

# Local Linear Regression

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## Estimating the conditional variance by local linear regression

```
# Libraries
library(sm)
library(KernSmooth)
source("locpolreg.R")
```

## Aircraft Data and log transformation

```
# Load data
data(aircraft)
attach(aircraft)

lgPower <- log(Power)
lgSpan <- log(Span)
lgLength <- log(Length)
lgWeight <- log(Weight)
lgSpeed <- log(Speed)
lgRange <- log(Range)
```

## Estimating the conditional variance

Consider the heteroscedastic regression model

$$Y = m(x) + \sigma(x)\varepsilon = m(x) + \epsilon,$$

where  $E(\varepsilon) = 0$ ,  $\text{Var}(\varepsilon) = 1$  and  $\sigma^2(x)$  is an unknown function that gives the conditional variance of  $Y$  given that the explanatory variable is equal to  $x$ . Let us define  $Z = \log((Y - m(x))^2) = \log \varepsilon^2$  and  $\delta = \log \varepsilon^2$ . Then

$$Z = \log \sigma^2(x) + \delta,$$

and  $\delta = \log \varepsilon^2$  is a random variable with expected value close to 0 (observe that  $E(\log \varepsilon^2)$  is close to  $\log E(\varepsilon^2) = \log \text{Var}(\varepsilon) = \log 1 = 0$ , at least when  $\text{Var}(\varepsilon^2)$  is small) taking the role of *noise* in the regression of  $Z$  against  $x$  (that is,  $Z$  is the response variable and  $x$  is the predicting variable).

Given that the values of  $\varepsilon_i^2$  are not observable, a way to estimate the function  $\sigma^2(x)$  is as follows:

### Part 1 - With function loc.pol.reg

1. Fit a nonparametric regression to data  $(x_i, y_i)$  and save the estimated values  $\hat{m}(x_i)$ .

```
## Function in ATENA to get loo-cv
h.cv.gcv <- function(x,y,h.v = exp(seq(log(diff(range(x)))/20),
```

```

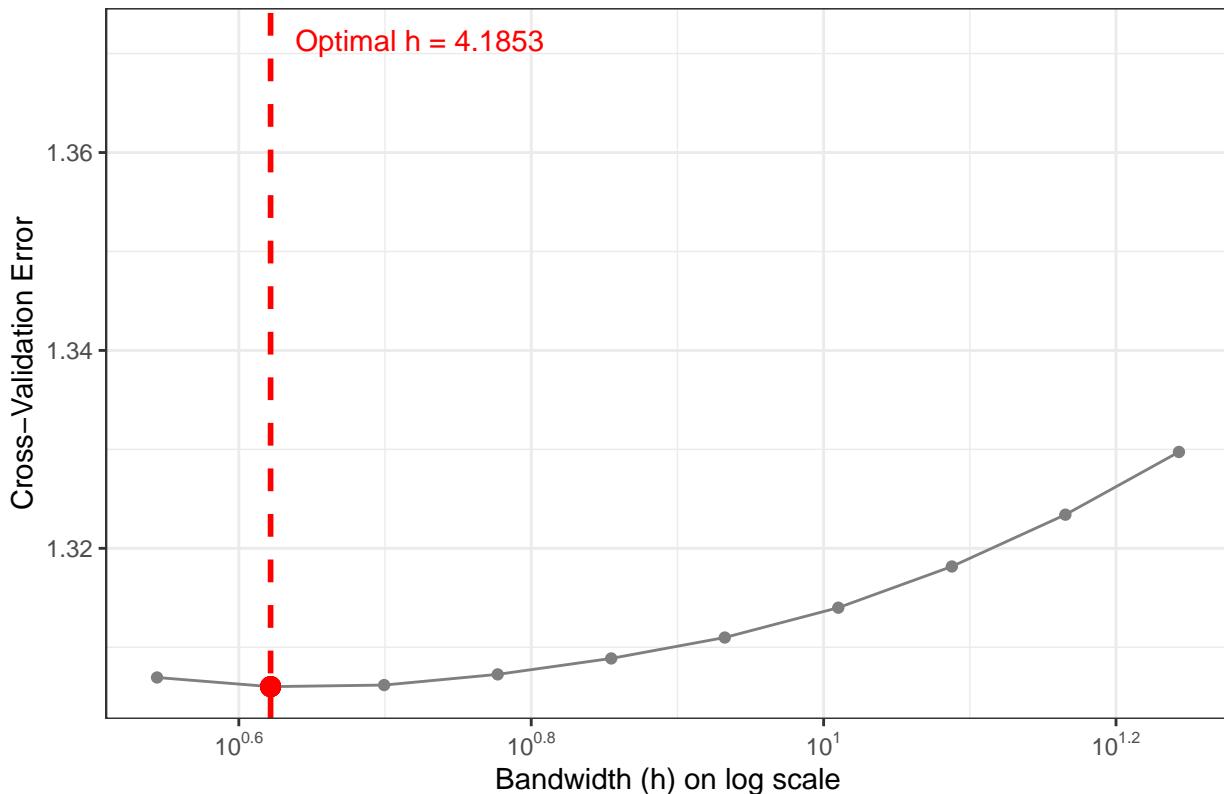
            log(diff(range(x))/4), l=10)),
p=1, type.kernel="normal"){
n <- length(x)
cv <- h.v*0
gcv <- h.v*0
for (i in 1:length(h.v)){
  h <- h.v[i]
  aux <- locpolreg(x=x, y=y, h=h, p=p, tg=x,
                     type.kernel=type.kernel, doing.plot=FALSE)
  S <- aux$S
  h.y <- aux$mtgr
  hii <- diag(S)
  av.hii <- mean(hii)
  cv[i] <- sum((y-h.y)/(1-hii))^2/n
  gcv[i] <- sum((y-h.y)/(1-av.hii))^2/n
}
return(list(h.v=h.v, cv=cv, gcv=gcv))
}

cv.res <- h.cv.gcv(Yr, lgWeight, type.kernel="normal")
h_cv <- cv.res$h.v[which.min(cv.res$cv)]
min_cv_error <- min(cv.res$cv)

