

Denoising Lévy Probabilistic Models - DLPM

Denoising Diffusion Models with Heavy Tails

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Introduction on Diffusion Models

DDPM – Overview

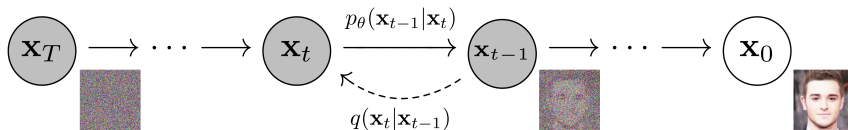


Figure: Forward/backward structure, discrete time [HJA20]

Setup, discrete time:

- **Forward process** $\{X_t\}_{t=0}^T$ is a Markov chain with Gaussian transition kernels $p_{t+1|t}(\cdot|\cdot)$, such that

$$X_0 \sim p_0 \text{ (the data)}, \quad X_T \sim p_T \approx \mathcal{N}(0, I_d) \text{ (the noise)} \quad (1)$$

- **Generative process** $\{\bar{X}_t^\theta\}_{t=0}^T$ will be a Markov chain running in reverse time
- **Training loss** Fit the joint distributions with an **ELBO loss**, like in VAEs.

DDPM – Overview

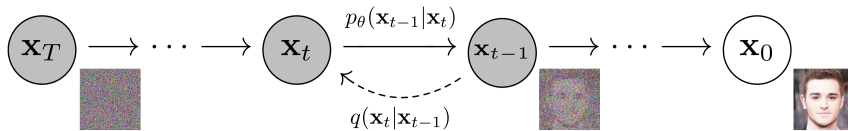


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DDPM – Forward Process

- **Forward process (Markov chain):**

$$X_{t+1} = \sqrt{\alpha_t} X_t + \sqrt{1 - \alpha_t} \epsilon_t, \quad (2)$$

where $\{\alpha_t\}_{t=0}^{T-1}$ is a noise schedule, $0 < \alpha_t < 1$.

- **Closed form for $X_t \mid X_0$, by stability of the Gaussian distribution:**

$$X_t \mid X_0 \stackrel{d}{=} \sqrt{\bar{\alpha}_t} X_0 + \sqrt{(1 - \bar{\alpha}_t)} \mathbf{I}_d \bar{\epsilon}_t, \quad (3)$$

with $\bar{\alpha}_t = \prod_{s=1}^t \alpha_s$, chosen such that X_T is approximately distributed as $\mathcal{N}(0, \mathbf{I}_d)$.

DDPM – Backward Process

- **Reformulating the forward process** Let us examine its joint distribution

$$\begin{aligned}
 p(x_0, \dots, x_T) &= p_0(x_0) \cdot \prod_{t=1}^T p_{t|t-1}(x_t|x_{t-1}) \\
 &= p_0(x_0) \cdot p_{1|0}(x_1|x_0) \cdot \prod_{t=2}^T p_{t|t-1}(x_t|x_{t-1}, x_0) \\
 &= p_0(x_0) \cdot p_{1|0}(x_1|x_0) \cdot \prod_{t=2}^T \frac{p_{t-1|t,0}(x_{t-1}|x_t, x_0) p_{t|0}(x_t|x_0)}{p_{t-1|0}(x_{t-1}|x_0)} \quad , \text{ by Bayes rule} \\
 &= \underbrace{p_0(x_0)}_{\text{data}} \cdot \underbrace{p_{T|0}(x_T|x_0)}_{\text{noise}} \cdot \prod_{t=2}^T \underbrace{p_{t-1|t,0}(x_{t-1}|x_t, x_0)}_{\text{Gaussian transitions}}
 \end{aligned}$$

- **Gaussian transitions** $p_{t-1|t,0}(\cdot|x_t, x_0)$ is the density of $\mathcal{N}(\tilde{m}_t(x_t, x_0), \tilde{\Sigma}_t)$.

DDPM – Backward Process

Gaussian transitions Again, by Bayes rule:

$$\begin{aligned}
 p_{t-1|t,0}(x_{t-1}|x_t, x_0) &= \frac{p_{t|t-1}(x_t|x_{t-1}, x_0)p_{t-1|0}(x_{t-1}|x_0)}{p_{t|0}(x_t|x_0)} \\
 &= \frac{p_{t|t-1}(x_t|x_{t-1})p_{t-1|0}(x_{t-1}|x_0)}{p_{t|0}(x_t|x_0)} \\
 &\propto \exp\left(-\frac{\|x_t - \sqrt{\alpha_t}x_{t-1}\|^2}{2(1-\alpha_t)} - \frac{\|x_{t-1} - \sqrt{\bar{\alpha}_{t-1}}x_0\|^2}{2(1-\bar{\alpha}_{t-1})} - \frac{\|x_t - \sqrt{\bar{\alpha}_t}x_0\|^2}{2(1-\bar{\alpha}_t)}\right) \\
 &\propto \dots \\
 &\propto \exp\left(-\frac{\|x_{t-1} - \tilde{m}_t(x_t, x_0)\|^2}{2\tilde{\Sigma}_t}\right)
 \end{aligned}$$

with

$$\tilde{m}_t(x_t, x_0) = \frac{\sqrt{\bar{\alpha}_{t-1}}(1-\alpha_t)}{1-\bar{\alpha}_t}x_0 + \frac{\sqrt{\alpha_t}(1-\bar{\alpha}_{t-1})}{1-\bar{\alpha}_t}x_t \quad \text{and} \quad \tilde{\Sigma}_t = \frac{1-\bar{\alpha}_{t-1}}{1-\bar{\alpha}_t}(1-\alpha_t). \quad (4)$$

DDPM – Generative Process

● Backward process

$$p(x_0, \dots, x_t) = \underbrace{p_0(x_0)}_{\text{data}} \cdot \underbrace{p_{T|0}(x_T|x_0)}_{\text{noise}} \cdot \prod_{t=2}^T \underbrace{p_{t-1|t,0}(x_{t-1}|x_t, x_0)}_{\text{Gaussian transitions}}, \quad (5)$$

where $p_{t-1|t,0}(\cdot|x_t, x_0) = \mathcal{N}(\cdot; \tilde{\mathbf{m}}_t(x_t, x_0), \tilde{\Sigma}_t)$.

- **Generative process** This suggests using the following structure for the generative model

$$p^\theta(x_0, \dots, x_t) = \underbrace{p_T^\theta(x_T)}_{\text{noise}} \cdot \prod_{t=1}^T \underbrace{p_{t-1|t}^\theta(x_{t-1}|x_t)}_{\text{Gaussian transitions}}, \quad (6)$$

with $p_{t-1|t}^\theta(\cdot|x_t) = \mathcal{N}(\cdot; \hat{\mathbf{m}}_t^\theta(x_t), \tilde{\Sigma}_t)$.

