

Topics

- Max Likelihood Estimation (MLE)
 - Max A posterior Estimation (MAP)
 - Kalman Filter
- prob perspectives
on LS.

Maximum Likelihood Estimation (MLE)

unknown parameters: x probability density

measurements: $\tilde{y} = \begin{bmatrix} \tilde{y}_1 \\ \vdots \\ \tilde{y}_m \end{bmatrix}$ $P(\tilde{y}|x)$

Gaussian example:

$$1D: x = [\mu \sigma^2]$$

$$P(\tilde{y}|x) = \left(\frac{1}{2\pi\sigma^2} \right)^{m/2} e^{-\sum_{i=1}^m (\tilde{y}_i - \mu)^2 / 2\sigma^2}$$

Given measurement \tilde{y} the max likelihood estimation \hat{x} value of x that maximizes

$$P(\tilde{y}|x)$$

Likelihood Function

prob of seeing all meas.

$$L(\tilde{y}|x) = \prod_{i=1}^m P(\tilde{y}_i|x) \quad \leftarrow L(\tilde{y}|x) \geq 0$$

prob of seeing meas \tilde{y}_i

$$\hat{x} = \max_x L(\tilde{y}|x) \quad \star$$

Problems:

$$\prod_{i=1}^m p(\tilde{y}_i|x) \leftarrow$$

cumbersome

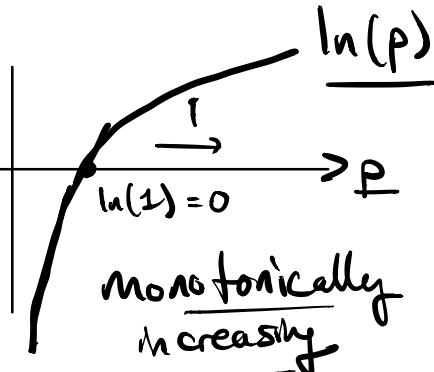
other of Gaussian
form contd by

Solution: $\ln(\cdot)$

Log-Likelihood

$$\hat{x} = \max_x \ln L(\tilde{y}|x) \quad \star\star$$

Optimizer of $\star\star$ also
optimizer of \star



$$\ln L(\tilde{y}|x) = \ln \left(\prod_{i=1}^m p(\tilde{y}_i|x) \right) = \sum_{i=1}^m \ln p(\tilde{y}_i|x)$$

Necessary: $\left[\frac{\partial}{\partial x} \ln L(\tilde{y}|x) \right]_{\hat{x}} = 0 \leftarrow \begin{matrix} \text{fixed pt.} \\ \text{if } x \text{ is a vector} \rightarrow \text{matrix} \end{matrix}$

Sufficient: $\left[\frac{\partial^2}{\partial x \partial x} \ln L(\tilde{y}|x) \right]_{\hat{x}} \leftarrow 0 \quad \begin{matrix} \text{negative} \\ \text{definite} \end{matrix}$

an extension:

Maximum a Posteriori Estimation (MAP)

specific case of
Bayesian Estimation

before: x
fixed
is unknown

review:

Conditional:
of \tilde{y} given x $P(\tilde{y}|x)$

$$P(\tilde{y}, x) = P(\tilde{y}|x) P(x)$$

prob of both = $P(x|\tilde{y}) P(\tilde{y})$

now:

$$x \sim P(x)$$

how do we incorporate this "prior" knowledge of x into estimation?

Bayes Rule:

$$P(x|\tilde{y}) P(\tilde{y}) = P(\tilde{y}, x) = P(\tilde{y}|x) P(x)$$

$$P(x|\tilde{y}) = \frac{P(\tilde{y}|x) P(x)}{P(\tilde{y})}$$

Relationship
between conditional
probabilities
 $P(\tilde{y}|x) \neq P(x|\tilde{y})$

$$\text{MLE} : \max p(\tilde{y}|x)$$

$$\text{MAP} : \max p(x|\tilde{y}) = \frac{p(\tilde{y}|x)p(x)}{p(\tilde{y})}$$

$$\max_x p(x|\tilde{y}) = \left[\frac{p(\tilde{y}|x)p(x)}{p(\tilde{y})} \right] \rightarrow \begin{matrix} \text{doesn't} \\ \text{depend on} \\ x \end{matrix}$$

