

# Job Transitions and Employee Earnings After Acquisitions: Linking Corporate and Worker Outcomes\*

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## Abstract

This paper connects changes in employer characteristics through job transitions to employee earnings following mergers and acquisitions. Using firm balance sheet data linked to individual earnings data in Canada and a matched difference-in-differences design, we find that earnings of workers at target firms decrease after M&As, largely driven by those who move to other firms. Workers leaving targets move to larger and more profitable firms, but experience wage declines potentially due to a loss of firm-specific human capital or backloaded contracts. It appears that losses of match-specific premiums from job transitions primarily explain the post-M&A earnings decline in our setting.

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# 1 Introduction

Mergers and acquisitions (M&As hereinafter) often result in important changes in corporate structure and worker composition for both target and acquiring firms. While there exists a large literature on how M&As affect firm performance and market value, there is growing interest in understanding labor market impacts of M&As.<sup>1</sup> In theory, M&As may have positive or negative impacts on worker earnings. If M&As lead to efficiency gains for the involved parties, incumbent workers at acquiring firms or target firms may experience a wage increase through rent sharing. By contrast, M&As may lead to layoffs through labor restructuring which could lead to earnings losses for displaced workers. Furthermore, M&As can result in increased labor-market or product-market power, again leading to lower wages. Since M&As may bring significant changes in both corporate and labor structure, a key challenge in answering this question is to link worker and firm outcomes following M&As to delve into potential mechanisms behind changes in worker earnings.

In this paper, we provide new empirical evidence on changes in workers' labor market outcomes following corporate M&As. While prior studies focus only on either firm-level outcomes or worker-level outcomes, in our work we are able to link detailed firm balance sheet data from corporate tax returns to worker-level earnings data from individual tax returns in Canada.<sup>2</sup> This allows us to not only assess employee-level outcomes, but also to study firm outcomes to connect and pin down potential mechanisms behind our findings. Specifically, the detailed joint information on firms and workers allows us to study the connection between changes in employer characteristics via job transitions and employee earnings after M&As.

We first examine the impacts of M&As on employment and average payrolls at the firm level. To do so, we implement a matched difference-in-differences design, where we match firms that go through M&A events to otherwise similar firms that never undergo M&As during our sample period. Importantly, in our setting, around 80 percent of M&A events involve acquisitions, where the target sells its assets, divisions, or only a part of its share, and remains operating as a separate entity following the M&A event. This allows us to keep track of the target's and the acquirer's outcomes after the event, so that we can assess the impacts on targets and on acquirers separately.<sup>3</sup>

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<sup>1</sup>There exists a large body of literature that examines the M&A impacts on firm performance, including [Schoar \(2002\)](#), [Hoberg and Phillips \(2010\)](#), [Boucly et al. \(2011\)](#), [Davis et al. \(2014\)](#), [Braguinsky et al. \(2015\)](#), [Blonigen and Pierce \(2016\)](#), [Malmendier et al. \(2018\)](#), and [Demirer and Karaduman \(2024\)](#). A growing body of studies assesses the impacts on workers, including [Brown and Medoff \(1988\)](#), [Siegel and Simons \(2010\)](#), [Dessaint et al. \(2017\)](#), [Ouimet and Zarutskie \(2020\)](#), [Prager and Schmitt \(2021\)](#), [Arnold \(2021\)](#), [Lagaras \(forthcoming\)](#), and [He and le Maire \(2022\)](#).

<sup>2</sup>While [Arnold \(2021\)](#) and [Lagaras \(forthcoming\)](#) also use employer-employee matched data in the U.S. and in Brazil, respectively, their datasets do not have detailed firm balance sheet information, such as profit margins, and therefore cannot be used to simultaneously evaluate firm-level and worker-level outcomes.

<sup>3</sup>As we discuss in detail later, our worker-level results depend substantially on whether the worker was originally at the target or the acquiring firm.

By contrast, in mergers, the entirety of the target is sold to (or merged with) the acquirer, meaning there is a definite change of control. Given that there can be qualitative or quantitative differences in firm-level impacts depending on whether there is a change of control, we separate our firm-level analyses into those involved in acquisitions and those involved in mergers, although we do not find meaningfully different results on firm-level outcomes in the aggregate between the two types.<sup>4</sup> Given that the vast majority of M&A activities in Canada (and in the United States) involve acquisitions, analyzing these activities in addition to mergers allows us to exploit more events and to examine a larger and broader sample of relevant firms and workers.

Focusing on these acquisitions, we find that acquirers expand, while targets shrink substantially after M&As. Acquirers' employment increases by 18.8 log points, without much change in average payrolls, on average relative to their matched control firms after M&As. By contrast, targets' employment and average payrolls decrease by 8.9 log points and 2.8 log points on average, respectively. At the aggregate (pooling targets and acquirers) level, however, we find almost no changes in employment or average payrolls in both mergers and acquisitions. These null results on firm size in the aggregate could be driven by reallocation of workers between targets and acquirers, so we use worker-level data to assess changes in their earnings and job transitions after M&As.

Since the changes in targets' average payrolls could be driven by changes in worker composition, the next part of our paper estimates the impact of M&As on worker-level earnings. Here, we do not make a distinction between acquisitions and mergers, not only because we can track individual worker transitions, but also because we find similar results in both cases (Appendix B). For workers at acquiring firms, we find relatively stable earnings after M&As. However, for target workers, we find their earnings decrease by 1.2 percent. This decline in worker earnings may stem from two channels. First, workers at target firms may face an increased risk of job displacement, implying the impacts on earnings are driven by job transitions. Second, workers at target firms may experience earnings losses due to changes in target firms' wage setting.

To test this, we first estimate the impact of M&As on job transitions. For target workers, we find there is a large increase in the probability of job transitions in the year following an M&A event. Given that we observe the universe of workers in the formal labor market in Canada, we check that most of these post-M&A job transitions are to different companies within the same industry, rather than to self-employment or unemployment. By contrast, for workers at acquiring firms, we find no change in the probability of job transitions.

For individuals who remain at target firms, we observe gradual and medium-run declines in

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<sup>4</sup>As discussed in Section 3, most acquisitions in our data likely result in a change of control through the acquisition of majority interest or assets. In both cases, however, the acquirer may not buy the entire target (known as partial acquisitions), so there could be still quantitative differences in firm-level impacts relative to mergers.

their earnings, resulting in a statistically insignificant decrease in earnings (0.8 percent) on average. However, target workers who move jobs after an M&A event suffer immediate earnings losses that continue to accumulate over time, resulting in earnings that are about 4 percent less on average after the M&A event. Given the relatively substantial losses for workers in this group, the next part of the paper explores the source of this decline in earnings.

First, we explore whether changes in the characteristics of the firm employing the worker can explain the drop in earnings, following a recent literature on job displacement ([Schmieder et al. forthcoming](#) and [Lachowska et al. 2020](#)). The firm balance-sheet dataset allows us to test whether there are observable differences between the target firm and the new firm that target workers move to. We find that the new firms are in fact larger (in terms of employment and revenue) and more profitable on average compared to the target firms.<sup>5</sup> This result suggests that the decline in job movers' earnings cannot be explained by these workers moving to observably worse firms. Importantly, most of these job movers from target firms move to non-acquiring firms, meaning that this transition to larger firms is not mechanically driven by worker reallocation toward acquirers.

Interestingly, these workers still suffer earnings losses despite moving to these observably better firms. This finding suggests there is a pay complementarity between an individual worker and a given firm, which can arise for two reasons. First, models of firm-specific human capital accumulation ([Lazear 2009](#)) imply that a high-tenure worker is especially productive at a given firm. Second, firm pay policies, such as backloaded contracts ([Lazear 1979](#) and [Saez et al. 2023](#)), imply that high-tenure workers within a firm will be making in excess of their marginal product of labor, meaning that layoffs will lower these workers' pay. Both of these models suggest that high-tenure (or high-wage) workers will be especially harmed by job displacement. To assess this channel empirically, we conduct heterogeneity analyses based on worker tenure measured one year before the event. We find that the declines are almost entirely driven by workers with longer tenure. Individuals with shorter tenure are not impacted by the M&A-induced job transition. Furthermore, we find that the earnings losses are much larger among workers at the top quintile of the within-firm earnings distribution. This result is also consistent with the result based on worker tenure, since workers at the top of the within-firm earnings distribution are more likely to be the ones who have accumulated firm-specific human capital.

While we interpret these results as a loss of productivity driven by worker-firm complementarities, an alternative interpretation is that the M&A event allows firms to displace overpaid workers

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<sup>5</sup>Additionally, we estimate firm-specific wage premiums (AKM) to study whether job movers transition to systematically lower-paying firms. We find that target workers move to higher-paying firms on average after M&As (see Appendix A). One concern with the AKM estimation is that the effects are residuals which attempt to explain why a given firm pays more relative to other firms; in other words, there is a potential for omitted variables driving the estimated AKM effects. This highlights the importance of using firm balance-sheet data to directly assess their observable characteristics, such as their profitability.

with backloaded contracts. For example, [Shleifer and Summers \(1988\)](#) conjecture that takeovers allow new managers to renege on implicit contracts and lay off workers paid more than their productivity. We cannot directly observe worker productivity in our data, but firms displacing overpaid workers may increase their profits. We find a decrease in target firms' profitability, which seems inconsistent with this mechanism. While it is possible for these firms to lose profits after M&As despite displacing overpaid workers, there is another piece of evidence against the backloaded contract channel. If M&As are targeting overpaid workers to renege on their contracts, we should expect to see larger job transitions among high-tenure and high-wage workers. However, we find that changes in job transitions are similar between low-tenure (or low-wage) workers and high-tenure (or high-wage) workers after M&As. Therefore, this finding, combined with the result on decreased profits, suggests that the earnings losses among target workers appear most consistent with a loss of firm-specific human capital.

This paper contributes to several distinct literatures. Most directly, we contribute to a literature on the impacts of M&As on the labor market ([Brown and Medoff, 1988](#); [Siegel and Simons, 2010](#)). Recently, both [Prager and Schmitt \(2021\)](#) and [Arnold \(2021\)](#) study how M&A events that lead to large changes in labor market concentration affect worker earnings. In these papers, the goal of the analysis is to argue that other changes, such as shifts in productivity or product market power, are stable across M&A events that differentially impact labor market concentration. In this paper, we seek to directly estimate these impacts using firm balance-sheet data. Additionally, the sample of M&A events in this paper is quite different. In both our analysis sample and the whole sample of M&A events, we find that shifts in labor market concentration are relatively rare mainly because most M&A events happen across different labor markets. However, we find that these M&A events still have important impacts on job transitions that translate to changes in worker outcomes.

Relative to a recent work documenting that M&A events create significant displacements of workers from target firms with higher wage premiums and reallocation to smaller firms with lower wage premiums ([Lagaras forthcoming](#)) in a developing country, our work shows that the majority of workers displaced from target firms moves to larger and higher-paying firms. This distinction is important because while the prior study implies that M&As are similar to mass layoffs to the extent that displaced workers are forced to move to lower-paying firms, our findings suggest that M&As are different from mass layoffs; thus, a mechanism through which displacements affect workers is distinct in our setting. Besides finding different mechanisms, our study differs from the prior work by additionally using firm balance sheet data in a developed country with a larger set of M&A events and by assessing important firm outcomes in a coherent setting, such as changes in profitability. This helps us rule out alternative mechanisms, such as backloaded contracts, which could also explain the earnings decline in the prior study. Our analysis complements the study by [He and](#)

le Maire (2022), which focuses on the role of replacing managers in target firms. Theoretically, high-wage managers may be replaced after M&A events (Shleifer and Summers, 1988) leading to wage declines in target firms. While we cannot distinguish managers from other workers in our data, job displacements at targets may have been led by changes in management. Moreover, our work complements recent studies (Dessaint et al. 2017; Ouimet and Zarutskie 2020; Tate and Yang 2024) documenting human-capital considerations and labor restructuring as important drivers of M&As. While these studies show that acquirers engage in M&As to hire more (productive) workers, we find that most workers at targets move to non-acquiring firms after the event, implying that acquirers engage in M&As to purchase particular assets or intellectual properties of target firms, rather than to poach productive workers.

Second, our paper relates to a literature on how job transitions impact wages. In particular, a large literature studies the role of firm-specific and match-specific components to explain changes in displaced workers' wages (Lachowska et al. 2020; Schmieder et al. forthcoming). Much of the recent work leverages mass layoffs to identify impacts of job transitions on workers earnings. In this paper, we find that M&A events can lead to significant job transitions for workers from target firms, leading to a decline in their earnings. In contrast to findings in the mass layoff literature, we find not only that the earning losses are smaller on average, but also that displaced workers move to larger and more profitable firms on average, likely because M&A events tend to be pro-cyclical in aggregate (Rhodes-Kropf and Viswanathan, 2004).

Lastly, there is a large literature in corporate finance and industrial organization that studies how M&As affect firm performance, finding mixed empirical results (Schoar 2002, Hoberg and Phillips 2010, Boucly et al. 2011, Davis et al. 2014, Braguinsky et al. 2015, Blonigen and Pierce 2016, and Malmendier et al. 2018). Declines in firm performance have been attributed to misaligned incentives, such as empire-building motives (Jensen, 1986) and CEO overconfidence (Malmendier and Tate, 2005a). Furthermore, acquiring firms' profitability may also decrease if the acquisition was covertly intended to reduce competition and retain market power in the long run at the cost of losing profits in the medium run (Cunningham et al. 2021). Our paper contributes to this literature by estimating the impact of M&As on a variety of firm-level outcomes and by finding results consistent with the set of studies reporting negative impacts of M&As on firm profitability. We then use these firm-level outcomes to rule out alternative mechanisms, such as changes in firms' market power or implicit contracts, behind the decline in worker earnings. Our ability to link these firm-level outcomes to the worker-level outcomes is unique to the literature.

This paper is organized as follows. Section 2 provides institutional details, Section 3 describes our data, Section 4 discusses potential impacts of M&As, Section 5 describes our design, Section 6 shows our main results, and Section 7 discusses potential mechanisms behind our findings.



## 2 Institutional Background

This section describes relevant institutional details about competition policy and labor regulation. Competition policy in Canada is administered by the Competition Bureau, which enforces a two-step process for merger reviews, similar to the American process with notification thresholds and a supplementary information request. The thresholds for a pre-merger notification requirement are set by the Competition Act. The two most relevant thresholds are the size of parties and the size of transaction. Both of these must be met to trigger a pre-merger notification.<sup>6</sup> In our analysis sample, most M&A deals and involved parties are not large enough to trigger any of these thresholds.

Until recently, the Competition Act featured an “efficiencies defence” of mergers, allowing anti-competitive mergers to proceed if potential cost savings outweighed the losses to consumers through higher prices. In the U.S. and many other jurisdictions, efficiency is also considered as a factor in antitrust decisions for mergers but is given less weight.<sup>7</sup> Recent legislation repealed the efficiencies defence for merger reviews after December 15, 2023. Our work, assessing changes in acquirers’ and targets’ performance after M&As through a retrospective study, provides evidence that can inform policy discussions about how to weigh potential efficiency gains from M&As.

The Canadian and the U.S. labor markets share many similarities. Labor regulation is the jurisdiction of the provinces for most industries, but unemployment insurance is national through the Employment Insurance program. Measures of employment protection from the OECD show Canada close to the U.S. with much lower levels of protection from dismissals than European countries (OECD, 2020). Dismissing a worker requires several weeks of advance notice to the worker and mandatory severance depending on years of service. Non-compete clauses have recently been banned in certain provinces, but were generally unenforceable even before the explicit banning (Hanson and Cohen, 2012). Therefore, non-compete clauses are unlikely to be a substantial issue for worker transitions after M&As in our setting. One difference for labor markets in Canada is the degree of unionization. During our sample period, the overall coverage in the private sector by union contracts is 18 percent on average, compared to 8 percent in the U.S.<sup>8</sup> While the higher union

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<sup>6</sup>The parties together must have aggregate assets in Canada, or aggregate annual gross revenue from sales in, from or into Canada, in excess of 400 million Canadian dollars. The aggregate value of the assets in Canada to be acquired, or the aggregate annual gross revenue from sales in or from Canada generated from those assets, must be greater than 93 million Canadian dollars. There is also a size of equity threshold that less frequently comes into play: the acquirer holds 20 percent of the shares of a public corporation, 35 percent of the share of a private corporation, or 35 percent of the profits or assets on dissolution of a non-corporate entity.

<sup>7</sup>Ware and Winter (2016) assert (p. 366) for Canada that “in no other jurisdiction in the world would a court accept evidence of substantial price effects from a merger and yet allow the merger.” A comparison of efficiency defences across many OECD countries is found in OECD (2013).

<sup>8</sup>The source for the Canadian statistics is Statistics Canada, Table 14-10-0070-01, “Union Coverage by Industry.” The source of the U.S. statistics is Bureau of Labor Statistics, Series ID LUU0204906700, “Percent of Employed, Private Wage and Salary Workers Represented by Unions.”

coverage in Canada may render it more difficult to lay off workers in general, high unionization may also make M&As an attractive opportunity to restructure the company (i.e., via layoffs).

### 3 Data

This section describes two datasets used for our analysis. First, we use the SDC Platinum database which contains information on M&A activities around the globe. Second, we use the Canadian Employer-Employee Dynamics Database, a matched employee-employer dataset that covers the universe of firms and workers in Canada from 2001 to 2017. To prevent disclosure of confidential information, Statistics Canada requires researchers to round estimates and observation counts.

#### 3.1 SDC Platinum

The SDC Platinum database allows detailed search on M&A activities around the world, covering the near universe of transactions involving either publicly traded or private companies in Canada.<sup>9</sup> This database contains variables that indicate the type of a transaction, such as merger or different kind of acquisition, and identify the target firm and the acquiring firm in each transaction.<sup>10</sup> Our main analyses include both mergers and acquisitions as well as show the results separately. Since the literature focuses on the change of control as a key determinant of post-M&A outcomes, we include acquisitions given that many of them may result in a change of control. To understand this choice, we provide details on how the SDC database categorizes each M&A transaction.

In our database, a merger is defined as an event in which either two parties become one entity or an acquirer buys 100 percent of a target's stock. By contrast, an acquisition is defined as taking over either a target's assets or a part of its stock. In the sample of all M&As that are matched with our firm-level data in Canada between 2001 and 2017, mergers account for 25 percent (20 percent in our analysis sample) of the cases, while acquisitions account for the rest. Acquisitions are further categorized into the following: (1) acquisition of majority interest, meaning deals in which the acquirer holds less than 50 percent and is seeking to purchase 50 percent or more, but less than 100 percent of the target company's stock, (2) acquisition of remaining interest, meaning deals

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<sup>9</sup>While [Barrios and Wollmann \(2024\)](#) notes that the SDC data may not cover all M&A deals in the U.S., identifying potentially missing M&As in Canada (which might be a small share) is beyond the scope of our study.

<sup>10</sup>In a relatively rare transaction with multiple ownership changes at the same time due to multiple acquiring firms involved for a given target firm, it identifies all the acquirers. Relatedly, if a transaction involves multiple ownership changes across different years by the same or a different acquiring firm, the SDC data keeps tracks of these transactions over time and separately lists each of these transactions for each year. We describe how we treat this type of transactions in our analysis in [Section 5](#).



in which the acquirer holds over 50 percent and is seeking to purchase 100 percent of the target company's stock, (3) acquisition of partial interest, meaning deals in which the acquirer holds less than 50 percent and is seeking to acquire less than 50 percent, or the acquirer holds over 50 percent and is seeking less than 100 percent of the target company's stock, (4) acquisition of assets, meaning deals in which the assets of a company, subsidiary, division, or branch are acquired, and (5) acquisition of certain assets, meaning deals in which certain assets of a company, subsidiary, division, or branch are acquired.<sup>11</sup> Except for (but possibly including some of) the acquisition of assets, these events are also known as partial acquisitions. While the acquirer may not purchase the entirety of the target in an acquisition, the acquirer gains the controlling stake of the target in the acquisition of majority interest (10 percent of all M&As). In this sense, an acquisition of majority interest is conceptually similar to a merger to the extent that there is a change of control.

By contrast, when the acquirer buys the remaining interest (3 percent of all M&As), there is no change of control even though the acquirer already had the controlling stake and now has the full control. Similarly, when the acquirer buys a partial interest (14 percent of all M&As), there is no change of control and the acquirer does not have the controlling stake unless it already had more than 50 percent of the target's stock prior to the event.

On the other hand, when an acquirer buys the assets of the target (47 percent of all M&As), the acquirer can buy the entire company's assets or its division's assets, so the acquirer may or may not have the controlling stake in the case of asset purchases. For example, if an acquirer buys two out of three factories owned by a target, then it appears that the acquirer now has the controlling stake. However, even if the acquirer buys only one factory, it still controls that specific division of the target. Therefore, it is difficult to say whether there is an absolute change of control in the acquisition of assets without more information from the data. If the acquisition of assets leads to a change of control, then again, it is conceptually similar to a merger. In the acquisition of certain assets (1 percent of all M&As), however, it is unlikely that there is a change of control.

Since M&As may lead to different outcomes depending on whether there is a change of control (see Section 4), we divide our firm-level analyses into those involving mergers and those involving acquisitions. Furthermore, M&As may have quantitatively different impacts depending on whether the acquirer buys a part or the entirety of the target, regardless of a change of control. Again, in the acquisition of majority interest, there is a definite change of control and most asset acquisitions likely result in a change of control as well, despite not buying the entire target.<sup>12</sup>

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<sup>11</sup>In fact, there is a separate category simply labeled as "Acquisition" in the data, in which 100 percent of a target is spun off, meaning that the acquirer takes over the entirety of the target and creates a new company. However, such an event is rare in our sample (0.4 percent of all M&A events) and therefore is dropped from our analysis sample.

<sup>12</sup>Our main results remain qualitatively similar even if we exclude the acquisition of remaining interest, partial interest, and certain assets (together 18 percent of all M&A deals in our data) where there is no change of control. These results can be provided upon request.

Finally, the SDC dataset includes names of the parties, NAICS industry codes, and other identifying information such as addresses and phone numbers. The dataset on M&A activities was merged with the firm-level data from the Canadian Employer-Employee Dynamics Database using all available identifying variables, such as firm names, addresses, and industry codes. The match rate is around 75 percent on average from 2001 to 2017.

### 3.2 Canadian Employer Employee Dynamics Database (CEEDD)

The Canadian Employer-Employee Dynamics Database draws information from both individual (T1) and corporate (T2) tax return records, merged with job-level information using T4 employee tax records (like a W-2 in the United States, with information on annual earnings) and Record of Employment (ROE) data with information on work history, and with firm-level information from the National Accounts Longitudinal Micro-data File (NALMF). This database has rich information on the universe of firms and workers in Canada from 2001 to 2017.

The main firm-level outcome variables are employment, average payrolls, and profitability. Employment is defined as the average number of employees reported on the NALMF. Average payrolls are defined as the total wage bill divided by the number of employees. We use three different measures of profitability: (1) profit margins, defined as total revenue minus total expenses scaled by total revenue, (2) returns on assets, defined as net income divided by total assets, and (3) EBITDA per worker. In Appendix B, we look at other outcomes such as markups and realized capital gains. We define markup as the elasticity of output with respect to variable costs as well as the variable costs share (De Loecker and Warzynski, 2012). For the elasticity of output with respect to variable costs, we use estimates from De Loecker et al. (2020) based on the U.S. data that allow for different elasticities across two-digit NAICS industry codes and years. Given the elasticity estimates, this allows us to estimate firm-level markups as the output elasticity multiplied by the inverse of the variable costs (total wage bills and material costs) share:  $\hat{\theta}_{st} * \frac{Sales}{Costs\ of\ Goods}$ . Furthermore, we link ownership data with the firm-level data to compute realized capital gains by owners for a given firm in each year. After merging individual tax returns data with the ownership data at the investor level, we aggregate owners' realized capital gains for each firm in each year to compute total realized capital gains by these investors at the firm level.

At the worker level, the key outcome variable is annual earnings, aggregated across all employers for that worker in a given year. While we include earnings across all employers, we associate workers with the “dominant” employer from which the employee receives the highest pay in the year. We also use information on workers' gender and age derived from the T1 income tax filing for creating a matched control group and for heterogeneity analyses (Appendix B).

## 4 Potential Impacts of M&As on Firms and Workers

In this section, we discuss potential impacts of M&As on firms and workers. The first subsection focuses on firm-level impacts, while the second one discusses how these firm-level impacts may spill over to workers, with a particular focus on the role of job displacement.

### 4.1 Firm-Level Impacts

There can be several motivations for a M&A. Acquiring firms often argue that M&As can lead to increased productivity through synergies. However, antitrust authorities often focus on the potential for M&As to increase prices. Furthermore, models in corporate finance discuss alternative, *profit-reducing* motivations for M&As, such as empire-building (Jensen, 1986) or CEO overconfidence (Malmendier and Tate, 2005b). Overall, these various motivations may lead to heterogeneous impacts of M&As on firm performance. In the next section, we discuss how these different motivations may impact firm outcomes. In particular, we focus on two core outcomes: profitability and firm size. As we discuss further in the next subsection, changes in these outcomes are crucial for understanding different mechanisms behind worker outcomes.

Another issue to keep in mind when interpreting firm-level impacts of M&As is that there are two main types of events that we consider: mergers and acquisitions. The discussion below applies to both types of events, although there may exist two key differences between these types, as described in Section 3. First, mergers result in a definite change of control, while acquisitions may not necessarily lead to a change of control. There can be qualitative or quantitative differences in firm outcomes depending on whether there is a change of control, which we describe below. Second, regardless of a change of control, there can be quantitative differences depending on whether the acquirer merges with the target or buys only a part of it. Furthermore, in most acquisitions, target firms continue to operate afterwards, whereas many target firms disappear from the data after mergers.<sup>13</sup> Therefore, we can evaluate firm outcomes separately by the target and the acquirer in the case of acquisitions, while we assess their joint outcomes in the case of mergers.

#### 4.1.1 Profits

There are three primary channels in which M&As can increase overall profits: changes in product-market power, changes in labor-market power, and changes in productivity. The primary focus of

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<sup>13</sup>Even if some target firms remain in the data after mergers, they are now subsidiaries of their acquirers, so it is difficult to assess their outcomes separately because their parents' financial decisions directly impact the subsidiaries' financial statements (which is not the case in the case of acquisitions).

the industrial organization literature on M&As has been the potential for price increases via increased product market power, which would lead to an overall increase in firm profits.<sup>14</sup> Similarly, increases in labor-market power will also lead to increased profits through decreased wages.

A common justification for M&As by firms is the possibility of increased profitability through productivity synergies. For example, [Braguinsky et al. \(2015\)](#) and [Demirer and Karaduman \(2024\)](#) find evidence of increased productivity after acquisitions. However, there are also prior studies that find negative impacts of M&As on firm performance. For example, [Malmendier et al. \(2018\)](#) finds that acquiring firms that win bidding contests for target firms suffer losses in stock returns relative to competitors. Possible motivations for productivity-reducing M&As include misaligned empire-building motives ([Jensen, 1986](#)) or CEO overconfidence ([Malmendier and Tate, 2005a](#)). Decreases in productivity will lead to declines in profits. This is a key distinction: increases in product-market power and labor-market power should not be associated with profit decreases, while decreases in productivity can lead to lower profits.