DDPM – Training Objective

- **Variational bound (ELBO)** We want to fit p^θ to p :

$$\begin{aligned}
 \log p_\theta(x_0) &= \log \left(\int p^\theta(X_{0:T}) dX_{1:T} \right) \\
 &\geq \log \left(\mathbb{E}_{p(X_{1:T}|x_0)} \frac{p^\theta(X_{0:T})}{p(X_{1:T}|X_0)} \right) \\
 &\geq \mathbb{E}_{p(X_{1:T})} \log \left(\frac{p^\theta(X_{0:T})}{p(X_{1:T}|X_0)} \right) \quad \text{By Jensen's ineq.} \\
 &= -\mathcal{L}_{\text{ELBO}}(\theta)
 \end{aligned}$$

Rearranging terms, we obtain

$$\mathcal{L}_{\text{ELBO}}(\theta) = \mathbb{E} \left[\underbrace{\text{KL}(p_{T|0}(\cdot|X_0) \parallel p_T^\theta(\cdot))}_{L_T} + \sum_{t=2}^T \underbrace{\text{KL}(p_{t-1|t,0}(\cdot|X_t, X_0) \parallel p_{t-1|t}^\theta(\cdot|X_t))}_{L_{t-1}} - \underbrace{\log p_{0|1}^\theta(X_0|X_1)}_{L_0} \right].$$

The terms L_T, L_0 are typically neglected.

DDPM – Training Objective

- **Analytical formula for L_{t-1}** KL between Gaussian distribution of equal variance $\tilde{\Sigma}_t$:

$$L_{t-1} = \frac{\|\tilde{\mathbf{m}}_t(X_t, X_0) - \hat{\mathbf{m}}_t^\theta(X_t)\|^2}{2\tilde{\Sigma}_t} .$$

- **ELBO loss**

$$\mathcal{L}(\theta) = \mathbb{E} \left[\frac{\|\tilde{\mathbf{m}}_t(X_t, X_0) - \hat{\mathbf{m}}_t^\theta(X_t)\|^2}{2\tilde{\Sigma}_t} \right] , \quad (7)$$

with a choice of time distribution ω (e.g., uniform, log-normal...).

- **Denoiser reparameterization** *We learn to predict (or remove) the noise added at each step.* Since

$$X_t \stackrel{d}{=} \sqrt{\bar{\alpha}_t} X_0 + \sqrt{1 - \bar{\alpha}_t} \bar{\epsilon}_t , \quad \bar{\epsilon}_t \sim \mathcal{N}(0, \mathbf{I}_d) , \quad (8)$$

we rewrite

$$\tilde{\mathbf{m}}_t(x_t, \bar{\epsilon}_t) = \frac{1}{\sqrt{\alpha_t}} \left(x_t - \frac{1 - \alpha_t}{\sqrt{1 - \bar{\alpha}_t}} \bar{\epsilon}_t \right) , \quad \hat{\mathbf{m}}_t^\theta(x_t) = \frac{1}{\sqrt{\alpha_t}} \left(x_t - \frac{1 - \alpha_t}{\sqrt{1 - \bar{\alpha}_t}} \hat{\epsilon}_t^\theta \right) . \quad (9)$$

Instead of optimizing the real ELBO, we optimize a simpler denoising loss

$$\mathcal{L}_{\text{simple}}(\theta) = \mathbb{E} [\|\bar{\epsilon}_t - \hat{\epsilon}_t^\theta(X_t)\|^2] . \quad (10)$$

DDPM – Recap

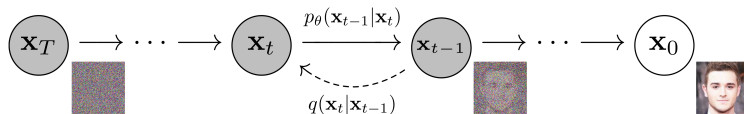


Figure: Forward/generative processes [HJA20]

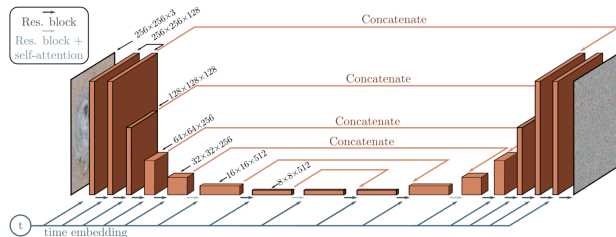


Figure: U-net architecture used for $\hat{\epsilon}_t^\theta$, predicting noise at each timestep [open_cv_ddpm]

Advantages

- High quality samples
- Stable/easy training (e.g., contrary to GANs)
- Equivalence between multiple approaches (continuous time with SDEs, flow matching etc.)

Disadvantages

- Lots of diffusion steps $T \gg 1$
- Mode collapse with high class imbalance
- What if initial data distribution is heavy tailed (no variance)?

Proposal – Change Noise Distribution

- Previous work:
 - Generalized Gaussian distributions ([DSL21])
 - Gamma distributions ([NRW21])
 - Lévy α -stable distribution ([Yoo+23])
- But show limitations:
 - No true time reversal, heuristics for sampling
 - Crude upper bound or unstable training
 - Hyper-parameters to tune
- We advocate for the α -stable **Lévy distributions**, which generalize Gaussian with heavy tails.
- Contrary to Lévy-Ito Models (LIM)([Yoo+23]), we employ a discrete time approach, which yields:
 - Distinct training and sampling equations
 - More stable training, with no **clipping hyper-parameters (!)** to tune
 - Improved performance

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Proposal - α -stable Heavy-tailed Distribution

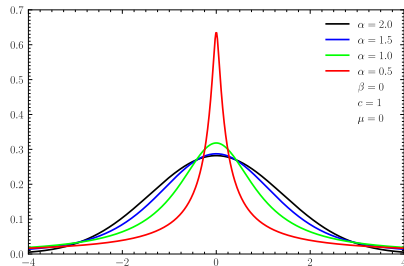
Explored solution: use heavy-tailed distributions for noising/denoising

- Better coverage of heavy-tailed data distribution
- Improvements on mode collapse especially in the context of class imbalance
- Less function evaluation

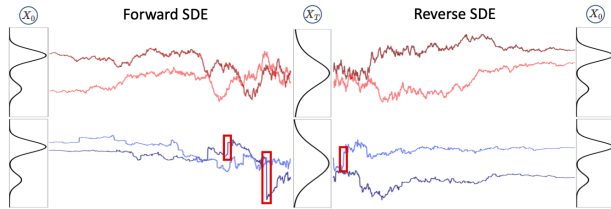
Large jumps benefit the exploration of the data space

α -stable Lévy distributions

α -stable Lévy distributions



(a) Symmetric α -Stable distribution, varying α [Wik24]



(b) Lévy Process vs Brownian Motion ($\alpha = 2$) [Yoo+23]

Definition and properties

The α -stable distributions $\mathcal{S}_{\alpha,\beta}(\mu, \sigma)$ are characterized by four parameters $(\alpha, \beta, \mu, \sigma)$:

- $\alpha \in (0, 2)$, the tail heaviness parameter
- $\beta \in (-1, 1)$, the skewness parameter
- μ , the location parameter
- σ , the scale parameter

This family of distributions is **stable** by addition, i.e.,

$$X_{\mathcal{S}_{\alpha,\beta_0}(\mu_0,\sigma_0)} + X_{\mathcal{S}_{\alpha,\beta_1}(\mu_1,\sigma_1)} \sim X_{\mathcal{S}_{\alpha,\beta}(\mu,\sigma)}$$

where

$$\sigma^\alpha = \sigma_0^\alpha + \sigma_1^\alpha, \quad \beta = \frac{\beta_0 \sigma_0^\alpha + \beta_1 \sigma_1^\alpha}{\sigma^\alpha}, \quad \mu = \mu_0 + \mu_1$$

