## Exercise 1 (Monte Carlo for Gaussians)

1. Let's prove that  $E[\phi(X)] = E[\phi(X + \theta)exp(\frac{-1}{2}\theta^T\theta - \theta^TX)].$ 

$$E[\phi(X+\theta)exp(\frac{-1}{2}\theta^T\theta - \theta^TX)] = \int_{\mathbb{R}^d} \phi(x+\theta)exp(\frac{-1}{2}\theta^T\theta - \theta^TX)\pi(x)dx_1...dx_d =$$

$$= \int_{\mathbb{R}^d} \phi(x+\theta)exp\left(\frac{-1}{2}\theta^T\theta - \theta^TX\right)exp(-x^Tx/2)\frac{1}{(\sqrt{2\pi})^d}dx_1...dx_d =$$

$$\int_{\mathbb{R}^d} \phi(x+\theta)exp\left(\frac{-1}{2}(x-\theta)^T(x-\theta)\right)\frac{1}{(\sqrt{2\pi})^d}dx_1...dx_d$$

Finally, making  $x - \theta = y$ ,

$$\int_{\mathbb{R}^d} \phi(y) exp\left(\frac{-1}{2}(y)^T(y)\right) \frac{1}{(\sqrt{2\pi})^d} dx_1...dx_d = E[\phi(Y)]$$

2. Let's show that

$$\sigma^{2}(\theta) = E\left[\phi^{2}(X)exp\left(\frac{-1}{2}X^{T}X + \frac{1}{2}(X - \theta)^{T}(X - \theta)\right)\right] - (E[\phi(X)]^{2}$$

Note that, using the result in the previous item we have:

$$\sigma^2(\theta) = V \left[ \phi(X + \theta) exp \left( \frac{-1}{2} \theta^T \theta - \theta^T X \right) \right] =$$

$$\begin{split} &= E\left[\left(\phi(X+\theta)exp\left(\frac{-1}{2}\theta^T\theta - \theta^TX\right)\right)^2\right] - E\left[\phi(X+\theta)exp\left(\frac{-1}{2}\theta^T\theta - \theta^TX\right)\right]^2 = \\ &= E\left[\left(\phi(X+\theta)exp\left(\frac{-1}{2}\theta^T\theta - \theta^TX\right)\right)^2\right] - E\left[\phi(X)\right]^2 \end{split}$$

Now, let's rearrange the first term in the variance.

$$\sigma^{2}(\theta) = \int_{\mathbb{R}^{d}} \phi(x+\theta)^{2} exp\left(-\theta^{T}\theta - 2\theta^{T}X\right) exp(-x^{T}x/2) \frac{1}{(\sqrt{2\pi})^{d}} dx_{1}...dx_{d} =$$

Make  $X + \theta = Y$ , then:

$$\int_{\mathbb{R}^d} \phi(y)^2 exp\left(-\theta^T \theta - 2\theta^T (y - \theta)\right) exp(-(y - \theta)^T (y - \theta)/2) \frac{1}{(\sqrt{2\pi})^d} dx_1 ... dx_d =$$

$$= \int_{\mathbb{R}^d} \phi(y)^2 exp\left(\frac{1}{2} (y - \theta)^T (y - \theta) - \frac{y^T y}{2}\right) exp\left(\frac{-y^T y}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx_1 ... dx_d =$$

$$= E\left[\phi^2(X) exp\left(\frac{-1}{2} X^T X + \frac{1}{2} (X - \theta)^T (X - \theta)\right)\right]$$

Therefore,

$$\sigma^{2}(\theta) = E\left[\phi^{2}(X)exp\left(\frac{-1}{2}X^{T}X + \frac{1}{2}(X - \theta)^{T}(X - \theta)\right)\right] - (E[\phi(X)]^{2}$$

3. Let's calculate  $\nabla^2 \sigma^2(\theta) = H(\theta)$ .

$$\frac{\partial \sigma^2(\theta)}{\partial \theta_i} = \frac{E[\phi(X)^2 exp(\frac{-X^T X + (X - \theta)^T (X - \theta)}{2})]}{\partial \theta_i} =$$

$$= \int_{\chi} \phi(x)^2 exp(-x^T x) \frac{\partial}{\partial \theta_i} exp\left(\frac{(x - \theta)^T (x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx =$$

$$= \int_{\chi} \phi(x)^2 exp(-x^T x) (\theta_i - x_i) exp\left(\frac{(x - \theta)^T (x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx$$