It is possible that there may not be any change in product-market power, labor-market power, or productivity unless there is a change of control (i.e., via mergers). However, even without a change of control, an acquirer may still increase its market share by buying a part of the target's businesses in the same industry. Furthermore, by purchasing particular divisions of the target with potential synergies, the acquirer may enhance its productivity. Therefore, the impacts on firms' profitability may be qualitatively similar in both mergers and acquisitions, although there could be still quantitative differences. Furthermore, if the acquirer buys only productive units of the target in the case of acquisitions, then the acquirer's profits may increase, while the target's profits would go down, so profitability may remain similar in the aggregate. By contrast, the acquirer can sell off unproductive units of the target after a merger, and the combined entity's profits may still increase.

#### **4.1.2 Firm Size**

The impacts of M&As on firm size, measured by the average number of employees and average payrolls, are ambiguous overall. Increases in product-market power will lead to declines in employment. This is because firms will increase prices, which will reduce demand, leading to a decrease in overall employment. By contrast, when the labor-market power increases, the impact on employment is ambiguous. In static monopsony models, increases in labor-market power will cause firms to decrease employment in order to decrease wages. However, in wage-bargaining models ([Abowd and Lemieux, 1993](#)), wages can decrease with no changes in employment.

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<sup>14</sup>See [Whinston \(2007\)](#) and [Asker and Nocke \(2021\)](#) for review articles on antitrust evaluation in the industrial organization literature.

Moreover, changes in productivity can lead to either higher or lower levels of employment. If the M&A leads to a more efficient production process, then we could see increases in employment as the firm expands. However, one potential source of synergies is through a reduction in redundant workers, which would lead to declines in employment. Lastly, if the M&A is motivated by empire-building motives (Jensen, 1986) or CEO overconfidence (Malmendier and Tate, 2005a), then we might expect to see productivity declines accompanied by employment losses if the acquiring firm is now less productive and loses market shares. Given the ambiguous impacts of M&As on productivity, the impacts on employment and payrolls are also ambiguous.

Similar to the predictions on firms' profitability, there could be both qualitative and quantitative differences on firm-size impacts between mergers and acquisitions. However, if either a merger or an acquisition leads to a pure reallocation of workers from the target to the acquirer, then the overall employment would stay the same in the aggregate in both cases.

In summary, there are overall ambiguous impacts of M&As on both firm size and profitability depending on the motivations and the type of transaction. Next, we discuss potential impacts of M&As on workers and how changes in firm outcomes may spill over to workers.

## 4.2 Impacts on Workers

In this section we discuss how the firm-level changes in profitability and sizes translate to impacts on workers. To help guide the discussion, we first set up a parsimonious conceptual framework.

Assume that the value of worker  $i$ 's output at firm  $j$  is given by  $p_{ij} = \rho_j + \xi_{ij}$ , which is composed of a firm component ( $\rho_j$ ) and a match component ( $\xi_{ij}$ ). The firm component of pay is common to all workers within a firm, while the match component is specific to the worker-firm match. The worker's outside option is given by  $O_i$ . A simple Nash bargaining model implies the wage paid to the worker is given by<sup>15</sup>:

$$w_{ij} = \gamma_j \cdot p_{ij} + (1 - \gamma_j) \cdot O_i \quad (1)$$

Following Lazear (1979) and Saez et al. (2023), we assume a firm pays a firing cost  $f$  in the event of a layoff. We will return to this firing cost when discussing the potential impact of M&As on job displacement, but first, we discuss the potential impacts on wages.

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<sup>15</sup>Many wage-determination models have a similar form. For example, in union-bargaining models (Abowd and Lemieux, 1993), standard static monopsony models (Manning, 2003), and job-differentiation models (Card et al., 2018), the wage is an average of productivity and the outside option of workers. The relevant measures of productivity, the determination of the weights, and the determination of the outside option vary across these models.

### 4.2.1 Wages

Given this framework, we can now discuss how M&As will impact wages, drawing on the previous discussion in the firm-level impacts section. On the one hand, the value of output,  $p_{ij}$ , may increase either through increased productivity or through higher prices via increased product-market power. The increase in value of output through this higher  $\rho_j$  would lead to higher wages.

On the other hand, M&As may also lead to decreased productivity, resulting in a decline in workers' wages through lower  $p_{ij}$ . Lastly, if the firm gains labor-market power, then wages may decrease through two channels. First, the bargaining parameter  $\gamma_j$  could decrease, implying that workers get less of the overall rents in the firm, leading to lower wages. Second, the outside option  $O_i$  could also decrease if M&As reduce the number of competing employers in a given labor market, leading to lower wages.

To the extent that there can be both qualitative and quantitative differences in impacts on firms' productivity between mergers and acquisitions, the subsequent impacts on workers' wages may be also different between the two types. By contrast, if the impacts on firms' profits are similar in both cases in the aggregate, the worker-level impacts would be also similar in both types. Furthermore, changes in  $p_{ij}$  would only have direct impacts on wages of workers who continue to stay at the firm going through a M&A. However, changes in firms' productivity or profitability may lead to layoffs, which can still affect workers leaving the M&A firm through job displacement.

### 4.2.2 Job Displacement

Changes in productivity may lead to job displacement for workers at either target or acquiring firms. For example, M&As may lead to layoffs due to shifts in the production process or organization structure, leading to changes in the worker-firm match component ( $\xi_{ij}$ ). If this falls far enough (or to zero, if the worker's job is now redundant), then this will lead to layoffs.

There are two potential channels through which  $\xi_{ij}$  may fall after an M&A event. First, there could be a set of overpaid workers at the firm whose marginal product is below their current wage due to backloaded contracts. Lazear (1979) illustrates a model in which contracts are structured in a way such that workers are eventually paid above their marginal product. These backloaded contracts are optimal in equilibrium as they solve a principal-agent problem. Importantly, firms have an incentive to lay off older workers whose marginal product is lower than their wage. However, laying off these workers involves a firing cost, which may be either explicit or implicit. Explicit firing costs stem from employment protection registration that may impact the cost of firing a worker, which is studied in Saez et al. (2023). Implicit costs involve reputation concerns: if the firm lays off



older workers, then younger workers would not accept the lower pay initially. However, [Shleifer and Summers \(1988\)](#) posit that M&As may allow new managers to displace overpaid workers and reduce salaries by breaching the trust developed between the previous manager and the workers.

Another source of the fall in  $\xi_{ij}$  is through the destruction of firm-specific human capital. Models of firm-specific human capital, such as [Lazear \(2009\)](#), predict that workers with high tenure will accumulate skills that make  $p_{ij}$  particularly large at their firm. The key difference relative to the previous channel is that in the case of backloaded contracts, firing overpaid workers with wages above their productivity will increase profits. By contrast, in the case of firm-specific human capital, the M&A may destroy this worker-firm match, leading to a drop in productivity both for the worker and the firm. Because the acquiring firm may not have full information on  $\xi_{ij}$ 's across all workers and the target firm may not have full control over which workers to retain, certain workers with high  $\xi_{ij}$  (concentrated among high-tenure and high-wage groups within a firm) may experience a drop in productivity when being forced to move to a different firm after a M&A. This may also have negative consequences on the target firm's performance as well, leading to lower profits. Therefore, the firm-level outcomes will be important for disentangling these two channels.

Interestingly, both channels predict that the fall in  $\xi_{ij}$  and the subsequent earnings losses would be larger among high-tenure and high-wage workers within a firm. However, the key distinction is that in the backloaded contract channel, post-M&A job transitions will also be higher among high-tenure and high-wage workers if the new manager is targeting overpaid workers, while that is not necessarily the case in the firm-specific human capital channel. In [Section 7](#), we conduct heterogeneity analyses based on workers' tenure and the within-firm earnings distribution, and assess changes in firms' profits to test which of these models seems most consistent with our findings.

Finally, a fall in the worker-firm match will cause short-run losses in worker earnings, which may or may not diminish over time as displaced workers seek to find a comparable match. A different source of long-run losses in earnings may stem from the firm-specific component of pay  $\rho_j$ , which can be proxied by their size or profitability. Therefore, we also examine whether workers who leave target firms after the event move to smaller or less profitable firms in [Section 7](#).

## 5 Empirical Strategy

This section describes our empirical design and provides descriptive statistics on our matched sample. To estimate the effects of M&As on firm-level outcomes, we implement a matched difference-

in-differences design by estimating a regression of the following form:

$$Y_{jt} = \sum_{k=-4}^5 \beta_k^{MA} \mathbb{1}(t_j = t^* + k) \times MA_j + \tau_t + \psi_j + u_{jt} \quad (2)$$

where  $Y_{jt}$  is an outcome variable for firm  $j$  at year  $t$ ,  $MA_j$  is an indicator for an M&A firm,  $\mathbb{1}(t_j = t^* + k)$  indicates an M&A event occurred  $k$  years in the past (or future) relative to the period of the M&A event  $t^*$ ,  $\tau_t$  are year fixed effects that vary by the year of the M&A event,  $\psi_j$  are firm fixed effects, and  $u_{jt}$  is an error term. To absorb any industry-specific shock affecting M&A activities in a given year (Maksimovic and Phillips 2001), we include 4-digit industry dummies interacted with year dummies as control variables. Furthermore, we include a quartic function in firm age to ensure that our results are not driven by differences in financial constraints of firms, although dropping these control variables does not qualitatively affect our results. The standard errors are clustered at the firm level. Note that for firms that go through M&As more than once, we focus on their first M&A event although we still allow the subsequent M&A event(s) to affect their outcomes in the post-event period.<sup>16</sup> Furthermore, most of the M&A firms and their matched control firms are private companies (96 percent) in our analysis sample, so our results remain nearly unchanged even if we focus our analyses exclusively on private firms (Appendix A).

In our matched sample described in the next subsection, roughly 80 percent of M&A events involve acquisitions. Relative to a merger, an acquisition results in the target firm continuing its operation as an independent entity, which allows us to keep track of the target's outcomes separately from the acquirer's outcomes after the event, as described in Sections 3 and 4. To have a clearer picture on how firm size changes after M&As, we examine our firm-level outcomes separately for those involved in acquisitions and for those involved in mergers. Specifically, when assessing acquisitions, we look at targets and acquirers separately. In contrast, when assessing mergers, we look at the aggregate (targets and acquirers pooled) outcomes, although we also look at the aggregate outcomes for acquisitions as a comparison. Additionally, we use the worker-level data to examine the extent to which workers reallocate from targets to acquirers.

To assess worker-level impacts, we estimate a similar matched difference-in-difference design of the following form:

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<sup>16</sup>When focusing on the first acquisition or merger, we still cover the subsequent event(s) in the post-event period for firms that go through M&As more than once. While the majority (roughly 70 percent) of firms goes through a merger or an acquisition only once during our sample period, there exists a handful of acquirers that engage in M&As more than once. Our main analysis sample still includes these acquirers involved in multiple M&As over time. In addition, we separately examine acquirers that go through multiple M&As in Appendix A.

$$y_{it} = \sum_{k=-4}^5 \beta_k^{MA} \mathbb{1}(t_i = t^* + k) \times MA_i + \tau_t + \omega_i + \mu_{it} \quad (3)$$

where  $y_{it}$  is an outcome variable for incumbent worker  $i$  at year  $t$ ,  $\omega_i$  are worker fixed effects, and all other variables are defined as in Equation (2). The standard errors are two-way clustered at the worker-by-firm level.<sup>17</sup>

The key identifying assumption is that outcomes for M&A firms and workers and for control firms and workers would have trended similarly in the absence of the M&A event. This assumption may be strong since M&As are the result of firms' decisions. For example, acquiring firms may target firms that will be profitable in the future, whose earnings may grow even absent the acquisition. By contrast, acquirers may target mismanaged and underperforming businesses whose employment and workers' earnings would fall regardless of M&As. One way to determine the direction of the potential bias is by comparing outcomes for M&A firms and workers to the control firms and workers in the years prior to the M&A event. Parallel pre-trends in firm-level and worker-level outcomes help alleviate the aforementioned concerns. In fact, we find that M&A firms' outcomes share parallel pre-trends with those of control firms, not only in terms of their size and profitability (Sections 6 – 7), but also with regards to their sales and markups (Appendix B).

While verifying common pre-trends is reassuring for a causal interpretation, contemporaneous shocks that occur with M&A events could still bias our results. For example, there could be a negative demand shock that hits a particular market and causes both a decline in wages and an increase in M&A activities as firms get purchased before they shut down. In this case, M&A activities are correlated with shocks that decrease labor demand. We can also have the opposite scenario; in fact, merger activities tend to be pro-cyclical in aggregate (Rhodes-Kropf and Viswanathan, 2004). We address this concern by also looking at M&As that are less likely to have been triggered by local economic conditions. Specifically, we consider the impact of national M&As that occur among domestic firms with multiple establishments across different commuting zones, and therefore are less likely to be driven by local economic conditions faced by the firms or workers (Appendix B).

## 5.1 Matched Samples

Before performing a matching procedure between M&A firms and potential control firms, we make the following restrictions. First, we require a firm to have at least 10 workers one year prior to the event. This choice focuses our sample on economically active firms with enough pre-period

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<sup>17</sup>Our results are robust to two-way clustering standard errors at the worker and the market (4-digit industry by commuting zone) level – see Appendix A.

observations, and drops most small businesses that are not comparable to either acquiring or target firms. Additionally, we drop a small share of firms that have missing observations for the key variables used for matching, such as employment, average payroll, and firm age, measured one year prior to the event. During our sample period, the total number of M&A events is 765 per year on average among this sample of firms eligible for matching (see Figure 1). We then match each firm one year prior to an M&A event to a control firm in the same province and 2-digit NAICS industry. A firm is a potential control firm for firm  $j$  if: (1) the firm is never involved in an M&A event during our sample period, and (2) the firm is in the same decile bin of average payrolls and is in the same 15-quantile bins in total revenue and firm age one year prior to the M&A event of firm  $j$ .<sup>18</sup> Of all the possible counterfactual firms for a given M&A firm, we choose the firm with the closest propensity score (one-to-one matching within a caliper), which is estimated by predicting treatment using a linear probability model with a quadratic function in average payrolls, total revenue, and firm age in year  $[t - 1]$ . This matching strategy is similar to a number of recent papers implementing a dynamic difference-in-differences design (Goldschmidt and Schmieder, 2017; Smith et al., 2019; Arnold, 2021). The matching strategy finds a control firm in about 80 percent of all cases among this eligible sample.

Choosing one counterfactual control firm per M&A firm in a given year ensures that the treated and control groups are comparable on the matched variables. We construct an unbalanced panel of firms which extends 4 years prior to and 5 years after the M&A event. M&A firms in our analysis sample are larger than those that fail to find a control group on average in part because we drop firms with fewer than 10 employees. Therefore, the M&A deals in our analysis sample are meaningful and larger than an average M&A deal in Canada during the sample period. Furthermore, there exist firms eligible for matching (i.e., with at least 10 employees), but not matched to control firms (i.e., similar in sizes to M&A firms but located in a different province). We provide descriptive statistics on the sample of unmatched M&A firms (the remaining 20 percent among the eligible firms) and show robustness test results including these excluded firms in Appendix A.

Matching on size, sector, and province finds treatment-control pairs that would plausibly exhibit common trends in the absence of M&A activity. While we do not explicitly match firms based on a market (defined at the 4-digit NAICS by commuting zone level), it is possible that firms are matched within the same market. This raises a potential concern if M&As have impacts on local labor markets through increased concentration. If M&As have negative effects on control firms in the same industry and commuting zone, then the impact of M&As on firms will be biased towards

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<sup>18</sup>While this specification yields the best parallel pre-trends on the key outcomes, which are important for our identification strategy, our main results remain qualitatively similar when we use different bin sizes or other related firm characteristics for matching. For example, we can make the matching criteria more restrictive so that the treated and control groups look even more comparable; we still find similar results with a smaller sample size. Results based on other variation in matching can be provided upon request.

zero. However, these potential spillover effects are minimal in our setting given that the number of M&A deals within the same market is small and the number of competitors in a given market is so large that most M&A deals do not contribute to a meaningful increase in concentration (see Appendix B). Nevertheless, to minimize this concern, we do a robustness check by matching firms within the same province, but in different markets (Appendix A).

To construct the worker-level sample, we extract all workers who were continuously employed in the matched firms during the entire pre-event period (4 years). This tenure restriction is chosen to obtain a sample of workers with attachment to the M&A firms and is similar to tenure restrictions used in the mass layoff literature (Jacobson et al., 1993; von Wachter et al., 2009; Lachowska et al., 2020). Additionally, we restrict workers to have at least 4,000 CAD in annual earnings and drop multiple jobholders to ensure that we study full-time workers with stable income and attached to their firms (Card et al., 2013; Sorkin, 2018), given that we do not observe work hours in our data. For each worker in a treated firm, we choose a worker in any of the matched control firms in the same sector, province, worker age (five-year) bin and gender. If more than one match is found, we choose the worker with the closest propensity score to the treated worker (one-to-one matching within a caliper), where the propensity score is estimated by predicting treatment using a linear probability model with a quadratic function in worker age. In total, about 60 percent of workers at treated firms are matched to control workers among the sample of eligible workers at matched control firms. We provide descriptive statistics for the unmatched workers and show robustness test results including these excluded workers in Appendix A. To compute earnings in the worker-level data, we aggregate annual earnings across all employers if a worker is employed at more than one firm in a given year (due to job transitions).

Since we use a matched control group that is never treated, the specifications above do not suffer from the identification issues that arise in conventional event-study designs (Borusyak et al., 2021) with staggered timing (Goodman-Bacon, 2021). Identification here comes from differences in always-treated and never-treated units over time, not from units coming in and out of treatment.<sup>19</sup>

Finally, we use only completed transactions in our analyses following the M&A literature. The number of withdrawn M&A events that meet the sample restrictions described above (or even without the restrictions) is very small in Canada. While using firms (and workers) involved in withdrawn M&A events as a placebo group yields null effects across the key outcomes, many of these results are also quite noisy given the small sample size.<sup>20</sup>

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<sup>19</sup>We let the first M&A event in our analysis sample start from 2005 to ensure that our matched control firms did not go through any M&A event at least 4 years prior to (or after) the first M&A event, given that our data starts from 2001. In this way, we ensure that none of the matched control firms went through any M&A event from 2001 to 2017.

<sup>20</sup>These placebo test results can be provided upon request.

## 5.2 Descriptive Statistics

We close this section with descriptive statistics of our analysis sample. Panel A of Table 1 shows the averages for key variables across firms, comparing M&A firms to the set of matched control firms one year before the event. On average, M&A firms are a bit larger than their control firms, in terms of total revenue, expenses, and employment. However, for average payrolls, leverage ratio, and markups, M&A firms are comparable to their control firms, suggesting that firms that go through M&As and firms that never get involved in M&As are comparable with regards to their average employee compensation and financing structure prior to the event. Importantly, as we show in Sections 6 – 7 and Appendix B, M&A firms and their control firms share parallel pre-trends on these variables, implying that they exhibited similar patterns not only in terms of employment and average payrolls, but also with regards to sales and profitability, prior to the M&A event. The dominant sectors are manufacturing, wholesale, and services (about 70 percent of our firm sample).

Panel C of Table 1 shows average worker characteristics in our analysis sample one year before the event. We distinguish between workers at acquiring firms and workers at target firms. Annual earnings are 70,046 CAD and 71,386 CAD among workers at acquiring firms and workers at the matched control firms, respectively. Annual earnings are 70,625 CAD and 72,317 CAD among workers at target firms and workers at the matched control firms, respectively. Therefore, annual earnings of treated workers are comparable to those of control workers on average. The difference between average payrolls in Panel A and annual earnings in Panel C arises because not all workers from treated firms are matched to workers at control firms, as we impose restrictions on tenure, worker age, and gender for matching. Hence, the average payrolls at a given firm may not equal the average annual earnings of workers at a given firm in our matched sample. The age and gender compositions are similar between treated workers and their matched control workers.

## 6 Results

This section reports the results from the estimation of the difference-in-differences model described in Section 6, and presents additional tests supporting the interpretations of our results.

### 6.1 Post-M&A Firm Size and Average Payrolls

Figure 2 plots estimates of  $\beta_k^{MA}$  from equation (2) across the main firm-level outcomes using our matched sample. As mentioned in Section 5, we show these results separately for targets and for acquirers involved in acquisitions (Panels (a) and (b)) and for the aggregate (targets and acquirers



pooled, Panels (c) and (d)). Panel (a) shows that acquiring firms' and target firms' employment followed a similar pattern as those of their matched control firms before the M&A event. This pre-event stability is important evidence in support of our empirical strategy. While acquirers' employment significantly increased after the event, targets' employment decreased after the event, compared to those of matched control firms. Panel (b) shows that the pre-event trends for average payrolls are also similar between M&A firms and their control firms. While acquirers' average payrolls stayed flat after the M&A event on average, targets' average payrolls significantly decreased after the event, compared to control firms' average payrolls.<sup>21</sup> In Panel (c) and Panel (d), we pool targets and acquirers and compare their outcomes with those of their matched control firms before and after the M&A event. Here, we simply drop the distinction between acquirers and targets when estimating the equation (2) so that acquirers, targets, and their matched control pairs are all in the same estimation.<sup>22</sup> The black line indicates the aggregate outcomes for those involved in mergers and the red line indicates the aggregate outcomes for those involved in acquisitions. In the case of acquisitions, employment and average payrolls did not change much after the event on average, except for the initial drop in average payrolls. Similarly, in the case of mergers, employment and average payrolls stayed roughly flat after the event on average, except for the initial increase in employment.

To interpret the magnitude of these results, Table 2 presents the difference-in-differences estimates on these outcomes, separately for targets and acquirers involved in acquisitions (Column 1 and 2) and for the aggregate (Column 3 and 4). Column (1) shows that target firms' employment and average payrolls fell by 8.9 log points and 2.8 log points, respectively, after the event. Column (2) shows that acquiring firms' employment increased by 18.8 log points, without significant changes in average payrolls, after the event. Columns (3) and (4) show these outcomes in the aggregate, separately for those involved in acquisitions and for those involved in mergers. Columns (3) and (4) show that employment and average payrolls did not change much on average in the aggregate in the case of either acquisitions or mergers.<sup>23</sup> Overall, our results show that M&As lead to mechanical firm-size adjustments when we look at targets and acquirers separately, but in

<sup>21</sup>To account for a small share of firms (especially among targets) that exit the sample after the M&A event, we repeat the same analysis using outcomes measured in levels instead of in logs, where we replace the missing observations with zeros. We find qualitatively similar results when we use the outcomes in levels (see Appendix A).

<sup>22</sup>An alternative approach would be to combine acquirers and targets into single entities (i.e., summing up their observations), re-do the firm-level matching to find their control firms, and estimate the equation (2) using the combined entities. However, this new matched sample would be likely different from the original analysis sample, given that finding matched control firms for the combined entities may result in a lower match rate and therefore a smaller matched sample.

<sup>23</sup>In Column (3), the average estimates are weighted by the relative sample sizes between targets (Column 1) and acquirers (Column 2), where the target sample size is much larger. Therefore, the negative coefficient estimates from the targets carry a larger weight than the positive coefficient estimates from the acquirers, leading to zero (or slightly negative and insignificant) effects on average at the aggregate level.

the aggregate, firm size remains relatively flat both in acquisitions and in mergers. To get a better sense of the impacts on worker reallocation and average payrolls, we next turn to worker-level data to assess the impacts of M&As on worker earnings and job transitions.

## 6.2 Post-M&A Worker-level Earnings and Job Transitions

Given the considerable turnover at target firms after M&As, changes in average payrolls may reflect changes in worker composition. For example, the decrease in firm-level wages could be driven by laying off high-wage workers or reducing wage growth for existing workers, or a combination of both. Therefore, we next turn to the worker-level data that allows us to flexibly control for composition by tracking the same workers over time. We focus on incumbent workers who were already employed at their firms before the event and follow their outcomes throughout our sample period. Here, we do not make a distinction between workers involved in acquisitions and workers involved in mergers, although the results are very similar when we assess them separately (Appendix B).

Panel (a) of Figure 3 shows that annual earnings for incumbent workers at target firms trend similarly to those of their matched control workers in the years prior to the event, but fall significantly afterwards. By contrast, annual earnings of incumbent workers at acquiring firms trend similarly to those of their matched control workers and stayed flat after the event. Column (1) of Table 3 shows that workers at target firms experience a decline of 1.2 log points in annual earnings after the M&A event on average. This decline could be due to M&A workers moving to lower-paying firms or M&A firms reducing earnings for their incumbent workers. Column (2) shows that the annual earnings of workers at acquiring firms did not change much on average.

The drop in employment at target firms, as shown in Table 2, suggests that job transitions could explain a part of the decline in worker-level earnings. The reduction in employment could come primarily through decreased hiring, implying incumbent workers may be relatively unaffected. We first consider the impact of M&As on the probability of worker transitions from a job. This transition could be to another firm or to non- or self-employment. In our data, most of the workers who leave their original employers do so involuntarily, but find a job afterwards.<sup>24</sup> Panel (b) of Figure 3 plots the estimates of equation (3) with an indicator for a job transition as the outcome, which equals one in year  $t$  if a worker moved to another firm in  $t - 1$ .<sup>25</sup> As shown in the figure, job

<sup>24</sup>Our data has an indicator for reasons for job separations, which can be broadly categorized into involuntary (i.e., shortage of work, takeovers, or retirement) or voluntary (i.e., personal or medical reasons). Roughly half of the observations in the relevant sample has these reasons as “unknown.” Among the other half, about three quarters of workers moving from target firms left involuntarily after the M&A event.

<sup>25</sup>This is because we impose that workers are continuously employed at their firms until the M&A event. Workers are allowed to move to other firms in the year of the event, but have to be employed at their original firm first at  $t = 0$  before the transition. Therefore, there is an one-year lag in the indicator for a job transition.

transitions spike in  $t = 1$ , meaning that many workers transitioned to another firm in the year of the M&A event. Workers in target firms are about 20 percentage points more likely to transition jobs in the year of the M&A event. Transition rates then drop, but remain elevated relative to workers in the control group. Given this notable increase in job separations especially in the year of the M&A event, part of the effects on earnings may be coming from departures from their employers rather than within-firm decreases in earnings. This also helps explain the immediate drop in target firms' average payrolls in  $t = 0$  (Panel (b) of Figure 2), which reflects a large compositional change due to the significant worker departures. By contrast, we find that workers at acquiring firms do not experience any increased probability of a job transition relative to their control counterparts.

To study the impacts solely due to within-firm changes in compensation, Panel (c) of Figure 3 restricts the analysis to workers who stay in the same firm in the years following the event. We make this restriction for both treated workers and control workers so that the treatment group does not mechanically contain workers who have more stable job histories (Walker, 2013). Column (3) of Table 3 shows that annual earnings for stayers in M&A firms do not change much at either target or acquiring firms on average in the years following the event.<sup>26</sup>

By contrast, Panel (a) of Figure 4 shows large and immediate drops in earnings of workers who move from target firms to other firms within the first two years after the event. Column (1) of Table 4 shows that these job movers experience a decline of 4.0 log points in their earnings on average. Panel (b) also shows that roughly 70 percent of these job transitions happen in the year of the event ( $t = 0$ ), with the rest in the second year ( $t = 1$ ).<sup>27</sup> Note that the reason why the earning drops in  $t = 1$  even though most of these workers moved to other firms in  $t = 0$  is because the annual earning in  $t = 0$  reflects the earnings from both the original employer and the new employer, and most job transitions at  $t = 0$  happened in the late calendar year, meaning the annual earning mostly reflects the earning from the original firm on average. Importantly, about 80 percent of these workers move to non-acquiring firms, meaning that most of these job transitions are not driven by a simple reallocation of workers from targets to acquirers. Interestingly, as Panel (c) and Columns (3) – (4) show, target workers who move to non-acquiring firms experience larger and more immediate drops in their earnings on average, relative to workers who move to acquiring firms. Panel (d) and Column (5) show that the vast majority of these workers are moving to other firms within the same industry, implying that switching an occupation via moving to a different industry is unlikely the reason why these job movers experience a decline in their earnings. In the next section, we explore potential mechanisms behind the post-M&A decline in worker earnings.