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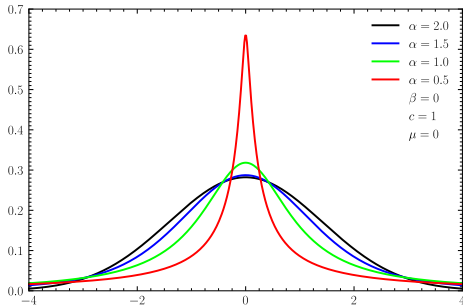
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$$\sigma^\alpha = \sigma_0^\alpha + \sigma_1^\alpha, \quad \beta = \frac{\beta_0 \sigma_0^\alpha + \beta_1 \sigma_1^\alpha}{\sigma^\alpha}, \quad \mu = \mu_0 + \mu_1$$

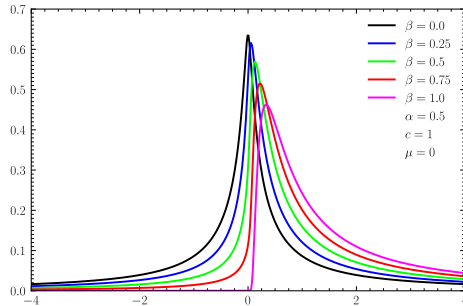
Notable special cases

- $(\beta = 0, \mu = 0)$ In this case, $\mathcal{S}_\alpha(0, \sigma)$ is symmetric and centered
- $(\alpha = 2)$ In this case, $\mathcal{S}_\alpha(0, \sigma)$ is the Gaussian distribution $\mathcal{N}(0, 2\sigma^2)$
- $(\alpha = 1)$ In this case, $\mathcal{S}_\alpha(0, 1)$ is the Cauchy distribution $\text{Cauchy}(0, 1)$

Definition and properties



(a) $\beta = 0, \mu = 0, \sigma = 1$, varying α [Wik24]



(b) $\alpha = 0.5, \mu = 0, \sigma = 1$, varying β [Wik24]

Gaussian Trick

Gaussian Trick

Let $A \sim \mathcal{S}_{\alpha/2,1}(0, c_A)$, and $G \sim \mathcal{N}(0, 1)$, where $c_A := \cos^{2/\alpha}(\pi\alpha/4)$. Then

$$A^{1/2}G \sim \mathcal{S}_\alpha(0, 1) . \quad (11)$$

- **Isotropic noise.** Draw $A \sim \mathcal{S}_{\alpha/2,1}(0, c_A)$, draw $G \sim \mathcal{N}(0, I_d)$, compute

$$A^{1/2} \cdot G . \quad (12)$$

- **Non-isotropic (independent) noise.** Draw $A = \{A_i\}_{i=1}^d$ i.i.d., draw $G \sim \mathcal{N}(0, I_d)$, compute

$$A^{1/2} \odot G . \quad (13)$$

Sampling an alpha-stable random variable

CMS algorithm (J.M. Chambers, C.L. Mallows and B.W. Stuck):

- Generate $U \sim \mathcal{U}([-\pi/2, \pi/2])$, and $W \sim \mathcal{E}(1)$.
- ($\alpha \neq 1$) Compute:

$$X = (1 + \zeta^2)^{1/2\alpha} \frac{\sin(\alpha(U + \xi))}{\cos(U)^{1/\alpha}} \left(\frac{\cos(U - \alpha(U + \xi))}{W} \right)^{(1-\alpha)/\alpha} \quad (14)$$

- ($\alpha = 1$) Compute:

$$X = \frac{1}{\xi} \left[\left(\frac{\pi}{2} + \beta U \right) \tan(U) - \beta \log \left(\frac{W \cos(u) \pi/2}{\zeta U + \pi/2} \right) \right] \quad (15)$$

with

$$\zeta = -\beta \tan \frac{\pi\alpha}{2}, \quad \xi = \begin{cases} \frac{1}{\alpha} \arctan(-\zeta) & \alpha \neq 1 \\ \frac{\pi}{2} & \alpha = 1 \end{cases} \quad (16)$$

- Then, $X \sim \mathcal{S}_{\alpha,\beta}(0, 1)$

When $\alpha = 2, \beta = 0$, this is the Box-Muller algorithm.

Different multidimensional heavy-tailed distributions

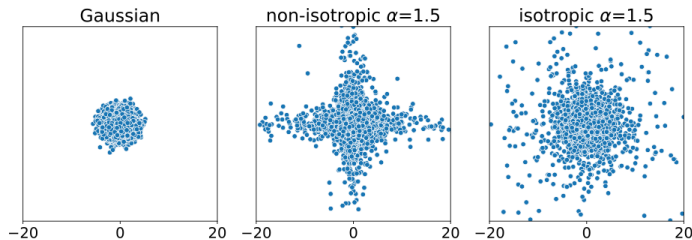


Figure: Different multidimensional heavy-tailed noise distributions, Gaussian vs $\alpha = 1.5$ [Yoo+23]

DLPM: Heavy-Tailed Denoising Diffusion

Forward Process - first approach

- **Forward process (Markov chain)** Consider $\{X_t\}_{t=0}^T$ defined by:

$$X_0 \sim p_0, \quad X_t = \gamma_t X_{t-1} + \sigma_t \epsilon_t^{(\alpha)}, \quad (17)$$

where $\{\epsilon_t^{(\alpha)}\}_{t=1}^T \sim \mathcal{S}_\alpha^i(0, I_d)^{\otimes T}$, and $\{(\gamma_t, \sigma_t)\}_{t=1}^T$ is the noising schedule.

- **Closed form for $X_t|X_0$**

$$X_t \stackrel{d}{=} \gamma_{1 \rightarrow t} X_0 + \sigma_{1 \rightarrow t} \epsilon_t^{(\alpha)}, \quad (18)$$

where $\epsilon_t^{(\alpha)} \sim \mathcal{S}_\alpha^i(0, I_d)$, and

$$\gamma_{1 \rightarrow t} = \prod_{i=1}^t \gamma_i, \quad \sigma_{1 \rightarrow t} = \left(\sum_{i=1}^t \left(\frac{\gamma_{1 \rightarrow t}}{\gamma_{1 \rightarrow i}} \sigma_i \right)^\alpha \right)^{1/\alpha}. \quad (19)$$

- **Variance Preserving (VP) schedule** Choose $0 < \gamma_t < 1$, $\sigma_t = (1 - \gamma_t^\alpha)^{1/\alpha}$. Then

$$\sigma_{1 \rightarrow t} = (1 - \gamma_{1 \rightarrow t}^\alpha)^{1/\alpha}.$$

- **Variance Exploding (VE) schedule** Choose $\gamma_t = 1$, $\sigma_t \nearrow_t \infty$

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Backward Process - first approach

- We want a similar structure for the generative process:

$$p_{0:T}^{\theta}(x_{0:T}) = p_T^{\theta}(x_T) \prod_{t=T}^1 p_{t-1|t}^{\theta}(x_{t-1}|x_t) , \quad (20)$$

- **Problem** No known techniques to characterize

$$p_{t-1|t}(x_{t-1}|x_t) , \quad p_{t-1|t,0}(x_{t-1}|x_t, x_0) . \quad (21)$$

Moreover, the KL between α -stable distributions is unavailable when $\alpha \neq 2, 1$.

