



Time Series Analysis and Forecasting

Chapter 4: SARIMA Models



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Learning Objectives

By the end of this chapter, you will be able to:

1. Identify seasonal patterns in time series data
2. Apply seasonal differencing to remove seasonal unit roots
3. Build and estimate SARIMA models with seasonal components
4. Interpret seasonal ACF/PACF patterns for model identification
5. Evaluate forecasts using rolling window methods for seasonal data
6. Apply the complete methodology on real data (airline passengers)



Outline

- Motivation
- Seasonality in Time Series
- Seasonal Differencing
- The SARIMA Model
- Seasonal ACF and PACF Patterns
- Estimation and Diagnostics
- Forecasting with SARIMA
- Case Study: Airline Passengers
- Practical Aspects
- AI Use Case
- Summary
- Quiz



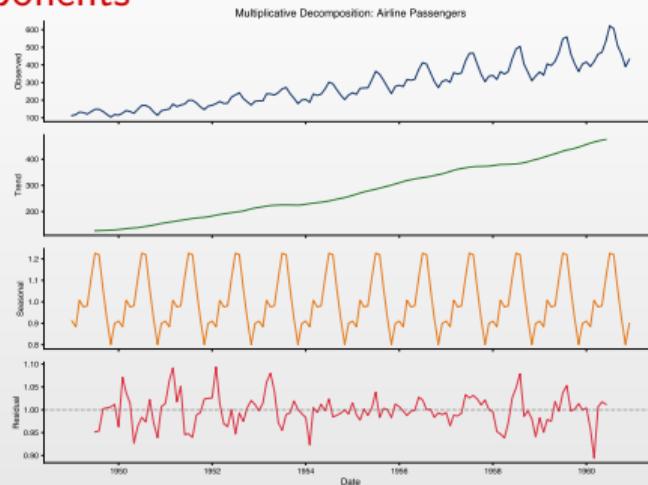
Why SARIMA? Seasonality Is Everywhere



- Retail sales exhibit clear **annual patterns**: December peaks, January troughs
- Standard ARIMA models cannot capture these **repeating seasonal cycles**
- Ignoring seasonality leads to systematic forecast errors



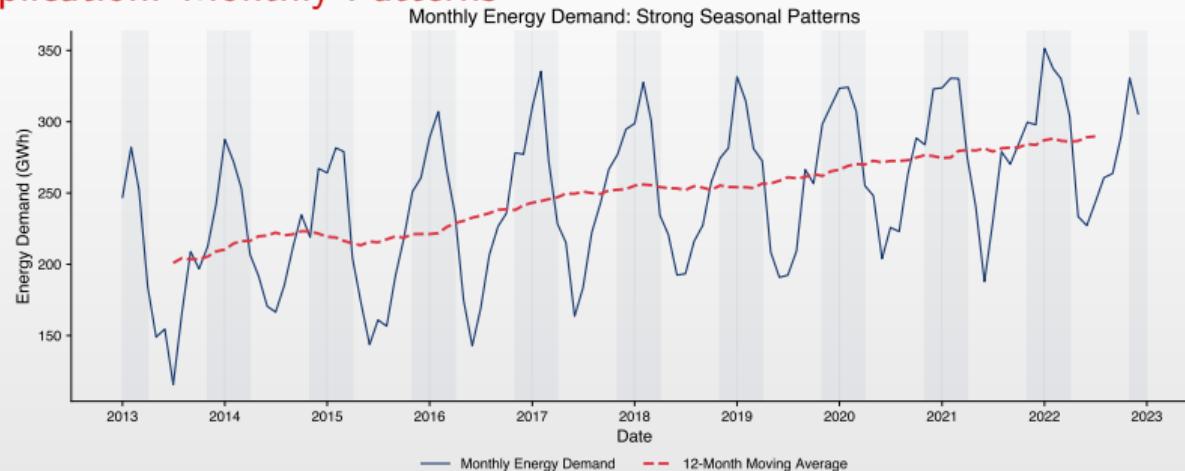
Understanding Seasonal Components



- ☐ Seasonal time series = **Trend + Seasonal Pattern + Residuals**
- ☐ Decomposition helps visualize each component separately
- ☐ SARIMA models capture both trend dynamics and seasonal behavior



Real-World Application: Monthly Patterns

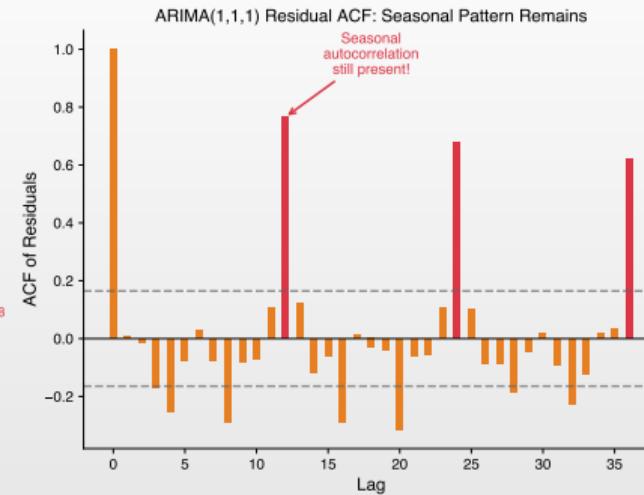
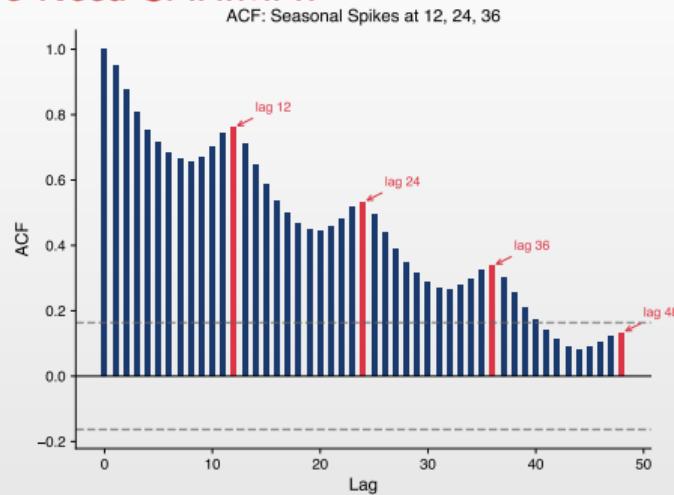


- Energy demand shows strong **monthly seasonality** (heating/cooling cycles)
- Pattern repeats predictably each year with slight variations
- Utility companies use SARIMA forecasts for capacity planning

TSA_ch4_motivation_monthly



Why Do We Need SARIMA?



- **Left:** Seasonal ACF shows spikes at lags 12, 24, 36... (annual pattern)
- **Right:** ARIMA residuals still show seasonal autocorrelation \Rightarrow model is incomplete
- **SARIMA adds seasonal AR and MA terms** to capture these patterns



What is Seasonality?

Definition 1 (Seasonality)

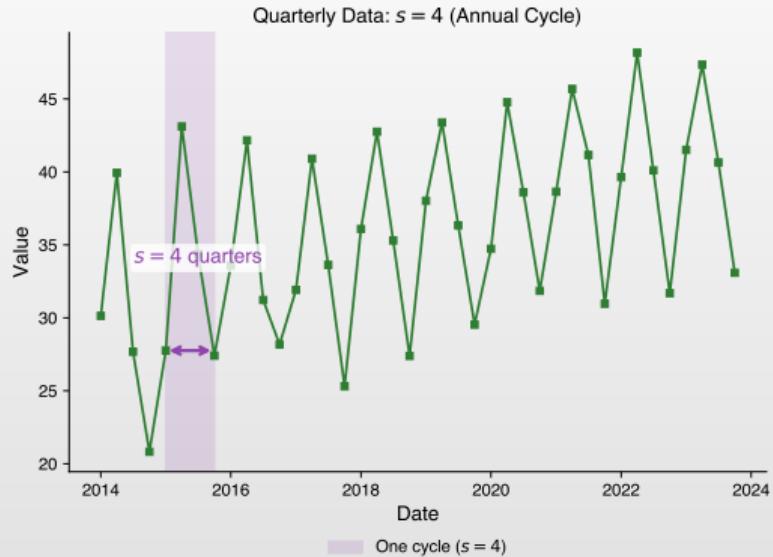
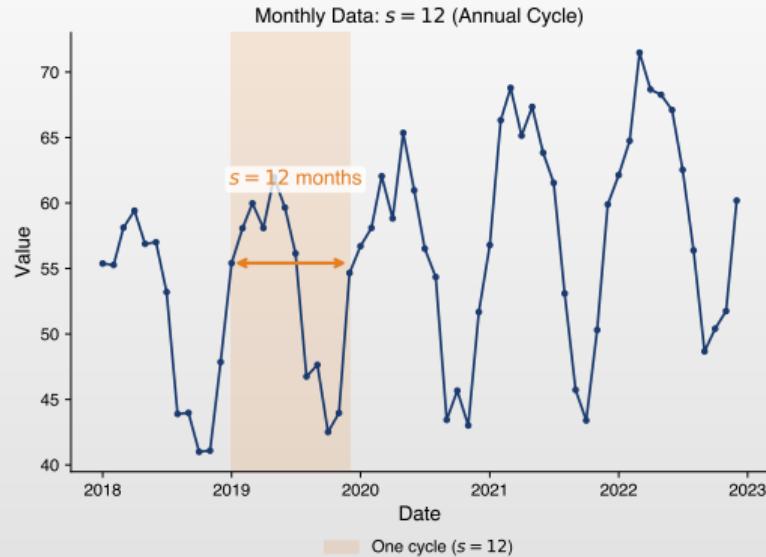
A time series exhibits seasonality when it shows regular, periodic fluctuations that repeat over a fixed period s (the seasonal period).

Common Seasonal Periods

- Monthly data: $s = 12$ (annual cycle)
- Quarterly data: $s = 4$ (annual cycle)
- Weekly data: $s = 52$ (annual) or $s = 7$ (weekly pattern)
- Daily data: $s = 7$ (weekly pattern)



What is Seasonality?



