

1 Forecasting biodiversity in breeding birds
2 using best practices

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Abstract

Biodiversity forecasts are important for conservation, management, and evaluating how well current models characterize natural systems. While the number of forecasts for biodiversity is increasing, there is little information available on how well these forecasts work. Most biodiversity forecasts are not evaluated to determine how well they predict future diversity, fail to account for uncertainty, and do not use time-series data that captures the actual dynamics being studied. We addressed these limitations by using best practices to explore our ability to forecast the species richness of breeding birds in North America. We used hindcasting to evaluate six different modeling approaches for predicting richness. Hindcasts for each method were evaluated annually for a decade at 1,237 sites distributed throughout the continental United States. While each model could explain most of the variance in richness, none of them consistently outperformed a baseline model that predicted constant richness at each site. In particular, we found no evidence that current methods (such as species distribution models) can successfully turn spatial data into useful temporal predictions about biodiversity at decadal time-scales. The best practices implemented in this study directly influence the forecasts, the relative performance of different modeling approaches, and the conclusions about the current state of biodiversity forecasting . To facilitate the rapid improvement of biodiversity forecasts, we emphasize the value of specific best practices in making forecasts and evaluating forecasting methods.

Introduction

Forecasting the future state of ecological systems is increasingly important for planning and management, and also for quantitatively evaluating how well ecological models capture the key processes governing natural systems (Clark et al. 2001, Dietze 2017, Houlahan et al. 2017). Forecasts regarding biodiversity are especially important, due to

35 biodiversity's central role in conservation planning and its sensitivity to anthropogenic
36 effects (Cardinale et al. 2012, Díaz et al. 2015, Tilman et al. 2017). High-profile studies
37 forecasting large biodiversity declines over the coming decades have played a large role
38 in shaping ecologists' priorities (as well as those of policymakers; e.g. IPCC 2014), but
39 it is inherently difficult to evaluate such long-term predictions before the projected
40 biodiversity declines have occurred.

41 Previous efforts to predict future patterns of species richness, and diversity more
42 generally, have focused primarily on building species distributions models (SDMs;
43 Thomas et al. 2004, Thuiller et al. 2011, Urban 2015). In general, these models
44 describe individual species' occurrence patterns as functions of the environment. Given
45 forecasts for environmental conditions, these models can predict where each species
46 will occur in the future. These species-level predictions are then combined ("stacked")
47 to generate forecasts for species richness (e.g. Calabrese et al. 2014). Alternatively,
48 models that directly relate spatial patterns of species richness to environment conditions
49 have been developed and generally perform equivalently to stacked SDMs (Algar et al.
50 2009, Distler et al. 2015). This approach is sometimes referred to as "macroecological"
51 modeling, because it models the larger-scale pattern (richness) directly (Distler et al.
52 2015).

53 Despite the emerging interest in forecasting species richness and other aspects of
54 biodiversity (Jetz et al. 2007, Thuiller et al. 2011), little is known about how effectively
55 we can anticipate these dynamics. This is due in part to the long time scales over which
56 many ecological forecasts are applied (and the resulting difficulty in assessing whether
57 the predicted changes occurred; Dietze et al. 2016). What we do know comes from a
58 small number of hindcasting studies, where models are built using data on species
59 occurrence and richness from the past and evaluated on their ability to predict
60 contemporary patterns (e.g., Algar et al. 2009, Distler et al. 2015). These studies are a
61 valuable first step, but lack several components that are important for developing

62 forecasting models with high predictive accuracy, and for understanding how well
63 different methods can predict the future. These “best practices” for effective forecasting
64 and evaluation (Box 1) broadly involve: 1) expanding the use of data to include
65 biological and environmental time-series (Tredennick et al. 2016); 2) accounting for
66 uncertainty in observations and processes, (Yu et al. 2010, Harris 2015); and 3)
67 conducting meaningful evaluations of the forecasts by hindcasting, archiving short-term
68 forecasts, and comparing forecasts to baselines to determine whether the forecasts are
69 more accurate than assuming the system is basically static (Perretti et al. 2013).

70 In this paper, we attempt to forecast the species richness of breeding birds at over 1,200
71 of sites located throughout North America, while following best practices for ecological
72 forecasting (Box 1). To do this, we combine 32 years of time-series data on bird
73 distributions from annual surveys with monthly time-series of climate data and
74 satellite-based remote-sensing. Datasets that span a time scale of 30 years or more have
75 only recently become available for large-scale time-series based forecasting. A dataset
76 of this size allows us to model and assess changes a decade or more into the future in
77 the presence of shifts in environmental conditions on par with predicted climate change.
78 We compare traditional distribution modeling based approaches to spatial models of
79 species richness, time-series methods, and two simple baselines that predict constant
80 richness for each site, on average (Figure 1). All of our forecasting models account for
81 uncertainty and observation error, are evaluated across different time lags using
82 hindcasting, and are publicly archived to allow future assessment. We discuss the
83 implications of these practices for our understanding of, and confidence in, the resulting
84 forecasts, and how we can continue to build on these approaches to improve ecological
85 forecasting in the future.

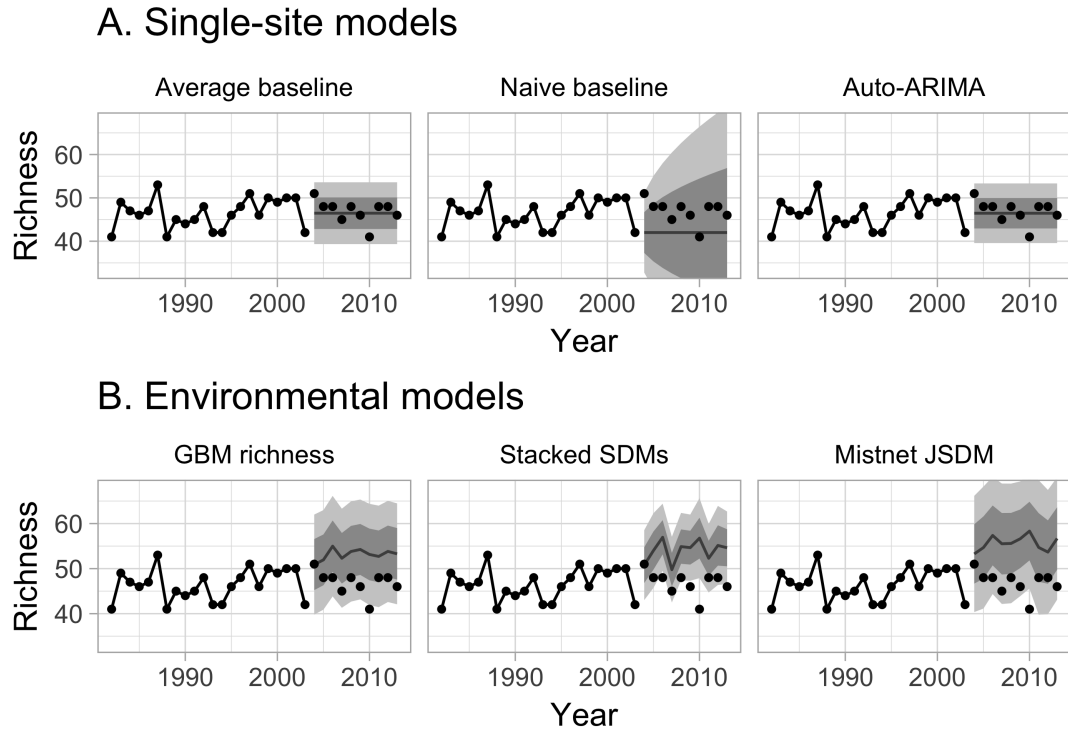


Figure 1: Example predictions from six forecasting models for a single site. Data from 1982 through 2003, connected by solid lines, were used for training the models; the remaining points were used for evaluating the models' forecasts. In each panel, point estimates for each year are shown with lines; the darker ribbon indicates the 68% prediction interval (1 standard deviation of uncertainty), and the lighter ribbon indicates the 95% prediction interval. **A.** Single-site models were trained independently on each site's observed richness values. The first two models ("average" and "naive") served as baselines. **B.** The environmental models were trained to predict richness based on elevation, climate, and NDVI; the environmental models' predictions change from year to year as environmental conditions change.

