

# BML. 1 Intro

1.2 Means. Normal-Normal model

DataLab CSIC

# Schedule

- Week 0. Props,
- Week 1. Means, Regression
- Week 2. DLMs, Recap.

## Lab1-1

Expo-Gamma (Reliability)

Beta-binomial (Fraud detection)

Normal-normal (Ecology)

# Normal-normal model. A typical example

Consider detecting fraud in fishing activities. Only sardines of size bigger than 5.5 cm are allowed to be captured by an EU regulation.

Inspectors catch the boat *Ayuso* right at the entrance of Cedeira port. We would like to assess whether the *Ayuso* crew has respected the EU regulation.

Numerical example in Lab. Here we just go through the concepts and methods.



# Normal-normal model. A typical example

Consider detecting fraud in fishing activities. Only sardines of size bigger than 5.5 cm are allowed to be captured by an EU regulation.

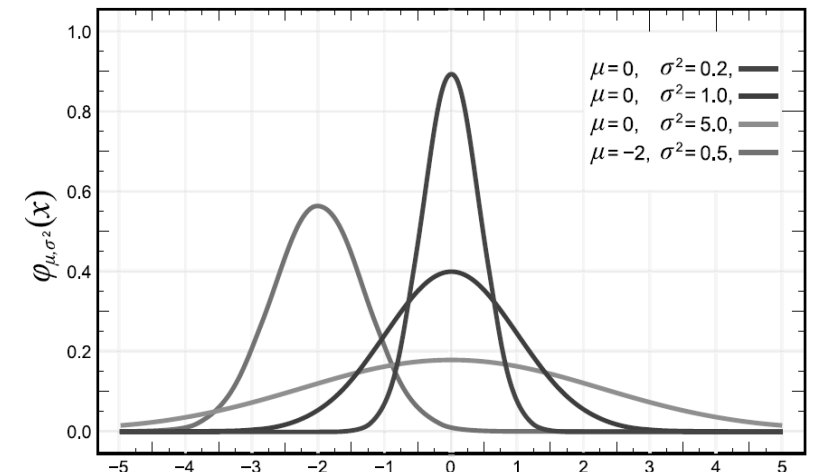
Inspectors catch the boat *Ayuso* right at the entrance of Cedeira port. We would like to assess whether the *Ayuso* crew has respected the EU regulation.

Numerical example in Lab. Here we just go through the concepts and methods.

Let's start with the model

Check

[https://en.wikipedia.org/wiki/Normal\\_distribution](https://en.wikipedia.org/wiki/Normal_distribution)



# Bayesian inference with the normal-normal model

Parameter

$$\theta$$

Model

$$X|\theta \sim N(\theta, \sigma^2) \quad f(x|\theta) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2}\left(\frac{x-\theta}{\sigma}\right)^2\right) \quad \begin{array}{l} E(X) = \theta \\ \text{Var}(X) = \sigma^2 \rightarrow \text{KNOWN!!} \end{array}$$

Data

$$x_1, x_2, \dots, x_n$$

Likelihood

$$f(x_1, x_2, \dots, x_n | \theta) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2}\left(\frac{x_i - \theta}{\sigma}\right)^2\right) \propto \exp\left(-\frac{1}{2\sigma^2} (n\theta^2 - 2\theta \sum x_i)\right) = \ell(\theta | \underline{x})$$

(MLE)

$$h(\theta) = -\log \ell(\theta | \underline{x}) = \frac{1}{2\sigma^2} (n\theta^2 - 2\theta \sum x_i) \quad h'(\theta) = 0 \Rightarrow \hat{\theta} = \frac{\sum_{i=1}^n x_i}{n} = \bar{x}$$

# Bayesian inference with the normal-normal model

Likelihood

$$l(\theta) \propto \exp\left(-\frac{1}{2\sigma^2} (n\theta^2 - 2\theta \sum x_i)\right)$$

Non-informative Prior

$$f(\theta) = 1 \rightarrow \text{IMPROPER PRIOR !!!}$$

$$\begin{aligned} f(\theta | x_1, \dots, x_n) &\propto \exp\left(-\frac{1}{2\sigma^2} (n\theta^2 - 2\theta \sum x_i)\right) \times 1 \\ &= \exp\left(-\frac{1}{2} \left(\theta^2 \frac{n}{\sigma^2} - 2\theta \frac{\sum x_i}{\sigma^2}\right)\right) \end{aligned}$$

Posterior

$$\theta \sim 1 \rightarrow N\left(\bar{x}, \frac{\sigma^2}{n}\right)$$

Sequential update

$$\mu_{n+1} = \frac{n\mu_n + x_{n+1}}{n+1} \quad \sigma_{n+1} = \frac{n\sigma_n}{n+1}$$

# Bayesian inference with the normal-normal model

Likelihood

$$l(\theta) \propto \exp\left(-\frac{1}{2\sigma^2} (n\theta^2 - 2\theta \sum x_i)\right)$$

Prior

$$\theta \sim N(\theta_0, \sigma_0^2) \quad f(\theta) = \frac{1}{\sqrt{2\pi}\sigma_0} \exp\left(-\frac{1}{2} \left(\frac{\theta - \theta_0}{\sigma_0}\right)^2\right) \propto \exp\left(-\frac{1}{2\sigma_0^2} (\theta^2 - 2\theta_0\theta)\right)$$

