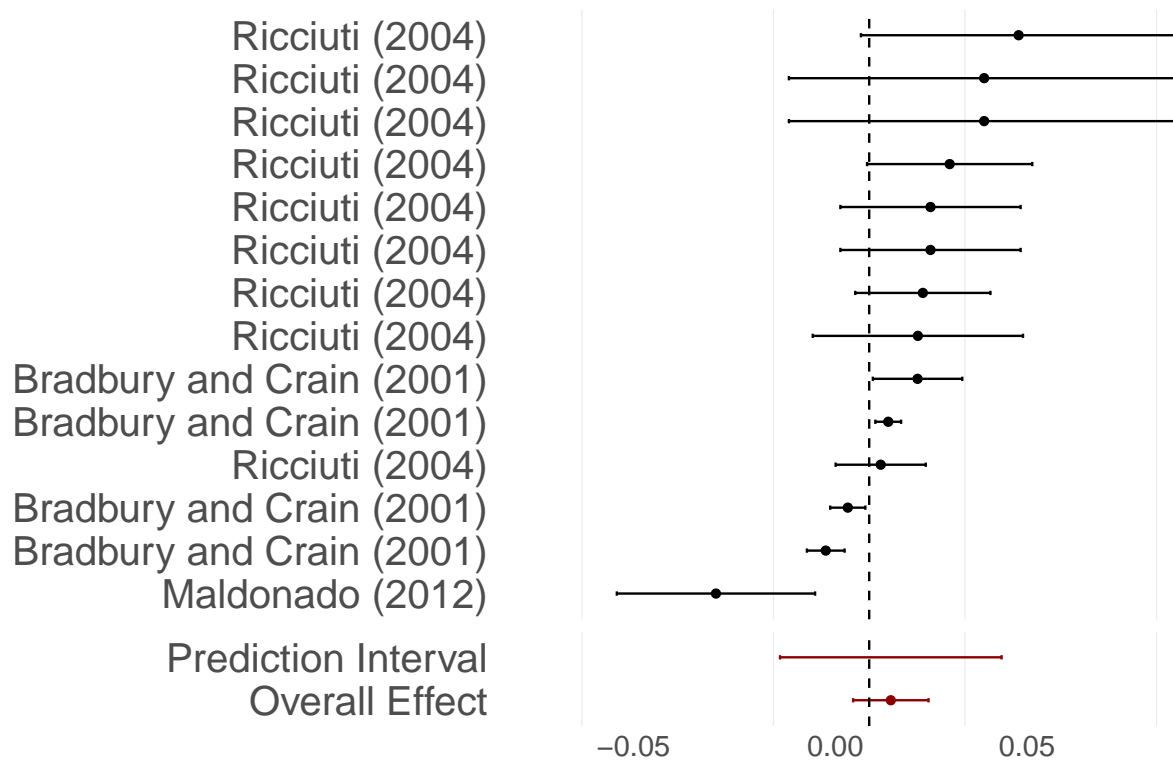


Figure 22: Effect of upper house size (K) on the public expenditure share of the GDP (PCTGDP)

Highlights:

1. The results are highly heterogeneous: $I^2 = 80.01$.
2. The estimated SMD in the random effects model is $g = 0.01$ ($SE = 0.005$).
3. The prediction interval ranges from -0.02 to 0.03. Therefore, it encompasses zero.

I Meta-Regressions

I.1 Meta-Regressions for Expenditure as a Percentage of the GDP

```
mod <- rma(yi = coef,
  sei = SE,
  data = dat,
  method = "REML",
  mods = ~indepvar2+year+published+elecsys2+method,
  test = "knha",
  subset = dat$depvar2=='PCTGDP',
  slab = dat$authoryear)
```

```

mod1 <- tibble(
  ` ` = c("Intercept",
    "Indepvar: N",
    "Indepvar: logN",
    "Year",
    "Published: No",
    "Elecsys: Non-Majoritarian",
    "Method: Panel"),
  Estimate = mod[["beta"]],
  SE = mod[["se"]],
  `T` = mod[["zval"]],
  `P-Value` = mod[["pval"]],
  CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
    round(mod[["ci.ub"]], digits = 4), ")"),
  model = "PCTGDP"
) %>%
  mutate_if(is.numeric, list(~round(., digits = 4)))

```

```
summary(mod)
```

```

##
## Mixed-Effects Model (k = 11; tau^2 estimator: REML)
##
##   logLik  deviance      AIC      BIC     AICc
##   8.9439  -17.8879   -1.8879   -6.7975  142.1121
##
## tau^2 (estimated amount of residual heterogeneity):    0.0000 (SE = 0.0000)
## tau (square root of estimated tau^2 value):           0.0009
## I^2 (residual heterogeneity / unaccounted variability): 3.13%
## H^2 (unaccounted variability / sampling variability):   1.03
## R^2 (amount of heterogeneity accounted for):           100.00%
##
## Test for Residual Heterogeneity:
## QE(df = 4) = 3.6928, p-val = 0.4492

```



```
##

## Test of Moderators (coefficients 2:7):

## F(df1 = 6, df2 = 4) = 16.9635, p-val = 0.0082

##

## Model Results:

##
```

	estimate	se			
## intrcpt	2.8561	1.7784			
## indepvar2N	-0.0093	0.0053			
## indepvar2logN	-0.0114	0.0169			
## year	-0.0004	0.0009			
## publishedNo	0.0625	0.0139			
## elecsys2Non-Majoritarian	-2.0557	0.2831			
## methodPANEL	0.0547	0.0129			

	tval	pval	ci.lb	ci.ub
## intrcpt	1.6060	0.1835	-2.0814	7.7937
## indepvar2N	-1.7367	0.1574	-0.0241	0.0055
## indepvar2logN	-0.6746	0.5369	-0.0583	0.0355
## year	-0.4825	0.6547	-0.0028	0.0020
## publishedNo	4.5057	0.0108	0.0240	0.1011 *
## elecsys2Non-Majoritarian	-7.2618	0.0019	-2.8416	-1.2697 **
## methodPANEL	4.2531	0.0131	0.0190	0.0903 *

```
##

## ---

## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

As we have considerable heterogeneity in our sample, we run a permutation test to ensure the validity of our estimates. The results follow below.

```
mod <- permutest(mod, progbar = F)
mod

##
## Test of Moderators (coefficients 2:7):
## F(df1 = 6, df2 = 4) = 16.9635, p-val* = 0.0050
##
## Model Results:
##
##               estimate      se      tval
## intrcpt          2.8561  1.7784   1.6060
## indepvar2N       -0.0093  0.0053  -1.7367
## indepvar2logN    -0.0114  0.0169  -0.6746
## year            -0.0004  0.0009  -0.4825
## publishedNo       0.0625  0.0139   4.5057
## elecsys2Non-Majoritarian -2.0557  0.2831  -7.2618
## methodPANEL       0.0547  0.0129   4.2531
##
##               pval*    ci.lb    ci.ub
## intrcpt          0.1250  -2.0814   7.7937
## indepvar2N       0.0570  -0.0241   0.0055  .
## indepvar2logN    0.5660  -0.0583   0.0355
## year            0.5090  -0.0028   0.0020
## publishedNo      0.0140   0.0240   0.1011  *
## elecsys2Non-Majoritarian 0.0050  -2.8416  -1.2697  **
## methodPANEL     0.0360   0.0190   0.0903  *
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

mod1_permu <- tibble(
  ` ` = c("Intercept",
    "Indepvar: N",
```

```

      "Indepvar: logN",
      "Year",
      "Published: No",
      "Elecsys: Non-Majoritarian",
      "Method: Panel"),
Estimate = mod[["beta"]],
SE = mod[["se"]],
`T` = mod[["zval"]],
`P-Value` = mod[["pval"]],
CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
            round(mod[["ci.ub"]], digits = 4), ")"),
model = "PCTGDP - Permutation"
) %>%
mutate_if(is.numeric, list(~round(., digits = 3)))

```

We have the following results for the meta-regressions of Expenditure as a Percentage of GDP:

1. Compared with K, models with N and logN find significantly negative coefficients.
2. Year has null effect.
3. Unpublished papers tend to have higher coefficients than published papers.
4. Passing from Majoritarian to Non-Majoritarian, decreases significantly the effects found in our models.
5. In terms of modelling, passing from OLS to PANEL increases the detected effects.
6. When passing from Country to Cross-Country, it has no effect on the estimated coefficients.

Below we also run the meta-regressions adding all coefficients in the papers.

