

**Exercise 1 (Monte Carlo for Gaussians)**

1. Let's prove that  $E[\phi(X)] = E[\phi(X + \theta) \exp(\frac{-1}{2}\theta^T \theta - \theta^T X)]$ .

$$\begin{aligned} E[\phi(X + \theta) \exp(\frac{-1}{2}\theta^T \theta - \theta^T X)] &= \int_{\mathbb{R}^d} \phi(x + \theta) \exp(\frac{-1}{2}\theta^T \theta - \theta^T X) \pi(x) dx_1 \dots dx_d = \\ &= \int_{\mathbb{R}^d} \phi(x + \theta) \exp\left(\frac{-1}{2}\theta^T \theta - \theta^T X\right) \exp(-x^T x/2) \frac{1}{(\sqrt{2\pi})^d} dx_1 \dots dx_d = \\ &\quad \int_{\mathbb{R}^d} \phi(x + \theta) \exp\left(\frac{-1}{2}(x - \theta)^T (x - \theta)\right) \frac{1}{(\sqrt{2\pi})^d} dx_1 \dots dx_d \end{aligned}$$

Finally, making  $x - \theta = y$ ,

$$\int_{\mathbb{R}^d} \phi(y) \exp\left(\frac{-1}{2}(y)^T (y)\right) \frac{1}{(\sqrt{2\pi})^d} dx_1 \dots dx_d = E[\phi(Y)]$$

□

2. Let's show that

$$\sigma^2(\theta) = E\left[\phi^2(X) \exp\left(\frac{-1}{2}X^T X + \frac{1}{2}(X - \theta)^T (X - \theta)\right)\right] - (E[\phi(X)])^2$$

Note that, using the result in the previous item we have:

$$\begin{aligned} \sigma^2(\theta) &= V\left[\phi(X + \theta) \exp\left(\frac{-1}{2}\theta^T \theta - \theta^T X\right)\right] = \\ &= E\left[\left(\phi(X + \theta) \exp\left(\frac{-1}{2}\theta^T \theta - \theta^T X\right)\right)^2\right] - E\left[\phi(X + \theta) \exp\left(\frac{-1}{2}\theta^T \theta - \theta^T X\right)\right]^2 = \\ &= E\left[\left(\phi(X + \theta) \exp\left(\frac{-1}{2}\theta^T \theta - \theta^T X\right)\right)^2\right] - E[\phi(X)]^2 \end{aligned}$$

Now, let's rearrange the first term in the variance.

$$\sigma^2(\theta) = \int_{\mathbb{R}^d} \phi(x + \theta)^2 \exp(-\theta^T \theta - 2\theta^T X) \exp(-x^T x/2) \frac{1}{(\sqrt{2\pi})^d} dx_1 \dots dx_d =$$

Make  $X + \theta = Y$ , then:

$$\begin{aligned}
& \int_{\mathbb{R}^d} \phi(y)^2 \exp(-\theta^T \theta - 2\theta^T(y - \theta)) \exp(-(y - \theta)^T(y - \theta)/2) \frac{1}{(\sqrt{2\pi})^d} dx_1 \dots dx_d = \\
& = \int_{\mathbb{R}^d} \phi(y)^2 \exp\left(\frac{1}{2}(y - \theta)^T(y - \theta) - \frac{y^T y}{2}\right) \exp\left(\frac{-y^T y}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx_1 \dots dx_d = \\
& = E\left[\phi^2(X) \exp\left(\frac{-1}{2}X^T X + \frac{1}{2}(X - \theta)^T(X - \theta)\right)\right]
\end{aligned}$$

Therefore,

$$\sigma^2(\theta) = E\left[\phi^2(X) \exp\left(\frac{-1}{2}X^T X + \frac{1}{2}(X - \theta)^T(X - \theta)\right)\right] - (E[\phi(X)])^2$$

□

3. Let's calculate  $\nabla^2 \sigma^2(\theta) = H(\theta)$ .

$$\begin{aligned}
\frac{\partial \sigma^2(\theta)}{\partial \theta_i} &= \frac{E[\phi(X)^2 \exp(\frac{-X^T X + (X - \theta)^T(X - \theta)}{2})]}{\partial \theta_i} = \\
&= \int_{\chi} \phi(x)^2 \exp(-x^T x) \frac{\partial}{\partial \theta_i} \exp\left(\frac{(x - \theta)^T(x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx = \\
&= \int_{\chi} \phi(x)^2 \exp(-x^T x) (\theta_i - x_i) \exp\left(\frac{(x - \theta)^T(x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx
\end{aligned}$$

We calculated the gradient, let's now calculate the second derivative. First the diagonal.

$$\begin{aligned}
& \frac{\partial}{\partial \theta_i} \int_{\chi} \phi(x)^2 \exp(-x^T x) (\theta_i - x_i) \exp\left(\frac{(x - \theta)^T(x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx = \\
& = E[\phi(X)^2] + \int_{\chi} \phi(x)^2 \exp(-x^T x) \exp\left(\frac{(x - \theta)^T(x - \theta)}{2}\right) (x_i - \theta_i)(x_i - \theta_i) \frac{1}{(\sqrt{2\pi})^d} dx
\end{aligned}$$

Now the rest:

$$\begin{aligned}
& \frac{\partial}{\partial \theta_j} \int_{\chi} \phi(x)^2 \exp(-x^T x) (\theta_i - x_i) \exp\left(\frac{(x - \theta)^T(x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx = \\
& = \int_{\chi} \phi(x)^2 \exp(-x^T x) \exp\left(\frac{(x - \theta)^T(x - \theta)}{2}\right) (x_i - \theta_i)(x_j - \theta_j) \frac{1}{(\sqrt{2\pi})^d} dx
\end{aligned}$$