Posterior

Sequential update

$$\begin{aligned} f(\theta | x_1, \dots, x_n) &= \frac{l(x_1, \dots, x_n | \theta) f(\theta)}{\int l(x_1, \dots, x_n | \theta) f(\theta) d\theta} \propto l(x_1, \dots, x_n | \theta) f(\theta) \\ &\propto \exp\left(-\frac{1}{2\sigma_0^2} (\theta^2 - 2\theta_0\theta)\right) \exp\left(-\frac{1}{2\sigma^2} (n\theta^2 - 2\theta \sum x_i)\right) \\ &= \exp\left(-\frac{1}{2} \left(\theta^2 \left(\frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}\right) - 2\theta \left(\frac{\sum x_i}{\sigma^2} + \frac{\theta_0}{\sigma_0^2}\right)\right)\right) \\ \theta \sim N(\theta_0, \sigma_0^2) &\longrightarrow \theta | x_1, \dots, x_n \sim N\left(\frac{\frac{\sum x_i}{\sigma^2} + \frac{\theta_0}{\sigma_0^2}}{\frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}}, \frac{1}{\frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}}\right) \end{aligned}$$

# Bayesian inference with the normal-normal model (analytic and simulation approaches)

Point estimation

Posterior mean

Mix of prior and data

What if n grows??

$$\frac{\frac{\sum x_i}{\sigma^2} + \frac{\theta_0}{\sigma_0^2}}{\frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}} = \frac{\frac{n}{\sigma^2}}{\frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}} \frac{\sum x_i}{n} + \frac{\frac{1}{\sigma_0^2}}{\frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}} \theta_0$$

$n \rightarrow \infty \quad \approx \quad \bar{x}$

Posterior median

SAME

Posterior mode

SAME

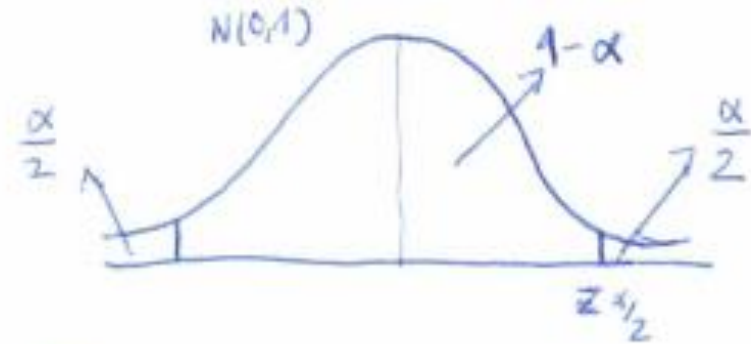


# Bayesian inference with the normal-normal model

Credible interval

Symmetric interval

HPD



$$\left[ \frac{\frac{z_{x_i}}{\sigma^2} + \frac{\theta_0}{\sigma_0^2}}{\frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{1}{\frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}}} \right]$$

SAME

# Bayesian inference with the normal-normal model

## Hypothesis testing

### Testing two lateral hypothesis

What if utility is not 0-1??

EVIDENCE SUPPORTS  $H_0$  IF

$$H_0: \theta \leq a \quad \Pr(\theta \leq a | x) \geq \Pr(\theta > a | x)$$
$$H_1: \theta > a \quad \Updownarrow$$
$$\Pr(\theta \leq a | x) \geq \frac{1}{2}$$

		IS	
		$\theta \leq a$	$\theta > a$
SAYS	$\theta \leq a$	1	0
	$\theta > a$	c	1

$$\begin{aligned} \theta \leq a &\rightarrow \Pr(\theta \leq a | x) \\ \theta > a &\rightarrow c \Pr(\theta \leq a | x) + \Pr(\theta > a | x) \end{aligned} \parallel \begin{array}{l} \text{SAY} \\ \theta \leq a \text{ IFF } \Pr(\theta \leq a | x) \geq \frac{1}{2-c} \end{array}$$

# Bayesian inference with the normal-normal model

Forecasting. The predictive distribution

$$y|\theta \sim N(\theta, \sigma^2) \quad \theta|\underline{x} \sim N(\mu_1, \sigma_1^2)$$

$$f(y|\underline{x}) = \int f(y|\theta) f(\theta|\underline{x}) d\theta$$

$$y|\underline{x} \sim N\left(\mu_1, \frac{1}{\frac{1}{\sigma^2} + \frac{1}{\sigma_1^2}}\right)$$

Summarising the predictive distribution

$$E(y|\underline{x}) = \mu_1$$

$$\text{Var}(y|\underline{x}) = \frac{1}{\frac{1}{\sigma^2} + \frac{1}{\sigma_1^2}}$$

# Pending issues

What if variance unknown?

Multivariate case?

Try them as exercises after seeing 1.3

# BML 1. Intro

## 1.3 Regression models

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# In this deck

Linear regression

(Low dimension) Logistic regression

Lab 1.2

Linear regression (ornitology)

Logistic regression (bioassay)

Dynamic linear models (hydrology)

# Linear regression model. A typical example

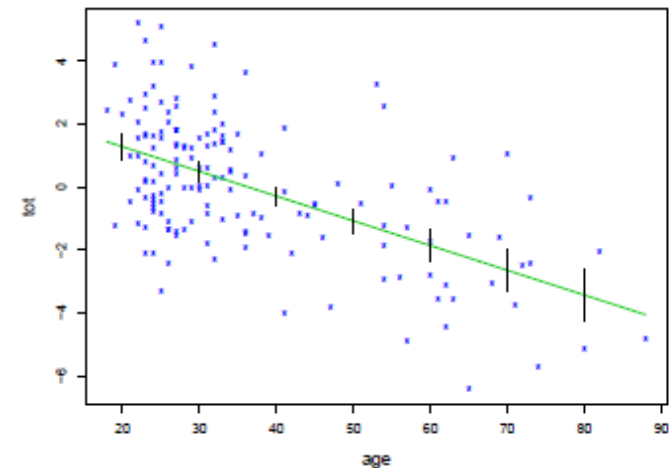
Consider a study of kidney function. The data represent ( $x$ =age of person,  $y$ =tot, a composite measure of the overall function). Kidney function declines with age and we need to provide additional information concerning decline rate. This is important in managing kidney transplant.

