

Problem Set #1

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Question 1

Let $(Y, X')'$ be a random vector, where $Y = X'\beta_0 \cdot U$, where $\mathbb{E}[U | X] = 1$, $\mathbb{E}[XX']$ is invertible, and $\mathbb{E}[Y^2 + ||X||^2] < \infty$.

- (i) Since $\mathbb{E}[U | X] = 1$, the expectation of Y , conditional on X , is $X'\beta_0$.
Then, $\frac{\partial}{\partial X} Y = \beta_0$.
- (ii) Define $V = U - 1$. Then, $\mathbb{E}[V | X] = \mathbb{E}[U - 1 | X] = 0$, and:

$$Y = X'\beta_0(V + 1) = X'\beta_0 V + X'\beta_0 = X'\beta_0 + \tilde{U}$$

Where:

$$\mathbb{E}[\tilde{U} | X] = \mathbb{E}[X'\beta_0 V | X] = X'\beta_0 \mathbb{E}[V | X] = 0$$

Thus, $Y = X'\beta_0 + \tilde{U}$, where $\mathbb{E}[\tilde{U} | X] = 0$.

- (iii) Let $\beta = \beta_0$. Then,
- (iv) Define $V = U - 1$. Then, $\mathbb{E}[V | X] = \mathbb{E}[U - 1 | X] = 0$, and:

$$\begin{aligned}\mathbb{E}[X(Y - X'\beta)] &= \mathbb{E}[X(X'\beta_0 \cdot U - X'\beta_0)] = \mathbb{E}[\mathbb{E}[X(X'\beta_0 \cdot U - X'\beta_0) | X]] \\ &= \mathbb{E}[XX'\beta_0 \mathbb{E}[(U - 1) | X]] = 0\end{aligned}$$

Thus, $\beta = \beta_0 \Rightarrow \mathbb{E}[X(Y - X'\beta)] = 0$. Now, Suppose $\mathbb{E}[X(Y - X'\beta)] = 0$.
Then,

$$\begin{aligned}\mathbb{E}[X(Y - X'\beta)] &= \mathbb{E}[X(X'\beta \cdot U - X'\beta_0)] = 0 \\ \mathbb{E}[XX' \mathbb{E}[\beta \cdot U - \beta_0 | X]] &= (\beta - \beta_0) \mathbb{E}[XX'] = 0\end{aligned}$$

We know that $\mathbb{E}[XX']$ is invertible, so $\mathbb{E}[XX'] \neq 0$. Thus, $\mathbb{E}[X(Y - X'\beta)] = 0 \Rightarrow \beta = \beta_0$.

$$\therefore \mathbb{E}[X(Y - X'\beta)] = 0 \iff \beta = \beta_0 \blacksquare$$

Knowing this, we can derive the method of moments estimator for β :

$$\begin{aligned}\mathbb{E}[X(Y - X'\beta)] &= \mathbb{E}[XY] - \mathbb{E}[XX']\beta = 0 \\ \mathbb{E}[XX']\beta &= \mathbb{E}[XY] \\ \beta &= \mathbb{E}[XX']^{-1} \mathbb{E}[XY] \\ \Rightarrow \hat{\beta} &= \left(\frac{1}{n} \sum_{i=1}^n X_i X_i' \right)^{-1} \frac{1}{n} \sum_{i=1}^n X_i Y_i = \hat{\beta}_{OLS}\end{aligned}$$

(v) We can simplify the final equation in (iii) to show:

$$\begin{aligned}\mathbb{E}[\hat{\beta} \mid X_1, \dots, X_n] &= \mathbb{E} \left[\left(\frac{1}{n} \sum_{i=1}^n X_i X_i' \right)^{-1} \frac{1}{n} \sum_{i=1}^n X_i Y_i \mid X_1, \dots, X_n \right] \\ &= \left(\frac{1}{n} \sum_{i=1}^n X_i X_i' \right)^{-1} \frac{1}{n} \sum_{i=1}^n X_i \mathbb{E}[X_i' \beta_0 \cdot U \mid X_1, \dots, X_n] \\ &= \beta_0 \left(\frac{1}{n} \sum_{i=1}^n X_i X_i' \right)^{-1} \left(\frac{1}{n} \sum_{i=1}^n X_i X_i' \right) \mathbb{E}[U \mid X_1, \dots, X_n] \\ &= \beta_0\end{aligned}$$

Thus, $\hat{\beta}$ is unbiased.

(vi) According to the weak law of large numbers (WLLN), random variables converge in probability to their expected value. Thus,

$$\hat{\beta} \rightarrow_p \mathbb{E}[\hat{\beta}] = \mathbb{E}[\mathbb{E}[\hat{\beta} \mid X_1, \dots, X_n]] = \mathbb{E}[\beta_0] = \beta_0$$

Thus, $\hat{\beta}$ is consistent.

Question 2

Question 3