

Review: Weighted LS

Model: $y = Hx$

Meas: $\tilde{y} = Hx + v$

$e = \tilde{y} - y$ W : diag. weights

$\min_x \frac{1}{2} e^T e = (\tilde{y} - Hx)^T (\tilde{y} - Hx) \Rightarrow \hat{x} = (H^T W H)^{-1} H^T W \tilde{y}$ solns

BATCH LS: $H_1 \in \mathbb{R}^{m_1 \times n}$ $H_2 \in \mathbb{R}^{m_2 \times n}$

2 sets of meas:

$\left\{ \begin{array}{l} \tilde{y}_1 = H_1 \boxed{x} + v_1 \end{array} \right. \} \text{ init. set of meas}$

$\left\{ \begin{array}{l} \tilde{y}_2 = H_2 \boxed{x} + v_2 \end{array} \right. \} \text{ add. set of meas}$

$$\begin{bmatrix} \tilde{y}_1 \\ \tilde{y}_2 \end{bmatrix} = \begin{bmatrix} H_1 \\ H_2 \end{bmatrix} \overset{\downarrow}{x} + \begin{bmatrix} v_1 \\ v_2 \end{bmatrix} \quad \begin{array}{l} W_1: \text{diag } w_1 > 0 \\ W_2: \text{diag } w_2 > 0 \end{array}$$

if we're already solved for $\hat{x}_1 = (H_1^T W_1 H_1)^{-1} H_1^T W_1 \tilde{y}_1$

can we leverage \hat{x}_1 to compute new est. w all data.
in particular:

H_1 : tall \rightarrow lot of data pts.

H_2 : fat \rightarrow just adding a few rows / data pts.

want to compute new estimate: \hat{x}_2

$$\hat{x}_2 = \left(\begin{bmatrix} H_1^T H_2 \\ 0_{m_2} \end{bmatrix} \begin{bmatrix} W_1 \\ 0_{m_2} \end{bmatrix} \right) \begin{bmatrix} H_1^T H_2 \\ 0_{m_2} \end{bmatrix}^{-1} \begin{bmatrix} w_1 \\ 0_{m_2} \end{bmatrix} \begin{bmatrix} \tilde{y}_1 \\ \tilde{y}_2 \end{bmatrix}$$

$$= \underbrace{\left[\begin{bmatrix} H_1^T W_1 H_1 + H_2^T W_2 H_2 \\ 0_{m_2} \end{bmatrix} \right]^{-1}}_{\text{need to invert.}} \begin{bmatrix} H_1^T W_1 \tilde{y}_1 + H_2^T W_2 \tilde{y}_2 \\ 0_{m_2} \end{bmatrix}$$

already computed $(H_1^T W_1 H_1)^{-1}$

can we leverage $\left[\begin{bmatrix} H_1^T W_1 H_1 + H_2^T W_2 H_2 \\ 0_{m_2} \end{bmatrix} \right]^{-1}$ to make inverting \star easier?

if H_2 fat: (just a few new data pts)

$$\underbrace{\left[\begin{bmatrix} H_1^T W_1 H_1 + H_2^T W_2 H_2 \\ 0_{m_2} \end{bmatrix} \right]^{-1}}_{n \times n} \leftarrow$$

$$\left[\begin{bmatrix} H_1^T W_1 H_1 \\ 0_{m_2} \end{bmatrix} \right] + \underbrace{\left[\begin{bmatrix} H_2^T \\ 0_{m_2} \end{bmatrix} \right]^{-1} \left[\begin{bmatrix} H_2 \\ 0_{m_2} \end{bmatrix} \right]}_{\text{adding a low rank piece}}$$

adding a low rank piece

Woodbury Matrix Identity. ✓

$$\underbrace{(A + UCV)^{-1}}_{\substack{\uparrow \\ \uparrow}} = A^{-1} - A^{-1}U \left(C^{-1} + V A^{-1} U \right)^{-1} V A^{-1} \substack{\uparrow \\ \uparrow}$$

$$\left[\mathbf{J} + \mathbf{u} \mathbf{u}^T \right]^{-1} = \left[\mathbf{A}^T \right] - \left[\mathbf{A}^T \left[\mathbf{C}^{-1} + \mathbf{V} \mathbf{A}^T \mathbf{u} \right]^{-1} \mathbf{V} \mathbf{A}^T \right] \leftarrow$$

