



Time Series Analysis and Forecasting

Chapter 5: GARCH and Volatility



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Learning Objectives

By the end of this chapter, you will be able to:

- Understand volatility clustering and its importance in financial data
- Model conditional heteroskedasticity using ARCH and GARCH models
- Estimate GARCH models and interpret their parameters
- Forecast volatility and apply it to risk management



Outline

Introduction to Volatility Modeling

The ARCH Model

The GARCH Model

Asymmetric GARCH Models

Model Selection and Diagnostics

Volatility Forecasting

Case Study: S&P 500

Case Study: Bitcoin

Summary



Why Model Volatility?

Empirical Observations in Financial Series

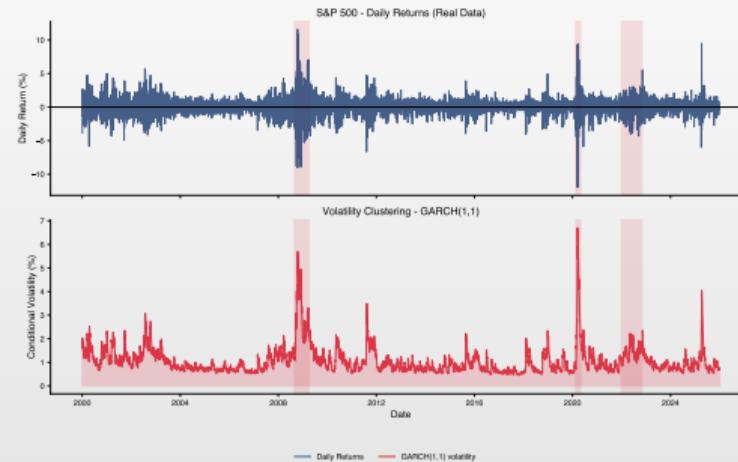
- Financial returns exhibit **volatility clustering** — periods of high volatility tend to be followed by periods of high volatility
- The distribution of returns has **fat tails** (leptokurtosis)
- Return correlation is nearly zero, but correlation of squares is significant
- Volatility responds **asymmetrically** to shocks (leverage effect)

Limitation of ARIMA Models

ARIMA models assume **constant variance** (homoskedasticity), which is not realistic for financial series!



Volatility Clustering

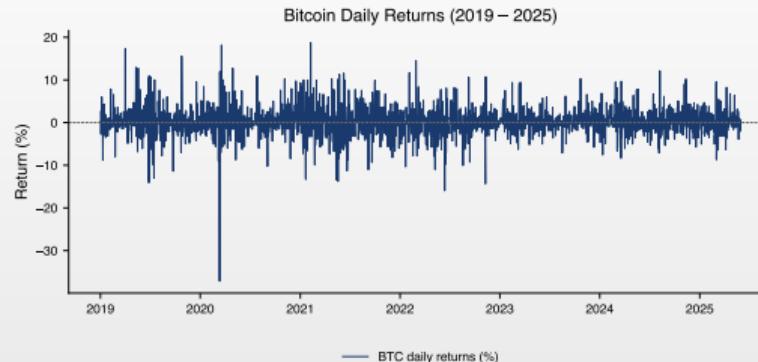


- High volatility periods are followed by high volatility; calm by calm
- This suggests that **conditional variance** is predictable

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Example: Bitcoin \succ Volatility Clustering



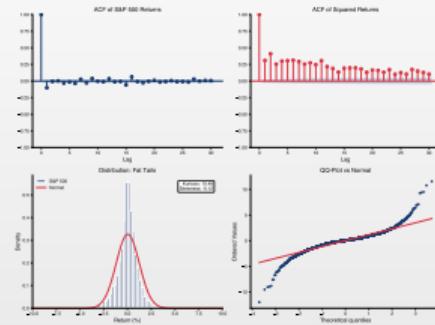
Observations

- ◻ Bitcoin daily returns (2019–2025): extremely pronounced volatility clustering
 - ▶ Returns of $\pm 20\%$ during crisis periods (COVID, Terra/Luna)
- ◻ Bitcoin volatility is significantly higher than traditional assets
 - ▶ Typical $\alpha \approx 0.10\text{--}0.20$ (fast reaction to news)

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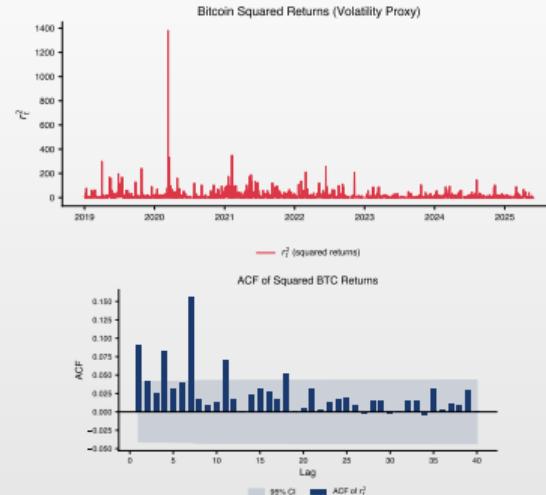
Stylized Facts of Financial Returns



Observed Properties

1. No autocorrelation in returns
2. Autocorrelation in r_t^2 , $|r_t|$
3. Fat tails ($\text{kurtosis} > 3$)
4. Leverage effect
5. Volatility clustering

Example: Bitcoin \succ Evidence for ARCH Effects



Interpretation

- Top: r_t^2 (volatility proxy) \succ peaks coincide with market crises
- Bottom: $ACF(r_t^2)$ significant \succ ARCH effects present, variance is predictable



Conditional Heteroskedasticity

Definition 1 (Conditional Variance)

For return series $\{r_t\}$, the **conditional variance** at time t is: $\sigma_t^2 = \text{Var}(r_t | \mathcal{F}_{t-1}) = \mathbb{E}[(r_t - \mu_t)^2 | \mathcal{F}_{t-1}]$ where \mathcal{F}_{t-1} is the information available up to time $t-1$.

General Model

$$r_t = \mu_t + \varepsilon_t, \quad \varepsilon_t = \sigma_t z_t, \quad z_t \sim \text{i.i.d.}(0, 1)$$

- ◻ μ_t = conditional mean (ARMA); σ_t^2 = conditional variance (GARCH)
- ◻ z_t = standardized innovations (Normal, Student-t, GED)



The ARCH(q) Model — Engle (1982)

Definition 2 (ARCH(q))

The Autoregressive Conditional Heteroskedasticity model of order q :

$$\varepsilon_t = \sigma_t z_t, \quad z_t \sim \text{i.i.d.}(0, 1), \quad \sigma_t^2 = \omega + \sum_{i=1}^q \alpha_i \varepsilon_{t-i}^2$$

Stationarity Restrictions

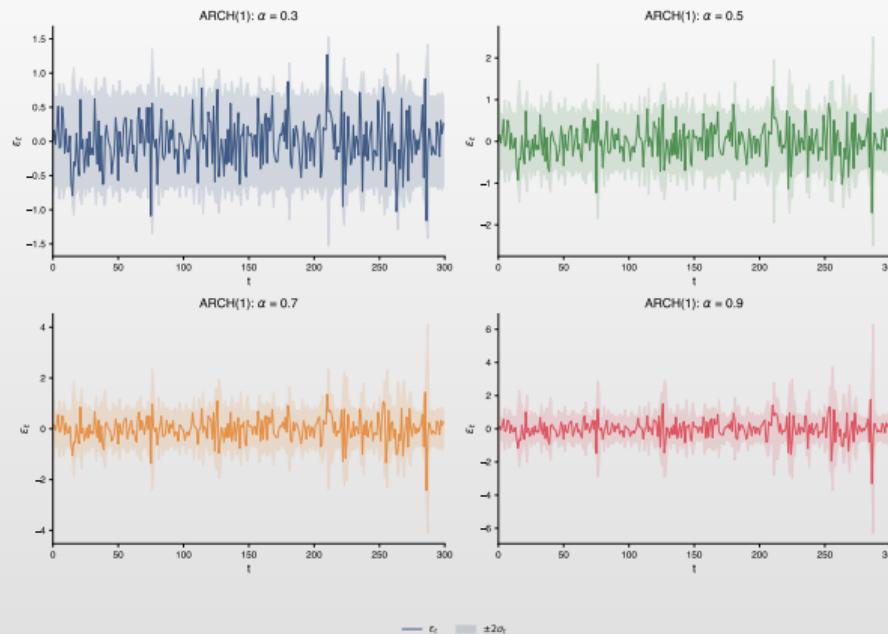
- $\omega > 0$ (positive base variance), $\alpha_i \geq 0$ (non-negativity)
- $\sum_{i=1}^q \alpha_i < 1$ (stationarity)

Remark 1

Robert Engle received the **Nobel Prize in Economics** in 2003 for developing the ARCH model!



ARCH(1) Simulation: Effect of α Parameter



Higher α means volatility reacts more strongly to recent shocks.

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Properties of the ARCH(1) Model

$$\text{ARCH}(1): \sigma_t^2 = \omega + \alpha_1 \varepsilon_{t-1}^2$$

- **Unconditional variance:** $\mathbb{E}[\varepsilon_t^2] = \frac{\omega}{1 - \alpha_1}$ (if $\alpha_1 < 1$)
- **Kurtosis:** $\kappa = 3 \cdot \frac{1 - \alpha_1^2}{1 - 3\alpha_1^2}$ (if $\alpha_1^2 < 1/3$)
- Kurtosis > 3 for $\alpha_1 > 0 \Rightarrow$ **fat tails!**

Numerical Example

If $\omega = 0.0001$ and $\alpha_1 = 0.3$:

- Unconditional variance: $\sigma^2 = \frac{0.0001}{1-0.3} = 0.000143$
- Kurtosis: $\kappa = 3 \cdot \frac{1-0.09}{1-0.27} = 3.74 > 3$



Derivation: Unconditional Variance of ARCH(1)

Derivation.

Let $\varepsilon_t = \sigma_t z_t$ with $z_t \sim N(0, 1)$ and $\sigma_t^2 = \omega + \alpha_1 \varepsilon_{t-1}^2$.