<sup>26</sup>Note that the coefficient estimates on stayers' earnings at target firms are negative and statistically significant in years 4 and 5. This result seems to suggest that post-event changes in target firms' characteristics, such as decreases in profitability (Section 7), may help partly explain the decline in stayers' earnings as discussed in Section 4.

<sup>27</sup>We allow these workers to move to other firms in subsequent years after their first transition, which explains why there exists a small share of workers moving after the second year.

## 7 Potential Mechanisms and Economic Interpretations

In this section, we discuss and empirically test potential mechanisms, as discussed in detail in Section 4, behind the responses in workers' earnings after M&As.

### 7.1 Market Power Channel

In principle, both increased labor-market power and increased product-market power may impact worker earnings. For example, [Prager and Schmitt \(2021\)](#) and [Arnold \(2021\)](#) study how M&As that generate large shifts in concentration generate market-level declines in wages. Similarly, shifts in product-market power could impact earnings through two channels. First, if product-market power increases, firms may cut quantity in order to increase price. A decrease in labor demand in an industry may therefore lower wages. However, in models of rent-sharing, increases in product-market power will increase wages for workers in merging firms.<sup>28</sup>

In our setting, we find limited evidence for these channels. In particular, it is important to note that most M&A events (around 90 percent) happen across different labor markets in our setting, implying that there is a limited scope for a given M&A to increase concentration in a given market. Even for the M&A events that happen within the same labor market, the number of competitors in that market is so large initially that any given M&A does not lead to a meaningful rise in concentration in our setting. To get direct evidence on the market power channels, we turn to the firm balance-sheet data. For both increases in product-market power and labor-market power, firms' profits should increase after M&As. We use three different measures of profitability: profit margins, returns on assets, and EBITDA per worker. Panel (a) of Figure 5 shows the results on profit margins, separately for targets and for acquirers involved in acquisitions. Both targets' and acquirers' profitability were in a parallel trend with those of their matched control firms prior to the event. We find that while acquirers' profit margins decrease significantly after the event, targets' profit margins spiked in the year of M&A, and decreased afterwards. The initial increase in profit margins for targets is likely mechanical, due to a larger saving in fixed costs from downsizing relative to the initial fall in sales at the event year. Panel (d) shows the aggregate outcomes (targets and acquirers pooled) separately for those involved in acquisitions and for those involved in mergers. In both cases, we find that profit margins decreased after the event in the aggregate. We find similar results on return on assets (Panel (b) and (e)) and EBITDA per worker (Panel (c) and (f)).<sup>29</sup>

<sup>28</sup>[Kroft et al. \(2020\)](#) study a setting featuring both labor market power and product market power to quantify mark-downs of wages and markups of prices simultaneously.

<sup>29</sup>These results on firms' profitability remain qualitatively similar even if they are weighted by firm size (which can be provided upon request).

To interpret the magnitude of these results, Table 5 presents the difference-in-differences estimates on these outcomes. Columns (1) and (2) show that targets' profit margins decrease by 0.7 percentage points (albeit not statistically significant), while acquirers' profit margins decrease by 1.9 percentage points on average after the event. Columns (3) and (4) show that in the aggregate, profit margins decrease by 0.9 percentage points in the case of acquisitions, while they decrease by 1.9 percentage points in the case of mergers. Columns (1) and (2) show that targets' returns on assets decrease by 3.3 percentage points, while acquirers' returns on assets decrease by 1.3 percentage points on average. Columns (3) and (4) show that in the aggregate, returns on assets decrease by 2.7 percentage points in the case of acquisitions, while they decrease by 2.9 percentage points in the case of mergers. Columns (1) and (2) show that targets' EBITDA decreases by 2,887 CAD per worker, while acquirers' EBITDA decreases by 4,570 CAD per worker on average. Columns (3) and (4) show that in the aggregate, EBITDA decreases by 3,428 CAD per worker in the case of acquisitions, while it decreases by 7,106 CAD per worker in the case of mergers.

While efficiency gains through synergies are often argued as potential benefits of M&As, we do not find any evidence of increased profitability for either acquiring or target firms after M&As. This has potentially important implications for pro-merger arguments that rely on predicted synergy gains. This result might be surprising if M&A events are undertaken by profit-maximizing firms, but this is consistent with a corporate finance literature that finds negative effects of M&A events on firm performance or market valuation. These results also have important implications for the impact on workers. If M&As did increase productivity, we might expect some workers to benefit through rent sharing. In this setting, however, the lack of productivity gains suggests that workers may not benefit from the M&A events if there are no increased rents to share. Note that our results capture short to medium run outcomes, so firms' profitability may improve in the long run. Whether or not profits actually increase in the long run, these negative results on firm performance suggest that increased market power via M&As seems unlikely to drive our worker-level results.

In Appendix B, we perform a variety of heterogeneity analyses to provide more evidence against the market power channel. For example, we do not find larger declines in target workers' earnings when an M&A event happens within the same labor market or in the markets with a high initial level of concentration or in non-tradable sectors or in within-industry (horizontal) M&As. The overall takeaway from these results is that we find declines in worker earnings even in M&A events that are unlikely to have any impact on labor or product market power.<sup>30</sup> Given these results, we focus on workers who experience immediate and significant declines in their earnings after M&As: job movers from target firms. We focus on those who move within the first two years

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<sup>30</sup>Note that we continue to find a lack of support for increased market power explaining our results even when we focus on a subset of large firms whose mergers or acquisitions may have larger impacts on market concentration. The results from this subsample analysis may be provided upon request.

after the event because most job transitions from target firms happen within the first two years and are likely induced by the M&A event. Note that control workers who were matched to these job movers are also allowed to move to other firms after the event, but empirically, we find that most of these control workers stay at their original firm.

## 7.2 Impacts on Job Movers

In this section, we empirically assess mechanisms through which job transitions impact workers. A large body of evidence studies job transitions through mass layoffs. M&A events provide a different and interesting source of variation to study job transitions. Both the motivations and impacts of job transitions may be very different in this setting. For example, while we find similar patterns of earnings losses, the magnitude is much lower in our setting, with a long-run earnings loss of around 4 percent. By contrast, [Jacobson et al. \(1993\)](#) find a long-run earnings loss of 25 percent. Furthermore, M&A events tend to be more procyclical on average ([Rhodes-Kropf and Viswanathan, 2004](#)) in contrast to mass layoffs. As we will show later, this may help explain why job movers from target firms were able to move to larger and more profitable firms. Therefore, while M&A events are similar to mass layoffs in the sense that they both create job displacements, the mechanism through which M&As affect job movers' earnings could be different relative to the channel through which mass layoffs affect displaced workers' earnings.

We consider two broad classes of mechanisms. First, job transitions may cause impacts purely from the type of firms individuals are employed at. If M&A events cause workers to move from highly productive and large firms to low productivity and small firms, then these compositional impacts could rationalize the results. For example, [Lachowska et al. \(2020\)](#) and [Schmieder et al. \(forthcoming\)](#) seek to understand whether the characteristics of the firm the individual is employed at can explain long-term earnings losses due to mass layoffs. Similarly, in the next section, we test whether the firm plays a role in earnings losses of job movers after M&A events. In contrast to their studies, our work looks at firm characteristics from the balance-sheet data, such as profitability, to directly assess whether target workers are moving to less productive firms. We show that target workers move to larger and more profitable firms on average after M&As, therefore finding limited evidence for this channel.<sup>31</sup> Therefore, we next consider heterogeneity in the types of workers. We find significant heterogeneity that can be rationalized by models of the labor market in which there is a complementarity between the worker and the firm that is specific to a given employment relationship, as discussed in Section 4. Then we empirically assess which of these models seems most consistent with our overall findings.

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<sup>31</sup>We also find that target workers move to firms with higher firm-specific wage premiums (Appendix A).



### 7.2.1 Compositional Effects

To understand the role of the firm, we estimate an analogue to equation (3), substituting in as the outcome variable the observable firm characteristics, such as employment, total revenue, and profit margins. The goal is to assess whether the earnings losses following job transitions can be attributed to a mover's reemployment by an employer with a different size or profitability than the employer from which the mover left. Similar to Figure 4 and Table 4, we compare job movers from targets to their matched workers at control firms who may or may not move after the event.

Figure 5 shows changes in average firm characteristics of target workers who move to other firms after the M&A event. Again, we define a job transition as occurring in time  $t$  if the worker is in a different firm at  $t$  relative to  $t - 1$ . Therefore, if a M&A occurs in time  $t = 0$ , and the worker appears in a new firm at  $t = 1$ , then we define the job transition as occurring at  $t = 1$ . This is why we observe a large spike in job transitions at  $t = 1$  in the data. Because of the tenure restriction, any change in pre-event firm characteristics is driven by yearly changes in target firms' characteristics. The change in firm characteristics in event year ( $t = 0$ ) still reflects the change in target firms' characteristics, given that the first job transition is reflected at  $t = 1$ . Starting from  $t = 1$ , changes in firm characteristics reflect both changes in new employers where target workers moved, and changes in target firms of workers who had not left yet, relative to changes in firm characteristics of their matched control workers. Figure 5 shows that workers transition to firms with more employees, higher sales, and higher profit margins. Table 5 presents the difference-in-differences results summarizing these impacts, finding workers from target firms move to firms with higher employment (51.8 log points), higher revenue (49.9 log points) and higher profit margins (1.8 percentage points).

To summarize, we find that workers who transition jobs after M&A events move to better firms along a number of dimensions. Despite this, their actual earnings are 4.0 log points lower on average. In the next section we explore potential mechanisms for this earnings decline.

### 7.2.2 Heterogeneity by Workers

In this section, we explore heterogeneity by workers in order to understand who is most impacted by these job transitions. This will allow us to understand various theoretical mechanisms through which workers are affected by M&As. In particular, it will allow us to understand why workers still suffer earnings losses despite moving to on-average larger and more profitable firms.

First, we consider heterogeneity by the worker's tenure. As discussed in Section 4, there are two conceptual models that would predict impacts on earnings increasing in tenure. First, a model

of firm-specific human capital (i.e., Lazear 2009) predicts that high-tenure workers would lose more from moving after an M&A event because the high-tenure worker’s productivity is much lower at other firms. Second, a model of backloaded contracts (i.e., Lazear 1979) also predicts high-tenure workers would lose more by moving to other firms, as these workers currently enjoy wages above their marginal product at the current employer.

To study impacts by tenure, we divide workers into two groups: those with 4 years of tenure (shorter tenure), and those with 7 or more years of tenure (longer tenure).<sup>32</sup> In Figure 6, for individuals with 4 years of tenure we find small, short-run negative impacts that quickly disappear 2 years after the event. Individuals with 7 or more years of tenure experience a large decline in their earnings after M&As. As shown in Table 6, on average, they suffer earnings losses of about 4.5 percent relative to workers with 4 years of tenure whose earnings stay similar on average.<sup>33</sup>

These results are consistent with models of the labor market that include some complementarity between the worker and the firm. To understand this channel further, we next exploit heterogeneity by within-firm earnings quintiles. Intuitively, both the productivity and bargaining channels may be more relevant for highly-skilled labor. For example, in hospital mergers, Prager and Schmitt (2021) finds impacts for nurses, but no impacts on unskilled workers such as cafeteria workers, whose skills were not tied to the hospital industry. In our context, we do not observe the worker’s occupation. Instead, we use within-firm earnings quintiles to capture a crude measure of the complexity of the job. Additionally, this analysis is also informative about distributional consequences of M&A events.

In Figure 7 we plot the impacts of M&As for individuals at the bottom quintile of the earnings distribution versus individuals at the top quintile of the earnings distribution.<sup>34</sup> For workers at the top quintile, we see a 4.6 log points larger decline in earnings when they move to other firms after the event, relative to job movers at the bottom quintile of the earnings distribution, as shown in Column (1) of Table 7. These differences are not due to the types of firms these workers move to.<sup>35</sup>

As discussed above, these results are consistent with either the worker suffering losses due to a loss of firm-specific human capital or the worker suffering losses due to a backloaded contract that pays higher wages later in an employment spell. Conceptually, one way to distinguish these two

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<sup>32</sup>For this analysis, we implement another matching strategy to ensure control workers are in the same tenure bin as the target workers. We find qualitatively similar results when we use a different cut to define short tenure and longer tenure workers. Results based on a different cut can be provided upon request.

<sup>33</sup>Furthermore, both short-tenure and long-tenure workers are moving to better-paying firms on average relative to their original firm. These results can be provided upon request.

<sup>34</sup>For this analysis, we implement the matching strategy again to ensure control workers are in the same quintile of the earnings distribution within their firm as the target workers.

<sup>35</sup>Similar to the results by workers’ tenure, both bottom-quintile workers and top-quintile workers are moving to better-paying firms on average. These results can be provided upon request.

would be to estimate worker productivity. This would allow one to compare the productivity of the worker to her wage, and therefore we would be able to understand whether a worker is overpaid.<sup>36</sup> However, in practice, even at a conceptual level, it is difficult to measure worker-level productivity for most jobs.

While we cannot observe worker-level productivity, one observable that is informative for this discussion is firms' profitability. Note that if firms are firing workers for whom  $p_{ij} - w_{ij} - O_j < 0$  then these firms should see increases in profit margins. Recall the simple model in Section 4. In this case, the outside option  $O_j$  is greater than the profits from worker  $i$  ( $p_{ij} - w_{ij}$ ). However, as shown earlier, we no evidence of increased profitability for either acquiring firms or target firms following the M&A event. Therefore, in order to rationalize the worker-level impacts through firing of overpaid workers, one would need to believe that there are multiple off-setting decisions being made by new management. In other words, in this story, management is laying off unproductive workers, so that overall productivity should increase. But they are not realizing any gains from this, so they must be making different, productivity-reducing decisions as well to rationalize the overall decline in profitability. M&As are of course heterogeneous events, so multiple channels may be at play.

While the decline in firms' profits alone is not definitive evidence against the backloaded contract channel, there is another result that potentially rules it out. As discussed in Section 4, we should expect to see larger job transitions among high-tenure and high-wage workers if overpaid workers are displaced through M&As, but we find similar changes in job transitions among short-tenure (or low-wage) workers. This result, combined with the finding on decreased profits, suggests that the earnings losses among workers who leave target firms after M&As are most likely driven by a loss of firm-specific human capital.

To summarize, we find that the decrease in earnings in our setting is primarily driven by target workers who transition jobs after an M&A event and move to different firms. While some M&A events may increase labor market or product market power, we find the number of such events is limited in our setting. Workers who transition to new firms after M&A events move to firms that have better observable characteristics along a number of dimensions. These firms have higher revenue, employment, and profit margins. Despite this, workers who transition to these firms still experience decreases in earnings overall. Our results are consistent with a story in which a complementarity between the worker and firm that results in higher wages is destroyed through M&As.

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<sup>36</sup>An alternative explanation could be that target workers move to larger and more profitable firms and accept lower wages due to non-wage amenities, such as geographical proximity or job stability. However, it is also likely that low-tenure or low-wage workers have preferences for non-wage amenities (i.e., good locations or job security), so this channel would be inconsistent with our heterogeneity results based on worker tenure or prior earnings.

## 8 Conclusion

In this paper, we use linked employer-employee data to connect the impact of M&A events on firms to the impact on their employees. Previous research has looked at the financial impact on firms and the impact on workers' outcomes separately, but our paper is the first to link these impacts directly using our administrative data from tax records on both firms and workers. This allows us to look deeper into firm-based mechanisms than has been possible in previous research.

Using a matched difference-in-differences design, we compare firms that went through an M&A to matched control firms of the same size bins, province, and sector. Our results show that acquiring firms expand, but target firms shrink significantly, without much change in the aggregate (targets and acquirers pooled). For workers at target firms, we find that their earnings decline and job separations increase significantly after an M&A event.

Leveraging detailed firm balance sheet data, we investigate different potential mechanisms for the decline in workers' earnings at target firms. Since we see no meaningful change in labor market concentration after M&As, we conclude that a rise in market concentration is unlikely the key mechanism behind post-M&A wage declines. Furthermore, since profit margins decrease at both targets and acquirers, a rise in product market power is unlikely to have driven the decline in worker earnings either. Instead, we find that the decrease in earnings of workers at targets is largely driven by those who move to other firms after the M&A event. While these workers move to larger firms with higher profit margins on average, they experience a decline in their earnings likely due to a loss of firm-specific human capital.

Our findings provide important context for research investigating the labor market consequences of corporate M&As. Whichever mechanism is under investigation, care should be taken to account for how firm-level outcomes, especially their profitability and growth, as well as worker-level outcomes, may change after M&As. Our results also matter for public policy. In Canada, recent changes to the Competition Act targeted both the impact of M&As on labor and how potential efficiency gains are weighed in merger decisions. Similar debates are happening in other countries, including the United States ([Naidu et al. 2018](#)). Our results provide new evidence of the negative impact of M&As on wages that add perspective to these policy debates, and our findings on post-M&A firm performance raise doubts about the efficiency arguments made in support of M&As.

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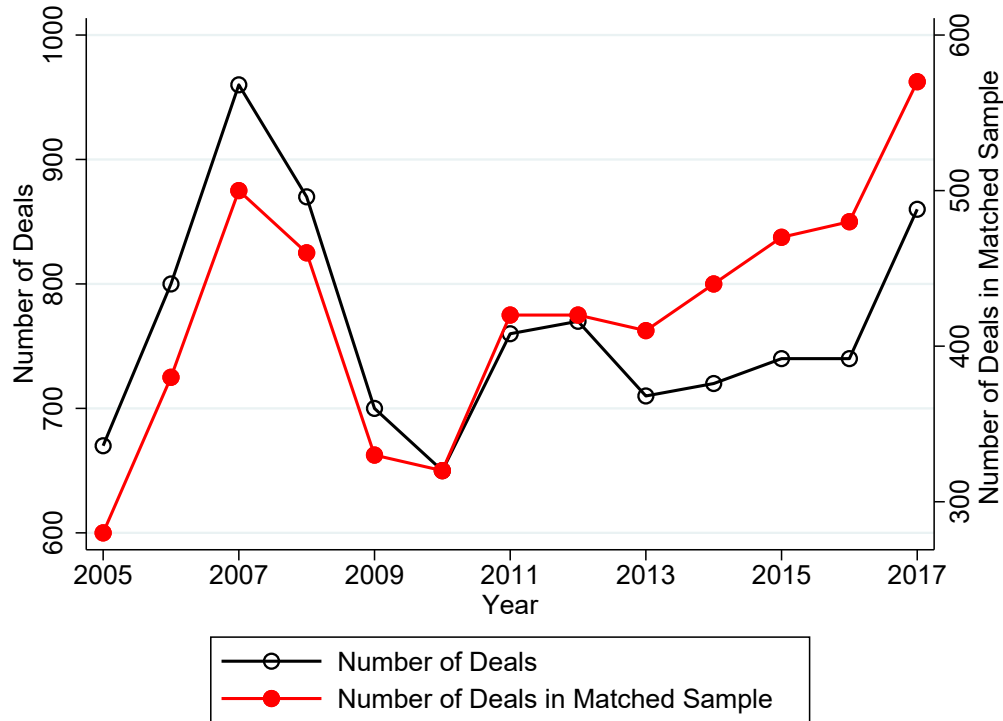
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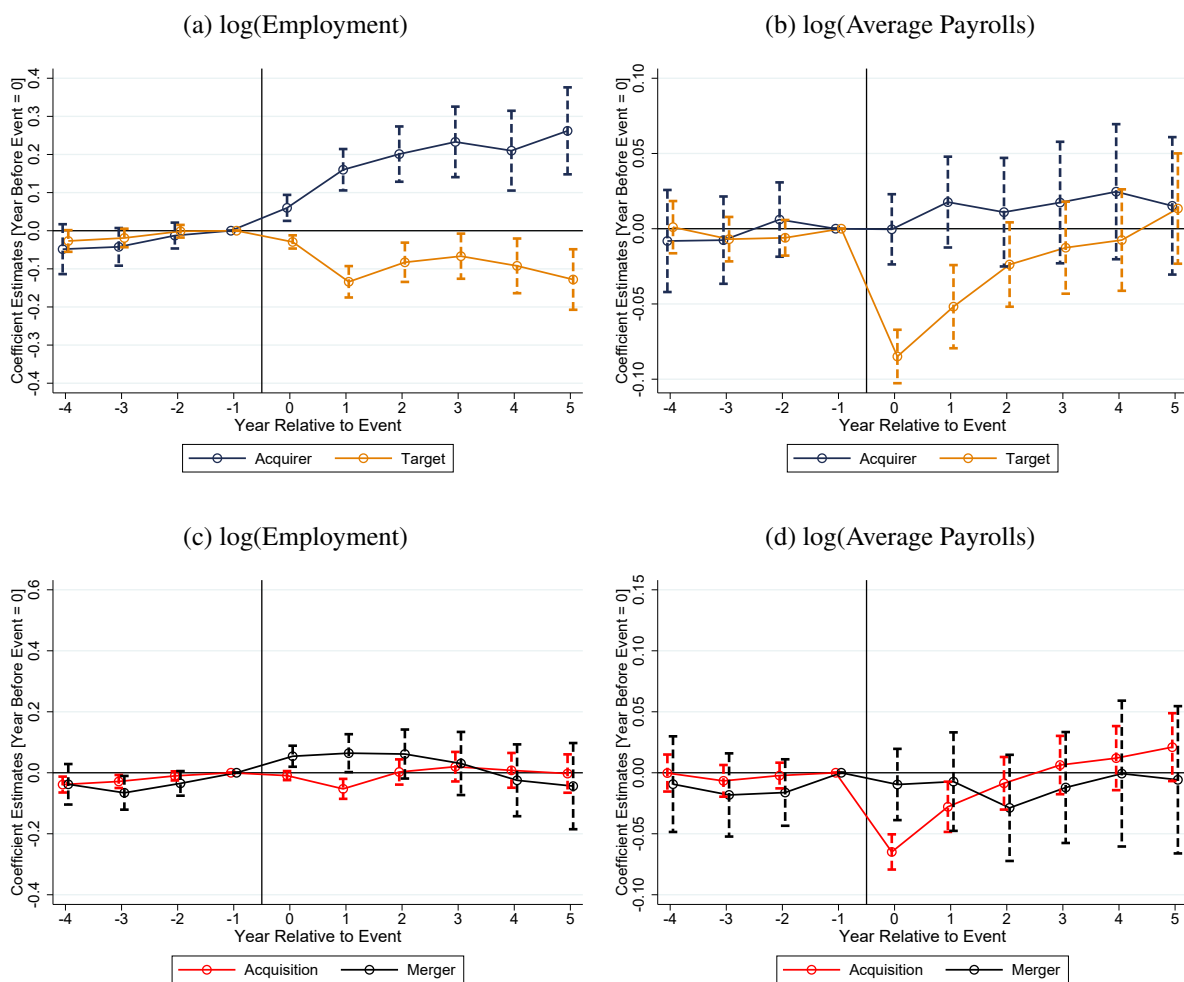
**Woodcock, Simon D.**, “Match effects,” *Research in Economics*, 2015, 69 (1), 100–121.

Figure 1: Number of M&A Events



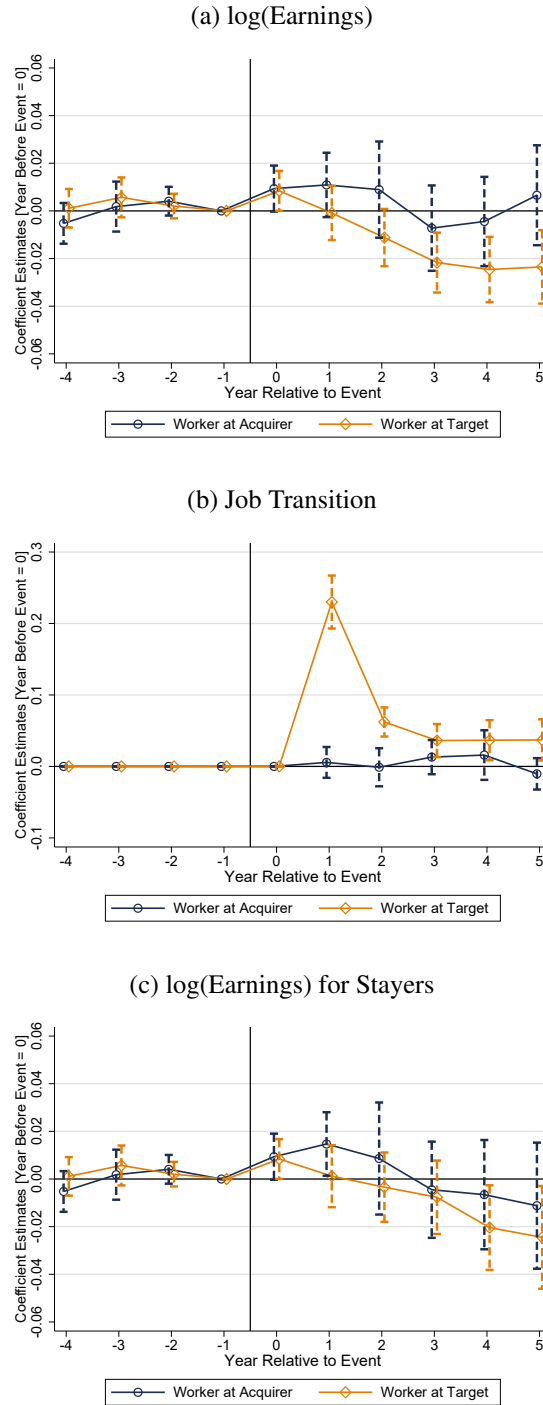
*Notes:* During our sample period, the number of M&A events is 765 per year (including multiple events for a given firm) on average among the sample of firms eligible for matching as explained in Section 5 (black line). Among these eligible firms, the number of M&A events in the matched sample is 422 per year (including multiple events for a given firm) on average (red line). Section 5 describes how we construct our matched sample of firms. During our sample period, the unique number of firms eligible for matching is 5,200. Among these firms, the unique number of matched firms (both targets and acquirers) is 4,100. Therefore, the matching strategy finds a counterfactual firm in about 79 percent of all cases among this sample of eligible firms.

Figure 2: Employment and Average Payrolls After M&As



*Notes:* These figures display event-study estimates for the impact of M&As on employment and average payrolls. Panel (a) and (b) show the estimates for the impact of acquisitions, separately for acquiring firms (navy line) and for target firms (orange line). Panel (c) and (d) show the estimates on the aggregate firm-level (targets and acquirers pooled) outcomes, separately for those involved in acquisitions (red line) and for those involved in mergers (black line). The dashed lines indicate 95 percent confidence intervals where the standard errors are clustered at the firm level. The M&A event is in year 0 and the coefficient is normalized to be zero in year -1.