```

## Bandwidth Selection via Leave-One-Out Cross-Validation

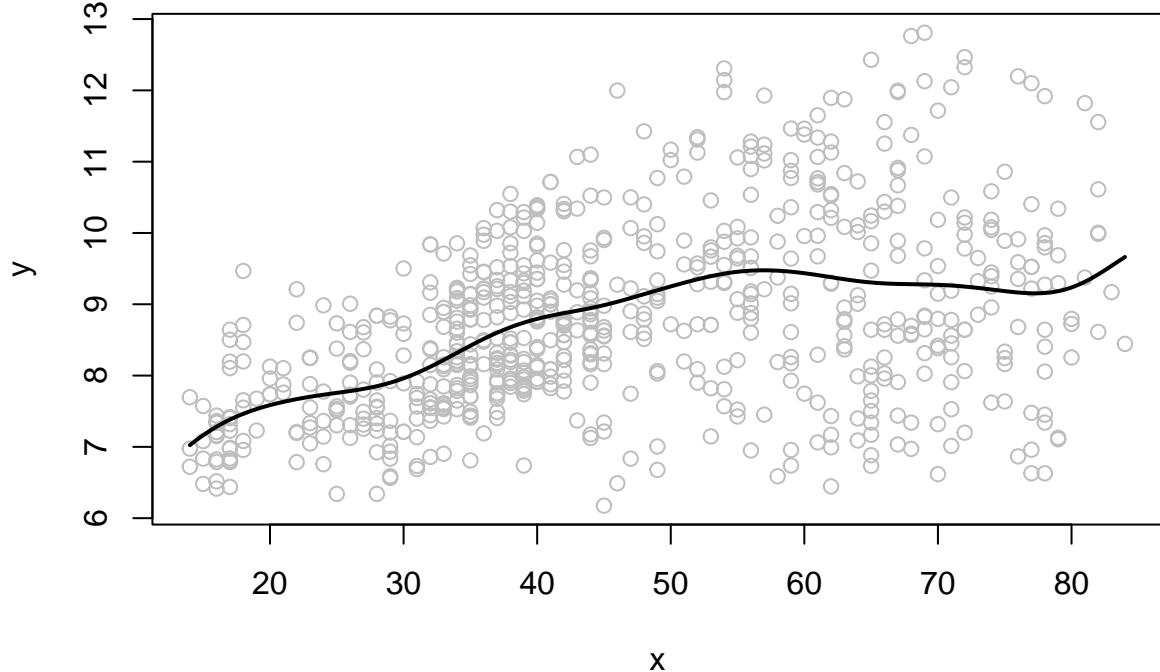
Bandwidth Selection via Leave–One–Out Cross–Validation



[1] “The optimal bandwidth selected by LOOCV is  $h = 4.1853$ .”

$$\hat{m}(x_i)$$

```
m_hat <- locpolreg(Yr, lgWeight, h=h_cv, q=1, tg=seq(min(Yr), max(Yr), length=200))
```