Step 1: Take unconditional expectation:

$$\mathbb{E}[\varepsilon_t^2] = \mathbb{E}[\sigma_t^2 z_t^2] = \mathbb{E}[\sigma_t^2] \cdot \mathbb{E}[z_t^2] = \mathbb{E}[\sigma_t^2]$$

Step 2: Apply expectation to variance equation:

$$\mathbb{E}[\sigma_t^2] = \mathbb{E}[\omega + \alpha_1 \varepsilon_{t-1}^2] = \omega + \alpha_1 \mathbb{E}[\varepsilon_{t-1}^2]$$

Step 3: By stationarity, $\mathbb{E}[\varepsilon_t^2] = \mathbb{E}[\varepsilon_{t-1}^2] = \sigma^2$:

$$\sigma^2 = \omega + \alpha_1 \sigma^2 \quad \Rightarrow \quad \sigma^2(1 - \alpha_1) = \omega$$

Result:
$$\boxed{\sigma^2 = \frac{\omega}{1 - \alpha_1}}$$
 (requires $\alpha_1 < 1$ for stationarity)



Derivation: Kurtosis of ARCH(1)

For $\varepsilon_t = \sigma_t z_t$ with $z_t \sim N(0, 1)$:

Step 1: $\mathbb{E}[\varepsilon_t^4] = \mathbb{E}[\sigma_t^4] \cdot \mathbb{E}[z_t^4] = 3\mathbb{E}[\sigma_t^4]$ (since $\mathbb{E}[z^4] = 3$)

Step 2: Using $\sigma_t^2 = \omega + \alpha_1 \varepsilon_{t-1}^2$:

$$\mathbb{E}[\sigma_t^4] = \mathbb{E}[(\omega + \alpha_1 \varepsilon_{t-1}^2)^2] = \omega^2 + 2\omega\alpha_1\sigma^2 + \alpha_1^2\mathbb{E}[\varepsilon_{t-1}^4]$$

Step 3: Solving the recursion yields:

$$\kappa = \frac{\mathbb{E}[\varepsilon_t^4]}{(\mathbb{E}[\varepsilon_t^2])^2} = 3 \cdot \frac{1 - \alpha_1^2}{1 - 3\alpha_1^2}$$

Interpretation

- ◻ $\kappa > 3$ for any $\alpha_1 > 0 \Rightarrow$ **fat tails** (leptokurtosis)
- ◻ Requires $\alpha_1 < 0.577$ for finite fourth moment
- ◻ ARCH naturally generates heavy-tailed distributions!



Testing for ARCH Effects

Engle's Test for ARCH Effects

Procedure:

1. Estimate the mean model and obtain residuals $\hat{\varepsilon}_t$
2. Calculate $\hat{\varepsilon}_t^2$
3. Regress $\hat{\varepsilon}_t^2$ on its lags:

$$\hat{\varepsilon}_t^2 = \beta_0 + \beta_1 \hat{\varepsilon}_{t-1}^2 + \cdots + \beta_q \hat{\varepsilon}_{t-q}^2 + u_t$$

4. Calculate the statistic $LM = T \cdot R^2 \sim \chi^2(q)$

Hypotheses

- H_0 : No ARCH effects ($\alpha_1 = \cdots = \alpha_q = 0$)
- H_1 : ARCH effects present (at least one $\alpha_i \neq 0$)



Limitations of the ARCH Model

Practical Problems

1. **High order** — many lags are usually needed (large q)
2. **Many parameters** — estimation difficulties
3. **Non-negativity constraints** — difficult to impose for large q
4. **Does not capture persistence** — observed volatility is very persistent

The Solution

The GARCH Model — introduces lags of conditional variance to capture persistence with fewer parameters!



The GARCH(p,q) Model — Bollerslev (1986)

Definition 3 (GARCH(p,q))

The Generalized ARCH model:

$$\varepsilon_t = \sigma_t z_t, \quad z_t \sim \text{i.i.d.}(0, 1)$$

$$\sigma_t^2 = \omega + \sum_{i=1}^q \alpha_i \varepsilon_{t-i}^2 + \sum_{j=1}^p \beta_j \sigma_{t-j}^2$$

Interpretation

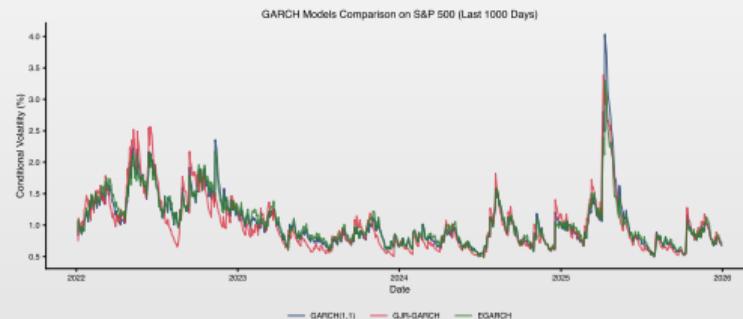
- ω = base level of volatility
- α_i = reaction to recent shocks (news coefficients)
- β_j = volatility persistence (memory)
- $\alpha + \beta$ = total persistence



The GARCH(1,1) Model

The Most Popular Volatility Model

$$\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \beta \sigma_{t-1}^2$$



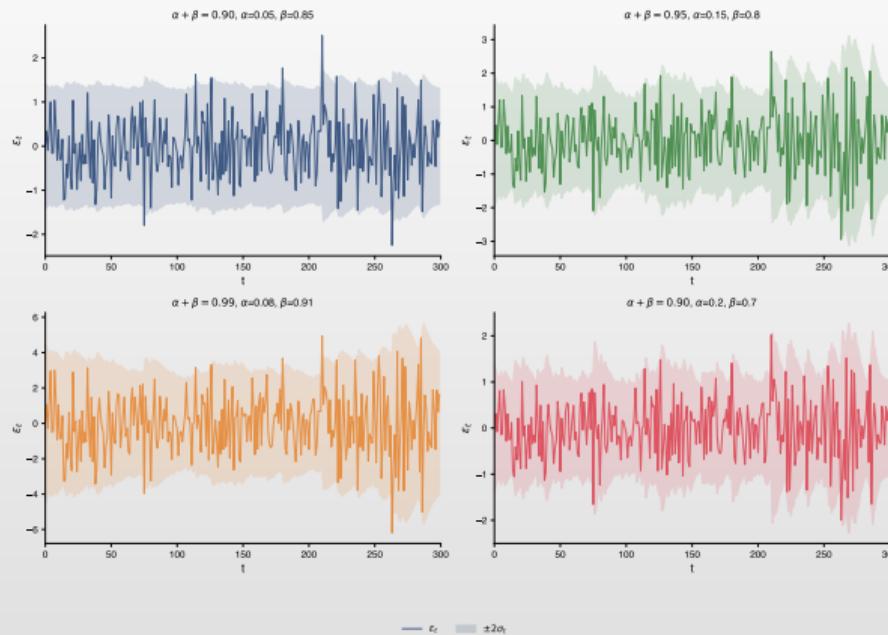
Restrictions & Properties

$\omega > 0, \alpha \geq 0, \beta \geq 0; \quad \alpha + \beta < 1$ (stationarity)

$\bar{\sigma}^2 = \frac{\omega}{1-\alpha-\beta}; \quad \text{Half-life: } HL = \frac{\ln(0.5)}{\ln(\alpha+\beta)}$



GARCH(1,1) Simulation: Persistence Effect



Parameter α controls reaction to shocks, β controls persistence. The sum $\alpha + \beta$ determines mean-reversion speed.

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Derivation: Unconditional Variance of GARCH(1,1)

Derivation.

For $\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \beta \sigma_{t-1}^2$:

Step 1: Take unconditional expectation: $\mathbb{E}[\sigma_t^2] = \omega + \alpha \mathbb{E}[\varepsilon_{t-1}^2] + \beta \mathbb{E}[\sigma_{t-1}^2]$

Step 2: By stationarity, $\mathbb{E}[\sigma_t^2] = \mathbb{E}[\sigma_{t-1}^2] = \bar{\sigma}^2$ and $\mathbb{E}[\varepsilon_t^2] = \bar{\sigma}^2$: $\bar{\sigma}^2 = \omega + (\alpha + \beta)\bar{\sigma}^2$

Step 3: Solve: $\bar{\sigma}^2(1 - \alpha - \beta) = \omega \Rightarrow \boxed{\bar{\sigma}^2 = \frac{\omega}{1 - \alpha - \beta}}$

□

Stationarity Condition

Requires $\alpha + \beta < 1$ for finite unconditional variance.



GARCH(1,1) as ARMA for ε_t^2

ARMA(1,1) Representation

Define $\nu_t = \varepsilon_t^2 - \sigma_t^2$ (variance shock). Then:

$$\varepsilon_t^2 = \omega + (\alpha + \beta)\varepsilon_{t-1}^2 + \nu_t - \beta\nu_{t-1}$$

This is an **ARMA(1,1)** for ε_t^2 !

Implications

- ACF of ε_t^2 decays exponentially (like ARMA)
- Persistence is given by $\alpha + \beta$
- PACF can help identify the order



Derivation: ARMA Representation of GARCH(1,1)

Derivation.

Step 1: Define variance shock: $\nu_t = \varepsilon_t^2 - \sigma_t^2$

- ◻ $\mathbb{E}[\nu_t | \mathcal{F}_{t-1}] = \mathbb{E}[\varepsilon_t^2 | \mathcal{F}_{t-1}] - \sigma_t^2 = \sigma_t^2 - \sigma_t^2 = 0$
- ◻ ν_t is a martingale difference sequence

Step 2: Substitute $\sigma_t^2 = \varepsilon_t^2 - \nu_t$ into GARCH equation:

$$\varepsilon_t^2 - \nu_t = \omega + \alpha \varepsilon_{t-1}^2 + \beta(\varepsilon_{t-1}^2 - \nu_{t-1})$$

Step 3: Rearrange:

$$\varepsilon_t^2 = \omega + (\alpha + \beta)\varepsilon_{t-1}^2 + \nu_t - \beta\nu_{t-1}$$

Result: ARMA(1,1) with AR coefficient $\phi = \alpha + \beta$ and MA coefficient $\theta = -\beta$. □



Volatility Persistence and Half-Life

Persistence

$\alpha + \beta$ measures mean reversion speed:

- ≈ 1 : very persistent
- $\ll 1$: quick reversion

Half-Life Formula

$$HL = \frac{-0.693}{\ln(\alpha + \beta)}$$

Example: S&P 500

$$\alpha = 0.09, \beta = 0.90$$

$$\alpha + \beta = 0.99$$

$$HL = \frac{-0.693}{\ln(0.99)} \approx 69$$

Shock halves in ~ 69 trading days!