Numerical example in the Lab. Here we just go through the concepts and methods.

Let's start with the model

Check

[https://en.wikipedia.org/wiki/Simple\\_linear\\_regression](https://en.wikipedia.org/wiki/Simple_linear_regression)



# Linear regression model

Data structure. Response

Explanatory variables

Model

Likelihood

(MLE)

$$y$$
$$X = (x_1, \dots, x_n)$$

$$y_i = \beta_0 + \beta_1 x_{1i} + \dots + \beta_k x_{ki} \quad , \quad i = 1, \dots, n$$
$$\varepsilon_i \sim N(0, \sigma^2) \quad \text{INDEPENDENT}$$

$$\theta = (\beta_0, \beta_1, \dots, \beta_k, \sigma)$$

$$p(\theta | \underline{x}) = \left( \frac{1}{\sqrt{2\pi} \sigma} \right)^n \prod_{i=1}^n \exp \left( -\frac{1}{2} \left( \frac{y_i - \beta x_i}{\sigma} \right)^2 \right)$$

$$\hat{\beta} = (X^T X)^{-1} X^T y$$

$$s^2 = \frac{1}{n-k} (y - X\hat{\beta})^T (y - X\hat{\beta})$$



# Bayesian inference with linear regression model

Model

$$y | \beta, \sigma, X \sim N(X\beta, \sigma^2 I)$$

Standard noninformative prior

$$p(\beta, \sigma^2) \propto \sigma^{-2}$$

$$p(\beta, \sigma^2 | y) = p(\beta | \sigma^2, y) p(\sigma^2 | y)$$

$$\bullet \beta | \sigma, y \sim N(\hat{\beta}, V_{\beta} \sigma^2) \quad \begin{aligned} \hat{\beta} &= V_{\beta} X^T y \\ V_{\beta} &= (X^T X)^{-1} \end{aligned}$$

Posterior

$$\bullet p(\sigma^2 | y) = \frac{p(\beta, \sigma^2 | y)}{p(\beta | \sigma^2, y)} \sim \text{Inv-}\chi^2(m-k, s^2) \quad s^2 = \frac{1}{(n-k)} (y - X\hat{\beta})^T (y - X\hat{\beta})$$

# Bayesian inference with linear regression model

## Sampling from the posterior

COMPUTE  $\mathbf{V}_\beta, \hat{\beta}, s^2$   
DRAW  $\sigma^2 \sim \text{INV-}\chi^2(n-K, s^2)$   
DRAW  $\beta | \sigma \sim N(\hat{\beta}, \mathbf{V}_\beta \sigma^2)$

COMPUTE  $\mathbf{X} = \mathbf{QR}$   $\mathbf{Q}_{n \times k}$  ORTH. COLUMNS,  $\mathbf{R}_{k \times k}$  UPPER TRIANGULAR  
COMPUTE  $\mathbf{R}^{-1} \rightarrow \mathbf{V}_\beta = \mathbf{R}^{-1} (\mathbf{R}^{-1})^T$   
COMPUTE  $\hat{\beta} : \mathbf{R} \hat{\beta} = \mathbf{Q}^T \mathbf{y}$

# Bayesian inference with linear regression model

## Hypothesis testing for coefficients

$$\beta_i = 0?? \quad \left| \begin{array}{l} \beta_i | \sigma, y \sim N(\hat{\beta}_i, V_{\beta_i} \sigma^2) \\ \sigma^2 | y \sim \text{INV-}\chi^2(m-K, s^2) \end{array} \right.$$

Check whether

```
FOR I = 1, 1000  
  GENERATE  $\sigma_I^2 \sim \sigma^2 | y$   
  GENERATE  $\beta_I \sim \beta_i | \sigma_I, y$   
  
 $\beta_{(I)} \leftarrow \text{SORT}(\beta_I)$   
  
 $0 \in [\beta_{(50)}, \beta_{(950)}]$ 
```

# Bayesian inference with the linear regression model

Forecasting. The predictive distribution

$$p(\tilde{y} | \tilde{X}, y) = \iint p(\tilde{y} | \beta, \sigma, \tilde{X}) p(\beta, \sigma | y) d\sigma d\beta$$

$$\text{GIVEN } \sigma, y \quad \tilde{y} \sim N(\tilde{X}\hat{\beta}, (I + \tilde{X}V_p\tilde{X})\sigma^2)$$

$$\text{GIVEN } y \quad \tilde{y} \sim t_{n-k}(\tilde{X}\hat{\beta}, s^2(I + \tilde{X}V_p\tilde{X}))$$

Sampling from the predictive

$$\text{GENERATE } \beta, \sigma \sim p(\beta | \sigma, y) p(\sigma | y)$$

$$\text{GENERATE } \tilde{y} \sim p(\tilde{y} | \beta, \sigma, \tilde{X})$$

# Bayesian inference with linear regression model

Model checking (Recall discussion in IntroML)

Examining plots of residuals againsts explanatory variables,...  
interpretable as posterior predictive checks

Advantage: Using simulation, compute the posterior predictive distribution for any data summary

Example in Lab

# Pending issues

Hierarchical models

Nonlinear regression (including NNs)