- **How to design the backward process, the generative process, and the training procedure?**
- **Our approach** Data augmentation and the Gaussian trick

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where $\{A_t\}_{t=1}^T \sim \mathcal{S}_{\alpha/2,1}(0, c_A)^{\otimes T}$ and $\{G_t\}_{t=1}^T \sim \mathcal{N}(0, I_d)^{\otimes T}$. This process satisfies

$$Y_t \stackrel{d}{=} X_t. \quad (23)$$

- **Closed form for $Y_t \mid Y_0, A_{1:t}$**

$$Y_t \mid Y_0, A_{1:t} \stackrel{d}{=} \gamma_{1 \rightarrow t} Y_0 + \sum_{1 \rightarrow t} (A_{1:t})^{1/2} \bar{G}_t, \quad (24)$$

where $\bar{G}_t \sim \mathcal{N}(0, I_d)$, and

$$\gamma_{1 \rightarrow t} = \prod_{k=1}^t \gamma_k, \quad \sum_{1 \rightarrow t} (A_{1:t}) = \sum_{k=1}^t \left(\frac{\gamma_{1 \rightarrow t}}{\gamma_{1 \rightarrow k}} \sqrt{A_k} \sigma_k \right)^2. \quad (25)$$

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$$Y_t \mid Y_0, A_{1:t} \stackrel{d}{=} \gamma_{1 \rightarrow t} Y_0 + \Sigma_{1 \rightarrow t}(A_{1:t})^{1/2} \bar{G}_t, \quad (24)$$

where $\bar{G}_t \sim \mathcal{N}(0, I_d)$, and

$$\gamma_{1 \rightarrow t} = \prod_{k=1}^T \gamma_k, \quad \Sigma_{1 \rightarrow t}(A_{1:t}) = \sum_{k=1}^t \left(\frac{\gamma_{1 \rightarrow t}}{\gamma_{1 \rightarrow k}} \sqrt{A_k} \sigma_k \right)^2. \quad (25)$$

Backward Process – Data Augmentation Approach

- **Conditioning on $\{A_t\}_{t=1}^T$** The joint distribution admits the decomposition

$$\begin{aligned}
 p(x_0, \dots, x_T, a_{1:T}) &= p_0(x_0) \cdot \prod_{t=1}^T p_{t|t-1}(x_t | x_{t-1}, a_{1:T}) \psi_{(\alpha)}^{\otimes T}(a_{1:T}) \\
 &= \underbrace{p_0(x_0)}_{\text{data}} \cdot \underbrace{p_{T|0}(x_T | x_0, a_{1:T})}_{\text{noise}} \cdot \prod_{t=2}^T \underbrace{p_{t-1|t,0}(x_{t-1} | x_t, x_0, a_{1:T})}_{\text{Gaussian transitions}} \psi_{(\alpha)}^{\otimes T}(a_{1:T}),
 \end{aligned}$$

where $\psi_{(\alpha)}$ is the density of $\mathcal{S}_{\alpha/2,1}(0, c_A)$.

- **Gaussian transitions** $p_{t-1|t,0,a_{1:T}}(\cdot | x_t, x_0, a_{1:T})$ is the density of $\mathcal{N}(\tilde{m}_t(x_t, x_0, a_{1:t}), \tilde{\Sigma}_t(a_{1:t}))$.

Backward process - data augmentation approach

Gaussian transitions $p_{t-1|t,0,a_{1:T}}(\cdot|x_t, x_0, a_{1:T})$ is the density of $\mathcal{N}(\tilde{m}_t(x_t, x_0, a_{1:t}), \tilde{\Sigma}_t(a_{1:t}))$, where

$$\begin{aligned}\tilde{m}_{t-1}(y_t, y_0, a_{1:t}) &= \frac{1}{\gamma_t} (y_t - \Gamma_t(a_{1:t})\sigma_{1 \rightarrow t}\epsilon_t(y_t, y_0)) , \\ \tilde{\Sigma}_{t-1}(a_{1:t}) &= \Gamma_t(a_{1:t})\Sigma_{1 \rightarrow t-1}(a_{1:t-1}) ,\end{aligned}\tag{26}$$

with

$$\underbrace{\epsilon_t(y_t, y_0) = \frac{y_t - \gamma_{1 \rightarrow t}y_0}{\sigma_{1 \rightarrow t}}}_{\text{noise}}, \quad \underbrace{\Sigma_{1 \rightarrow t}(a_{1:t}) = \sum_{k=1}^t \left(\frac{\gamma_{1 \rightarrow t}}{\gamma_{1 \rightarrow k}} \sqrt{a_k} \sigma_k \right)^2}_{\text{variance}}, \quad \underbrace{\Gamma_t(a_{1:t}) = 1 - \frac{\gamma_t^2 \Sigma_{1 \rightarrow t-1}(a_{1:t-1})}{\Sigma_{1 \rightarrow t}(a_{1:t})}}_{\text{stochastic scaling}} .\tag{27}$$

Note that Γ_t is bounded: $0 \leq \Gamma_t \leq 1$.

Backward process - model

Generative process

$$p^\theta(x_0, \dots, x_t, a_{1:T}) = \underbrace{p_T^\theta(x_T)}_{\text{noise}} \cdot \prod_{t=1}^T \underbrace{p_{t-1|t,a}^\theta(x_{t-1}|x_t, a_{1:t})}_{\text{Gaussian transitions}} \psi_{(\alpha)}^{\otimes T}(a_{1:T}) , \quad (28)$$

where $\psi_{(\alpha)}$ is the density of the $\mathcal{S}_{\alpha/2,1}(0, c_A)$ distribution, and

$$p_{t-1|t,a}^\theta(\cdot|x_t, a_{1:t}) = \mathcal{N}(\cdot ; \hat{\mathbf{m}}_t^\theta(x_t, a_{1:t}), \tilde{\Sigma}_t(a_{1:t})) . \quad (29)$$

Loss function - alpha-stable case

Reminder: ELBO loss, Gaussian case

$$\mathcal{L}(\theta) = \mathbb{E} \left[\frac{\|\tilde{\mathbf{m}}_t(X_t, X_0) - \hat{\mathbf{m}}_t^\theta(X_t)\|^2}{2\tilde{\Sigma}_t} \right], \quad (30)$$

- **A naive solution:** by Jensen's inequality:

$$\text{KL}(p_0 \| p_0^\theta) \leq \mathbb{E} \left(\text{KL} [p_0(\cdot) \| p_{0|a}^\theta(\cdot | A_{1:T})] \right). \quad (31)$$

- As we see in (30), this expression would involve taking expectation of A_t
- However, A_t is distributed as $\mathcal{S}_{\alpha/2,1}(0, c_A)$, and does not admit a first order moment.