```

mod <- rma(yi = coef,
           sei = SE,
           data = fulldat,
           method = "REML",
           mods = ~ indepvar2+year+published+elecsys2+method,
           test = "knha",
           subset = fulldat$depvar2=='PCTGDP',
           slab = fulldat$authoryear)

```

```

mod2 <- tibble(
  ` ` = c("Intercept",
    "Indepvar: N",
    "Indepvar: logN",
    "Year",
    "Published: No",
    "Elecsys: Non-Majoritarian",
    "Method: Panel"),
  Estimate = mod[["beta"]],
  SE = mod[["se"]],
  `T` = mod[["zval"]],
  `P-Value` = mod[["pval"]],
  CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
    round(mod[["ci.ub"]], digits = 4), ")"),
  model = "PCTGDP - All coefs"
) %>%
  mutate_if(is.numeric, list(~round(., digits = 3)))

```

```
summary(mod)
```

```

##
## Mixed-Effects Model (k = 41; tau^2 estimator: REML)
##
##      logLik   deviance      AIC      BIC
##  86.8787  -173.7575  -157.7575  -145.5466
##
##      AICc
## -151.9975
##
## tau^2 (estimated amount of residual heterogeneity):    0.0001 (SE = 0.0000)
## tau (square root of estimated tau^2 value):           0.0110
## I^2 (residual heterogeneity / unaccounted variability): 94.65%
## H^2 (unaccounted variability / sampling variability):   18.70
## R^2 (amount of heterogeneity accounted for):           99.91%
##

```

```

## Test for Residual Heterogeneity:

## QE(df = 34) = 1064.2732, p-val < .0001

##

## Test of Moderators (coefficients 2:7):

## F(df1 = 6, df2 = 34) = 33.5853, p-val < .0001

##

## Model Results:

##

##               estimate      se
## intrcpt          2.5076  2.4669
## indepvar2N        0.0028  0.0050
## indepvar2logN      0.0020  0.0142
## year             -0.0002  0.0012
## publishedNo        0.0601  0.0158
## elecsys2Non-Majoritarian -2.1692  0.1658
## methodPANEL        0.0582  0.0173
##               tval      pval
## intrcpt          1.0165  0.3166
## indepvar2N        0.5728  0.5705
## indepvar2logN      0.1381  0.8910
## year             -0.1589  0.8747
## publishedNo        3.8002  0.0006
## elecsys2Non-Majoritarian -13.0856 <.0001
## methodPANEL        3.3702  0.0019
##               ci.lb      ci.ub
## intrcpt          -2.5057  7.5209
## indepvar2N       -0.0073  0.0129
## indepvar2logN    -0.0269  0.0308
## year             -0.0027  0.0023
## publishedNo        0.0280  0.0923 ***
## elecsys2Non-Majoritarian -2.5061 -1.8323 ***
## methodPANEL        0.0231  0.0932 **
##

```

```
## ---

## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
mod <- permutest(mod, progbars = F)
```

```
mod
```

```
##
```

```
## Test of Moderators (coefficients 2:7):
```

```
## F(df1 = 6, df2 = 34) = 33.5853, p-val* = 0.0010
```

```
##
```

```
## Model Results:
```

```
##
```

	estimate	se	tval	
intrcpt	2.5076	2.4669	1.0165	
indepvar2N	0.0028	0.0050	0.5728	
indepvar2logN	0.0020	0.0142	0.1381	
year	-0.0002	0.0012	-0.1589	
publishedNo	0.0601	0.0158	3.8002	
elecsys2Non-Majoritarian	-2.1692	0.1658	-13.0856	
methodPANEL	0.0582	0.0173	3.3702	
	pval*	ci.lb	ci.ub	
intrcpt	0.1530	-2.5057	7.5209	
indepvar2N	0.5170	-0.0073	0.0129	
indepvar2logN	0.9310	-0.0269	0.0308	
year	0.7960	-0.0027	0.0023	
publishedNo	0.0110	0.0280	0.0923	*
elecsys2Non-Majoritarian	0.0010	-2.5061	-1.8323	***
methodPANEL	0.0130	0.0231	0.0932	*

```
##
```

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
mod2_permu <- tibble(
```

```
  ` ` = c("Intercept",
    "Indepvar: N",
```

```

      "Indepvar: logN",
      "Year",
      "Published: No",
      "Elecsys: Non-Majoritarian",
      "Method: Panel"),
Estimate = mod[["beta"]],
SE = mod[["se"]],
`T` = mod[["zval"]],
`P-Value` = mod[["pval"]],
CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
            round(mod[["ci.ub"]], digits = 4), ")"),
model = "PCTGDP - All coefs - Permutation"
) %>%
mutate_if(is.numeric, list(~round(., digits = 3)))

```

For all the coefficients, we have the following results:

1. Compared with K, models with N and logN tend to have significantly negative coefficients.
2. Year has a positive effect: the younger the publication, the higher the detected coefficient.
3. Unpublished papers tend to have higher coefficients than published papers.
4. Passing from Majoritarian to Non-Majoritarian significantly decreases the effects found in our models.
5. In terms of the modeling, passing from OLS to PANEL increases the detected effects.

I.2 Meta-Regressions for Expenditure Per Capita

```

mod <- rma(yi = coef,
           sei = SE,
           data = dat,
           method = "REML",
           mods = ~indepvar2+year+published+elecsys2+method,
           test = "knha",
           subset = dat$depvar2=='ExpPC',
           slab = dat$authoryear,
           control=list(maxiter=1000))

```

```

mod3 <- tibble(
  ` ` = c("Intercept",
    "Indepvar: N",
    "Year",
    "Elecsys: Non-Majoritarian",
    "Method: Panel",
    "Method: IV"),
  Estimate = mod[["beta"]],
  SE = mod[["se"]],
  `T` = mod[["zval"]],
  `P-Value` = mod[["pval"]],
  CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
    round(mod[["ci.ub"]], digits = 4), ")"),
  model = "ExpPC"
) %>%
  mutate_if(is.numeric, list(~round(., digits = 3)))

```

```
summary(mod)
```

```

##
## Mixed-Effects Model (k = 18; tau^2 estimator: REML)
##
##   logLik  deviance      AIC      BIC      AICc
## -35.6479   71.2957   85.2957   88.6901  113.2957
##
## tau^2 (estimated amount of residual heterogeneity):      0.4079 (SE = 0.3165)
## tau (square root of estimated tau^2 value):             0.6387
## I^2 (residual heterogeneity / unaccounted variability): 91.13%
## H^2 (unaccounted variability / sampling variability):    11.27
## R^2 (amount of heterogeneity accounted for):             6.03%
##
## Test for Residual Heterogeneity:
## QE(df = 12) = 96.9469, p-val < .0001
##

```



```

## Test of Moderators (coefficients 2:6):

## F(df1 = 5, df2 = 12) = 0.5468, p-val = 0.7381

##

## Model Results:

##

##               estimate          se
## intrcpt        -117.9871    203.7167
## indepvar2N       -2.3072     1.4898
## year             0.0591     0.1014
## elecsys2Non-Majoritarian  0.4283     0.8274
## methodPANEL      1.2027     1.5749
## methodIV         1.3170     1.8074
##
##               tval    pval
## intrcpt        -0.5792  0.5732
## indepvar2N     -1.5487  0.1474
## year            0.5826  0.5709
## elecsys2Non-Majoritarian  0.5177  0.6141
## methodPANEL     0.7637  0.4598
## methodIV        0.7287  0.4802
##
##               ci.lb    ci.ub
## intrcpt       -561.8476  325.8734
## indepvar2N     -5.5531   0.9387
## year          -0.1619   0.2801
## elecsys2Non-Majoritarian -1.3743  2.2310
## methodPANEL    -2.2286  4.6341
## methodIV       -2.6210  5.2551
##
## ---

## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

As we have considerable heterogeneity in our sample, we run a permutation test to ensure the validity of our estimates. The results follow below.