To be seen

# Logistic regression. A typical example

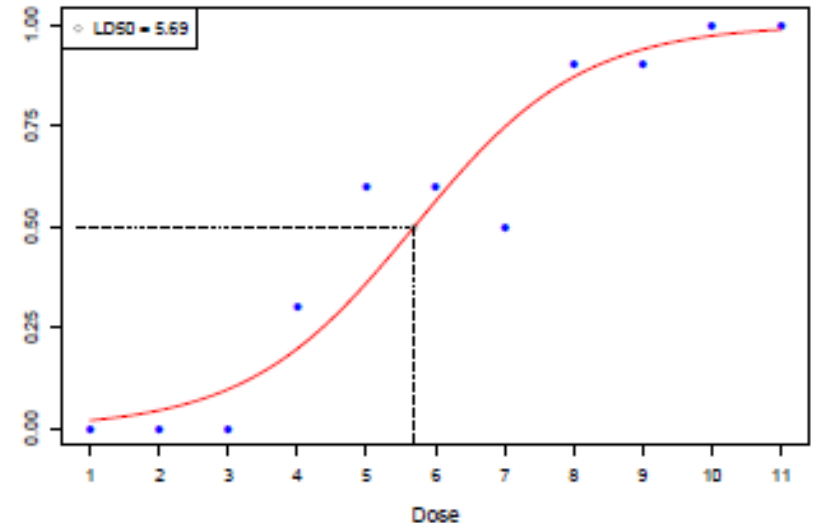
A new anti-cancer drug is being developed. Before human testing can begin, animal studies are needed to determine safe dosages. A bioassay or dose-response experiment is carried out: 11 groups of 10 mice are treated with an increasing dose of drug and the proportion of deaths are observed.

Numerical example in the Lab. Here we just go through the concepts and methods.

Let's start with the model

Check

[https://en.wikipedia.org/wiki/Logistic\\_regression](https://en.wikipedia.org/wiki/Logistic_regression)



# Bayesian inference with logit regression

Data

$$\{x_i, n_i, y_i\} \quad i=1, \dots, K$$

$x_i$ , EXP. VARIABLE  
 $n_i$ , TRIALS  
 $y_i$ , SUCCESSES.

Model

$$y_i | \theta_i \sim \text{Bin}(n_i, \theta_i) \quad \text{logit}(\theta_i) = \alpha + \beta x_i \quad \text{logit}(\theta_i) = \log \frac{\theta_i}{1-\theta_i} \quad \theta_i = \text{logit}^{-1}(\alpha + \beta x_i)$$

Likelihood

$$\ell(\alpha, \beta | x, n, y) = \prod_{i=1}^K \left[ \binom{n_i}{y_i} \text{logit}^{-1}(\alpha + \beta x_i)^{y_i} (1 - \text{logit}^{-1}(\alpha + \beta x_i))^{n_i - y_i} \right]$$

(MLE)

$$\max_{\alpha, \beta} \sum_{i=1}^K y_i \left[ \log(\text{logit}^{-1}(\alpha + \beta x_i)) + (n_i - y_i) \log(1 - \text{logit}^{-1}(\alpha + \beta x_i)) \right]$$

$\hookrightarrow \hat{\alpha}, \hat{\beta}$



# Bayesian inference with logistic regression

Likelihood

Generic prior

$$p(\alpha, \beta | y, n, \mathbf{X}) \propto p(\alpha, \beta) \prod_{i=1}^K \left[ \left( \text{logit}^{-1}(\alpha + \beta x_i) \right)^{y_i} \left( 1 - \text{logit}^{-1}(\alpha + \beta x_i) \right)^{n_i - y_i} \right]$$

Generic posterior

# Bayesian inference with logistic regression

Sampling from the joint posterior. Piecewise constant approximation

- Get some finite region where posterior is concentrated

- Choose grid over region

- Evaluate density over grid

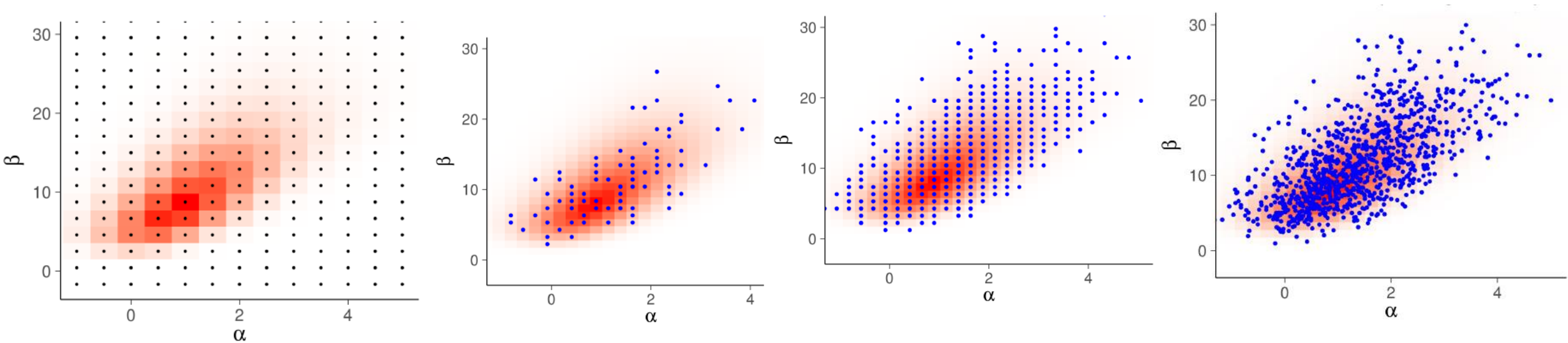
- (Normalised) Density  $\times$  Cell area approximates cell prob. in each cell