$$\max_x p(\tilde{y}|x)p(x) \quad \begin{matrix} \text{added term} \\ \downarrow \quad \text{based on} \\ \text{applying } \ln(\cdot) \quad \text{prior of } x \end{matrix}$$

$$\max_x \ln(p(\tilde{y}|x)) + \ln p(x) \quad \begin{matrix} \text{give a bonus} \\ \text{for } x \text{ 's } \bar{w} \\ \text{high probability} \end{matrix}$$

Ex. Gaussians.

Maximum Likelihood: MLE

$$\tilde{y} = Hx + v \quad v \sim N(0, R) \quad \tilde{y} \in \mathbb{R}^m \quad H \in \mathbb{R}^{m \times n}$$

$$v = \tilde{y} - Hx \quad -\frac{1}{2} [\tilde{y} - Hx]^T R^{-1} [\tilde{y} - Hx]$$

$$P(\tilde{y}|x) = \frac{1}{(2\pi)^{m/2} \det(R)^{1/2}} e^{-\frac{1}{2} [\tilde{y} - Hx]^T R^{-1} [\tilde{y} - Hx]}$$

$$\max_x \ln L(\tilde{y}|x) = \ln p(\tilde{y}|x) \quad \ln c^x = x$$

$$= -\underbrace{\ln \left(\frac{1}{(2\pi)^{m/2} \det(R)^{1/2}} \right)}_{\text{constant}} - \frac{1}{2} [\tilde{y} - Hx]^T R^{-1} [\tilde{y} - Hx]$$

$$\min_x + \frac{1}{2} [\tilde{y} - Hx]^T R^{-1} [\tilde{y} - Hx] \quad \leftarrow \begin{array}{l} \text{Same} \\ \text{as min} \\ \text{variance} \\ \text{estimator} \end{array}$$

$$\Rightarrow \hat{x} = (H^T R^{-1} H)^{-1} H^T R^{-1} \tilde{y}$$

Max. a posteriori (MAP)

Now a prior... $\underline{x} \sim N(\underline{x}_a, \underline{Q})$

$$P(x) = \frac{1}{(2\pi)^{\frac{N}{2}} \det(Q)} e^{-\frac{1}{2} [\underline{x}_a - \underline{x}]^T \underline{Q}^{-1} [\underline{x}_a - \underline{x}]}$$

$$\max_x \ln P(\tilde{y}|x) + \ln P(x)$$

$$\min_x + \frac{1}{2} [\tilde{y} - Hx]^T R^{-1} [\tilde{y} - Hx] + \frac{1}{2} (\underline{x}_a - x)^T \underline{Q}^{-1} (\underline{x}_a - x)$$

$$\frac{\partial}{\partial x} = 0 \Rightarrow \hat{x} = \underbrace{(H^T R^{-1} H + \bar{Q}^{-1})^{-1}}_{\substack{\uparrow \\ \text{minimum variance}}} \underbrace{(H^T R^{-1} \tilde{y} + \bar{Q}^{-1} \underline{x}_a)}_{\substack{\uparrow \\ \text{estimate w/ a prior on } x}}$$

efficient

Efficiency:

MAP estimator: new term from prior distribution

Cramer Rao bound: $\rho(x)$

$$P = E[(\hat{x} - x)(\hat{x} - x)^T] \geq [F + E \frac{\partial}{\partial x} \ln(\rho(x)) \frac{\partial}{\partial x} \ln(\rho(x))^T]^{-1}$$

now introduce estimation for parameters in dynamics ...

Before: Reality: Estimator

static $\tilde{y} = Hx + v$

$$\hat{y} = H\hat{x}$$

Now
dynamic $\dot{x} = Fx + Bu$

$$\begin{aligned}\hat{x} &= F\hat{x} + Bu + K[\tilde{y} - H\hat{x}] \\ \hat{y} &= H\hat{x}\end{aligned}$$

feedback
observer feedback gain

NEW NOTATION

$$\tilde{x} = e = \hat{x} - x$$

$$\dot{e} = \dot{\hat{x}} - \dot{x} = (F - KH)e + KV + Bu - Bu$$

LINSYS

Question:

how to select

char poly /

eigenvalues...