```

mod <- permutest(mod, progbar = F)

mod

##

## Test of Moderators (coefficients 2:6):

## F(df1 = 5, df2 = 12) = 0.5468, p-val* = 0.2850

##

## Model Results:

##

##               estimate          se      tval
## intrcpt          -117.9871    203.7167   -0.5792
## indepvar2N         -2.3072     1.4898   -1.5487
## year               0.0591     0.1014    0.5826
## elecsys2Non-Majoritarian  0.4283     0.8274    0.5177
## methodPANEL        1.2027     1.5749    0.7637
## methodIV           1.3170     1.8074    0.7287
##
##               pval*      ci.lb      ci.ub
## intrcpt          0.4040  -561.8476  325.8734
## indepvar2N        0.0560   -5.5531    0.9387
## year             0.4040   -0.1619    0.2801
## elecsys2Non-Majoritarian 0.4360   -1.3743    2.2310
## methodPANEL       0.2900   -2.2286    4.6341
## methodIV          0.3000   -2.6210    5.2551
##

## ---

## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

```

mod3_permu <- tibble(
  ` ` = c("Intercept",
    "Indepvar: N",
    "Year",
    "Elecsys: Non-Majoritarian",
    "Method: Panel",
    "Method: IV"),

```

```

Estimate = mod[["beta"]],
SE = mod[["se"]],
`T` = mod[["zval"]],
`P-Value` = mod[["pval"]],
CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
            round(mod[["ci.ub"]], digits = 4), ")"),
model = "ExpPC - Permutation"
) %>%
mutate_if(is.numeric, list(~round(., digits = 3)))

```

We have the following results for the meta-regressions of Expenditure Per Capita:

1. Compared with K, models with N tend to detect significantly smaller effects.
2. Year has null effect.
3. Majoritarian, when compared to Non-Majoritarian electoral systems, significantly increase per capita expenditure.
4. In terms of the modeling, passing from OLS to PANEL or IV increases the detected effects.

We also run the meta-regressions adding all coefficients in the papers. The results follow below:

```

mod <- rma(yi = coef,
           sei = SE,
           data = fulldat,
           method = "REML",
           mods = ~ indepvar2+year+published+elecsys2+method,
           test = "knha",
           subset = fulldat$depvar2=='ExpPC',
           slab = fulldat$authoryear)

mod4 <- tibble(
  ` ` = c("Intercept",
          "Indepvar: N",
          "Year",
          "Elecsys: Non-Majoritarian",
          "Method: Panel",

```

```

      "Method: IV"),
  Estimate = mod[["beta"]],
  SE = mod[["se"]],
  `T` = mod[["zval"]],
  `P-Value` = mod[["pval"]],
  CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
              round(mod[["ci.ub"]], digits = 4), ")"),
  model = "ExpPC - All coefs"
) %>%

mutate_if(is.numeric, list(~round(., digits = 3)))

```

```
summary(mod)
```

```

##
## Mixed-Effects Model (k = 60; tau^2 estimator: REML)
##
##      logLik    deviance      AIC      BIC
## -142.4363    284.8726    298.8726    312.7955
##      AICc
##   301.3073
##
## tau^2 (estimated amount of residual heterogeneity):    1.6376 (SE = 0.4634)
## tau (square root of estimated tau^2 value):           1.2797
## I^2 (residual heterogeneity / unaccounted variability): 99.79%
## H^2 (unaccounted variability / sampling variability):   472.09
## R^2 (amount of heterogeneity accounted for):           42.34%
##
## Test for Residual Heterogeneity:
## QE(df = 54) = 500.8874, p-val < .0001
##
## Test of Moderators (coefficients 2:6):
## F(df1 = 5, df2 = 54) = 7.1695, p-val < .0001
##
## Model Results:

```

```
##
##               estimate      se
## intrcpt      -300.7894  157.5549
## indepvar2N    -5.3436   0.9293
## year          0.1521   0.0785
## elecsys2Non-Majoritarian  0.9846   0.7229
## methodPANEL  -0.1346   0.7966
## methodIV      0.1861   0.8023
##               tval    pval
## intrcpt      -1.9091  0.0616
## indepvar2N    -5.7503  <.0001
## year          1.9364  0.0581
## elecsys2Non-Majoritarian  1.3619  0.1789
## methodPANEL  -0.1690  0.8665
## methodIV      0.2319  0.8175
##               ci.lb    ci.ub
## intrcpt      -616.6680  15.0892  .
## indepvar2N    -7.2066  -3.4805  ***
## year          -0.0054   0.3096  .
## elecsys2Non-Majoritarian  -0.4648   2.4340
## methodPANEL  -1.7316   1.4624
## methodIV      -1.4224   1.7946
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
mod <- permutest(mod, progbars = F)
```

```
mod
```

```
##
## Test of Moderators (coefficients 2:6):
## F(df1 = 5, df2 = 54) = 7.1695, p-val* = 0.0010
##
## Model Results:
```

```
##
##               estimate      se      tval
## intrcpt      -300.7894  157.5549  -1.9091
## indepvar2N    -5.3436   0.9293  -5.7503
## year          0.1521   0.0785   1.9364
## elecsys2Non-Majoritarian  0.9846   0.7229   1.3619
## methodPANEL  -0.1346   0.7966  -0.1690
## methodIV      0.1861   0.8023   0.2319
##
##           pval*      ci.lb      ci.ub
## intrcpt      0.0150 -616.6680  15.0892   *
## indepvar2N    0.0010  -7.2066 -3.4805  ***
## year          0.0130  -0.0054   0.3096   *
## elecsys2Non-Majoritarian  0.0820  -0.4648   2.4340   .
## methodPANEL    0.8110  -1.7316   1.4624
## methodIV      0.7720  -1.4224   1.7946
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
mod4_permu <- tibble(
  ` ` = c("Intercept",
    "Indepvar: N",
    "Year",
    "Elecsys: Non-Majoritarian",
    "Method: Panel",
    "Method: IV"),
  Estimate = mod[["beta"]],
  SE = mod[["se"]],
  `T` = mod[["zval"]],
  `P-Value` = mod[["pval"]],
  CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
    round(mod[["ci.ub"]], digits = 4), ")"),
  model = "ExpPC - All coefs - Permutation"
) %>%
```

```
mutate_if(is.numeric, list(~round(., digits = 3)))
```

With all coefficients, the results of the effect sizes on the Expenditure Per Capita regressions are the following:

1. Compared with κ , models with N tend to detect significantly smaller effects.
2. Year has now a positive effect on coefficient sizes.
3. Majoritarian, when compared to Non-Majoritarian electoral systems, significantly increase per capita expenditure.
4. In terms of the modeling, passing from OLS to PANEL decreases the detected effects.
5. All other coefficients were not significant.

