

Beyond mean similarity: reconsidering homogenization research

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Abstract

Studies of beta diversity often use the average Jaccard similarity among pairs of sites to quantify a landscape's biotic homogeneity. Over more than a decade, community ecologists have suggested a range of ecological hypotheses and frameworks to explain the variation in species' effects on this average similarity. However, many of these results can be explained simply by counting the proportion of sites occupied by each species and adding a small correction for non-independent assortment. Specifically, average Jaccard similarity can be decomposed exactly into an occupancy-based component that explains most of the variance and a covariance-based component that explains the rest. To the extent that homogenization research can be summarized by these two terms, it contains little information about ecological processes. Future ecological studies on beta diversity and biotic homogenization should therefore focus on 1) explaining changes in occupancy, 2) explaining changes in the covariance terms, and 3) explaining changes in local similarity at sub-landscape scales.

Introduction

As human influences spread some species globally and drive others extinct, community ecologists have become increasingly concerned about the loss of local biotic distinctiveness, or *biotic homogenization* [McKinney and Lockwood (1999); Rooney et al. (2007); refs]. In general, homogenization occurs when landscapes become dominated by cosmopolitan invaders and lose local endemics, although ecologists have observed a great deal of context dependence (McKinney 2004, Olden and Rooney 2006, Rooney et al. 2007). While several frameworks have been proposed for understanding these effects at the level of individual site pairs, they are cumbersome to use in practice. For example, Olden and Poff (2003)'s framework involves classifying biotic turnover events into 14 categories, while Rosenblad and Sax (2016)'s framework involves six. In either framework, a researcher may need to keep track of thousands of such events across thousands of pairs of sites. Moreover, these pairs of sites will not be statistically independent of one another, as each site will belong to many pairs. Worse, many types of turnover event are compatible with either homogenization or differentiation, depending on the prior state of the community. This complexity makes it difficult to test researchers' claims that their framework can explain the observed context dependence and make useful predictions about future biotic changes.

Homogenization is typically quantified as an increase in the average Jaccard similarity

among pairs of sites on a landscape (\bar{J} , defined more precisely in Equation 1 below). While many papers provide verbal arguments about the behavior of this metric, Harris et al. (2011) showed that it is not sensitive to the site-level patterns emphasized by most of these papers. As discussed below, nearly all of the variance in mean Jaccard similarity is determined by species' occupancy rates, and the remaining variance is determined by a simple covariance term; there is no room for any additional ecological information in the metric. This decomposition can provide researchers a more intuitive and less cumbersome way to study biotic homogenization at the landscape level than previous frameworks. As the behavior of landscape-level mean similarity has been fully characterized, homogenization research can now focus on other factors (discussed at the end of this paper).

Average Jaccard similarity

The Jaccard similarity between site i and j (\bar{J}) is defined as the proportion of species that are shared between them. More precisely, it is the number of species that are shared between the two sites (S_{ij}) divided by the number of species that occur in at least one (T_{ij}). The landscape-level mean Jaccard similarity value that homogenization research has largely focused on is therefore given by

$$\bar{J} = \frac{1}{\binom{n}{2}} \sum_{i \neq j} \left(\frac{S_{ij}}{T_{ij}} \right), \quad (1)$$

where n is the number of sites on the landscape (e.g. islands in an archipelago) and $\binom{n}{2}$ is the number of distinct site pairs, equal to $n(n-1)/2$. Most existing frameworks focus on tracking biotic turnover's effects on individual S_{ij} and T_{ij} values (e.g. Olden and Poff 2003, Rosenblad and Sax 2016). However, there is good reason to believe that this level of detail is not necessary for understanding \bar{J} (Harris et al. 2011). Using 47 large data sets from the USDA PLANTS database (one for each of the contiguous US states except Delaware; USDA NRCS 2010), I found that independently shuffling the list of sites occupied by every species did not appreciably affect \bar{J} (mean absolute deviation of 0.006 on a scale from 0 to 1; Appendix 1).

The central role of occupancy

If similarity and homogenization do not depend strongly on which sites are occupied by which species, then what does matter? In Harris et al. (2011), two colleagues and I showed

(empirically and with an appeal to the law of large numbers) that average similarity depends primarily on the proportion of sites occupied by each species. Specifically, we defined an approximation to mean Jaccard similarity, J^* , given by the average value of S_{ij} divided by the average value of T_{ij} . Substituting in formulas for these two averages, we derived

$$J^* = \sum_k \binom{p_k n}{2} / \sum_k \left[\binom{n}{2} - \binom{(1-p_k)n}{2} \right], \quad (2)$$

where p_k is the proportion of sites occupied by species k .