86 **Methods**

87 We evaluated 6 types of forecasting models (Table 1) by dividing the 32 years of data
88 into 22 years of training data and 10 years of data for evaluating forecasts using
89 hindcasting. Here we use definitions from meteorology, where a hindcast is generally
90 any prediction for an event that has already happened, while forecasts are predictions
91 for actual future events (Jolliffe and Stephenson 2003). We also made long term
92 forecasts by using the full data set for training and making forecasts through the year
93 2050. For both time frames, we made forecasts using each model with and without
94 correcting for observer effects, as described below.

95 **Data**

96 **Richness data.** Bird species richness was obtained from the North American Breeding
97 Bird Survey (BBS) (Pardieck et al. 2017) using the Data Retriever Python package
98 (Morris and White 2013) and rdataretriever R package (McGlinn et al. 2017). BBS
99 observations are three-minute point counts made at 50 fixed locations along a 40km
100 route. Here we denote each route as a site and summarize richness as the total species
101 observed at all 50 locations in each surveyed year. Prior to summarizing the data was
102 filtered to exclude all nocturnal, crepuscular, and aquatic species (since these species are
103 not well sampled by BBS methods; Hurlbert and White 2005), as well as unidentified
104 species, and hybrids. All data from surveys that did not meet BBS quality criteria were
105 also excluded.

106 We used observed richness values from 1982 (the first year of complete environmental
107 data) to 2003 to train the models, and from 2004 to 2013 to test their performance. We
108 only used BBS routes from the continental United States (i.e. routes where climate data
109 was available PRISM Climate Group (2004)), and we restricted the analysis to routes
110 that were sampled during 70% of the years in the training period (i.e., routes with at

111 least 16 annual observations). The resulting dataset included 34,494 annual surveys of
112 1,279 unique sites, and included 385 species. Site-level richness varied from 8 to 91
113 with an average richness of 51 species.

114 **Past environmental data.** Environmental data included a combination of elevation,
115 bioclimatic variables and a remotely sensed vegetation index (the normalized difference
116 vegetation index; NDVI), all of which are known to influence richness and distribution
117 in the BBS data (Kent et al. 2014). For each year in the dataset, we used the 4 km
118 resolution PRISM data (PRISM Climate Group 2004) to calculate eight bioclimatic
119 variables identified as relevant to bird distributions (Harris 2015): mean diurnal range,
120 isothermality, max temperature of the warmest month, mean temperature of the wettest
121 quarter, mean temperature of the driest quarter, precipitation seasonality, precipitation
122 of the wettest quarter, and precipitation of the warmest quarter. Satellite-derived NDVI,
123 a primary correlate of richness in BBS data (Hurlbert and Haskell 2002), was obtained
124 from the NDIV3g dataset with an 8 km resolution (Pinzon and Tucker 2014) and was
125 available from 1981-2013. Average summer (May, June, July) and winter (December,
126 January, February) NDVI values were used as predictors. Elevation was from the SRTM
127 90m elevation dataset (Jarvis et al. 2008) obtained using the R package raster (Hijmans
128 2016). Because BBS routes are 40-km transects rather than point counts, we used the
129 average value of each environmental variable within a 40 km radius of each BBS route's
130 starting point.

131 **Future environmental projections.** We made long term forecasts from 2014-2050
132 using the CMIP5 multi-model ensemble dataset as the source for climate variables
133 (Brekke et al. 2013). Precipitation and temperature from 37 downscaled model runs
134 (Brekke et al. 2013, see Table S1) using the RCP6.0 scenario were averaged together to
135 create a single ensemble used to calculate the bioclimatic variables for North America.
136 For NDVI we used the per-site average values from 2000-2013 as a simple forecast. For
137 observer effects (see below) each site was set to have zero observer bias.

138 **Accounting for observer effects**

139 Observer effects are inherent in large data sets collected by different observers, and are
140 known to occur in BBS (Sauer et al. 1994). For each forecasting approach, we trained
141 two versions of the corresponding model: one with corrections for differences among
142 observers, and one without (Figure 2). We estimated the observer effects (and
143 associated uncertainty about those effects) with a linear mixed model, with observer as
144 a random effect, built in the Stan probabilistic programming language (Carpenter et al.
145 2017). Because observer and site are strongly related (observers tend to repeatedly
146 sample the same site), site was also included as a random effect to ensure that inferred
147 deviations were actually observer-related (as opposed to being related to the sites that a
148 given observer happened to see). The resulting model partitions the variance in
149 observed richness values into site-level variance, observer-level variance, and residual
150 variance (e.g. variation within a site from year to year). The site-level estimates can also
151 be used directly as the “average” baseline model (see below). The estimated observer
152 effects can be subtracted from the richness values for a particular observer to provide an
153 estimate of how many species would have been found by a “typical” observer. To
154 incorporate uncertainty in these “corrected” richness values into the forecasting models
155 we collected 500 Monte Carlo samples from the model’s posterior distribution, and fit
156 each of the downstream models with each of the Monte Carlo samples. Each Monte
157 Carlo sample represented a different possible set of observer-level and site-level random
158 effect values across the full 32-year dataset.

159 **Models: site-level models**

160 Three of the models used in this study were fit to each site separately, with no
161 environmental information (Table 1). These models were fit to each BBS route twice:
162 once using the residuals from the observer model, and once using the raw richness
163 values. When correcting for observer effects, we averaged across 500 models that were

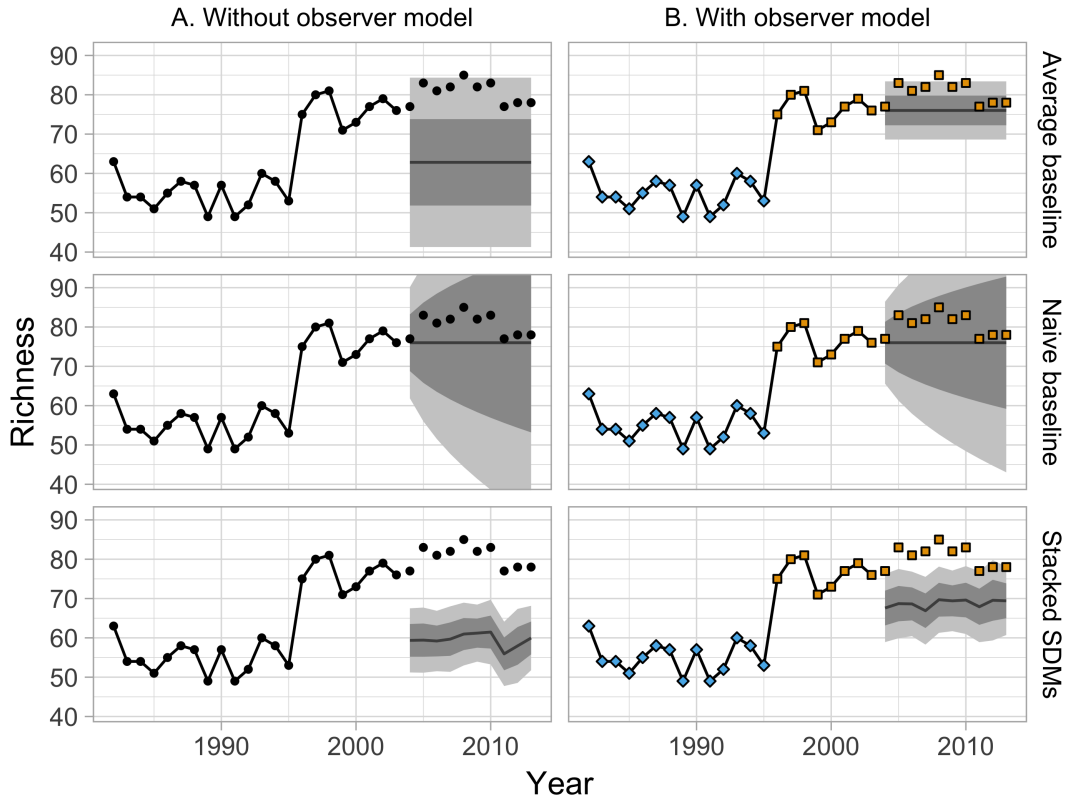


Figure 2: **A.** Model predictions for Pennsylvania route 35 when all observers are treated the same (black points). **B.** Model predictions for the same route when accounting for systematic differences between observers (represented by the points' colors). In this example most models are made more robust to observer turnover. Note that the “naive” model is less sensitive to observer turnover, and does not benefit as much from modeling it.