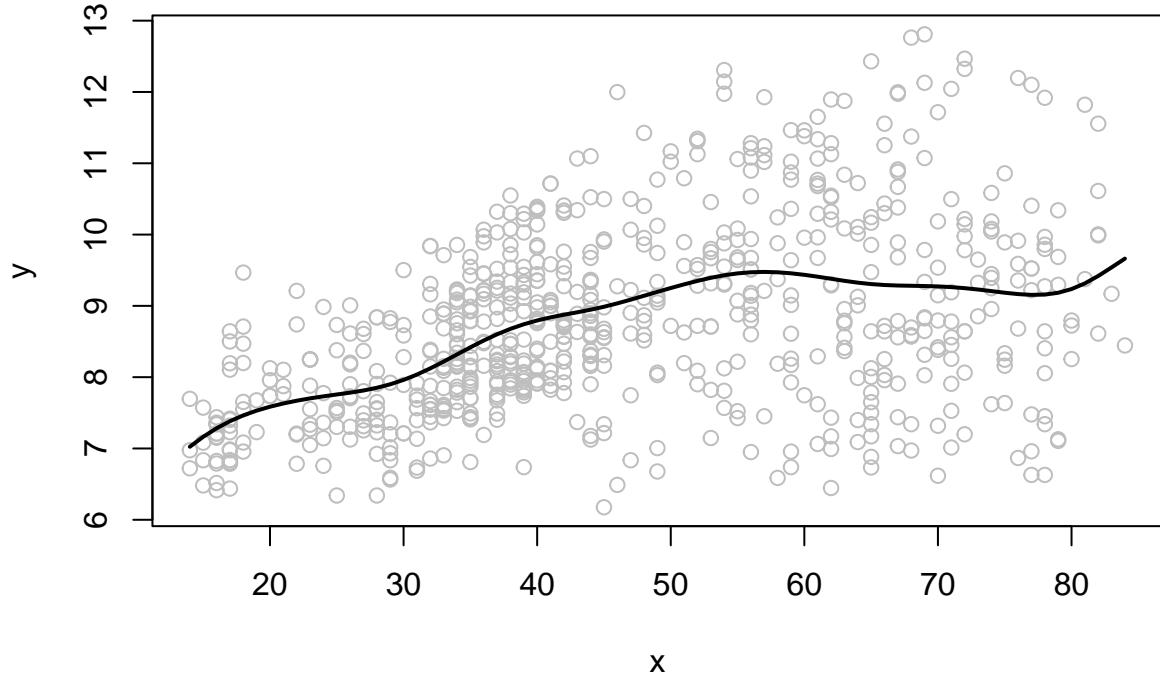


This plot displays the local linear regression fit (solid black line) over the raw data (grey circles). The smooth curve effectively captures the non-linear trend in the data.

**2. Transform the estimated residuals  $\hat{\epsilon} = y_i - \hat{m}(x_i)$ :**

$$z_i = \log(\hat{\epsilon}_i^2) = \log((y_i - \hat{m}(x_i))^2).$$

```
eps_hat <- lgWeight - locpolreg(Yr, lgWeight, h=h_cv, q=1, r=0, tg=Yr)$mtgr
```



```

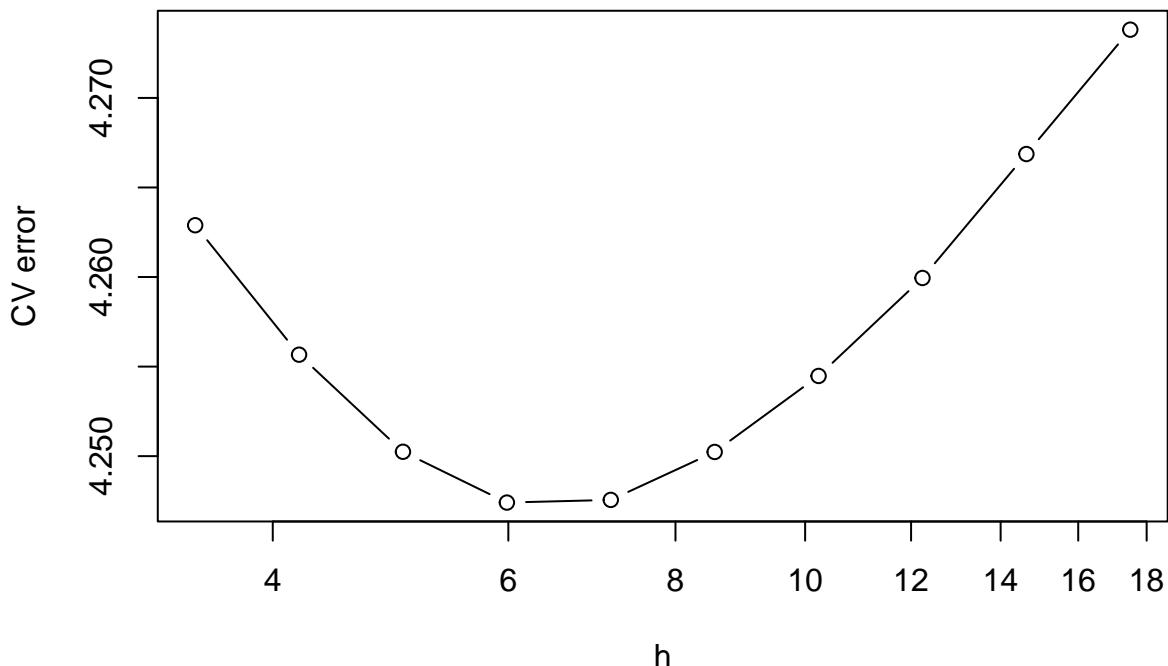
z <- log(eps_hat^2)

3. Fit a nonparametric regression to data  $(x_i, z_i)$  and call the estimated function  $\hat{q}(x)$ . Observe that  $\hat{q}(x)$  is an estimate of  $\log(\sigma^2(x))$ .
# Select bandwidth for  $z_i$  regression by LOOCV again
cv.res.var <- h.cv.gcv(Yr, z, type.kernel="normal")

# Plot CV curve for variance regression
plot(cv.res.var$h.v, cv.res.var$cv, type="b", log="x",
      main="LOOCV for log(residual^2)", xlab="h", ylab="CV error")