I.3 Meta-Regressions for the Log of Expenditure Per Capita

```
mod <- rma(yi = coef,
           sei = SE,
           data = dat,
           method = "REML",
           mods = ~indepvar2+year+published+elecsys2+method,
           test = "knha",
           subset = dat$depvar2=='logExpPC',
           slab = dat$authoryear)
```

```
mod5 <- tibble(
  ` ` = c("Intercept",
           "Indepvar: N",
           "Year",
           "Published: No",
           "Method: Panel"),
  Estimate = mod[["beta"]],
  SE = mod[["se"]],
  `T` = mod[["zval"]],
  `P-Value` = mod[["pval"]],
  CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
              round(mod[["ci.ub"]], digits = 4), ")"),
```

```

model = "logExpPC"

) %>%

mutate_if(is.numeric, list(~round(., digits = 3)))

```

```
summary(mod)
```

```

##
## Mixed-Effects Model (k = 7; tau^2 estimator: REML)
##
##   logLik  deviance      AIC      BIC      AICc
##   2.1772   -4.3544    7.6456   -0.1956   91.6456
##
## tau^2 (estimated amount of residual heterogeneity):    0.0060 (SE = 0.0071)
## tau (square root of estimated tau^2 value):           0.0776
## I^2 (residual heterogeneity / unaccounted variability): 87.73%
## H^2 (unaccounted variability / sampling variability):   8.15
## R^2 (amount of heterogeneity accounted for):           78.05%
##
## Test for Residual Heterogeneity:
## QE(df = 2) = 21.6330, p-val < .0001
##
## Test of Moderators (coefficients 2:5):
## F(df1 = 4, df2 = 2) = 5.7815, p-val = 0.1529
##
## Model Results:
##
##           estimate      se    tval    pval
## intrcpt      -7.3634  25.4584  -0.2892  0.7996
## indepvar2N   -0.2806   0.1621  -1.7309  0.2256
## year          0.0037   0.0127   0.2941  0.7964
## publishedNo   0.2001   0.1269   1.5766  0.2556
## methodPANEL   0.1884   0.0928   2.0296  0.1795
##
##           ci.lb    ci.ub
## intrcpt    -116.9019  102.1751

```



```
## indepvar2N      -0.9780    0.4169
## year            -0.0509    0.0584
## publishedNo     -0.3459    0.7461
## methodPANEL     -0.2110    0.5879
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

As we have considerable heterogeneity in our sample, we run a permutation test to ensure the validity of our estimates. The results follow below.

```
mod <- permutest(mod, progbar = F)
```

```
mod
```

```
##
## Test of Moderators (coefficients 2:5):
## F(df1 = 4, df2 = 2) = 5.7815, p-val* = 0.1570
##
## Model Results:
##
##           estimate      se    tval  pval*
## intrcpt      -7.3634  25.4584  -0.2892  0.8120
## indepvar2N   -0.2806   0.1621  -1.7309  0.2030
## year          0.0037   0.0127   0.2941  0.8100
## publishedNo   0.2001   0.1269   1.5766  0.2680
## methodPANEL   0.1884   0.0928   2.0296  0.2060
##
##           ci.lb    ci.ub
## intrcpt    -116.9019  102.1751
## indepvar2N   -0.9780    0.4169
## year        -0.0509    0.0584
## publishedNo  -0.3459    0.7461
## methodPANEL -0.2110    0.5879
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```

mod5_permu <- tibble(
  ` ` = c("Intercept",
    "Indepvar: N",
    "Year",
    "Published: No",
    "Method: Panel"),
  Estimate = mod[["beta"]],
  SE = mod[["se"]],
  `T` = mod[["zval"]],
  `P-Value` = mod[["pval"]],
  CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
    round(mod[["ci.ub"]], digits = 4), ")"),
  model = "logExpPC - Permutation"
) %>%

mutate_if(is.numeric, list(~round(., digits = 3)))

mod5_permu %>%

select(-model) %>%

kable(booktabs = T, align = "c", linesep = '') %>%

kable_styling(c("striped", "bordered", "scale_down"), position = "center")

```

	Estimate	SE	T	P-Value	CI
Intercept	-7.363	25.458	-0.289	0.812	(-116.9019 ; 102.1751)
Indepvar: N	-0.281	0.162	-1.731	0.203	(-0.978 ; 0.4169)
Year	0.004	0.013	0.294	0.810	(-0.0509 ; 0.0584)
Published: No	0.200	0.127	1.577	0.268	(-0.3459 ; 0.7461)
Method: Panel	0.188	0.093	2.030	0.206	(-0.211 ; 0.5879)

We have the following results for the meta-regressions of Log of Expenditure Per Capita:

1. Unpublished papers report significantly higher coefficients.
2. In terms of the modeling, passing from OLS to PANEL increases the detected effects.
3. All other coefficients remained insignificant.

Below we also run the meta-regressions adding all coefficients in the papers.

```
mod <- rma(yi = coef,
  sei = SE,
  data = fulldat,
  method = "REML",
  mods = ~ indepvar2+year+published+elecsys2+method,
  test = "knha",
  subset = fulldat$depvar2=='logExpPC',
  slab = fulldat$authoryear)
```

```
mod6 <- tibble(
  ` ` = c("Intercept",
    "Indepvar: N",
    "Year",
    "Published: No",
    "Method: Panel",
    "Method: RDD"),
  Estimate = mod[["beta"]],
  SE = mod[["se"]],
  `T` = mod[["zval"]],
  `P-Value` = mod[["pval"]],
  CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
    round(mod[["ci.ub"]], digits = 4), ")"),
  model = "logExpPC - All coeffs"
) %>%
  mutate_if(is.numeric, list(~round(., digits = 3)))
```

```
summary(mod)
```

```
##
## Mixed-Effects Model (k = 27; tau^2 estimator: REML)
##
##   logLik  deviance      AIC      BIC     AICc
##  23.4038  -46.8076  -32.8076  -25.4959  -24.1922
##
```

```

## tau^2 (estimated amount of residual heterogeneity):      0.0049 (SE = 0.0019)
## tau (square root of estimated tau^2 value):             0.0703
## I^2 (residual heterogeneity / unaccounted variability): 86.48%
## H^2 (unaccounted variability / sampling variability):    7.40
## R^2 (amount of heterogeneity accounted for):             83.01%
##
## Test for Residual Heterogeneity:
## QE(df = 21) = 102.3018, p-val < .0001
##
## Test of Moderators (coefficients 2:6):
## F(df1 = 5, df2 = 21) = 20.8366, p-val < .0001
##
## Model Results:
##
##           estimate      se    tval    pval
## intrcpt      -8.2960  12.0371  -0.6892  0.4982
## indepvar2N   -0.0577   0.0733  -0.7879  0.4396
## year          0.0042   0.0060   0.7070  0.4874
## publishedNo   0.1030   0.0651   1.5833  0.1283
## methodPANEL  -0.2524   0.0682  -3.6997  0.0013
## methodRDD    -0.2851   0.0617  -4.6212  0.0001
##           ci.lb    ci.ub
## intrcpt     -33.3285  16.7365
## indepvar2N   -0.2102   0.0947
## year         -0.0082   0.0167
## publishedNo  -0.0323   0.2383
## methodPANEL  -0.3942  -0.1105  **
## methodRDD    -0.4134  -0.1568  ***
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

```

mod <- permutest(mod, progbar = F)

mod

##

## Test of Moderators (coefficients 2:6):

## F(df1 = 5, df2 = 21) = 20.8366, p-val* = 0.0010

##

## Model Results:

##

##           estimate      se    tval   pval*
## intrcpt      -8.2960  12.0371  -0.6892  0.5160
## indepvar2N   -0.0577   0.0733  -0.7879  0.4460
## year          0.0042   0.0060   0.7070  0.4980
## publishedNo   0.1030   0.0651   1.5833  0.1360
## methodPANEL  -0.2524   0.0682  -3.6997  0.0010
## methodRDD    -0.2851   0.0617  -4.6212  0.0010
##
##           ci.lb    ci.ub
## intrcpt    -33.3285  16.7365
## indepvar2N  -0.2102   0.0947
## year        -0.0082   0.0167
## publishedNo -0.0323   0.2383
## methodPANEL -0.3942  -0.1105 ***
## methodRDD   -0.4134  -0.1568 ***

##

## ---

## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