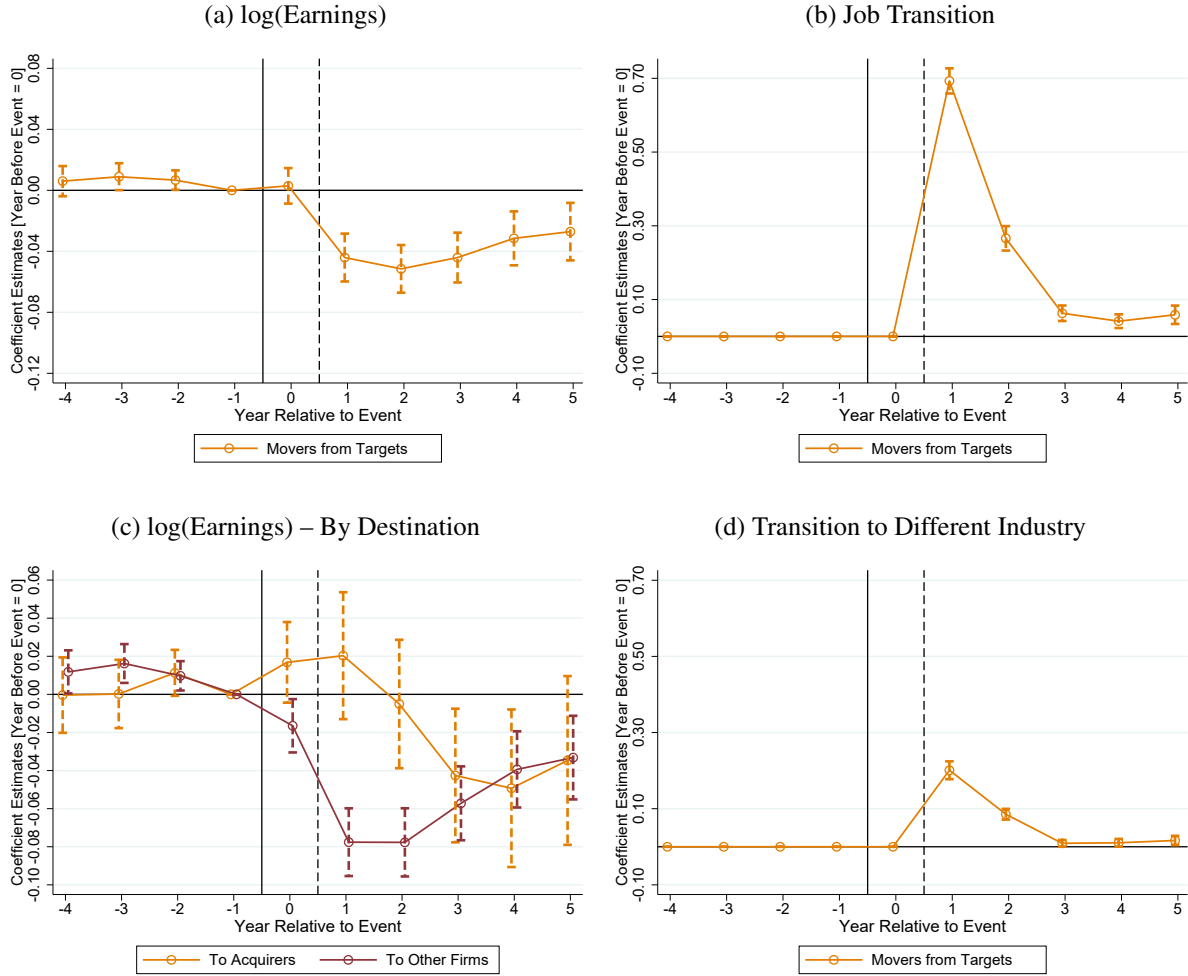
Figure 3: Worker Earnings and Job Transitions After M&As



*Notes:* These figures display event-study estimates for the impact of M&As on worker-level outcomes, separately for workers at acquiring firms (navy line) and for workers at target firms (orange line). Panel (a) shows the estimates for log of total earnings. Panel (b) shows the estimates for job transition probabilities. Panel (c) shows the estimates for log of total earnings for firm stayers. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker and firm level. The M&A event is in year 0 and the coefficient is normalized to be zero in year -1.

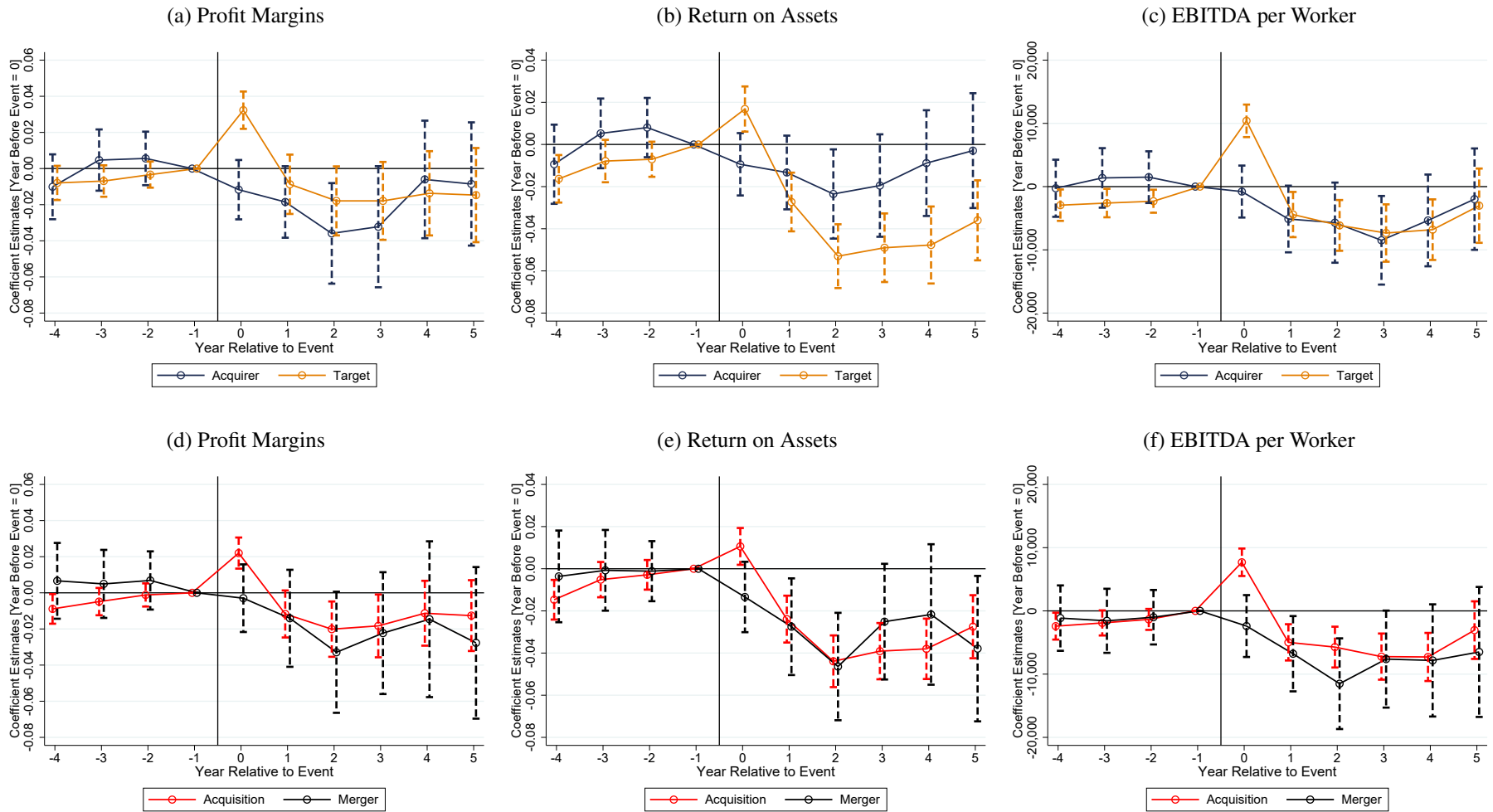


Figure 4: Workers Moving from Targets



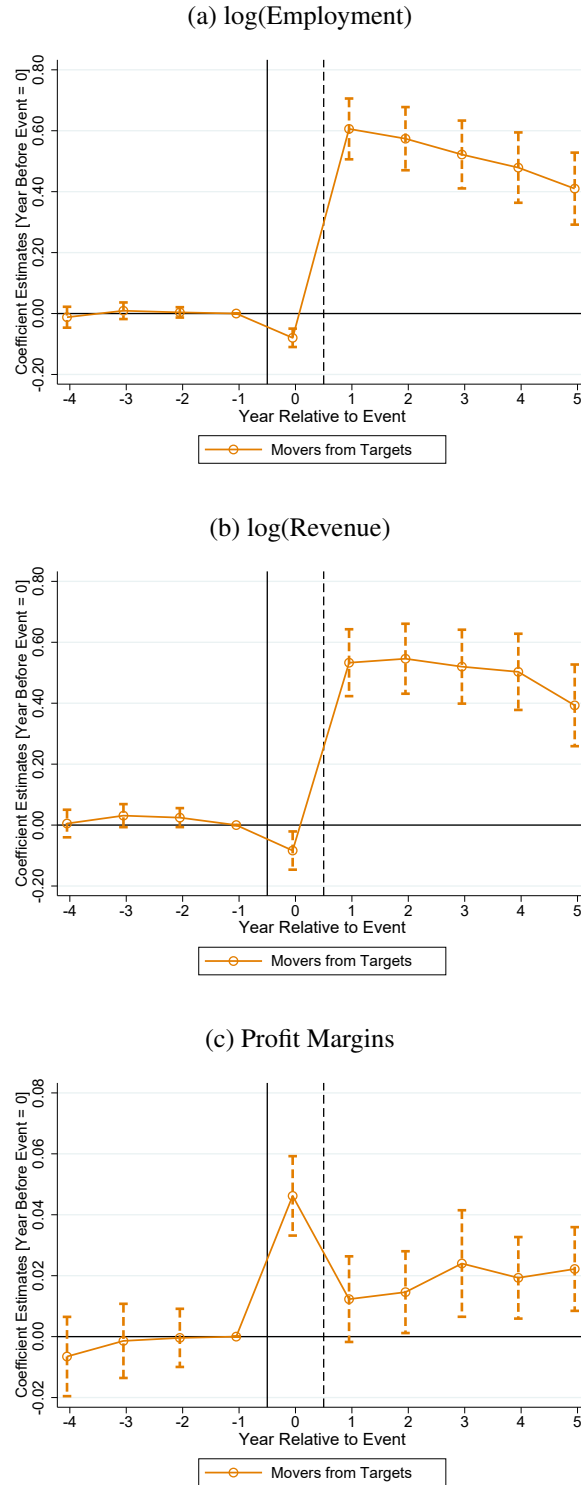
*Notes:* These figures display event-study estimates for the impact of M&As on workers moving from target firms within the first two years after the event, as described in Section 6. Panel (a) shows the estimates for log of total earnings. Panel (b) shows the estimates for job transition probabilities. Panel (c) shows the estimates for log of total earnings based on their destination. A small share of workers also moves within their original parent company; however, we do not observe any impact for these workers, so we do not report their estimates here. Panel (d) shows the estimates for the probability of transition to a different industry. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker and firm level. The M&A event is in year 0 and the coefficient is normalized to be zero in year -1.

Figure 5: Firms' Profitability After M&As



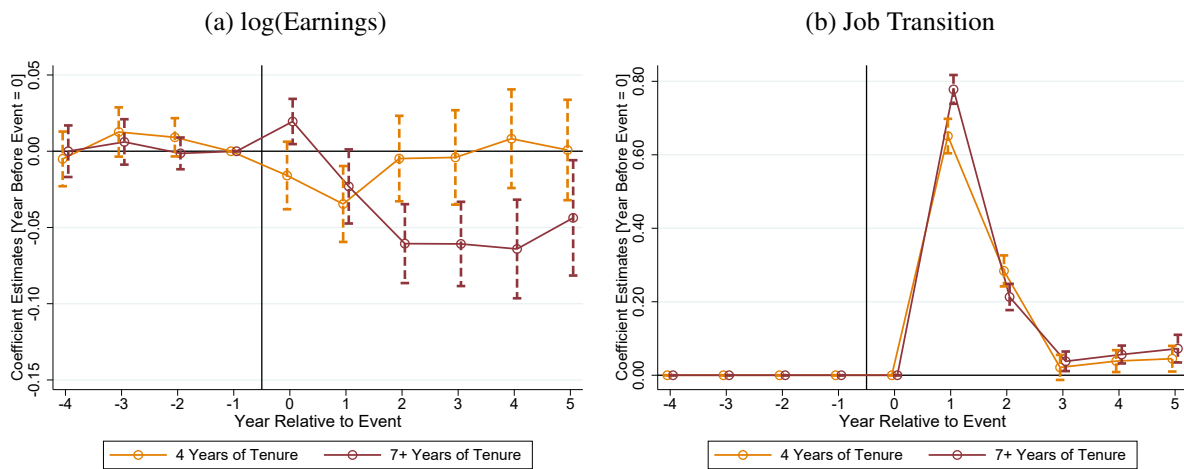
*Notes:* Panels (a) – (c) display event-study estimates for the impact of M&As on firm-level profitability for firms involved in acquisitions, separately for acquiring firms (navy line) and for target firms (orange line). Panel (a) shows the estimates for profit margins. Panel (b) shows the estimates for return on assets. Panel (c) shows the estimates for EBDITA per worker. Panels (d) – (f) display event-study estimates for the impact of M&As on the aggregate firm-level (targets and acquirers pooled) profitability, separately for those involved in acquisitions (red line) and for those involved in mergers (black line). Panel (d) shows the estimates for profit margins. Panel (e) shows the estimates for return on assets. Panel (f) shows the estimates for EBDITA per worker. The dashed lines indicate 95 percent confidence intervals where the standard errors are clustered at the firm level. The M&A event is in year 0 and the coefficient is normalized to be zero in year -1.

Figure 6: Firm Characteristics of Workers Moving from Targets



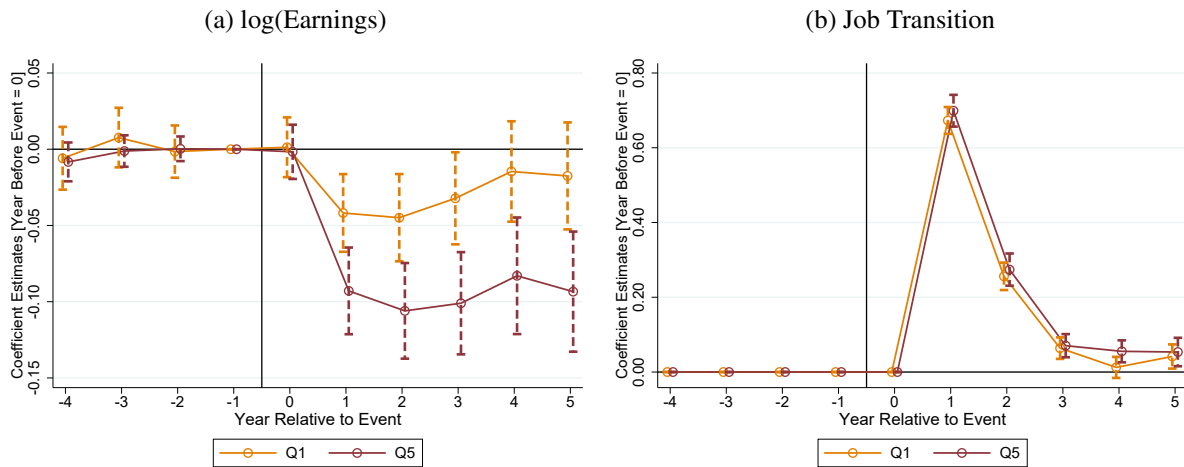
*Notes:* These figures display event-study estimates for changes in average firm characteristics of workers moving from target firms within the first two years after the event, as described in Section 6. Panel (a) shows the estimates for log of employment. Panel (b) shows the estimates for log of total revenue. Panel (c) shows the estimates for profit margins. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker and firm level. The M&A event is in year 0 and the coefficient is normalized to be zero in year -1.

Figure 7: Workers Moving from Targets – By Worker Tenure



*Notes:* These figures display event-study estimates for the impact of M&As for workers moving from target firms within the first two years after the event, separately for those with 4 years of tenure and for those with 7 or more years of tenure by one year before the event. Panel (a) shows the estimates for log of total earnings. Panel (b) shows the estimates for job transition probabilities. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker and firm level. The M&A event is in year 0 and the coefficient is normalized to be zero in year -1.

Figure 8: Workers Moving from Targets – By Within-Firm Earnings Distribution



*Notes:* These figures display event-study estimates for the impact of M&As for workers moving from target firms within the first two years after the event, separately for those at the bottom quintile (Q1) and for those at the top quintile (Q5) of the within-firm earnings distribution. Panel (a) shows the estimates for log of total earnings. Panel (b) shows the estimates for job transition probabilities. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker and firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Table 1: Descriptive Statistics on the Matched Sample of Firms and Workers

|  | (1)<br>Acquirer | (2)<br>Control | (3)<br>Target | (4)<br>Control |
|--|-----------------|----------------|---------------|----------------|
| <i>Panel A: Firm Characteristics</i>   |                 |                |               |                |
| Total Revenue (in millions)            | 62              | 47             | 34            | 29             |
| Total Expenses (in millions)           | 58              | 44             | 32            | 28             |
| Profit Margins                         | 0.02            | 0.05           | 0.02          | 0.05           |
| Number of Employees                    | 187             | 125            | 106           | 87             |
| Average Wage Bill                      | 52,472          | 51,956         | 50,380        | 49,175         |
| Leverage Ratio                         | 0.65            | 0.67           | 0.7           | 0.65           |
| Markups                                | 1.88            | 1.68           | 1.88          | 1.68           |
| Number of Firms                        | 1,040           | 1,040          | 3,060         | 3,060          |
| <i>Panel B: Sectors (Firms)</i>        |                 |                |               |                |
| Construction                           | 0.06            | 0.06           | 0.05          | 0.05           |
| Manufacturing                          | 0.23            | 0.23           | 0.26          | 0.26           |
| Wholesale                              | 0.15            | 0.15           | 0.14          | 0.14           |
| Retail                                 | 0.03            | 0.03           | 0.05          | 0.05           |
| Transportation                         | 0.03            | 0.03           | 0.04          | 0.04           |
| Information                            | 0.04            | 0.04           | 0.04          | 0.04           |
| Services                               | 0.33            | 0.33           | 0.32          | 0.32           |
| Other Sectors                          | 0.13            | 0.13           | 0.11          | 0.11           |
| <i>Panel C: Worker Characteristics</i> |                 |                |               |                |
| Total Earnings                         | 70,046          | 71,386         | 70,625        | 72,317         |
| Age                                    | 46.7            | 46.6           | 47.2          | 47.0           |
| Female                                 | 0.33            | 0.33           | 0.31          | 0.31           |
| Number of Workers                      | 42,780          | 42,780         | 64,520        | 64,520         |
| <i>Panel D: Sectors (Workers)</i>      |                 |                |               |                |
| Construction                           | 0.03            | 0.03           | 0.02          | 0.02           |
| Manufacturing                          | 0.49            | 0.49           | 0.5           | 0.5            |
| Wholesale                              | 0.16            | 0.16           | 0.1           | 0.1            |
| Retail                                 | 0.01            | 0.01           | 0.03          | 0.03           |
| Transportation                         | 0.03            | 0.03           | 0.03          | 0.03           |
| Information                            | 0.02            | 0.02           | 0.02          | 0.02           |
| Services                               | 0.19            | 0.19           | 0.21          | 0.21           |
| Other Sectors                          | 0.05            | 0.05           | 0.08          | 0.08           |

*Notes:* This table reports descriptive statistics on the matched sample of firms and workers, measured one year prior to the event. Panel A reports firm characteristics such as total revenue, total expenses, profit margins, number of employees, average payrolls, leverage ratio, and markups. Columns (1) and (3) report these statistics for acquiring firms and for target firms, respectively, and column (2) and (4) report these statistics for their respective matched control firms. Panel B reports the distribution of firms in the matched sample across 2-digit NAICS sectors. Panel C reports worker characteristics such as total annual earnings, age, and gender. Columns (1) and (3) report these statistics for workers at acquirers and for workers at targets, respectively, and column (2) and (4) report these statistics for their respective matched control workers. Panel D reports the distribution of workers in the matched sample across 2-digit NAICS sectors. Other sectors include (1) Agriculture, forestry, and fishing, (2) Mining, quarrying, and oil and gas extraction, (3) Utilities, (4) Real estate and rental and leasing, (5) Arts, entertainment and recreation, (6) Accommodation and food services, (7) Other services, and (8) Public administration.



Table 2: Difference-in-differences Estimates on Firm Outcomes

|                       | (1)                  | (2)                 | (3)               | (4)               |
|-----------------------|----------------------|---------------------|-------------------|-------------------|
|                       | <b>Acquisition</b>   |                     | <b>Aggregate</b>  |                   |
|                       | Target               | Acquirer            | Acquisition       | Merger            |
| log(Employment)       | -0.089***<br>(0.023) | 0.188***<br>(0.033) | -0.006<br>(0.018) | 0.024<br>(0.038)  |
| Mean at t = -1        | 4.00                 | 4.55                | 4.14              | 4.02              |
| Adj. R squared        | 0.875                | 0.891               | 0.879             | 0.880             |
| Firm-Year Obs.        | 66,680               | 21,930              | 89,230            | 19,100            |
| log(Average Payrolls) | -0.028***<br>(0.011) | 0.014<br>(0.014)    | -0.010<br>(0.008) | -0.011<br>(0.017) |
| Mean at t = -1        | 10.68                | 10.73               | 10.70             | 10.73             |
| Adj. R squared        | 0.795                | 0.817               | 0.798             | 0.799             |
| Firm-Year Obs.        | 66,590               | 21,870              | 89,080            | 19,040            |

*Notes:* This table reports the difference-in-differences estimates for the impact of M&As on employment and average payrolls. Columns (1) and (2) report the estimates for the impact of acquisitions on target firms and on acquiring firms, respectively. Columns (3) and (4) report the estimates on the aggregate firm-level (targets and acquirers pooled) outcomes for those involved in acquisitions and for those involved in mergers, respectively. The standard errors are clustered at the firm level.

Table 3: Difference-in-differences Estimates on Worker Outcomes

|                     | (1)                 | (2)                 | (3)                     |
|---------------------|---------------------|---------------------|-------------------------|
|                     | log(Earnings)       | Transition          | log(Earnings) - Stayers |
| Workers at Target   | -0.012**<br>(0.005) | 0.067***<br>(0.007) | -0.008<br>(0.006)       |
| Mean at t = -1      | 11.01               | 0.00                | 11.01                   |
| Adj. R squared      | 0.739               | 0.187               | 0.797                   |
| Worker-Year Obs.    | 2,023,130           | 2,0264,40           | 1,581,880               |
| Workers at Acquirer | 0.004<br>(0.007)    | 0.004<br>(0.006)    | 0.002<br>(0.008)        |
| Mean at t = -1      | 11.02               | 0.00                | 11.02                   |
| Adj. R squared      | 0.732               | 0.171               | 0.786                   |
| Worker-Year Obs.    | 1,343,370           | 1,345,330           | 1,084,030               |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As on workers' outcomes, separately for those at acquiring firms and for those at target firms. The dependent variables in columns (1) and (2) are log of total earnings and the probability of job transitions. Column (3) reports the estimates on log of total earnings for firm stayers. The standard errors are two-way clustered at the worker and firm level.

Table 4: Workers Moving from Targets

|                             | (1)                                   | (2)                 | (3)               | (4)                  | (5)                 |
|-----------------------------|---------------------------------------|---------------------|-------------------|----------------------|---------------------|
|                             | <b>log(Earnings) - By Destination</b> |                     |                   |                      |                     |
|                             | log(Earnings)                         | Transition          | To Acquirer       | To Other Firms       | To Diff Industry    |
| Workers Moving from Targets | -0.040***<br>(0.007)                  | 0.224***<br>(0.006) | -0.016<br>(0.013) | -0.050***<br>(0.007) | 0.064***<br>(0.004) |
| Mean at t = -1              | 10.98                                 | 0.00                | 10.96             | 10.93                | 0.00                |
| Adj. R squared              | 0.739                                 | 0.317               | 0.765             | 0.736                | 0.167               |
| Worker-Year Obs.            | 689,900                               | 691,120             | 144,550           | 430,750              | 691,120             |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As for workers moving from target firms within the first two years after the event. Column (1) displays the estimates for log of total earnings. Column (2) displays the estimates for the job transition probabilities. Columns (3) and (4) display the estimates for log of total earnings based on workers' destination (to acquiring firms and to other firms, respectively). A small share of workers also moves within their original parent company; however, we do not observe any impact for these workers, so we do not report the estimates here. Column (5) displays the estimates for the probability of transition to a different industry. The standard errors are two-way clustered at the worker and firm level.

Table 5: Difference-in-differences Estimates on Firms' Profitability

|                   | (1)                      | (2)                       | (3)                        | (4)                        |
|-------------------|--------------------------|---------------------------|----------------------------|----------------------------|
|                   | <b>Acquisition</b>       |                           | <b>Aggregate</b>           |                            |
|                   | Target                   | Acquirer                  | Acquisition                | Merger                     |
| Profit Margins    | -0.007<br>(0.007)        | -0.019**<br>(0.009)       | -0.009*<br>(0.005)         | -0.019<br>(0.012)          |
| Mean at t = -1    | 0.02                     | 0.02                      | 0.02                       | 0.00                       |
| Adj. R squared    | 0.356                    | 0.355                     | 0.352                      | 0.380                      |
| Firm-Year Obs.    | 68,910                   | 22,200                    | 91,730                     | 19,520                     |
| Return on Assets  | -0.033***<br>(0.006)     | -0.013<br>(0.008)         | -0.027***<br>(0.005)       | -0.029***<br>(0.010)       |
| Mean at t = -1    | 0.08                     | 0.05                      | 0.07                       | 0.04                       |
| Adj. R squared    | 0.384                    | 0.420                     | 0.387                      | 0.433                      |
| Firm-Year Obs.    | 69,900                   | 22,390                    | 92,900                     | 19,690                     |
| EBITDA per worker | -2886.774*<br>(1484.173) | -4570.210**<br>(2329.311) | -3427.987***<br>(1204.799) | -7105.831***<br>(2730.995) |
| Mean at t = -1    | 15,880.58                | 17,906.32                 | 16,375.09                  | 14,431.66                  |
| Adj. R squared    | 0.441                    | 0.458                     | 0.436                      | 0.449                      |
| Firm-Year Obs.    | 65,810                   | 21,600                    | 88,050                     | 18,690                     |

*Notes:* This table reports the difference-in-differences estimates for the impact of M&As on different measures of firms' profitability. Columns (1) and (2) report the estimates for target firms and for acquiring firms involved in acquisitions, respectively. Columns (3) and (4) report the estimates on the aggregate firm-level (targets and acquirers pooled) outcomes, for those involved in acquisitions and for those involved in mergers, respectively. The standard errors are clustered at the firm level.