```

## LOOCV for $\log(\text{residual}^2)$



```

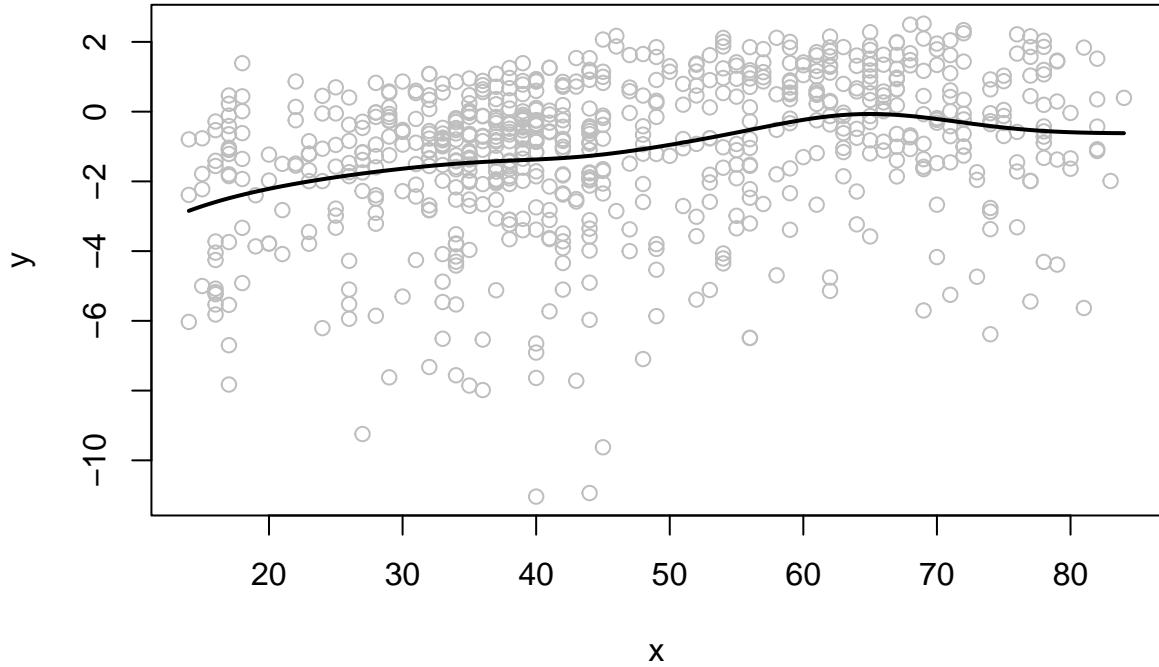
# Optimal bandwidth for  $\hat{q}(x)$ 
h_cv2 <- cv.res.var$h.v[which.min(cv.res.var$cv)]
h_cv2

## [1] 5.984916

# Define evaluation grid for smooth plots
tg <- seq(min(Yr), max(Yr), length=200)

# Fit  $\hat{q}(x) = E[\log(\text{residual}^2) | x]$ 
q_hat_result <- locpolreg(Yr, z, h=h_cv2, q=1, r=0, tg=tg, type.kernel="normal")

