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Lower Relative Contribution of Positive Family History to Colorectal Cancer Risk with Increasing Age: A Systematic Review and Meta-Analysis of 9.28 Million Individuals

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- OBJECTIVES:** Existing algorithms predicting the risk of colorectal cancer (CRC) assign a fixed score for family history of CRC. Whether the increased CRC risk attributed to family history of CRC was higher in younger patients remains inconclusive. We examined the risk of CRC associated with family history of CRC in first-degree relative (FDR) according to the age of index subjects (<40 vs. ≥40; <50 vs. ≥50; and <60 vs. ≥60 years).
- METHODS:** Ovid Medline, EMBASE, and gray literature from the reference lists of all identified studies were searched from their inception to March 2017. We included case-control/cohort studies that investigated the relationship between family history of CRC in FDR and prevalence of CRC. Two reviewers independently selected articles according to the PRISMA guideline. A random effects meta-analysis pooled relative risks (RR).
- RESULTS:** We analyzed 9.28 million subjects from 63 studies. A family history of CRC in FDR confers a higher risk of CRC (RR = 1.76, 95% CI = 1.57–1.97, $p < 0.001$). This increased risk was higher in younger individuals (RR = 3.29, 95% CI = 1.67–6.49 for <40 years versus RR = 1.42, 95% CI = 1.24–1.62 for ≥40 years, $p = 0.017$; RR = 2.81, 95% CI = 1.94–4.07 for <50 years versus RR = 1.47, 95% CI = 1.28–1.69 for ≥50 years, $p = 0.001$). No publication bias was identified, and the findings are robust in subgroup analyses.
- CONCLUSIONS:** The increase in relative risk of CRC attributed to family history was found to be higher in younger individuals. Family history of CRC could be assigned a higher score for younger subjects in CRC risk prediction algorithms. Future studies should examine if such approach may improve their predictive capability.

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INTRODUCTION

Globally, colorectal cancer (CRC) is the third commonest diagnosed cancer, and the second leading cause of cancer-related mortality [1]. It accounted for 10% of all malignancies and 8% of all cancer deaths. Its incidence and mortality are rising rapidly in Asia Pacific countries, and poses substantial public health challenges in terms of its healthcare costs incurred [2]. Most CRC are sporadic with one-fifth associated with familial clustering and 6% attributed

to familial adenomatous polyposis syndrome and hereditary nonpolyposis colorectal cancer [3].

In the past decade, the concept of personalized medicine and targeted screening has become increasingly popular [4]. Risk scores devised and validated to predict the risk of CRC could improve detection yield and optimize screening efficiency. Clinically, it allows individuals to be more aware of their own risk, make informed decisions on choices of screening tests, and tailor the

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intensity of screening or prevention approaches to the predicted level of risk. From a public health perspective, the implementation of risk stratification strategies may better justify allocation and utilization of colonoscopic resources, facilitate resource planning in the formulation of population-based screening programs, and reduce healthcare costs. There are several at-risk groups who should receive earlier screening, and colonoscopy is more preferred among the increased risk individuals.

A recent systematic review provided a comprehensive analysis of risk prediction tools for risk of CRC in asymptomatic individuals within the general population [5]. A total of 52 risk models have been identified, and among them 33 included family history of CRC as a major predictor. Among these predictors, the risk increase is strongest for subjects with first-degree relatives (FDRs) with CRC, where their risk of CRC could be increased by two- to three-fold [6, 7]. However, current risk scoring systems assigned a fixed score for a positive family history of CRC, irrespective of the age of the screening participant. There have been three meta-analyses that assessed the association between familial CRC risks in FDR of CRC according to the nature of family history [8–10]. A 2001 meta-analysis was limited by the modest number of studies, making comparison of CRC risk according to age of screening subjects who reported family history of CRC underpowered [8]. Two subsequent meta-analyses showed the relative risk (RR) of CRC among FDRs decreased with age of CRC diagnosis in the proband [9, 10]. Since then, many large-scale studies have been published, including a multinational, multicenter study from our expert panels in the Asia Pacific Working Group for CRC [11]. It is still uncertain if the increased risk of CRC conferred by a family history of CRC in FDR is significantly higher in younger individuals than older subjects. An updated meta-analysis is needed to address this important knowledge gap.