Despite the omission of any information at the level of individual sites or site pairs, we showed that the approximation explained 99.8% of the variance in \bar{J} across the 47 USDA PLANTS data sets. It also explained an average of 98.8% of the variance in species-level effects on average similarity. The paper introduced the *blender* package (Harris 2014) for R (R Core Team 2015), which allows users to easily perform these calculations. Similar results were presented around the same time by Chase et al. (2011) and by Vergara et al. (2011).

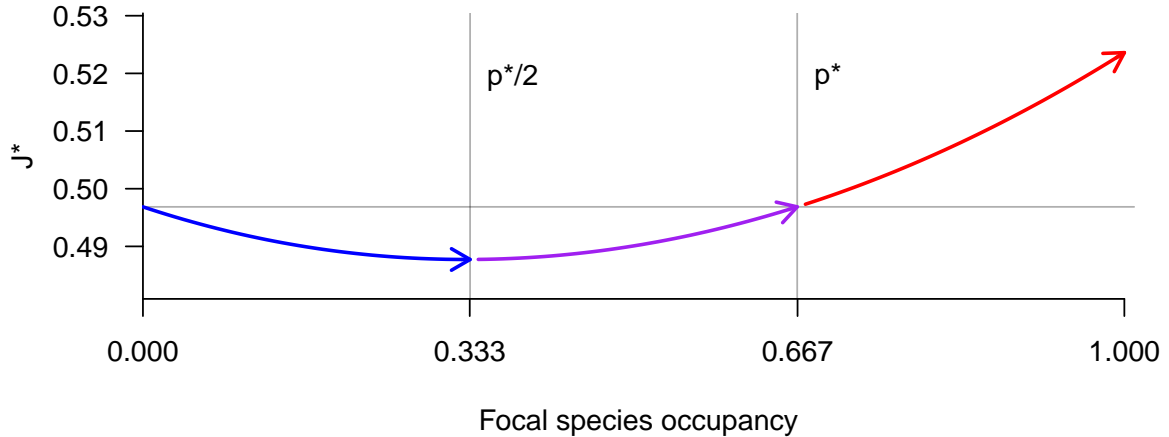


Figure 1: Effect of an exotic species on J^* , given a landscape with 120 sites and 20 native species whose occupancy rates produce $p^* = 2/3$. The exotic species’ effects on J^* occur in three phases as it spreads across the landscape. When an exotic species begins spreading, it will decrease mean similarity until its occupancy rate reaches $p^*/2$ (blue arrow). As exotic occupancy continues to increase toward p^* , and J^* will return to its initial value (purple arrow). Only once the species becomes more widespread than its native counterparts will it be able to raise J^* above its initial value (red arrow).

The above equation can be simplified further using Harris et al. (2011)’s notion of “effective occupancy,” denoted p^* . By interpolating between the species’ p_k values, p^* acts as a “center of gravity” for average similarity: species whose p_k values exceed p^* pull its value up, while species with smaller p_k values pull p^* down. As a result, the addition of exotic species

will generally cause net differentiation until they occupy at least as many sites as their native counterparts; from then on, their spread will make the landscape more homogenous than it would have been if they hadn't invaded at all (Figure 1). The existence of such a threshold has been clear to homogenization researchers for more than a decade (McKinney 2004, Rooney et al. 2007, Rosenblad and Sax 2016), but its location has not been discussed outside of Harris et al. (2011) (nor has the location of an equally important point at $p^*/2$, shown in Figure 1).

A large portion of the homogenization literature can be explained by this single value. Species that have had longer to spread, and species whose local occupancy rates did not start at zero do not have any unique biological property that causes them to exert larger homogenizing effects on landscapes; they will simply tend to be farther along the path shown in Figure 1 [(c.f. refs]. Precisely the same argument could be made for species with high dispersal rates [refs], [etc.].

The role of covariance

One point that Harris et al. (2011) made but did not emphasize is that the residuals from the J^* approximation are given by an identity from Welsh et al. (1988):

$$\bar{J} = J^* - \frac{\text{cov}(T_{ij}, S_{ij}/T_{ij})}{\text{mean}(T_{ij})}, \quad (3)$$

where cov refers to the population covariance (rather than the more familiar sample covariance).