```



```
q_hat <- q_hat_result$mtgr
```

4. Estimate  $\sigma^2(x)$  by

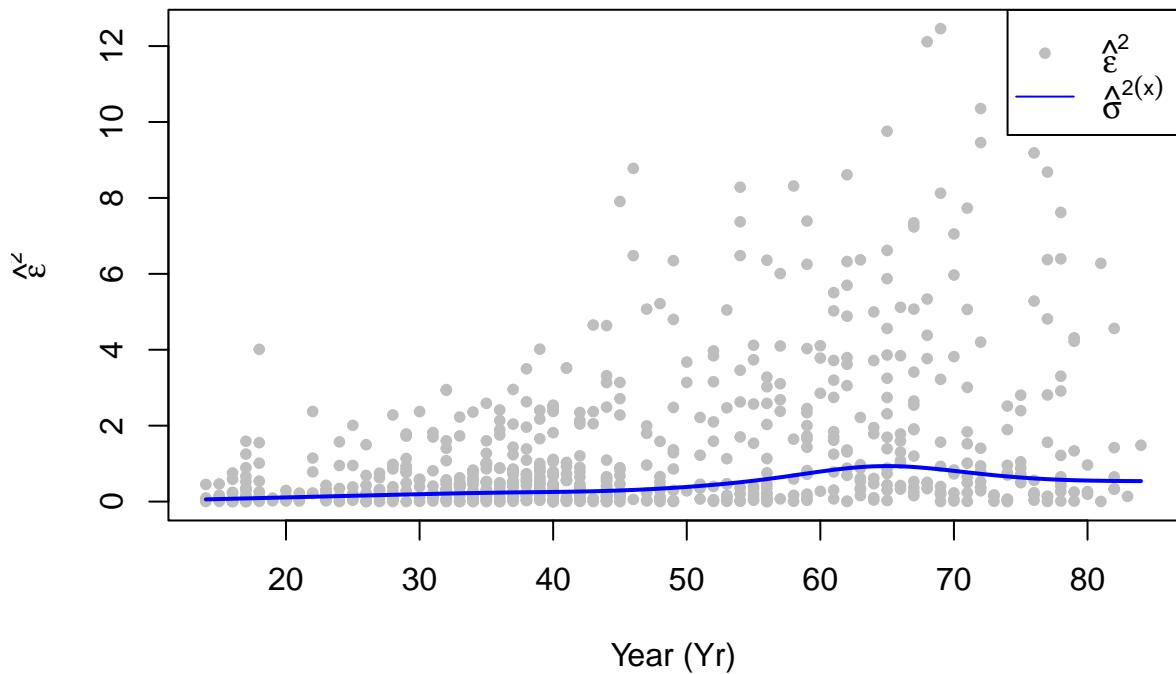
$$\hat{\sigma}^2(x) = e^{\hat{q}(x)}.$$

```
sigma2_hat <- exp(q_hat)
sigma_hat <- sqrt(sigma2_hat)
```

Apply this procedure to estimate the conditional variance of lgWeigth (variable  $Y$ ) given Yr (variable  $x$ ). Draw a graphic of  $\hat{\epsilon}_i^2$  against  $x_i$  and superimpose the estimated function  $\hat{\sigma}^2(x)$ . Lastly draw the function  $\hat{m}(x)$  and superimpose the bands  $\hat{m}(x) \pm 1,96\hat{\sigma}(x)$ .

```
plot(Yr, eps_hat^2, col="grey", pch=20,
      xlab="Year (Yr)", ylab=expression(hat(epsilon)^2),
      main=expression(paste(hat(epsilon)^2, " vs Yr with ", hat(sigma)^2(x))))
lines(tg, sigma2_hat, col="blue", lwd=2)
legend("topright", legend=c(expression(hat(epsilon)^2), expression(hat(sigma)^2(x))),
      col=c("grey","blue"), lty=c(NA,1), pch=c(20,NA))
```