The objective of this systematic review and meta-analysis is to quantify the risk of CRC in asymptomatic subjects with a FDR with CRC compared with those with no such family history. We tested the *a priori* hypothesis that the increased CRC risk associated with a family history of CRC in FDR was not higher in younger subjects compared to older subjects. If this null hypothesis is rejected, the finding would serve as an evidence to recommend that a higher score should be assigned to family history of CRC in younger individuals.

METHODS

Search strategy and selection criteria

This systematic review and meta-analysis was performed with adherence to the PRISMA (Preferred Reporting Items for Systematic reviews and meta-Analyses) statement [12], performed according to a pre-determined protocol. Ovid Medline and EMBASE were searched from their inception to March 2017, as were gray literature from the reference lists of all identified studies. Three main categories of the search terms were used: “colorectal cancer”, “family”, and “study design”. All searches were limited to English language. The following search terms were used: (1) AND (2) AND (3) AND (4):

(1) “colon”, “colonic”, “rectal”, “rectum”, “colorectal”, “colorectum”, “bowel”

(2) “neoplas”, “cancer”, “carcinoma”, “adenoma”, “adenocarcinoma”, “polyp”

(3) “family”, “familial”, “parent”, “father”, “mother”, “paternal”, “maternal”, “offspring”, “child”, “children”, “son”, “daughter”, “sibling”, “brother”, “sister”

(4) “case control”, “cohort”, “cross-sectional”, “observational”, “epidemiologic”, “prospective”, and “retrospective”. Supplementary Table 1 shows detailed search strategy.

All articles extracted from the databases were filtered. Reference lists of eligible studies and related meta-analyses were hand searched to identify further relevant studies. Two reviewers (CHC and JL) independently screened all abstracts identified in the initial search and excluded studies not fulfilling the eligible criteria. The search was restricted to case-control and cohort studies that investigated the relationship between family history of CRC in FDR and incidence/prevalence of CRC. The following types of studies were excluded:

1. Studies that recruited symptomatic patients;
2. Studies with duplicated data;
3. Studies not specifying the type of family history;
4. Studies that did not define subjects in their “non-exposure group” as those without family history of CRC in FDRs.

A third reviewer (JLWH) reviewed all selected studies to ensure they met the inclusion criteria.

Data extraction

Data from summary estimates of all eligible studies were obtained. CHC and JL extracted data from all selected full-text articles reviewed in duplicate, and in the case of disagreement, consensus was reached via referral to a third reviewer (MCSW). The information on study population and study design was extracted into a database. All estimations of the RR and their precision were collected. If they were not provided, the estimations of RR were calculated as frequencies of the exposed and non-exposed among cases and controls. The odds ratios (ORs) were extracted directly with an assumption that ORs are good estimates of RR [13], since CRC is relatively rare among asymptomatic subjects (prevalence <1%) [14]. Adjusted risk estimates were used if they were adjusted for relevant effect moderators; otherwise unadjusted estimates or raw data were collected to calculate the relative risks. When data were shown separately in strata, Mantel-Haenszel method was used to pool raw data [15]. The inverse variance method was used for data presented as relative risks and confidence intervals (CIs). Stratified data were also recorded and categorized according to potential effect modifiers for subgroup analysis.

Quality assessment

The Newcastle-Ottawa Scale (NOS) is used as a quality assessment tool for observational studies [16, 17]. It is a 9-point scale with eight items on the selection of the study groups, the comparability of the groups, and the ascertainment of either the exposure or outcome of interest for case-control or cohort studies, respectively. The quality of studies was scored based on these eight items. The main items for study quality scoring were as follows: (1) representativeness of the exposed cohort: one point was assigned if the subjects represent the general population; no points were

assigned if samples are special population groups (e.g., veteran) or not mentioned; (2) selection of the non-exposed cohort: one point was assigned if the non-exposed cohort (i.e., subjects without family history) was drawn from the same community as the exposed cohort (i.e., subject with family history); (3) ascertainment of exposure: one point was assigned if family history was ascertained by healthcare professionals; (4) demonstration that outcome of interest was not present at study commencement: one point was assigned for stating exclusion of subjects with CRA (colorectal adenoma)/advanced CRA/CRC or stating subjects have no history of CRA/advanced CRA/CRC; (5) comparability of cohorts on the basis of the design or analysis: two points were assigned for studies that adjusted for covariates in the statistical analysis, and one point was given if there was no adjustment; (6) assessment of outcome by colonoscopy and histological examination: one point was assigned if it was based on medical records; (7) follow-up duration: one point was assigned for all eligible studies if the follow-up period is long enough to detect the outcome of interest; (8) adequacy of follow-up among cohorts: one point was assigned for completion of at least 90% follow-up. Scores ranged from 0 (lowest) to 9 (highest). Similar to a previous meta-analysis [18], studies with scores ≥ 7 and < 7 were classified as “high” and “low” quality, respectively.