Loss function - alpha-stable case

- **Loss function** We aim to minimize the following KL divergence:

$$\mathcal{L}(\theta) := \mathbb{E} \left[\text{KL}(p_{0|a}(\cdot | A_{1:T}) \| p_{0|a}^{\theta}(\cdot | A_{1:T}))^{1/2} \right] . \quad (32)$$

To obtain our loss, we employ the usual derivations:

$$\begin{aligned} \mathcal{L}(\theta) &\leq \mathbb{E} \left[L_T(\theta, A_{1:T}) + \sum_{t \geq 2} L_{t-1}(\theta, A_{1:T}) + L_0(\theta, A_{1:T}) \right]^{1/2} && \text{(ELBO)} , \\ &\leq \mathbb{E} \left[L_T(\theta, A_{1:T})^{1/2} + \sum_{t \geq 2} L_{t-1}(\theta, A_{1:T})^{1/2} + L_0(\theta, A_{1:T})^{1/2} \right] && (\sqrt{a+b} < \sqrt{a} + \sqrt{b}) . \end{aligned} \quad (33)$$

Again, we neglect L_T, L_0 . Since $L_{t-1}(\theta, A_{1:T}) = \text{KL}(p_{t-1|t,0,a}(\cdot | Y_t, Y_0, A_{1:T}) \| p_{t-1|t,a}^{\theta}(\cdot | Y_t, A_{1:T}))$:

$$\mathcal{L}^L(\theta) = \mathbb{E} \left[\mathbb{E} \left[\frac{1}{2\hat{\Sigma}_{t-1}^{\theta}(A_{1:t})} \|\tilde{m}_{t-1}(Y_t, Y_0, A_{1:t}) - \hat{m}_{t-1}^{\theta}(Y_t, A_{1:t})\|^2 \mid A_{1:t} \right]^{1/2} \right] . \quad (34)$$

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Loss function - design choices **D1**

- **D1 (Fixed variance)** We set $\hat{\Sigma}_t^\theta = \tilde{\Sigma}_t$.

Loss function - design choice **D2**

- **D2 (Denoiser Reparameterization)** We predict the *injected noise* $\epsilon_t(y_t, y_0)$ rather than $\tilde{m}_{t-1}(y_t, y_0, a_{1:t})$. Since

$$\tilde{m}_{t-1}(Y_t, Y_0, A_{1:t}) = \frac{1}{\gamma_t} (Y_t - \sigma_{1 \rightarrow t} \Gamma_t(A_{1:t}) \epsilon_t(Y_t, Y_0)) , \quad (35)$$

we re-parameterize \hat{m}_{t-1}^θ as

$$\hat{m}_{t-1}^\theta(Y_t, A_{1:t}) = \frac{1}{\gamma_t} \left(Y_t - \sigma_{1 \rightarrow t} \Gamma_t(A_{1:t}) \hat{\epsilon}_t^\theta(Y_t) \right) . \quad (36)$$

with $\hat{\epsilon}_t^\theta$ the output of the model.

- The model $\hat{\epsilon}_t^\theta$ does not take any heavy-tailed $A_{1:t}$ as input.
- Assuming **D1**, the loss \mathcal{L}^L becomes

$$\mathcal{L}^L(\theta) = \mathbb{E} \left[\lambda_{t, A_{1:t}}^2 \|\hat{\epsilon}_t^\theta(Y_t, A_{1:t}) - \epsilon_t(Y_t, Y_0)\|^2 \right] , \quad (37)$$

$$\lambda_{t, a_{1:t}} = \frac{\Gamma_t(a_{1:t}) \sigma_{1 \rightarrow t}}{2 \gamma_t \tilde{\Sigma}_{t-1}} , \quad \epsilon_t(Y_t, Y_0) = \frac{(Y_t - \gamma_{1 \rightarrow t} Y_0)}{\sigma_{1 \rightarrow t}} . \quad (38)$$

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Loss function - design choice **D3**

- **D3 (Simple loss)** With design choices **D1**, **D2**, the loss \mathcal{L}^L is

$$\mathcal{L}^L(\theta) = \mathbb{E} \left[\lambda_{t, A_{1:t}}^2 \|\hat{\epsilon}_t^\theta(Y_t, A_{1:t}) - \epsilon_t(Y_t, Y_0)\|^2 \right]. \quad (39)$$

We choose to set $\lambda_{t, A_{1:t}} = 1$, which improves performance, and draws similarities to the continuous α -stable score-based perspective.

We obtain a simplified denoising objective function

$$\mathcal{L}^{\text{Simple}}(\theta) = \mathbb{E} \left[\mathbb{E} \left(\|\hat{\epsilon}_t^\theta(Y_t) - \epsilon_t(Y_t, Y_0)\|^2 \mid A_{1:t} \right)^{1/2} \right]. \quad (40)$$

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Bonus - faster sampling

Assume design choices **D1**, **D2**, **D3** are satisfied. Then one can obtain the following simplified denoising objective function:

$$\mathcal{L}_{t-1}^{\text{SimpleLess}}(\theta) = \mathbb{E} \left[\mathbb{E} \left(\left\| \hat{\epsilon}_t^\theta(Y_t^{\text{Less}}) - \epsilon_t(Y_t^{\text{Less}}, Y_0^{\text{Less}}) \right\|^2 \mid \bar{A}_t \right) \right]^{1/2}, \quad t \in \{2, \dots, T\} \quad (41)$$

where

$$Y_t^{\text{Less}} = \gamma_{1 \rightarrow t} Y_0^{\text{Less}} + \sigma_{1 \rightarrow t} \bar{A}_t^{1/2} G_t, \quad \epsilon_t(Y_t^{\text{Less}}, Y_0^{\text{Less}}) = \frac{Y_t^{\text{Less}} - \gamma_{1 \rightarrow t} Y_0^{\text{Less}}}{\sigma_{1 \rightarrow t}}. \quad (42)$$

with $G_t \sim \mathcal{N}(0, I_d)$, $\bar{A}_t \sim \mathcal{S}_{\alpha/2, 1}(0, c_A)$.

- Idea: sufficient statistic, as

$$Y_t \stackrel{d}{=} \gamma_{1 \rightarrow t} Y_0 + \Sigma_{1 \rightarrow t} (A_{1:t})^{1/2} \bar{G}_t \stackrel{d}{=} \gamma_{1 \rightarrow t} Y_0 + \sigma_{1 \rightarrow t} \epsilon_t^{(\alpha)} \stackrel{d}{=} \gamma_{1 \rightarrow t} Y_0 + \sigma_{1 \rightarrow t} \bar{A}_t^{1/2} G_t \quad (43)$$

- Cheaper than sampling a list $A_{1:t}$ for each datapoint.
- The final denoising loss is similar to LIM (continuous α -stable case), but guaranteed to be finite.