## $\hat{\varepsilon}^2$ vs Yr with $\hat{\sigma}^{2(x)}$

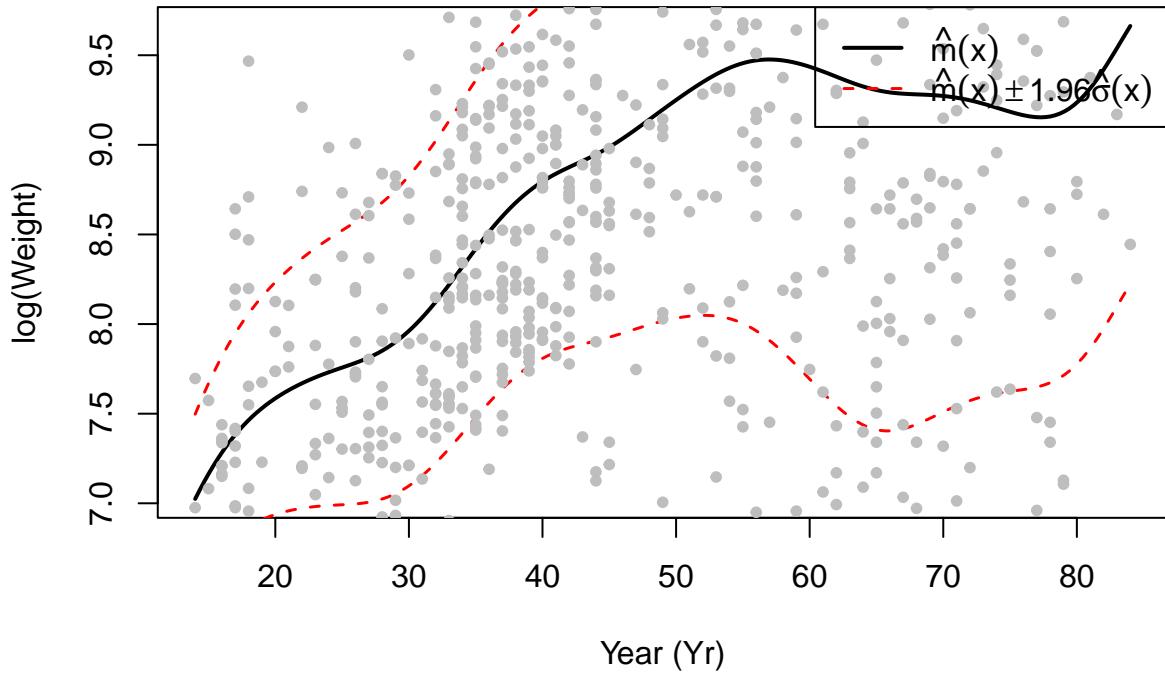


```

plot(tg, m_hat$mtgr, type="l", lwd=2, col="black",
      main=expression(paste(hat(m)(x), " ± 1.96 ", hat(sigma)(x))),
      xlab="Year (Yr)", ylab="log(Weight)")
lines(tg, m_hat$mtgr + 1.96*sigma_hat, col="red", lty=2, lwd=1.5)
lines(tg, m_hat$mtgr - 1.96*sigma_hat, col="red", lty=2, lwd=1.5)
points(Yr, lgWeight, col="grey", pch=20)
legend("topright",
      legend=c(expression(hat(m)(x)), expression(hat(m)(x) %+-% 1.96*hat(sigma)(x))),
      lty=c(1,2), col=c("black","red"), lwd=c(2,1.5))