Table 1: Six forecasting models. Single-site models were trained site-by-site, without environmental data. Environmental models were trained using all sites together, without information regarding which transects occurred at which site or during which year. Most of the models were trained to predict richness directly. This mirrors the standard application of these techniques. Separate random forest SDMs were fit for each species and used to predict the probability of that species occurring at each site. The species-level probabilities at a site were summed to predict richness. The mistnet JSDM was trained to predict the full species composition at each site, and the number of species in its predictions was used as an estimate of richness.

Model	Response variable	Predictors		
		Site id	Time	Environment
Single-site models				
Average baseline	richness	✓		
Naive baseline	richness	✓	✓	
Auto-ARIMA	richness	✓	✓	
Environmental models				
GBM richness	richness			✓
Stacked SDMs	species-level presence			✓
Mistnet JSDM	species composition			✓

164 fit separately to the 500 Monte Carlo estimates of the observer effects, to account for
165 our uncertainty in the true values of those effects. All of these models use a Gaussian
166 error distribution (rather than a count distribution) for reasons discussed below (see
167 “Model evaluation”).

168 **Baseline models.** We used two simple baseline models as a basis for comparison with
169 the more complex models (Figure 2A). These baselines treated site-level richness
170 observations either as uncorrelated noise around a site-level constant (the “average”
171 model) or as an autoregressive model with a single year of history (the “naive” model,
172 Hyndman and Athanasopoulos 2014). Predictions from the “average” model are
173 centered on the average richness observed during training, and the confidence intervals
174 are narrow and constant-width. The “naive” model, in contrast, predicts that future
175 observations will be similar to the final observed value (e.g., in our hindcasts the value
176 observed in 2003), and the confidence intervals expand rapidly as the predictions extend

177 farther into the future. Both models' richness predictions are centered on a constant
178 value, so neither model can anticipate any trends in richness or any responses to future
179 environmental changes.

180 **Time series models.** We used Auto-ARIMA models (based on the `auto.arima`
181 function in the package `forecast`; Hyndman 2017) to represent an array of different
182 time-series modeling approaches. These models can include an autoregressive
183 component (as in the “naive” model, but with the possibility of longer-term
184 dependencies in the underlying process), a moving average component (where the noise
185 can have serial autocorrelation) and an integration/differencing component (so that the
186 analysis could be performed on sequential differences of the raw data, accommodating
187 more complex patterns including trends). The `auto.arima` function chooses whether
188 to include each of these components (and how many terms to include for each one)
189 using AICc (Hyndman 2017). Since there is no seasonal component to the BBS
190 time-series, we did not include a season component in these models. Otherwise we used
191 the default settings for this function (Hyndman 2017).

192 **Models: environmental models**

193 In contrast to the single-site models, most attempts to predict species richness focus on
194 using correlative models based on environmental variables. We tested three common
195 variants of this approach: direct modeling of species richness; stacking individual
196 species distribution models; and joint species distribution models (JSDMs). Following
197 the standard approach, site-level random effects were not included in these models as
198 predictors, meaning that this approach implicitly assumes that two sites with identical
199 Bioclim, elevation, and NDVI values should have identical richness distributions. As
200 above, we included observer effects and the associated uncertainty by running these
201 models 500 times (once per MCMC sample).

202 **“Macroecological” model: richness GBM.** We used a boosted regression tree model

203 using the `gbm` package (Ridgeway *et al.* 2017) to directly model species richness as a
204 function of environmental variables. Boosted regression trees are a form of tree-based
205 modeling that work by fitting thousands of small tree-structured models sequentially,
206 with each tree optimized to reduce the error of its predecessors. They are flexible
207 models that are considered well suited for prediction (Elith *et al.* 2008). This model was
208 optimized using a Gaussian likelihood, with a maximum interaction depth of 5,
209 shrinkage of 0.015, and up to 10,000 trees. The number of trees used for prediction was
210 selected using the “out of bag” estimator; this number averaged 6,700 for the
211 non-observer data and 7,800 for the observer-corrected data.

212 **Species Distribution Model: stacked random forests.** Species distribution models
213 (SDMs) predict individual species’ occurrence probabilities using environmental
214 variables. Species-level models are used to predict richness by summing the predicted
215 probability of occupancy across all species at a site. This avoids known problems with
216 the use of thresholds for determining whether or not a species will be present at a site
217 (Pellissier *et al.* 2013, Calabrese *et al.* 2014). Following Calabrese *et al.* (2014), we
218 calculated the uncertainty in our richness estimate by treating richness as a sum over
219 independent Bernoulli random variables: $\sigma_{richness}^2 = \sum_i p_i(1 - p_i)$, where i indexes
220 species. By itself, this approach is known to underestimate the true community-level
221 uncertainty because it ignores the uncertainty in the species-level probabilities
222 (Calabrese *et al.* 2014). To mitigate this problem, we used an ensemble of 500 estimates
223 for each of the species-level probabilities instead of just one, propagating the
224 uncertainty forward. We obtained these estimates using random forests (Liaw and
225 Wiener 2002), a common approach in the species distribution modeling literature.
226 Random forests are constructed by fitting hundreds of independent regression trees to
227 randomly-perturbed versions of the data (Cutler *et al.* 2007, Caruana *et al.* 2008). When
228 correcting for observer effects, each of the 500 trees in our species-level random forests
229 used a different Monte Carlo estimate of the observer effects as a predictor variable.

230 **Joint Species Distribution Model: mistnet.** Joint species distribution models
 231 (JSDMs) are a new approach that makes predictions about the full composition of a
 232 community instead of modeling each species independently as above (Warton et al.
 233 2015). JSDMs remove the assumed independence among species and explicitly account
 234 for the possibility that a site will be much more (or less) suitable for birds in general (or
 235 particular groups of birds) than one would expect based on the available environmental
 236 measurements alone. As a result, JSDMs do a better job of representing uncertainty
 237 about richness than stacked SDMs (Harris 2015, Warton et al. 2015). We used the
 238 `mistnet` package (Harris 2015) because it is the only JSDM that describes species'
 239 environmental associations with nonlinear functions.

240 **Model evaluation**

241 We defined model performance for all models in terms of continuous Gaussian errors,
 242 instead of using discrete count distributions. Variance in species richness within sites
 243 was lower than predicted by several common count models, such as the Poisson or
 244 binomial (i.e. richness was underdispersed for individual sites), so these count models
 245 would have had difficulty fitting the data (cf. Calabrese et al. 2014). The use of a
 246 continuous distribution is adequate here, since richness had a relatively large mean (51)
 247 and all models produce continuous richness estimates. When a model was run multiple
 248 times for the purpose of correcting for observer effects, we used the mean of those runs'
 249 point estimates as our final point estimate and we calculated the uncertainty using the
 250 law of total variance (i.e. $\text{Var}(\bar{y}) + \mathbb{E} [\text{Var}(y)]$, or the variance in point estimates plus
 251 the average residual variance).