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Denoising Lévy Implicit Models: Deterministic Generation

Deterministic Generation – Gaussian case (DDIM)

- **Directly define the bridges** For DDPM, we did not need the forward process to be Markovian, and only benefited from the following decomposition:

$$p(x_0, \dots, x_t) = \underbrace{p_0(x_0)}_{\text{data}} \cdot \underbrace{p_{T|0}(x_T|x_0)}_{\text{noise}} \cdot \prod_{t=2}^T \underbrace{p_{t-1|t,0}(x_{t-1}|x_t, x_0)}_{\text{Gaussian transitions}}. \quad (44)$$

- **Non-necessarily Markovian process** Sample endpoints first

$$\bar{X}_0 \sim p_0, \quad \bar{X}_T | \bar{X}_0 \sim \mathcal{N}(\sqrt{\bar{\alpha}_T} \bar{X}_0, (1 - \bar{\alpha}_T) \text{I}_d), \quad (45)$$

and then the bridges

$$\bar{X}_{t-1} = \sqrt{\bar{\alpha}_{t-1}} \bar{X}_0 + \sqrt{1 - \bar{\alpha}_{t-1} - \sigma_t^2} \cdot \underbrace{\frac{\bar{X}_t - \sqrt{\bar{\alpha}_t} \bar{X}_0}{\sqrt{1 - \bar{\alpha}_t}}}_{\text{injected noise between 0 and } t} + \underbrace{\sigma_t \epsilon_t}_{\text{stochasticity}}, \quad (46)$$

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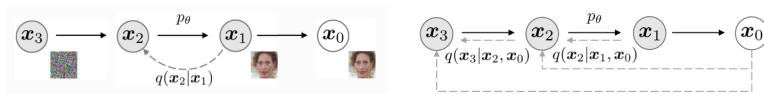


Figure: Non-Markovian forward process [SME20]

Distribution of $X_t|X_0$ Same as DDPM. Informal proof:

$$\bar{X}_{t-1} = \sqrt{\bar{\alpha}_{t-1}}\bar{X}_0 + \sqrt{1 - \bar{\alpha}_{t-1} - \sigma_t^2} \cdot \underbrace{\frac{\bar{X}_t - \sqrt{\bar{\alpha}_t}\bar{X}_0}{\sqrt{1 - \bar{\alpha}_t}}}_{\equiv \text{injected noise between 0 and } t} + \sigma_t \epsilon_t$$

$$\stackrel{d}{=} \sqrt{\bar{\alpha}_{t-1}}\bar{X}_0 + \sqrt{1 - \bar{\alpha}_{t-1} - \sigma_t^2 + \sigma_t^2}\hat{\epsilon}_t, \quad \hat{\epsilon}_t \sim \mathcal{N}(0, I_d) \quad (\text{Stability of Gaussian})$$

$$\stackrel{d}{=} \sqrt{\bar{\alpha}_{t-1}}\bar{X}_0 + \sqrt{1 - \bar{\alpha}_{t-1}}\hat{\epsilon}_t.$$

DLIM - Denoising Lévy Implicit Models

- **Forward process** $\{Z_t\}_{t=0}^T$ is such that:

$$Z_0 \sim p_0, \quad Z_T \sim \mathcal{S}_\alpha(\gamma_{1 \rightarrow T} Z_0, \sigma_{1 \rightarrow T} \mathbf{I}_d), \quad \text{and} \quad (47)$$

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with $\{G_t\}_{t=1}^T \sim \mathcal{N}(0, \mathbf{I}_d)^{\otimes T}$, $\{A_t\}_{t=1}^T \sim \mathcal{S}_{\alpha/2, 1}(0, c_A)^{\otimes T}$.

- Closed form expression for $p_{t|0} Z_t | Z_0 \stackrel{d}{=} Y_t | Y_0$ for $t \in \{1, \dots, T\}$.

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- **Closed form expression for** $p_{t|0} Z_t | Z_0 \stackrel{d}{=} Y_t | Y_0$ for $t \in \{1, \dots, T\}$.

DLIM - Denoising Lévy Implicit Models

- **Always recovers DLPM loss** We derive the same KL loss with the same techniques; re-use $\hat{\epsilon}_t^\theta(Z_t)$ trained for DLPM
- **Possibly better loss** Since we directly specify $p_{t-1|t,0}$, we can bypass the need for $A_{1:T}$ if a closed-form KL exists between $\mathcal{S}(\mu_1, \sigma_1)$ and $\mathcal{S}(\mu_2, \sigma_2)$; it is the case for Cauchy ($\alpha = 1$).
- **Deterministic generation** We obtain a deterministic sampling process, with the same techniques as in DDIM, as $\rho \rightarrow 0$.

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Lévy-Itô Models (LIM)

Lévy-Itô Models (LIM) vs DLPM

- **LIM** Levy-Itô Models are the continuous time version of α -stable generative models. They **extend the SDE formulation to Levy processes**.
- **LIM vs DLPM**
 - DLPM has much simpler and accessible theory, without any need for **complicated fractional stochastic calculus**
 - DLPM leverages the flexibility of the discrete formulation for diffusion. Example: possibility to learn variance.
 - Both approaches yield different training losses and sampling procedures



Figure: Illustration of available methods.

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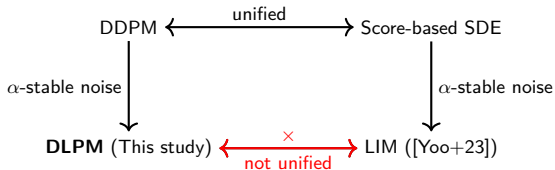


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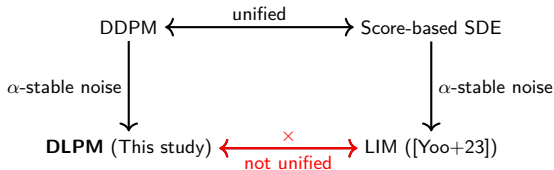


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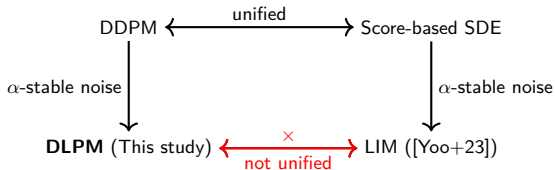


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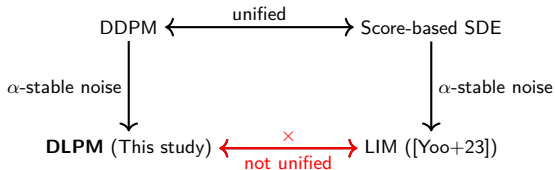


Figure: Illustration of available methods.

LIM - forward

- **Forward process** The forward process $\{X_t\}_{0 \leq t \leq T}$, with $X_0 \sim p_0$, is obtained with

$$dX_t = \gamma(t, X_{t-})dt + \sigma(t)dL_t^\alpha, \quad (49)$$

where X_{t-} denotes the left limit of X at time t . **LIM only defines scale-preserving schedule:**

$$\gamma(t, x) = -\frac{\beta_t}{\alpha}x, \quad \sigma(t) = \beta_t^{1/\alpha}. \quad (50)$$

- **Closed-form expression of $X_t|X_0$**

$$X_t \stackrel{d}{=} \gamma_{1 \rightarrow t} X_0 + \sigma_{1 \rightarrow t} \bar{\epsilon}, \quad (51)$$

where $\bar{\epsilon}_t \sim \mathcal{S}_\alpha^i(0, I_d)$. The values of $\gamma_{1 \rightarrow t}$ and $\sigma_{1 \rightarrow t}$ match with the DLPM definition on integer timesteps.