```

mod6_permu <- tibble(
  ` ` = c("Intercept",
    "Indepvar: N",
    "Year",
    "Published: No",
    "Method: Panel",
    "Method: RDD"),

```

```

Estimate = mod[["beta"]],
SE = mod[["se"]],
`T` = mod[["zval"]],
`P-Value` = mod[["pval"]],
CI = paste0("(", round(mod[["ci.lb"]], digits = 4), " ; ",
            round(mod[["ci.ub"]], digits = 4), ")"),
model = "logExpPC - All coefs - Permutation"
) %>%

mutate_if(is.numeric, list(~round(., digits = 3)))

```

With all coefficients, the results of the effect sizes on the Log of Expenditure Per Capita Regressions are the following:

1. In terms of the modeling, passing from OLS to PANEL or RDD decreases the detected effects.
2. All other coefficients remained insignificant.

```

coefs <- tibble(
  ` ` = c("Intercept",
          "Indepvar: N",
          "Indepvar: logN",
          "Year",
          "Published: No",
          "Elecsys: Non-Majoritarian",
          "Method: Panel",
          "Method: IV",
          "Method: RDD"))

aux = list(mod3, mod5, mod1) %>%

map(~{
  left_join(coefs, .x) %>%
    mutate(model = ifelse(is.na(model), model[1], model))
}) %>%

reduce(bind_cols) %>%

mutate(
  order = ifelse(` ` == "Intercept", 1, NA),

```

```

order = ifelse(` ` == "Indepvar: N", 2, order),
order = ifelse(` ` == "Indepvar: logN", 3, order),
order = ifelse(` ` == "Year", 4, order),
order = ifelse(` ` == "Published: No", 5, order),
order = ifelse(` ` == "Elecsys: Non-Majoritarian", 6, order),
order = ifelse(` ` == "Method: Panel", 7, order),
order = ifelse(` ` == "Method: IV", 8, order),
order = ifelse(` ` == "Method: RDD", 9, order)) %>%

filter(` ` != "Method: RDD") %>%

arrange(order) %>%

mutate_at(vars(contains("Value")), list(
  ~(ifelse(is.na(.), " ", ifelse(. < 0.001, "***",
    ifelse(. < 0.01, "**",
      ifelse(. < 0.05, "*",
        ifelse(. >= 0.05, " ",
          ifelse(is.na(.), " ", .)))))))))) %>%

mutate_at(vars(contains("Estimate") | contains("SE")),
  list(~(ifelse(is.na(.), " ", .))) %>%

select(-` 1`, -` 2`, -contains("model"), -order) %>%

mutate(mod = "Base") %>%

select(mod, everything())

aux2 = list(mod4, mod6, mod2) %>%

map(~{
  left_join(coefs, .x) %>%
  mutate(model = ifelse(is.na(model), model[1], model))
}) %>%

reduce(bind_cols) %>%

mutate(
  order = ifelse(` ` == "Intercept", 1, NA),
  order = ifelse(` ` == "Indepvar: N", 2, order),
  order = ifelse(` ` == "Indepvar: logN", 3, order),
  order = ifelse(` ` == "Year", 4, order),

```

```

order = ifelse(` ` == "Published: No", 5, order),

order = ifelse(` ` == "Elecsys: Non-Majoritarian", 6, order),

order = ifelse(` ` == "Method: Panel", 7, order),

order = ifelse(` ` == "Method: IV", 8, order),

order = ifelse(` ` == "Method: RDD", 9, order)) %>%

arrange(order) %>%

mutate_at(vars(contains("Value")), list(

  ~(ifelse(is.na(.), " ", ifelse(. < 0.001, "***",

    ifelse(. < 0.01, "**",

      ifelse(. < 0.05, "*",

        ifelse(. >= 0.05, " ",

          ifelse(is.na(.), " ", .)))))))))) %>%

mutate_at(vars(contains("Estimate") | contains("SE")),

  list(~ifelse(is.na(.), " ", .))) %>%

select(-`1`, -`2`, -contains("model"), -order) %>%

mutate(mod = "All Coefs") %>%

select(mod, everything())

aux <- bind_rows(aux, aux2) %>%

select(-contains("CI"), -`T`, -T1, -T2)

```

J Robustness: Full Model Meta-Regressions Combined

In this section, we aggregate all the coefficients and run a multivariate meta-regression, controlling for:

1. The type of the dependent variable in the study (expenditure per capita, log of the expenditure per capita, and share of government expenditure in the GDP)
2. The type of the independent variable in the study (N , K , $\log(N)$);
3. The electoral system (Majoritarian, Proportional Representation, and Mixed).

The results follow below, and show null effects for all variables, including the intercept.

```

mod <- rma(yi = coef,

  sei = SE,

  data = dat,

```



```

method = "REML",

mods = ~ depvar2+indepvar2+year+published+elecsys2+method,

test = "knha")

```

```
summary(mod)
```

```

##

## Mixed-Effects Model (k = 36; tau^2 estimator: REML)

##

##   logLik  deviance      AIC      BIC      AICc
## -51.1108  102.2217  126.2217  140.8482  152.2217

##

## tau^2 (estimated amount of residual heterogeneity):      0.2752 (SE = 0.1068)
## tau (square root of estimated tau^2 value):              0.5246
## I^2 (residual heterogeneity / unaccounted variability): 99.97%
## H^2 (unaccounted variability / sampling variability):    3029.13
## R^2 (amount of heterogeneity accounted for):              0.00%

##

## Test for Residual Heterogeneity:

## QE(df = 25) = 201.5647, p-val < .0001

##

## Test of Moderators (coefficients 2:11):

## F(df1 = 10, df2 = 25) = 0.2816, p-val = 0.9796

##

## Model Results:

##

##               estimate      se    tval
## intrcpt          -62.7057  87.3038  -0.7182
## depvar2PCTGDP         0.0662  0.6243   0.1060
## depvar2logExpPC       -0.3290  0.6889  -0.4776
## indepvar2N           -0.5079  0.5344  -0.9503
## indepvar2logN         0.4261  0.7846   0.5431
## year                0.0313  0.0434   0.7216
## publishedNo          0.3589  0.6034   0.5947

```

```
## elecsys2Non-Majoritarian -0.1527 0.5199 -0.2936
## methodPANEL 0.3697 0.5768 0.6409
## methodIV 0.3435 0.8881 0.3868
## methodRDD 0.4214 0.8556 0.4925
## pval ci.lb ci.ub
## intrcpt 0.4793 -242.5111 117.0998
## depvar2PCTGDP 0.9164 -1.2196 1.3520
## depvar2logExpPC 0.6371 -1.7479 1.0899
## indepvar2N 0.3511 -1.6085 0.5928
## indepvar2logN 0.5919 -1.1897 2.0419
## year 0.4772 -0.0581 0.1207
## publishedNo 0.5574 -0.8840 1.6017
## elecsys2Non-Majoritarian 0.7715 -1.2234 0.9181
## methodPANEL 0.5274 -0.8183 1.5577
## methodIV 0.7022 -1.4856 2.1726
## methodRDD 0.6266 -1.3407 2.1835
##
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