Table 6: Changes in Average Firm Characteristics of Workers Moving from Targets

|                  | (1)             | (2)          | (3)            |
|------------------|-----------------|--------------|----------------|
|                  | log(Employment) | log(Revenue) | Profit Margins |
| Workers Moving   | 0.518***        | 0.499***     | 0.018***       |
| From Targets     | (0.052)         | (0.055)      | (0.007)        |
| Mean at t = -1   | 5.74            | 18.22        | 0.05           |
| Adj. R squared   | 0.790           | 0.803        | 0.466          |
| Worker-Year Obs. | 688,280         | 649,220      | 650,510        |

*Notes:* This table reports the difference-in-differences estimates for changes in average firm characteristics of workers moving from target firms within the first two years after the event. Column (1) displays the estimates for log of employment. Column (2) displays the estimates for log of revenue. Column (3) displays the estimates for profit margins. The standard errors are two-way clustered at the worker and firm level.

Table 7: Workers Moving from Targets – By Worker Tenure

|   | (1)                  | (2)                 |
|---|----------------------|---------------------|
|   | log(Earnings)        | Transition          |
| Post $\times$ Treated                             | 0.006<br>(0.011)     | 0.180***<br>(0.008) |
| Post $\times$ Treated $\times$ 7+ Years of Tenure | -0.045***<br>(0.014) | 0.012<br>(0.010)    |
| Mean at t = -1 (4 Years of Tenure)                | 10.96                | 0.00                |
| Mean at t = -1 (7+ Years of Tenure)               | 11.03                | 0.00                |
| Adj. R squared                                    | 0.745                | 0.404               |
| Worker-Year Obs. (4 Years of Tenure)              | 186,170              | 186,660             |
| Worker-Year Obs. (7+ Years of Tenure)             | 251,020              | 251,340             |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As for workers moving from targets within the first two years after the event, separately for those with 4 years of tenure measured one year prior to the event. The triple interaction term captures the triple-difference estimates for those with 7 or more years of tenure measured one year prior to the event. Column (1) displays the estimates for log of total earnings. Column (2) displays the estimates for the job transition probabilities. The standard errors are two-way clustered at the worker and firm level.

Table 8: Workers Moving from Targets – By Within-Firm Earnings Distribution

|                                   | (1)                 | (2)                 |
|-----------------------------------|---------------------|---------------------|
|                                   | log(Earnings)       | Transition          |
| Post $\times$ Treated             | -0.028**<br>(0.011) | 0.181***<br>(0.006) |
| Post $\times$ Treated $\times$ Q5 | -0.046**<br>(0.021) | 0.007<br>(0.005)    |
| Mean at t = -1 (Q1 = 1)           | 10.45               | 0.00                |
| Mean at t = -1 (Q5 = 1)           | 11.56               | 0.00                |
| Adj. R squared                    | 0.745               | 0.324               |
| Worker-Year Obs. (Q1 = 1)         | 126,710             | 126,980             |
| Worker-Year Obs. (Q5 = 1)         | 119,950             | 120,160             |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As for workers moving from targets within the first two years after the event, separately for those at the bottom quintile (Q1) of the within-firm earnings distribution measured one year prior to the event. The triple interaction term captures the triple-difference estimates for workers at the top quintile (Q5) of the within-firm earnings distribution. Column (1) displays the estimates log of total earnings. Column (2) displays the estimates for the probability of transition. The standard errors are two-way clustered at the worker-firm level.



## ONLINE APPENDIX:

### Job Transitions and Employee Earnings After Acquisitions: Linking Corporate and Worker Outcomes

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## **A Robustness Checks**

In Appendix A, we provide results from robustness tests discussed in Sections 5 – 7.

### **A.1 Different Clustering**

Our main firm-level results are based on clustering at the firm level and our main worker-level results are based on two-way clustering at both firm level and worker level. We also do robustness tests on the key firm-level and worker-level outcomes, where standard errors are clustered at the market level (defined at the four digit NAICS by commuting zone) for firm-level results, and are two-way clustered at the worker and market level for worker-level results. Figure A1 and Table A1 show that the results on employment and average payrolls, and worker-level earnings are similar to the main estimates. Note that the coefficient estimates on these outcomes are slightly different from our main estimates, even though we only change the way we cluster our standard errors. This is because some firms and workers have missing commuting zone information, so they are dropped from our analysis sample when we cluster the standard errors at the market level.

### **A.2 Using Matched Control Firms in Different Markets**

Matching on size, sector, and province finds firms that would plausibly exhibit common trends in the absence of an M&A activity. However, it is possible that firms can be matched within the same market (defined at the four-digit NAICS industry by commuting zone), which is potentially concerning if M&As have impacts on local labor markets through increased concentration. If M&As have negative effects on firms in the same market, then the impact of M&As on firms and workers will be biased towards zero. To minimize this concern, we do a robustness check by matching firms within the same province, but across different markets, and find similar results to our main results where we allow M&A firms to be matched with control firms within the same market. Figure A2 and Table A2 show that the effects on the key firm-level and worker-level outcomes from this approach are qualitatively similar to our main estimates.

### **A.3 Outcomes in Levels (Replacing Missing Observations with Zeros)**

We show results on the key outcomes, such as employment, average payrolls, and worker-level earnings, in logs. A potential concern with this approach is that we cannot account for firms and workers exiting the sample after M&As. This concern is especially relevant for target firms and

for displaced workers from targets that might exit our sample. Although we find the vast majority of target firms continue to operate independently and most displaced workers find a job at a different firm, we run our main analyses in levels, replacing missing observations with zeros, to account for a small share of firms and workers who exit the sample after M&As. Figure A3 and Table A3 show that the main results are qualitatively robust to the specification where the outcomes are measured in levels.

#### **A.4 By One-time vs. Repeat Acquirers**

Prior research points out that a part of the motive behind M&As involves empire-building, which could result in losses in efficiency and profitability after the event (Jensen, 1986). Even though it is practically difficult to discern whether a particular acquirer has an empire-building motive in our data, we test whether the effects on firm size and average payrolls are different depending on whether an acquirer engages in multiple M&A transactions (roughly 30 percent of our analysis sample). Specifically, we compare the outcomes of acquirers involved in a single acquisition relative to the outcomes of acquirers involved in multiple acquisitions during our sample period. Figure A4 and Table A4 show that acquirers involved in multiple M&A transactions tend to grow a bit larger, in terms of the number of employees and average payrolls after the event, compared to acquirers involved in a single M&A deal. While these results provide suggestive evidence consistent with the empire-building story, these differences could be also simply driven by the fact that acquirers involved in multiple deals might mechanically increase in size more.

#### **A.5 Using Private Firms Only**

Prior research points out that publicly listed firms engage in M&As more than private firms during merger waves, and that acquisitions can be efficiency improving, especially when buyers and sellers are publicly listed firms during on-the-wave mergers (Maksimovic et al., 2013). While the vast majority (96 percent) of firms that go through M&As in our matched sample are private, we repeat our analysis focusing only on private firms, since the effects of M&As on firm-level outcomes might be different between listed firms and private firms. Figure A5 and Table A5 show that the results on firm and worker outcomes are similar to our main results where we include publicly listed companies, implying that our results are robust to just focusing on private firms in our sample.

## A.6 Use Sample of Unmatched Firms and Workers

In our main analysis sample, roughly 20 percent of M&A firms eligible for matching (as described in Section 5) are not matched to their control group either because of their sector, location, or firm characteristics. Correspondingly, about 68 percent of workers in M&A firms (eligible for matching) are not matched to their control group either because their firms were not matched or there are not enough control workers to be matched (as we impose that only one worker is matched with a treated worker). In Table A8, we show characteristics of these unmatched firms and workers. Panel A shows that unmatched firms are quite similar to our matched control firms on average in terms of total revenue, total expenses, and leverage ratio, but are much larger in terms of the number of employees, average payrolls, and markups. Panel C shows that unmatched workers have much lower earnings relative to our matched control workers on average. Note that the number of workers eligible for matching includes those at unmatched M&A firms eligible for matching. Therefore, the match rate for workers decreases from about 60 percent (as indicated in Section 5) to 32 percent once we include all eligible workers across all eligible firms. The match rate for our main worker sample (around 60 percent) is computed using the sample of all eligible workers among the matched sample of firms only.

An external validity concern is whether our matched sample of M&A firms and workers are representative within Canada. We argue that our matched sample of firms and workers are representative of overall M&A activities in Canada, given that we have a good match rate among firms and workers eligible for matching. Still, we test whether our results significantly change once we incorporate these previously unmatched set of firms and workers that were eligible for matching.

Panel (a) of Figure A6 shows the results on employment, separately for targets and for acquirers involved in acquisitions, including previously unmatched M&A firms. Here, we just add these unmatched acquirers and targets as part of the treated (M&A) group, without matching them with possible control firms, so we still use the same set of matched control firms in this analysis. As shown in the figure, the parallel pre-trend appears to get a bit weaker once we include previously unmatched firms, but we still find qualitatively similar results on employment after M&As. We find similar results on average payrolls in Panel (b). Panel (c) shows the results on worker-level earnings, separately for workers at targets and for workers at acquirers, including previously unmatched workers from M&A firms. While the parallel pre-trends are not as good as those in our main results (Panel (a) of Figure 3), we still find qualitatively similar results on worker-level earnings after the event. Therefore, these results suggest that our results are not driven by a particular sample of matched M&A firms and workers in our data.

## A.7 AKM and Match Effects

### A.7.1 AKM effects

To understand the role of the firm in explaining the earnings decline for target workers, we compare the observable characteristics of the targets firms and of the new firms that target workers move to after M&As in Section 7. In this section, we estimate firm-specific pay premiums using standard AKM models. We then characterize a firm-specific wage premium of the old employer and of the new employer for each worker who undergoes a separation following an M&A event to understand the decline in earnings. Our implementation of the AKM model regresses log earnings observed for individual  $i$  working at firm  $j$  in year  $t$  ( $y_{ijt}$ ) on employer-specific fixed effects which reflect firm characteristics that result in above- or below-average earnings for all workers at firm  $j$  ( $\phi_{j(i,t)}$ ), individual fixed effects ( $\omega_i$ ), and year effects ( $\tau_t$ ):

$$y_{ijt} = \phi_{j(i,t)} + \omega_i + \tau_t + u_{ijt} \quad (\text{A1})$$

We can then assess the role played by employers by estimating an analogue to equation (3), substituting in as the outcome variable the estimated firm fixed effects  $\hat{\phi}_j$ . The goal is to estimate the share of earnings losses following job transitions that can be attributed to a mover's reemployment by an employer with a different  $\hat{\phi}_j$  than the employer from which the mover left. Note that we omit M&A firms in the year of the event to avoid changes in composition affecting firm effects estimation, although including them does not affect our estimates by much.

Panel (a) of Figure A7 shows firm-specific wage premiums of workers at target firms who move to other firms after the M&A event. Relative to their matched control workers, job movers from target firms show a significant increase in their firm fixed effects after the event, implying that on average, they move to employers with higher wage premiums. Column (1) of Table A7 shows that workers who move from target firms experience 3.2 log points increase in firm-specific wage premiums after the event.

This suggests that workers are actually moving to firms with higher wage premiums after M&As. Importantly, most of these workers who leave target firms move to non-acquiring firms, meaning that this transition to larger firms with higher pays is not mechanically driven by worker reallocation toward acquirers. In practice, many factors may contribute to a firm having a higher wage premium. One common issue in this literature is that AKM effects may capture whether some firms pay more relative to others, but cannot control for other aspects of the firm. Therefore, to get into the black box of the AKM effects, we take advantage of our firm balance sheet data to characterize the types of firms that target workers transition to after M&As (see Section 7).

We test the main assumptions underlying the estimation of the AKM specification following [Lachowska et al. \(2020\)](#).

### **A.7.2 AKM Assumption 1: Sufficient Mobility**

The firm wage premium in the AKM model is identified by workers who move between firms. For this reason, the sample formation strategy (and the underlying mobility pattern) need to exhibit sufficient mobility to allow the firm wage premiums to be estimated. In the average year during our sample period, roughly 18 workers per employer move to other firms. Among full-time workers, the average number of movers per employer is about 8. Across the entire sample, 77 percent of workers make at least one move to a different firm during our sample period. Therefore, the mobility rates in the sample we use to estimate the AKM model appear to be high and comparable to the mobility rate in the sample used by [Lachowska et al. \(2020\)](#).

### **A.7.3 AKM Assumption 2: Exogenous Mobility**

Since the firm wage premium in the AKM model is identified by workers who move between firms, the model requires an assumption of exogenous mobility of workers between firms. If this assumption fails, then the firm wage premium would be biased because the workers who move would be different than those who do not move. We test this using a similar test as in [Lachowska et al. \(2020\)](#) who builds off an exercise in [Card et al. \(2013\)](#). We group firms into quartiles by their estimated firm fixed effects and study the wages of workers who move between firms. If wages of movers are determined by the quartile of the firm effects symmetrically both when moving from high to low and low to high firms, this symmetry supports the assumption that mobility is exogenous. In contrast, if movers show systematic wage gains regardless of the fixed effects of the origin firms and destination firms, then the assumption of exogenous mobility could be violated.

In Figure A8, we plot the log wages of job movers for eight different quartile-to-quartile transitions. The top of the figure shows workers moving from the top (fourth) to the top quartile of firms. The wages of these workers are high and stay high. The same stable pattern can be seen for those going from the bottom (first) to the bottom quartile; their wages are low and stay relatively low. In contrast, for workers going from the fourth to the first quartile of firm fixed effects, their wages drop significantly. Symmetrically, those going from the first to the fourth quartile of firm fixed effects see a strong increase in wages. Because the wage quality of the firm drives the wage change of the moving worker, this provides support for the exogenous mobility assumption.

#### A.7.4 Match Effects

Our results shown in Section 7 point to the idea that the decline in target workers' earnings is likely driven by these workers losing either firm-specific human capital or backloaded contracts. Following prior research that draws a connection between this mechanism and a loss in “match effects”, we estimate match effects based on Lachowska et al. (2020) which implements a strategy based on Woodcock (2015). For each employee-employer spell, we first calculate the average of residualized log earnings ( $\overline{y_{ij}}$ ) by removing calendar-year effects and regressing this adjusted log earnings on years of job tenure and worker-employer match indicators. We then compute within-match averages of the outcome after subtracting the contribution of job tenure. Then we estimate a model similar to the AKM model in equation (A1), but using within-match averages as the dependent variable:

$$\overline{y_{ij}} = \alpha_i + \pi_{j(i,t)} + e_{ij} \quad (\text{A2})$$

where  $\alpha_i$ ,  $\pi_{j(i,t)}$ , and  $e_{ij}$  denote the worker fixed effects, employer fixed effects, and an error term independent of individual and firm fixed effects, respectively.

We then calculate the residuals from equation (A2) and interpret them as worker-employer match effects averaged over the years we observe a given worker-employee match:

$$\hat{e}_{ij} = \overline{y_{ij}} - \hat{\alpha}_i - \hat{\pi}_{j(i,t)} \quad (\text{A3})$$

We proceed to take the estimated  $\hat{e}_{ij}$  terms relevant for the employee in each time period and use them as the dependent variable in equation (3) to see the contribution of match effects in explaining the earnings loss of target workers who move to other firms after the event.

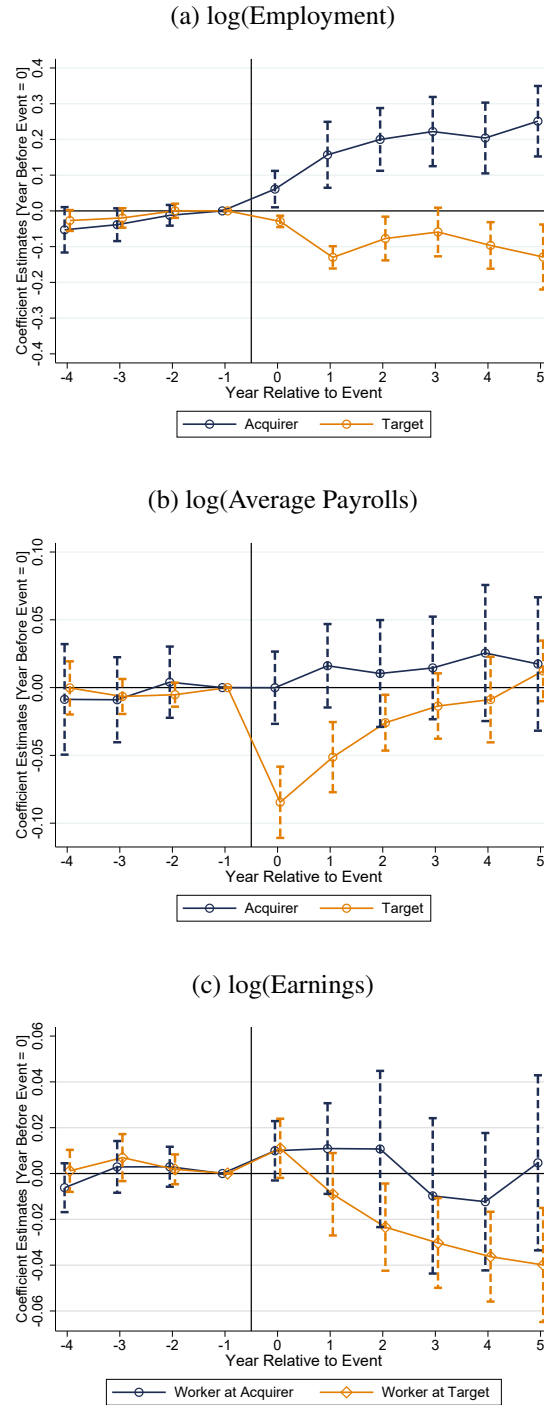
Panel (b) of Figure A7 shows match effects of workers at target firms who move to other firms after the M&A event. Relative to their control workers, movers from target firms experience a significant and large decrease in their match effects after the event, implying that on average, they move to employers with a lower match quality. The decrease in match effects in each post-M&A year comes both from workers who had already left target firms and from workers who leave in each subsequent year. Column (2) of Table A7 shows that workers who move from target firms within the first two years after the event experience 9 log points decrease in match effects.

The decline in match effects may imply that these workers lose the benefit of a specific employee skill set that fits better with the previous employer. Also, the decrease in match effects could simply indicate that these workers lose an employer-specific contract that yields a better work environment or amenity. Taken together, it appears that after target firms shrink through an M&A



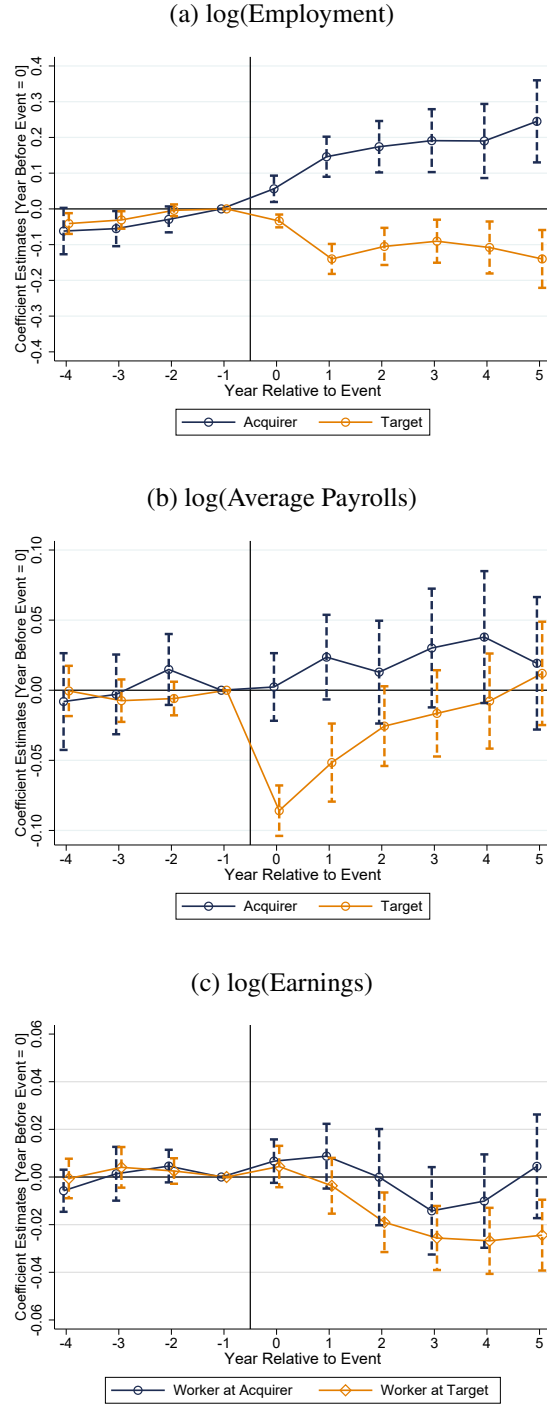
event, workers move to larger firms with higher wage premiums and take a modest wage cut in the medium run due to worse matches with new employers. Relative to previous work ([Lagaras forthcoming](#); [He and le Maire 2022](#)) documenting that M&A events create significant displacements of workers from target firms with higher wage premiums and reallocation to firms with lower wage premiums, we find that a significant share of workers' earnings losses after M&As stems primarily from a decline in match quality.

Figure A1: Different Clustering



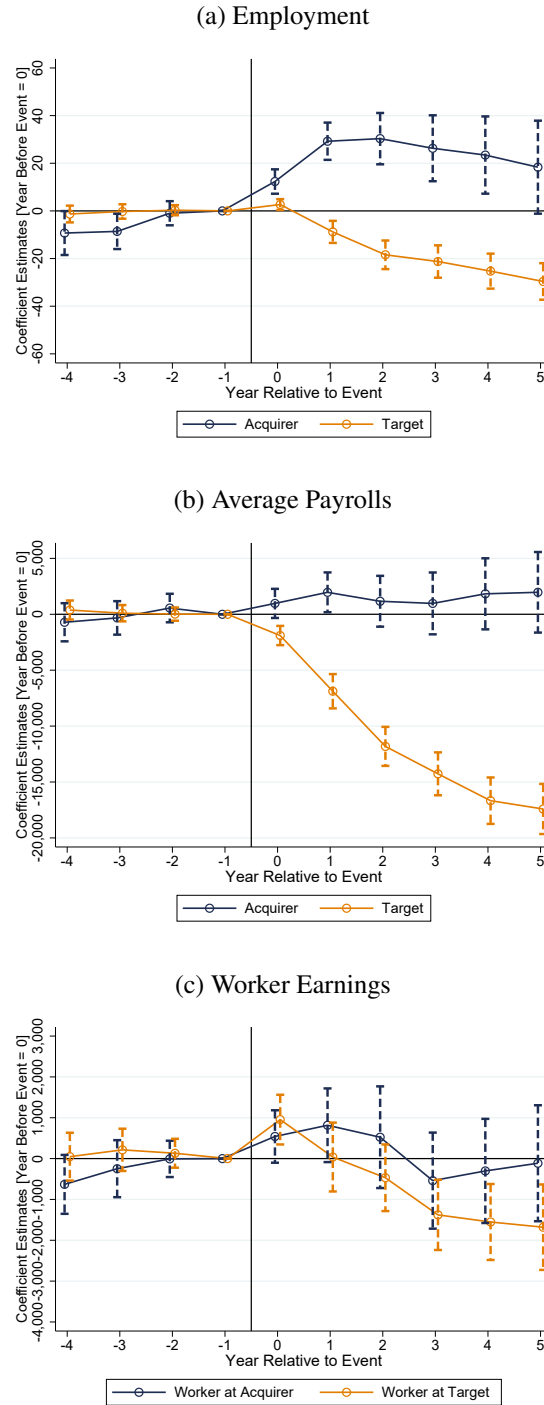
*Notes:* These figures display event-study estimates for the impact of M&As on the main firm-level and worker-level outcomes. Panel (a) shows the estimates for log of employment for those involved in acquisitions. Panel (b) shows the estimates for log of average payrolls for those involved in acquisitions. Panel (c) shows the estimates for log of worker-level earnings. The dashed lines indicate 95 percent confidence intervals where the standard errors are clustered at the market-level (by 4-digit NAICS  $\times$  commuting zone) for Panels (a) and (b), and at the worker and market level for Panel (c). The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure A2: Matched Control Firms in Different Markets



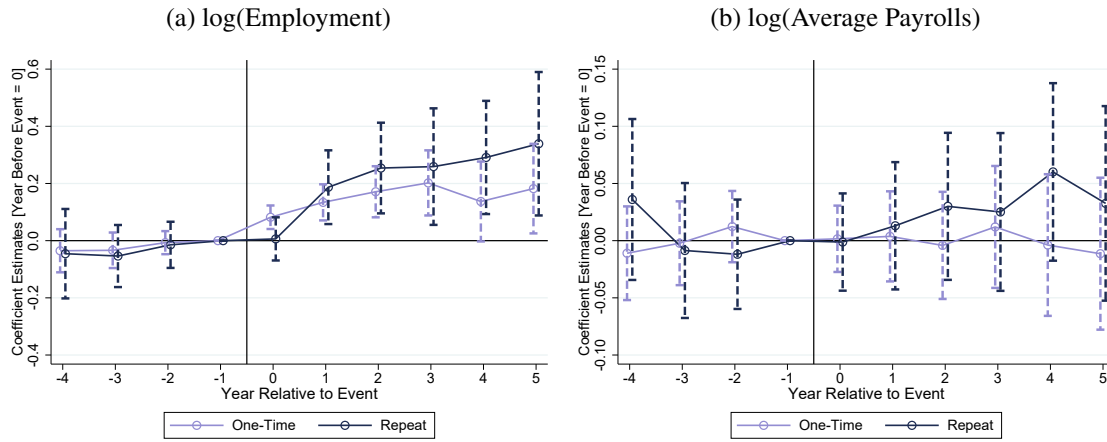
*Notes:* These figures display event-study estimates for the impact of M&As on the main firm-level and worker-level outcomes. For this analysis, we implement the same matching procedure (Section 5), except that we restrict M&A firms to be at different markets (defined by 4-digit NAICS  $\times$  commuting zone) from matched control firms. Panel (a) shows the estimates for log of employment for those involved in acquisitions. Panel (b) shows the estimates for log of average payrolls for those involved in acquisitions. Panel (c) shows the estimates for log of worker-level earnings. The dashed lines indicate 95 percent confidence intervals where the standard errors are clustered at the firm level for Panels (a) and (b), and at the worker and firm level for Panel (c). The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure A3: Outcomes in Levels (Replacing Missing with Zeros)