Equation 2 shows how \bar{J} can be exactly decomposed into an occupancy component and a covariance component¹. In 2011, we largely disregarded the covariance component because it was usually small (and approaches zero as the size of the data set increases under some fairly realistic assumptions), but it has important consequences for the way we think about homogenization. In general, it implies that landscapes whose local species compositions are determined by a small number of important factors (which will tend to produce negative covariances) will have higher mean similarity values than would be implied by their occupancy rates (Figure 2a). The covariance term also implies a possible discrepancy between an intuitive view of beta diversity and the one given by \bar{J} (Figure 2); this apparent conflict

¹Note that our 2011 treatment of this decomposition switched J^* and \bar{J} in Appendix 2's Equation B1 and some of the subsequent discussion in that Appendix is backwards.

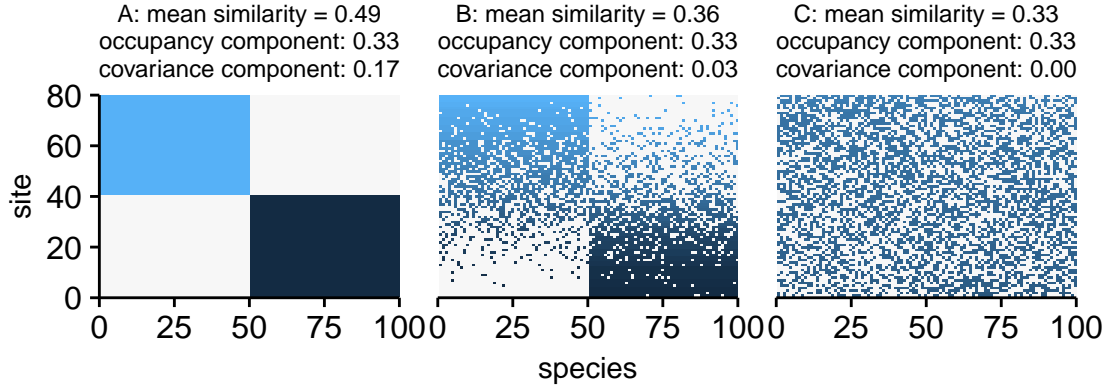


Figure 2: The effect of covariance on \bar{J} (Equation 3), holding each species' occurrence rate constant at 50% and the occupancy component constant at 0.33. In **A**, there are only two kinds of communities: light blue communities that contain species 1-50 and dark blue communities that contain species 51-100. Here, the covariance between T_{ij} and S_{ij}/T_{ij} is strongly negative and covariance contributes substantially to \bar{J} . In **B**, the two community types are not as distinct, and the covariance effect is much smaller. In **C**, there are no distinct community types and the covariance is not distinguishable from zero. Counter-intuitively, randomly smearing species across the landscape causes anti-homogenization, in terms of mean similarity.

deserves further scrutiny. If \bar{J} remains an important metric in future research, then a deeper understanding of this covariance term will be important.

Avenues for future research

If most of the variance studied in homogenization research can be explained by p^* or J^* , and the rest can be explained by a simple covariance term, where does this leave the field? At the end of Harris et al. (2011), we listed three paths forward, each of which remains promising five years later.

1. To the extent that p^* explains most of the variance, homogenization researchers should focus on explaining and predicting changes in species' occupancy rates. These are already important for other reasons in basic and applied community ecology, so predictive methods from those fields (Pearse et al. 2013, Harris 2015) could help make homogenization research more predictive.
2. Researchers should increase their focus on the covariance effects that cause deviations from J^* . Welsh et al. (1988)'s approximation involving the variance of T_{ij} will likely

be a fruitful starting point for these analyses². Alternatively, researchers could focus on the observed deviations from permutation-based null distributions (Chase 2007, Chase et al. 2011). Recent research on [ordination, JSDM; Harris (2015); refs].

3. If \bar{J} can be calculated exactly from occupancy rates and a covariance term, then a great deal of local information must be lost during the averaging process. To use that information effectively, homogenization researchers will need to think about similarity at scales below the landscape level. As the focus of homogenization research shifts toward sub-landscape scales, a deeper understanding of biotic turnover on individual pairs of sites (as provided by the frameworks of Olden and Poff (2003) and Rosenblad and Sax (2016)) will become increasingly valuable.

Finally, the potential discrepancy between \bar{J} 's behavior and the intuitive idea that randomly shuffling species should be a homogenizing process (Figure 2) implies that community ecologists should consider the use of other metrics for landscape-level homogenization.

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²Omitting subscripts, the approximation is $-\text{cov}(S/T, T)/\text{mean}(T) \approx \text{mean}(S)\text{var}(T)/\text{mean}(T)^3$. As in Equation 2, the population variance and covariance should be used rather than the sample variance/covariance.

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