Sample according to grid cell probabilities

Jitter for visualisation

# Bayesian inference with logistic regression

## Sampling from the joint posterior



# Bayesian inference with logistic regression

Posteriors of derived quantities. The case of LD50, dose level at which death probability is 50%

$$\begin{aligned} E\left(\frac{y}{n}\right) &= \theta = \text{logit}^{-1}(\alpha + \beta x) = 0.5 \\ \alpha + \beta x &= \text{logit}(0.5) = 0 \\ x_{\text{LD50}} &= -\frac{\alpha}{\beta} \\ E(x_{\text{LD50}} \mid y, n, x) &\approx \frac{1}{S} \sum_{s=1}^S \frac{\alpha^{(s)}}{\beta^{(s)}} \\ (\alpha^{(s)}, \beta^{(s)})_{s=1}^S &\sim p(\alpha, \beta \mid y, n, x) \end{aligned}$$

# Pending issues

Logistic multiple regression

Hierarchical models

Other GLMs (e.g., Probit)

To be seen

# BML. 1 Intro

## 1.4 Dynamic Linear Models

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# In this deck

## Dynamic Linear Models

### Lab 1.2

Linear regression (ornitology)

Logistic regression (bioassay)

## Dynamic linear models (hydrology)

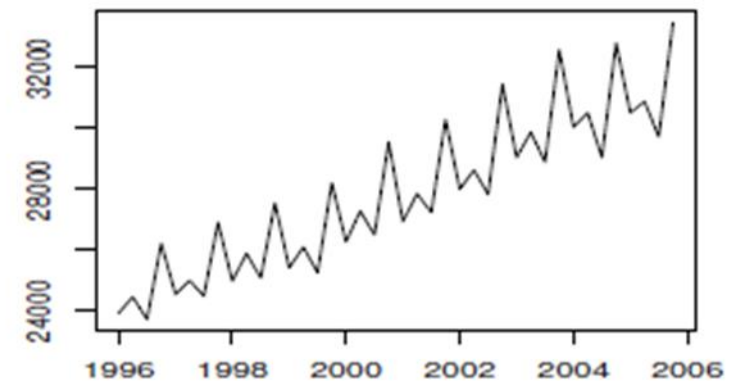
# Dynamic linear models. A typical example

We have available quarterly consumptions of gas in Spain. We need to forecast consumption over the next year to plan production.

Numerical example in the Lab. Here we just go through the concepts and methods.

Check e.g.

<https://www.math.unm.edu/~ghuerta/tseries/dlmch2.pdf>





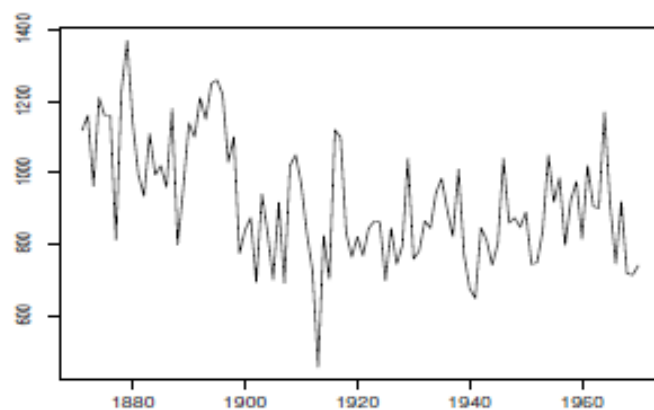
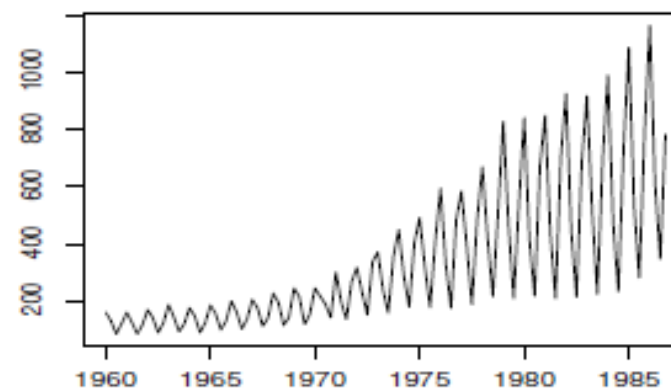
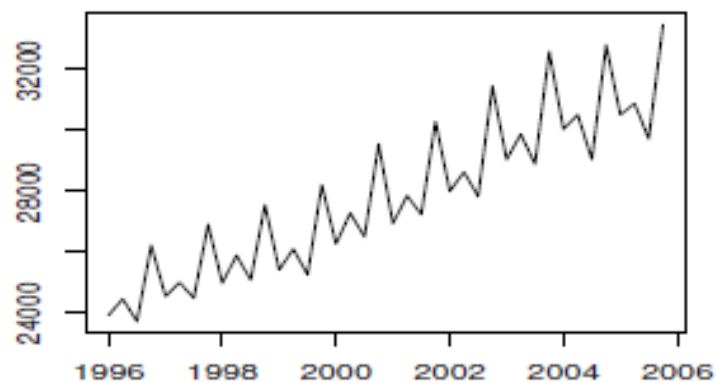
# Time series: three approaches

- Models in time domain: ARIMA. Box Jenkins
  - Models in frequency domain. Spectral analysis.
  - State space models. Series as output of dynamic systems subject to random perturbations.
- 
- Recurrent NNs, LSTMs, GRUs, Transformers,...

