

Confidence Interval

$$P\left(\theta \in \left[\underline{L}(x), \underline{U}(x)\right]\right) \geq 1 - \alpha$$

Pivot Exact Pivot

~~Asymptotical~~ Pivot

$$X_i \sim \text{Bernoulli}(p) \quad \hat{p} = \frac{\sum X_i}{n}$$

$$\hat{p} \sim N\left(p, \frac{p(1-p)}{n}\right)$$

$$\frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}} \sim Z \quad \text{CI: } \hat{p} \pm Z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

7.22, USI Variance Stabilization Transform

$$X_i \sim \text{Bernoulli}(p)$$

$$\hat{p} = \frac{\sum X_i}{n} \quad E \hat{p} = p \quad V(\hat{p}) = \frac{p(1-p)}{n}$$

Assume that X has the mean μ , and $V(x) = v(\mu)$.

$$\text{then } g(x) = \int_0^x \frac{1}{\sqrt{v(\mu)}} d\mu$$

$$\text{Let } Y = g(x).$$

$$X = \hat{p} \quad V(\hat{p}) = \boxed{\frac{p(1-p)}{n}}$$

$$g(x) = \int_0^x \frac{1}{\sqrt{\frac{p(1-p)}{n}}} dp = \sqrt{n} \int_0^x \frac{1}{\sqrt{p(1-p)}} dp$$

$$= \sqrt{n} \quad 2 \arcsin \sqrt{x}.$$

$$\dots, \quad 1 \quad 1 \quad n^{-\frac{1}{2}} \quad 1 \quad 1 \quad 1 \quad 1$$

$$-\sqrt{n} < \arcsin \sqrt{x}$$

$$2 (\arcsin \sqrt{x})' = 2 \cdot \frac{1}{\sqrt{1-(\sqrt{x})^2}} \cdot \frac{1}{2} x^{-\frac{1}{2}} = \frac{1}{\sqrt{x(1-x)}}$$

VST. Let $Y = g(\hat{p}) = 2\sqrt{n} \arcsin \sqrt{\hat{p}}$

Delta Method

Thm: For a sequence of r.v. W_n if

$$a_n(W_n - b) \xrightarrow{D} N(0,1)$$

Then
$$\frac{a_n(g(W_n) - g(b))}{|g'(b)|} \xrightarrow{D} N(0,1)$$

a_n
// For the Bernoulli:

$$\sqrt{\frac{n}{p(1-p)}} (\hat{p} - p) \xrightarrow{D} N(0,1) \quad \frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}} \rightarrow N(0,1)$$

$$\frac{\sqrt{\frac{n}{p(1-p)}} (2\sqrt{n} \arcsin \sqrt{\hat{p}} - 2\sqrt{n} \arcsin \sqrt{p})}{\sqrt{\frac{n}{p(1-p)}}} \rightarrow N(0,1)$$

$$g(p) = 2\sqrt{n} \arcsin \sqrt{p} \quad g'(p) = 2\sqrt{n} \cdot \frac{1}{\sqrt{1-(\sqrt{p})^2}} \cdot \frac{1}{2} \frac{1}{\sqrt{p}} = \frac{\sqrt{n}}{\sqrt{p(1-p)}}$$

$$2\sqrt{n} (\arcsin \sqrt{\hat{p}} - \arcsin \sqrt{p}) \rightarrow N(0,1)$$

$$P \left(-Z_{\frac{\alpha}{2}} < 2\sqrt{n} (\arcsin \sqrt{\hat{p}} - \arcsin \sqrt{p}) < Z_{\frac{\alpha}{2}} \right) \geq 1 - \alpha$$

$$\square \text{ for } \arcsin \sqrt{p} \text{ is } \arcsin \sqrt{p} \pm Z_{\frac{\alpha}{2}} \cdot \frac{1}{2\sqrt{n}}$$

$$CI \text{ for } p \text{ is. } \left[\sin^2 \left(\arcsin \sqrt{\hat{p}} - z_{\frac{\alpha}{2}} \frac{1}{\sqrt{n}} \right), \sin^2 \left(\arcsin \sqrt{\hat{p}} + z_{\frac{\alpha}{2}} \frac{1}{\sqrt{n}} \right) \right]$$

Example 7.2.3. $CI = (0.38, 0.44)$ based on VST

For discrete r.v. mean and variance are dependent.

$$p \sqrt{\frac{p(1-p)}{n}}$$

7.2.3: Bootstrap Confidence Interval.

population mean / median

$X_i \stackrel{\text{ind}}{\sim} \text{Normal}$

$\hat{p} \quad X_i \sim \text{Bernoulli}(p)$

Nonparametric Confidence Interval

If we have all the data,

Sample 1000 students. \mapsto

Sample 1000 students \mapsto

10,000 times

$$\begin{array}{c} \overline{X}_1 \\ \overline{X}_1 \\ \vdots \\ \overline{X}_{10,000} \end{array}$$

$$CI. [X_{(250)}, X_{(9750)}] \mapsto \underline{95\%}$$

Bradley Efron Bootstrap Treat the data as the full population

we sample with replacement

For each Bootstrap sample, calculate the mean $\hat{\mu}_j^B$

Replicate $B = 10,000$ times

① $\hat{\mu}_1^B, \hat{\mu}_2^B, \hat{\mu}_3^B, \dots, \hat{\mu}_B^B$

~~$\hat{\mu} = \frac{\hat{\mu}_1 + \hat{\mu}_2 + \dots + \hat{\mu}_B}{B}$~~

② $V^B(\hat{\mu}) = \text{sample variance of } \hat{\mu}_j^B$

③a $\hat{\mu} \pm \underline{Z_{\frac{\alpha}{2}} \sqrt{V^B(\hat{\mu})}}$ $\left. \vphantom{\hat{\mu} \pm Z_{\frac{\alpha}{2}} \sqrt{V^B(\hat{\mu})}} \right\} (1-\alpha) - \text{CI.}$