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LIM - backward

- **Backward process** The following backward process \bar{X}_t is obtained:

$$d\bar{X}_t = (-\gamma(t, \bar{X}_{t+}) + \alpha\sigma^\alpha(t, \bar{X}_{t+})s_t(\bar{X}_{t+})) dt + \sigma(t)d\bar{L}^\alpha_t + d\bar{Z}_t \quad (52)$$

where

- \bar{Z}_t is the backward version of a Lévy-type stochastic integral Z_t s.t $\mathbb{E}[Z_t] = 0$ with finite variation
- s_t is the fractional score function:

$$s_t(x) = \frac{\Delta^{\frac{\alpha-2}{2}} \nabla p_t(x)}{p_t(x)}, \quad (53)$$

where $\Delta^{\eta/2}$ is the fractional Laplacian of order $\eta/2$, defined with Fourier transform \mathcal{F} :

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LIM - training

- The true score $s_t(x_t|x_0)$ can be expressed as

$$s_t(x_t|x_0) = -\frac{1}{\alpha\sigma_{1\rightarrow t}^{\alpha-1}(t)}\epsilon_t(x_t, x_0), \quad (55)$$

where $\epsilon_t(x_t, x_0) = \frac{x_t - \gamma_{1\rightarrow t}x_0}{\sigma_{1\rightarrow t}}$, thus we re-parametrize

$$s_\theta(x_t, t) = -\frac{1}{\alpha\sigma_{1\rightarrow t}^{\alpha-1}(t)}\hat{\epsilon}_t^\theta(x_t, x_0), \quad (56)$$

so that we rather work with $\hat{\epsilon}_t^\theta$.

- Training loss obtained using denoising score matching technique:

$$L : \theta \mapsto \mathbb{E}\|s_\theta(X_t, t) - s_t(X_t)\|^2, \quad L' : \theta \mapsto \mathbb{E}\|s_\theta(X_t, t) - s_t(X_t|x_0)\|^2, \quad (57)$$

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LIM vs DLPM - forward/backward

With $\{G'_t\}_{t=T}^1$ i.i.d. $\mathcal{N}(0, I_d)$, $\{\epsilon'_t\}_{t=T}^1$ i.i.d. $\mathcal{S}_\alpha^i(0, I_d)$, and $\hat{\epsilon}_t^\theta$ the model at time t :

	Stochastic	Deterministic
Continuous (LIM)	$\frac{\bar{X}_t^\theta}{\gamma_t} - \frac{\alpha(1/\gamma_t - 1)}{\sigma_{1 \rightarrow t}^{\alpha-1}} \hat{\epsilon}_t^\theta + \left(\frac{1}{\gamma_t^\alpha} - 1\right)^{1/\alpha} \epsilon'_t$	$\frac{\bar{X}_t^\theta}{\gamma_t} - \left(\frac{\sigma_{1 \rightarrow t}^{1-\alpha}}{\gamma_t} - \sigma_{1 \rightarrow t}^{1-\alpha}\right) \hat{\epsilon}_t^\theta$
Denoising (DLPM)	$\frac{\bar{Y}_t^\theta}{\gamma_t} - \Gamma_t \sigma_{1 \rightarrow t} \hat{\epsilon}_t^\theta + \Gamma_t \Sigma_{1 \rightarrow t-1} G'_t$	$\frac{\bar{Y}_t^\theta}{\gamma_t} - \left(\frac{\sigma_{1 \rightarrow t}}{\gamma_t} - \sigma_{1 \rightarrow t-1}\right) \hat{\epsilon}_t^\theta$

- **Stochastic sampling** Different sampling procedures. Moreover:
 - When $\alpha = 2$, $0 \leq \Gamma_t \leq 1$ becomes deterministic, and one recovers DDPM formulas
 - Γ_t brings additional stochasticity
 - Γ_t scales (i) the noise added at time $t - 1$ (ii) the output of the noise model.
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LIM vs DLPM - training

- Alike the Gaussian case ($\alpha = 2$), the score $s_t(x_t|x_0)$ is a linear expression of the noise term:

$$s_t(x_t|x_0) = -\frac{1}{\alpha\sigma_{1\rightarrow t}^{\alpha-1}(t)}\epsilon_t(x_t, x_0) , \quad (58)$$

leading to a similar denoising loss:

$$\mathcal{L}_{t-1} : \theta \mapsto \mathbb{E} \left(\|\hat{\epsilon}_t^\theta(X_t) - \epsilon_t(X_t, X_0)\|_{\substack{\eta \\ p}} \right) . \quad (59)$$

- DLPM: use $p = 2$ and $\eta = 1$.
- LIM (theory): use $p = 2$ and $\eta = 2$, for denoising score matching loss equivalence. But $\epsilon_t(X_t, X_0)$ is heavy-tailed: no variance!
- LIM (experiments): use $p = 1$ and $\eta = 1$. Indicates potential shortcoming of the theoretical approach.

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Experiments

Experiments – Setup

- The loss function

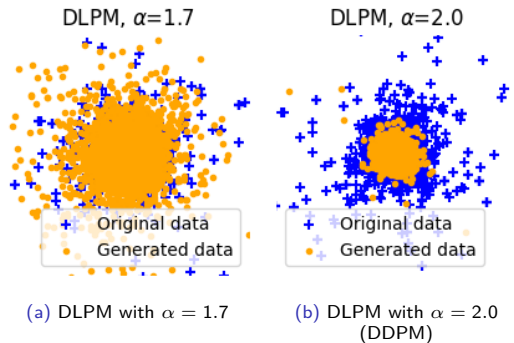
$$\mathcal{L}^{\text{SimpleLess}}(\theta) = \sum_{t=1}^T \mathbb{E} \left[\mathbb{E} \left(\|\hat{\epsilon}_t^\theta(Y_t) - \epsilon_t(Y_t, Y_0)\|^2 \mid \bar{A}_t \right)^{1/2} \right] \quad (60)$$

involves an expectation with respect to \bar{A}_t . We propose the *median-of-means* estimator ([LM19]), denoted by DLPM₅ ($M = 5$).

- We experiment with non-isotropic diffusion DLPMⁿⁱ
- We consider the range $1.5 \leq \alpha \leq 2.0$
- We use CIFAR10_LT (long tail), unbalanced modification of the CIFAR10 ([Yoo+23]).
 - Class count: [5000, 2997, 1796, 1077, 645, 387, 232, 139, 83, 50].

2D data - covering the dataset and capturing heavy-tails

- **Dataset** 20000 samples of $\mathcal{S}_\alpha^i(0, 0.05 \cdot I_2)$, with $\alpha = 1.7$.
- **Main challenge:** cover the dataset and correctly capture the tails.



- The lighter tailed process fails to capture the distribution's tail.

2D data - covering the dataset and capturing heavy-tails

- Drawing inspiration from [AGG22], we define the MSLE:

$$\text{MSLE}(\xi) = \int_{\xi}^1 \left(\log \hat{F}^{-1}(p) - \log \hat{F}^{\theta^{-1}}(p) \right)^2 dp, \quad (61)$$

where $\hat{F}, \hat{F}^{\theta}$ denote respectively the cdf of the true data and the generated data.