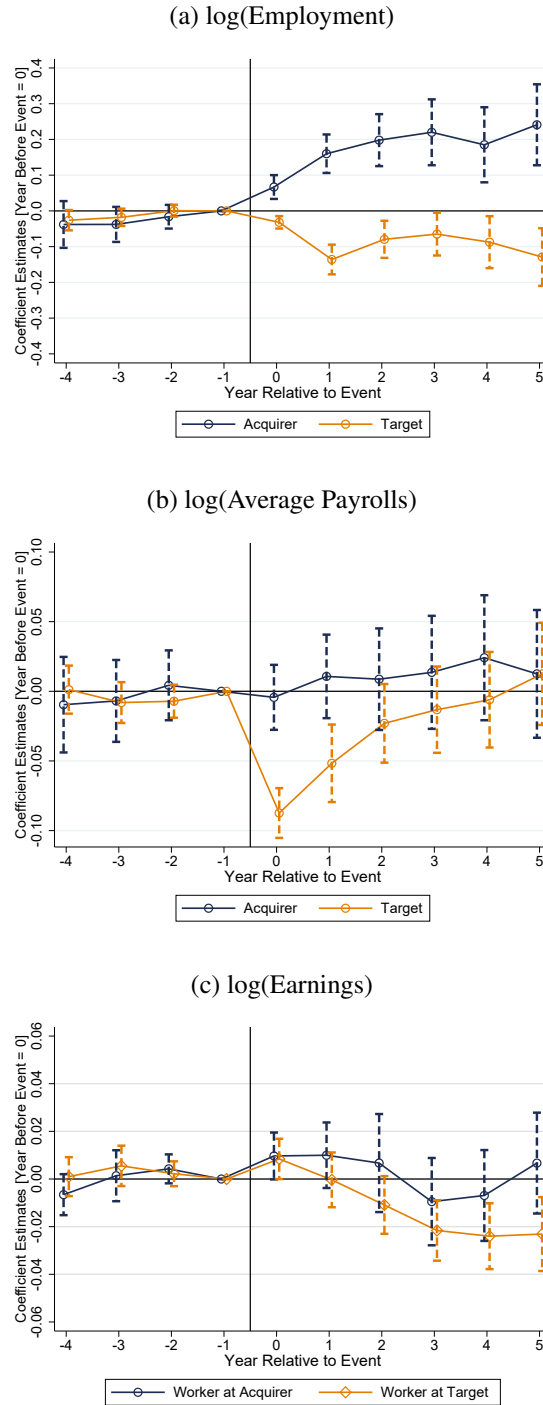
*Notes:* These figures display event-study estimates for the impact of M&As on the main firm-level and worker-level outcomes in levels, where we replace missing observations with zeros if the firms or workers exit the sample. Panel (a) shows the estimates for employment for those involved in acquisitions. Panel (b) shows the estimates for average payrolls for those involved in acquisitions. Panel (c) shows the estimates for worker-level earnings. The dashed lines indicate 95 percent confidence intervals where the standard errors are clustered at the firm level for Panels (a) and (b), and at the worker and firm level for Panel (c). The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure A4: By One-time vs. Repeat M&As (Acquirers)



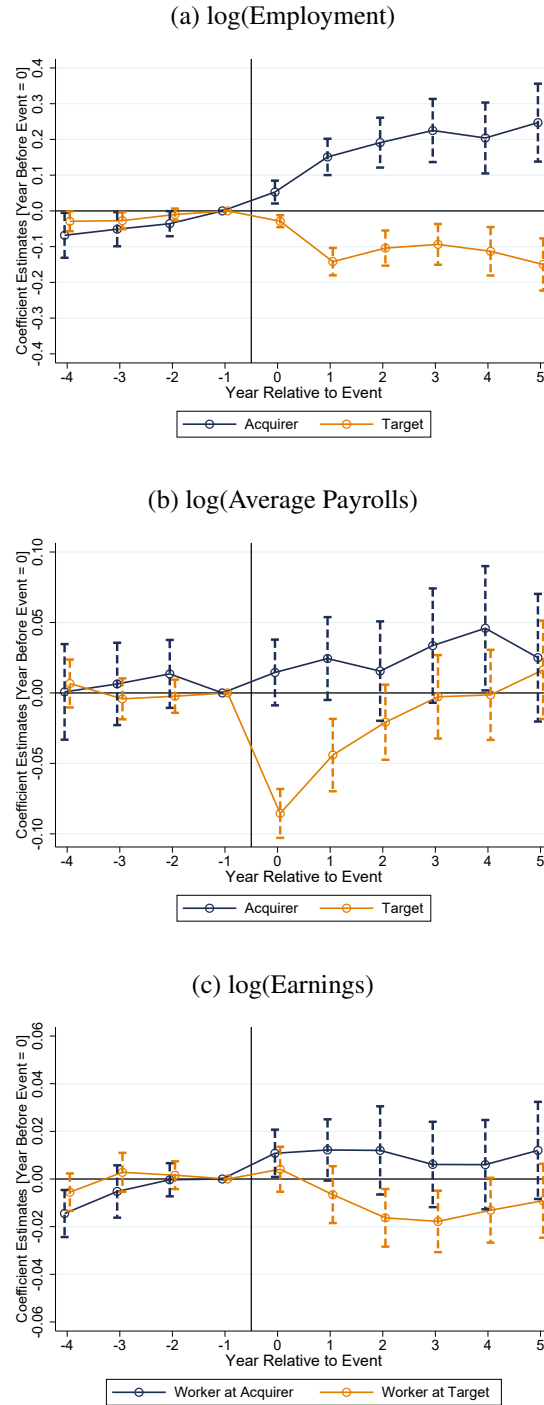
*Notes:* These figures display event-study estimates for the impact of M&As on the main firm-level outcomes, separately for acquiring firms that go through an M&A only once throughout our sample period and for acquiring firms with multiple M&A events. Panel (a) shows the estimates for log of employment. Panel (b) shows the estimates for log of average payrolls. The dashed lines indicate 95 percent confidence intervals where the standard errors are clustered at the firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure A5: Using Private Firms Only



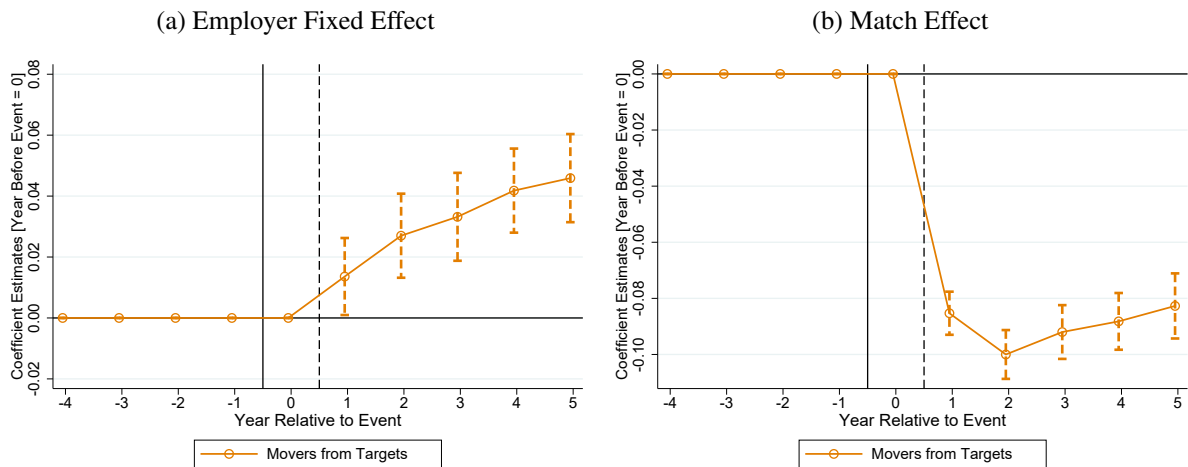
*Notes:* These figures display event-study estimates for the impact of M&As on the main firm-level and worker-level outcomes using only private firms/workers involved in acquisitions in the matched sample. In our matched sample, 96 percent of firms are private during our sample period. Panel (a) shows the estimates for log of employment. Panel (b) shows the estimates for log of average payrolls. Panel (c) shows the estimates for log of worker-level earnings. The dashed lines indicate 95 percent confidence intervals where the standard errors are clustered at the firm level for Panels (a) and (b), and at the worker and firm level for Panel (c). The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure A6: Including Unmatched Eligible Firms and Workers



*Notes:* Panel (a) and (b) display event-study estimates for the impact of M&As on the main firm-level outcomes separately for targets and for acquirers involved in acquisitions, including the unmatched sample of M&A firms eligible for matching. Panel (c) displays event-study estimates for the impact of M&As on worker-level earnings for workers at targets and workers at acquirer, including the unmatched sample of eligible workers at M&A firms. The dashed lines indicate 95 percent confidence intervals where the standard errors are clustered at the firm level for Panels (a) and (b), and at the worker and firm level for Panel (c). The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

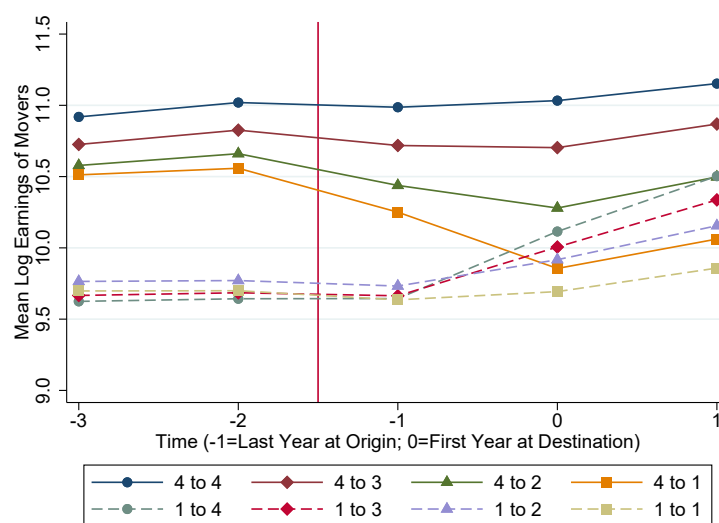
Figure A7: Workers Moving from Targets – Employer FE and Match Effects (Targets)



*Notes:* This figure displays event-study estimates of the impact of M&As for workers moving from target firms within the first two years after the event. Panel (a) shows the estimates for employer fixed effects. Panel (b) shows the estimates for worker-employer match effects. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker and firm level. The M&A event is in year 0 and the coefficient is normalized to be zero in year -1.

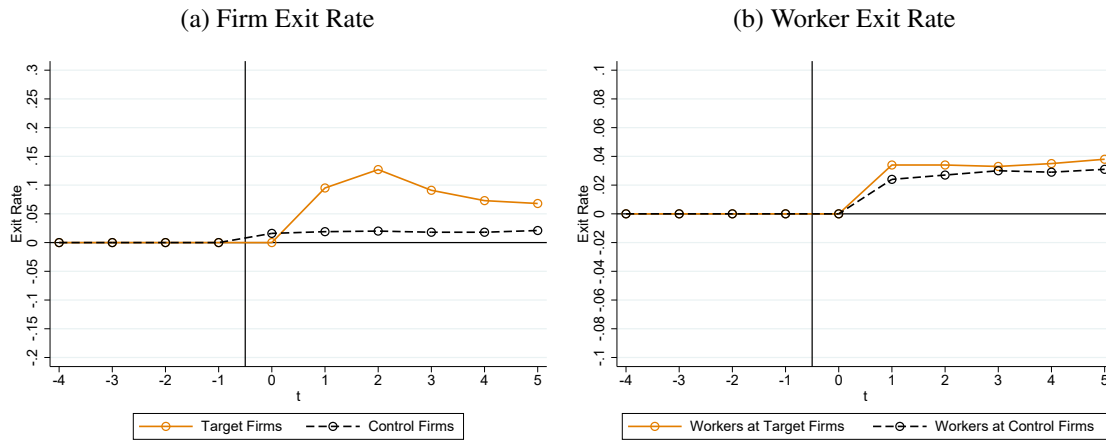


Figure A8: Exogenous Mobility Assumption



*Notes:* This figure displays the estimates of the average log earnings of job movers during the four-year period around their transition. The estimates are classified based on the quartile of firm effects at origin firms ( $t = -1$ ) and destination firms ( $t = 0$ ).

Figure A9: Probability of Exit



*Notes:* Panel (a) displays the exit rate of firms in the matched sample, separately for target firms and their control firms. After the event, the probability of exiting the sample increases by 2.4 percentage points ( $SE = 0.006$ ) for target firms (mostly those involved in mergers) relative to control firms on average. Panel (b) displays the exit rate of workers in the matched sample, separately for workers at target firms and for their control workers. After the event, the probability of exiting the sample increases by 0.6 percentage points ( $SE = 0.002$ ) for target workers relative to their control workers on average.

Table A1: Different Clustering

|                       | (1)                  | (2)                   | (3)                  |
|-----------------------|----------------------|-----------------------|----------------------|
|                       | log(Employment)      | log(Average Payrolls) | log(Earnings)        |
| Target                | -0.087***<br>(0.023) | -0.029***<br>(0.008)  | -0.021***<br>(0.008) |
| Mean at t = -1        | 4.00                 | 10.68                 | 11.02                |
| Adj. R squared        | 0.876                | 0.797                 | 0.65                 |
| Firm/Worker-Year Obs. | 66,380               | 66,260                | 1,954,480            |
| Acquirer              | 0.183***<br>(0.038)  | 0.014<br>(0.013)      | 0.002<br>(0.012)     |
| Mean at t = -1        | 4.55                 | 10.73                 | 11.03                |
| Adj. R squared        | 0.892                | 0.819                 | 0.642                |
| Firm/Worker-Year Obs. | 21,830               | 21,760                | 1,296,770            |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As on the main firm-level and worker-level outcomes. The outcome variables in Columns (1) to (3) are log of employment, log of average payrolls, and log of worker-level earnings. Columns (1) and (2) show the results for firms involved in acquisitions. The standard errors of the firm-level estimates are clustered at the market level (defined by 4-digit NAICS  $\times$  commuting zone). The standard errors of the worker-level estimates are two-way clustered at the worker and market level. Note that the coefficient estimates also changed slightly from clustering at the market-level because the information on commuting zone is missing for a small share of firms and workers, resulting in a slightly different sample compared to our main analysis sample.

Table A2: Matched Control Firms in Different Markets

|                       | (1)<br>log(Employment) | (2)<br>log(Average Payrolls) | (3)<br>log(Earnings) |
|-----------------------|------------------------|------------------------------|----------------------|
| Target                | -0.103***<br>(0.023)   | -0.029***<br>(0.011)         | -0.016***<br>(0.005) |
| Mean at t = -1        | 4.00                   | 10.69                        | 11.01                |
| Adj. R squared        | 0.877                  | 0.797                        | 0.739                |
| Firm/Worker-Year Obs. | 65,740                 | 65,620                       | 1,960,640            |
| Acquirer              | 0.167***<br>(0.033)    | 0.021<br>(0.014)             | -0.001<br>(0.007)    |
| Mean at t = -1        | 4.56                   | 10.73                        | 11.02                |
| Adj. R squared        | 0.891                  | 0.815                        | 0.731                |
| Firm/Worker-Year Obs. | 21,620                 | 21,560                       | 1,282,810            |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As on the main firm-level and worker-level outcomes. For this analysis, we implement the same matching procedure (Section 5), except that we restrict M&A firms to be at different markets (defined by 4-digit NAICS  $\times$  commuting zone) from matched control firms. The outcome variables in Columns (1) to (3) are log of employment, log of average payrolls, and log of worker-level earnings. Columns (1) and (2) show the results for firms involved in acquisitions. The standard errors of firm-level estimates are clustered at the firm level. The standard errors of worker-level estimates are two-way clustered at the worker and firm level.

Table A3: Outcomes in Levels (Replacing Missing with Zeros)

|                       | (1)<br>log(Employment) | (2)<br>log(Average Payrolls) | (3)<br>log(Earnings) |
|-----------------------|------------------------|------------------------------|----------------------|
| Target                | -16.784***<br>(2.535)  | -11492.158***<br>(730.927)   | -681*<br>(353)       |
| Mean at t = -1        | 105.69                 | 49060.32                     | 70,625               |
| Adj. R squared        | 0.845                  | 0.695                        | 0.79                 |
| Firm/Worker-Year Obs. | 73,430                 | 73,360                       | 2,026,430            |
| Acquirer              | 23.330***<br>(5.172)   | 1477.942<br>(1008.155)       | 154<br>(465)         |
| Mean at t = -1        | 177.75                 | 50885.79                     | 70,046               |
| Adj. R squared        | 0.86                   | 0.732                        | 0.793                |
| Firm/Worker-Year Obs. | 23,300                 | 23,260                       | 1,345,330            |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As on the firm-level and worker-level outcomes in levels, where we replace missing observations with zeros if the firms or the workers exit the sample. The outcome variables in Columns (1) to (3) are log of employment, log of average payrolls, and log of worker-level earnings. Columns (1) and (2) show the results for firms involved in acquisitions. The standard errors of firm-level estimates are clustered at the firm level. The standard errors of worker-level estimates are two-way clustered at the worker and firm level.

Table A4: By One-time vs. Repeat M&amp;A (Acquirer)

|  | (1)<br>log(Employment) | (2)<br>log(Average Payrolls) |
|--|------------------------|------------------------------|
| Post $\times$ Treated                  | 0.130***<br>(0.038)    | 0.009<br>(0.016)             |
| Post $\times$ Treated $\times$ Repeat  | 0.130**<br>(0.066)     | 0.010<br>(0.027)             |
| Mean at t = -1 (One-Time Acquirer = 1) | 4.40                   | 10.73                        |
| Mean at t = -1 (Repeat Acquirer = 1)   | 4.85                   | 10.74                        |
| Adj. R squared                         | 0.892                  | 0.817                        |
| Firm-Year Obs. (One-Time Acquirer = 1) | 15,400                 | 15,360                       |
| Firm-Year Obs. (Repeat Acquirer = 1)   | 7,110                  | 7,100                        |

*Notes:* This table reports the difference-in-differences estimates of the impact of M&As on the main firm-level outcomes for acquiring firms with only one M&A event during our sample period. The triple interaction term captures the triple-difference estimates for acquiring firms with repeated M&A events. Column (1) displays the estimates for log of employment. Column (2) displays the estimates for log of average payrolls. The standard errors are clustered at the firm level.

Table A5: Using Private Firms Only

|                  | (1)<br>log(Employment) | (2)<br>log(Average Payrolls) | (3)<br>log(Earnings) |
|------------------|------------------------|------------------------------|----------------------|
| Target           | -0.088***<br>(0.023)   | -0.028***<br>(0.011)         | -0.012**<br>(0.005)  |
| Mean at t = -1   | 4.01                   | 10.68                        | 11.01                |
| Adj. R squared   | 0.876                  | 0.796                        | 0.739                |
| Firm/Worker-Year | 65,560                 | 65,450                       | 1,990,940            |
| Acquirer         | 0.178***<br>(0.033)    | 0.011<br>(0.014)             | 0.003<br>(0.007)     |
| Mean at t = -1   | 4.56                   | 10.73                        | 11.02                |
| Adj. R squared   | 0.894                  | 0.818                        | 0.732                |
| Firm/Worker-Year | 21,040                 | 20,990                       | 1,308,270            |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As on the main firm-level and worker-level outcomes using only private firms/workers in the matched sample. In our matched sample, 96 percent of firms are private during our sample period. The outcome variables in Columns (1) to (3) are log of employment, log of average payrolls, and log of worker-level earnings. Columns (1) and (2) show the results for firms involved in acquisitions. The standard errors of firm-level estimates are clustered at the firm level. The standard errors of worker-level estimates are two-way clustered at the worker and firm level.

Table A6: Including Unmatched Eligible Firms and Workers

|                  | (1)                  | (2)                   | (3)                |
|------------------|----------------------|-----------------------|--------------------|
|                  | log(Employment)      | log(Average Payrolls) | log(Earnings)      |
| Target           | -0.105***<br>(0.021) | -0.023**<br>(0.010)   | -0.010*<br>(0.005) |
| Mean at t = -1   | 4.00                 | 10.64                 | 10.83              |
| Adj. R squared   | 0.872                | 0.793                 | 0.732              |
| Firm/Worker-Year | 73,670               | 73,580                | 3,815,300          |
| Acquirer         | 0.178***<br>(0.032)  | 0.027*<br>(0.014)     | 0.010<br>(0.007)   |
| Mean at t = -1   | 4.53                 | 10.68                 | 10.93              |
| Adj. R squared   | 0.890                | 0.816                 | 0.697              |
| Firm/Worker-Year | 24,610               | 24,580                | 3,143,520          |

*Notes:* Columns (1) and (2) report the difference-in-differences estimates for the impacts of M&As on the main firm-level outcomes separately for targets and for acquirers involved in acquisitions, including the unmatched sample M&A firms eligible for matching. Column (3) reports the difference-in-differences estimates for the impact of M&As on worker-level earnings, including the unmatched sample of eligible workers at M&A firms. The standard errors of firm-level estimates are clustered at the firm level. The standard errors of worker-level estimates are two-way clustered at the worker and firm level.



Table A7: Changes in Firm Effects and Worker-Employer Match Effects (Targets)

|                            | (1)                 | (2)                  |
|----------------------------|---------------------|----------------------|
|                            | Employer FE         | Match Effect         |
| Workers Moving From Target | 0.032***<br>(0.006) | -0.090***<br>(0.004) |
| Mean at t = -1             | 0.20                | 0.06                 |
| Adj. R squared             | 0.881               | 0.205                |
| Worker-Year Obs.           | 684,800             | 677,750              |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As for workers moving from target firms within the first two years after the event. Column (1) displays the estimates for the employer fixed effects. Column (2) displays the estimates for worker-employer match effects. The standard errors are two-way clustered at the worker and firm level.

Table A8: Mean Log Wage Before and After Employer Change by Quartile of Mean Coworkers' Wages at Origin and Destination Firms

| Origin to<br>Destination<br>Quartile | (1)                       | (2)                     | (3)                | (4)               | (5)               | (6)                        | (7)      |
|--------------------------------------|---------------------------|-------------------------|--------------------|-------------------|-------------------|----------------------------|----------|
|                                      | Number of<br>Observations | Mean Log Wage of Movers |                    |                   |                   | Change from t = -2 to t=+1 |          |
|                                      |                           | Two Years<br>Before     | One Year<br>Before | One Year<br>After | Two Year<br>After | Unadjusted                 | Adjusted |
| 1 to 1                               | 1,355,190                 | 9.70                    | 9.64               | 9.70              | 9.87              | 0.16                       | 0.00     |
| 1 to 2                               | 834,130                   | 9.77                    | 9.73               | 9.92              | 10.15             | 0.38                       | 0.22     |
| 1 to 3                               | 429,280                   | 9.68                    | 9.66               | 10.00             | 10.33             | 0.65                       | 0.49     |
| 1 to 4                               | 306,100                   | 9.64                    | 9.65               | 10.12             | 10.51             | 0.86                       | 0.70     |
| 2 to 1                               | 384,560                   | 10.05                   | 9.89               | 9.80              | 9.99              | -0.06                      | -0.18    |
| 2 to 2                               | 1,110,570                 | 10.30                   | 10.22              | 10.26             | 10.42             | 0.12                       | 0.00     |
| 2 to 3                               | 777,090                   | 10.38                   | 10.33              | 10.44             | 10.63             | 0.26                       | 0.14     |
| 2 to 4                               | 383,820                   | 10.27                   | 10.23              | 10.48             | 10.77             | 0.51                       | 0.39     |
| 3 to 1                               | 160,360                   | 10.30                   | 10.05              | 9.82              | 10.03             | -0.27                      | -0.38    |
| 3 to 2                               | 429,880                   | 10.51                   | 10.35              | 10.31             | 10.50             | -0.01                      | -0.12    |
| 3 to 3                               | 1,009,290                 | 10.66                   | 10.61              | 10.65             | 10.77             | 0.11                       | 0.00     |
| 3 to 4                               | 956,990                   | 10.72                   | 10.70              | 10.79             | 10.92             | 0.20                       | 0.09     |
| 4 to 1                               | 97,140                    | 10.57                   | 10.26              | 9.87              | 10.08             | -0.49                      | -0.62    |
| 4 to 2                               | 187,550                   | 10.68                   | 10.45              | 10.30             | 10.51             | -0.17                      | -0.30    |
| 4 to 3                               | 413,630                   | 10.85                   | 10.72              | 10.71             | 10.88             | 0.03                       | -0.10    |
| 4 to 4                               | 1,250,900                 | 11.02                   | 10.99              | 11.04             | 11.15             | 0.13                       | 0.00     |

*Notes:* This table reports the average of log of earnings for workers who move and are observed for at least two years prior to a job transition, and for two years after. We define quartiles based on the average earnings of coworkers at the origin firm in the year prior to transition, and at the destination group in the year after. The adjusted earnings change is estimated as the average change for each origin-destination group, minus the average change for switchers from the same origin quartile who remain in the same quartile.

Table A9: Descriptive Statistics on Unmatched Eligible Firms and Workers

|  | (1)<br>Acquirer | (2)<br>Target |
|--|-----------------|---------------|
| <i>Panel A: Firm Characteristics</i>   |                 |               |
| Total Revenue (in millions)            | 42              | 24            |
| Total Expense (in millions)            | 40              | 24            |
| Number of Employees                    | 238             | 155           |
| Average Wage Bill                      | 64,637          | 64,653        |
| Leverage Ratio                         | 0.67            | 0.73          |
| Markups                                | 2.97            | 2.87          |
| Number of Firms                        | 290             | 810           |
| <i>Panel B: Sectors (Firms)</i>        |                 |               |
| Construction                           | 0.01            | 0.02          |
| Manufacturing                          | 0.14            | 0.13          |
| Wholesale                              | 0.05            | 0.08          |
| Retail                                 | 0.04            | 0.03          |
| Transportation                         | 0.03            | 0.06          |
| Information                            | 0.15            | 0.19          |
| Services                               | 0.12            | 0.19          |
| Other Sectors                          | 0.46            | 0.3           |
| <i>Panel C: Worker Characteristics</i> |                 |               |
| Total Earnings                         | 61,611          | 55,731        |
| Age                                    | 47.0            | 45.0          |
| Female                                 | 0.39            | 0.43          |
| Number of Workers                      | 115,866         | 112,539       |
| <i>Panel D: Sectors (Workers)</i>      |                 |               |
| Construction                           | 0.03            | 0.01          |
| Manufacturing                          | 0.15            | 0.23          |
| Wholesale                              | 0.07            | 0.08          |
| Retail                                 | 0.10            | 0.43          |
| Transportation                         | 0.37            | 0.02          |
| Information                            | 0.06            | 0.02          |
| Services                               | 0.08            | 0.08          |
| Other Sectors                          | 0.13            | 0.13          |

*Notes:* This table reports descriptive statistics on the the unmatched sample of M&A firms and workers eligible for matching, measured one year prior to the event. Panel A reports firm characteristics such as total revenue, total expenses, number of employees, average payrolls, leverage ratio, and markups. Panel B reports the distribution of firms in the matched sample across 2-digit NAICS sectors. Panel C reports worker characteristics such as total annual earnings, age, and gender. Panel D reports the distribution of workers in the matched sample across 2-digit NAICS sectors. Other sectors include (1) Agriculture, forestry, and fishing, (2) Mining, quarrying, and oil and gas extraction, (3) Utilities, (4) Real estate and rental and leasing, (5) Arts, entertainment and recreation, (6) Accommodation and food services, (7) Other services, and (8) Public administration.

Table A10: Additional Descriptive Statistics on Workers at Targets

|                   | (1)    | (2)     | (3)                        | (4)                    | (5)                   | (6)            |
|-------------------|--------|---------|----------------------------|------------------------|-----------------------|----------------|
|                   |        |         | <b>Movers from Targets</b> |                        | <b>By Destination</b> |                |
|                   | Movers | Stayers | Voluntary Separation       | Involuntary Separation | To Acquirer           | To Other Firms |
| Total Earnings    | 68,023 | 72,551  | 65,800                     | 64,531                 | 67,820                | 65,082         |
| Age               | 46.0   | 48.6    | 40.9                       | 46.6                   | 46.7                  | 45.7           |
| Female            | 0.31   | 0.31    | 0.35                       | 0.32                   | 0.36                  | 0.32           |
| Number of Workers | 21,890 | 33,130  | 2,510                      | 7,880                  | 4,560                 | 13,660         |

*Notes:* This table reports descriptive statistics on the matched sample of workers at target firms, measured one year prior to the event. Columns (1) and (2) report these statistics, respectively, for workers moving from target firms within the first two years after the event and for workers who stay at target firms throughout the entire post-event period. Column (3) reports these statistics for workers moving voluntarily from target firms and column (4) reports these statistics for workers displaced from targets. Column (5) reports these statistics for workers who move to acquiring firms after the M&A event and column (6) reports the statistics for workers who move to other firms after the M&A event.