Estimation of GARCH Models

Maximum Likelihood Estimation (MLE)

$$\text{Log-likelihood (normal): } \ell(\theta) = -\frac{T}{2} \ln(2\pi) - \frac{1}{2} \sum_{t=1}^T \left[\ln(\sigma_t^2) + \frac{\varepsilon_t^2}{\sigma_t^2} \right]$$

Alternative Distributions for z_t

- Student-t**: captures fat tails — most common choice
- GED**: flexibility for kurtosis
- Skewed Student-t**: asymmetry and fat tails

Practical Note

Student-t distribution typically provides better fit for financial returns due to fat tails ($\text{kurtosis} > 3$).



Typical Values for GARCH(1,1)

Series	α	β	$\alpha + \beta$
S&P 500 daily	0.05–0.10	0.85–0.95	0.95–0.99
EUR/USD daily	0.03–0.08	0.90–0.95	0.95–0.99
Bitcoin daily	0.10–0.20	0.75–0.85	0.90–0.98
Bonds	0.02–0.05	0.90–0.97	0.95–0.99

Observations

- $\alpha + \beta$ close to 1 \Rightarrow very persistent volatility
- Small α , large β \Rightarrow slow reaction to shocks, long memory
- Bitcoin: larger α \Rightarrow faster reaction to news



IGARCH — Integrated GARCH

Definition 4 (IGARCH(1,1))

When $\alpha + \beta = 1$:

$$\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + (1 - \alpha) \sigma_{t-1}^2$$

Properties

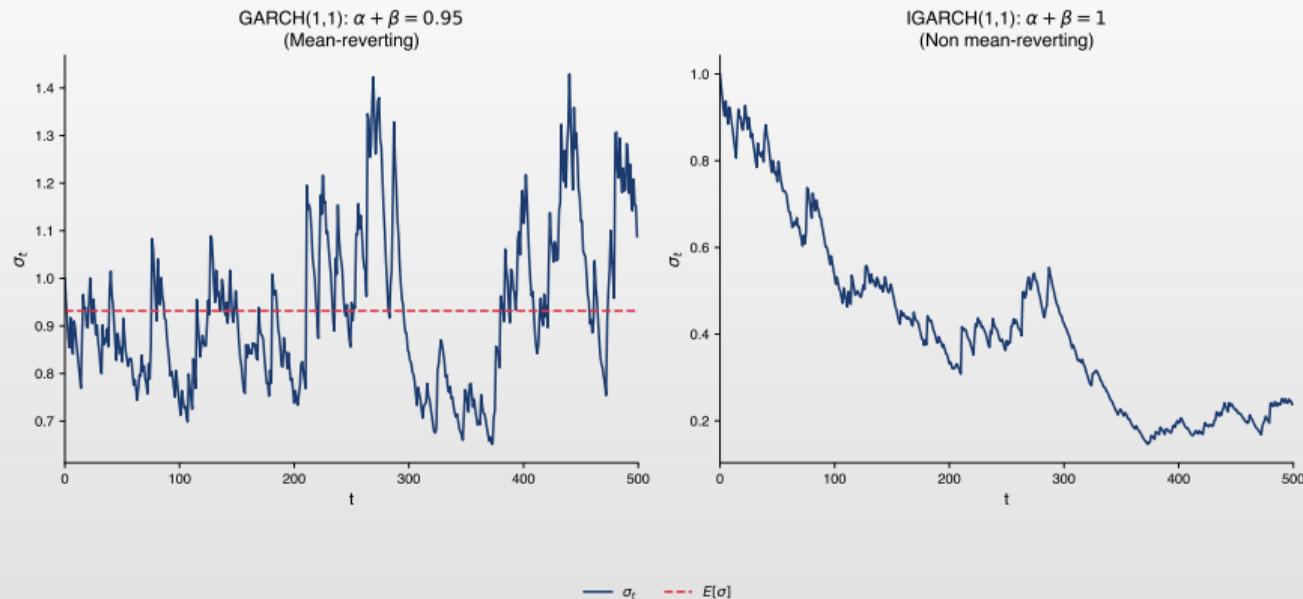
- Unconditional variance does not exist (infinite)
- Shocks have **permanent** effect on volatility
- Used for series with extreme persistence
- Useful for **RiskMetrics** (J.P. Morgan): $\alpha = 0.06$, $\beta = 0.94$

Remark 2

IGARCH is analogous to a unit root in variance!



GARCH vs IGARCH: Persistence Comparison

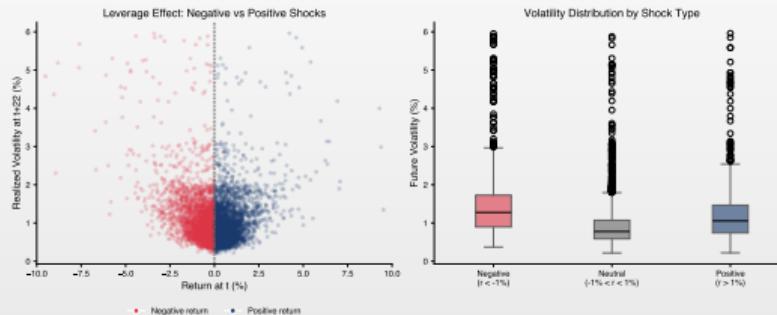


Standard GARCH reverts to unconditional mean, while IGARCH has no finite mean and shocks persist indefinitely.

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Leverage Effect



Definition

Leverage effect: Negative shocks increase volatility **more** than positive shocks of the same magnitude.

Problem with GARCH

Standard GARCH: $\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \beta \sigma_{t-1}^2$ — only ε_{t-1}^2 matters, sign is lost! Economic intuition: Bad news \Rightarrow stock price falls \Rightarrow debt/equity ratio rises \Rightarrow volatility increases.



The EGARCH Model — Nelson (1991)

Definition 5 (EGARCH(1,1))

Exponential GARCH:

$$\ln(\sigma_t^2) = \omega + \alpha(|z_{t-1}| - \mathbb{E}[|z_{t-1}|]) + \gamma z_{t-1} + \beta \ln(\sigma_{t-1}^2)$$

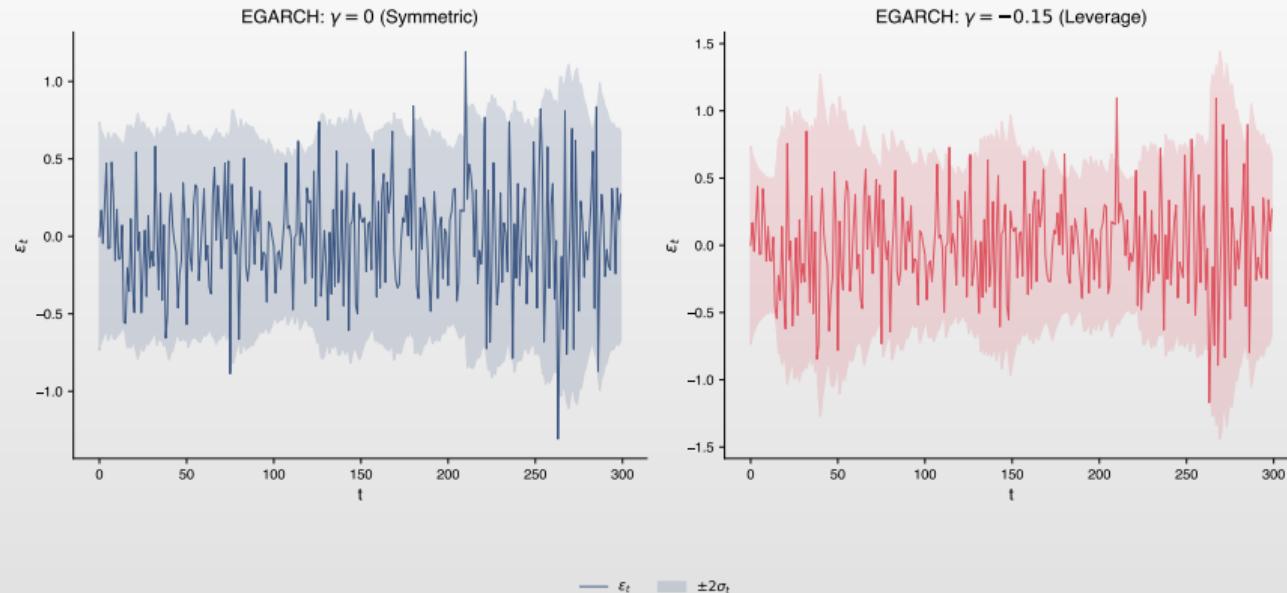
where $z_t = \varepsilon_t / \sigma_t$.

EGARCH Advantages

- **No non-negativity constraints required** — models $\ln(\sigma_t^2)$
- **Captures leverage effect** through parameter γ
 - ▶ $\gamma < 0$: negative shocks \Rightarrow higher volatility
 - ▶ $\gamma = 0$: symmetric effect (like GARCH)
- Persistence is given by β



EGARCH Simulation: Symmetric vs Asymmetric

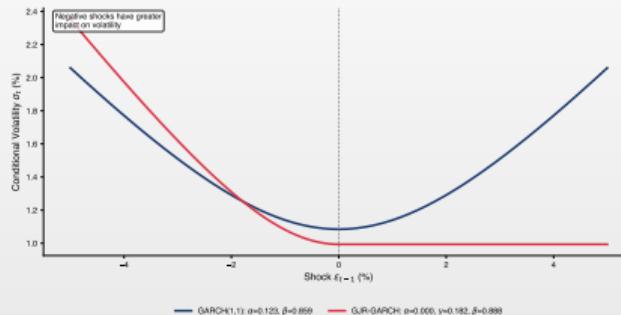


When $\gamma < 0$, negative shocks (bad news) increase volatility more than positive shocks of the same magnitude.

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News Impact Curve — EGARCH



Definition

News Impact Curve: σ_{t+1}^2 as function of ε_t , holding σ_t^2 constant.

- **GARCH:** Symmetric V-shape (parabola); **EGARCH:** Asymmetric — steeper for negative shocks; **GJR:** Piecewise linear with kink at zero
- The asymmetry captures the leverage effect: bad news has larger impact on volatility than good news.

The GJR-GARCH Model

Definition 6 (GJR-GARCH(1,1))

Glosten, Jagannathan & Runkle (1993): $\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \gamma \varepsilon_{t-1}^2 \cdot I_{t-1} + \beta \sigma_{t-1}^2$ where $I_{t-1} = 1$ if $\varepsilon_{t-1} < 0$, else 0.