Example: Airline Passengers Data



- Monthly international airline passengers (1949–1960)
- Clear **upward trend and growing seasonal amplitude**
- Summer peaks reflect vacation travel patterns

 TSA_ch4_airline_data



Examples of Seasonal Data

Economic Series

- Retail sales (holiday peaks)
- Tourism (summer/winter)
- Agricultural production
- Energy consumption
- Employment (hiring cycles)

Other Domains

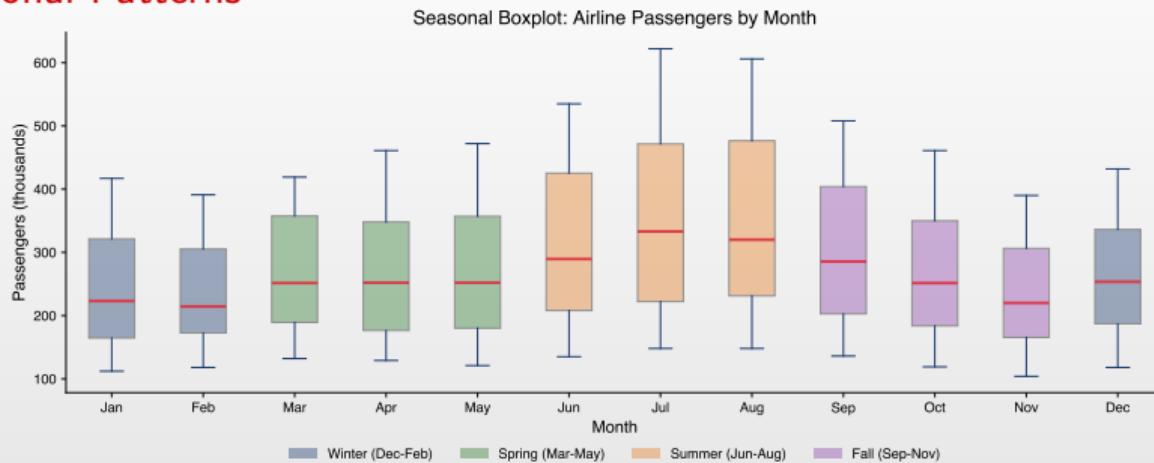
- Weather/temperature
- Website traffic
- Hospital admissions
- Transportation usage
- Electricity demand

Why It Matters

Ignoring seasonality leads to biased forecasts and invalid inference!



Visualizing Seasonal Patterns



- Box plot reveals consistent seasonal pattern across years
- July–August: highest passenger counts (summer travel)
- November–February: lowest counts (winter months)

 TSA_ch4_seasonal_boxplot



Deterministic vs Stochastic Seasonality

Deterministic Seasonality

- Fixed pattern:** $Y_t = \sum_{j=1}^s \gamma_j D_{jt} + \varepsilon_t$
 - D_{jt} are seasonal dummies
- Pattern constant over time
- Same amplitude every year
- Removed by regression on dummies
- ACF: sharp cutoff at seasonal lags
- Example:** University enrollment peaks every September by the same amount

Stochastic Seasonality

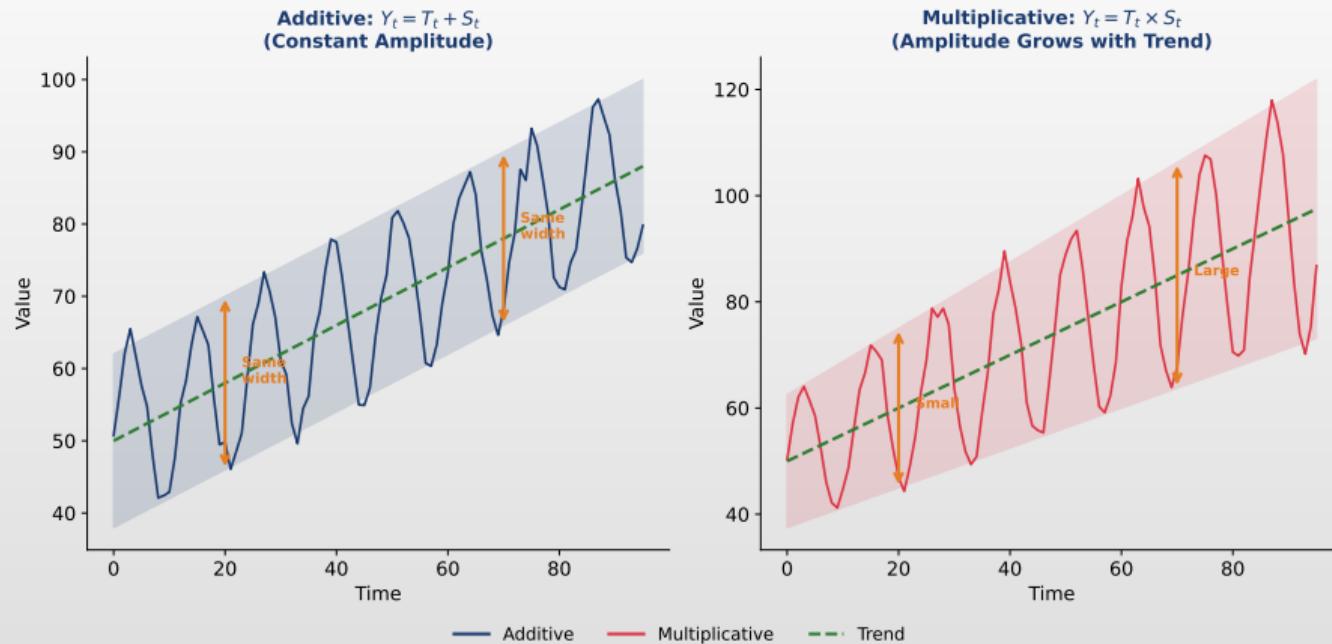
- Evolving pattern:** $\Delta_s Y_t = Y_t - Y_{t-s}$
 - Exhibits dependence structure
- Pattern evolves over time
- Amplitude may grow or shrink
- Requires seasonal differencing
- ACF: slow decay at seasonal lags
- Example:** Retail sales peaks grow larger each December

How to decide?

- Slow ACF decay at lags $s, 2s, 3s, \dots \Rightarrow$ stochastic (use Δ_s)
- Sharp cutoff \Rightarrow deterministic (use dummies)
- Use HEGY or Canova-Hansen tests to confirm



Additive vs Multiplicative Seasonality



Additive vs Multiplicative Seasonality

$$\text{Additive: } Y_t = T_t + S_t + \varepsilon_t$$

- Seasonal amplitude **constant**
- No transformation needed
- Ex: temperatures, university enrollment

$$\text{Multiplicative: } Y_t = T_t \cdot S_t \cdot \varepsilon_t$$

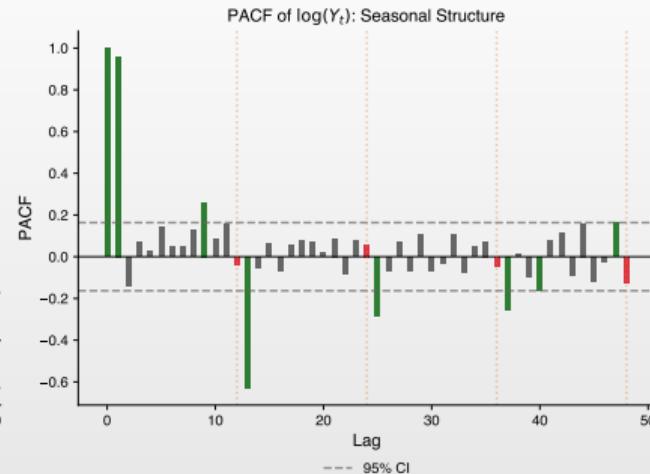
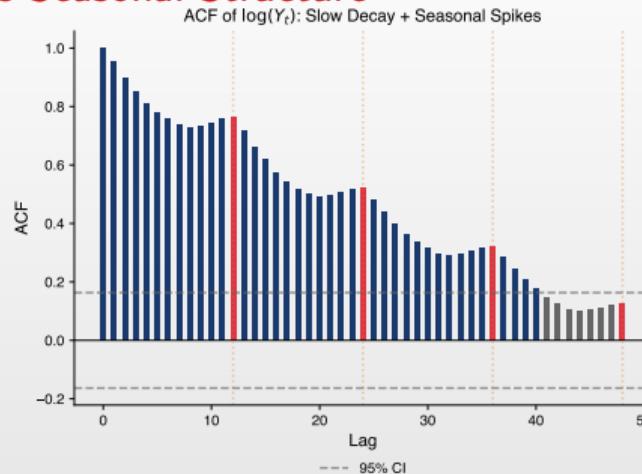
- Amplitude **grows with level**
- Requires log transform (Box-Cox)
- Ex: Airline, retail sales, GDP

First practical decision

- Amplitude grows with the trend? \Rightarrow multiplicative \Rightarrow apply log/Box-Cox *before* differencing



ACF Reveals Seasonal Structure



- ☐ Slow decay at all lags indicates non-stationarity (trend)
- ☐ Spikes at lags 12, 24, 36 confirm seasonal pattern ($s = 12$)
- ☐ ACF at seasonal lags: slow decay \Rightarrow needs seasonal differencing



Detecting Seasonality

Visual Methods

- Time series plot – look for repeating patterns
- Seasonal subseries plot – compare same seasons across years
- ACF plot – spikes at seasonal lags ($s, 2s, 3s, \dots$)

Statistical Tests

- Seasonal unit root tests (HEGY, Canova-Hansen, OCSB^a)
- F-test for seasonal dummy variables
- Kruskal-Wallis test (non-parametric)

ACF Signature

- Strong seasonality: ACF shows significant spikes at lags $s, 2s, 3s, \dots$

^aOsborn-Chui-Smith-Birchenhall — the default test in `auto_arima`



F-Test for Seasonal Dummy Variables: Intuition

What does this test do?