# State space models: Advantages

- Natural interpretation through combination of components: trend, seasonal, regression, autoregression
- Powerful probabilistic structure facilitating flexible modelling in numerous domains
- Recursive computation
- Bayesian treatment naturally
- Uni and multivariate
- Non stationarity, structural changes, irregular patterns naturally
- Linear (Kalman filter) and non-linear (MCMC).
- Very useful in applications

# Time series features



# State space models: simple example SLAM

Object position from measurements subject to random errors

$$\theta \quad (Y_t : t = 1, 2, \dots)$$

$$Y_t = \theta + \epsilon_t, \quad \epsilon_t \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$$

$$Y_1, Y_2, \dots | \theta \stackrel{iid}{\sim} \mathcal{N}(\theta, \sigma^2).$$

$$\theta \sim \mathcal{N}(m_0, C_0).$$

$$\theta | y_{1:n} \sim \mathcal{N}(m_n, C_n),$$

$$m_n = E(\theta | y_{1:n}) = \frac{C_0}{C_0 + \sigma^2/n} \bar{y} + \frac{\sigma^2/n}{C_0 + \sigma^2/n} m_0$$

$$C_n = \text{Var}(\theta | y_{1:n}) = \left( \frac{n}{\sigma^2} + \frac{1}{C_0} \right)^{-1} = \frac{\sigma^2 C_0}{\sigma^2 + n C_0}.$$

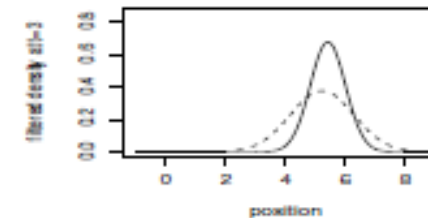
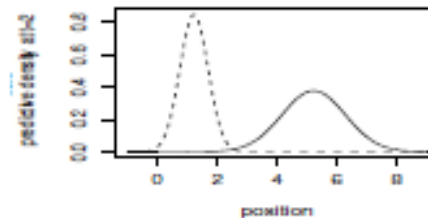
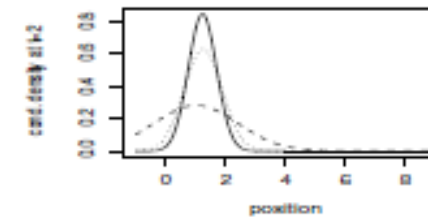
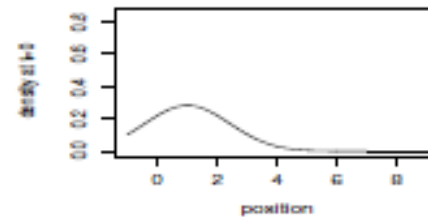
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$$\begin{aligned} m_n &= \frac{C_{n-1}}{C_{n-1} + \sigma^2} y_n + \left( 1 - \frac{C_{n-1}}{C_{n-1} + \sigma^2} \right) m_{n-1} \\ &= m_{n-1} + \frac{C_{n-1}}{C_{n-1} + \sigma^2} (y_n - m_{n-1}) \end{aligned}$$

$$C_n = \left( \frac{1}{\sigma^2} + \frac{1}{C_{n-1}} \right)^{-1} = \frac{\sigma^2 C_{n-1}}{\sigma^2 + C_{n-1}}.$$

# State space models: Simple example

Time	Observation	Mean	Variance
0	-	1	2
1	1.3	1.24	0.4
2	1.2	1.222	0.222
3	.....	.....	.....



# Simple example

- Introduce dynamic component. From  $t=2$ , object adopts a velocity

With in  $\theta_t = \theta_{t-1} + \nu + w_t, \quad w_t \sim \mathcal{N}(0, \sigma_w^2), \quad Y_t = \theta_t + \epsilon_t, \quad \epsilon_t \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$ .

- Initial step:

$$\theta_2|y_{1:2} \sim \mathcal{N}(m_2 = 1.222, C_2 = 0.222).$$

- Prediction:

$$\theta_3|y_{1:2} \sim \mathcal{N}(a_3, R_3)$$

$$Y_3|y_{1:2} \sim \mathcal{N}(f_3, Q_3).$$

$$a_3 = E(\theta_2 + \nu + w_3|y_{1:2}) = m_2 + \nu = 5.722$$

$$f_3 = E(\theta_3 + \epsilon_3|y_{1:2}) = a_3 = 5.722$$

$$R_3 = \text{Var}(\theta_2 + \nu + w_3|y_{1:2}) = C_2 + \sigma_w^2 = 1.122 \quad Q_3 = \text{Var}(\theta_3 + \epsilon_3|y_{1:2}) = R_3 + \sigma^2 = 1.622$$

# Simple example

- Initial step:

$$\theta_2|y_{1:2} \sim \mathcal{N}(m_2 = 1.222, C_2 = 0.222).$$

- Forecast:

$$\theta_3|y_{1:2} \sim \mathcal{N}(a_3, R_3)$$

$$Y_3|y_{1:2} \sim \mathcal{N}(f_3, Q_3).$$

$$a_3 = E(\theta_2 + \nu + w_3|y_{1:2}) = m_2 + \nu = 5.722$$

$$f_3 = E(\theta_3 + \epsilon_3|y_{1:2}) = a_3 = 5.722$$

$$R_3 = \text{Var}(\theta_2 + \nu + w_3|y_{1:2}) = C_2 + \sigma_w^2 = 1.122$$

$$Q_3 = \text{Var}(\theta_3 + \epsilon_3|y_{1:2}) = R_3 + \sigma^2 = 1.622$$

- Filter:

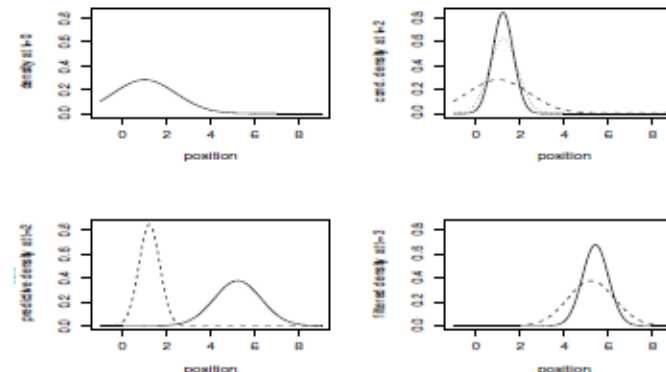
$$Y_3 = 5 \quad e_t = y_t - f_t = -0.722$$

$$\theta_3|y_1, y_2, y_3 \sim \mathcal{N}(m_3, C_3),$$

$$m_3 = a_3 + \frac{R_3}{R_3 + \sigma^2}(y_3 - f_3) = 5.568$$

$$C_3 = \frac{\sigma^2 R_3}{\sigma^2 + R_3} = R_3 - \frac{R_3}{R_3 + \sigma^2} R_3 = 0.346$$

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# Lessons from introductory example

- Observable process determined by latent process (except for normal error)
  - Latent process depends on previous state linearly (except for normal error)
  - Forecast and estimation sequentially as data arrives
- 
- Linearity and normality define dynamic linear models (DLMs)
  - Time dependence structure, define state space models



# State space models. Definition

- p-variate and m-variate time series  $(\theta_t)_{t \in \mathbb{N}} \cup Y_t$  that

- $(\theta_t)$  is Markov chain

- Given  $(\theta_t)$   $Y_t$  are independent and just depends on

$\theta_t$

$$\begin{array}{ccccccc}
 \theta_0 & \longrightarrow & \theta_1 & \longrightarrow & \theta_2 & \longrightarrow & \cdots \longrightarrow \theta_{t-1} \longrightarrow \theta_t \longrightarrow \theta_{t+1} \longrightarrow \cdots \\
 & & \downarrow & & \downarrow & & \downarrow & \downarrow & \downarrow \\
 & & Y_1 & & Y_2 & & Y_{t-1} & Y_t & Y_{t+1}
 \end{array}$$

$$\pi(\theta_{0:t}, y_{1:t}) = \pi(\theta_0) \cdot \prod_{j=1}^t \pi(\theta_j | \theta_{j-1}) \pi(y_j | \theta_j)$$

# Dynamic linear models. Definition

- State space model with

- Observation equation

$$Y_t = F_t \theta_t + v_t,$$

$$v_t \sim \mathcal{N}_m(0, V_t)$$

- State equation

$$\theta_t = G_t \theta_{t-1} + w_t,$$

$$w_t \sim \mathcal{N}_p(0, W_t)$$

- Prior

$$\theta_0 \sim \mathcal{N}_p(m_0, C_0)$$

# DLM: Random walk with noise

## AKA: Local level model

- State space with

$$\begin{aligned} Y_t &= \mu_t + v_t, & v_t &\sim \mathcal{N}(0, V) \\ \mu_t &= \mu_{t-1} + w_t, & w_t &\sim \mathcal{N}(0, W) \end{aligned}$$

# DLM: Linear growth model

- State space with

$$Y_t = \mu_t + v_t, \quad v_t \sim \mathcal{N}(0, V),$$

$$\mu_t = \mu_{t-1} + \beta_{t-1} + w_{t,1}, \quad w_{t,1} \sim \mathcal{N}(0, \sigma_\mu^2)$$

$$\beta_t = \beta_{t-1} + w_{t,2}, \quad w_{t,2} \sim \mathcal{N}(0, \sigma_\beta^2)$$

# Problems

Target

$$\pi(\theta_s | y_{1:t})$$

Filtering

$$s = t$$

Forecasting

$$s > t$$

Smooth

$$s < t$$

$$\begin{array}{ccccccc} \theta_0 & \longrightarrow & \theta_1 & \longrightarrow & \theta_2 & \longrightarrow & \dots \longrightarrow \theta_{t-1} & \longrightarrow & \theta_t & \longrightarrow & \theta_{t+1} & \longrightarrow & \dots \\ & & \downarrow & & \downarrow & & & & \downarrow & & \downarrow & & \downarrow \\ & & Y_1 & & Y_2 & & & & Y_{t-1} & & Y_t & & Y_{t+1} \end{array}$$

# Estimation and Forecasting

- Predictive density, for state one step ahead

$$\pi(\theta_t|y_{1:t-1}) = \int \pi(\theta_t|\theta_{t-1})\pi(\theta_{t-1}|y_{1:t-1}) d\theta_{t-1}$$

- Predictive density

$$\pi(y_t|y_{1:t-1}) = \int \pi(y_t|\theta_t)\pi(\theta_t|y_{1:t-1}) d\theta_t$$

- Filtering density

$$\pi(\theta_t|y_{1:t}) = \frac{\pi(y_t|\theta_t)\pi(\theta_t|y_{1:t-1})}{\pi(y_t|y_{1:t-1})}$$

# Estimation and forecasting with DLM (Kalman filter)

$$\theta_{t-1}|y_{1:t-1} \sim \mathcal{N}(m_{t-1}, C_{t-1})$$

- Currently,
- State predictive density one step ahead, normal

$$a_t = \mathbb{E}(\theta_t|y_{1:t-1}) = G_t m_{t-1},$$
$$R_t = \text{Var}(\theta_t|y_{1:t-1}) = G_t C_{t-1} G_t' + W_t$$

- Predictive density, normal

$$f_t = \mathbb{E}(Y_t|y_{1:t-1}) = F_t a_t,$$
$$Q_t = \text{Var}(Y_t|y_{1:t-1}) = F_t R_t F_t' + V_t$$

- Filtering density

$$m_t = \mathbb{E}(\theta_t|y_{1:t}) = a_t + R_t F_t' Q_t^{-1} e_t,$$
$$C_t = \text{Var}(\theta_t|y_{1:t}) = R_t - R_t F_t' Q_t^{-1} F_t R_t$$
$$e_t = Y_t - f_t$$