252 We evaluated each model's forecasts using the data for each year between 2004 and
 253 2013. We used three metrics for evaluating performance: 1) root-mean-square error
 254 (RMSE) to determine how far, on average, the models' predictions were from the
 255 observed value; 2) the 95% prediction interval coverage to determine how well the

models predicted the range of possible outcomes; and 3) deviance (i.e. negative 2 times the Gaussian log-likelihood) as an integrative measure of fit incorporating good point estimates, precision, and coverage. In addition to evaluating forecast performance in general, we evaluated how performance changed as the time horizon of forecasting increased by plotting performance metrics against year. Finally, we decomposed each model's squared error into two components: the squared error associated with site-level means and the squared error associated with annual fluctuations in richness within a site. This decomposition describes the extent to which each model's error depends on consistent differences among sites versus changes in site-level richness from year to year.

All analyses were conducted using R (R Core Team 2017). Primary R packages used in the analysis included dplyr (Wickham et al. 2017), tidyr (Wickham 2017), gimms (Detsch 2016), sp (Pebesma and Bivand 2005, Bivand et al. 2013), raster (Hijmans 2016), prism (PRISM Climate Group 2004), rdataretriever (McGlinn et al. 2017), forecast (Hyndman and Khandakar 2008, Hyndman 2017), git2r (Widgren and others 2016), ggplot (Wickham 2009), mistnet (Harris 2015), viridis (Garnier 2017), rstan (Stan Development Team 2016), yaml (Stephens 2016), purrr (Henry and Wickham 2017), gbm (Ridgeway *et al.* 2017), randomForest (Liaw and Wiener 2002). Code to fully reproduce this analysis is available on GitHub (<https://github.com/weecology/bbs-forecasting>) and archived on Zenodo (Harris et al. 2017).

Results

The site-observer mixed model found that 70% of the variance in richness in the training set could be explained by differences among sites, and 21% could be explained by differences among observers. The remaining 9% represents residual variation, where a given observer might report a different number of species in different years. In the

282 training set, the residuals had a standard deviation of about 3.6 species. After correcting
283 for observer differences, there was little temporal autocorrelation in these residuals
284 (i.e. the residuals in one year explain 1.3% of the variance in the residuals of the
285 following year), suggesting that richness was approximately stationary between 1982
286 and 2003.

287 When comparing forecasts for richness across sites all methods performed well (Figure
288 3; all $R^2 > 0.5$). However SDMs (both stacked and joint) and the macroecological
289 model all failed to successfully forecast the highest-richness sites, resulting in a notable
290 clustering of predicted values near ~60 species and the poorest model performance
291 ($R^2=0.52-0.78$, versus $R^2=0.67-0.87$ for the within-site methods).

292 While all models generally performed well in absolute terms (Figure 3), none
293 consistently outperformed the “average” baseline (Figure 4). The auto-ARIMA was
294 generally the best-performing non-baseline model, but in many cases (67% of the time),
295 the auto.arima procedure selected a model with only an intercept term (i.e. no
296 autoregressive terms, no drift, and no moving average terms), making it similar to the
297 “average” model. All five alternatives to the “average” model achieved lower error on
298 some of the sites in some years, but each one had a higher mean absolute error and
299 higher mean deviance (Figure 4).

300 Most models produced confidence intervals that were too narrow, indicating
301 overconfident predictions (Figure 5C). The random forest-based SDM stack was the
302 most overconfident model, with only 72% of observations falling inside its 95%
303 confidence intervals. This stacked SDM’s narrow predictive distribution caused it to
304 have notably higher deviance (Figure 5B) than the next-worst model, even though its
305 point estimates were not unusually bad in terms of RMSE (5A). As discussed elsewhere
306 (Harris 2015), this overconfidence is a product of the assumption in stacked SDMs that
307 errors in the species-level predictions are independent. The GBM-based
308 “macroecological” model and the mistnet JSMD had the best calibrated uncertainty

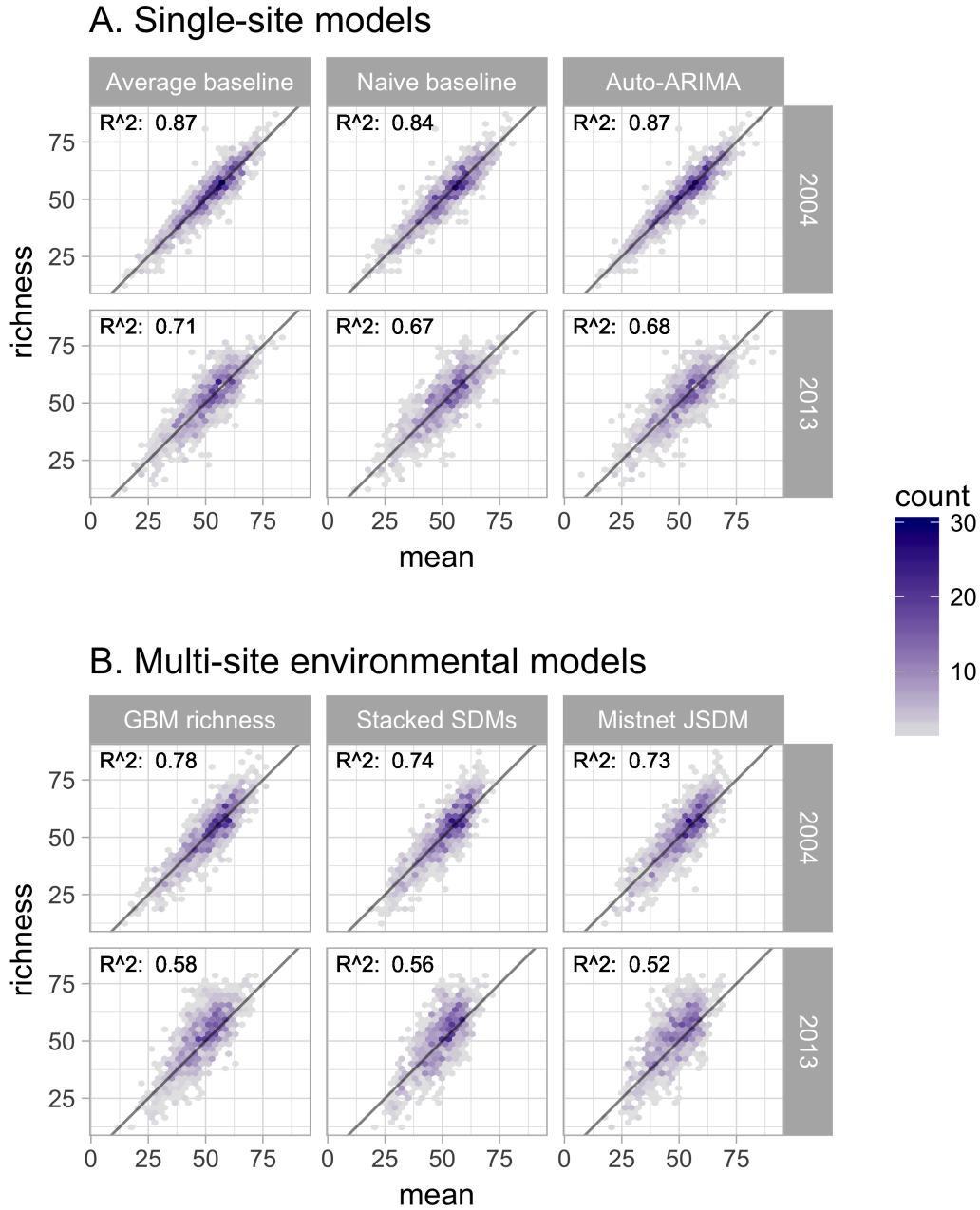


Figure 3: Performance of six forecasting models for predicting species richness one year (2004) and ten years into the future (2013). Plots show observed vs. predicted values for species richness. Models were trained with data from 1982-2003. In general, the single-site models (**A**) outperformed the environmental models (**B**). The accuracy of the predictions generally declined as the timescale of the forecast was extended from 2004 to 2013.