- Goal:** test whether mean values differ significantly across seasons
- Logic:** if the mean in January \neq February $\neq \dots \neq$ December \Rightarrow seasonality
- Method:** compare a model WITH seasonal dummy variables vs. a model WITHOUT

Models compared

- Restricted:** $Y_t = \alpha + \varepsilon_t$ **Unrestricted:** $Y_t = \alpha + \sum_{j=1}^{s-1} \gamma_j D_{jt} + \varepsilon_t$
- where $D_{jt} = 1$ if observation t is in season j , 0 otherwise

Key idea

- If adding seasonal dummy variables **significantly reduces** prediction errors, then seasonality is present



F-Test for Seasonal Dummy Variables: Formula and Example

F-statistic formula

- **Formula:** $F = \frac{(SSR_R - SSR_U)/(s-1)}{SSR_U/(n-s)} \sim F_{s-1, n-s}$
 - ▶ SSR_R : sum of squared residuals from the restricted model (no dummies)
 - ▶ SSR_U : sum of squared residuals from the unrestricted model (with dummies)
 - ▶ $s - 1$: number of restrictions (monthly: 11, quarterly: 3)

Numerical example (Monthly data, $n=120$)

- $SSR_R = 15000, SSR_U = 8500, s = 12$
- $F = \frac{(15000 - 8500)/11}{8500/108} = \frac{590.9}{78.7} = 7.51$
- Critical value $F_{0.05, 11, 108} \approx 1.87$. Since $7.51 > 1.87$: **Reject H_0** \Rightarrow Seasonality present!



Kruskal-Wallis Test: Intuition

What does this test do?

- ☐ **Non-parametric test:** checks whether observations from different seasons come from the same distribution (no distributional assumptions needed)
- ☐ **Mechanism:** ranks all observations 1 to N , then checks whether ranks are uniformly distributed across seasons
- ☐ **Conclusion:** if one season consistently has higher/lower ranks \Rightarrow seasonality (e.g., summer always has high ranks \Rightarrow seasonal energy consumption)

Why use it instead of the F-test?

- ☐ **No normality assumption** – works with any distribution; ideal for heavy-tailed or skewed data
- ☐ **Robust to outliers** – ranks bound the influence of any single observation

Limitation

- ☐ Less powerful than the F-test when data ARE normally distributed



Kruskal-Wallis Test: Formula and Example

Test statistic

- $\square H = \frac{12}{N(N+1)} \sum_{j=1}^s \frac{R_j^2}{n_j} - 3(N + 1)$ where N = total obs., n_j = obs. in season j , R_j = rank sum

Example: Quarterly sales ($n=20$, $s=4$)

- \square Data ranked 1–20. Rank sums: Q1: $R_1 = 15$, Q2: $R_2 = 35$, Q3: $R_3 = 70$, Q4: $R_4 = 90$
- $\square H = \frac{12}{20 \times 21} \left(\frac{15^2}{5} + \frac{35^2}{5} + \frac{70^2}{5} + \frac{90^2}{5} \right) - 3(21) = 19.6$
- \square Critical value $\chi^2_{0.05,3} = 7.81$. Since $19.6 > 7.81$: **Reject $H_0 \Rightarrow$ Seasonality!**

In Python

- \square Implementation: `scipy.stats.kruskal(q1, q2, q3, q4)`



HEGY Test: What Problem Does It Solve?

Key question

- **Problem:** given a seasonal series, we need to determine the type of differencing (wrong choice \Rightarrow over- or under-differencing)
- **Regular differencing** ($1 - L$)? \Rightarrow set $d = 1$; **Seasonal differencing** ($1 - L^s$)? \Rightarrow set $D = 1$
- **HEGY** (Hylleberg, Engle, Granger, Yoo, 1990): tests for both types of unit roots simultaneously!

Why not just use ADF?

- **ADF:** tests only for a regular unit root at frequency zero — completely ignores seasonal components
- **Limitation:** seasonal data may have unit roots at seasonal frequencies that ADF misses!

HEGY tests multiple frequencies

- **Quarterly:** tests at $0, \pi, \pm\pi/2$ (3 freq. \Rightarrow one regular + two seasonal roots)
- **Monthly:** tests at $0, \pi, \pm\pi/6, \pm\pi/3, \pm\pi/2, \pm2\pi/3, \pm5\pi/6$ (6 freq. \Rightarrow one regular + five seasonal roots)



HEGY Test: Auxiliary Regression (Quarterly)

HEGY auxiliary regression

- **Quarterly data ($s = 4$):** $\Delta_4 y_t = \pi_1 z_{1,t-1} + \pi_2 z_{2,t-1} + \pi_3 z_{3,t-2} + \pi_4 z_{4,t-2} + \sum_{j=1}^k \phi_j \Delta_4 y_{t-j} + \varepsilon_t$
 - Lags k chosen by AIC/BIC to eliminate residual autocorrelation

Transformed variables

- z_{1t} : $(1 + L + L^2 + L^3)y_t$ — captures the zero-frequency component (trend)
- z_{2t} : $-(1 - L + L^2 - L^3)y_t$ — captures the frequency- π component (semi-annual)
- z_{3t}, z_{4t} : capture the frequency $\pm\pi/2$ components (annual cycle)

Hypotheses

- $H_0 : \pi_1 = 0$: unit root at frequency 0 \Rightarrow requires $d = 1$
- $H_0 : \pi_2 = 0$: unit root at frequency π \Rightarrow requires $D = 1$
- $H_0 : \pi_3 = \pi_4 = 0$: unit root at frequency $\pm\pi/2$ \Rightarrow requires $D = 1$



HEGY Test: Decision Rules with Examples

HEGY critical values (5%, n=100, with constant)

Test	Statistic	Critical value	If NOT rejected...
$t_1 (\pi_1 = 0)$	t-stat	-2.88	Requires $d = 1$
$t_2 (\pi_2 = 0)$	t-stat	-2.88	Requires $D = 1$
$F_{34} (\pi_3 = \pi_4 = 0)$	F-stat	6.57	Requires $D = 1$

Example: Quarterly GDP

- **HEGY results:** $t_1 = -1.52$, $t_2 = -4.21$, $F_{34} = 2.15$
- $t_1 = -1.52 > -2.88$: Cannot reject \Rightarrow requires $d = 1$
- $t_2 = -4.21 < -2.88$: Reject \Rightarrow no unit root at π
- $F_{34} = 2.15 < 6.57$: Cannot reject \Rightarrow requires $D = 1$
- **Conclusion:** Use SARIMA with $d = 1, D = 1$



Canova-Hansen Test: The Opposite of HEGY

HEGY vs Canova-Hansen: Different null hypotheses!

	HEGY	Canova-Hansen
H_0	Seasonal unit root	No seasonal unit root
H_1	No seasonal unit root	Seasonal unit root
Reject H_0	Use seasonal dummies	Use differencing ($1 - L^s$)
Do not reject	Use differencing ($1 - L^s$)	Use seasonal dummies

Why does it matter?

- HEGY: "Prove there is NO unit root" (conservative towards differencing)
- CH: "Prove there IS a unit root" (conservative towards dummies)
- Use **both** tests for robust conclusions!



Canova-Hansen Test: Formula

Testing procedure

- **Step 1:** Regress y_t on seasonal dummies: $y_t = \sum_{j=1}^s \gamma_j D_{jt} + u_t$
- **Step 2:** Compute partial sums at seasonal frequency λ_i :
 - $S_{it}^{(c)} = \sum_{j=1}^t \hat{u}_j \cos(\lambda_i j)$, $S_{it}^{(s)} = \sum_{j=1}^t \hat{u}_j \sin(\lambda_i j)$

LM test statistic

- $LM_i = \frac{1}{T^2 \hat{\omega}_i} \left[\sum_{t=1}^T (S_{it}^{(c)})^2 + \sum_{t=1}^T (S_{it}^{(s)})^2 \right]$
- where $\hat{\omega}_i$ = consistent estimator of the spectral density at frequency λ_i

Decision

- **Rule:** reject H_0 (stationarity) if $LM >$ critical value \Rightarrow seasonal differencing is needed



Summary: Choosing the Right Seasonality Test

Test	H_0	If rejected	Best for
F-test	No seasonality	Seasonality exists	Normal data
Kruskal-Wallis	No difference across seasons	Seasonality exists	Non-normal, outliers
HEGY	Unit root exists	Use dummies	Determining d, D
Canova-Hansen	No unit root	Use $(1 - L^s)$	Confirming stability

Key idea

- F-test / Kruskal-Wallis: “Does seasonality exist?”
- HEGY / Canova-Hansen: “What type?” (deterministic vs stochastic)



Box-Cox Transformation: Variance Stabilization

Box-Cox Family of Transformations

- ☐ **Formula:** $Y_t^{(\lambda)} = (Y_t^\lambda - 1)/\lambda$ if $\lambda \neq 0$; $\ln(Y_t)$ if $\lambda = 0$
- ☐ **Special cases:** $\lambda=1$ (none), $\lambda=0$ (log), $\lambda=0.5$ ($\sqrt{\cdot}$). $\lambda < 1$ compresses large values; $\lambda > 1$ amplifies

Automatic Selection of λ

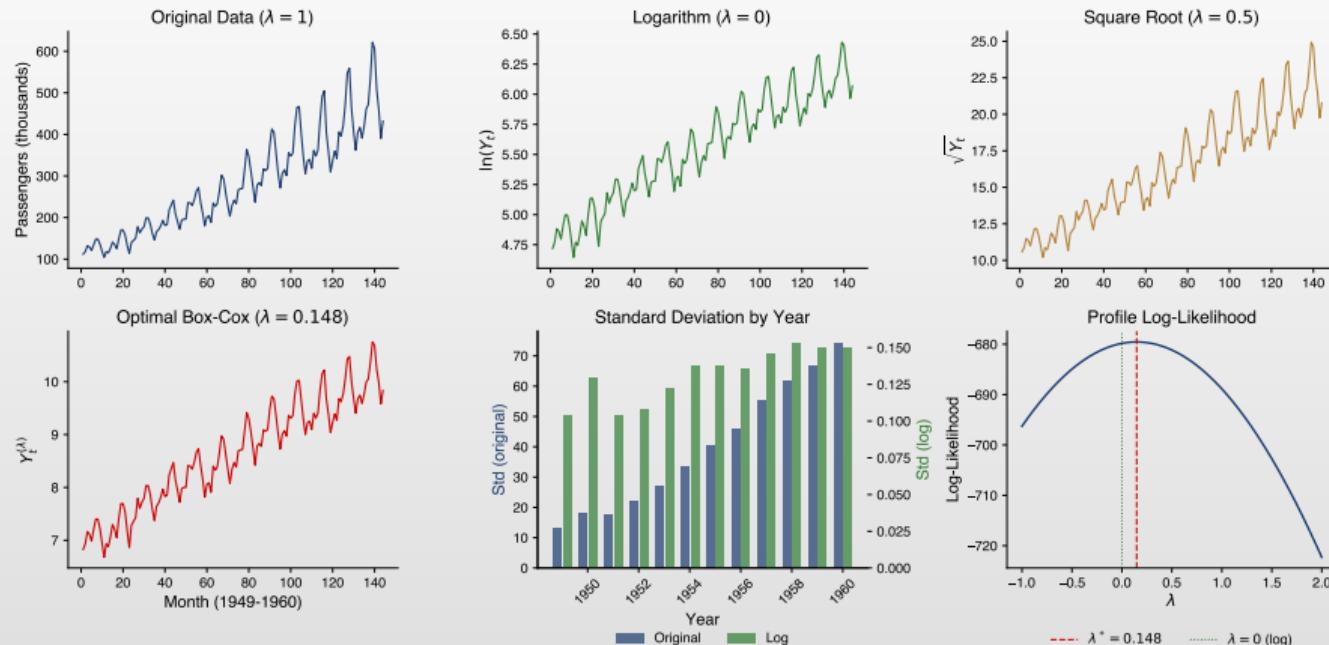
- ☐ **Profile likelihood:** maximizes log-likelihood as a function of λ (most common; provides CI)
- ☐ **Guerrero (1993):** minimizes CV of seasonal sub-series (preferred for varying amplitude)
- ☐ **Python:** `boxcox(y)` from `scipy.stats` or `BoxCox.lambda_(y)` in R

Why not just logarithm?