Interpretation

- Positive shocks: impact = α ; Negative shocks: impact = $\alpha + \gamma$
- Leverage effect present if $\gamma > 0$
- Stationarity: $\alpha + \gamma/2 + \beta < 1$



TGARCH — Threshold GARCH

Definition 7 (TGARCH(1,1))

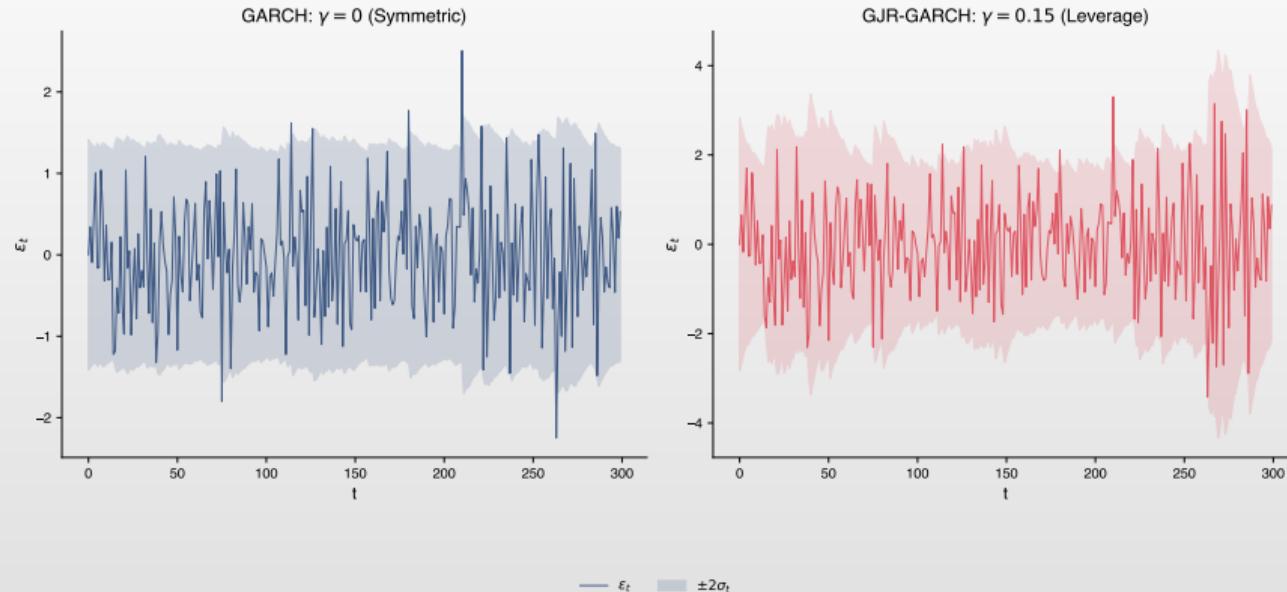
Zakoian (1994) models standard deviation: $\sigma_t = \omega + \alpha^+ \varepsilon_{t-1}^+ + \alpha^- \varepsilon_{t-1}^- + \beta \sigma_{t-1}$

Comparison of Asymmetric Models

Model	Specification	Leverage
GARCH	σ_t^2	No
EGARCH	$\ln(\sigma_t^2)$	Yes ($\gamma < 0$)
GJR-GARCH	σ_t^2 with indicator	Yes ($\gamma > 0$)
TGARCH	σ_t	Yes ($\alpha^- > \alpha^+$)



GJR-GARCH/TGARCH Simulation

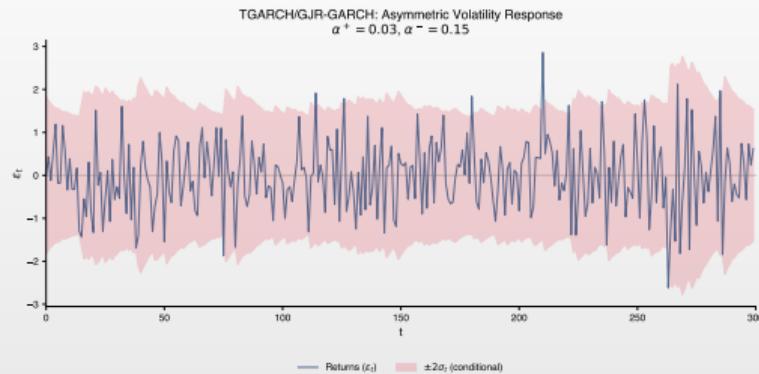


GJR-GARCH adds an indicator term to capture asymmetric response to negative shocks.

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TGARCH Simulation: Asymmetric Volatility Response



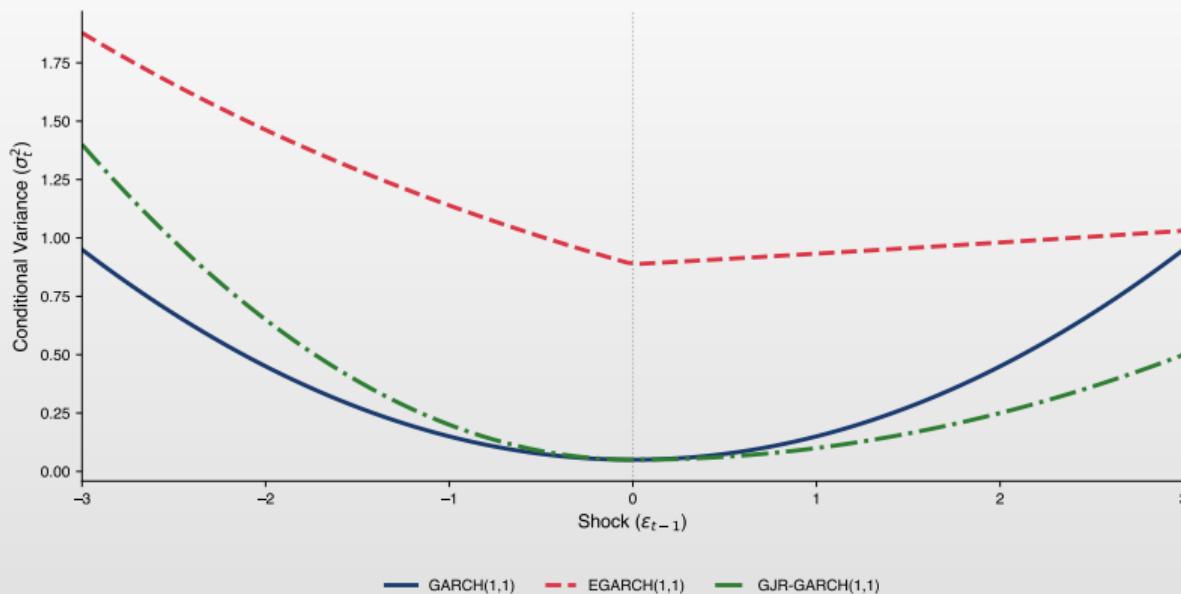
Interpretation

- ◻ TGARCH with $\alpha^+ = 0.03$ and $\alpha^- = 0.15 \succ$ negative shocks amplify volatility by 5×
- ◻ Volatility bands $\pm 2\sigma$ widen asymmetrically during crisis periods

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News Impact Curves Comparison

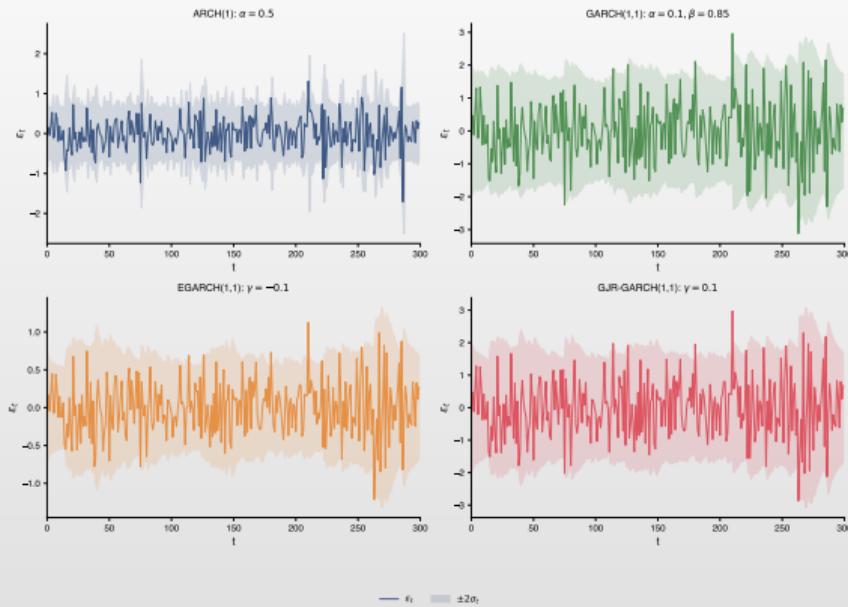


Standard GARCH is symmetric, while EGARCH and GJR-GARCH capture asymmetry (leverage effect).

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GARCH Family Comparison



All models capture volatility clustering, but differ in how they model asymmetry.

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GARCH-M: GARCH-in-Mean Model

Definition 8 (GARCH-M)

The **GARCH-in-Mean** model: $r_t = \mu + \lambda\sigma_t + \varepsilon_t$, where λ is the **risk premium**.

Interpretation

- $\lambda > 0$: Higher risk \Rightarrow higher expected return
- $\lambda = 0$: Reduces to standard GARCH
- $\lambda < 0$: Higher risk \Rightarrow lower return (rare)

Financial Intuition

Investors demand compensation for bearing risk — GARCH-M captures this **risk-return tradeoff**.



GARCH-M: Alternative Specifications

Common Specifications

Risk premium can enter in different forms: (1) $r_t = \mu + \lambda\sigma_t + \varepsilon_t$; (2) $r_t = \mu + \lambda\sigma_t^2 + \varepsilon_t$; (3) $r_t = \mu + \lambda \ln(\sigma_t^2) + \varepsilon_t$

Typical Results for Equity Markets

- Estimated λ often positive but small (0.01–0.10)
- Significance varies across markets and periods
- Variance specification yields larger λ estimates

Remark 3

GARCH-M is used in asset pricing, portfolio optimization, and CAPM testing.



Order Selection

Information Criteria

- AIC** = $-2\ell + 2k$
- BIC** = $-2\ell + k \ln(T)$
- HQIC** = $-2\ell + 2k \ln(\ln(T))$

where ℓ = maximized log-likelihood, k = number of parameters.