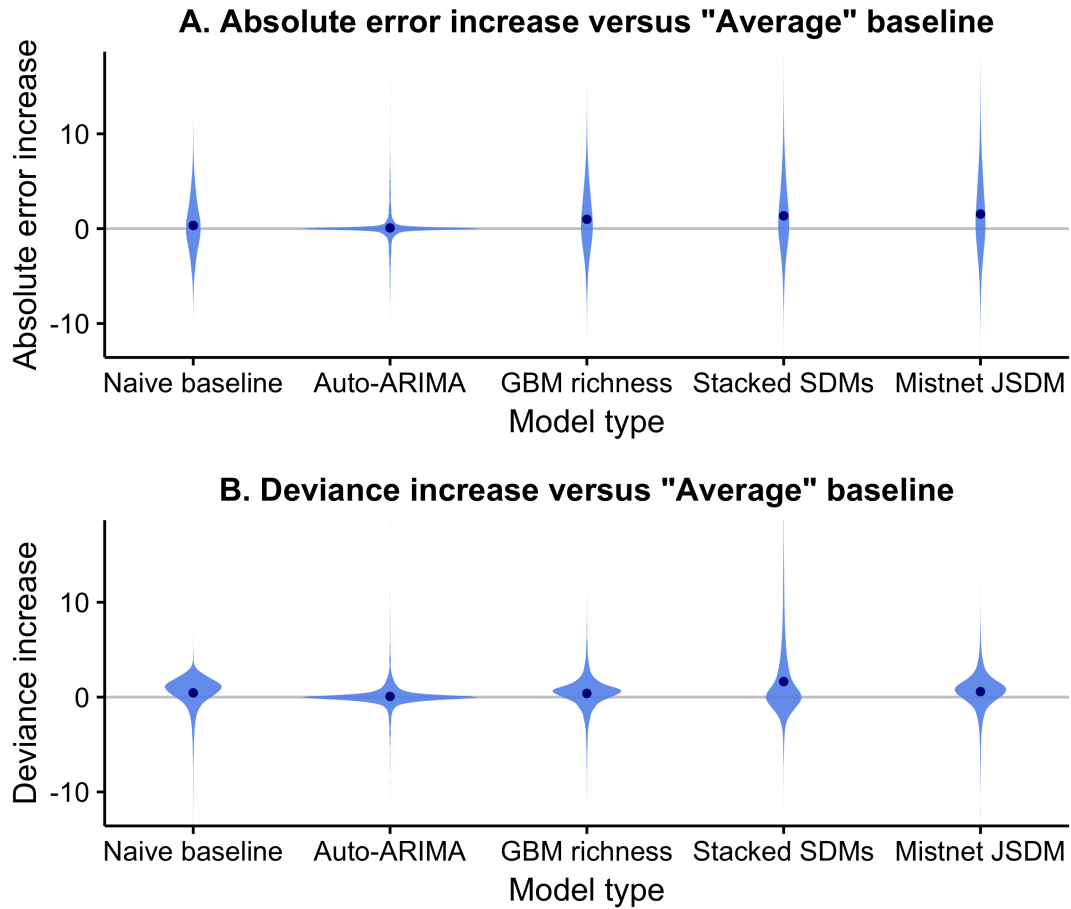


Figure 4: Difference between the forecast error of models and the error of the average baseline using both absolute error (A.) and deviance (B.). Differences are taken for each site and testing year so that errors for the same forecast are directly compared. The error of the average baseline is by definition zero and is indicated by the horizontal gray line. None of the five models provided a consistent improvement over the average baseline. The absolute error of the models was generally similar or larger than that of the “average” model, with large outliers in both directions. The deviance of the models was also generally higher than the “average” baseline.

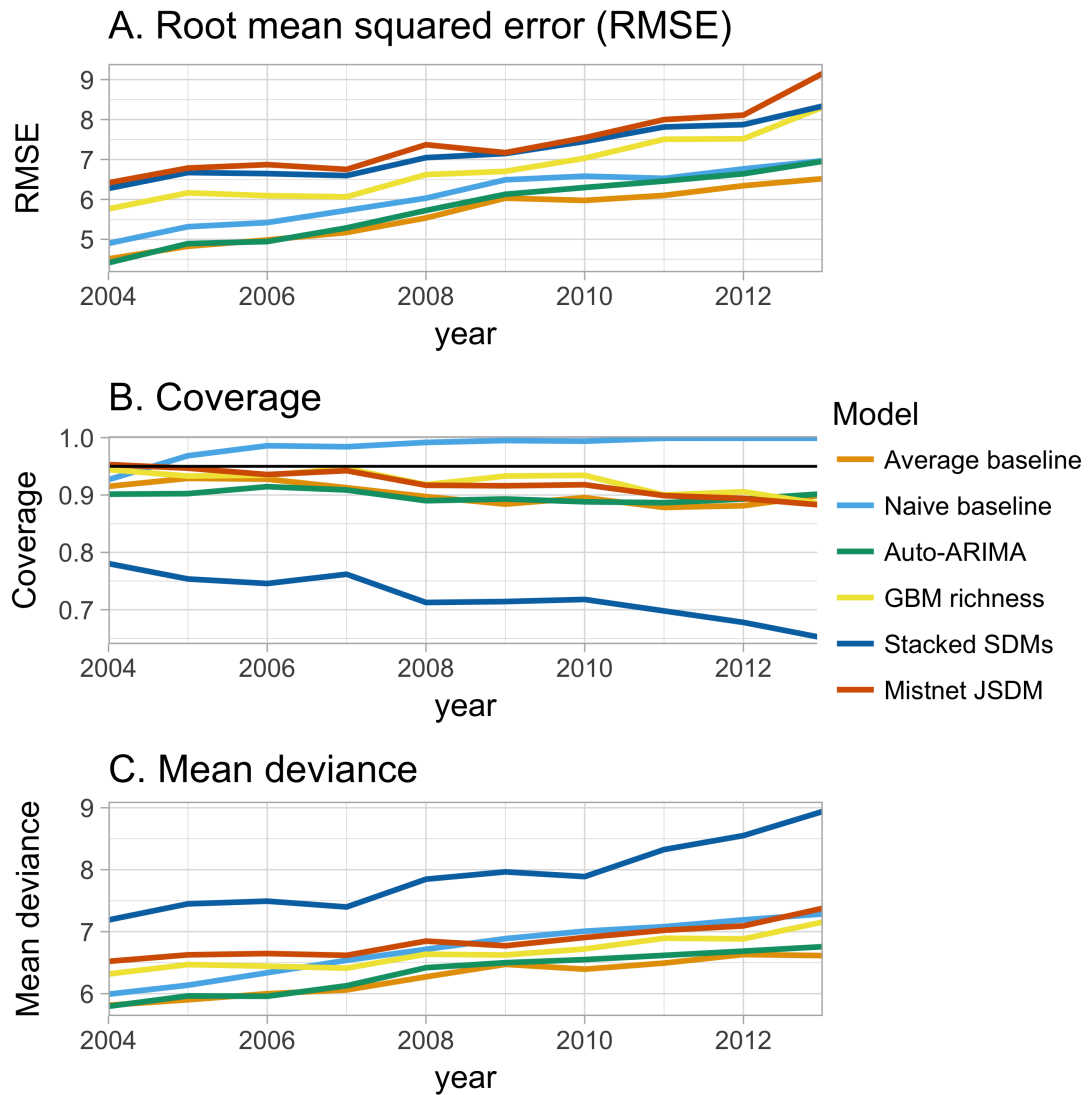


Figure 5: Change in performance of the six forecasting models with the time scale of the forecast (1-10 years into the future). **A.** Root mean square error (rmse; the error in the point estimates) shows the three environmental models tending to show the largest errors at all time scales and the models getting worse as they forecast further into the future at approximately the same rate. **B.** Deviance (lack of fit of the entire predictive distribution) shows the stacked species distribution models with much higher error than other models and shows that the “naive” model’s deviance grows relatively quickly. **C.** Coverage of a model’s 95% confidence intervals (how often the observed values fall inside the predicted range; the black line indicates ideal performance) shows that the “naive” model’s predictive distribution is too wide (capturing almost all of the data) and the stacked SDM’s predictive distribution is too narrow (missing almost a third of the observed richness values by 2014).

estimates (Figure 5B) and therefore their relative performance was higher in terms of deviance than in terms of RMSE. The “naive” model was the only model whose confidence intervals were too wide (Figure 5C), which can be attributed to the rapid rate at which these intervals expand (Figure 1).