- ☐ Log ($\lambda=0$) assumes variance \propto level — not always the case (e.g., may need $\lambda = 0.3$)
- ☐ Box-Cox chooses the optimal transformation based on data, not assumptions



Box-Cox on the Airline Data: Complete Example



Box-Cox on the Airline Data: Complete Example

Result for Airline Passengers

- $\hat{\lambda} = 0.148 \approx 0 \Rightarrow \log$ is nearly optimal
- Standard deviation per year: from increasing (original) to stable (log)

Bias Correction in Back-Transformation

- On log scale: \hat{y}_{T+h} is the **median**, not the mean
- Correction: $\hat{Y}_{T+h} = \exp\left(\hat{y}_{T+h} + \frac{\sigma_h^2}{2}\right)$
- Without correction: systematically under-estimated forecasts!



STL Decomposition: Modern Alternatives

STL: Seasonal-Trend Decomposition using Loess (Cleveland et al., 1990)

- Advantages:** time-varying seasonality, robust to outliers, any period s
- Algorithm:** iterative locally weighted regression (loess)

Key Parameters

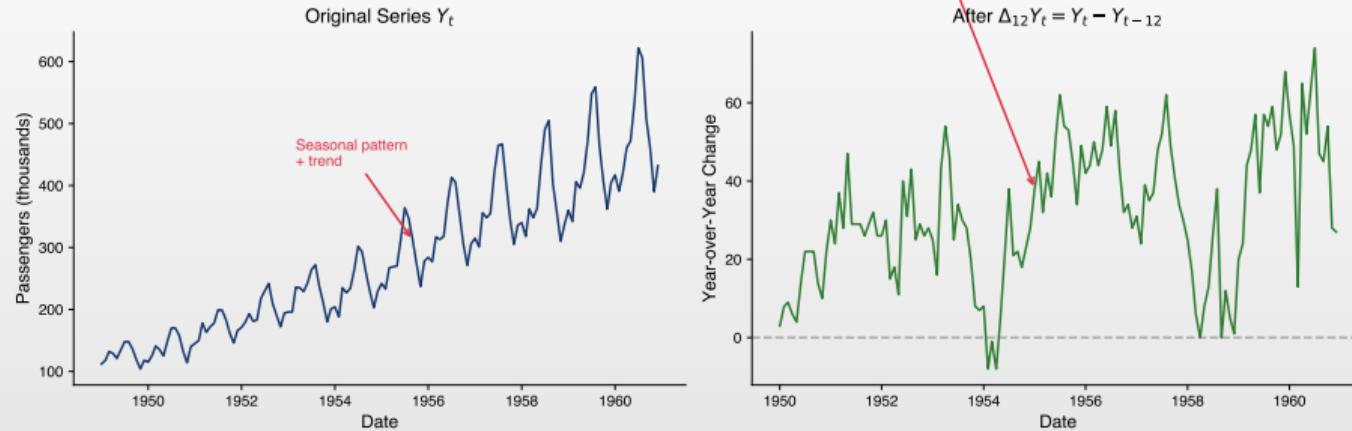
- Seasonal window** (`seasonal`): controls how quickly seasonality changes
- Trend window** (`trend`): smoothing of the trend component
- Robustness** (`robust=True`): reduces influence of outliers

Practical Usage

- STL for exploration and preprocessing; SARIMA for modeling and forecasting
- Python: `STL(y, period=12).fit()` from `statsmodels`



Seasonal Difference: Visual Illustration



- Left:** Original series with clear seasonal pattern
- Right:** After $\Delta_{12} = (1 - L^{12})$, the seasonal pattern is removed
 - ▶ Year-over-year comparison eliminates seasonal effects



The Seasonal Difference Operator

Definition 2 (Seasonal Difference)

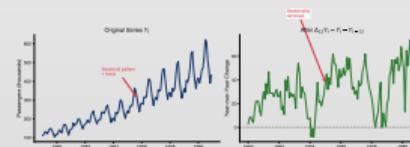
The seasonal difference operator Δ_s is defined as:

$$\Delta_s Y_t = (1 - L^s)Y_t = Y_t - Y_{t-s}$$

where $L^s Y_t = Y_{t-s}$ is the seasonal lag operator.

Examples

- Monthly data ($s = 12$): $\Delta_{12} Y_t = Y_t - Y_{t-12}$
Compares each month to the same month last year
- Quarterly data ($s = 4$): $\Delta_4 Y_t = Y_t - Y_{t-4}$
Compares each quarter to the same quarter last year



Proof: Seasonal Differencing Removes Deterministic Seasonality

Claim: If $Y_t = \mu_t + \varepsilon_t$ where $\mu_t = \mu_{t-s}$ (periodic mean), then $\Delta_s Y_t$ removes the seasonal mean.

- The condition $\mu_t = \mu_{t-s}$ means the mean repeats exactly every s periods

Proof: Let $Y_t = \mu_t + \varepsilon_t$ where μ_t has period s . Apply seasonal difference:

$$\begin{aligned}\Delta_s Y_t &= Y_t - Y_{t-s} = (\mu_t + \varepsilon_t) - (\mu_{t-s} + \varepsilon_{t-s}) \\ &= \mu_t - \mu_{t-s} + \varepsilon_t - \varepsilon_{t-s} \\ &= 0 + \varepsilon_t - \varepsilon_{t-s} \quad (\text{since } \mu_t = \mu_{t-s})\end{aligned}$$

Properties of $\Delta_s Y_t = \varepsilon_t - \varepsilon_{t-s}$:

- $\mathbb{E}[\Delta_s Y_t] = 0$ (constant mean); $\text{Var}(\Delta_s Y_t) = 2\sigma^2$ (constant variance)
 - ▶ Both properties are required for second-order stationarity
- Autocovariance: $\gamma(s) = -\sigma^2$, $\gamma(k) = 0$ for $k \neq 0, s$
 - ▶ ACF shows a single spike at lag s , then zero

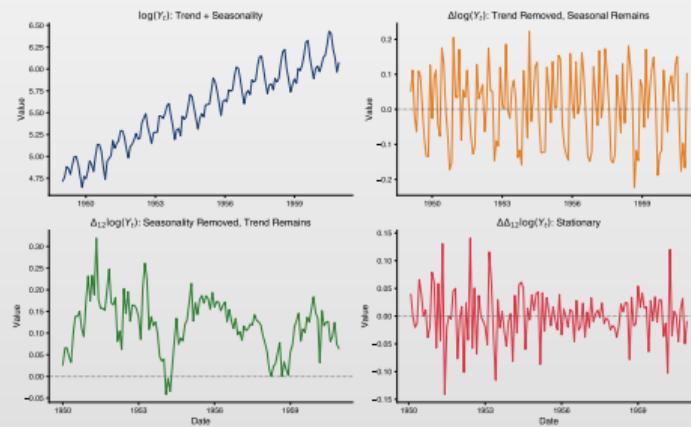
Result

Seasonal differencing transforms periodic seasonal pattern into MA(1) at seasonal lag.



Effect of Differencing Operations

- Regular differencing removes trend but seasonal pattern remains
- Seasonal differencing removes seasonality but trend pattern remains
- Both differences needed to achieve stationarity**



Q TSA_ch4_differencing_effect



Combining Regular and Seasonal Differencing

Full Differencing

For series with both trend and seasonality:

$$\Delta\Delta_s Y_t = (1 - L)(1 - L^s)Y_t$$

Expansion

$$(1 - L)(1 - L^s)Y_t = Y_t - Y_{t-1} - Y_{t-s} + Y_{t-s-1}. \text{ For monthly: } \Delta\Delta_{12} Y_t = Y_t - Y_{t-1} - Y_{t-12} + Y_{t-13}$$

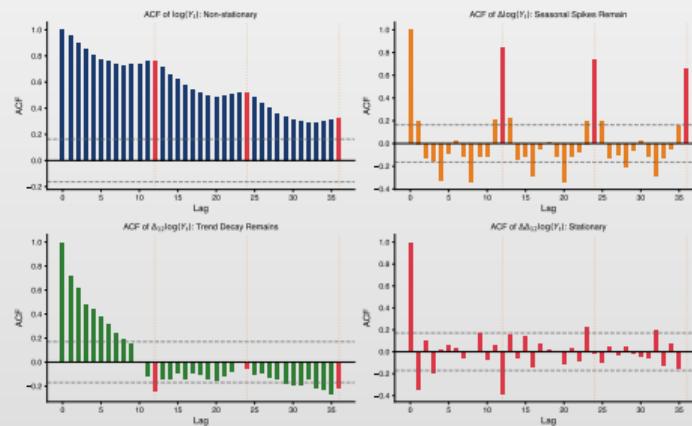
Order of Differencing

d : regular differences (trend removal); D : seasonal differences (seasonal trend removal)