## **B Additional Heterogeneity Results**

In Appendix B, we provide results from heterogeneity analyses in addition to those discussed in Sections 5 – 7. Here, for worker-level analyses, we focus on target workers to see whether alternative mechanisms can explain the decline in their earnings after M&As. For most of the heterogeneity results, we separately look at stayers at targets and job movers from targets to get a better sense where the effects are concentrated.

### **B.1 By Local vs. National M&As**

While checking common pre-trends is reassuring for a causal interpretation when implementing a difference-in-differences design, contemporaneous shocks that occur with M&A events could still bias our results. For example, there could be a negative demand shock that affects a commuting zone and causes both a decline in employment and wages and an increase in M&A activities as firms get purchased before they shut down. In this case, M&A activities are correlated with shocks that decrease labor demand. We can also have the opposite scenario.

We address this concern by also looking at M&As that are less likely to have been triggered by local economic conditions of the firm. Specifically, we consider the impact of national M&As that occur among domestic firms with multiple establishments across different commuting zones. The intuition is that these changes in ownership are less likely to be driven by the local economic conditions of the firms or workers. Figure B1 and Table B1 show that the effect of local M&As on earnings of target workers is economically indistinguishable from the effect of national M&As on the same outcome. Therefore, the decline in earnings of workers at target firms is unlikely driven by local economic conditions of firms where M&A activities occur.

### **B.2 By Acquisition vs. Merger**

In our analysis sample, roughly 80 percent of M&A events are acquisitions (75 percent among the entire M&A events). In other words, the vast majority of M&A events in Canada (and in North America) involves an acquirer purchasing a part of a target’s businesses. As discussed in Sections 3 and 4, it is possible that a wage decline is larger (or smaller) in the case of a merger, where there is a complete transfer of ownership. We explore whether impacts on worker earnings are larger in the case of a merger, compared to a acquisition. We examine these impacts separately for stayers and job movers, so that we can see in which group the effects are concentrated. In Figure B2 and Table B2, we find that decreases in workers’ earnings in target firms in the case of mergers

were not larger than decreases in workers' earnings in the case of acquisitions, suggesting that our results are not driven by the fact the majority of our M&A events involves acquisitions. In other words, both mergers and acquisitions create job separations of workers from target firms, resulting in a wage loss that we observe in the data.

### **B.3 By Within vs. Across Labor Markets**

We explore whether impacts on worker earnings are larger in markets where merging firms are located in the same labor market (defined at the four-digit industry by commuting zone level), following [Prager and Schmitt \(2021\)](#). Once again, we look at these impacts separately for stayers and job movers, so that we can see where the effects are concentrated. In Figure B3 and Table B3, we find that decreases in workers' earnings in target firms where M&As occur within the same market were not larger than decreases in workers' earnings where M&As occur across different markets, suggesting a limited role for the change in concentration in explaining the change in worker earnings after M&As.

### **B.4 Distribution of HHI Across Markets**

Before we assess the heterogeneity based on the initial level of concentration (measured in Herfindahl-Hirschman Index) in labor markets, we examine the distribution of labor market concentration across different markets (defined at the four-digit industry by commuting zone level) in our data. Panel (a) in Figure B4 shows that the distribution of HHI across the entire labor market is a bit skewed towards the left, with some spikes around the middle and at the very right tail. Column (1) in Table B4 shows that the average and median HHI are 0.38 and 0.28 across the entire labor market in Canada during our sample, with roughly 12 percent of labor markets have a single employer in each market. Focusing on labor markets in the entire sample with any M&A event during our sample period, however, the distribution becomes a lot more left-skewed, with only 3 percent of labor markets having a single firm, as shown in panel (b) of Figure B3. Column (2) of Table B4 shows that the average and median HHI are 0.264 and 0.175. These statistics become even smaller (to 0.138 and 0.069, respectively, as shown in Column (3) of Table B4) when we zoom in on labor markets with any M&A event in our matched analysis sample, with less than one percent of labor markets having a single employer. Therefore, while the labor markets in the entire sample are reasonably concentrated, the markets in our analysis sample are not very concentrated on average to begin with, which can explain why M&A events in our analysis sample have little impacts on concentration in our setting.

## B.5 By Initial Level of Concentration

Our findings on the decreases in employment and worker-level earnings in target firms can be potentially rationalized by an increase in monopsony power of firms through increased concentration. [Prager and Schmitt \(2021\)](#) and [Arnold \(2021\)](#) study this channel in the U.S. context and find that M&A events that generate large shifts in concentration result in market-level declines in earnings. While this channel may still be a factor for a subset of the M&A events we study in Canada, we rule it out as being the only factor determining wage losses.

In particular, we find that 99 percent of M&A events have a zero predicted change in local labor-market concentration, and even among the remaining 1 percent, the predicted change is not large enough to generate any meaningful variation. This is consistent with the description on the distribution of HHI in our analysis sample in the previous subsection. In contrast, [Prager and Schmitt \(2021\)](#) and [Arnold \(2021\)](#) isolate M&A events with statistically significant increases in concentration.<sup>37</sup>

To examine this in more detail, in Figure B5 and Table B5, we turn to a metric that antitrust authorities consider: the initial concentration level. To study this channel, we split the analysis sample by quintiles in the HHI measured one year before the event. In both the high-concentration markets (fifth quintile) and the low-concentration markets (first quintile), we find similar levels of declines in earnings of either stayers at or job movers from target firms.<sup>38</sup> This is not surprising given most of these events do not actually increase concentration in the local labor market.

Taken together, these results, along with the results from Section B3, suggest that increased monopsony power due to changes in local labor-market concentration or outside options is unlikely to be the primary driver of declines in worker earnings in our setting.<sup>39</sup>

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<sup>37</sup>In particular, when conducting market-level analyses, [Arnold \(2021\)](#) focuses on the top 5 percent of the U.S. labor markets that experience a sufficiently large predicted change in concentration after M&As, implying that the remaining 95 percent do not experience any meaningful (predicted) change in concentration.

<sup>38</sup>Given that the change in concentration after M&As was close to zero for about 99 percent of labor markets in our sample, cutting our analysis sample based on the predicted change in HHI after M&As, as in [Arnold \(2021\)](#), is not meaningful in our setting due to the lack of variation in the predicted change in HHI.

<sup>39</sup>Even if the results are not driven by changes in concentration, changes in bargaining power of workers may be independent of concentration changes. For example, [He and le Maire \(2022\)](#) finds that M&A events in Denmark result in high-wage managers being replaced in target firms. Such a change in management may result in shifts in bargaining power of workers at target firms. However, a change in bargaining power through a change in management is only relevant for incumbent or new workers at target firms, and thus is unlikely to explain the decline in earnings of workers leaving target firms.

## B.6 By Initial Level of (flows-adjusted) Concentration

In the previous subsection, we show our estimates on worker earnings, separately for markets with low level (below the first quintile) of HHI and for markets with high level (above the last quintile) of HHI measured one year before the event. A standard Herfindahl-Hirschmann Index (HHI) takes as given the definition of the market and then computes

$$HHI = \sum_j s_j^2, \quad (B1)$$

as the measure of concentration. We describe a flows-adjusted concentration measure that takes into account transitions across markets, following [Arnold \(2021\)](#). To begin, let market  $m$  be defined by the interaction between 4-digit NAICS and commuting zone. The flows-adjusted concentrated measure (denoted  $C$ ) requires computing transition rates across markets. While, in theory, transition rates across markets may change, we instead choose to pool the entire sample in order to retrieve a consistent and more precise measure of the rate of transitions across markets. The share of firm  $j$  in market  $m$  is given by:

$$\tilde{s}_{jm} = \frac{l_{jm}}{\sum_k \alpha_{m \rightarrow k} L_k} \quad (B2)$$

where

$$\alpha_{m \rightarrow k} = \frac{P(k|m)}{P(m|m)} \frac{L_m}{L_k} \quad (B3)$$

where  $P(k|m)$  is the probability an individual from market  $m$  transitions to market  $k$  conditional on experiencing a transition. The intuition behind this formulation is that jobs in other markets likely provide viable options for workers. [Arnold \(2021\)](#) shows that one can use a discrete choice model and empirical flows across markets (*i.e.*,  $\alpha_{k \rightarrow m}$ ) to measure the value a worker from a given market places on another market.

Intuitively, if we observe a large number of flows from market  $m$  to  $k$ , then  $k$  likely serves as a viable outside option. Additionally, we need to take into account the relative sizes of the markets. For example, if  $k$  is a relatively small market, but we still observe high rates of flows to this market, it must provide a particularly good option for the workers. This is why the relative size between  $m$  and  $k$  is taken into account when determining the value individuals from  $m$  place in receiving a job in market  $k$ . The concentration in market  $m$  is given by:

$$\tilde{C} = \sum_j \tilde{s}_{jm}^2 \quad (B4)$$



One key difference in this formulation relative to [Arnold \(2021\)](#) is that transitions across both industries and locations are taken into account. In [Arnold \(2021\)](#), the market shares depend only on employment in other industries within the same commuting zone. In this more general version, the market share depends on employment in other commuting zones as well.

Second, the relative size term in  $\alpha_{m \rightarrow k}$  (i.e.  $\mathbb{E}[\frac{L_k}{L_m}]$ ) is now the expected relative size of industries across commuting zones. To understand this factor, imagine there are two equally sized industries that use similar workers but are generally located in different areas. For example, imagine plastic manufacturing and rubber manufacturing plants hire similar workers, but plastic manufacturing primarily takes place in Texas while rubber manufacturing primarily takes place in Ohio. In this case, the aggregate relative size of the industries will be quite different than the expected relative size within a commuting zone given the two industries primarily operate in different commuting zones. Therefore, a low volume of flows between the two industries does not necessarily reflect low substitutability, but rather they are generally located in different areas.

The flows-adjusted local labor market concentration measure,  $C_{mc}$ , is defined as:

$$C_{mc} = \sum_{j \in c} (\tilde{s}_{jmc})^2 \quad (\text{B5})$$

Figure B6 and Table B6 show the effects of M&As on earnings of workers at target firms, separately for markets with high (flow-adjusted) HHI and for markets with low (flow-adjusted) HHI. Similar to our main results based on the regular HHI, these results show that the decline in worker earnings is not larger for high-HHI markets.

## B.7 Results on Revenue, Markups, and Realized Capital Gains

To get a better sense of why acquirers' and targets' profitability decreases after M&As (discussed in Section 7), we next examine what happens to their revenue, markups, and cash-out responses. Panel (a) of Figure B7 shows the results on total revenue, separately for targets and for acquirers involved in acquisitions. Both targets' and acquirers' sales were in a parallel trend with those of their matched control firms prior to the event. We find that while acquirers' revenue increases significantly after the event, targets' revenue declines after the event. Even though acquirers' sales increase after M&As, their profit margins can decrease if the costs of acquisitions outweigh the benefits in the short to medium run. Furthermore, while the decline in targets' revenue is mechanical in the case of acquisitions, the decrease in profitability happens if they sold a profitable part of their businesses to acquirers. Panel (d) shows the aggregate outcomes (targets and acquirers pooled) separately for those involved in acquisitions and for those involved in mergers. In both

cases, revenue significantly decreases after the event in the aggregate.

Panel (b) of Figure B7 shows the results on markups, separately for targets and for acquirers involved in acquisitions. Both targets' and acquirers' markups were in a parallel trend with those of their matched control firms prior to the event, and remained relatively unchanged after the event. Panel (e) shows the aggregate outcomes (targets and acquirers pooled) separately for those involved in acquisitions and for those involved in mergers. In the case of acquisitions, markups do not change much after the event. In the case of mergers, markups decrease after the event in the aggregate. These results, in conjunction with the results on profitability, provide more direct evidence that firms' product market power did not increase after the event at least in the short to medium run.

Panel (c) of Figure B7 shows the results on realized capital gains, separately for targets and for acquirers involved in acquisitions. We find that while acquirers' realized capital gains decreased slightly after the event, targets' realized capital gains increase significantly after the event. The increase in realized capital gains at targets is driven by their initial investors selling a part of their shares ("cashing out") to those in acquiring firms, while the slight decrease in realized capital gains at acquiring firms indicate that their investors were net buyers of shares after the event. Panel (f) shows the aggregate outcomes (targets and acquirers pooled) separately for those involved in acquisitions and for those involved in mergers. In both cases, we see increases in realized capital gains. These results imply that even though targets' profitability did not improve after the event, initial investors still benefitted from the transactions by selling a part of their shares to acquirers.

To interpret the magnitude of these results, Table B7 presents the difference-in-differences estimates on these outcomes. Columns (1) and (2) show that targets' revenue decreases by 54.5 log points, while acquirers' revenue increases by 26.9 log points on average. Columns (3) and (4) show that in the aggregate, revenue decreases by 30.7 log points in the case of acquisitions, while it decreases by 10.3 log points in the case of mergers. Columns (1) and (2) show that acquirers' and targets' markups do not change much after the event on average. Columns (3) and (4) show that in the aggregate, markups do not change in the case of acquisitions, while they decrease by 4.2 log points in the case of mergers. Columns (1) and (2) show that targets' realized capital gains increase by 27,558 CAD, while acquirers' realized capital gains decrease by 13,071 CAD on average. Columns (3) and (4) show that in the aggregate, realized capital gains increase by 21,646 CAD in the case of acquisitions, while they increase by 28,002 CAD in the case of mergers. Overall, the decrease in profitability, without much change in markups, suggests that neither targets nor acquirers experienced an increase in market power after M&As at least in the short to medium run (De Loecker et al. 2020).

## **B.8 By Tradable vs. Non-Tradable Sectors**

There are several results that suggest product market power is not the main driver of the observed labor-market impacts. First, we find that profitability decreased for both target and acquiring firms. If product market power increased on average after M&A events, we would expect to see a rise in profits, not a decline (De Loecker et al. 2020). Additionally, we estimate markups and do not find any statistically significant impacts of M&A events on markups for either target or acquiring firms.

To explore this channel further, we also conduct our analysis separately by tradable and non-tradable goods sectors. The intuition is that an M&A would have a larger impact on a firm's market power if the firm does not face competition outside its geographical (i.e., international) markets. We define firms as active in tradable good sectors if they fall under Agriculture, Forestry, Fishing, Mining, Quarrying, and Oil and Gas Extraction, and Manufacturing. Firms active in other sectors (i.e., Construction, Retail, Real Estate, Services, etc) are defined as falling under non-tradable sectors (Berger et al., 2022; Delgado et al., 2014). Figure B8 and Table B8 show that the declines in worker earnings at target firms are not larger for non-tradable sectors. Therefore, this finding, along with the results on decreased profit margins without much change in markups, suggests that an increase in product market concentration seems to be unlikely the key driver behind the decline in target workers' earnings.

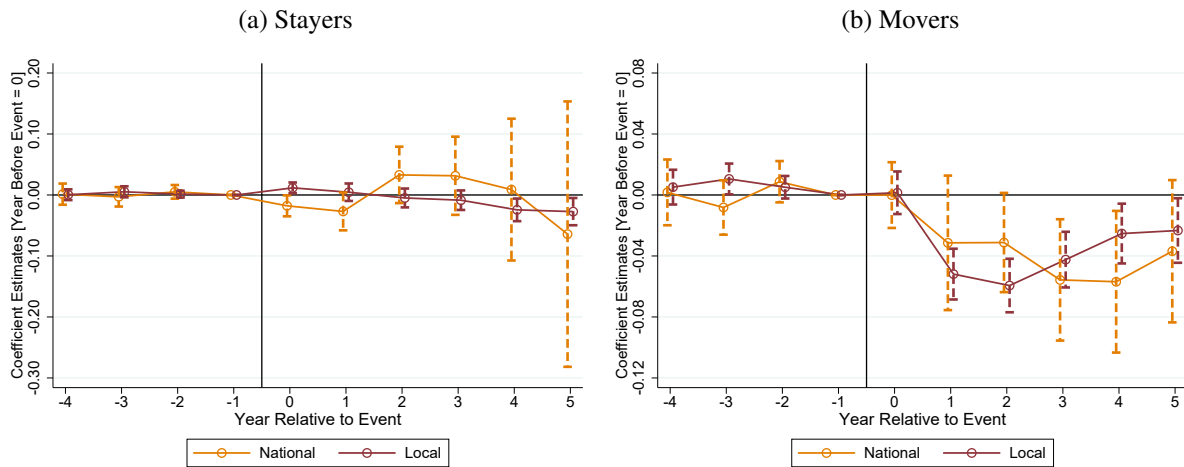
## **B.9 By Within vs. Between Industry M&As**

In the previous subsection, we show that the decline in earnings of workers at target firms is similar between M&As that happen in tradable sectors and those that happen in non-tradable sectors, suggesting a limited role for the change in product market power in explaining the change in worker earnings. To further support this conclusion, we also conduct our analysis separately by within-industry M&As and between-industry M&As. The intuition is that an M&A would have a larger impact on firms' market power if the acquirer buys another firm within the same industry (i.e., horizontal mergers). We divide our sample of all M&A firms based on the industries of the parties involved in a transaction. An M&A is within-industry if the industries (4-digit NAICS) of both parties are identical and it is between-industry (i.e., vertical mergers) if the industries are different. For firms with one transaction, we define a firm as "within" if it participated in a within-industry M&A and as "between" if it participated in a between-industry M&A. For firms with multiple M&A deals, we consider the majority of transactions to determine the within- and between-indicator. Figure B9 and Table B9 show that the decline in worker earnings is larger for those involved in within-industry M&As. These results imply that a rise in product market power is unlikely the main driver behind the results on worker earnings.

## B.10 By Worker Age

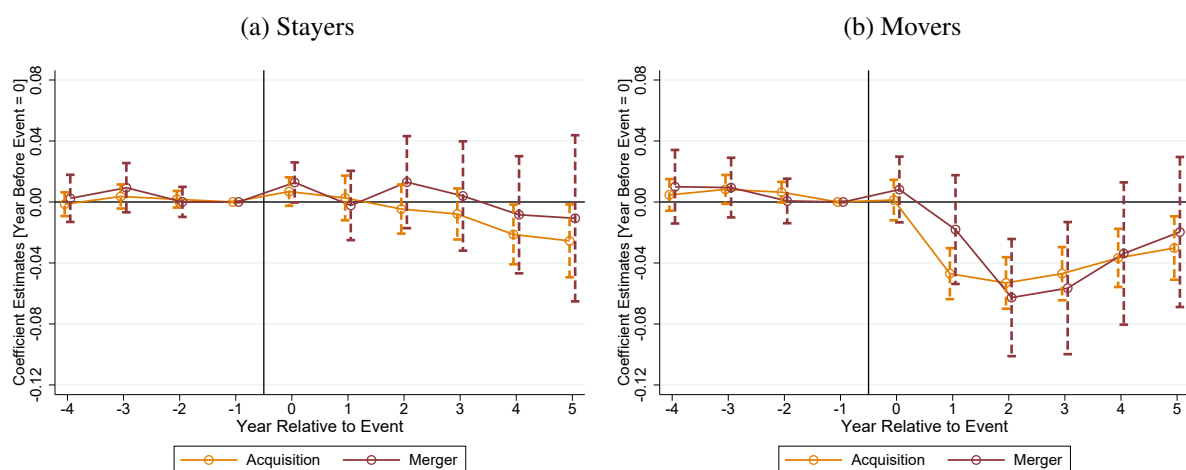
We additionally explore whether the decline in earnings of workers at target firms is different based on workers' age. Prior studies have found differential impacts of firm-level shocks on worker earnings depending on their age (Kline et al. 2019; Saez et al. 2019). As shown in Section 6, the decline in workers' earnings at target firms is entirely driven by those who move to other firms after the M&A event. Therefore, we focus on job movers from target firms, and estimate what happens to their earnings, separately across different age groups. Figure B10 and Table B10 show that while we observe declines in earnings across all age groups for workers moving from target firms, the decline in earnings is largest among movers who are at least 50 years old before the event. Taken together, these results imply that there exists a substantial degree of heterogeneity across age groups for changes in worker earnings after the M&A event. These results are consistent with our main results based on worker tenure in Section 7.

Figure B1: Worker-level Earnings By National M&As vs. Local M&As (Targets)



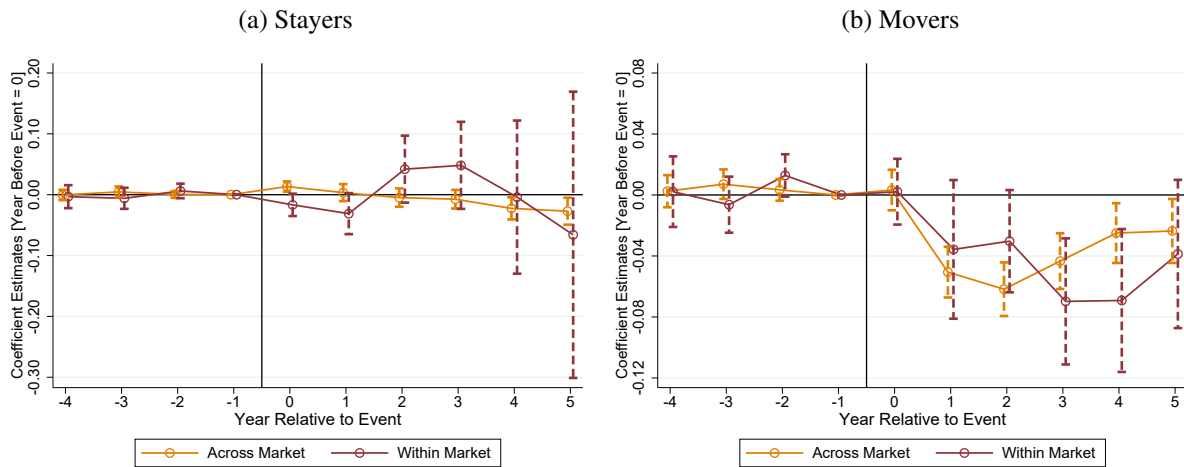
*Notes:* These figures display event-study estimates for the impact of M&As on log of total earnings for workers at target firms, separately for those involved in national M&A deals and for those involved in local M&A deals. Panel (a) displays the estimates for stayers and panel (b) displays the estimates for workers moving from target firms within the first two years after the event. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker-firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure B2: Worker-level Earnings By Acquisition vs. Merger (Targets)



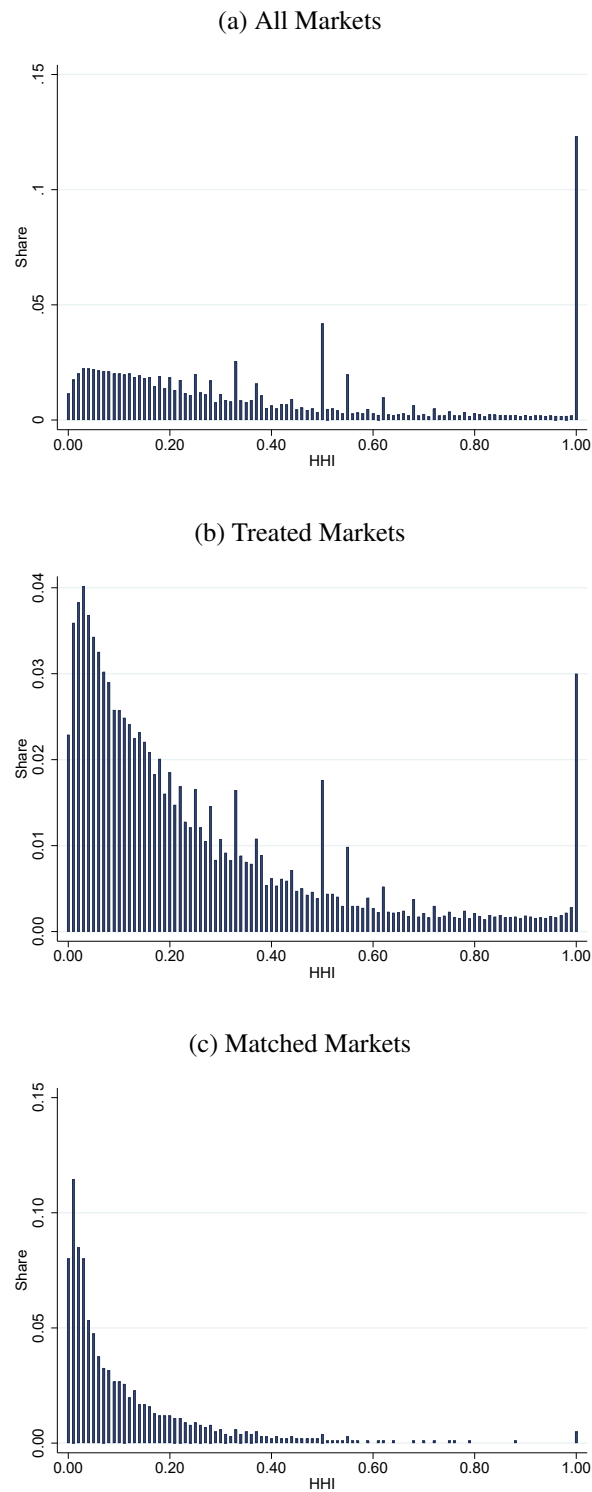
*Notes:* These figures display event-study estimates for the impact of M&As on log of total earnings for workers at target firms, separately for those involved in acquisitions and for those involved in mergers. Panel (a) displays the estimates for stayers and panel (b) displays the estimates for workers moving from target firms within the first two years after the event. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker-firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure B3: Worker-level Earnings By Within Market M&As vs. Across Market M&As (Targets)



*Notes:* These figures display event-study estimates for the impact of M&As on log of total earnings for workers at target firms, separately for M&As that happen within the same labor market and for M&As that happen across different labor markets. Panel (a) displays the estimates for stayers and panel (b) displays the estimates for workers moving from target firms within the first two years after the event. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker-firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

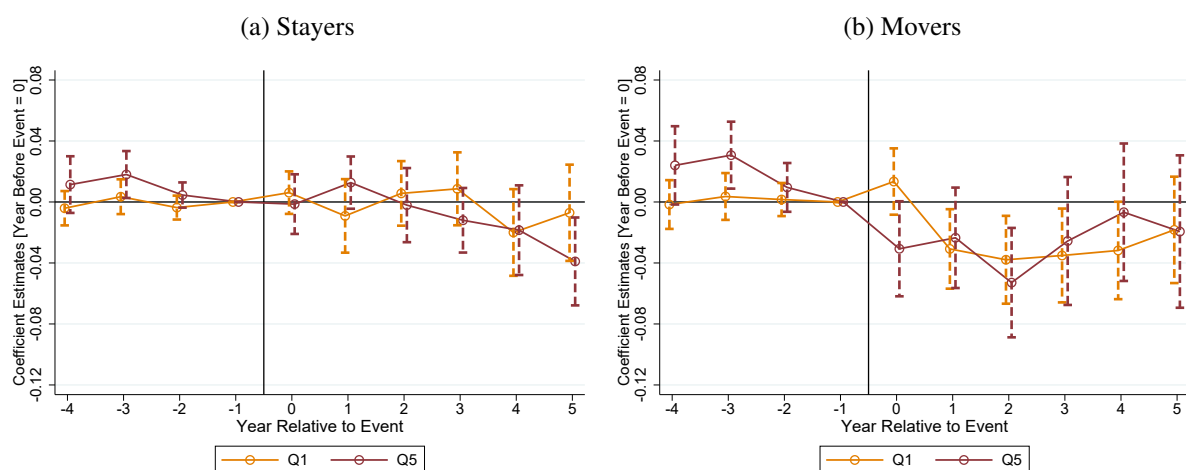
Figure B4: Distribution of HHI across markets



*Notes:* These figures display the distribution of concentration (measured by HHI) across labor markets. Panel (a) displays the distribution across all labor markets in the entire sample. Panel (b) displays the distribution across labor markets with at least one M&A deal in the entire sample during our sample period. Panel (c) displays the distribution across labor markets with at least one M&A deal in our matched sample (used for the main analyses).