one $n \times n$
inverse

order $\overbrace{n^3}^{(n^2 \cdot 3)}$

for $2 m_2 \times m_2$ and some
matrix multiplications

if $n \gg m_2$

\Rightarrow on order $\overbrace{n^2}^{(-)}$

$$\hat{\mathbf{x}}_2 = \left(\mathbf{H}_1^T \mathbf{W}_1 \mathbf{H}_1 + \mathbf{H}_2^T \mathbf{W}_2 \mathbf{H}_2 \right)^{-1} \left[\mathbf{H}_1^T \mathbf{W}_1 \tilde{\mathbf{y}}_1 + \mathbf{H}_2^T \mathbf{W}_2 \tilde{\mathbf{y}}_2 \right]$$

$$\text{let } P_1 = \left(\mathbf{H}_1^T \mathbf{W}_1 \mathbf{H}_1 \right)^{-1}, \quad P_2 = \left(\mathbf{H}_1^T \mathbf{W}_1 \mathbf{H}_1 + \mathbf{H}_2^T \mathbf{W}_2 \mathbf{H}_2 \right)^{-1}$$

Woodbury:

$$P_2 = P_1 - \underbrace{P_1 \mathbf{H}_2^T \left(\mathbf{W}_2^{-1} + \mathbf{H}_2 P_1 \mathbf{H}_2^T \right)^{-1} \mathbf{H}_2 P_1}_{\leftarrow K_2} \leftarrow$$

Let

$$K_2 := P_1 \mathbf{H}_2^T \left(\mathbf{W}_2^{-1} + \mathbf{H}_2 P_1 \mathbf{H}_2^T \right)^{-1} \stackrel{*}{=} P_2 \mathbf{H}_2^T \mathbf{W}_2 \quad \text{not obvs.}$$

Proof of \star :

$$\begin{aligned} P_2 \mathbf{H}_2^T \mathbf{W}_2 &= P_1 \mathbf{H}_2^T \mathbf{W}_2 - P_1 \mathbf{H}_2^T \left(\mathbf{W}_2^{-1} + \mathbf{H}_2 P_1 \mathbf{H}_2^T \right)^{-1} \mathbf{H}_2 P_1 \mathbf{H}_2^T \mathbf{W}_2 \\ &= P_1 \mathbf{H}_2^T \left(\mathbf{W}_2 - \left(\mathbf{W}_2^{-1} + \mathbf{H}_2 P_1 \mathbf{H}_2^T \right)^{-1} \mathbf{H}_2 P_1 \mathbf{H}_2^T \mathbf{W}_2 \right) \\ &= P_1 \mathbf{H}_2^T \underbrace{\left(\mathbf{W}_2^{-1} + \mathbf{H}_2 P_1 \mathbf{H}_2^T \right)^{-1}}_{K_2} \left[\left(\mathbf{W}_2^{-1} + \mathbf{H}_2 P_1 \mathbf{H}_2^T \right)^{-1} - \cancel{\mathbf{H}_2 P_1 \mathbf{H}_2^T} \right] \mathbf{W}_2 \end{aligned}$$

$$K_2 = P_1 H_2^T (\omega_2^{-1} + H_2 P_1 H_2^T)^{-1}$$

like
KF gain

$$P_2 = (I - K_2 H_2) P_1$$

like
covariance
update

$$\hat{x}_2 = \underbrace{P_2 H_1^T \omega_1 \tilde{y}_1}_{\sim} + \underbrace{P_2 H_2^T \omega_2 \tilde{y}_2}_{\sim}$$

$$= (I - K_2 H_2) \underbrace{P_1 H_1^T \omega_1 \tilde{y}_1}_{\hat{x}_1} + K_2 \tilde{y}_2$$

$$\hat{x}_2 = \hat{x}_1 + K_2 (\tilde{y}_2 - \underbrace{H_2 \hat{x}_1}_{\text{update gain}}) \leftarrow$$

like state
update

if \hat{x}_1 is correct
what will \tilde{y}_2 be?
i.e. prediction of \tilde{y}_2

modification to the estimate

SEQUENTIAL LS:

have \hat{x}_k , K_k , P_k

$$K_{k+1} = P_k H_{k+1}^T (\omega_{k+1}^{-1} + H_{k+1} P_k H_{k+1}^T)^{-1}$$

$$P_{k+1} = (I - K_{k+1} H_{k+1}) P_k$$

$$\hat{x}_{k+1} = \hat{x}_k + K_{k+1} (\tilde{y}_{k+1} - \underbrace{H_{k+1} \hat{x}_k}_{\checkmark})$$

$$x \quad k_F \quad x^+ = Ax + Bu.$$

Nonlinear LS: (Gauss-Newton?)