ACF Before and After Differencing

- ◻ Original ACF: slow decay indicates non-stationarity
- ◻ After Δ : seasonal spikes remain at lags 12, 24, 36
- ◻ After Δ_{12} : trend decay remains at early lags
- ◻ After $\Delta\Delta_{12}$: ACF cuts off \Rightarrow stationary



Seasonal Integration

Definition 3 (Seasonally Integrated Process)

A series Y_t is **seasonally integrated** of order $(d, D)_s$, written $Y_t \sim I(d, D)_s$, if:

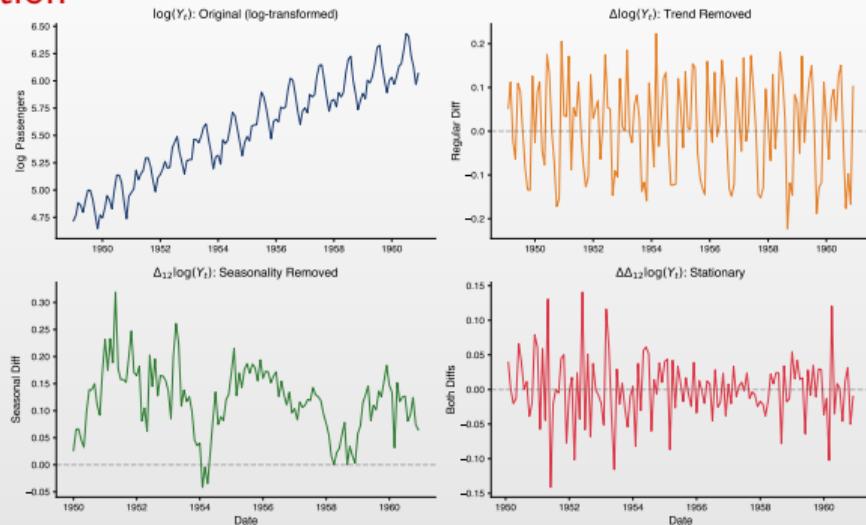
$$(1 - L)^d(1 - L^s)^D Y_t$$

is stationary.

Common Cases

- $I(1, 0)_{12}$: Regular unit root only (monthly)
- $I(0, 1)_{12}$: Seasonal unit root only
- $I(1, 1)_{12}$:
 - ▶ Both regular and seasonal unit roots

SARIMA: Visual Illustration



- Original \Rightarrow regular difference (removes trend) \Rightarrow seasonal difference (removes seasonality)
- Apply minimum differencing needed to achieve stationarity



SARIMA Model Definition

Definition 4 (SARIMA(p, d, q) \times (P, D, Q) $_s$)

The **Seasonal ARIMA** model is:

$$\phi(L)\Phi(L^s)(1 - L)^d(1 - L^s)^D Y_t = c + \theta(L)\Theta(L^s)\varepsilon_t$$

Components

- $\phi(L) = 1 - \phi_1L - \dots - \phi_pL^p$: Non-seasonal AR
- $\Phi(L^s) = 1 - \Phi_1L^s - \dots - \Phi_PL^{Ps}$: Seasonal AR
- $\theta(L) = 1 + \theta_1L + \dots + \theta_qL^q$: Non-seasonal MA
- $\Theta(L^s) = 1 + \Theta_1L^s + \dots + \Theta_QL^{Qs}$: Seasonal MA
- $(1 - L)^d$:
 - ▶ Regular differencing; $(1 - L^s)^D$: Seasonal differencing



Proof: Multiplicative Seasonal Structure

Why multiplicative? Consider SARIMA(1, 0, 0) \times (1, 0, 0)_s:

$$(1 - \phi L)(1 - \Phi L^s) Y_t = \varepsilon_t$$

Expand: $(1 - \phi L)(1 - \Phi L^s) Y_t = Y_t - \phi Y_{t-1} - \Phi Y_{t-s} + \phi \Phi Y_{t-s-1}$

Interpretation (Monthly, $s = 12$)

Y_t depends on: Y_{t-1} (last month), Y_{t-12} (same month last year), Y_{t-13} (interaction).

Parsimony: Multiplicative form uses 2 parameters (ϕ, Φ); additive would need 3+.



SARIMA Notation

Full Specification

SARIMA(p, d, q) \times (P, D, Q) $_s$ has 7 parameters to specify:

Parameter	Meaning
p	Non-seasonal AR order
d	Non-seasonal differencing order
q	Non-seasonal MA order
P	Seasonal AR order
D	Seasonal differencing order
Q	Seasonal MA order
s	Seasonal period

Example

SARIMA(1, 1, 1) \times (1, 1, 1) $_{12}$: Monthly data with AR(1), MA(1), seasonal AR(1), seasonal MA(1), and both regular and seasonal differencing.



Common SARIMA Models

Airline Model: $\text{SARIMA}(0, 1, 1) \times (0, 1, 1)_s$

$$(1 - L)(1 - L^s)Y_t = (1 + \theta L)(1 + \Theta L^s)\varepsilon_t \text{ -- Classic model (Box & Jenkins, 1970)}$$

$\text{SARIMA}(1, 0, 0) \times (1, 0, 0)_s$

$$(1 - \phi L)(1 - \Phi L^s)Y_t = \varepsilon_t \text{ -- Pure seasonal and non-seasonal AR}$$

$\text{SARIMA}(0, 1, 1) \times (0, 1, 0)_s$

$$(1 - L)(1 - L^s)Y_t = (1 + \theta L)\varepsilon_t \text{ -- Random walk + seasonal diff + MA(1)}$$



ACF/PACF for Seasonal Models

Key Insight

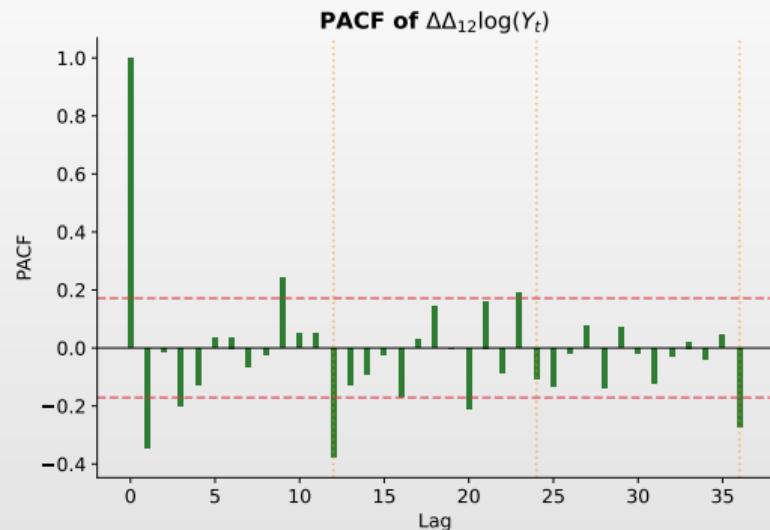
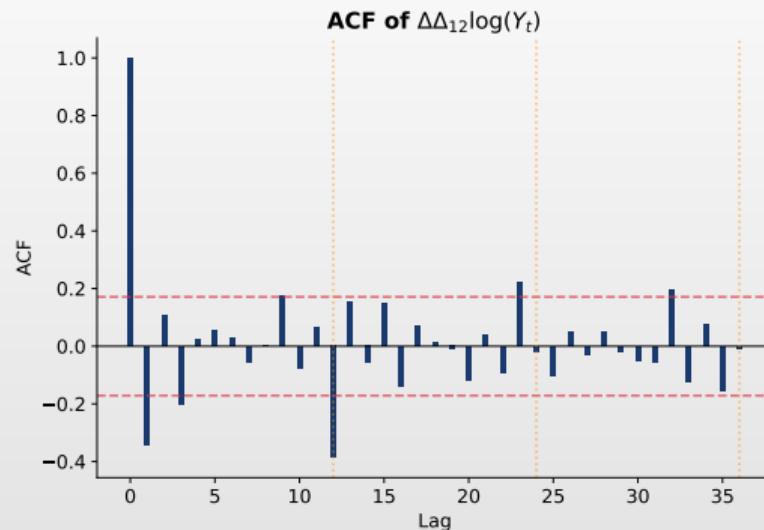
Seasonal models show patterns at both:

- Non-seasonal lags: $1, 2, 3, \dots$
- Seasonal lags: $s, 2s, 3s, \dots$

Model	ACF	PACF
SAR(P)	Decays at $s, 2s, \dots$	Cuts off after Ps
SMA(Q)	Cuts off after Qs	Decays at $s, 2s, \dots$
SARMA	Decays at seasonal lags	Decays at seasonal lags



Example: Airline Model ACF/PACF



Example: Airline Model ACF/PACF

ACF: $\Delta\Delta_{12} \log(Y_t)$

- Spike at lag 1 \leftarrow MA(1), θ
- Spike at lag 12 \leftarrow SMA(1), Θ
- Rest \approx zero

PACF: exponential decay

- Decays at lags 1, 2, 3, ...
- Decays at lags 12, 24, 36
- \Rightarrow MA, not AR

-
- Conclusion: ACF cuts off \Rightarrow MA; PACF decays \Rightarrow not AR. Model: $(0, 1, 1) \times (0, 1, 1)_{12}$



Model Identification Guidelines

Step-by-Step Process

1. Examine ACF for slow decay at seasonal lags \Rightarrow seasonal differencing
2. After differencing, check ACF/PACF patterns
3. Non-seasonal behavior at lags $1, 2, \dots, s - 1$
4. Seasonal behavior at lags $s, 2s, 3s, \dots$

Practical Tips

- Start with $d \leq 1$ and $D \leq 1$
- Usually $P, Q \leq 2$ is sufficient
- Use information criteria (AIC, BIC) for final selection
- Auto-SARIMA algorithms can help



Estimation Methods

Maximum Likelihood Estimation

Standard approach for SARIMA:

- Conditional MLE (conditional on initial values)
- Exact MLE (via Kalman filter)

Computational Considerations

- More parameters than ARIMA \Rightarrow more data needed
- Seasonal parameters estimated from lags $s, 2s, \dots$
- Need sufficient seasonal cycles (at least 3-4 years of monthly data)



Exact Likelihood: Prediction Error Decomposition

Why the Kalman Filter?

