

Homework assignment 1

1. (15 points) Suppose we toss a coin n times. Let X be a random variable denoting the number of heads among n coin tosses. We simply assume that X follows the binomial distribution with parameter θ ,

$$\mathbb{P}(X = x|\theta) = \binom{n}{x} \theta^x (1 - \theta)^{n-x}. \quad (1)$$

- (a) (5 points) We assume that θ itself is a random variable following the beta distribution with Probability Density Function (PDF) with parameters $a, b > 0$,

$$f(\theta; a, b) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} \mathbb{1}_{\{0 \leq \theta \leq 1\}}, \quad (2)$$

where $\Gamma(t) = \int_0^\infty z^{t-1} e^{-z} dz$ is the gamma function. Compute the marginal likelihood of x under this prior setting. That is, compute

$$\mathbb{P}(X = x; a, b) = \int_0^1 \mathbb{P}(X = x|\theta) f(\theta; a, b) d\theta. \quad (3)$$

- (b) (10 points) Assume we toss the coin $n = 10$ times and see $x = 9$ heads. To see if the coin is fair, we compare two prior distributions for θ :

$$\mathcal{M}_1 : \mathbb{P}(\theta) = \delta_{1/2}(\theta), \quad \mathcal{M}_2 : f(\theta; a, b) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} \mathbb{1}_{\{0 \leq \theta \leq 1\}}. \quad (4)$$

Set $a = b = 1$ and compute the Bayes factor

$$\frac{\mathbb{P}(X = x|\mathcal{M}_1)}{\mathbb{P}(X = x|\mathcal{M}_2)}. \quad (5)$$

(Hint) Use the fact that $\Gamma(n) = (n-1)!$ for all $n \in \mathbb{N}$.

2. (15 points) Let E be a set and \mathcal{A} be a collection of subsets of E . The σ -algebra generated by \mathcal{A} , denoted by $\sigma(\mathcal{A})$, is the intersection of all σ -algebras containing \mathcal{A} , or equivalently, the smallest σ -algebra containing \mathcal{A} .

- (a) (5 points) Let $E = \{1, 2, 3, 4\}$ and $\mathcal{A} = \{\{1\}, \{2\}\}$. Find $\sigma(\mathcal{A})$.

- (b) (10 points) Let E be a set and \mathcal{A}, \mathcal{C} be collections of subsets of E . Prove the followings.

1. $\mathcal{A} \subset \mathcal{C} \implies \sigma(\mathcal{A}) \subset \sigma(\mathcal{C})$.
2. $\mathcal{A} \subset \sigma(\mathcal{C}), \mathcal{C} \subset \sigma(\mathcal{A}) \implies \sigma(\mathcal{A}) = \sigma(\mathcal{C})$.
3. $\mathcal{A} \subset \mathcal{C} \subset \sigma(\mathcal{A}) \implies \sigma(\mathcal{A}) = \sigma(\mathcal{C})$.

3. (25 points) Let X be a \mathbb{R} -valued random variable. The Moment Generating Function (MGF) of X is defined as

$$M_X(t) = \mathbb{E}[e^{tX}]. \quad (6)$$

An important property of MGF is that it uniquely determines a distribution of a random variable. That is, given two \mathbb{R} -valued random variables X and Y (assume that there exist MGF for both X and Y),

$$\forall t, M_X(t) = M_Y(t) \implies \forall x, F_X(x) = F_Y(x). \quad (7)$$

One can also show that, if $(X_n)_{n \geq 1}$ is a sequence of random variables whose MGF $M_{X_n}(t)$ converges to the MGF of a random variable X , then $(X_n)_{n \geq 1}$ converges in distribution to X .

- (a) (5 points) Compute the MGF of a standard normal random variable X with PDF,

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}. \quad (8)$$

- (b) (5 points) For $n \in \mathbb{N}$, let X_1, \dots, X_n be i.i.d. \mathbb{R} -valued random variables, and let $M_{X_1}(t)$ be the MGF of X_1 . Derive the MGF of $Y_n := c(X_1 + \dots + X_n)$.
- (c) (15 points) Let X be a \mathbb{R} -valued random variable with mean μ and variance σ^2 . For $n \in \mathbb{N}$, define

$$Y_i := \frac{X_i - \mu}{\sigma}, \quad Z_n := \frac{1}{\sqrt{n}} \sum_{i=1}^n Y_i. \quad (9)$$

Show that Z_n converges in distribution to a standard normal distribution as $n \rightarrow \infty$.

(Hint) Show that the MGF of Z_n converges to the MGF of a standard normal distribution as $n \rightarrow \infty$. Use the Taylor's theorem to analyze the limit.

4. (25 points) Let X be a random variable with finite mean μ and variance σ^2 . Let X_1, \dots, X_n be i.i.d. copies of X , and consider the following estimator of σ :

$$\hat{\sigma}_n^2 := \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2, \quad \bar{X}_n := \frac{1}{n} \sum_{i=1}^n X_i. \quad (10)$$

- (a) (10 points) Show that $\hat{\sigma}_n^2$ is consistent; that is, show that $\hat{\sigma}_n^2 \xrightarrow{P} \sigma^2$ as $n \rightarrow \infty$.
- (b) (10 points) Show that $\hat{\sigma}_n^2$ is biased; that is, show that $\mathbb{E}[\hat{\sigma}_n^2] \neq \sigma^2$.
- (c) (5 points) Propose a simple fix to make $\hat{\sigma}_n^2$ unbiased.
5. (20 points) Let X be a \mathbb{R} -valued random variable. Show that $X_n \xrightarrow{d} c$ where c is a constant implies $X_n \xrightarrow{P} c$.