$$\text{Model: } y = f(x)$$

$$\text{Meas } \tilde{y} = f(x) + v$$

Levenberg - Marquadt
(another version)

Iterative linearization process

$$x = x_c + \underline{\Delta x}$$

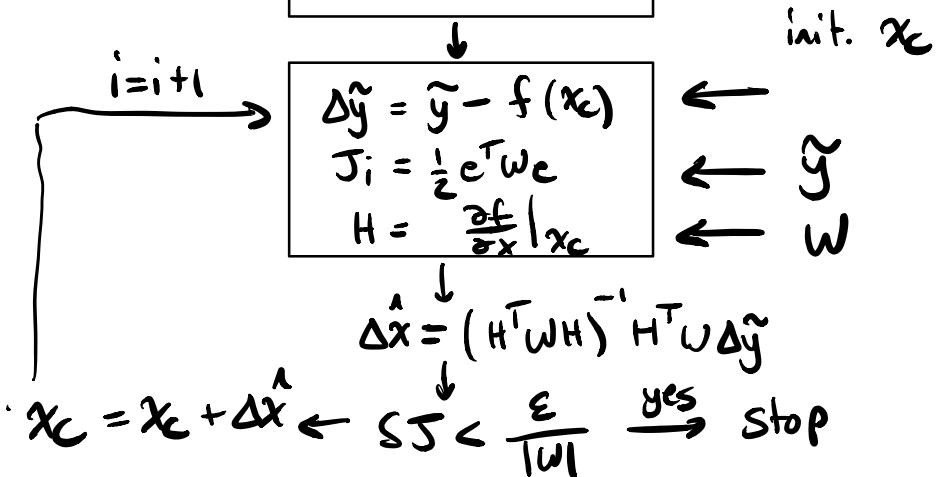
$$f(x) = \underline{f(x_c + \Delta x)} \approx f(x_c) + \left. \frac{\partial f}{\partial x} \right|_{x_c} \Delta x$$

$$c = \tilde{y} - f(x) = \tilde{y} - \underline{f(x_c)} - H \Delta x$$

$$\min_{\Delta x} \frac{1}{2} c^T W c = \frac{1}{2} (\Delta \tilde{y} - H \Delta x)^T (\Delta \tilde{y} - H \Delta x)$$

$$\text{solution: } \hat{\Delta x} = (H^T W H)^{-1} H^T W \Delta \tilde{y}$$

MODEL $f(x), \frac{\partial f}{\partial x}$



MINIMUM VARIANCE ESTIMATION

$$\rightarrow \hat{y} = Hx + v \quad \begin{matrix} \text{give} \\ v \text{ more} \\ \text{structure} \end{matrix} \quad v \sim N(0, R)$$

$$\rightarrow \quad \quad \quad R = R^T > 0$$

what is an estimator...

$\hat{x}(\hat{y}) \Leftarrow$ function of meas

Assume some class of functions for $\hat{x}(\cdot)$
find best $\hat{x}(\hat{y})$ w/in that class of functions

Estimator Bias "is the expected value
of the estimated parameters
consistent w/ true parameters"

$$\text{BIAS: } E[\hat{x}(\hat{y}) - x]$$

$$\text{UNBIASED} \Rightarrow E[\hat{x}(\hat{y})] = x$$

Linear Estimator (assume)

$$\hat{x} = M\hat{y} + n \quad \begin{matrix} \text{picking } \hat{x}(\cdot) \text{ is} \\ \text{now picking } M, n \end{matrix}$$

What cons do we need on M, n for an
unbiased estimator?

$$E[\hat{x}] = E[M\hat{y} + n] = E[MHx + Mv + n]$$

$$= MHx + n + M E[v] \stackrel{\text{want.}}{=} x$$

\Rightarrow imposes constraints
on $M \notin n \dots$

- $\mathbf{H} = \mathbf{O}$
 - $\mathbf{M}\mathbf{H} = \mathbf{I}$
- } → guarantee that our unbiased

NOTE: $\mathbf{M} = \mathbf{H}^{-1}$?

\mathbf{H} is tall. can't invert \mathbf{H} ...

if $\mathbf{M}\mathbf{H} = \mathbf{I}$ and $\mathbf{Z}\mathbf{H} = \mathbf{O} \Rightarrow (\mathbf{M} + \mathbf{Z})\mathbf{H} = \mathbf{I}$
since \mathbf{H} is tall multiple \mathbf{M} 's can work.