Statistical analysis

The relative risk was used as the effect measure of outcomes. Random effects meta-analysis was performed using the inverse variance method. Between-study variance in the random effects model was estimated by a restricted maximum-likelihood estimator [19]. Heterogeneity between the included studies was assessed using the I^2 statistic from the standard chi-square test [20]. This describes the percentage of the variability in the effect estimates resulting from heterogeneity. $I^2 > 50\%$ was considered to reflect significant statistical heterogeneity and in such cases the random effects model was used, otherwise the fixed effects model was used.

Owing to heterogeneity and possible between-factor interactions, subgroup analyses were conducted by random effects models with separate estimates for between-study variance across different subgroups [21]. The subgroup differences were tested by the Q-tests. Meta-regression was used to investigate the effect of any potential confounders. Univariate meta-regression analysis was used and followed by multivariate modeling for covariates with $p < 0.2$. We compared the risk of CRC in subjects at different age with affected FDRs, namely, ≥ 40 years versus < 40 years; ≥ 50 years versus < 50 years; and ≥ 60 years versus < 60 years. To explore possible publication bias, a funnel plot was used as a graphical presentation of the data. Begg's and Egger's regression tests were performed to examine funnel plot asymmetry by Comprehensive Meta Analysis (version 2.2, Biostat, Inc., 2011) [22]. The funnel plot is a scatter plot of the effect estimate from every study enrolled in the meta-analysis against the measure of its precision (1/standard error). If either the Begg's or Egger's test indicated publication bias, the Trim and Fill method would be used to estimate the number of missing studies and the treatment effect after adjustment for small-study effect [23]. All of these

analyses were conducted by using R version 3.3.2 with package meta ver 4.6-0 and metafor ver 1.9-9.

Subgroup analysis

Since the data were expected to be heterogeneous, we performed 10 additional subgroup analyses on the risk of CRC according to the following characteristics: (1) number of FDR affected: one versus ≥ 1 versus ≥ 2 FDRs; (2) age of affected relatives; (3) type of family history exposure: parents versus siblings; (4) site of CRC in the proband: colon versus rectum; (5) gender of the proband; (6) gender of exposed relatives; (7) study design: cohort studies versus case-control studies versus nested case-control studies at different settings; (8) outcomes: incidence versus mortality; (9) geographic region of study: Western Pacific versus America vs. Eastern Mediterranean versus Europe; and (10) method of assessment of family history: medical records versus self-reported history. These subgroup analyses are important as we perceived them as potential effect modifiers of the present meta-analysis. The authors had full access to all the data in the study and had final responsibility for the decision to submit for publication.

RESULTS

We identified 4119 articles from our search strategy (Fig. 1). The following articles were excluded: duplicates ($n = 863$) or irrelevance to the present research question ($n = 2444$). The full texts of the remaining 812 articles were reviewed, with 45 articles fulfilling our eligibility criteria being selected, and 18 additional studies selected from reference lists of eligible articles. Therefore, a total of 63 articles were included in this meta-analysis, including 9,284,074 patients (Supplementary Table 2) [24–86]. The earliest calendar year of subject enrollment was 1982. Among these, 43, 13, and 7 studies were conducted in western countries, Western Pacific nations, and eastern Mediterranean, respectively. Most of the selected studies were case-control studies ($n = 56$), examined risk of CRC in at least one FDR affected by CRC ($n = 63$), assessed family history of CRC based on self-reports or surveys ($n = 54$), and provided incidence estimates of CRC ($n = 62$). Based on NOS score of ≥ 7 , a total of 30 studies were classified as high quality (Supplementary Table 3).

From available data among the selected studies, subjects with a FDR with CRC were significantly more likely to have CRC ($RR = 1.76$, 95% $CI = 1.57$ – 1.97 , $p < 0.001$) compared to subjects with no family history (Fig. 2). A total of 50 studies showed significantly positive association. The heterogeneity was high ($I^2 = 95.7\%$; Q ($df = 62$) = 756.9 , $p < 0.001$). No publication bias was observed (Begg's regression test $p = 0.133$; Egger's regression test $p = 0.078$; see funnel plot in Fig. 3).