```

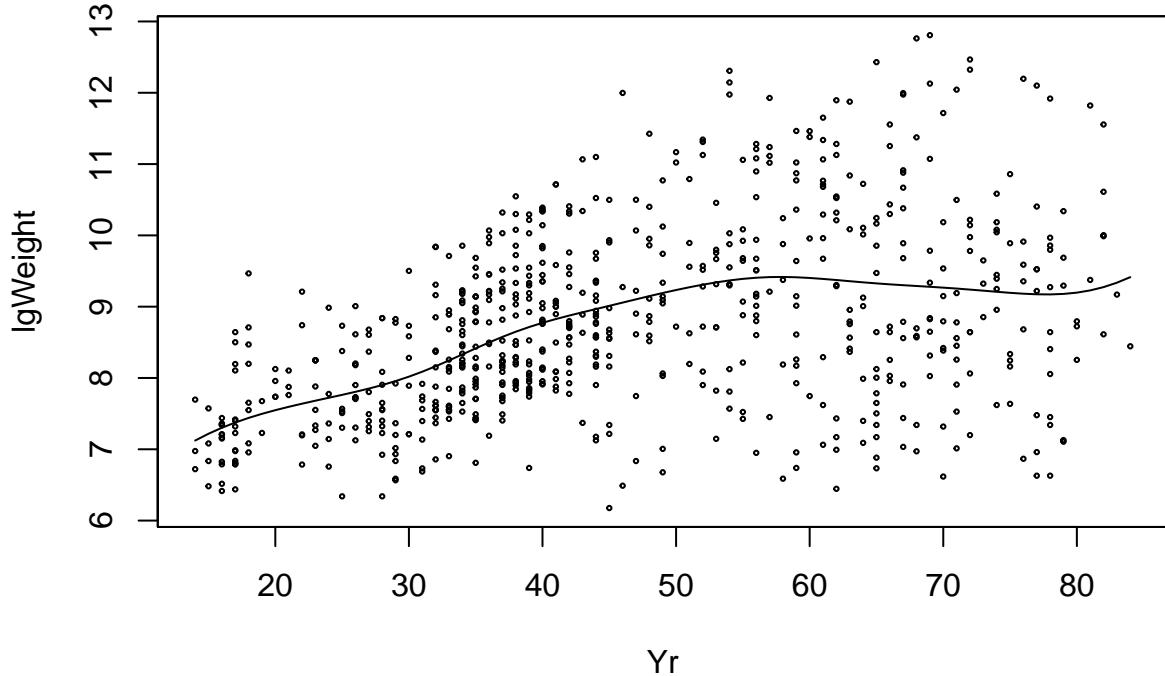
$$\hat{m}(x) \pm 1.96 \hat{\sigma}(x)$$



### With sm.regression

1. Fit a nonparametric regression to data  $(x_i, y_i)$  and save the estimated values  $\hat{m}(x_i)$ .

```
h1 <- dpill(Yr, lgWeight)
sm.regression(Yr, lgWeight, h=h1, eval.points=Yr, model="none") -> sm_m
```



```
m_hat_sm <- sm_m$estimate
```

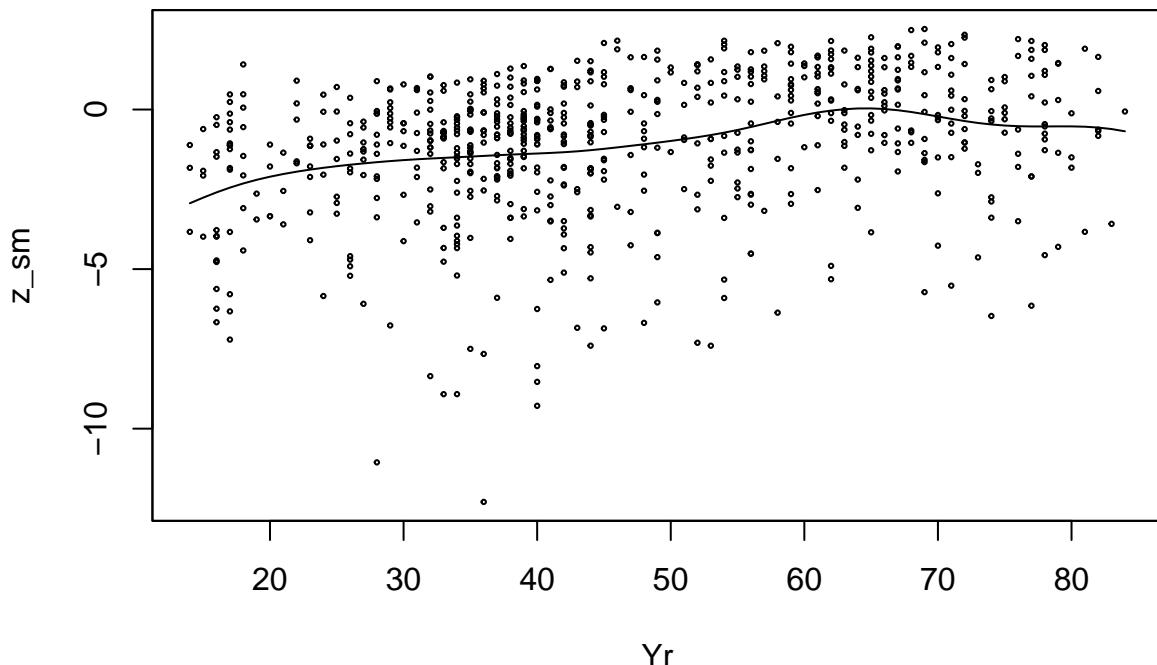
2. Transform the estimated residuals  $\hat{\epsilon} = y_i - \hat{m}(x_i)$ :

$$z_i = \log(\hat{\epsilon}_i^2) = \log((y_i - \hat{m}(x_i))^2).$$

```
eps_hat_sm <- lgWeight - sm_m$estimate
z_sm <- log(eps_hat_sm^2)
```

3. Fit a nonparametric regression to data  $(x_i, z_i)$  and call the estimated function  $\hat{q}(x)$ . Observe that  $\hat{q}(x)$  is an estimate of  $\log(\sigma^2(x))$ .

```
h2 <- dpill(Yr, z_sm)
sm.regression(Yr, z_sm, h=h2, model="none", eval.points=Yr) -> sm_q
```