Partitioning each model’s squared error shows that the majority of the residual error was attributed to errors in estimating site-level means, rather than errors in tracking year-to-year fluctuations (Figure 6). The “average” model, which was based entirely on site-level means, had the lowest error in this regard. In contrast, the three environmental models showed larger biases at the site level, though they still explained most of the variance in this component. This makes sense, given that they could not explicitly distinguish among sites with similar climate, NDVI, and elevation. Interestingly, the environmental models had higher squared error than the baselines did for tracking year-to-year fluctuations in richness as well.

Accounting for differences among observers generally improved measures of model fit (Figure 7). Improvements primarily resulted from a small number of forecasts where observer turnover caused a large shift in the reported richness values. The naive baseline was less sensitive to these shifts, because it largely ignored the richness values reported by observers that had retired by the end of the training period (Figure 1). The average model, which gave equal weight to observations from the whole training period, showed a larger decline in performance when not accounting for observer effects – especially in terms of coverage. The performance of the mistnet JSDM was notable here, because its prediction intervals retained good coverage even when not correcting for observer differences, which we attribute to the JSDM’s ability to model this variation with its latent variables.

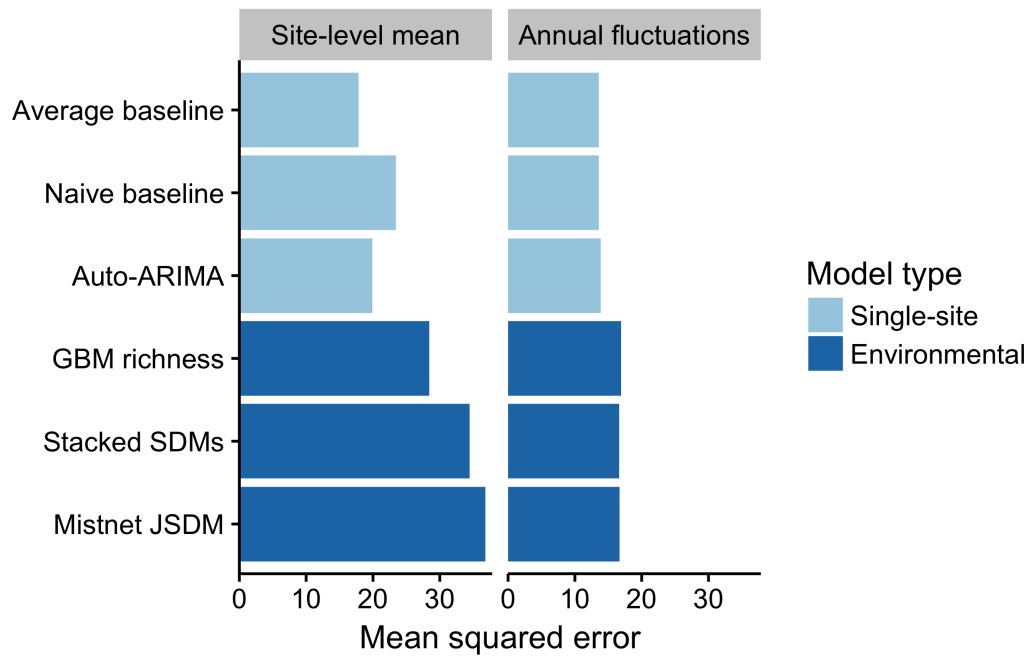


Figure 6: Partitioning of the squared error for each model into site and year components. The site-level mean component shows consistent over or under estimates of richness at a site across years. The annual fluctuation component shows errors in predicting fluctuations in a site's richness over time. Both components of the mean squared error were lower for the single-site models than for the environmental models.

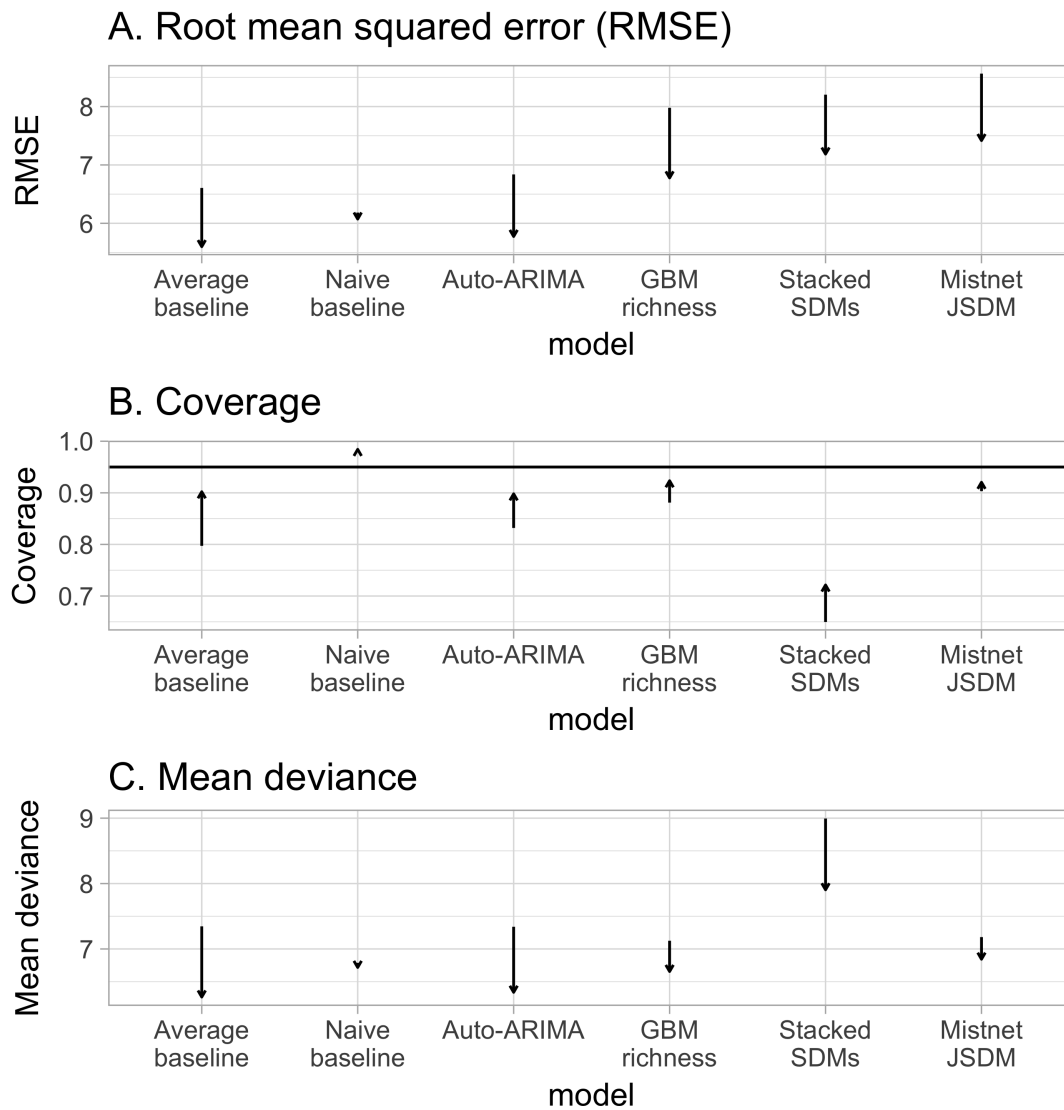


Figure 7: Controlling for differences among observers generally improved each model's predictions, on average. The magnitude of this effect was negligible for the Naive baseline, however.