- $\mathbf{M} = \underbrace{(\mathbf{H}^T \mathbf{H})^{-1}}_{\mathbf{H}^T \mathbf{H}} \mathbf{H}^T \quad (\mathbf{H}^T \mathbf{H})^{-1} \mathbf{H}^T \times \mathbf{H} = \mathbf{I}$
- $\mathbf{M} = \underbrace{(\mathbf{H}^T \mathbf{W} \mathbf{H})^{-1}}_{\mathbf{H}^T \mathbf{W} \mathbf{H}} \mathbf{H}^T \mathbf{W} \quad (\mathbf{H}^T \mathbf{W} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{W} \times \mathbf{H} = \mathbf{I}$
- take any \mathbf{C} s.t. \mathbf{CH} is invertible...

$$\mathbf{M} = (\mathbf{CH})^{-1} \mathbf{C} \Rightarrow \underbrace{(\mathbf{CH})^{-1} \mathbf{C}}_{\mathbf{C}^{-1} \mathbf{H}^{-1}} \mathbf{C} \times \mathbf{H} = \mathbf{I}$$

$$\mathbf{C}^{-1} \mathbf{H}^{-1}$$

unbiased \Rightarrow gave us constraints --

now choose $\mathbf{M}, (\mathbf{C})$ to minimize variance

$$\begin{aligned} J &= \frac{1}{2} E \left[(\hat{\mathbf{x}} - \mathbf{x})^T (\hat{\mathbf{x}} - \mathbf{x}) \right] \\ &= \frac{1}{2} E \sum_i (\hat{x}_i - x_i)^2 \end{aligned}$$

$$= \text{Tr} \frac{1}{2} E \left[(\hat{x} - x)^T (\hat{x} - x) \right] \quad \text{Tr}(AB) = \text{Tr}(BA)$$

$$= \frac{1}{2} \text{Tr} E \left[(\hat{x} - x) (\hat{x} - x)^T \right]$$

min trace of covariance of \hat{x}

$$\begin{aligned} \hat{x} &= M \hat{y} + \underbrace{v}_0 \\ &= \underbrace{M H x}_{\text{needs to be } I} + \underbrace{M v}_0 \\ &= x + M v \\ &\qquad v \sim N(0, R) \end{aligned}$$

$$J = \frac{1}{2} \text{Tr} E \left[M v v^T M^T \right] = \frac{1}{2} \text{Tr} (M R M^T)$$

optimization:

$$\begin{array}{l} \min_M J = \frac{1}{2} \text{Tr} (M R M^T) \\ \text{s.t. } M H = I \end{array} \quad \xrightarrow{\quad} \quad \text{HOMEWORK}$$

$$M = (H^T R^{-1} H)^{-1} H^T R^{-1}$$

$$\Rightarrow \hat{x} = M \hat{y} = (H^T R^{-1} H)^{-1} H^T R^{-1} \hat{y}$$

this is weighted LS with $W = R^{-1}$

\tilde{R}^{-1} not diag, $\tilde{R}^{-1} > 0$

How do I weight my measurements properly
to account for noise v ?

deps on structure of v $v \sim N(0, R)$

\Rightarrow use $W = \tilde{R}^{-1}$ ← meas. w more noise
get weighted less.

Solving HWI

$$L(M, V) = \frac{1}{2} \underbrace{\text{Tr}(M R M^T)} + \underbrace{\text{Tr}(V^T (M H - I))}$$

for matrices

$$\langle A, B \rangle = \sum_{ij} A_{ij} B_{ij} = \text{Tr}(A^T B)$$

$$\underbrace{\frac{\partial L}{\partial M}}_{} = 0 \quad \underbrace{\frac{\partial L}{\partial V}}_{} = 0$$

$$\frac{\partial}{\partial M} \frac{1}{2} \text{Tr}(M R M^T)$$

$$f(M) = \text{Tr}(M R M^T)$$

$$\Delta f = \text{Tr}(\Delta M R M^T) + \text{Tr}(\underbrace{M R \Delta M^T}_{\text{transpose}})$$

$$\begin{aligned}
&= \text{Tr}(\Delta M R M^T) + \text{Tr}(\Delta M R^T M) \\
&= \text{Tr}(\Delta M (R + R^T) M) \\
&= \langle \Delta M^T, \underline{(R + R^T) M} \rangle = \langle (R + R^T) M, \Delta M^T \rangle
\end{aligned}$$

$$f(x) = \langle c, x \rangle \Rightarrow \frac{\partial f}{\partial x} = c$$

$$\frac{\partial f}{\partial M} = M^T (R + R^T)$$