□

4. We already know that the Hessian is positive definite. Hence, we only need to show that the derivative is equal to zero at  $\theta^*$ .

$$\begin{aligned}
 \nabla \sigma^2(\theta) &= \int_{\mathcal{X}} \phi(x)^2 \exp(-x^T x) (\theta - x) \exp\left(\frac{-(x - \theta)^T (x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx = \\
 &= \int_{\mathcal{X}} \phi(x)^2 \exp(-x^T x) (\theta - x) \exp\left(\frac{-(x - \theta)^T (x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx = \\
 &= \int_{\mathcal{X}} \phi(x)^2 (\theta - x) \exp\left(\frac{-x^T x}{2} - \theta^T x + \frac{-\theta^T \theta}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx = \\
 &= \int_{\mathcal{X}} \phi(x)^2 (\theta - x) \exp(-\theta^T x) \exp(-x^T x/2) \frac{\exp(\theta^T \theta/2)}{(\sqrt{2\pi})^d} dx = 0
 \end{aligned}$$

Finally, since the last term doesn't depend on  $X$ , we can eliminate it, obtaining:

$$E[\phi(X)^2 (\theta - X) \exp(-\theta^T X)] = 0$$

□

5. Couldn't solve.

**Exercise 2 (Metropolis-Hastings)**

1. First, generate  $Y \sim q(X_t, \cdot)$ . With probability equal to  $\alpha(x, y)$ , make  $X_{t+1} = Y$ , and with probability  $1 - \sum_{z \in \mathbb{Z}} \alpha(x, z)q(x, z)$  make  $X_{t+1} = X_t$ . Then, repeat the process.
2. We want to show that  $\pi(x)T(x, y) = \pi(y)T(y, x)$ . For  $y = x$ , this is trivial. Now, assume  $y \neq x$ . Therefore:

$$\begin{aligned}\pi(x)T(x, y) &= \pi(x)\alpha(x, y)q(x, y) = \pi(x)\frac{\gamma(x, y)}{\pi(x)q(x, y)} = \gamma(y, x) = \\ &= \pi(y)\frac{\gamma(y, x)q(y, x)}{\pi(y)q(y, x)} = \pi(y)T(y, x)\end{aligned}$$

□

3. Using Metropolis-Hastings, one has:

$$\alpha = \min \left\{ 1, \frac{\pi(x^*)q(x_{t-1} | x^*)}{\pi(x_{t-1})q(x^* | x_{t-1})} \right\}$$

So,  $\alpha(x, y) = \frac{\gamma(x, y)}{\pi(x)q(x, y)}$ , for  $x = x_{t-1}, x^* = y, q(x, y) = q(x^* | x_{t-1})$ .

Make:

$$\begin{aligned}\gamma(x, y) &= \max \{ \pi(y)q(y, x), \pi(x)q(x, y) \} \\ &\quad \therefore \\ \alpha(x, y) &= \frac{\max \{ \pi(y)q(y, x), \pi(x)q(x, y) \}}{\pi(x)q(x, y)} = \\ &= \min \left\{ 1, \frac{\pi(y)q(y, x)}{\pi(x)q(x, y)} \right\}\end{aligned}$$

□

4. First, let's rewrite the estimate as a function of  $Y$ . Note that  $Y^{(k)} = X^{(\tau_k)}$ , hence:

$$\frac{1}{\tau_k - 1} \sum_{t=1}^{\tau_k-1} \phi(X^{(t)}) = \frac{1}{k - 1} \sum_{t=1}^{k-1} \phi(Y^{(t)})$$

Now, let's prove it has the transition desired transition kernel.

$$K(x, y) = P(Y^{(k)} = y | Y^{(k-1)} = x) = P(X^{(\tau_k)} = y | X^{(\tau_k-1)} = x)$$

The event  $X^{(\tau_k)} = y | X^{(\tau_k-1)} = x$  is equivalent to selecting  $y$  starting from  $x$  and then accepting  $y$ , hence  $q(x, y) \cdot \alpha(x, y) = P(X^{(\tau_k)} = y | X^{(\tau_k-1)} = x)$ .

Therefore,  $P(Y^{(k)} = y | Y^{(k-1)} = x) = \frac{q(x, y)\alpha(x, y)}{\sum_{z \in \mathbb{Z}} \alpha(x, z)q(x, z)}$ .

□

5. Let's show that  $\tilde{\pi}(x)K(x, y) = \tilde{\pi}(y)K(y, x)$ .

$$\frac{\pi(x)m(x)}{\sum_{z \in \mathbb{Z}} \pi(z)m(z)} \cdot \frac{q(x, y)\alpha(x, y)}{\sum_{z \in \mathbb{Z}} \alpha(x, z)q(x, z)} = \frac{\pi(x)q(x, y)\alpha(x, y)}{\sum_{z \in \mathbb{Z}} \pi(z)m(z)}$$

Note that  $\alpha(x, y) = \frac{\gamma(x, y)}{\pi(x)q(x, y)} = \frac{\gamma(y, x)}{\pi(x)q(x, y)}$ . Therefore:

$$\frac{\pi(x)q(x, y)\gamma(x, y)}{\alpha(x, z)q(x, z) \sum_{z \in \mathbb{Z}} \pi(z)m(z)} = \frac{\gamma(y, x)}{\sum_{z \in \mathbb{Z}} \pi(z)m(z)} = \frac{\alpha(y, x)\pi(y)q(y, x)}{\sum_{z \in \mathbb{Z}} \pi(z)m(z)} = \tilde{\pi}(y)K(y, x)$$

□

6. Couldn't solve.