Method	$\alpha = 1.5$	$\alpha = 1.6$	$\alpha = 1.7$	$\alpha = 1.8$	$\alpha = 1.9$	$\alpha = 2.0$
DLPM	0.160 \pm 0.128	0.081 \pm 0.078	0.071 \pm 0.028	0.099 \pm 0.044	0.132 \pm 0.101	0.798 \pm 0.601
DDPM	-	-	-	-	-	0.528 \pm 0.400
						1.0e-1
LIM	0.743 \pm 0.290	0.497 \pm 0.311	0.267 \pm 0.077	0.653 \pm 0.413	2.444 \pm 1.067	1.239 \pm 0.240
	1.0e-08	8.6e-06	1.3e-10	8.8e-06	7.9e-09	5.0e-3

Table: $\text{MSLE}_{\xi=0.95} \downarrow$ averaged over 20 runs. Figures below scores corresponds to p -values from Welch's t -test (assuming unequal variances), comparing the mean of DLPM with the given method.

2D data - managing class imbalance

- **Dataset** Mixture of nine Gaussian distributions arranged in a grid

$$\sum_{i=1}^9 w_i \mathcal{N}(\mu_i, 0.05^2 \cdot \mathbf{I}_2) . \quad (62)$$

Mixture weights range from .01 to .3: { .01, .02, .02, .05, .05, .1, .1, .15, .2, .3 }.

- **Main challenge:** correctly guess the mixture weights

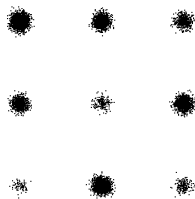


Figure: Gaussian grid

Method	$\alpha = 1.5$	$\alpha = 1.6$	$\alpha = 1.7$	$\alpha = 1.8$	$\alpha = 1.9$	$\alpha = 2.0$
DLPM	0.933 ± 0.018	0.923 ± 0.005	0.933 ± 0.028	0.923 ± 0.024	0.907 ± 0.034	0.862 ± 0.028
DLPM _s	0.944 ± 0.013	0.943 ± 0.021	0.943 ± 0.010	0.941 ± 0.014	0.928 ± 0.016	-
	<i>9.0e-3</i>	<i>1.6e-05</i>	<i>7.4e-2</i>	<i>9.0e-4</i>	<i>3.9e-3</i>	
LIM	0.842 ± 0.039	0.850 ± 0.046	0.868 ± 0.034	0.874 ± 0.030	0.884 ± 0.017	0.874 ± 0.027
	<i>1.7e-14</i>	<i>1.3e-09</i>	<i>5.7e-11</i>	<i>3.9e-09</i>	<i>1.9e-3</i>	<i>9.6e-2</i>
DDPM	-	-	-	-	-	0.867 ± 0.029
						<i>5.0e-1</i>

Table: $F_1^{\text{pr}} \uparrow$ score, averaged over 30 runs. Figures below scores corresponds to p -values from Welch's t -test (assuming unequal variances), comparing the mean of DLPM with the given method.

2D data - faster convergence

- DLIM vs LIM-ODE with varying total diffusion steps T , on the Gaussian grid.
- **Main challenge: get to the data distribution with the smallest T possible**

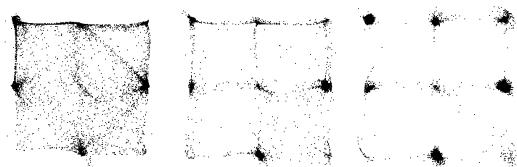


Figure: DLIM with $T = 5, 10, 25$ diffusion steps on the Gaussian grid

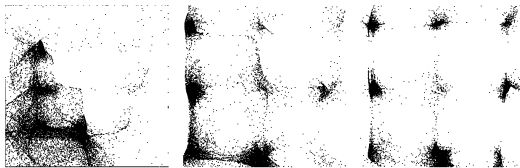


Figure: LIM-ODE with $T = 5, 10, 25$ diffusion steps on the Gaussian grid

Image data - LIM vs DLPM

- **Dataset** MNIST and CIFAR10_LT.
- Convergence speed for the different methods, varying total number of diffusion steps T .

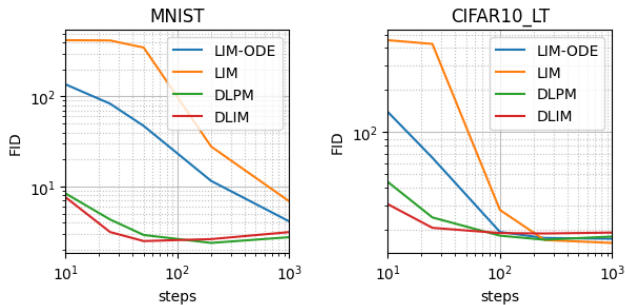


Figure: FID↓ with varying step size, $\alpha = 1.7$

Image data - LIM vs DLPM

MNIST	$\alpha = 1.5$	$\alpha = 1.7$	$\alpha = 1.8$	$\alpha = 1.9$	$\alpha = 2.0$
DDPM	-	-	-	-	3.43
LIM	14.37	11.54	11.18	13.75	11.69
<i>w/ clipping</i>	<i>4.08</i>	<i>5.17</i>	<i>6.81</i>	<i>11.20</i>	
DLPM ₅	3.80	3.03	2.51	2.71	-
DLPM	5.39	2.94	2.93	3.24	3.63
DDIM	-	-	-	-	5.16
LIM-ODE	49.63	78.59	92.93	109.48	29.04
<i>w/ clipping</i>	<i>45.72</i>	<i>68.15</i>	<i>85.09</i>	<i>113.20</i>	
DLIM ₅	3.37	2.93	3.44	4.31	-
DLIM	3.38	2.81	3.18	3.27	5.18
CIFAR10_LT					
DDPM	-	-	-	-	19.05
LIM	75.38	35.15	31.14	21.68	21.56
<i>w/ clipping</i>	<i>16.13</i>	<i>16.21</i>	<i>17.67</i>	<i>19.24</i>	
DLPM	16.10	18.00	19.94	20.21	21.07
DDIM	-	-	-	-	23.44
LIM-ODE	42.07	91.64	105.95	407.79	32.00
<i>w/ clipping</i>	<i>30.17</i>	<i>65.78</i>	<i>84.55</i>	<i>101.70</i>	
DLIM	20.69	20.77	21.96	22.79	23.99

- Better performance of DLPM as compared to LIM.
- Better performance with smaller α .

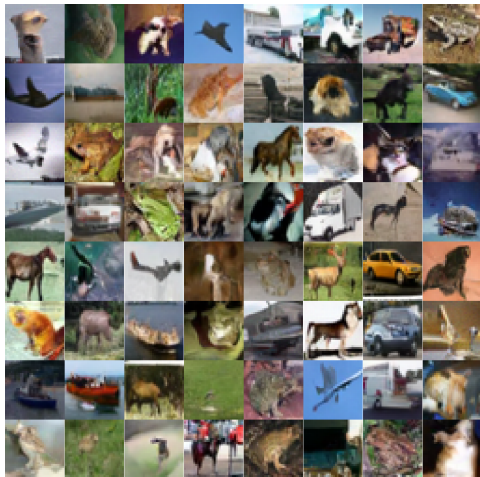
Reference I

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Some images - DLPM



(a) CIFAR10, $T = 4000$



(b) MNIST, $T = 1000$

Some images - DLIM



(a) CIFAR10, $T = 200$



(b) MNIST, $T = 50$