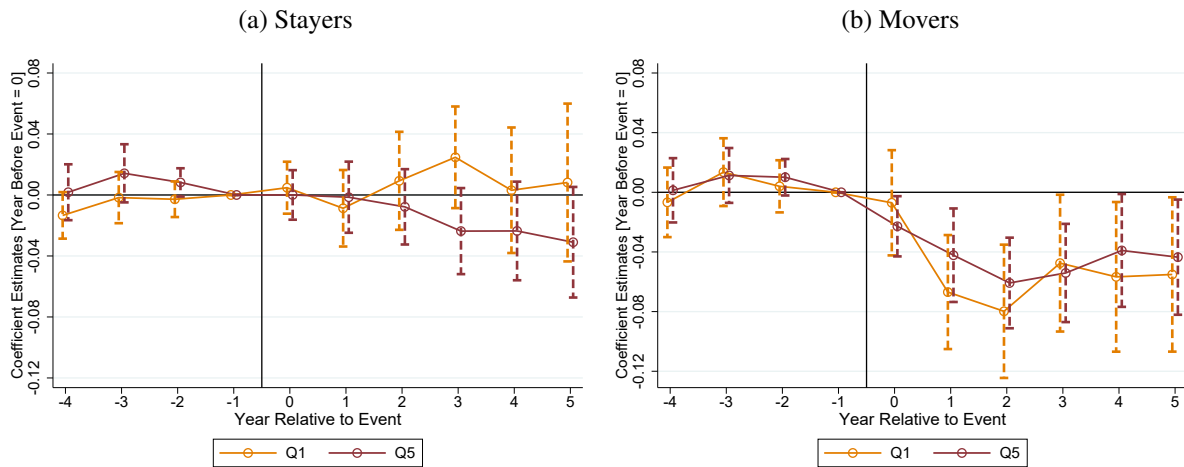


Figure B5: Worker-level Earnings By Initial Level of HHI (Targets)



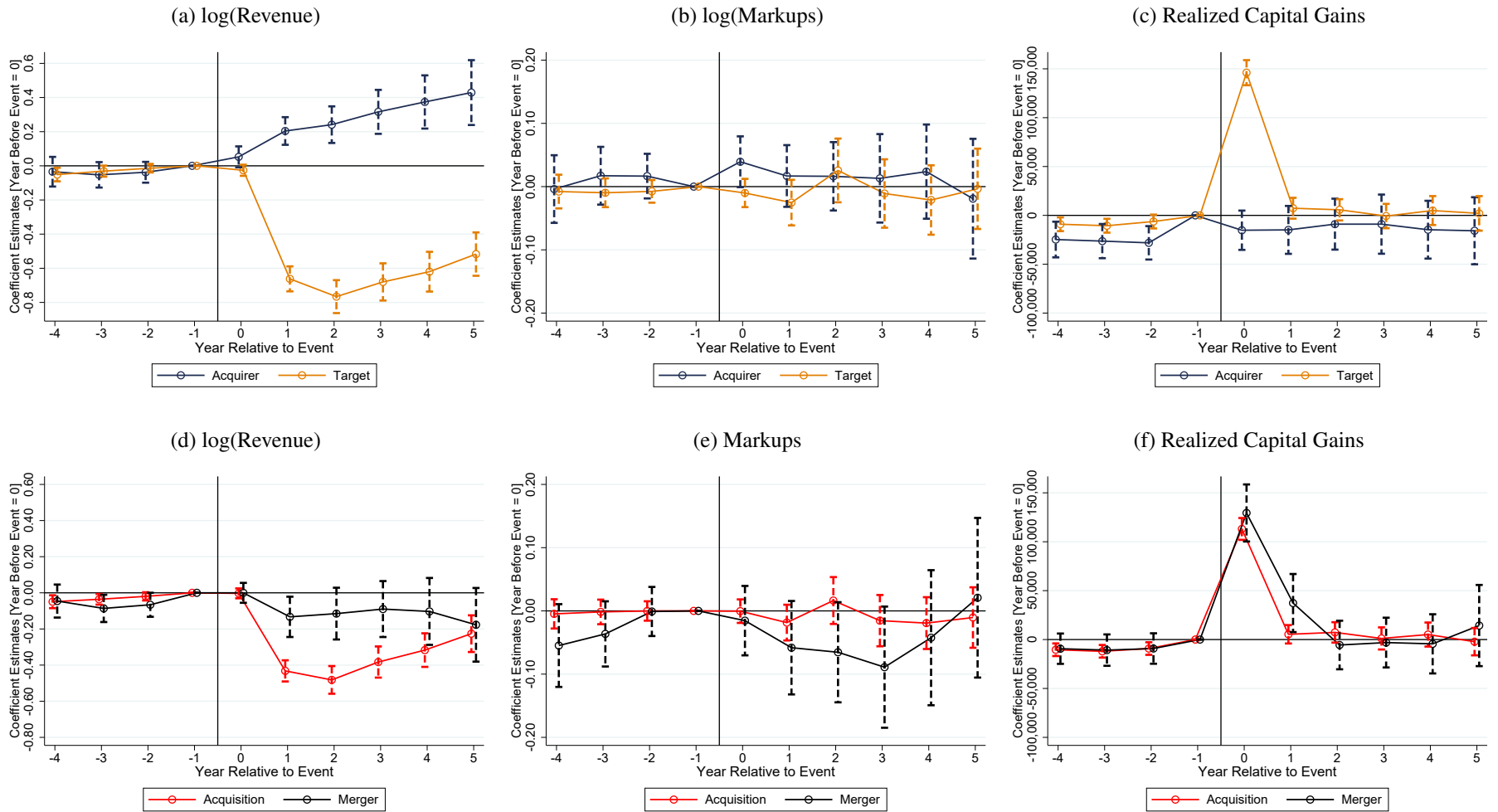
*Notes:* These figures display event-study estimates for the impact of M&As on log of total earnings for workers at target firms, separately for those in markets with low initial level of concentration (first quintile) and for those in markets with high initial level of concentration (fifth quintile). Panel (a) displays the estimates for stayers and panel (b) displays the estimates for workers moving from target firms within the first two years after the event. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker-firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure B6: Worker-level Earnings By Initial Level of Flows-adjusted HHI (Targets)



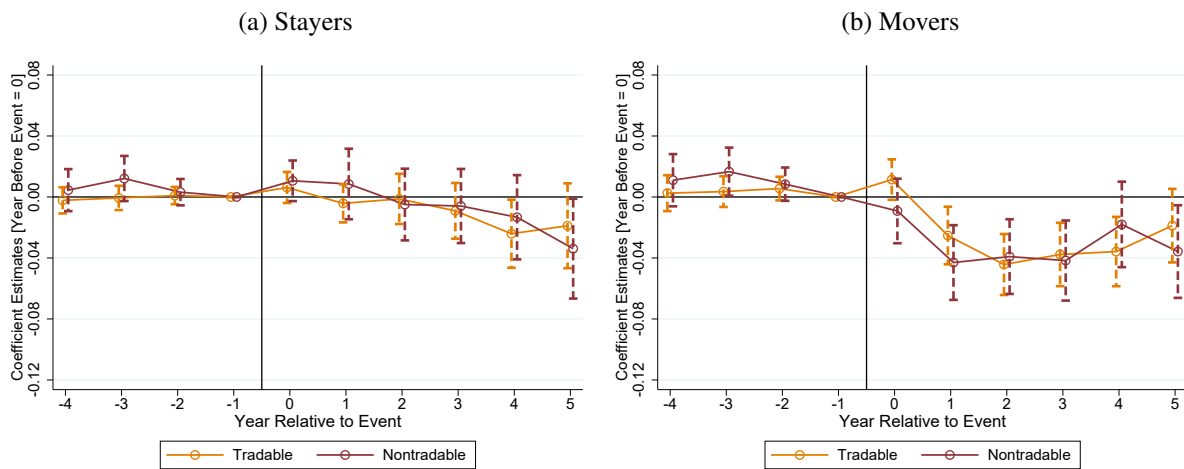
*Notes:* These figures display event-study estimates for the impact of M&As on log of total earnings for workers at target firms, separately for those in markets with low initial level of concentration (first quintile) and for those in markets with high initial level of concentration (fifth quintile). Panel (a) displays the estimates for stayers and panel (b) displays the estimates for workers moving from target firms within the first two years after the event. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker-firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure B7: Difference-in-differences Estimates on Firms' Sales, Markups, and Payouts



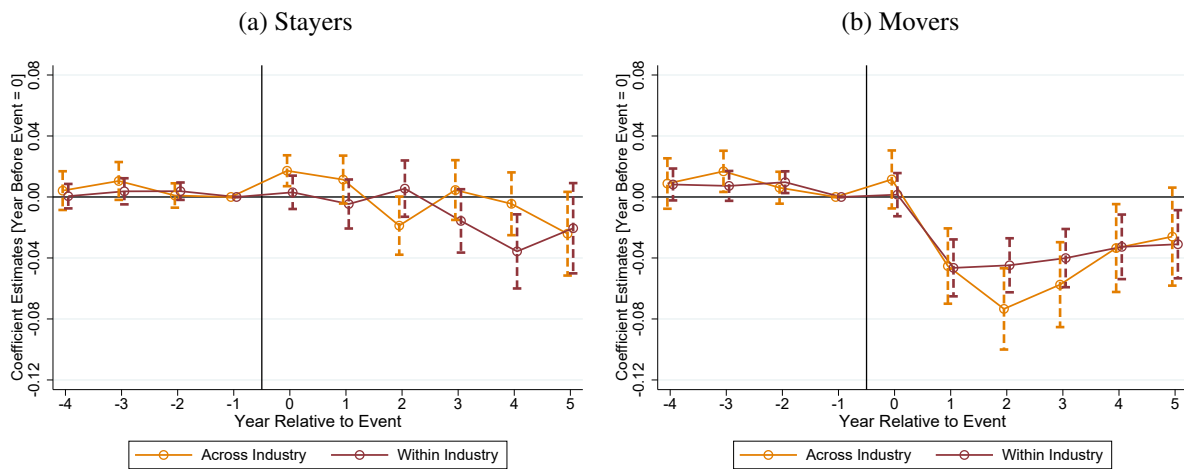
*Notes:* Panels (a) – (c) display event-study estimates for the impact of M&As on other firm-level outcomes for firms involved in acquisitions, separately for acquiring firms (navy line) and for target firms (orange line). Panel (a) shows the estimates for log of total revenue. Panel (b) shows the estimates for log of markups. Panel (c) shows the estimates for owners' realized capital gains aggregated at the firm level. Panels (d) – (f) display event-study estimates for the impact of M&As on the aggregate firm-level (targets and acquirers pooled) outcomes, separately for those involved in acquisitions (red line) and for those involved in mergers (black line). Panel (d) shows the estimates for log of total revenue. Panel (e) shows the estimates for log of markups. Panel (f) shows the estimates for owners' realized capital gains aggregated at the firm level. The dashed lines indicate 95 percent confidence intervals where the standard errors are clustered at the firm level. The M&A event is in year 0 and the coefficient is normalized to be zero in year -1.

Figure B8: Worker-level Earnings By Tradable Sectors vs. Non-tradable Sectors (Targets)



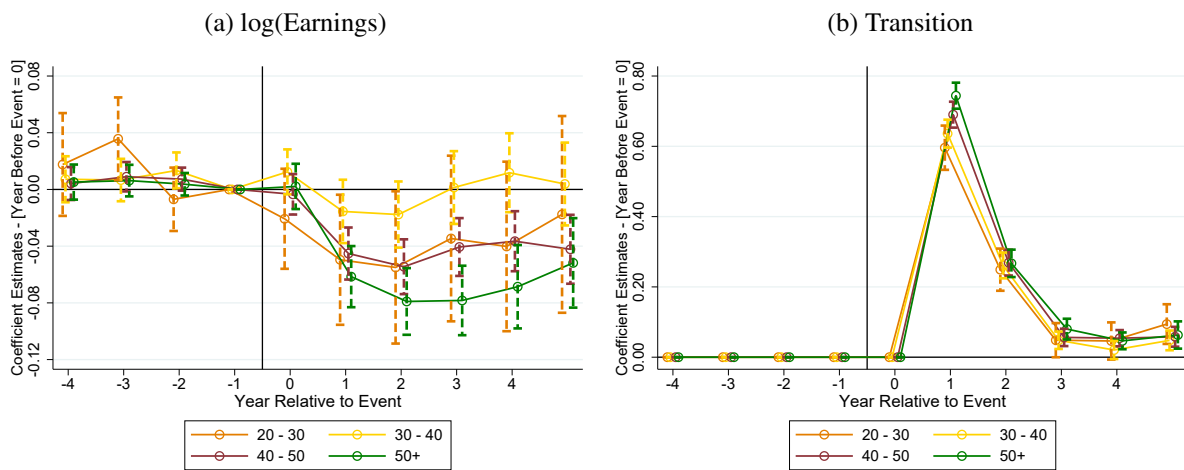
*Notes:* These figures display event-study estimates for the impact of M&As on log of total earnings for workers at target firms, separately for those in tradable sectors and for those in non-tradable sectors. Panel (a) displays the estimates for stayers and panel (b) displays the estimates for workers moving from target firms within the first two years after the event. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker-firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure B9: Worker-level Earnings By Within Industry vs. Across Industry (Targets)



*Notes:* These figures display event-study estimates for the impact of M&As on log of total earnings for workers at target firms, separately for those involved in (horizontal) M&As that happen within the same industry and for those involved in (vertical) M&As that happen across different industries. Panel (a) displays the estimates for stayers and panel (b) displays the estimates for workers moving from target firms within the first two years after the event. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker-firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Figure B10: Workers Moving from Targets – By Worker Age (Targets)



*Notes:* These figures display event-study estimates of the impact of M&As on worker-level outcomes for workers moving from target firms within the first two years after the event, separately for various age groups. Panel (a) shows the estimates for log of total earnings. Panel (b) shows the estimates for job transition probabilities. Panel (c) shows the estimates for employer fixed effects. The dashed lines indicate 95 percent confidence intervals where the standard errors are two-way clustered at the worker and firm level. The M&A event is in year 0 and the coefficient estimate is normalized to be zero in year -1.

Table B1: Worker-level Earnings By National M&As vs. Local M&As (Targets)

|                  | (1)               | (2)                  |
|------------------|-------------------|----------------------|
|                  | Stayers           | Movers               |
| National         | -0.008<br>(0.006) | -0.033***<br>(0.007) |
| Mean at t = -1   | 11.00             | 10.95                |
| Adj. R squared   | 0.799             | 0.741                |
| Worker-Year Obs. | 1,373,290         | 549,560              |
| Local            | -0.006<br>(0.030) | -0.035**<br>(0.014)  |
| Mean at t = -1   | 11.09             | 11.08                |
| Adj. R squared   | 0.790             | 0.744                |
| Worker-Year Obs. | 193,180           | 134,420              |

*Notes:* This table reports the difference-in-differences estimates of the impact of M&As on log of total earnings for workers at target firms, separately for those involved in national M&As and for those involved in local M&As. Columns (1) and column (2) display the estimates for stayers and for workers moving from target firms within the first two years after the event, respectively. The standard errors are two-way clustered at the worker and firm level.

Table B2: Worker-level Earnings By Acquisition vs. Merger (Targets)

|                  | (1)               | (2)                  |
|------------------|-------------------|----------------------|
|                  | Stayers           | Movers               |
| Acquisition      | -0.008<br>(0.006) | -0.035***<br>(0.007) |
| Mean at t = -1   | 11.01             | 10.97                |
| Adj. R squared   | 0.800             | 0.743                |
| Worker-Year Obs. | 1,365,580         | 601,180              |
| Merger           | 0.001<br>(0.013)  | -0.030*<br>(0.016)   |
| Mean at t = -1   | 11.01             | 11.02                |
| Adj. R squared   | 0.785             | 0.726                |
| Worker-Year Obs. | 216,050           | 88,190               |

*Notes:* This table reports the difference-in-differences estimates of the impact of M&As on log of total earnings for workers at target firms, separately for those involved in acquisitions and for those involved in mergers. Columns (1) and column (2) display the estimates for stayers and for workers moving from target firms within the first two years after the event, respectively. The standard errors are two-way clustered at the worker and firm level.



Table B3: Worker-level Earnings By Within Market M&As vs. Across Market M&As (Targets)

|                  | (1)               | (2)                  |
|------------------|-------------------|----------------------|
|                  | Stayers           | Movers               |
| Across Market    | -0.007<br>(0.006) | -0.034***<br>(0.007) |
| Mean at t = -1   | 11.00             | 10.95                |
| Adj. R squared   | 0.799             | 0.741                |
| Worker-Year Obs. | 1,387,150         | 547,370              |
| Within Market    | -0.005<br>(0.033) | -0.040***<br>(0.015) |
| Mean at t = -1   | 11.1              | 11.09                |
| Adj. R squared   | 0.786             | 0.743                |
| Worker-Year Obs. | 170,870           | 120,900              |

*Notes:* This table reports the difference-in-differences estimates of the impact of M&As on log of total earnings for workers at target firms, separately for those involved in M&As that happen across different labor markets (defined at the four-digit NAICS by commuting zone) and for those involved in M&As that happen within the same labor market. Columns (1) and column (2) display the estimates for stayers and for workers moving from target firms within the first two years after the event, respectively. The standard errors are two-way clustered at the worker and firm level.

Table B4: Distribution of HHI across labor markets

|                  | (1)<br>All Markets | (2)<br>Markets with M&A<br>(Entire Sample) | (3)<br>Markets with M&A<br>(Matched Sample) |
|------------------|--------------------|--|---|
| Average          | 0.3827             | 0.264                                      | 0.1383                                      |
| Median           | 0.28               | 0.1753                                     | 0.0694                                      |
| Min              | 0.0002             | 0.0003                                     | 0.0004                                      |
| Max              | 1                  | 1  | 1   |
| Market-Year Obs. | 745,849            | 177,629                                    | 10,120                                      |

*Notes:* This table reports summary statistics (average, median, min and max) on concentration, measured by HHI, (1) across all labor markets in the entire sample, (2) across labor markets with at least one M&A event in the entire sample, and (3) across labor markets with at least one M&A event in the matched analysis sample.

Table B5: Worker-level Earnings By Initial Level of HHI (Targets)

|                  | (1)               | (2)                 |
|------------------|-------------------|---------------------|
|                  | Stayers           | Movers              |
| Q1               | -0.003<br>(0.009) | -0.023**<br>(0.011) |
| Mean at t = -1   | 11.04             | 11.03               |
| Adj. R squared   | 0.814             | 0.748               |
| Worker-Year Obs. | 384,470           | 186,790             |
| Q5               | -0.010<br>(0.008) | -0.026*<br>(0.015)  |
| Mean at t = -1   | 11.03             | 11.00               |
| Adj. R squared   | 0.777             | 0.744               |
| Worker-Year Obs. | 379,700           | 122,390             |

*Notes:* This table reports the difference-in-differences estimates of the impact of M&As on log of total earnings for workers at target firms, separately for those in markets with initially low level of concentration (first quintile) and for those in markets with initially high level of concentration (fifth quintile). Columns (1) and column (2) display the estimates for stayers and for workers moving from target firms within the first two years after the event, respectively. The standard errors are two-way clustered at the worker and firm level.

Table B6: Worker Earnings - By Initial Level of Flows-Adjusted HHI (Targets)

|                  | (1)               | (2)                  |
|------------------|-------------------|----------------------|
|                  | Stayers           | Movers               |
| Q1               | 0.007<br>(0.012)  | -0.052***<br>(0.016) |
| Mean at t = -1   | 10.97             | 10.95                |
| Adj. R squared   | 0.812             | 0.735                |
| Worker-Year Obs. | 166,590           | 66,840               |
| Q5               | -0.015<br>(0.010) | -0.044***<br>(0.012) |
| Mean at t = -1   | 11.06             | 10.97                |
| Adj. R squared   | 0.792             | 0.747                |
| Worker-Year Obs. | 395,720           | 182,240              |

*Notes:* This table reports the difference-in-differences estimates of the impact of M&As on log of total earnings for workers at target firms, separately for those in markets with initially low level of concentration (first quintile) and for those in markets with initially high level of concentration (fifth quintile). Columns (1) and column (2) display the estimates for stayers and for workers moving from target firms within the first two years after the event, respectively. The standard errors are two-way clustered at the worker and firm level.

Table B7: Difference-in-differences Estimates on Firms' Sales, Markups, and Payouts

|                        | (1)                          | (2)                        | (3)                          | (4)                          |
|------------------------|------------------------------|----------------------------|------------------------------|------------------------------|
|                        | <b>Acquisition</b>           |                            | <b>Aggregate</b>             |                              |
|                        | Target                       | Acquirer                   | Acquisition                  | Merger                       |
| log(Revenue)           | -0.545***<br>(0.038)         | 0.269***<br>(0.048)        | -0.307***<br>(0.030)         | -0.103*<br>(0.058)           |
| Mean at t = -1         | 16.37                        | 17.02                      | 16.53                        | 16.45                        |
| Adj. R squared         | 0.820                        | 0.861                      | 0.829                        | 0.840                        |
| Firm-Year Obs.         | 68,460                       | 22,060                     | 91,140                       | 19,400                       |
| log(Markups)           | -0.008<br>(0.019)            | 0.015<br>(0.024)           | -0.008<br>(0.014)            | -0.042<br>(0.034)            |
| Mean at t = -1         | 0.42                         | 0.40                       | 0.42                         | 0.39                         |
| Adj. R squared         | 0.758                        | 0.811                      | 0.769                        | 0.746                        |
| Firm-Year Obs.         | 40,800                       | 12,760                     | 54,150                       | 11,170                       |
| Realized Capital Gains | 27,557.747***<br>(4,309.922) | -13,071.405<br>(9,385.609) | 21,645.781***<br>(3,735.874) | 28,002.089***<br>(9,603.254) |
| Mean at t = -1         | 37,938.34                    | 51,147.49                  | 40,770.09                    | 35,448.88                    |
| Adj. R squared         | 0.282                        | 0.209                      | 0.245                        | 0.304                        |
| Firm-Year Obs.         | 42,150                       | 10,820                     | 53,670                       | 9,410                        |

*Notes:* This table reports the difference-in-differences estimates for the impact of M&As on firms' sales, markups, and realized capital gains. Columns (1) and (2) report the estimates for target firms and for acquiring firms involved in acquisitions, respectively. Columns (3) and (4) report the estimates on the aggregate firm-level (targets and acquirers pooled) outcomes, for those involved in acquisitions and for those involved in mergers, respectively. The standard errors are clustered at the firm level.

Table B8: Worker-level Earnings By Tradable Sectors vs. Non-tradable Sectors (Targets)

|                  | (1)               | (2)                  |
|------------------|-------------------|----------------------|
|                  | Stayers           | Movers               |
| Tradable         | -0.006<br>(0.009) | -0.031***<br>(0.010) |
| Mean at t = -1   | 10.97             | 10.95                |
| Adj. R squared   | 0.788             | 0.739                |
| Worker-Year Obs. | 826,140           | 320,350              |
| Nontradable      | -0.009<br>(0.007) | -0.025***<br>(0.008) |
| Mean at t = -1   | 11.06             | 11.00                |
| Adj. R squared   | 0.804             | 0.748                |
| Worker-Year Obs. | 755,700           | 369,300              |

*Notes:* This table reports the difference-in-differences estimates of the impact of M&As on log of total earnings for workers at target firms, separately for those in tradable sectors and for those in non-tradable sectors. Columns (1) and column (2) display the estimates for stayers and for workers moving from target firms within the first two years after the event, respectively. The standard errors are two-way clustered at the worker and firm level.

Table B9: Worker-level Earnings By Within Industry vs. Across Industry (Targets)

|                  | (1)               | (2)                  |
|------------------|-------------------|----------------------|
|                  | Stayers           | Movers               |
| Across Industry  | -0.002<br>(0.007) | -0.037***<br>(0.010) |
| Mean at t = -1   | 10.97             | 10.88                |
| Adj. R squared   | 0.798             | 0.755                |
| Worker-Year Obs. | 679,760           | 233,960              |
| Within Industry  | -0.011<br>(0.007) | -0.032***<br>(0.007) |
| Mean at t = -1   | 11.04             | 11.03                |
| Adj. R squared   | 0.798             | 0.732                |
| Worker-Year Obs. | 901,990           | 455,670              |

*Notes:* This table reports the difference-in-differences estimates of the impact of M&As on log of total earnings for workers at target firms, separately for those involved in M&As that happen across different industries and for those involved in M&As that happen within the same industry. Columns (1) and column (2) display the estimates for stayers and for workers moving from target firms within the first two years after the event, respectively. The standard errors are two-way clustered at the worker and firm level.

Table B10: Workers Moving from Targets – By Worker Age (Targets)

|                  | (1)                  | (2)                 |
|------------------|----------------------|---------------------|
|                  | log(Earnings)        | Transition          |
| 20s              | -0.036*<br>(0.021)   | 0.172***<br>(0.011) |
| Mean at t = -1   | 10.69                | 0.00                |
| Adj. R squared   | 0.764                | 0.282               |
| Worker-Year Obs. | 32,390               | 32,410              |
| 30s              | -0.001<br>(0.009)    | 0.169***<br>(0.006) |
| Mean at t = -1   | 10.98                | 0.00                |
| Adj. R squared   | 0.73                 | 0.267               |
| Worker-Year Obs. | 154,250              | 154,370             |
| 40s              | -0.037***<br>(0.007) | 0.188***<br>(0.006) |
| Mean at t = -1   | 11.05                | 0.00                |
| Adj. R squared   | 0.778                | 0.333               |
| Worker-Year Obs. | 246,160              | 246,510             |
| 50+              | -0.056***<br>(0.009) | 0.200***<br>(0.007) |
| Mean at t = -1   | 10.95                | 0.00                |
| Adj. R squared   | 0.761                | 0.369               |
| Worker-Year Obs. | 256,090              | 256,800             |

*Notes:* This table reports the difference-in-differences estimates for the impacts of M&As on workers moving from target firms within the first two years after the event across various age groups. Column (1) displays the estimates log of total earnings. Column (2) displays the estimates for the transition probabilities. Column (3) displays the estimates for the employer fixed effects. The standard errors are two-way clustered at the worker and firm level.