## Problem Set #2

#### Danny Edgel Econ 710: Economic Statistics and Econometrics II Spring 2021

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Collaborated with Sarah Bass, Emily Case, Michael Nattinger, and Alex Von Hafften

### Question 1

(i) Yes,  $\hat{\beta}_1^{IV} \to_p \beta_1$ . By the Weak Law of Large Numbers (WLLN) and the recognition that the law of iterated expectation (LIE) implies  $\mathbb{E}[U] = \mathbb{E}[\mathbb{E}[U|Z]] = 2$ ,

$$\begin{split} \hat{\beta}_{1}^{IV} &\to_{p} \frac{\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(Y - \mathbb{E}\left[Y\right])\right]}{\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(X - \mathbb{E}\left[X\right])\right]} \\ &= \frac{\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(\beta_{0} + X\beta_{1} + U - \mathbb{E}\left[\beta_{0} + X\beta_{1} + U\right])\right]}{\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(X - \mathbb{E}\left[X\right])\right]} \\ &= \frac{\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(\beta_{0} + X\beta_{1} + U - \beta_{0} - \beta_{1}\mathbb{E}\left[X\right] - \mathbb{E}\left[U\right])\right]}{\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(X - \mathbb{E}\left[X\right])\right]} \\ &= \frac{\mathbb{E}\left[\beta_{1}(Z - \mathbb{E}\left[Z\right])(X - \mathbb{E}\left[X\right]) + (Z - \mathbb{E}\left[Z\right])(U - \mathbb{E}\left[U\right])\right]}{\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(X - \mathbb{E}\left[X\right])\right]} \\ &= \frac{\beta_{1}\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(X - \mathbb{E}\left[X\right])\right] + \mathbb{E}\left[ZU - U\mathbb{E}\left[Z\right] - Z\mathbb{E}\left[U\right] + \mathbb{E}\left[Z\right]\mathbb{E}\left[U\right]}{\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(X - \mathbb{E}\left[X\right])\right]} \\ &= \beta_{1} + \frac{2\mathbb{E}\left[Z\right] - 2\mathbb{E}\left[Z\right] + 2\mathbb{E}\left[Z\right] - 2\mathbb{E}\left[Z\right]}{\mathbb{E}\left[(Z - \mathbb{E}\left[Z\right])(X - \mathbb{E}\left[X\right])\right]} \\ &= \beta_{1} \end{split}$$

(ii) Yes,  $\hat{\beta}_0^{IV} \to_p \beta_0$ . Given (i), we can calculate:

$$\hat{\beta}_{0}^{IV} = \overline{Y} - \overline{X}\hat{\beta}_{1} \rightarrow_{p} \mathbb{E}\left[Y\right] - \mathbb{E}\left[X\right]\beta_{1} = \beta_{0}$$

## Question 2

- (i) Z is a valid instrument if  $Cov(Z, X) \neq 0$ , i.e., if  $\pi_1 \neq 0$ .
- (ii) We can derive  $\gamma_0$ ,  $\gamma_1$  and  $\varepsilon$  as functions of the structural parameters by first deriving the reduced form of the model:

$$Y = \beta_0 + (\pi_0 + Z\pi_1 + V) \beta_1 + U$$
  

$$Y = \beta_0 + \pi_0 \beta_1 + Z\pi_1 \beta_1 + V\beta_1 + U$$
  

$$Y = \gamma_0 + Z\gamma_1 + \varepsilon$$

Where:

$$\gamma_0 = \beta_0 + \pi_0 \beta_1, \ \gamma_1 = \pi_1 \beta_1, \ \varepsilon = V \beta_1 + U$$

(iii) The IV estimator of  $\beta_1$  is

$$\hat{\beta}_1^{IV} = \frac{\widehat{Cov(Z, Y)}}{\widehat{Cov(Z, X)}}$$

And the two OLS estimators, of  $\gamma_1$  and  $\pi_1$  respectively, are

$$\hat{\pi}_1 = \frac{\sum_{i=1}^n (Z_i - \overline{Z})(X_i - \overline{X})}{\sum_{i=1}^n (Z_i - \overline{Z})^2}, \, \hat{\gamma}_1 = \frac{\sum_{i=1}^n (Z_i - \overline{Z})(Y_i - \overline{Y})}{\sum_{i=1}^n (Z_i - \overline{Z})^2}$$

Then the indirect least squares estimator of  $\beta_1$  is

$$\frac{\hat{\gamma}_1}{\hat{\pi}_1} = \frac{\sum_{i=1}^n (Z_i - \overline{Z})(Y_i - \overline{Y}) \sum_{i=1}^n (Z_i - \overline{Z})^2}{\sum_{i=1}^n (Z_i - \overline{Z})^2 \sum_{i=1}^n (Z_i - \overline{Z})(X_i - \overline{X})} = \frac{\sum_{i=1}^n (Z_i - \overline{Z})(Y_i - \overline{Y})}{\sum_{i=1}^n (Z_i - \overline{Z})(X_i - \overline{X})} = \hat{\beta}_1^{IV}$$

(iv)

(v)

# Question 3

- (i)
- (ii)
- (iii)
- (iv)
- (v)
- (vi)