• place cmd. selected

• observability char poly / eigenvals

→ controllability canonical form for stability

• Ackermann's formula

Discrete time version:

$$\dot{x}_t = F_t x + B_t u \quad \Phi = e^{F\Delta t}$$

$$\tilde{y}_t = H_t x + V_t \quad \Gamma = \int_0^{\Delta t} e^{F(\Delta t - t)} B dt$$

Reality

$$x_{k+1} = \Phi_k x_k + \Gamma u_k$$

$$y_k = H_k x_k + V_k$$

Estimator

$$\hat{x}_{k+1} = \Phi_k \hat{x}_k + \Gamma u_k + K_k [\tilde{y}_{k+1} - H_k \hat{x}_k]$$

Instead problem

$$\textcircled{1} \quad \underline{\hat{x}_{k+1}^-} = \underline{\Phi_k} \underline{\hat{x}_k^+} + \underline{\Gamma u_k} \quad \text{prediction or propagation}$$

$$\textcircled{2} \quad \underline{\hat{x}_{k+1}^+} = \underline{\hat{x}_{k+1}^-} + \underline{K} [\underline{\tilde{y}_{k+1}} - \underline{H} \underline{\hat{x}_{k+1}^-}] \quad \text{update or correction}$$

Full update:

$$\hat{x}_{k+1}^- = \Phi_k \hat{x}_k^- + \Gamma u_k + \Phi_k K [\tilde{y}_k - H \hat{x}_k^-]$$

$$\hat{x}_k^- = e_k^- = \hat{x}_k^- - x_k$$

$$\hat{x}_k^+ = e_k^+ = \hat{x}_k^+ - x_k$$

$$e_{k+1}^- = \Phi [I - \underline{K} H] e_k^- \leftarrow$$

$$e_{k+1}^+ = [I - \underline{K} H] \Phi e_k^+ \leftarrow$$

Note:

$$\Phi [I - K H] \in$$

[I - KH] Φ
have the same
eigenvalues

(square invertible
 A, B)

eigenvalues evals
 $AB = BA$

$$\underline{B} (\underline{AB})^{-1} = \underline{BA}$$

DISCRETE TIME KALMAN FILTER:

Reality:

$$\begin{aligned} \hat{x}_{k+1} &= \Phi_k \hat{x}_k + \Gamma_k u_k + \gamma_k w_k && \begin{matrix} \downarrow \\ \text{process noise} \end{matrix} \\ &\quad \begin{matrix} \uparrow \\ \text{state} \end{matrix} \quad \begin{matrix} \uparrow \\ \text{control} \end{matrix} \\ \tilde{y}_k &= H_k \hat{x}_k + v_k && \begin{matrix} \leftarrow \\ \text{measurement noise} \end{matrix} \end{aligned}$$

NOISE v_k, w_k zero mean Gaussian

$$\begin{aligned} v_k &\sim N(0, R_k) \\ w_k &\sim N(0, Q_k) \end{aligned} \quad \begin{matrix} \leftarrow \text{ static version} \\ \text{white noise version} \end{matrix} \quad \begin{matrix} \text{white noise processes} \\ \rightarrow v_k \perp w_k \\ \text{are not correlated over time.} \end{matrix}$$

$$\begin{aligned} E[v_k] &= 0 & E[v_k v_j^T] &= \begin{cases} 0 & k \neq j \\ R_k & k = j \end{cases} \\ E[w_k] &= 0 & E[w_k w_j^T] &= \begin{cases} 0 & k \neq j \\ Q_k & k = j \end{cases} \end{aligned}$$