- **SARIMA:** has the structure of a state-space model
- **Kalman filter:** recursively computes prediction errors v_t and their variances f_t , without conditioning on initial values

Exact Log-Likelihood (Prediction Error Decomposition)

- **Formula:** $\ell(\theta) = -\frac{T}{2} \ln(2\pi) - \frac{1}{2} \sum_{t=1}^T \left[\ln(f_t) + \frac{v_t^2}{f_t} \right]$
- $v_t: Y_t - \hat{Y}_{t|t-1}$ (innovation); $f_t: \text{Var}(v_t)$ (innovation variance)

Advantages over Conditional MLE

- Does not require choosing initial values
- Each term $\ln(f_t)$ weights observations differently (variable variance at start)
- Essential for short series where initial values matter
- Implemented by default in `statsmodels.tsa.SARIMAX()` with `method='mle'`



Stationarity and Invertibility

Stationarity Conditions

Both non-seasonal and seasonal AR polynomials must have roots outside the unit circle:

- $\phi(z) = 0 \Rightarrow |z| > 1$
- $\Phi(z^s) = 0 \Rightarrow |z| > 1$

Invertibility Conditions

Both non-seasonal and seasonal MA polynomials must have roots outside the unit circle:

- $\theta(z) = 0 \Rightarrow |z| > 1$
- $\Theta(z^s) = 0 \Rightarrow |z| > 1$



Diagnostic Checking

Residual Analysis

After fitting SARIMA, check that residuals are white noise:

1. Plot residuals over time (no patterns)
2. ACF of residuals (no significant spikes)
3. Ljung-Box test at multiple lags including seasonal
4. Normality tests (Q-Q plot, Jarque-Bera)

Important

Check ACF at **both** non-seasonal and seasonal lags!

Significant ACF at lag 12 suggests inadequate seasonal modeling.



Model Selection Criteria

Information Criteria

Compare competing SARIMA models using:

- $AIC = -2 \ln(L) + 2k$
- $BIC = -2 \ln(L) + k \ln(n)$
- $AICc = AIC + \frac{2k(k+1)}{n-k-1}$ (corrected for small samples)

where $k = p + q + P + Q + 1$ (plus 1 for variance).

Auto-SARIMA

Python's `pmdarima.auto_arima()` with `seasonal=True` automatically searches for optimal $(p, d, q) \times (P, D, Q)_s$.



Hyndman-Khandakar Algorithm (auto_arima)

How does automatic selection work? (Hyndman & Khandakar, 2008)

1. d : successive KPSS tests ($d = 0, 1, 2$); D : OCSB or Canova-Hansen test ($D = 0, 1$)
2. **Stepwise search**: starts from initial model, explores neighboring models
3. **Criterion**: AICc (correct for small samples)

Search Strategy

- **Initial model**: SARIMA(2, d , 2)(1, D , 1)_s or SARIMA(0, d , 0)(0, D , 0)_s
- **Variations tested**: ± 1 for each order (p, q, P, Q); stops when no neighbor improves AICc
- **Complexity**: $O(20-30)$ models evaluated (vs. $O(k^4)$ for grid search)

Python: `pm.auto_arima(y, seasonal=True, m=12, stepwise=True, trace=True)`

- Set `stepwise=False` for exhaustive search (slower, sometimes better)



Point Forecasts

Forecast Computation

SARIMA forecasts are computed recursively:

- Replace future ε_{T+h} with 0
- Replace future Y_{T+h} with forecasts $\hat{Y}_{T+h|T}$
- Use known past values Y_T, Y_{T-1}, \dots

Seasonal Pattern in Forecasts

SARIMA forecasts naturally capture seasonality:

- Short-term: influenced by recent values
- Long-term: revert to seasonal pattern



Forecast Intervals

Uncertainty Quantification

$(1 - \alpha)\%$ prediction interval:

$$\hat{Y}_{T+h|T} \pm z_{\alpha/2} \sqrt{\text{Var}(e_{T+h})}$$

Variance computed from MA(∞) representation.

Key Properties

- Intervals widen with forecast horizon
- For $I(1, 1)_s$ series: intervals grow without bound
- Seasonal pattern visible in point forecasts
- Uncertainty captures both trend and seasonal variation



Long-Horizon Forecasts

Behavior as $h \rightarrow \infty$

- Point forecasts converge to deterministic seasonal pattern
- If drift present: linear trend + seasonal pattern
- Forecast intervals continue to widen

Practical Implication

- Short-term: SARIMA captures both short-term dynamics and season
- Medium-term: Good seasonal forecasts, growing uncertainty
- Long-term: Mainly reflects seasonal pattern, wide intervals



The Seasonal Naive Benchmark

Definition: Seasonal Naive Forecast

- Formula:** $\hat{Y}_{T+h} = Y_{T+h-s}$ (last observed season)
- Monthly example:** Forecast for March 2025 = value from March 2024
- Interpretation:** "The simplest model that respects seasonality"

Why is it essential?

- Any SARIMA model **must** outperform the seasonal naive benchmark
- If it doesn't \Rightarrow the model complexity is not justified
- Surprisingly effective for many series with stable seasonality

Golden Rule

- Always** report SARIMA performance relative to seasonal naive
- This is the **first thing** a reviewer or manager checks



The MASE Metric: Proper Evaluation for Seasonal Series

MASE — Mean Absolute Scaled Error (Hyndman & Koehler, 2006)

- **Formula:**
$$\text{MASE} = \frac{\frac{1}{H} \sum_{h=1}^H |e_{T+h}|}{\frac{1}{T-s} \sum_{t=s+1}^T |Y_t - Y_{t-s}|}$$
- **Numerator:** mean absolute error of the model
- **Denominator:** mean absolute error of seasonal naive (on training data)

Interpretation

- $\text{MASE} < 1$: Model is **better** than seasonal naive
- $\text{MASE} = 1$: Model is **equivalent** to seasonal naive
- $\text{MASE} > 1$: Model is **worse** — abandon it!

Why MASE and not MAPE?

- MAPE: undefined for $Y_t = 0$; asymmetric; scale-dependent
- MASE: works with any data; symmetric; comparable across different series



Forecast Evaluation: Rolling Forecast Origin

Cross-Validation for Seasonal Time Series

- Principle:** re-estimate model → forecast h steps → advance 1 step → repeat
- Fixed window:** training on last w observations (constant size)
- Expanding window:** training from beginning to $T + i$ (grows)

Step-by-step procedure

1. Train SARIMA on Y_1, \dots, Y_T ; forecast $\hat{Y}_{T+1}, \dots, \hat{Y}_{T+h}$
2. Train SARIMA on Y_1, \dots, Y_{T+1} ; forecast $\hat{Y}_{T+2}, \dots, \hat{Y}_{T+h+1}$
3. ... repeat N times; compute RMSE, MAE, MASE on all N forecasts

Important

- Minimum $N \geq 2s$ origins (2 complete seasonal cycles) for reliable results
- Never “look ahead” — test data is strictly after training data



SARIMA vs Holt-Winters/ETS: When to Use Which?

Comparison

Criterion	SARIMA	ETS / Holt-Winters
Approach	Box-Jenkins (ACF/PACF)	Exponential smoothing
Seasonality	Stochastic (differencing)	Additive or multiplicative
Interpretation	AR/MA coefficients	Smoothing weights α, β, γ
Flexibility	Very flexible (7 params.)	Less flexible
Automation	auto_arima	ets() / ExponentialSmoothing

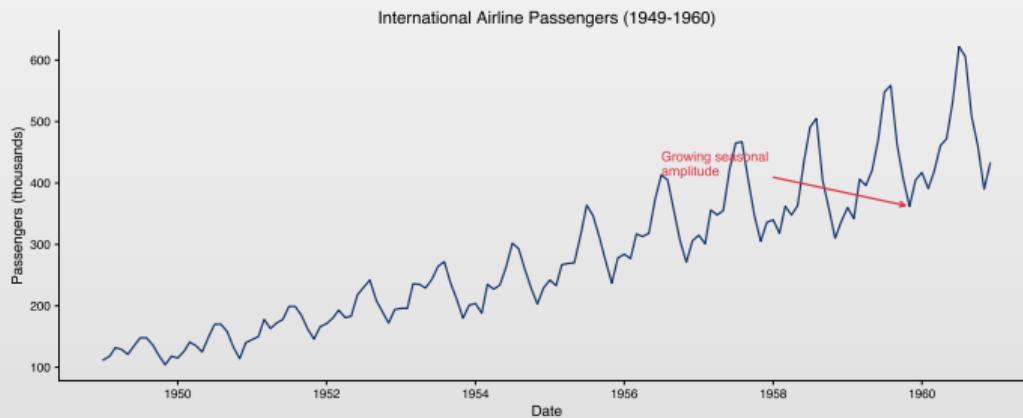
Practical Selection Guide

- SARIMA preferred:** series with complex autocorrelation, stochastic seasonality, ARMA components
- ETS preferred:** short series, stable seasonality, quick forecasts without diagnostics
- Best:** compare both on out-of-sample data and choose the winner



Case Study: Airline Passengers Data

- Classic Box-Jenkins dataset: monthly airline passengers (1949-1960)
- Clear upward trend and increasing seasonal amplitude
- Multiplicative seasonality suggests log transformation



 TSA_ch4_case_raw_data



Data Splitting Strategy

Time Series Train/Validation/Test Split

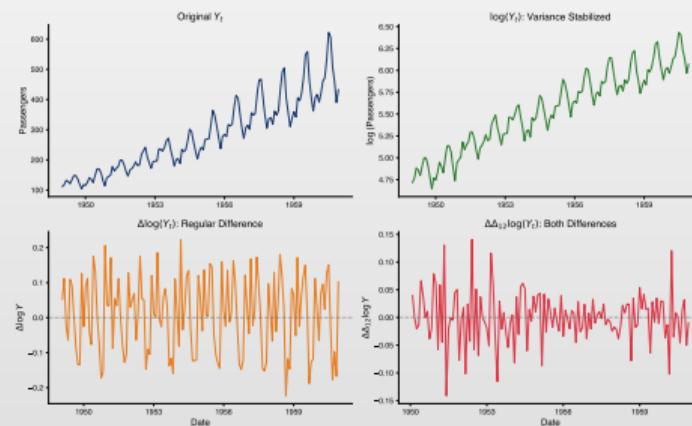


- **Training set (70%)** — Fit model parameters; estimate SARIMA coefficients ($\phi, \theta, \Phi, \Theta$)
- **Validation set (15%)** — Select best model; compare candidates, choose lowest validation error
- **Test set (15%)** — Final evaluation; unbiased out-of-sample performance, never used during development