We calculated the gradient, let's now calculate the second derivative. First the diagonal.

$$\frac{\partial}{\partial \theta_i} \int_{\chi} \phi(x)^2 exp(-x^T x) (\theta_i - x_i) exp\left(\frac{(x - \theta)^T (x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx =$$

$$= E[\phi(X)^2] + \int_{\chi} \phi(x)^2 exp(-x^T x) exp\left(\frac{(x - \theta)^T (x - \theta)}{2}\right) (x_i - \theta_i) (x_i - \theta_i) \frac{1}{(\sqrt{2\pi})^d} dx$$

Now the rest:

$$\frac{\partial}{\partial \theta_j} \int_{\chi} \phi(x)^2 exp(-x^T x) (\theta_i - x_i) exp\left(\frac{(x - \theta)^T (x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^d} dx =$$

$$= \int_{\chi} \phi(x)^2 exp(-x^T x) exp\left(\frac{(x - \theta)^T (x - \theta)}{2}\right) (x_i - \theta_i) (x_j - \theta_j) \frac{1}{(\sqrt{2\pi})^d} dx$$

4. We already know that the Hessian is positive definite. Hence, we only need to show that the derivative is equal to zero at  $\theta^*$ .

$$\nabla \sigma^{2}(\theta) = \int_{\chi} \phi(x)^{2} exp(-x^{T}x)(\theta - x) exp\left(\frac{-(x - \theta)^{T}(x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^{d}} dx =$$

$$= \int_{\chi} \phi(x)^{2} exp(-x^{T}x)(\theta - x) exp\left(\frac{-(x - \theta)^{T}(x - \theta)}{2}\right) \frac{1}{(\sqrt{2\pi})^{d}} dx =$$

$$= \int_{\chi} \phi(x)^{2}(\theta - x) exp\left(\frac{-x^{T}x}{2} - \theta^{T}x + \frac{-\theta^{T}\theta}{2}\right) \frac{1}{(\sqrt{2\pi})^{d}} dx =$$

$$= \int_{\chi} \phi(x)^{2}(\theta - x) exp(-\theta^{T}x) exp(-x^{T}x/2) \frac{exp(\theta^{T}\theta/2)}{(\sqrt{2\pi})^{d}} dx = 0$$

Finally, since the last term doens't depend on X, we can eliminate it, obtaining:

$$E[\phi(X)^{2}(\theta - X)exp(-\theta^{T}X)] = 0$$

5. Couldn't solve.

## Exercise 2 (Metropolis-Hastings)

- 1. First, generate  $Y \sim q(X_t, \cdot)$ . With probability equal to  $\alpha(x, y)$ , make  $X_{t+1} = Y$ , and with probability  $1 \sum_{z \in \mathbb{Z}} \alpha(x, z) q(x, z)$  make  $X_{t+1} = X_t$ . Then, repeat the process.
- 2. We want to show that  $\pi(x)T(x,y)=\pi(y)T(y,x)$ . For y=x, this is trivial. Now, assume  $y\neq x$ . Therefore:

$$\pi(x)T(x,y) = \pi(x)\alpha(x,y)q(x,y) = \pi(x)\frac{\gamma(x,y)}{\pi(x)q(x,y)} = \gamma(y,x) =$$
$$= \pi(y)\frac{\gamma(y,x)q(y,x)}{\pi(y)q(y,x)} = \pi(y)T(y,x)$$

3. Using Metropolis-Hastings, one has:

$$\alpha = \min \left\{ 1, \frac{\pi(x^*)q(x_{t-1}) \mid x^*}{\pi(x_{t-1})q(x^* \mid x_{t-1})} \right\}$$

So,  $\alpha(x,y) = \frac{\gamma(x,y)}{\pi(x)q(x,y)}$ , for  $x = x_{t-1}, x^* = y, q(x,y) = q(x^* \mid x_{t-1})$ .