Optimal estimator derivation:

$$\begin{aligned} \hat{\bar{x}}_{k+1} &= \Phi_k \hat{x}_k^+ + \Gamma_k u_k && \leftarrow \text{propagation} \\ \hat{x}_k^+ &= \hat{\bar{x}}_k + K_k [\tilde{y}_k - H_k \hat{\bar{x}}_k^-] && \leftarrow \text{correction} \end{aligned}$$

Error covariance

$$\begin{aligned} e_k^- &= \hat{\bar{x}}_k - \underline{x}_k & e_{k+1}^- &= \hat{\bar{x}}_{k+1} - \underline{x}_{k+1} \\ e_k^+ &= \hat{x}_k^+ - \underline{x}_k & e_{k+1}^+ &= \hat{x}_{k+1}^+ - \underline{x}_{k+1} \end{aligned}$$

$$P_k^- = E[e_k^- e_k^{-T}] \quad P_{k+1}^- = E[e_{k+1}^- e_{k+1}^{-T}]$$

$$P_k^+ = E[e_k^+ e_k^{+T}] \quad P_{k+1}^+ = E[e_{k+1}^+ e_{k+1}^{+T}]$$

$$\underline{e}_{k+1}^- = \underline{\Phi}_k \underline{e}_k^+ - \underline{Y}_k \underline{w}_k \\ + \cancel{\underline{\Gamma}_{k+1} \underline{u}_k} - \cancel{\underline{\Gamma}_k \underline{u}_k}$$

Covariance update:

DISCRETE TIME KALMAN FILTER:

Reality:

$$\underline{x}_{k+1} = \underline{\Phi}_k \underline{x}_k + \underline{\Gamma}_k \underline{u}_k + \underline{Y}_k \underline{w}_k$$

↑ state ↑ control ↓ process noise

$$\tilde{y}_k = H_k \underline{x}_k + V_k$$

measurement noise

$$P_{k+1}^- = E[e_{k+1}^- e_{k+1}^{-T}] \\ = E[\underline{\Phi}_k \underline{e}_k^+ \underline{e}_k^{+T} \underline{\Phi}_k^T] - E[\underline{\Phi}_k \underline{e}_k^+ \underline{w}_k \cancel{Y}_k^T] \\ - E[\cancel{Y}_k \underline{w}_k \underline{e}_k^{+T} \underline{\Phi}_k^T] + E[\underline{Y}_k \underline{w}_k \underline{w}_k^T \cancel{Y}_k^T] \\ = \underline{\Phi}_k E[e_k^+ e_k^{+T}] \underline{\Phi}_k^T + \underline{Y}_k E[w_k w_k^T] \cancel{Y}_k^T$$

→ $P_{k+1}^- = \underline{\Phi}_k \underline{P}_k^+ \underline{\Phi}_k^T + \underline{Y}_k Q_k \cancel{Y}_k^T$ ↪

$$\underline{e}_k^+ = \hat{\underline{x}}_k^+ - \underline{x}_k = (I - K_k H_k) \hat{\underline{x}}_k^- + K_k H_k \underline{x}_k + K_k V_k - \underline{x}_k \\ = (I - K_k H_k) \underline{e}_k^- + K_k V_k$$

$$P_k^+ = E(e_k^+ e_k^{+T}) = E[(I - K_k H_k) \underline{e}_k^- \underline{e}_k^{-T} (I - K_k H_k)^T] \\ + E[(I - K_k H_k) \underline{e}_k^- \cancel{V}_k \cancel{K}_k^T] + E[K_k V_k \cancel{e}_k^{-T} (I - K_k H_k)^T] \\ E[K_k V_k V_k^T K_k^T] \\ = (I - K_k H_k) E[\underline{e}_k^- \underline{e}_k^{-T}] (I - K_k H_k)^T + K_k E[V_k V_k^T] K_k^T$$

$$P_k^+ = (I - K_k H_k) P_k^- (I - K_k H_k)^T + K_k R_k K_k^T \quad \leftarrow$$

Initial cond:

$$P_0^- = E[e_0 e_0^T] = E[(\hat{x}_0 - x_0)(\hat{x}_0 - x_0)^T]$$

\downarrow doesn't have to be precise

Choosing K_k :