Practical Recommendations

- GARCH(1,1) is sufficient in **90% of cases**
- Check if asymmetric model significantly improves fit
- Choose innovation distribution that minimizes AIC/BIC



GARCH Model Diagnostics

Standardized Residuals

$$\hat{z}_t = \frac{\hat{\epsilon}_t}{\hat{\sigma}_t}$$

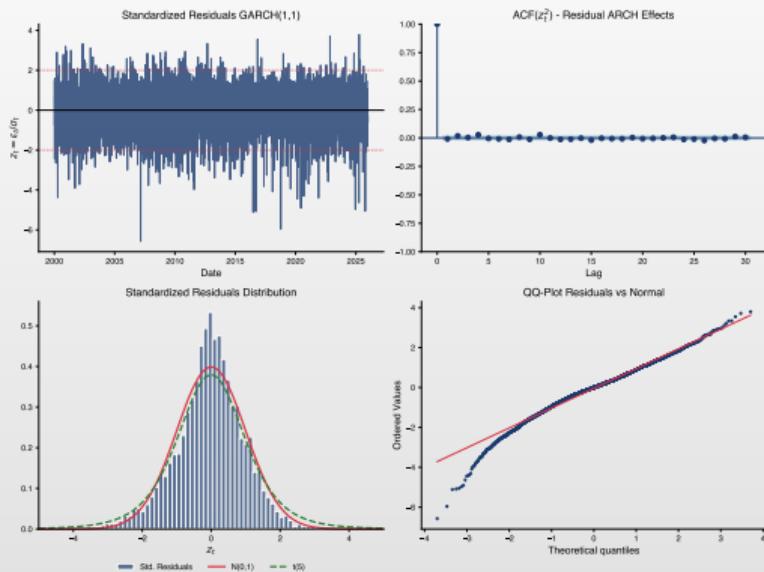
If the model is correctly specified, \hat{z}_t should be i.i.d.(0,1).

Diagnostic Checks

1. **Ljung-Box on \hat{z}_t :** check absence of autocorrelation in mean
2. **Ljung-Box on \hat{z}_t^2 :** check absence of residual ARCH effects
3. **ARCH-LM test on \hat{z}_t :** confirm absence of heteroskedasticity
4. **Histogram + QQ-plot:** verify assumed distribution



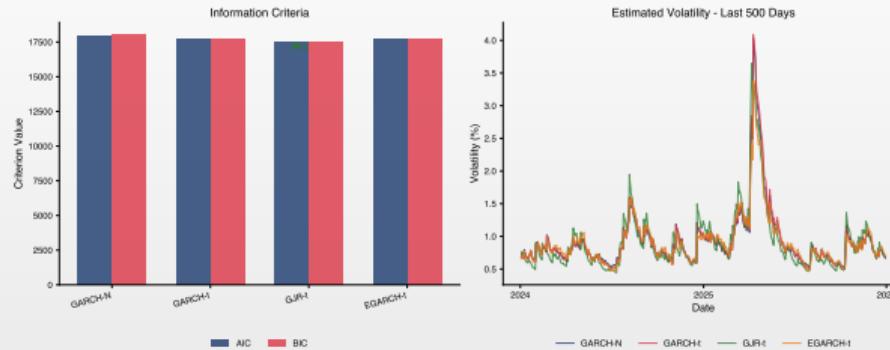
Diagnostic Example



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GARCH Model Comparison: Validation



Interpretation

- GARCH(1,1) achieves the lowest MAE on the validation set
 - ▶ More parsimonious and stable than higher-order models
- GARCH(2,1) and GJR-GARCH: similar performance, but more parameters
- **Conclusion:** simplicity wins \succ GARCH(1,1) is hard to beat



Forecasting with GARCH(1,1)

One-Step-Ahead Forecast

$$\hat{\sigma}_{T+1}^2 = \omega + \alpha \varepsilon_T^2 + \beta \sigma_T^2$$

Multi-Step Forecast

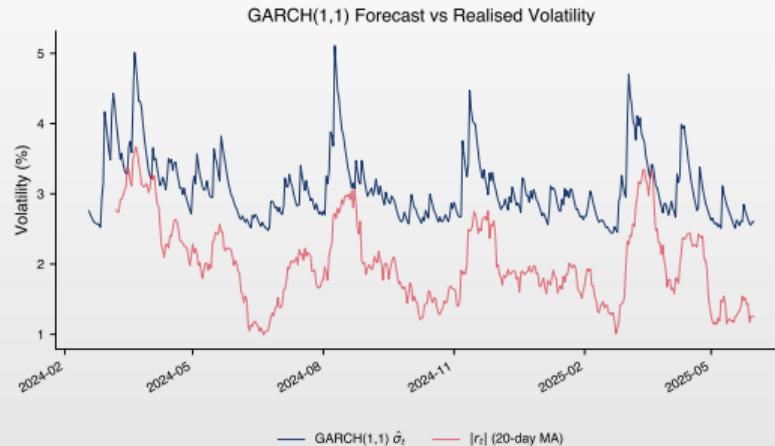
For $h > 1$: $\mathbb{E}_T[\sigma_{T+h}^2] = \bar{\sigma}^2 + (\alpha + \beta)^{h-1}(\sigma_{T+1}^2 - \bar{\sigma}^2)$ where $\bar{\sigma}^2 = \frac{\omega}{1-\alpha-\beta}$ = unconditional variance.

Convergence

$\lim_{h \rightarrow \infty} \mathbb{E}_T[\sigma_{T+h}^2] = \bar{\sigma}^2$ — forecast converges to unconditional variance!



Volatility Forecast — Visualization

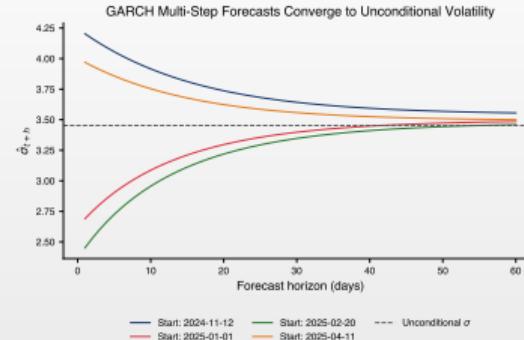


- ◻ Forecast converges exponentially to $\bar{\sigma}^2$; speed depends on $\alpha + \beta$
- ◻ The closer $\alpha + \beta$ is to 1, the slower the convergence

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GARCH Forecast Convergence to Unconditional Variance



Interpretation

- Multi-step forecast converges exponentially to $\bar{\sigma}^2 = \frac{\omega}{1-\alpha-\beta}$
- The closer $\alpha + \beta$ is to 1, the slower the convergence
 - ▶ S&P 500: $\alpha + \beta \approx 0.99 \succ$ convergence in ~ 50 days
 - ▶ Bitcoin: $\alpha + \beta \approx 0.95 \succ$ faster convergence

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Applications of Volatility Forecasting

Value at Risk (VaR)

$$\text{VaR}_\alpha = z_\alpha \cdot \sigma_{T+1}$$

Maximum loss with probability $1 - \alpha$.

Expected Shortfall (ES)

$$\text{ES}_\alpha = \mathbb{E}[-r | r < -\text{VaR}_\alpha]$$

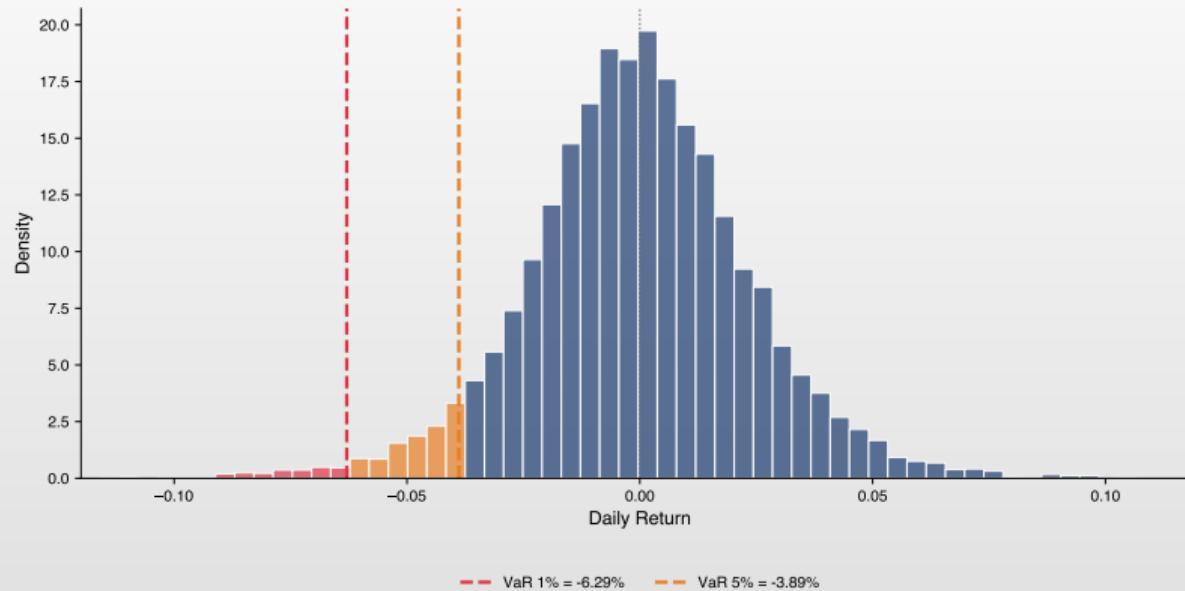
Average loss when VaR is exceeded.