As we have considerable heterogeneity in our sample, we run a permutation test to ensure the validity of our estimates. The results follow below.

```
permutest(mod, progbars = F)
```

```
##
## Test of Moderators (coefficients 2:11):
## F(df1 = 10, df2 = 25) = 0.2816, p-val* = 0.7400
##
## Model Results:
##
## estimate se
## intrcpt -62.7057 87.3038
## depvar2PCTGDP 0.0662 0.6243
## depvar2logExpPC -0.3290 0.6889
```

## indepvar2N	-0.5079	0.5344
## indepvar2logN	0.4261	0.7846
## year	0.0313	0.0434
## publishedNo	0.3589	0.6034
## elecsys2Non-Majoritarian	-0.1527	0.5199
## methodPANEL	0.3697	0.5768
## methodIV	0.3435	0.8881
## methodRDD	0.4214	0.8556
##	tval	pval*
## intrcpt	-0.7182	0.2990
## depvar2PCTGDP	0.1060	0.8810
## depvar2logExpPC	-0.4776	0.5070
## indepvar2N	-0.9503	0.1780
## indepvar2logN	0.5431	0.4530
## year	0.7216	0.2980
## publishedNo	0.5947	0.3790
## elecsys2Non-Majoritarian	-0.2936	0.6370
## methodPANEL	0.6409	0.3720
## methodIV	0.3868	0.5600
## methodRDD	0.4925	0.4710
##	ci.lb	ci.ub
## intrcpt	-242.5111	117.0998
## depvar2PCTGDP	-1.2196	1.3520
## depvar2logExpPC	-1.7479	1.0899
## indepvar2N	-1.6085	0.5928
## indepvar2logN	-1.1897	2.0419
## year	-0.0581	0.1207
## publishedNo	-0.8840	1.6017
## elecsys2Non-Majoritarian	-1.2234	0.9181
## methodPANEL	-0.8183	1.5577
## methodIV	-1.4856	2.1726
## methodRDD	-1.3407	2.1835
##		

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

In the main text, we selected coefficients based on the regressions that had most observations and presented a full model (with fixed effects or intermediate bandwidth in RDD). We also run the meta-regressions adding all coefficients in the papers. The results follow below:

```
mod <- rma(yi = coef,
           sei = SE,
           data = fulldat,
           method = "REML",
           mods = ~ depvar2+indepvar2+year+published+elecsys2+method,
           test = "knha")
```

```
summary(mod)
```

```
##
```

```
## Mixed-Effects Model (k = 128; tau^2 estimator: REML)
```

```
##
```

```
##      logLik   deviance      AIC      BIC
```

```
## -211.8306   423.6613   447.6613   480.8073
```

```
##      AICc
```

```
##  450.6613
```

```
##
```

```
## tau^2 (estimated amount of residual heterogeneity):      0.1613 (SE = 0.0264)
```

```
## tau (square root of estimated tau^2 value):              0.4016
```

```
## I^2 (residual heterogeneity / unaccounted variability): 99.99%
```

```
## H^2 (unaccounted variability / sampling variability):     7225.40
```

```
## R^2 (amount of heterogeneity accounted for):              13.60%
```

```
##
```

```
## Test for Residual Heterogeneity:
```

```
## QE(df = 117) = 2134.3400, p-val < .0001
```

```
##
```

```
## Test of Moderators (coefficients 2:11):
```

```
## F(df1 = 10, df2 = 117) = 1.2469, p-val = 0.2689
```

```
##

## Model Results:

##
##              estimate      se
## intrcpt      -39.7005  39.1677
## depvar2PCTGDP  -0.0842  0.2622
## depvar2logExpPC -0.4715  0.2675
## indepvar2N     -0.2192  0.2088
## indepvar2logN   0.6596  0.3602
## year           0.0199  0.0195
## publishedNo     -0.0751  0.2327
## elecsys2Non-Majoritarian 0.1848  0.2468
## methodPANEL    -0.1587  0.2112
## methodIV       -0.0272  0.3285
## methodRDD       0.0134  0.3158
##
##              tval    pval
## intrcpt      -1.0136  0.3129
## depvar2PCTGDP -0.3213  0.7486
## depvar2logExpPC -1.7625  0.0806
## indepvar2N     -1.0494  0.2962
## indepvar2logN   1.8312  0.0696
## year           1.0213  0.3092
## publishedNo     -0.3226  0.7476
## elecsys2Non-Majoritarian 0.7486  0.4556
## methodPANEL    -0.7515  0.4538
## methodIV       -0.0827  0.9342
## methodRDD       0.0423  0.9663
##
##              ci.lb    ci.ub
## intrcpt      -117.2702  37.8691
## depvar2PCTGDP  -0.6034  0.4350
## depvar2logExpPC -1.0014  0.0583 .
## indepvar2N     -0.6328  0.1944
## indepvar2logN   -0.0538  1.3730 .
```

```
## year                -0.0187   0.0585
## publishedNo         -0.5360   0.3858
## elecsys2Non-Majoritarian -0.3041   0.6736
## methodPANEL         -0.5769   0.2595
## methodIV            -0.6778   0.6235
## methodRDD           -0.6121   0.6388
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
permutest(mod, progbars = F)
```

```
##
## Test of Moderators (coefficients 2:11):
## F(df1 = 10, df2 = 117) = 1.2469, p-val* = 0.0030
##
## Model Results:
##
##               estimate      se
## intrcpt       -39.7005  39.1677
## depvar2PCTGDP   -0.0842   0.2622
## depvar2logExpPC -0.4715   0.2675
## indepvar2N      -0.2192   0.2088
## indepvar2logN    0.6596   0.3602
## year           0.0199   0.0195
## publishedNo     -0.0751   0.2327
## elecsys2Non-Majoritarian 0.1848   0.2468
## methodPANEL     -0.1587   0.2112
## methodIV        -0.0272   0.3285
## methodRDD       0.0134   0.3158
##
##               tval    pval*
## intrcpt       -1.0136  0.1390
## depvar2PCTGDP -0.3213  0.6100
## depvar2logExpPC -1.7625  0.0140
```

```
## indepvar2N          -1.0494  0.1290
## indepvar2logN       1.8312  0.0110
## year               1.0213  0.1350
## publishedNo        -0.3226  0.6330
## elecsys2Non-Majoritarian  0.7486  0.2730
## methodPANEL        -0.7515  0.2640
## methodIV           -0.0827  0.8850
## methodRDD          0.0423  0.9510
##
##                  ci.lb   ci.ub
## intrcpt           -117.2702  37.8691
## depvar2PCTGDP      -0.6034  0.4350
## depvar2logExpPC    -1.0014  0.0583  *
## indepvar2N         -0.6328  0.1944
## indepvar2logN      -0.0538  1.3730  *
## year              -0.0187  0.0585
## publishedNo        -0.5360  0.3858
## elecsys2Non-Majoritarian -0.3041  0.6736
## methodPANEL        -0.5769  0.2595
## methodIV           -0.6778  0.6235
## methodRDD          -0.6121  0.6388
##
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

K Auxiliary Functions

K.1 Function to Generate Meta-Analytic Figures

This function receives the meta-analysis results and builds a forest plot using `ggplot2`.