$$\min_{K_k} J(K_k) = \text{Tr}(P_k^+) = E(\text{Tr}(e_k^+ e_k^{+T})) \\ = E(\text{Tr}(e_k^+ e_k^+))$$

$$\frac{\partial J}{\partial K_k} = 0 = -2(I - K_k H_k) P_k^- H_k^T + 2 K_k R_k \quad \begin{matrix} \downarrow \text{sum of} \\ \text{squared} \\ \text{errors} \end{matrix}$$

$$\Rightarrow K_k = P_k^- H_k^T [H_k P_k^- H_k^T + R_k]^{-1} \quad \leftarrow$$

plugging back in to P_k^+ step...

$$P_k^+ = (I - K_k H_k) P_k^- (I - K_k H_k)^T + K_k R_k K_k^T$$

$$= P_k^- - K_k H_k P_k^- - P_k^- H_k^T K_k^T + K_k R_k K_k^T$$

$$+ K_k H_k P_k^- H_k^T K_k^T$$

$$- P_k^- H_k^T [H_k P_k^- H_k^T + R_k]^{-1} H_k P_k^-$$

$$K_k [H_k P_k^- H_k^T + R_k] K_k^T$$

$$+ P_k^- H_k^T [H_k P_k^- H_k^T + R_k]^{-1} H_k P_k^-$$

$$P_k^+ = P_k^- - K_k H_k P_k^-$$

$$P_k^+ = [I - K_k H_k] P_k^-$$

Summary:

MODEL: $\begin{aligned} \hat{x}_{k+1} &= \Phi_k \hat{x}_k + \Gamma_k u_k + Y_k w_k & w_k \sim N(0, Q_k) \\ \tilde{y}_k &= H_k \hat{x}_k + v_k & v_k \sim N(0, R_k) \end{aligned}$

Initialize: $\begin{aligned} \hat{x}_0^- &= \hat{x}(t_0) \\ P_0^- &= E[e_0 e_0^T] = E[(\hat{x}_0 - \bar{x}_0)(\hat{x}_0 - \bar{x}_0)^T] \end{aligned}$

Gain
(optimal) $K_k = P_k^- H_k^T [H_k P_k^- H_k^T + R_k]^{-1}$

Update $\begin{aligned} \hat{x}_k^+ &= \hat{x}_k^- + K_k [\tilde{y}_k - H_k \hat{x}_k^-] \\ P_k^+ &= [I - K_k H_k] P_k^- \end{aligned}$

Propagation $\hat{x}_{k+1}^- = \Phi_k \hat{x}_k^+ + \Gamma_k u_k$

$$P_{k+1}^- = \Phi_k P_k^+ \Phi_k^T + Y_k Q_k Y_k^T$$

$$\bar{P}_{k+1} = \Phi_k P_k^+ \Phi_k' + Y_k Q_k Y_k^+$$

$$P_k^+ = [I - K_k H_k] \bar{P}_k$$

$$P_k^+ = \bar{P}_k - \bar{P}_k H_k^T \underbrace{[H_k \bar{P}_k H_k^T + R_k]}^{-1} H_k \bar{P}_k$$

$$P_k^+ = \underbrace{(P_k^-)^{-1} + H_k^T R_k^{-1} H_k}_{\text{More computationally expensive}}^{-1}$$

↑ ↑

woodbury matrix identity

matrix inversion lemma

more computationally
expensive

$$(A + u C V)^{-1} = \bar{A}^{-1} - \bar{A}^{-1} \bar{C} \bar{u} (\bar{C}^T V \bar{A})^T \bar{V} \bar{A}^{-1}$$

Wikipedia

$$\bar{P}_{k+1} = \Phi_k P_k^+ \Phi_k' + Y_k Q_k Y_k^+$$

pos def

pos def

sort of depending on Φ_k

dynamics

increases size of P_k

$$P_k^+ = \bar{P}_k - \bar{P}_k H_k^T \underbrace{[H_k \bar{P}_k H_k^T + R_k]}^{-1} H_k \bar{P}_k$$

pos def

pos def

meas. correction

decreases size of P_k

improve the covariance