The relative risk of CRC among those with a FDR with CRC was further analyzed according to the age of the index subject, age of affected relatives, number of FDRs affected, type of family history exposure (parents versus siblings), site of CRC (colon versus rectum) in the affected FDR, gender of the index subject and affected FDR, assessment method of family history, study design, clinical outcome of CRC (incidence versus mortality), geographic area of

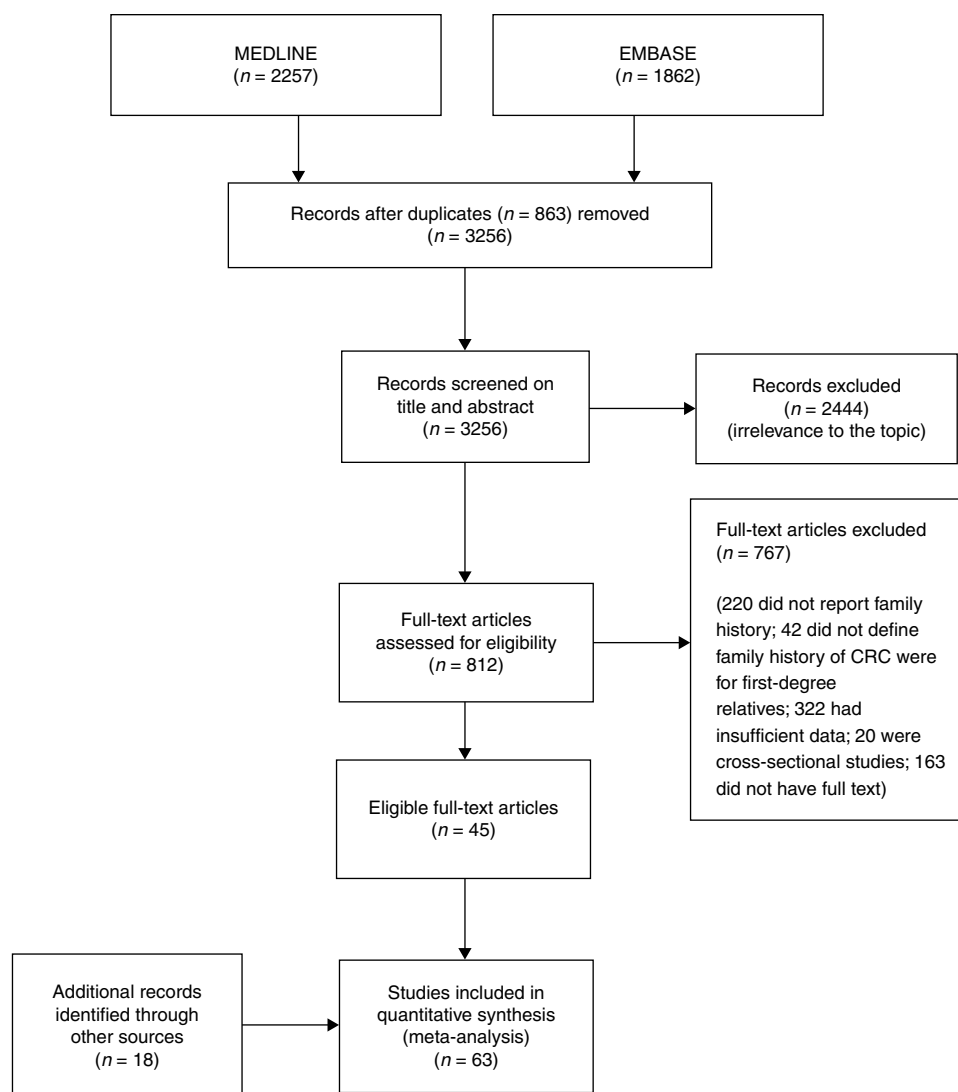


Fig. 1 Flow diagram of study selection

study, and study quality (Fig. 4). The impact of family history on risk of CRC was significantly stronger when the index subject was aged <40 years ($RR = 3.29$, 95% $CI = 1.67$ – 6.49 , $I^2 = 78.4\%$ versus $RR = 1.42$, 95% $CI = 1.24$ – 1.62 , $I^2 = 95\%$, $p = 0.017$ for ≥ 40 years) and <50 years ($RR = 2.81$, 95% $CI = 1.94$ – 6.07 , $I^2 = 81.0\%$ versus $RR = 1.47$, 95% $CI = 1.28$ – 1.69 , $I^2 = 91.8\%$, $p < 0.001$ for ≥ 50 years). Individuals with ≥ 2 FDRs affected had higher risk of CRC than those with one FDR affected ($RR = 2.68$, 95% $CI = 1.92$ – 3.74 , $I^2 = 75.0\%$ versus $RR = 1.82$, 95% $CI = 1.51$ – 2.18 , $I^2 = 76.3\%$), but the difference did not reach statistical significance as demonstrated by an overlap of their 95% CI s. Other subgroup analysis did not identify inter-group differences. No significant difference was observed between studies that reported relative risks and odds ratios ($p = 0.12$); between crude and adjusted effect size ($p = 0.12$); and studies that were classified as high (NOS score ≥ 7) and low quality (NOS score < 7). From meta-regression analysis, it was found that age (coefficient = 0.366 , 95% $CI = 0.148$, 0.585 , $p = 0.001$) and body mass index (coefficient = -0.449 , 95%

$CI = -0.753$, -0.145 , $p = 0.004$) accounted for the heterogeneity detected in this study (Supplementary Table 4).

DISCUSSION

The main findings of this systematic review of 63 studies involving more than 9 million patients show that: (1) a FDR with a history of CRC conferred a 1.76-fold increased risk of CRC compared to those without such a family history; (2) the increased risk of CRC associated with a family history of CRC in FDR was significantly higher when the index subject is younger than 50 years old than in subjects older than 50 years; and (3) the higher risk of CRC for index subjects with two or more affected FDRs compared to those with only one affected FDR was marginal.

Thus far, there were three published meta-analyses that investigated the association between family history and risk of CRC [8–10]. The pooled risk estimate of developing CRC given at least one affected FDR ranged between 2.26 and 2.28, which was higher