4. Estimate  $\sigma^2(x)$  by

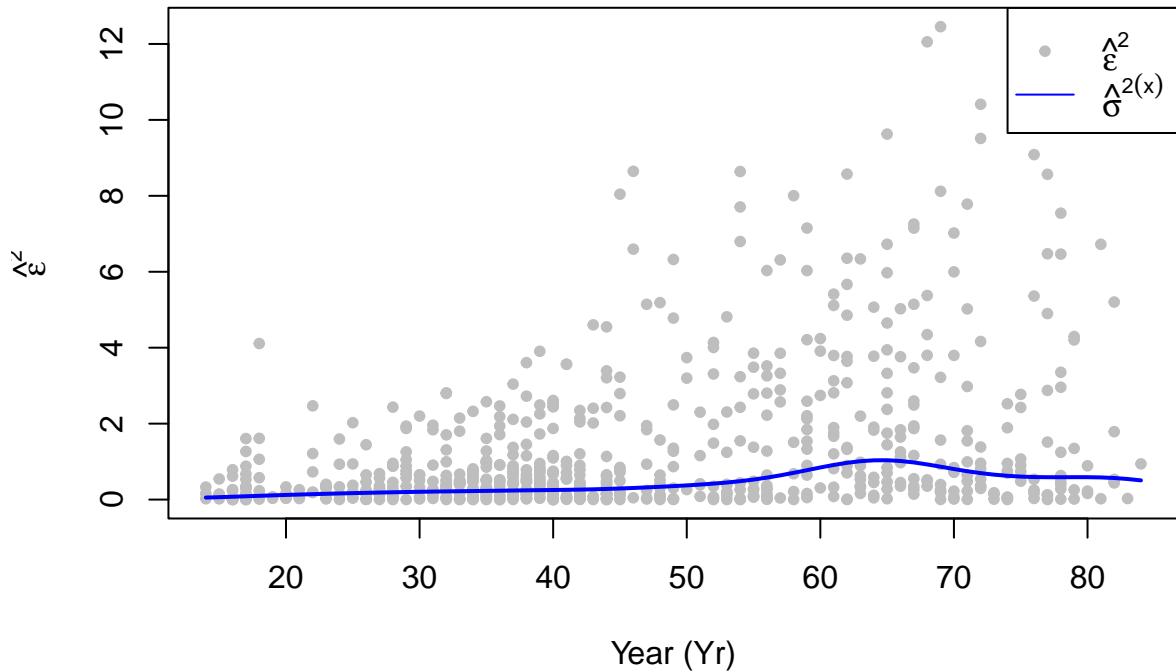
$$\hat{\sigma}^2(x) = e^{\hat{q}(x)}.$$

```
sigma2_hat_sm <- exp(sm_q$estimate)
sigma_hat_sm <- sqrt(sigma2_hat_sm)
```

Apply this procedure to estimate the conditional variance of lgWeigth (variable Y) given Yr (variable x). Draw a graphic of  $\hat{\epsilon}_i^2$  against  $x_i$  and superimpose the estimated function  $\hat{\sigma}^2(x)$ . Lastly draw the function  $\hat{m}(x)$  and superimpose the bands  $\hat{m}(x) \pm 1.96\hat{\sigma}(x)$ .

```
plot(Yr, eps_hat_sm^2, col="grey", pch=20,
      xlab="Year (Yr)", ylab=expression(hat(epsilon)^2),
      main=expression(paste(hat(epsilon)^2, " vs Yr with ", hat(sigma)^2(x))))
lines(Yr, sigma2_hat_sm, col="blue", lwd=2)
legend("topright", legend=c(expression(hat(epsilon)^2), expression(hat(sigma)^2(x))),
       col=c("grey","blue"), lty=c(NA,1), pch=c(20,NA))
```

## $\hat{\varepsilon}^2$ vs Yr with $\hat{\sigma}^2(x)$



```
plot(Yr, m_hat_sm, type="l", lwd=2, col="black",
      main=expression(paste(hat(m)(x), " ± 1.96 ", hat(sigma)(x))),
      xlab="Year (Yr)", ylab="log(Weight)")
lines(Yr, m_hat_sm + 1.96*sigma_hat_sm, col="red", lty=2, lwd=1.5)
lines(Yr, m_hat_sm - 1.96*sigma_hat_sm, col="red", lty=2, lwd=1.5)
points(Yr, lgWeight, col="grey", pch=20)
legend("topright",
       legend=c(expression(hat(m)(x)), expression(hat(m)(x) %+-% 1.96*hat(sigma)(x))),
       lty=c(1,2), col=c("black","red"), lwd=c(2,1.5))
```