Other Applications

- Option pricing
- Dynamic hedging
- Portfolio allocation
- Stress testing



VaR and ES: Graphical Illustration

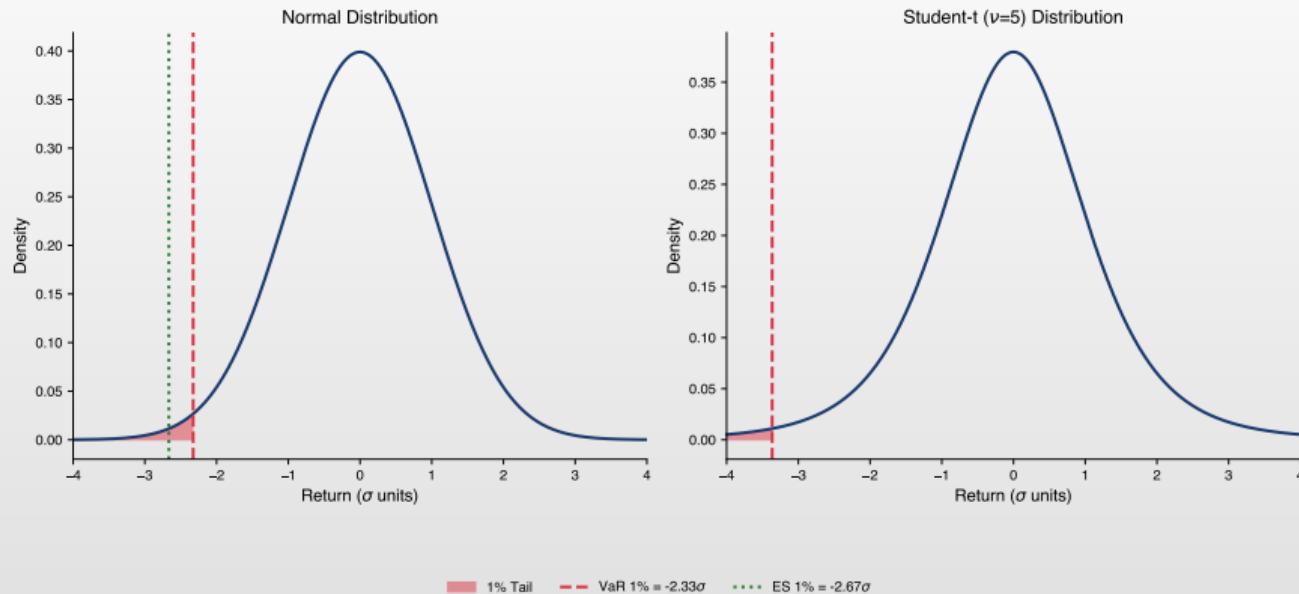


VaR 1% = loss exceeded only in 1% of cases. Red area = left tail (extreme losses).

Q TSA_ch5_var_plot



VaR vs Expected Shortfall: Normal vs Student-t



ES (green line) measures average loss when VaR is exceeded. Student-t has heavier tails \Rightarrow larger VaR and ES. Q

TSA_ch5_var_es



Value at Risk — Numerical Example

VaR Calculation

Portfolio: 1,000,000 EUR, forecasted volatility $\hat{\sigma}_{T+1} = 1.5\%$

VaR with Normal Distribution

Level	z_α	VaR (%)	VaR (EUR)
5% (1 day)	1.645	2.47%	24,675
1% (1 day)	2.326	3.49%	34,890

Scaling for Longer Periods

$$\text{VaR}_{h \text{ days}} = \text{VaR}_{1 \text{ day}} \cdot \sqrt{h} \text{ — assumes i.i.d. returns}$$



Value at Risk — Student-t Distribution

Why Student-t?

Normal distribution **underestimates** tail risk. Student-t with ν degrees of freedom better captures fat tails (kurtosis > 3).

VaR 1% (1 day) Comparison: $\sigma = 1.5\%$, Portfolio = 1M EUR

Distribution	Quantile	VaR (EUR)
Normal	2.326	34,890
Student-t ($\nu = 6$)	3.143	47,145
Student-t ($\nu = 4$)	3.747	56,205

Observation

With $\nu = 6$ (typical for stocks), VaR is **35% higher** than normal!



VaR — Complete Example with GARCH

VaR Calculation Procedure

1. Estimate GARCH(1,1) model with Student-t distribution
2. Obtain volatility forecast: $\hat{\sigma}_{T+1}$
3. Calculate VaR: $\text{VaR}_\alpha = t_\alpha(\nu) \cdot \hat{\sigma}_{T+1} \cdot \sqrt{\frac{\nu-2}{\nu}}$

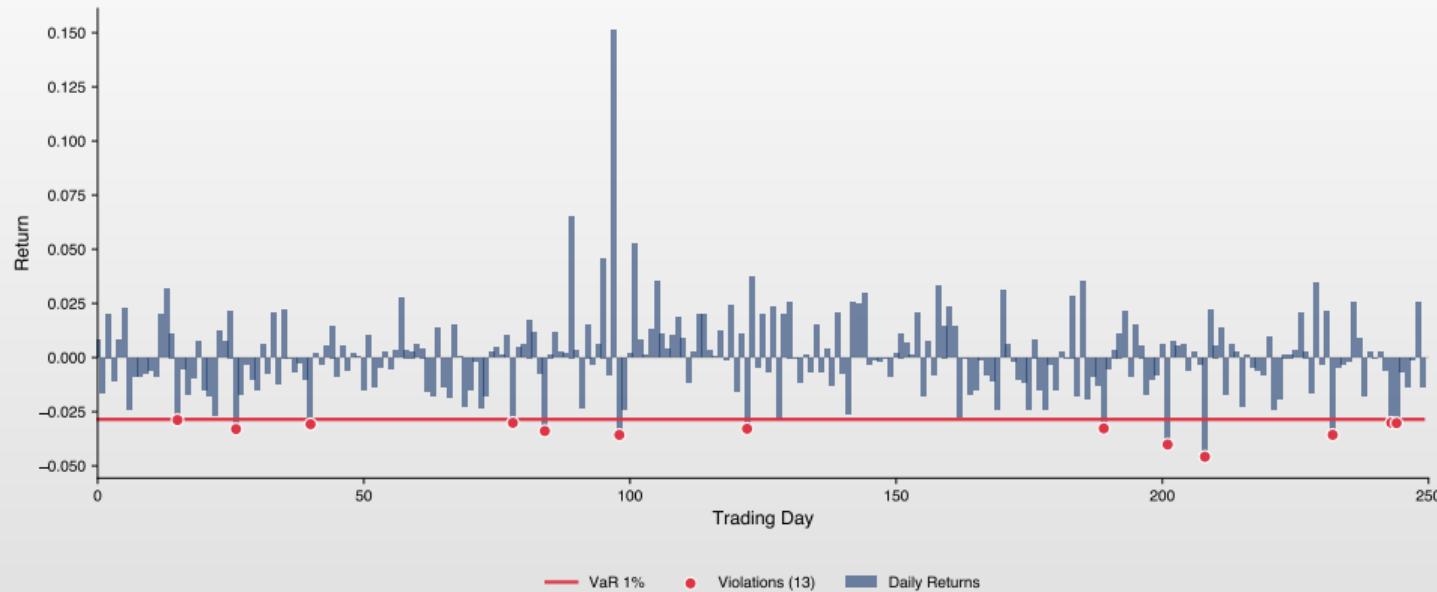
Example: S&P 500

- Estimated parameters: $\alpha = 0.088$, $\beta = 0.900$, $\nu = 6.4$
- Forecasted volatility: $\hat{\sigma}_{T+1} = 1.2\%$
- Portfolio: 10,000,000 EUR

VaR 1% (1 day): $\text{VaR} = 3.05 \times 0.012 \times 10,000,000 = \mathbf{366,000 \text{ EUR}}$



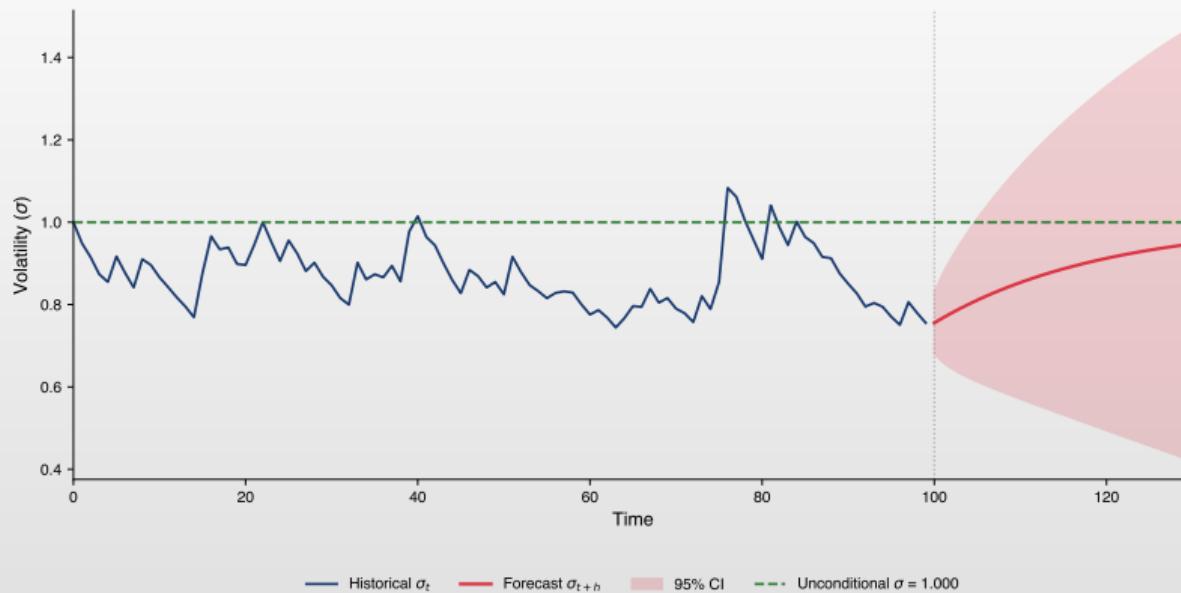
VaR Backtesting: Visual Overview



Backtesting checks if VaR violations match expected rate (e.g., 2.5 violations/year for VaR 1%). [Q_TSA_ch5_backtest](#)



Volatility Forecast with Confidence Intervals

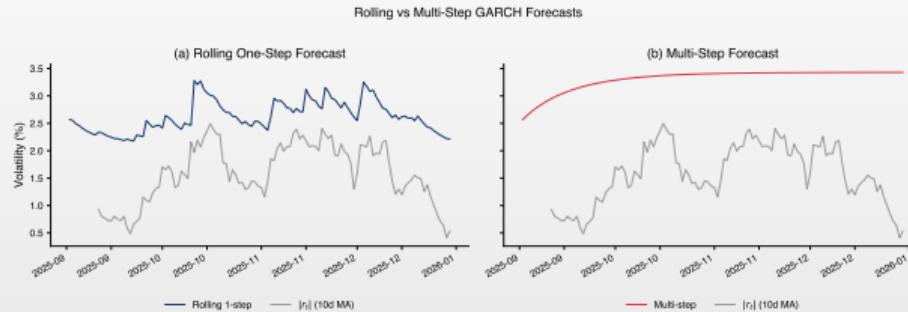


Forecast converges to unconditional volatility $\bar{\sigma}$. Uncertainty increases with forecast horizon.

Q TSA_ch5_vol_ci



Rolling Forecast: Step-by-Step Prediction



Procedure

S&P 500, W=500, GARCH(1,1)-t

- Re-estimate GARCH on $[t-W, t-1]$; forecast $\hat{\sigma}_{t|t-1}$
- Compare with realized vol. (20-day rolling std.)