333 Discussion

334 Forecasting is an emerging imperative in ecology; as such, the field needs to develop
335 and follow best practices for conducting and evaluating ecological forecasts (Clark et al.
336 2001). We have used a number of these practices (Box 1) in a single study that builds
337 and evaluates forecasts of biodiversity in the form of species richness. The results of
338 this effort are both promising and humbling. When comparing forecasts across sites,
339 many different approaches to forecasting produce reasonable forecasts (Figure 3). If a
340 site is predicted to have a high number of species in the future, relative to other sites, it
341 generally does. However, none of the methods evaluated reliably determined how
342 site-level richness changes over time (Figure 6), which is generally the stated purpose of
343 these forecasts. As a result, baseline models, which did not attempt to anticipate
344 changes in richness over time, generally provided the best forecasts for future
345 biodiversity. While this study is restricted to breeding birds in North America, its results
346 are consistent with a growing literature on the limits of ecological forecasting, as
347 discussed below.

348 The most commonly used methods for forecasting future biodiversity, SDMs and
349 macroecological models, both produced worse forecasts than time-series models and
350 simple baselines. This weakness suggests that predictions about future biodiversity
351 change should be viewed with skepticism unless the underlying models have been
352 validated temporally, via hindcasting and comparison with simple baselines. Since
353 site-level richness is relatively stable, spatial validation is not enough: a model can have
354 high accuracy across spatial gradients without being able to predict changes over time.
355 This gap between spatial and temporal accuracy is known to be important for
356 species-level predictions (Rapacciuolo et al. 2012, Oedekoven et al. 2017); our results
357 indicate that it is substantial for higher-level patterns like richness as well. SDMs' poor
358 temporal predictions are particularly sobering, as these models have been one of the
359 main foundations for estimates of the predicted loss of biodiversity to climate change

over the past decade or so (Thomas et al. 2004, Thuiller et al. 2011, Urban 2015). Our results also highlight the importance of comparing multiple modeling approaches when conducting ecological forecasts, and in particular, the value of comparing results to simple baselines to avoid over-interpreting the information present in these forecasts [Box 1]. Disciplines that have more mature forecasting cultures often do this by reporting “forecast skill”, i.e., the improvement in the forecast relative to a simple baseline (Jolliffe and Stephenson 2003). We recommend following the example of Ye et al. (2015) and adopting this approach in future ecological forecasting research.

When comparing different methods for forecasting our results demonstrate the importance of considering uncertainty (Box 1; Clark et al. 2001, Dietze et al. 2016). Previous comparisons between stacked SDMs and macroecological models reported that the methods yielded equivalent results for forecasting diversity (Algar et al. 2009, Distler et al. 2015). While our results support this equivalence for point estimates, they also show that stacked SDMs dramatically underestimate the range of possible outcomes; after ten years, more than a third of the observed richness values fell outside the stacked SDMs’ 95% prediction intervals. Consistent with Harris (2015) and Warton et al. (2015), we found that JSMDs’ wider prediction intervals enabled them to avoid this problem. Macroecological models appear to share this advantage, while being considerably easier to implement.

We have only evaluated annual forecasts up to a decade into the future, but forecasts are often made with a lead time of 50 years or more. These long-term forecasts are difficult to evaluate given the small number of century-scale datasets, but are important for understanding changes in biodiversity at some of the lead times relevant for conservation and management. Two studies have assessed models of species richness at longer lead times (Algar et al. 2009, Distler et al. 2015), but the results were not compared to baseline or time-series models (in part due to data limitations) making them difficult to compare to our results directly. Studies on shorter time scales, such as

387 ours, provide one way to evaluate our forecasting methods without having to wait
388 several decades to observe the effects of environmental change on biodiversity (Petchey
389 et al. 2015, Dietze et al. 2016, Tredennick et al. 2016), but cannot fully replace
390 longer-term evaluations (Tredennick et al. 2016). In general, drivers of species richness
391 can differ at different temporal scales (Rosenzweig 1995, White 2004, 2007, Blonder et
392 al. 2017), so different methods may perform better for different lead times. In particular,
393 we might expect environmental and ecological information to become more important
394 at longer time scales, and thus for the performance of simple baseline forecasts to
395 degrade faster than forecasts from SDMs and other similar models. We did observe a
396 small trend in this direction: deviance for the auto-ARIMA models and for the average
397 baseline grew faster than for two of the environmental models (the JSMD and the
398 macroecological model), although this growth was not statistically significant for the
399 average baseline.

400 While it is possible that models that include species' relationships to their environments
401 or direct environmental constraints on richness will provide better fits at longer lead
402 times, it is also possible that they will continue to produce forecasts that are worse than
403 baselines that assume the systems are static. This would be expected to occur if richness
404 in these systems is not changing over the relevant multi-decadal time scales, which
405 would make simpler models with no directional change more appropriate. Recent
406 suggestions that local scale richness in some systems is not changing directionally at
407 multi-decadal scales supports this possibility (Brown et al. 2001, Ernest and Brown
408 2001, Vellend et al. 2013, Dornelas et al. 2014). A lack of change in richness may be
409 expected even in the presence of substantial changes in environmental conditions and
410 species composition at a site due to replacement of species from the regional pool
411 (Brown et al. 2001, Ernest and Brown 2001). On average, the Breeding Bird Survey
412 sites used in this study show little change in richness (site-level SD of 3.6 species, after
413 controlling for differences among observers; see also La Sorte and Boecklen 2005). The

414 absence of rapid change in this dataset is beneficial for the absolute accuracy of
415 forecasts across different sites: when a past year's richness is already known, it is easy
416 to estimate future richness. Ward et al. (2014) found similar patterns in time series of
417 fisheries stocks, where relatively stable time series were best predicted by simple
418 models and more complex models were only beneficial with dynamic time series. The
419 site-level stability of the BBS data also explains why SDMs and macroecological
420 models perform relatively well at predicting future richness, despite failing to capture
421 changes in richness over time. However, this stability also makes it difficult to improve
422 forecasts relative to simple baselines, since those baselines are already close to
423 representing what is actually occurring in the system. These results suggest that
424 single-site models should be actively considered for forecasts of richness and other
425 stable aspects of biodiversity. Our results also suggest that future efforts to understand
426 and forecast biodiversity should incorporate species composition, since lower-level
427 processes are expected to be more dynamic (Ernest and Brown 2001, Dornelas et al.
428 2014) and contain more useful information (Harris 2015).

429 Future biodiversity forecasting efforts also need to address the uncertainty introduced
430 by the error in forecasting the environmental conditions that are used as predictor
431 variables. In this, and other hindcasting studies, the environmental conditions for the
432 “future” are known because the data has already been observed. However, in real
433 forecasts the environmental conditions themselves have to be predicted, and
434 environmental forecasts will also have uncertainty and bias. Ultimately, ecological
435 forecasts that use environmental data will therefore be more uncertain than our current
436 hindcasting efforts, and it is important to correctly incorporate this uncertainty into our
437 models (Clark et al. 2001, Dietze 2017). Limitations in forecasting future
438 environmental conditions—particularly at small scales—will present continued
439 challenges for models incorporating environmental variables, and this may result in a
440 continued advantage for simple single-site approaches.