```
# Build plot function for forest plots

build_forest <- function(mod, capt, lsize = 22, ttl = NULL) {

  # Build dataset for plot

  mod2 <- tibble(

    TE = mod$TE,
```

```

seTE = mod$seTE,

studlab = mod$studlab,

lower = mod$lower,

upper = mod$upper,

group = "A") %>%

bind_rows(.,

  aux = tibble(

    TE = c(mod$TE.random, NA),

    seTE = c(mod$seTE.random, NA),

    studlab = c("Overall Effect", "Prediction Interval"),

    lower = c(mod$lower.random, mod$lower.predict),

    upper = c(mod$upper.random, mod$upper.predict),

    group = "B")) %>%

group_by(studlab) %>%

mutate(studlab2 = paste0(studlab, "_", 1:n())) %>%

ungroup()

# Graph limits

limg <- max(abs(c(mod2$lower, mod2$upper)))

# Build plot

p <- mod2 %>%

ggplot(aes(y = reorder(studlab2, TE),

  x = TE, xmin = lower, xmax = upper)) +

geom_point(aes(color = group)) +

geom_errorbarh(aes(color = group), height = 0.1) +

scale_color_manual(values = c("#000000", "#8b0000")) +

scale_x_continuous(limits=c(-1.1*limg, 1.1*limg)) +

scale_y_discrete(

  labels = function(x) str_replace(x, "_[0-9]*$", "")) +

geom_vline(xintercept = 0,

  color = "#000000", linetype = "dashed") +

labs(x = "",

```



```

    y = "") +

  facet_grid(group~., scales = "free", space = "free") +

  labs(caption = capt,

       title = ttl) +

  theme_minimal() %+replace%

  theme(strip.text.y = element_blank(),

        legend.position = "none",

        axis.text.y = element_text(size = .8 * lsize,

                                     hjust = 1),

        axis.text.x = element_text(size = .6 * lsize,

                                     hjust = 1.1),

        plot.caption = element_text(size = lsize),

        plot.title.position = "plot",

        plot.title = element_text(hjust = 0.5,

                                    face = "bold",

                                    margin = margin(0, 0, 10, 0)),

        panel.grid.major = element_blank())

  return(p)
}

```

K.2 Webscraping Code

```

#### PACKAGES CONFIGURATION ####

# Needed packages

pkgs <- c("tidyverse", "rvest", "RSelenium")

# Install if not already installed

installIfNot <- function(x) {

  if (x %in% rownames(installed.packages()) == FALSE)

    install.packages(x, dependencies = T,

                     repos = "http://cran.us.r-project.org")

}

```

```

lapply(pkgs, installIfNot)

# Load packs

lapply(pkgs, require, character.only = T)

rm(pkgs, installIfNot)

#### SETTING UP SELENIUM ####

# Alternative 1: Setting up Selenium (head)

rsD <- rsDriver(port = 1114L, browser = c("firefox"))

remDr <- rsD$client

remDr$open()

#### Google scholar ####

# site: https://scholar.google.com

remDr$navigate("https://scholar.google.com/

               scholar?cites=13117579863846712459&as_sdt=2005&scioldt=0,5")

articles_weingast <- tibble(

  value = NA,

  term = NA,

  page = NA

)

k <- 0

for (j in 1:213) { # we had to manually choose the number of pages here

  Sys.sleep(rpois(1, 5))

  # Getting articles basic information

  k <- k + 1

```

```

webElem <- remDr$findElement("css", "body")

title <- read_html(remDr$getPageSource()[[1]]) %>%

  html_nodes(
    xpath = '//*[contains(concat( " ", @class, " " ),
      concat( " ", "gs_rt", " " ))]')
  ) %>%

  html_text() %>%

  enframe(name = NULL) %>%

  rename("title" = "value") %>%

  mutate(page = k)

articles_partial <- read_html(remDr$getPageSource()[[1]]) %>%

  html_nodes(xpath = '//*[contains(concat( " ", @class, " " ), concat( " ", "gs_a", " " ))]') %>%

  html_text() %>%

  enframe(name = NULL) %>%

  bind_rows(
    tibble(
      value = "delete",
      term = NA,
      page = NA
    ),
    .
  ) %>%

  bind_cols(., title)

# Binding articles
articles_weingast <- bind_rows(articles_weingast, articles_partial)

# Changing Pages

```

```

next_button <- remDr$findElement(using = "xpath", "/html/body/div/div[11]/div[2]/div[2]/div[3]/
                                div[2]/center/table/tbody/tr/td[12]/a/b")

next_button$clickElement()

# Deleting cookies

remDr$deleteAllCookies()

}

write_csv(articles_weingast, "scholar_weingast_raw.csv")

articles_weingast <- articles_weingast %>%

  select(-term, -page) %>%

  filter(value != "delete") %>%

  slice(2:nrow(.)) %>%

  filter(!grepl("books.google.com", value)) %>%

  filter(!grepl("BOOK", title)) %>%

  separate(., value, into = c("author", "value"),
           sep = " -", remove = T, extra = "merge", fill = "right") %>%

  separate(., value, into = c("journal", "year"),
           sep = ",", remove = T, extra = "merge", fill = "right") %>%

  mutate(

    year_2 = ifelse(is.na(year), journal, year),

    journal = ifelse(is.na(year), NA, journal),

    year = year_2

  ) %>%

  select(-year_2) %>%

  separate(., year, into = c("year", "site"),
           sep = "-", remove = T, extra = "merge", fill = "right") %>%

  mutate(

    year = gsub("[^0-9 ]", "", value),

    year = gsub("[0-9]{5,}", "", year),

    year = gsub("[^0-9]", "", year)

  ) %>%

  separate(., year, into = c("year", "junk"),

```

```

      sep = " ", remove = T, extra = "merge", fill = "right") %>%
separate(., value, into = c("journal", "value"),
      sep = "-", remove = T, extra = "merge", fill = "right") %>%

mutate(
  journal_untidy = year,
  year = gsub("[^0-9 ]", "", year)
)

articles <- articles %>%
  na.omit() %>%
  distinct(., value, .keep_all = T)

write_csv(articles, "google_scholar_clean.csv")

#### SCOPUS ####

# url: https://www.scopus.com/home.uri

# Scraping Scopus requires a bit more manual labor.
# You can login on scopus through your university/institution
# and download the metadata of the article(s) you
# want directly from there.
# All we need to do after that is scrape the information of
# every link from the .csv file downloaded previously

scopus <- read_csv("scopus.csv")

scopus <- scopus %>%
  mutate(article = map_chr(Link, ~ {
    remDr$navigate(.x)
    read_html(remDr$getPageSource()[[1]]) %>%
      html_nodes(xpath = '//*[@id = "abstractSection"]//p') %>%
      html_text() %>%

```

```

paste(., collapse = "\r\n")
))) %>%
mutate(
  article = gsub(
    '\r\n\nUse this section.*Topics\n\n\n"',
    "",
    article
  ),
  article = gsub(
    "Topics are unique.*onwards.",
    "",
    article
  ),
  article = gsub(
    "Use this section.*documents.",
    "",
    article
  ),
  article = gsub(
    "Learn more about these Topics",
    "",
    article
  ),
  article = gsub(
    ". 20.*, Springer Science\\+Business Media, LLC, part of Springer Nature.",
    "",
    article
  ),
  article = gsub("\r", "", article),
  article = gsub("\n", "", article),
  article = gsub(" {2,}", " ", article),
  article = gsub(" {3,}", "", article)
)

```

```

write_csv(scopus, "scopus_clean.csv")

#### Microsoft Academic ####

# url: https://academic.microsoft.com/home
articles <- list()
k <- 0
remDr$navigate("https://academic.microsoft.com/paper/
                2076316673/citedby/search?q=The%20Political%
                20Economy%20of%20Benefits%20and%20Costs%3A%20A%
                20Neoclassical%20Approach%20to
                %20Distributive%20Politics&qe=RIId%253D2076316673&f=&orderBy=0")