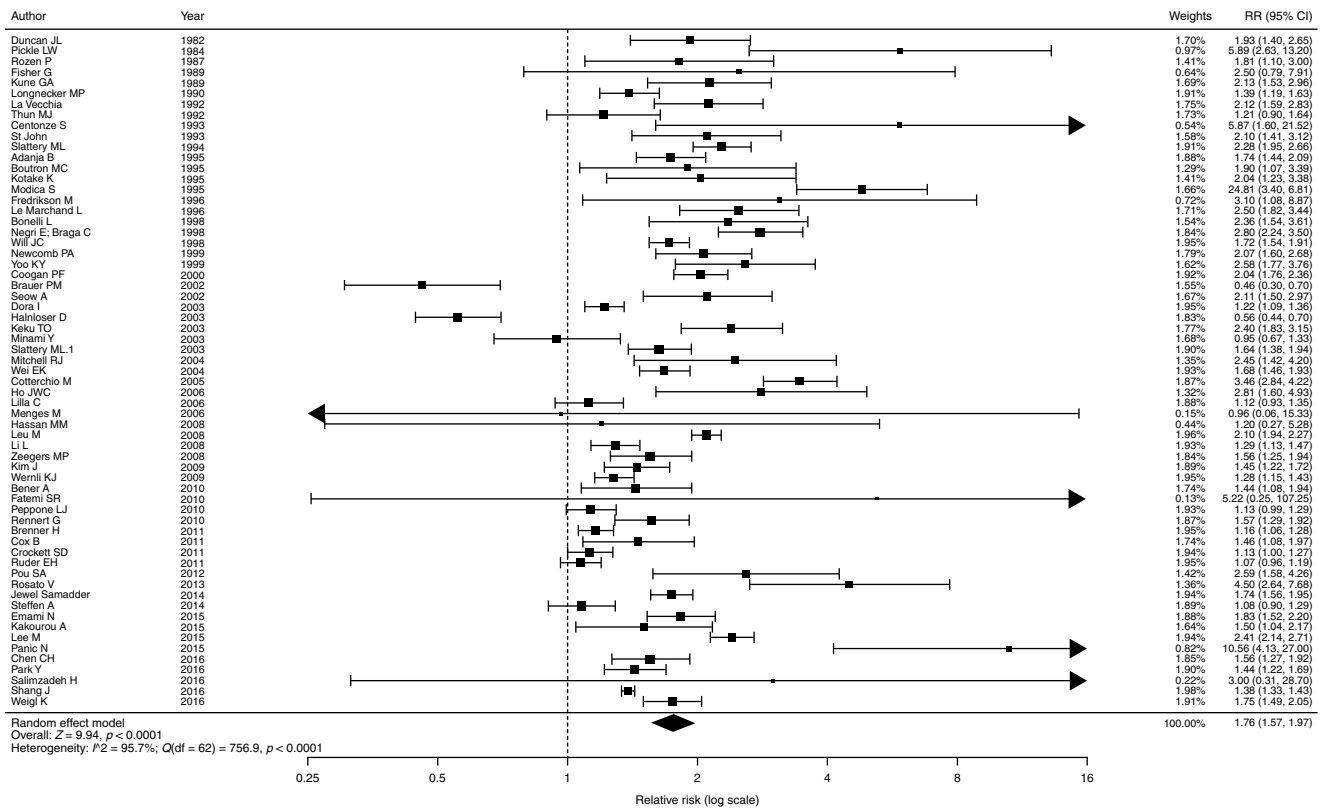


Fig. 2 Pooled odds ratios of colorectal cancer in individuals with family history of colorectal cancer compared with control

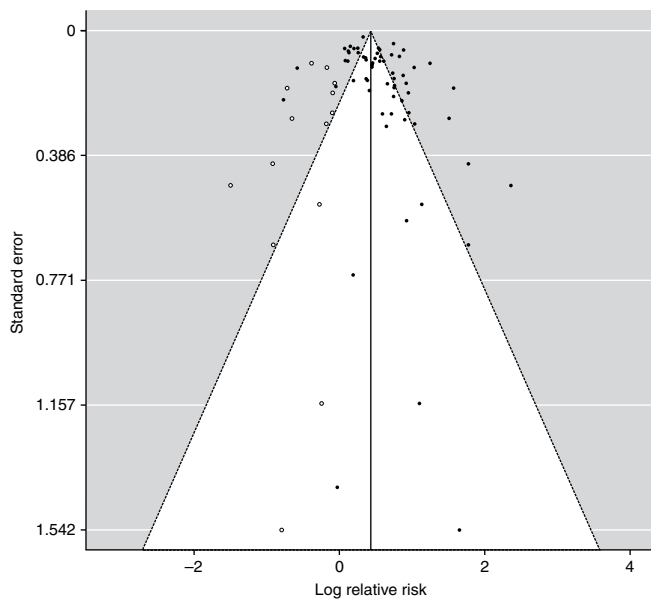


Fig. 3 Funnel plot for publication bias

than the relative risk (1.76) reported in this study. This may be due to the smaller number of studies ($n = 27$ and 20 , respectively) included in the two previous meta-analyses [8, 10] compared with this study, difference in calendar years where the included studies were published (up to late 1990 and early 2000 in previous meta-analyses), and the inclusion of cross-sectional studies in

one meta-analysis [9]. It should also be noted that some studies presented type I relative risk, which refers to risk of CRC in an index person given a specific family history of CRC, whereas some reported type II relative risk, which signifies risk of CRC in a relative given a specific index person is affected [10].

In our meta-analysis, 13 out of 63 studies showed no significantly increased risk of CRC among individuals with family history of CRC [28, 39, 40, 42, 43, 56, 58, 59, 66, 72, 73, 79, 80], compared with other 50 studies that reported increased relative risk. Some possible explanations for this difference include the accuracy of self-reported family history, varying quality of colonoscopy and experience of endoscopists across studies, the number of risk factors of CRC among subjects included in these studies, and differences in their region of residence. Most of the studies included in the present meta-analysis did not match between those having family history and those without, recruited consecutive patients who had CRC symptoms as controls, and included individuals' FDRs who were not examined by colonoscopies for ascertainment of the presence/absence of CRC. In addition, differences in the methodological design of these studies could lead to differences in conclusions drawn. For instance, relative risk estimates from prospective cohort studies are in general lower than that from retrospective studies due to the presence of recall bias in the latter [9].