441 In addition to comparing and improving the process models used for forecasting it is
442 important to consider the observation models. When working with any ecological
443 dataset, there are imperfections in the sampling process that have the potential to
444 influence results. With large scale surveys and citizen science datasets, such as the
445 Breeding Bird Survey, these issues are potentially magnified by the large number of
446 different observers and by major differences in the habitats and species being surveyed
447 (Sauer et al. 1994). Accounting for differences in observers reduced the average error in
448 our point estimates and also improved the coverage of the confidence intervals. In
449 addition, controlling for observer effects resulted in changes in which models performed
450 best, most notably improving most models' point estimates relative to the naive baseline.
451 This demonstrates that modeling observation error can be important for properly
452 estimating and reducing uncertainty in forecasts and can also lead to changes in the best
453 methods for forecasting [Box 1]. This suggests that, prior to accounting for observer
454 effects, the naive model performed well largely because it was capable of
455 accommodating rapid shifts in estimated richness introduced by changes in the observer.
456 These kinds of rapid changes were difficult for the other single-site models to
457 accommodate. Another key aspect of an ideal observation model is imperfect detection.
458 In this study, we did not address differences in detection probability across species and
459 sites (Boulinier et al. 1998) since there is no clear way to address this issue using North
460 American Breeding Bird Survey data without making strong assumptions about the data
461 (i.e., assuming there is no biological variation in stops along a route; White and Hurlbert
462 2010), but this would be a valuable addition to future forecasting models.

463 The science of forecasting biodiversity remains in its infancy and it is important to
464 consider weaknesses in current forecasting methods in that context. In the beginning,
465 weather forecasts were also worse than simple baselines, but these forecasts have
466 continually improved throughout the history of the field (McGill 2012, Silver 2012,
467 Bauer et al. 2015). One practice that lead to improvements in weather forecasts was that

large numbers of forecasts were made publicly, allowing different approaches to be regularly assessed and refined (McGill 2012, Silver 2012). To facilitate this kind of improvement, it is important for ecologists to start regularly making and evaluating real ecological forecasts, even if they perform poorly, and to make these forecasts openly available for assessment (McGill 2012, Dietze et al. 2016). These forecasts should include both short-term predictions, which can be assessed quickly, and mid- to long-term forecasts, which can help ecologists to assess long time-scale processes and determine how far into the future we can successfully forecast (Dietze et al. 2016, Tredennick et al. 2016). We have openly archived forecasts from all six models through the year 2050 (White and Harris 2017), so that we and others can assess how well they perform. We plan to evaluate these forecasts and report the results as each new year of BBS data becomes available, and make iterative improvements to the forecasting models in response to these assessments.

Making successful ecological forecasts will be challenging. Ecological systems are complex, our fundamental theory is less refined than for simpler physical and chemical systems, and we currently lack the scale of data that often produces effective forecasts through machine learning. Despite this, we believe that progress can be made if we develop an active forecasting culture in ecology that builds and assesses forecasts in ways that will allow us to improve the effectiveness of ecological forecasts more rapidly (Box 1; McGill 2012, Dietze et al. 2016). This includes expanding the scope of the ecological and environmental data we work with, paying attention to uncertainty in both model building and forecast evaluation, and rigorously assessing forecasts using a combination of hindcasting, archived forecasts, and comparisons to simple baselines.

Acknowledgments

This research was supported by the Gordon and Betty Moore Foundation's Data-Driven Discovery Initiative through Grant GBMF4563 to E.P. White. We thank the developers

494 and providers of the data and software that made this research possible including: the
495 PRISM Climate Group at Oregon State University, the staff at USGS and volunteer
496 citizen scientists associated with the North American Breeding Bird Survey, NASA, the
497 World Climate Research Programme's Working Group on Coupled Modelling and its
498 working groups, the U.S. Department of Energy's Program for Climate Model
499 Diagnosis and Intercomparison, and the Global Organization for Earth System Science
500 Portals. A. C. Perry provided valuable comments that improved the clarity of this
501 manuscript.

502 **Box 1: Best practices for making and evaluating ecological forecasts**

503 **1. Compare multiple modeling approaches**

504 Typically ecological forecasts use one modeling approach or a small number of related
505 approaches. By fitting and evaluating multiple modeling approaches we can learn more
506 rapidly about the best approaches for making predictions for a given ecological quantity
507 (Clark et al. 2001, Ward et al. 2014). This includes comparing process-based (e.g.,
508 Kearney and Porter 2009) and data-driven models (e.g., Ward et al. 2014), as well as
509 comparing the accuracy of forecasts to simple baselines to determine if the modeled
510 forecasts are more accurate than the naive assumption that the world is static (Jolliffe
511 and Stephenson 2003, Ye et al. 2015).

512 **2. Use time-series data when possible**

513 Forecasts describe how systems are expected to change through time. While some areas
514 of ecological forecasting focus primarily on time-series data (Ward et al. 2014), others
515 primarily focus on using spatial models and space-for-time substitutions (Blois et al.
516 2013). Using ecological and environmental time-series data allows the consideration of

517 actual dynamics from both a process and error structure perspective (Tredennick et al.
518 2016).

519 **3. Pay attention to uncertainty**

520 Understanding uncertainty in a forecast is just as important as understanding the
521 average or expected outcome. Failing to account for uncertainty can result in
522 overconfidence in uncertain outcomes leading to poor decision making and erosion of
523 confidence in ecological forecasts (Clark et al. 2001). Models should explicitly include
524 sources of uncertainty and propagate them through the forecast where possible (Clark et
525 al. 2001, Dietze 2017). Evaluations of forecasts should assess the accuracy of models'
526 estimated uncertainties as well as their point estimates (Dietze 2017).

527 **4. Use predictors related to the question**

528 Many ecological forecasts use data that is readily available and easy to work with.
529 While ease of use is a reasonable consideration it is also important to include predictor
530 variables that are expected to relate to the ecological quantity being forecast.
531 Time-series of predictors, instead of long-term averages, are also preferable to match
532 the ecological data (see #2). Investing time in identifying and acquiring better predictor
533 variables may have at least as many benefits as using more sophisticated modeling
534 techniques (Kent et al. 2014).

535 **5. Address unknown or unmeasured predictors**

536 Ecological systems are complex and many biotic and abiotic aspects of the environment
537 are not regularly measured. As a result, some sites may deviate in consistent ways from
538 model predictions. Unknown or unmeasured predictors can be incorporated in models

539 using site-level random effects (potentially spatially autocorrelated) or by using latent
540 variables that can identify unmeasured gradients (Harris 2015).

541 **6. Assess how forecast accuracy changes with time-lag**

542 In general, the accuracy of forecasts decreases with the length of time into the future
543 being forecast (Petchey et al. 2015). This decay in accuracy should be considered when
544 evaluating forecasts. In addition to simple decreases in forecast accuracy the potential
545 for different rates of decay to result in different relative model performance at different
546 lead times should be considered.

547 **7. Include an observation model**

548 Ecological observations are influenced by both the underlying biological processes
549 (e.g. resource limitation) and how the system is sampled. When possible, forecasts
550 should model the factors influencing the observation of the data (Yu et al. 2010,
551 Hutchinson et al. 2011, Schurr et al. 2012).

552 **8. Validate using hindcasting**

553 Evaluating a model's predictive performance across time is critical for understanding if
554 it is useful for forecasting the future. Hindcasting uses a temporal out-of-sample
555 validation approach to mimic how well a model would have performed had it been run
556 in the past. For example, using occurrence data from the early 20th century to model
557 distributions which are validated with late 20th century occurrences. Dense time series,
558 such as yearly observations, are desirable to also evaluate the forecast horizon (see #6),
559 but this is not a strict requirement.

560 **9. Publicly archive forecasts**

561 Forecast values and/or models should be archived so that they can be assessed after new
562 data is generated (McGill 2012, Silver 2012, Dietze et al. 2016). Enough information
563 should be provided in the archive to allow unambiguous assessment of each forecast's
564 performance (Tetlock and Gardner 2016).

565 **10. Make both short-term and long-term predictions**

566 Even in cases where long-term predictions are the primary goal, short-term predictions
567 should also be made to accommodate the time-scales of planning and management
568 decisions and to allow the accuracy of the forecasts to be quickly evaluated (Dietze et al.
569 2016, Tredennick et al. 2016).

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