# 1) Getting hyperlinks from articles
for (j in 1:100) {
  k <- k + 1
  print(k)

  # Navigating Website

  Sys.sleep(rpois(2, 5))

  # Getting articles' links
  articles[[k]] <- read_html(remDr$pageSource()[[1]]) %>%
    html_nodes(xpath = "//a") %>%
    html_attr("href") %>%
    enframe(name = NULL) %>%
    filter(
      grepl("paper/", value),
      !grepl("citedby", value)
    ) %>%

```

```

mutate(value = paste0("https://academic.microsoft.com/", value))

# Changing Page
next_page_others <- remDr$findElement(using = "xpath", "/html/body/div/div/div/router-view/router-view/
compose/div/div[2]/ma-pager/div/i[2]")
next_page_others$clickElement()
}

articles <- articles %>%
  reduce(bind_rows) %>%
  distinct(value, .keep_all = T)

# 2) Navigating through articles and scraping them

articles_links <- articles %>%
  mutate(
    abstract = NA,
    title = NA,
    year = NA,
    journal = NA,
    authors = NA,
    tags = NA
  )

for (i in 1:nrow(articles_links)) {
  remDr$navigate(articles_links$value[i])
  Sys.sleep(rpois(1, 4))

  articles_links$abstract[i] <- read_html(remDr$getPageSource()[[1]]) %>%
    html_nodes(., xpath = "//html/body/div/div/div/router-view/compose[1]/
      div/div/ma-entity-detail-info/compose/div/div/div[1]/p") %>%
    html_text() %>%

```



```

paste(., collapse = " ")

articles_links$title[i] <- read_html(remDr$getPageSource()[[1]]) %>%
  html_nodes(.,
    xpath = '//*[contains(concat( " ", @class, " " ),
      concat( " ", "name", " " ))]') %>%
  html_text() %>%
  paste(., collapse = " ")

articles_links$year[i] <- read_html(remDr$getPageSource()[[1]]) %>%
  html_nodes(.,
    xpath = '//*[contains(concat( " ", @class, " " ),
      concat( " ", "name-section", " " ))]
      /*[contains(concat( " ", @class, " " ), concat( " ", "year", " " ))]') %>%
  html_text() %>%
  paste(., collapse = " ")

articles_links$journal[i] <- read_html(remDr$getPageSource()[[1]]) %>%
  html_nodes(., xpath = '//*[contains(concat( " ", @class, " " )
    , concat( " ", "pub-name", " " ))]') %>%
  html_text() %>%
  paste(., collapse = " ")

articles_links$authors[i] <- read_html(remDr$getPageSource()[[1]]) %>%
  html_nodes(., xpath = "/html/body/div/div/div/router-view/compose[1]/div/div/
    ma-entity-detail-info/compose/div/div/div[1]/
    ma-author-string-collection") %>%
  html_text() %>%
  paste(., collapse = " ")

articles_links$tags[i] <- read_html(remDr$getPageSource()[[1]]) %>%
  html_nodes(., xpath = "/html/body/div/div/div/router-view/compose[1]/
    div/div/ma-entity-detail-info/compose/

```

```


## L Session Information



```

sessionInfo()

R version 4.0.0 (2020-04-24)
Platform: x86_64-apple-darwin17.0 (64-bit)
Running under: macOS Mojave 10.14.6
##
Matrix products: default
BLAS: /Library/Frameworks/R.framework/Versions/4.0/Resources/lib/libRblas.dylib
LAPACK: /Library/Frameworks/R.framework/Versions/4.0/Resources/lib/libRlapack.dylib
##
locale:
[1] en_US.UTF-8/en_US.UTF-8/en_US.UTF-8/C/en_US.UTF-8/en_US.UTF-8
##
attached base packages:
[1] grid parallel stats graphics grDevices
[6] utils datasets methods base
##
other attached packages:
[1] magick_2.3 kableExtra_1.1.0
[3] ggpubr_0.2.5 magrittr_1.5
[5] gridExtra_2.3 gridGraphics_0.5-0
[7] knitr_1.28 compareGroups_4.4.1
[9] SNPAssoc_1.9-2 mvtnorm_1.1-0
[11] survival_3.1-12 haplo.stats_1.7.9

```



114


```

```

## [13] readxl_1.3.1      metafor_2.4-0
## [15] Matrix_1.2-18     meta_4.11-0
## [17] forcats_0.5.0     stringr_1.4.0
## [19] dplyr_0.8.5       purrr_0.3.4
## [21] readr_1.3.1       tidyr_1.0.2
## [23] tibble_3.0.1      ggplot2_3.3.0
## [25] tidyverse_1.3.0   rmarkdown_2.1
## [27] nvimcom_0.9-88
##
## loaded via a namespace (and not attached):
## [1] TH.data_1.0-10     minqa_1.2.4
## [3] colorspace_1.4-1   ggsignif_0.6.0
## [5] ellipsis_0.3.0     flextable_0.5.9
## [7] htmlTable_1.13.3   base64enc_0.1-3
## [9] fs_1.4.1           rstudioapi_0.11
## [11] mice_3.8.0         farver_2.0.3
## [13] MatrixModels_0.4-1 fansi_0.4.1
## [15] lubridate_1.7.8    xml2_1.3.2
## [17] codetools_0.2-16   splines_4.0.0
## [19] Formula_1.2-3      jsonlite_1.6.1
## [21] nloptr_1.2.2.1     broom_0.5.6
## [23] cluster_2.1.0      dbplyr_1.4.3
## [25] png_0.1-7          compiler_4.0.0
## [27] httr_1.4.1         backports_1.1.6
## [29] assertthat_0.2.1   cli_2.0.2
## [31] acepack_1.4.1      htmltools_0.4.0
## [33] quantreg_5.55      tools_4.0.0
## [35] gtable_0.3.0       glue_1.4.0
## [37] tinytex_0.22       Rcpp_1.0.4.6
## [39] cellranger_1.1.0   vctrs_0.2.4
## [41] writexl_1.2        nlme_3.1-147
## [43] xfun_0.13          lme4_1.1-23
## [45] rvest_0.3.5        CompQuadForm_1.4.3

```

```

## [47] lifecycle_0.2.0      statmod_1.4.34
## [49] polspline_1.1.17     MASS_7.3-51.5
## [51] zoo_1.8-7            scales_1.1.0
## [53] hms_0.5.3            sandwich_2.5-1
## [55] SparseM_1.78         RColorBrewer_1.1-2
## [57] HardyWeinberg_1.6.3  yaml_2.2.1
## [59] gdtools_0.2.2        rms_5.1-4
## [61] rpart_4.1-15         latticeExtra_0.6-29
## [63] stringi_1.4.6        highr_0.8
## [65] checkmate_2.0.0      zip_2.0.4
## [67] boot_1.3-24          truncnorm_1.0-8
## [69] chron_2.3-55         systemfonts_0.2.0
## [71] rlang_0.4.5          pkgconfig_2.0.3
## [73] Rsolnp_1.16          evaluate_0.14
## [75] lattice_0.20-41      labeling_0.3
## [77] htmlwidgets_1.5.1    tidyselect_1.0.0
## [79] R6_2.4.1             generics_0.0.2
## [81] Hmisc_4.4-0          multcomp_1.4-13
## [83] DBI_1.1.0            pillar_1.4.3
## [85] haven_2.2.0          foreign_0.8-78
## [87] withr_2.2.0          nnet_7.3-13
## [89] modelr_0.1.6         crayon_1.3.4
## [91] uuid_0.1-4           officer_0.3.8
## [93] jpeg_0.1-8.1         data.table_1.12.8
## [95] webshot_0.5.2        reprex_0.3.0
## [97] digest_0.6.25        munsell_0.5.0
## [99] viridisLite_0.3.0

```