# K steps ahead forecasts

- Looking ahead in the future

$$\begin{array}{ccccccc} \theta_t & \longrightarrow & \theta_{t+1} & \longrightarrow & \dots & \longrightarrow & \theta_{t+k} \\ | & & & & & & | \\ Y_{1:t} & & & & & & Y_{t+k} \end{array}$$

- State k steps ahead

$$\pi(\theta_{t+k}|y_{1:t}) = \int \pi(\theta_{t+k}|\theta_{t+k-1})\pi(\theta_{t+k-1}|y_{1:t}) d\theta_{t+k-1}$$

- Observation k steps ahead

$$\pi(y_{t+k}|y_{1:t}) = \int \pi(y_{t+k}|\theta_{t+k})\pi(\theta_{t+k}|y_{1:t}) d\theta_{t+k}$$



# K steps ahead prediction with DLM

- State, normal with mean and variance

$$\begin{aligned}a_t(k) &= G_{t+k}a_{t,k-1}, \\ R_t(k) &= G_{t+k}R_{t,k-1}G'_{t+k} + W_{t+k}\end{aligned}$$

- Observation, normal with mean and variance

$$\begin{aligned}f_t(k) &= F_{t+k}a_t(k), \\ Q_t(k) &= F_{t+k}R_t(k)F'_{t+k} + V_t\end{aligned}$$

# Model validation

- Prediction errors

$$e_t = Y_t - \mathbb{E}(Y_t | y_{1:t-1}) = Y_t - f_t$$
$$e_t = Y_t - F_t a_t = F_t \theta_t + v_t - F_t a_t$$
$$= F_t (\theta_t - a_t) + v_t.$$

- Properties facilitating validation

- Expected value is 0
- Observations from gaussian process
- White noise gaussian process

$$\tilde{e}_t = e_t / \sqrt{Q_t}$$

# Model specification. Superposition principle

- The sum of independent DLMs is a DLM

$$Y_t = Y_{1,t} + \cdots + Y_{h,t}$$

$$\begin{aligned} Y_{i,t} &= F_{i,t}\theta_{i,t} + v_{i,t}, & v_{i,t} &\sim \mathcal{N}(0, V_{i,t}), \\ \theta_{i,t} &= G_{i,t}\theta_{i,t-1} + w_{i,t}, & w_{i,t} &\sim \mathcal{N}(0, W_{i,t}) \end{aligned}$$

results in

$$\begin{aligned} Y_t &= F_t\theta_t + v_t, & v_t &\sim \mathcal{N}(0, V_t), \\ \theta_t &= G_t\theta_{t-1} + w_t, & w_t &\sim \mathcal{N}(0, W_t), \end{aligned}$$

$$\theta_t = \begin{bmatrix} \theta_{1,t} \\ \vdots \\ \theta_{h,t} \end{bmatrix} \quad F_t = [F_{1,t} \mid \cdots \mid F_{h,t}] \quad G_t = \begin{bmatrix} G_{1,t} & & \\ & \ddots & \\ & & G_{h,t} \end{bmatrix} \quad W_t = \begin{bmatrix} W_{1,t} & & \\ & \ddots & \\ & & W_{h,t} \end{bmatrix} \quad V_t = \sum_{i=1}^h V_{i,t}$$

Model building strategy based on blocks

# Trend models (polynomial)

- Order n

$$\begin{cases} Y_t = \theta_{t,1} + v_t \\ \theta_{t,j} = \theta_{t-1,j} + \theta_{t-1,j+1} + w_{t,j} & j = 1, \dots, n-1 \\ \theta_{t,n} = \theta_{t-1,n} + w_{t,n}. \end{cases}$$

$$\begin{aligned} F &= (1, 0, \dots, 0) \\ G &= \begin{bmatrix} 1 & 1 & 0 & \dots & 0 \\ 0 & 1 & 1 & 0 & \dots & 0 \\ \vdots & & & \ddots & \vdots \\ 0 & \dots & 0 & 1 & 1 \\ 0 & \dots & & 0 & 1 \end{bmatrix} \\ W &= \text{diag}(W_1, \dots, W_n). \end{aligned}$$

- Predictive function

$$f_t(k) = E(Y_{t+k} | y_{1:t}) = a_{t,0} + a_{t,1}k + \dots + a_{t,n-1}k^{n-1}$$

- Consider cases n=1,2,3

# Seasonal models

- State vector dimension  $s-1$

$$F = (1, 0, \dots, 0)$$

$$G = \begin{bmatrix} -1 & -1 & \dots & -1 & -1 \\ 1 & 0 & & 0 & 0 \\ 0 & 1 & & 0 & 0 \\ & & \ddots & & \\ 0 & 0 & & 1 & 0 \end{bmatrix}.$$

# (Dynamic) regression models

- From the standard regression model

$$Y_t = \beta_1 + \beta_2 x_t + \epsilon_t, \quad \epsilon_t \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$$

- To the dynamic regression model

$$Y_t = \beta_{1,t} + \beta_{2,t} x_t + \epsilon_t, \quad \epsilon_t \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$$

$$\begin{aligned} Y_t &= x_t' \theta_t + v_t, & v_t &\sim \mathcal{N}(0, \sigma_t^2) \\ \theta_t &= G_t \theta_{t-1} + w_t, & w_t &\sim \mathcal{N}_p(0, W_t) \end{aligned}$$

# Modelling strategy

- Trend + Seasonal + Regression term+ Low order AR term (1)

$\theta_t$

`d1mModPoly(2) + d1mModSeas(4)`

# Some pending issues

- What if  $V$ ,  $W$  unknown?
- What if non-normal?
- What if non-linear?

Some discussion in Lab1-2. A lot of discussion in later chapters