Step 1: Transformations

- Log transform stabilizes variance (multiplicative \Rightarrow additive)
- First difference removes trend; seasonal difference removes seasonality
- Double-differenced series appears stationary

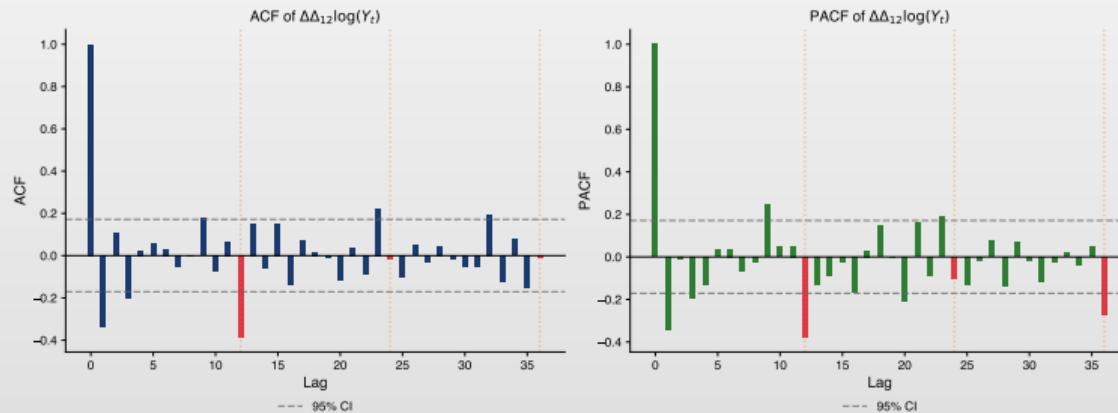


 TSA_ch4_case_transformations



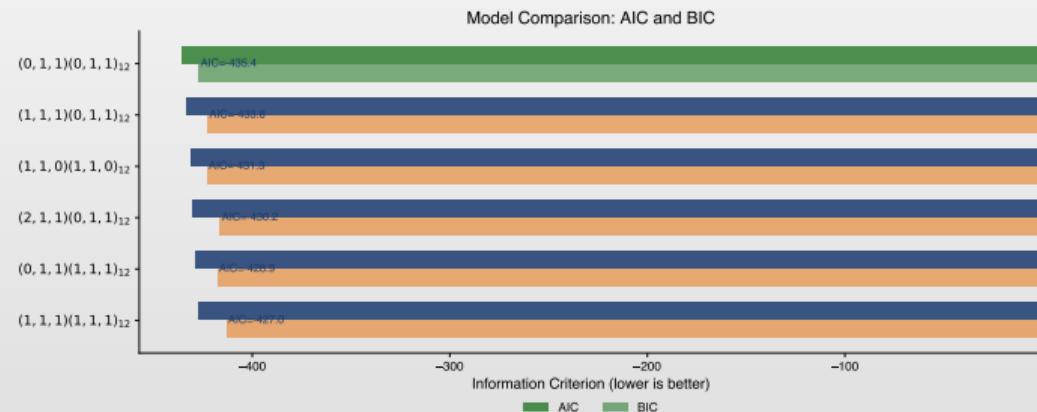
Step 2: ACF/PACF Analysis

- ACF: Significant spike at lag 1 and lag 12 \Rightarrow MA(1), SMA(1)
- PACF: Exponential decay pattern confirms MA structure
- Suggests SARIMA(0, 1, 1) \times (0, 1, 1)₁₂ (airline model)



Step 3: Model Comparison

- Compare candidate SARIMA models using AIC criterion
- SARIMA(0, 1, 1) × (0, 1, 1)₁₂ provides best fit (lowest AIC)
- This is the famous “airline model” identified by Box & Jenkins

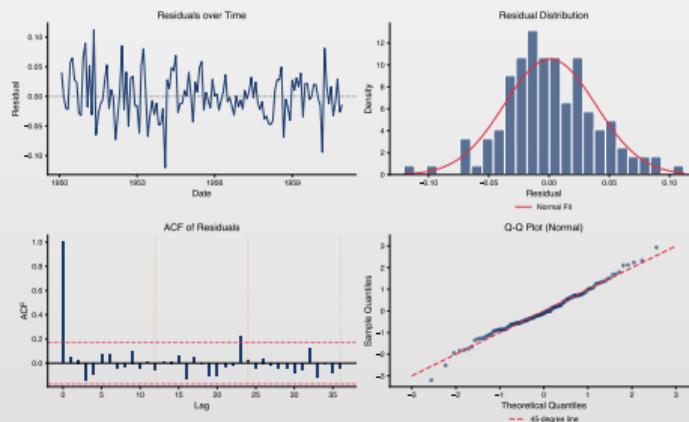


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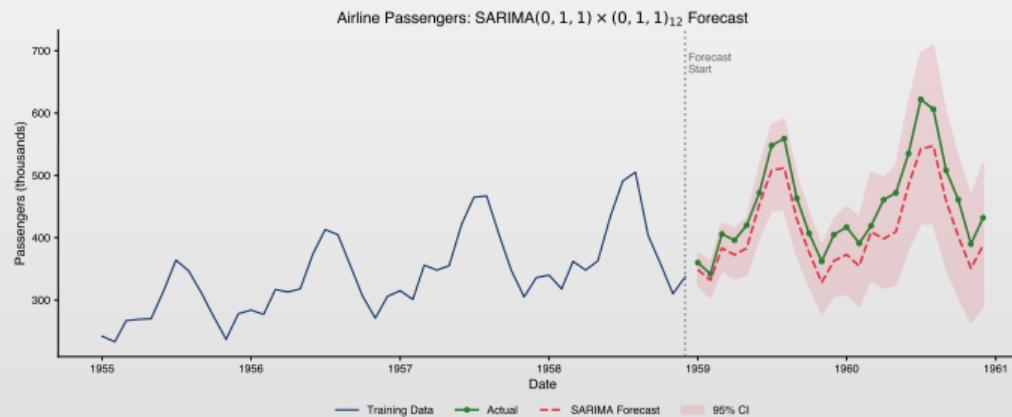
Step 4: Residual Diagnostics

- Residuals appear random with no remaining autocorrelation
- Q-Q plot shows approximate normality
- Model adequately captures both trend and seasonal structure



Step 5: Forecasting

- 24-month forecast with 95% confidence interval
- Model captures seasonal pattern and upward trend
- Prediction intervals widen appropriately with forecast horizon



Q TSA_ch4_case_forecast



Practical Pitfalls in SARIMA Modeling

1. Over-differencing

- Symptom:** ACF at lag 1 ≈ -0.5 (regular) or at lag $s \approx -0.5$ (seasonal)
- Cause:** applying $(1 - L)$ or $(1 - L^s)$ too many times
- Solution:** reduce d or D by 1 and re-examine ACF/PACF

2. Insufficient Data

- Minimum:** 3–4 complete seasonal cycles (36–48 monthly obs.); **recommended:** 5+ cycles
- Seasonal parameters Φ, Θ are estimated from lags $s, 2s, 3s, \dots$

3. Other Common Pitfalls

- Root cancellation:** $\phi \approx \theta$ suggests over-parameterization
- Parameters at invertibility boundary:** $|\theta| \approx 1$ or $|\Theta| \approx 1$ indicates problems
- Forgetting inverse transformation:** forecasts on log scale must be back-transformed!



X-13ARIMA-SEATS: Official Seasonal Adjustment

What is seasonal adjustment?

- Goal:** remove the seasonal component to reveal the true trend
- Users:** Eurostat, US Census Bureau, central banks, national statistical offices
- Example:** "GDP grew 0.3% compared to previous quarter" (seasonally adjusted data)

X-13ARIMA-SEATS (US Census Bureau)

- Step 1:** Identify and estimate a regARIMA model (SARIMA + calendar effects)
- Step 2:** Extract the seasonal component via SEATS or X-11 filters
- Step 3:** $Y_t^{\text{adjusted}} = Y_t - \hat{S}_t$ (additive) or $Y_t^{\text{adjusted}} = Y_t / \hat{S}_t$ (multiplicative)

Why does it matter for economists?

- Published macroeconomic data is almost always seasonally adjusted
- Misinterpreting unadjusted data can lead to erroneous conclusions



AI Exercise: Critical Thinking

Prompt to test in ChatGPT / Claude / Copilot

"I have the AirPassengers dataset from statsmodels (monthly data, international airline passengers, 1949–1960, 144 obs.). Identify seasonality, apply Box-Cox transform if needed, estimate a SARIMA model, and forecast 12 months. Give me complete Python code with plots."

Exercise:

1. Run the prompt in an LLM of your choice and critically analyze the response.
2. Does it check seasonality with ACF at lags $s, 2s, 3s$? Does it use STL decomposition?
3. Does it apply Box-Cox *before* differencing? Does it justify the choice of λ ?
4. How does it choose orders $(p, d, q) \times (P, D, Q)_s$? Only auto_arima or also ACF/PACF?
5. Does it evaluate with MASE relative to seasonal naïve? Does it use rolling forecast?

Warning: AI-generated code may run without errors and look professional. *That does not mean it is correct.*



Summary

What We Learned in This Chapter

- Seasonality in time series
 - ▶ Repetitive patterns at regular intervals; additive vs multiplicative
- Seasonal differencing and Box-Cox transformation
 - ▶ $(1 - L^s)$ removes stochastic seasonality; Box-Cox stabilizes variance
- SARIMA($p, d, q) \times (P, D, Q)_s$ models
 - ▶ Extend ARIMA with seasonal components; automatic selection via `auto_arima`
- Forecasting and evaluation
 - ▶ Benchmark: MASE relative to seasonal naive; rolling forecast out-of-sample

Key Idea

- **Parsimony principle:** The Airline Model $(0, 1, 1) \times (0, 1, 1)_{12}$ with only 2 parameters is remarkably effective for many seasonal economic series.



What's Next?

Chapter 5: Volatility Modeling — GARCH

- **Volatility:** conditional variation of financial returns
- **ARCH/GARCH:** models for conditional variance
- **Asymmetric extensions:** GJR-GARCH, EGARCH (leverage effect)
- **VaR:** Value-at-Risk based on GARCH models
- **Case study:** S&P 500 returns volatility

Questions?