③b $\left[\hat{\mu}_{(\frac{\alpha}{2} \times B)}, \hat{\mu}_{(1-\frac{\alpha}{2}) \times B} \right]$

$\hat{\mu} = \bar{X}$

Example (Law School Data)

LSAT GPA

What is the correlation? What is the CI for the correlation?

$$\hat{\rho} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum (x_i - \bar{x})^2 \cdot \sum (y_i - \bar{y})^2}}$$

Fisher Transformation

$$g = 0.5 \log \frac{1+\rho}{1-\rho} \quad \hat{g} = 0.5 \log \frac{1+\hat{\rho}}{1-\hat{\rho}}$$

$$\hat{\theta} \sim N(\theta, v(\hat{\theta}))$$

7.3. CI for two populations

Independent samples μ_1, μ_2

(i) Equal Variances ($\sigma_1^2 = \sigma_2^2$)

$$S_p^2 = \frac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1 + n_2 - 2}$$

CI $\mu_1 - \mu_2$ as

$$(\bar{y}_1 - \bar{y}_2) \pm t_{\frac{\alpha}{2}, n_1+n_2-2} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

(ii) Unequal variance

$$v(\bar{y}_1 - \bar{y}_2) = v(\bar{y}_1) + v(\bar{y}_2) = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$$

CI for $\mu_1 - \mu_2$ is

$$(\bar{y}_1 - \bar{y}_2) \pm t_{\frac{\alpha}{2}, df} \sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}$$

Example: Microarray. 1st and 2nd gene

7.3.2 Variances CI for $\frac{\sigma_1^2}{\sigma_2^2}$

$$\frac{S_2^2/\sigma_2^2}{S_1^2/\sigma_1^2} \sim \frac{X_{n_2-1}^2/n_{2-1}}{X_{n_1-1}^2/n_{1-1}} \sim F_{n_2-1, n_1-1}$$

$$F_{\frac{\alpha}{2}, n_1-1, n_2-1} < \frac{S_2^2/\sigma_2^2}{S_1^2/\sigma_1^2} < F_{1-\frac{\alpha}{2}, n_1-1, n_2-1}$$

$$\frac{S_1^2}{S_2^2} F_{\frac{\alpha}{2}, n_2-1, n_1-1} < \frac{\sigma_1^2}{\sigma_2^2} < \frac{S_1^2}{S_2^2} F_{1-\frac{\alpha}{2}, n_1-1, n_2-1}$$

$$[\quad , \quad]$$

$$t_{\frac{\alpha}{2}, n_2-1, n_1-1} = \frac{1}{F_{1-\frac{\alpha}{2}, n_1-1, n_2-1}}$$

7.3.3 Two sample proportions.

One sample, CI: $\hat{p} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$

populations 1 and 2

n_1, x_1	$\hat{p}_1 = \frac{x_1}{n_1}$
n_2, x_2	$\hat{p}_2 = \frac{x_2}{n_2}$

$P_1 :$

$P_2 :$

CI for $P_1 - P_2$ is

$$(\hat{p}_1 - \hat{p}_2) \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$$

$$2\sqrt{n} \arcsin \sqrt{p_1} - 2\sqrt{n} \arcsin \sqrt{p_2}$$

$$P_1 - P_2 ?$$

Agresti and Caffo: $\tilde{p}_i = \frac{x_i + 1}{n_i + 2}$

Then the CI for $P_1 - P_2$ is

$$\tilde{p}_1 - \tilde{p}_2 \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\tilde{p}_1(1-\tilde{p}_1)}{n_1+2} + \frac{\tilde{p}_2(1-\tilde{p}_2)}{n_2+2}}$$

Example 7.3.3. (Exit Poll)

Sample Size Calculation

$$\underline{1-\alpha}$$

$$\text{Length / Width} \mid \text{MOE} = \frac{\text{Length}}{2}$$

$$(1-\alpha) \nearrow, \quad \text{Length} \nearrow \\ \text{Length / MOE} \searrow, \quad (1-\alpha) \searrow$$

Enough Sample size

CI for the mean σ^2 , $(1-\alpha)$, MOE

$$\bar{x} \pm Z_{1-\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}$$

$$Z_{1-\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}} = \text{MOE}$$

$$n > \frac{\sigma^2 (Z_{1-\frac{\alpha}{2}})^2}{(\text{MOE})^2}$$

Sample size for proportion:

$$\hat{p} \pm Z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n+2}}$$

$$Z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n+2}} = \text{MOE}$$

$$Z_{\frac{\alpha}{2}}^2 \frac{\hat{p}(1-\hat{p})}{n+2} = \text{MOE}^2 \Rightarrow n+2 = \frac{\hat{p}(1-\hat{p}) \cdot Z_{\frac{\alpha}{2}}^2}{\text{MOE}^2}$$

Worst case: $\hat{p} = 1 - \hat{p} = \frac{1}{2}$

$$n \geq \frac{\frac{1}{4} Z_{\frac{\alpha}{2}}^2}{\text{MOE}^2} - 2 = \frac{Z_{\frac{\alpha}{2}}^2}{4 \text{MOE}^2} - 2$$