Results (2015 days OOS)

- $\rho = 0.938 \succ$ excellent tracking; MAE = 0.15%, RMSE = 0.24%

Time ~~COVID-19~~ system ~~temporally~~ under prediction, rapid adaptation



GARCH Estimation in Python: arch Package

Python Code

```
pip install arch
from arch import arch_model

model = arch_model(returns,
    vol='Garch', p=1, q=1,
    dist='normal')
results = model.fit(disp='off')
print(results.summary())
```

Key Parameters

- **vol:** model type
 - ▶ 'Garch', 'EGARCH'
- **p, q:** GARCH order
 - ▶ p=1, q=1 standard
- **dist:** distribution
 - ▶ 'normal', 't'

Q TSA_ch5_garch



Asymmetric Models in Python

EGARCH and GJR-GARCH

```
# EGARCH
model_egarch = arch_model(returns, vol='EGARCH', p=1, q=1)
# GJR-GARCH (o=1 adds the asymmetric term)
model_gjr = arch_model(returns, vol='Garch', p=1, o=1, q=1)
```

Alternative Distributions

```
# Student-t for fat tails
model_t = arch_model(returns, vol='Garch', p=1, q=1, dist='t')
# Skewed Student-t for asymmetry and fat tails
model_skewt = arch_model(returns, vol='Garch', p=1, q=1,
                         dist='skewt')
```

Q TSA_ch5_egarch



Forecasting and Diagnostics

Volatility Forecast

```
forecasts = results.forecast(horizon=10)
vol_forecast = np.sqrt(forecasts.variance.values[-1,:])
```

Diagnostics and VaR

```
std_resid = results.std_resid
lb_test = acorr_ljungbox(std_resid**2, lags=10)
sigma = np.sqrt(forecasts.variance.values[-1, 0])
VaR_5pct = 1.645 * sigma
```

 TSA_ch5_forecast



VaR Backtesting: Kupiec Test

Unconditional Coverage Test

Tests whether the observed violation rate equals the expected rate p (e.g., 1% for VaR 1%).

Let N = number of VaR violations, T = total observations, $\hat{p} = N/T$.

Likelihood Ratio Statistic:

$$LR_{uc} = -2 \ln \left[\frac{(1-p)^{T-N} p^N}{(1-\hat{p})^{T-N} \hat{p}^N} \right] \sim \chi^2(1)$$

Hypotheses

- $H_0: \hat{p} = p$ (VaR model is correctly calibrated)
- $H_1: \hat{p} \neq p$ (VaR model under- or over-estimates risk)



VaR Backtesting: Christoffersen Test

Conditional Coverage Test

Tests both **unconditional coverage** and **independence** of violations.

Violations should be independent — no clustering of exceptions!

Test Components

- Independence test (LR_{ind}): Tests if violations are serially independent
- Conditional coverage: $LR_{cc} = LR_{uc} + LR_{ind} \sim \chi^2(2)$

Interpretation

Reject LR_{uc} : wrong frequency; Reject LR_{ind} : clustered violations; Reject LR_{cc} : model fails



VaR Backtesting: Python Implementation

Kupiec Test Implementation

```
import numpy as np
from scipy import stats
def kupiec_test(violations, T, p=0.01):
    N = np.sum(violations)
    p_hat = N / T
    if N == 0 or N == T:
        return np.nan, np.nan
    LR = -2 * (np.log((1-p)**(T-N) * p**N) -
               np.log((1-p_hat)**(T-N) * p_hat**N))
    return LR, 1 - stats.chi2.cdf(LR, df=1)
```

Usage

```
LR, pval = kupiec_test(violations, T=250, p=0.01)
```

Full Backtesting: Results and Decision

Application S&P 500 (T=500, VaR 1%)

```
violations = returns[window:] < -VaR_series
n_viol = violations.sum()
T = len(violations)
rate = n_viol / T
print(f"Violations: {n_viol}/{T} (rate = {rate:.2%})")
LR_uc, p_uc = kupiec_test(violations, T, alpha=0.01)
LR_ind, p_ind = christoffersen_test(violations)
LR_cc = LR_uc + LR_ind # combined test ~ chi2(2)
p_cc = 1 - stats.chi2.cdf(LR_cc, df=2)
```

Typical Output

```
Violations: 13/500 (rate = 2.60%)
Kupiec    LR = 5.83,  p-value = 0.0157 => Rejected (p<0.05)
Independ. LR = 0.42,  p-value = 0.5171 => Accepted
Combined  LR = 6.25,  p-value = 0.0439 => Rejected
Basel Zone: RED (>=10 violations) => Inadequate model
```



ARMA-GARCH: Joint Mean and Variance Modeling

Why Joint Modeling?

Serial correlation \Rightarrow ARMA for mean; **Volatility clustering** \Rightarrow GARCH for variance.

Definition 9 (ARMA(p,q)-GARCH(r,s))

Mean equation: $r_t = \mu + \sum_{i=1}^p \phi_i(r_{t-i} - \mu) + \sum_{j=1}^q \theta_j \varepsilon_{t-j} + \varepsilon_t$

Variance equation: $\varepsilon_t = \sigma_t z_t, \quad \sigma_t^2 = \omega + \sum_{i=1}^r \alpha_i \varepsilon_{t-i}^2 + \sum_{j=1}^s \beta_j \sigma_{t-j}^2$



ARMA-GARCH: Model Selection Strategy

Step-by-Step Approach

1. **Identify mean model:** Check ACF/PACF of returns for ARMA structure
2. **Test for ARCH effects:** Apply ARCH-LM test to residuals
3. **Specify variance model:** Usually GARCH(1,1) is sufficient
4. **Joint estimation:** Estimate both equations via MLE
5. **Diagnostic checking:** Standardized residuals should be i.i.d.

Common Specifications

- Stock returns:** AR(1)-GARCH(1,1) or ARMA(1,1)-GARCH(1,1)
- Exchange rates:** Often just GARCH(1,1) (no mean dynamics)
- Interest rates:** AR(1)-EGARCH(1,1) for leverage effects



ARMA-GARCH: Python Implementation

Using the arch Package

```
from arch import arch_model
model = arch_model(returns,
                    mean='ARX',
                    lags=1,
                    vol='Garch',
                    p=1, q=1,
                    dist='t')
result = model.fit(disp='off')
print(result.summary())
```

Parameters

mean='ARX': ARMA mean; lags=1: AR(1); dist='t': Student-t



ARMA-GARCH: Complete Example

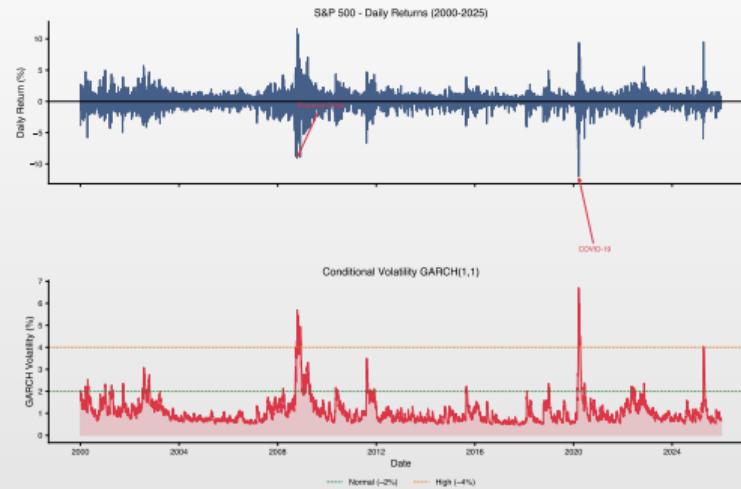
```
from arch import arch_model
model = arch_model(returns,
                    mean='ARX', lags=[1],
                    vol='EGARCH', p=1, q=1,
                    dist='skewt')
result = model.fit(update_freq=0)
cond_mean = result.conditional_mean
cond_vol = result.conditional_volatility
forecasts = result.forecast(horizon=5)
```

Note

For MA terms, use `mean='HARX'` or pre-filter with statsmodels ARIMA.



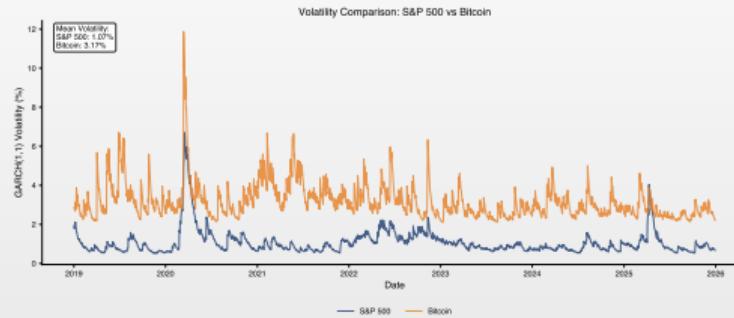
S&P 500 Volatility Analysis



- S&P 500 daily returns (2000–2024) — volatility clustering visible
- Crisis periods: 2008 (financial), 2020 (COVID-19), 2022 (inflation)



GARCH(1,1) Estimation — S&P 500

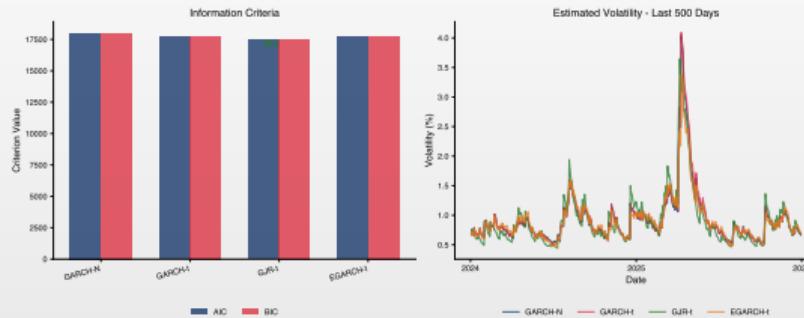


Estimation Results

Parameter	Value
ω	0.0108
α	0.0883
β	0.9002
$\alpha + \beta$	0.9885
ν (df)	6.42



GARCH vs EGARCH Comparison — S&P 500



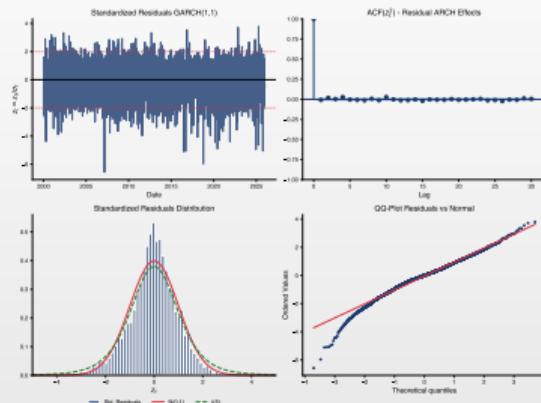
Leverage Effect Confirmed

EGARCH: $\gamma = -0.12$ (significantly negative) — negative shocks amplify volatility more than positive shocks. Both models capture volatility clustering, but EGARCH better fits crisis periods (2008, 2020).