Question 1

Question

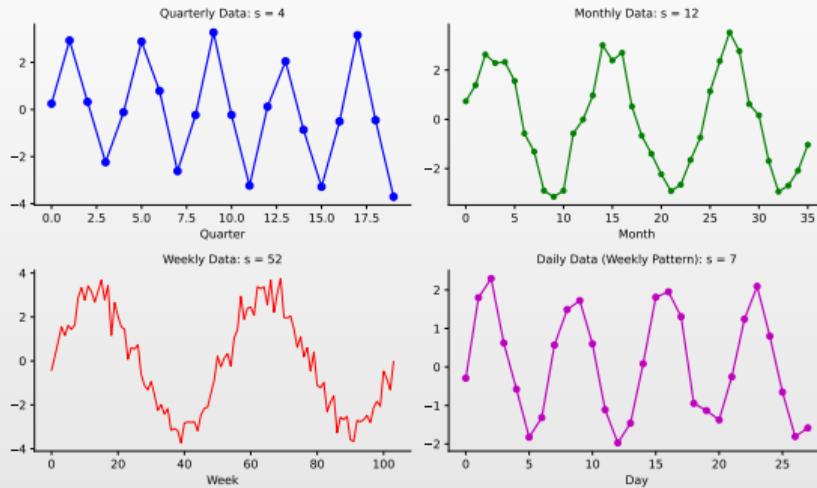
- For monthly data with annual seasonality, what is the seasonal period s ?

Answer Choices

- (A)** $s = 4$
- (B)** $s = 7$
- (C)** $s = 12$
- (D)** $s = 52$



Question 1: Answer



Answer: (C)

- $s = 12$ (12 months per year). Common periods: Quarterly=4, Monthly=12, Weekly=52, Daily=7, Hourly=24



Question 2

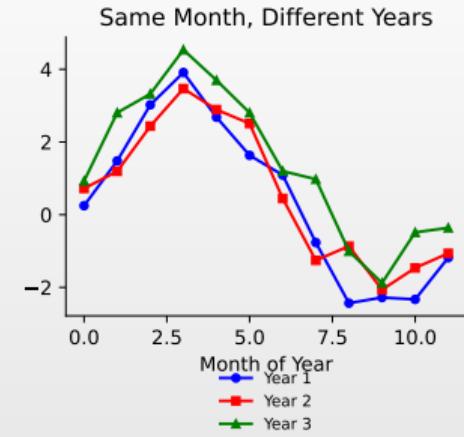
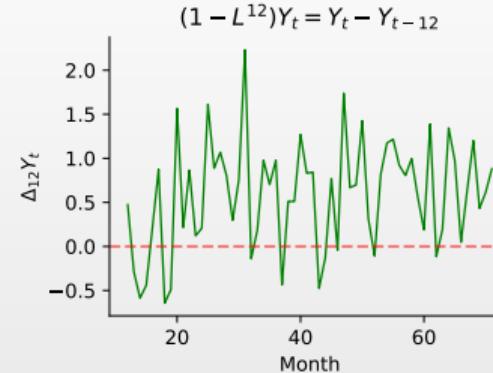
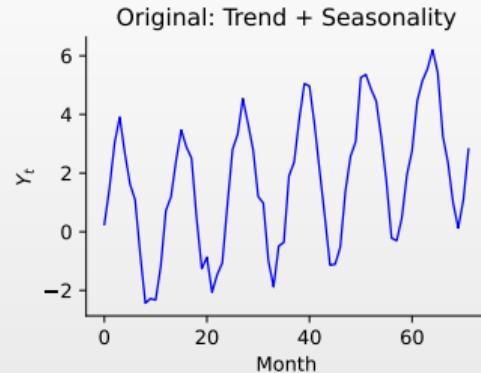
Question

- What does the seasonal difference operator $(1 - L^{12})$ do to a monthly series?

Answer Choices

- (A) Computes $Y_t - Y_{t-1}$ (month-to-month change)
- (B) Computes $Y_t - Y_{t-12}$ (year-over-year change)
- (C) Computes the 12-month moving average
- (D) Removes the trend component only

Question 2: Answer



Answer: (B)

- (1 - L¹²)Y_t = Y_t - Y_{t-12} removes the seasonal pattern by comparing same months.

Q TSA_ch4_quiz2_seasonal_diff



Question 3

Question

- In SARIMA(1, 1, 1) \times (1, 1, 1)₁₂ notation, what does the (1, 1, 1)₁₂ part represent?

Answer Choices

- (A) AR(1), differencing once, MA(1) at the regular level
- (B) Seasonal AR(1), seasonal differencing once, seasonal MA(1)
- (C) 12 AR terms, 12 differences, 12 MA terms
- (D) The model has 12 parameters in total

Question 3: Answer

Answer: (B)

- Seasonal AR(1), seasonal differencing once, seasonal MA(1)

SARIMA Notation Breakdown

SARIMA(p, d, q) \times (P, D, Q)_s:

- (p, d, q) Non-seasonal: AR(p), d differences, MA(q)
(P, D, Q)_s Seasonal: SAR(P), D seasonal diffs, SMA(Q)

For $(1, 1, 1) \times (1, 1, 1)_{12}$:

- Non-seasonal: AR(1), one regular difference, MA(1)
- Seasonal: SAR(1) at lag 12, one Δ_{12} , SMA(1) at lag 12



Question 4

Question

- The “Airline Model” is SARIMA(0, 1, 1) \times (0, 1, 1)₁₂. How many parameters need to be estimated (excluding variance)?

Answer Choices

- (A) 1
- (B) 2
- (C) 4
- (D) 12

Question 4: Answer

Answer: (B)

- 2 parameters: SARIMA(0, 1, 1) \times (0, 1, 1)₁₂: $(1 - L)(1 - L^{12})Y_t = (1 + \theta_1 L)(1 + \Theta_1 L^{12})\varepsilon_t$
- Parameters: θ_1 (non-seasonal MA) and Θ_1 (seasonal MA), plus σ^2 .

Why “Airline Model”?

Box & Jenkins (1970) used this model to forecast international airline passengers. Remarkably effective for many seasonal economic series!



Question 5

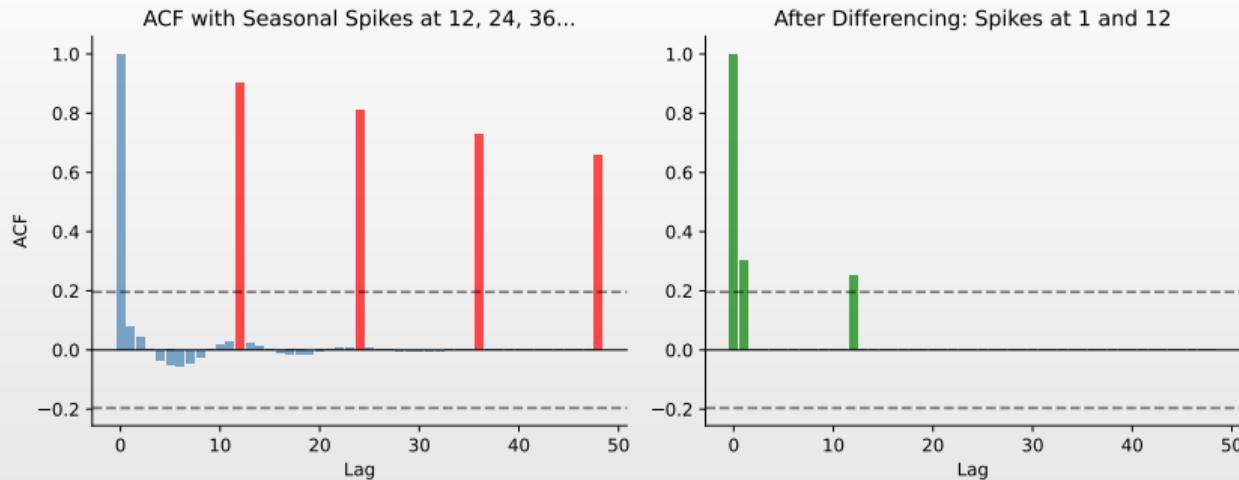
Question

- You observe significant ACF spikes at lags 12, 24, and 36 in a monthly series. What does this suggest?

Answer Choices

- (A) The series has a unit root
- (B) The series has annual seasonality that needs seasonal differencing
- (C) The series follows an AR(36) process
- (D) The series is already stationary

Question 5: Answer



Answer: (B)

- ACF spikes at 12, 24, 36 = stochastic seasonality. Apply $(1 - L^{12})$ to remove it.

Q TSA_ch4_quiz5_seasonal_acf



Question 6

Question

- After applying $(1 - L)(1 - L^{12})$ to a monthly series, the ACF shows a significant spike only at lag 1 and lag 12. What SARIMA model is suggested?

Answer Choices

- (A)** SARIMA(1, 1, 0) \times (1, 1, 0)₁₂
- (B)** SARIMA(0, 1, 1) \times (0, 1, 1)₁₂
- (C)** SARIMA(1, 1, 1) \times (1, 1, 1)₁₂
- (D)** SARIMA(0, 1, 0) \times (0, 1, 0)₁₂

Question 6: Answer

Answer: (B)

- Model:** SARIMA(0, 1, 1) \times (0, 1, 1)₁₂ (The Airline Model)

ACF/PACF Identification Rules

- Rule:** for MA processes, ACF cuts off after lag q
- ACF spike at lag 1:** MA(1) for non-seasonal part
- ACF spike at lag 12:** SMA(1) for seasonal part
- Combined:** MA(1) \times SMA(1) = (0, d , 1) \times (0, D , 1)₁₂
- With** $d = 1$, $D = 1$: (0, 1, 1) \times (0, 1, 1)₁₂



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- Hyndman, R.J., & Khandakar, Y. (2008). Automatic Time Series Forecasting: The forecast Package for R, *Journal of Statistical Software*, 27(3), 1–22.

Online Resources and Code

- Quantlet:** <https://quantlet.com> – Code platform for quantitative methods
- Quantinar:** <https://quantinar.com> – Learning platform for quantitative methods
- GitHub TSA:** https://github.com/QuantLet/TSA/tree/main/TSA_ch4 – Python code for this chapter



Thank You!

Questions?

Course materials available at: <https://danpele.github.io/Time-Series-Analysis/>