Make:

$$\gamma(x,y) = \max\left\{\pi(y)q(y,x), \pi(x)q(x,y)\right\}$$

 $\alpha(x,y) = \frac{\max\{\pi(y)q(y,x),\pi(x)q(x,y)\}}{\pi(x)q(x,y)} = \min\left\{1,\frac{\pi(y)q(y,x)}{\pi(x)q(x,y)}\right\}$ 

4. First, let's rewrite the estimate as a function of Y. Note that  $Y^{(k)} = X^{(\tau_k)}$ , hence:

$$\frac{1}{\tau_k - 1} \sum_{t=1}^{\tau_k - 1} \phi(X^{(t)}) = \frac{1}{k - 1} \sum_{t=1}^{k - 1} \phi(Y^{(t)})$$

Now, let's prove it has the transition desired transition kernel.

$$K(x,y) = P(Y^{(k)} = y \mid Y^{(k-1)} = x) = P(X^{(\tau_k)} = y \mid X^{(\tau_k-1)} = x)$$

The event  $X^{(\tau_k)} = y \mid X^{(\tau_k-1)} = x$  is equivalent to selecting y starting from x and then accepting y, hence  $q(x,y) \cdot \alpha(x,y) = P(X^{(\tau_k)} = y \mid X^{(\tau_k-1)} = x)$ .

Therefore, 
$$P(Y^{(k)} = y \mid Y^{(k-1)} = x) = \frac{q(x,y)\alpha(x,y)}{\sum_{z \in \mathbb{Z}} \alpha(x,z)q(x,y)}$$
.

5. Let's show that  $\tilde{\pi}(x)K(x,y) = \tilde{\pi}(y)K(y,x)$ .

$$\frac{\pi(x)m(x)}{\sum_{z\in\mathbb{Z}}\pi(z)m(z)}\cdot\frac{q(x,y)\alpha(x,y)}{\sum_{z\in\mathbb{Z}}\alpha(x,z)q(x,z)}=\frac{\pi(x)q(x,y)\alpha(x,y)}{\sum_{z\in\mathbb{Z}}\pi(z)m(z)}$$

Note that  $\alpha(x,y) = \frac{\gamma(x,y)}{\pi(x)q(x,y)} = \frac{\gamma(y,x)}{\pi(x)q(x,y)}$ . Therefore:

$$\frac{\pi(x)q(x,y)\gamma(x,y)}{\alpha(x,z)q(x,z)\sum_{z\in\mathbb{Z}}\pi(z)m(z)} = \frac{\gamma(y,x)}{\sum_{z\in\mathbb{Z}}\pi(z)m(z)} = \frac{\alpha(y,x)\pi(y)q(y,x)}{\sum_{z\in\mathbb{Z}}\pi(z)m(z)} = \tilde{\pi}(y)K(y,x)$$

6. Couldn't solve.

## Exercise 3 (Metropolis-Hastings)

1. Let's show that  $\int_{\chi} \pi(x)(\alpha(x)q(y)(1-\alpha(x)))\delta_x(y)dx = \pi(y)$ . Note that,  $\pi(x) = \frac{q(x)}{\alpha(x)Z_{\pi}}$ , hence:

$$\int_{\chi} \pi(x)(\alpha(x)q(y)(1-\alpha(x)))\delta_x(y)dx = \int_{\chi} \frac{q(x)}{\alpha(x)Z_{\pi}}(\alpha(x)q(y)(1-\alpha(x)))\delta_x(y)dx =$$

$$= \int_{\chi} \frac{q(x)}{\alpha(x)Z_{\pi}}\alpha(x)q(y)dx + \pi(y)(1-\alpha(y)) = \frac{q(y)}{Z_{\pi}}\int_{\chi} q(x)dx + \pi(y)(1-\alpha(y)) =$$

$$= \pi(y)\alpha(y) + \pi(y)(1-\alpha(y)) = \pi(y)$$

2. Couldn't solve.

## Exercise 4 (Gibbs Sampler)

1.

$$\pi(x \mid y) = \frac{\pi(x,y)}{\pi(y)} \propto \frac{exp((x-1)^2(y-2)^2)/2}{(y-2)^{-1}} \sim N(1,(y-2)^{-2})$$

Now, follow the same procedure to for  $\pi(y \mid x)$ , hence:

$$\pi(y \mid x) \sim N(2, (x-1)^{-2})$$

2. The sampler doesn't make sense, because:

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \pi(x, y) dx dy \propto \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp((x - 1)^{2} (y - 2)^{2}) / 2 dx dy =$$

$$= \int_{-\infty}^{\infty} \sqrt{2\pi} ((y - 2)^{-2})^{\frac{1}{2}} dy = \infty$$