The reasons why the additional risk of CRC conferred by family history was higher in younger subjects remain speculative. Early age at diagnosis has been thought to be indicative of inherited genetic mutations. One might hypothesize that low-penetrant

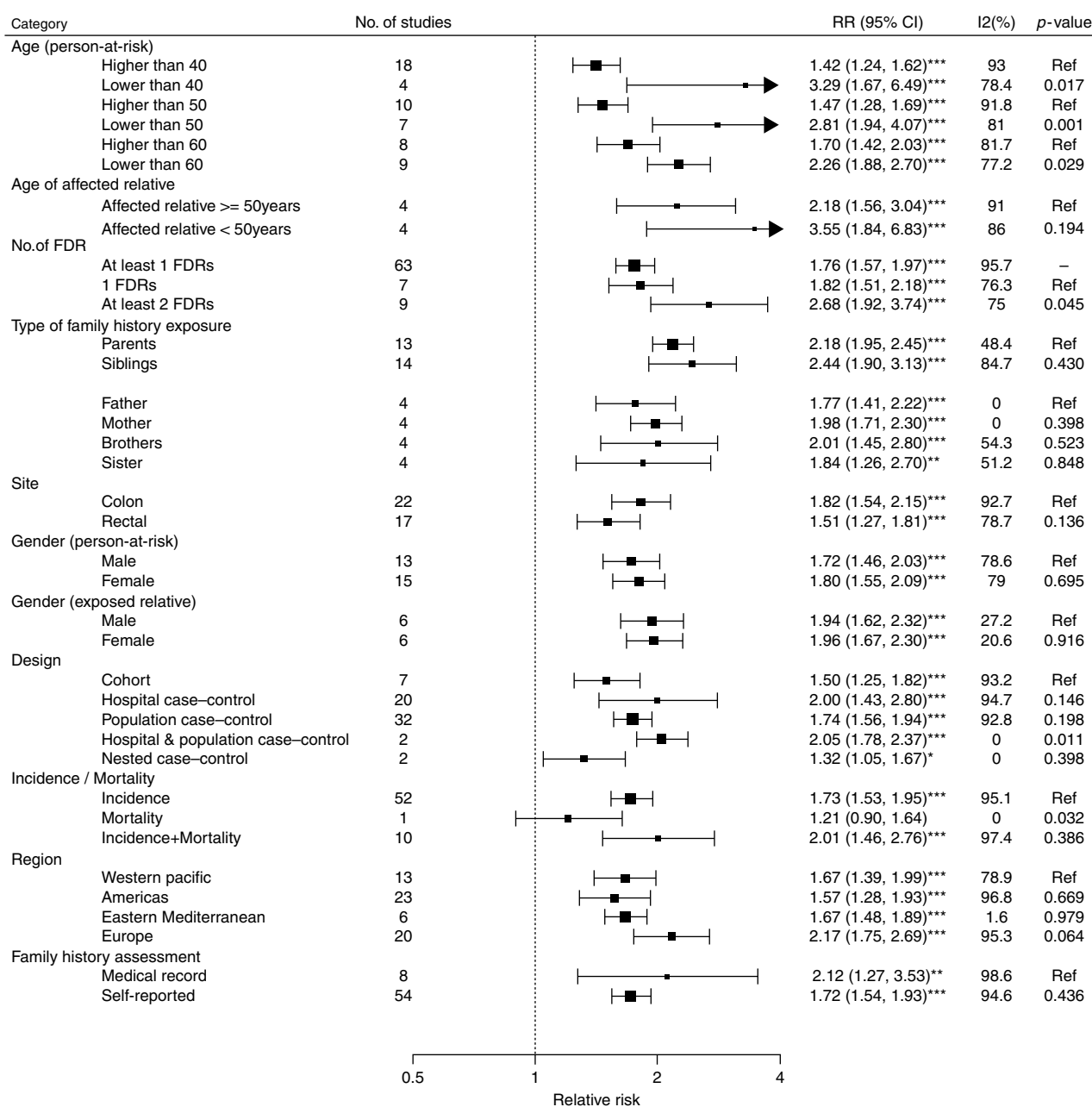


Fig. 4 Subgroup analysis—association between family history of colorectal cancer in a first-degree relative and incidence/prevalence of colorectal cancer. The “*” indicate different significant levels of z tests in meta-analysis. * $P \leq 0.05$; ** $P \leq 0.01$; *** $P \leq 0.001$

genes interacting with diet and lifestyle factors contribute differentially to individuals of different ages when the family history of CRC is reported.

Study limitations

Our meta-analysis has a number of strengths. Firstly, it included asymptomatic, average-risk subjects pooled from all selected studies conducted in both western and Asian countries. In the literature, the retrospective observational studies of the effectiveness of screening on CRC incidence reduction are all struggling with the contamination of screening and diagnostic examinations. Hence, the application of its findings based on asymptomatic participants is more generalizable to screening practices

when compared with previous colonoscopy studies. Secondly, this meta-analysis includes a large number of screening participants recruited from all eligible studies. Thirdly, various effect modifiers were addressed in subgroup analyses. Also, quality assessment was based on an internationally recognized NOS scale for all the selected articles in a systematic manner.