Q TSA_ch5_sp500_comp



Step 5: Diagnostics — EGARCH(1,1)-t

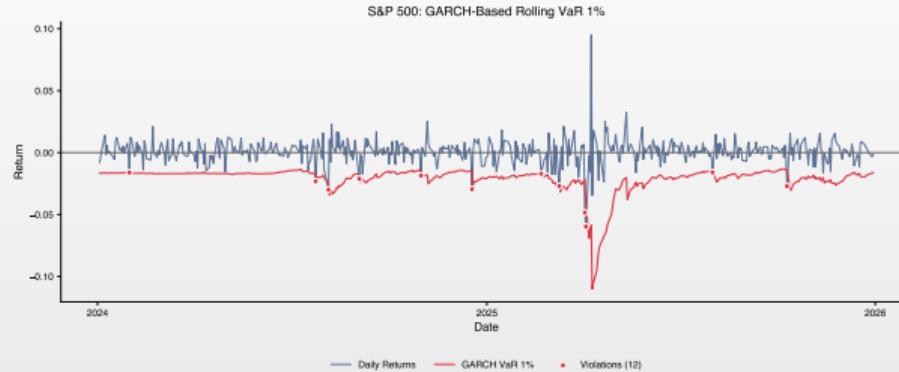


Checks on Standardized Residuals $z_t = \varepsilon_t / \hat{\sigma}_t$

- **Ljung-Box** on z_t : p-value = 0.38 — no residual autocorrelation
- **Ljung-Box** on z_t^2 : p-value = 0.52 — **ARCH effects eliminated**
- **Q-Q plot**: points follow the theoretical Student-t line
- **Conclusion**: EGARCH(1,1)-t adequately captures volatility dynamics



Step 6: Backtesting Rolling VaR — S&P 500

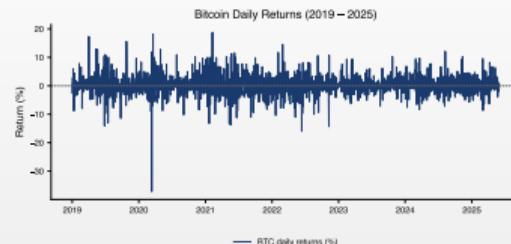


Kupiec + Christoffersen Results (2015 days out-of-sample)

Test	Statistic	p-value	Decision
Violations	27/2015 ($\hat{p} = 1.34\%$)	—	Green zone
Kupiec (uc)	2.13	0.145	Accepted
Christoffersen (ind)	0.79	0.375	Accepted
Combined (cc)	2.91	0.233	Accepted



Step 1: Data — Bitcoin Daily Returns



Data Description

- Source: Yahoo Finance (BTC-USD), daily data 2018–2024
- Log returns: mean $\approx 0.05\%$, volatility $\approx 3.5\%$

Descriptive Statistics

Mean	Std. Dev.	Skewness	Kurtosis	Min	Max
0.05%	3.48%	-0.72	12.1	-46.5%	+22.5%

- Volatility $\sim 3\times$ higher than S&P 500
- Extreme kurtosis — high risk of large losses

Steps 3–4: Estimation and Model Selection — Bitcoin

Estimated Parameters

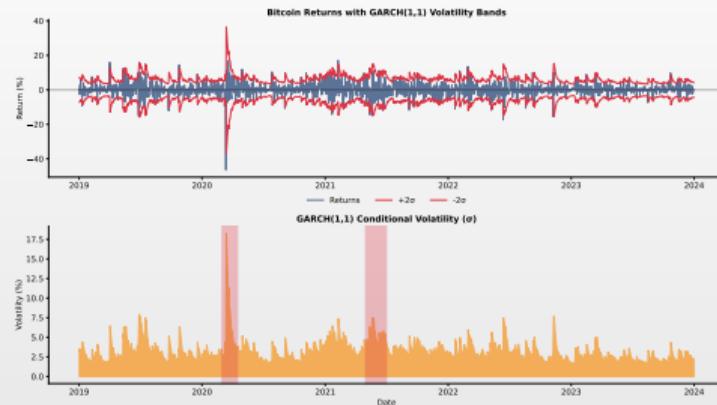
Model	ω	α	β	γ	$\alpha+\beta$	ν	AIC
GARCH-t	0.42	0.131	0.848	—	0.979	4.82	9284
EGARCH-t	0.08	0.184	0.976	-0.061	—	4.79	9276
GJR-t	0.40	0.088	0.854	0.078	0.976	4.85	9271

Interpretation

- GJR-GARCH-t wins** (lowest AIC)
- $\nu \approx 4.8$: **much heavier tails** than S&P 500 ($\nu = 6.4$)
- $\alpha = 0.131$ (BTC) vs 0.088 (S&P) — Bitcoin reacts faster to news
- Leverage effect weaker than for stocks ($\gamma_{\text{BTC}} = 0.078$ vs 0.126)



Step 5: Conditional Volatility — Bitcoin

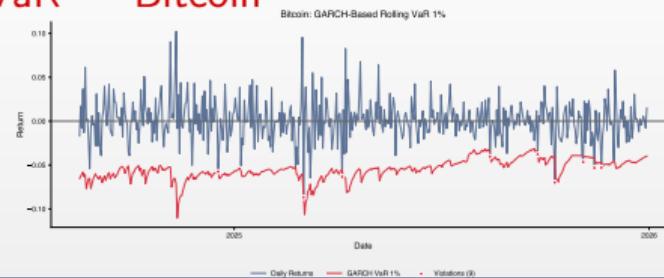


GJR-GARCH(1,1)-t Diagnostics

- ◻ Ljung-Box on z_t^2 : p-value = 0.41 — ARCH effects eliminated
- ◻ Volatility peaks: March 2020 (COVID), May 2022 (Terra/Luna)
- ◻ Daily volatility: from 1% (calm periods) to >15% (crises)



Step 6: Backtesting Rolling VaR — Bitcoin



Statistical Tests (2421 days out-of-sample)

Test	Statistic	p-value	Decision
Violations	28/2421 ($\hat{\rho} = 1.16\%$)	—	Green zone
Kupiec (uc)	0.57	0.450	Accepted
Christoffersen (ind)	0.94	0.333	Accepted
Combined (cc)	1.51	0.471	Accepted

Interpretation

- Volatility ranges from 3% to 38% — rolling window is **essential**
- All tests **accepted**: model valid for risk management



Final Comparison: S&P 500 vs Bitcoin

Comparative Summary

	S&P 500	Bitcoin
Average volatility	1.2%	3.5%
Kurtosis	13.8	12.1
Student-t ν	6.42	4.82
Best model	EGARCH(1,1)-t	GJR-GARCH(1,1)-t
Leverage effect	Strong ($\gamma = -0.12$)	Moderate ($\gamma = 0.078$)
Half-life	~60 days	~42 days
Rolling VaR 1% mean	2.53%	9.34%
Rolling VaR 1% max	22.02% (COVID)	37.54% (COVID)
Kupiec	Accepted ($p=0.145$)	Accepted ($p=0.450$)
Christoffersen (ind)	Accepted ($p=0.375$)	Accepted ($p=0.333$)

General Conclusion

- ☐ Re-estimating GARCH at each step: Kupiec + Christoffersen **accepted**
- ☐ Rolling window VaR: **mandatory** — static VaR is completely inadequate
- ☐ Student-t + asymmetric model: **essential** for both markets



Key Formulas

Volatility Models

- **ARCH(q):** $\sigma_t^2 = \omega + \sum_{i=1}^q \alpha_i \varepsilon_{t-i}^2$
- **GARCH(1,1):** $\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \beta \sigma_{t-1}^2$
- **EGARCH:** $\ln(\sigma_t^2) = \omega + \alpha(|z_{t-1}| - \mathbb{E}[|z|]) + \gamma z_{t-1} + \beta \ln(\sigma_{t-1}^2)$
- **GJR-GARCH:** $\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \gamma \varepsilon_{t-1}^2 I_{t-1} + \beta \sigma_{t-1}^2$

Properties and Measures

- **Unconditional variance:** $\bar{\sigma}^2 = \frac{\omega}{1-\alpha-\beta}$ **Half-life:** $HL = \frac{\ln(0.5)}{\ln(\alpha+\beta)}$
- **VaR:** $\text{VaR}_\alpha = z_\alpha \cdot \sigma_{T+1}$ **Stationarity:** $\alpha + \beta < 1$
- **ARCH-LM:** $LM = T \cdot R^2 \sim \chi^2(q)$



Summary — Chapter 5: Volatility Models

Key Concepts

- **ARCH(q)**: conditional variance depends on past squared errors
- **GARCH(p,q)**: adds variance lags for persistence
- **EGARCH/GJR-GARCH**: capture leverage effect (asymmetric response)

Applications

Risk measurement (VaR, ES), derivative pricing, portfolio management

Practical Tip

Start with GARCH(1,1), check for leverage, choose distribution minimizing AIC/BIC!



References

-  Engle, R.F. (1982). *Autoregressive Conditional Heteroscedasticity with Estimates of the Variance of United Kingdom Inflation*. *Econometrica*, 50(4), 987-1007.
-  Bollerslev, T. (1986). *Generalized Autoregressive Conditional Heteroskedasticity*. *Journal of Econometrics*, 31(3), 307-327.
-  Nelson, D.B. (1991). *Conditional Heteroskedasticity in Asset Returns: A New Approach*. *Econometrica*, 59(2), 347-370.
-  Glosten, L.R., Jagannathan, R., & Runkle, D.E. (1993). *On the Relation between the Expected Value and the Volatility of the Nominal Excess Return on Stocks*. *The Journal of Finance*, 48(5), 1779-1801.
-  Tsay, R.S. (2010). *Analysis of Financial Time Series*. 3rd Edition, Wiley.

Online Resources and Code

- **Quantlet:** <https://quantlet.com> → Code repository for statistics
- **Quantinar:** <https://quantinar.com> → Learning platform for quantitative methods
- **GitHub TSA_ch5:** https://github.com/QuantLet/TSA/tree/main/TSA_ch5



Thank You!

Questions?

Course materials available at: <https://danpele.github.io/Time-Series-Analysis/>