Several limitations should be noted. Firstly, informal reports might have been omitted by our search strategy. Secondly, a high level of heterogeneity was observed, but this could be partially explained by adjusting for age and body mass index. Thirdly, not all studies have been adjusted for potential confounders such as dietary habits and physical inactivity. In addition, prevalence and distribution of colorectal neoplasia varied in people of different

ethnicities yet the present studies recruited subjects from certain continents only. The heterogeneity of the study was reduced by subgroup analysis based on ethnicity. Ascertainment of both family history and the outcomes might have improved over time. Finally, the data on the age for the diagnosis of the FDR were limited in the current analysis. One of the most important questions in clinical practice is when to start CRC screening for individuals who have family history of CRC in a FDR. The current consideration that may trigger a colonoscopy in an individual aged <40 years with a family history of CRC is that the proband should receive colonoscopy 10 years younger than the FDR with CRC. The relationship between the age of diagnosis of the FDR and CRC risk is also important, and few studies have examined this issue. For instance, Lee et al. [53] reported that CRC risks were highest in siblings who were diagnosed at a younger age, with incident rate ratios of 9.1, 6.0, 2.6, 2.2, and 1.9 for the 30–40, 50–60, 60–70, and 70–75 age groups, respectively. Also, Slattery et al. [77] also found greater associations when the FDR was diagnosed with CRC at the age of less than 50 years or younger. The odds ratios were 2.1, 1.6, and 1.0, respectively, for age of diagnosis at <50 years, 51–64 years, and ≥65 years, respectively. These findings imply that the younger the FDR with CRC, the earlier the proband should be investigated with colonoscopy. Nevertheless, there have been few studies that have directly addressed whether probands of age 40–50 years who had CRC were related in age to the FDR who were in their fifth decade of life. The correlation between the ages of FDRs and the ages of the probands should be studied in future evaluations.

Our meta-analysis shows that among subjects with family history of CRC in FDR, the increased risk of CRC associated with the family history was higher in young individuals (<50 years) than in older individuals (>50 years). This finding implies that family history of CRC in FDR could be assigned a higher score for younger subjects in CRC risk prediction algorithms. Future studies should examine if such approach may improve their predictive capability.

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CONFLICT OF INTEREST

Guarantor of the article: CH Chan and Jiayan Lin.

Specific author contributions: Martin CS Wong had full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis. Study concept and design: Martin CS Wong, John CT Wong, Justin CY Wu, Francis KL Chan; acquisition, analysis, or interpretation of data: Martin CS Wong, CH Chan, Jiayan Lin, Jason LW Huang, Junjie Huang; drafting of the manuscript: Martin CS Wong, Justin CY Wu, Francis KL Chan; critical revision of the manuscript for important intellectual content: John CT Wong, Gary Tse, Justin CY Wu, Francis KL Chan; statistical analysis: CH Chan, Jiayan Lin, Jason LW Huang, Junjie Huang, Yuan Fang, Wilson WL Cheung; administrative, technical, or material support: Martin CS Wong, CH Chan, Jiayan Lin, Jason LW Huang, Junjie Huang; study supervision: Justin CY Wu, Francis KL Chan.

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Study Highlights

WHAT IS CURRENT KNOWLEDGE

- Current risk scoring systems assigned a fixed score for a positive family history of colorectal cancer (CRC), irrespective of the age of the screening participant.
- However, whether the increased CRC risk associated with family history was higher in younger patients remains inconclusive.
- Previous meta-analyses are limited by the modest number of studies, making comparison of CRC risk according to age of screening participants who reported family history of CRC underpowered.
- We examined the risk of CRC associated with family history of CRC in first-degree relative (FDR) according to the age of index subjects.

WHAT IS NEW HERE

- In this meta-analysis including 63 studies (9.28 million asymptomatic screening participants), the increased CRC risk associated with family history was higher in younger individuals.
- No publication bias was observed and the findings are robust in subgroup analysis.
- Family history of CRC could be assigned a higher score for younger subjects in CRC risk prediction algorithms.
- Future studies should examine if such approach may improve their predictive capability.

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