

VI-Unified f(.) Functions Documentation

0.1 exact_linear_posterior()

Purpose: Computes the analytical conjugate posterior for linear regression under Normal-Inverse-Gamma prior.

Arguments:

- **X**: Design matrix ($N \times K$)
- **y**: Response vector ($N \times 1$)
- **prior_mu**: Prior mean for β (default: $\mathbf{0}_K$)
- **prior_sigma**: Prior standard deviation for β (default: 10)
- **prior_a**: Inverse-Gamma shape parameter for σ^2 (default: 2)
- **prior_b**: Inverse-Gamma rate parameter for σ^2 (default: 2)

Calculations:

Prior: $\beta \sim N(\mu_0, \sigma^2 \Lambda_0^{-1})$, $\sigma^2 \sim \text{IG}(a_0, b_0)$

Where $\Lambda_0 = \text{diag}(K)/\sigma_{\text{prior}}^2$

Posterior parameters:

$$\begin{aligned}\Sigma_n &= (\Lambda_0 + \mathbf{X}^\top \mathbf{X})^{-1} \\ \mu_n &= \Sigma_n (\Lambda_0 \mu_0 + \mathbf{X}^\top \mathbf{y}) \\ a_n &= a_0 + N/2 \\ b_n &= b_0 + \frac{1}{2} \left(\mathbf{y}^\top \mathbf{y} - \mu_n^\top \Sigma_n^{-1} \mu_n \right)\end{aligned}$$

Returns:

- **mu_n**: Posterior mean μ_n for β (stored as **mu_n**)
- **Sigma_n**: Posterior covariance Σ_n (conditional on σ^2 , stored as **Sigma_n**)
- **a_n**: Posterior Inverse-Gamma shape a_n (stored as **a_n**)
- **b_n**: Posterior Inverse-Gamma rate b_n (stored as **b_n**)
- **df**: Degrees of freedom $2a_n$
- **scale_matrix**: $(b_n/a_n)\Sigma_n$ for multivariate-t marginal

0.2 `sample_exact_posterior()`

Purpose: Draws samples from the Normal-Inverse-Gamma posterior using conditional sampling.

Arguments:

- `exact_post`: List returned by `exact_linear_posterior()`
- `n_samples`: Number of posterior samples (default: 2000)

Calculations:

Two-stage sampling:

1. Draw $\sigma^2 \sim \text{IG}(a_n, b_n)$ via $\sigma^2 = 1/\text{Gamma}(a_n, b_n)$
2. Draw $\boldsymbol{\beta} \mid \sigma^2 \sim N(\boldsymbol{\mu}_n, \sigma^2 \boldsymbol{\Sigma}_n)$

Repeat for `n_samples` iterations.

Returns:

- Matrix ($n_{\text{samples}} \times K + 1$) with columns: $[\beta_1, \dots, \beta_K, \sigma]$

0.3 run_laplace()

Purpose: Laplace approximation via maximum a posteriori (MAP) estimation and inverse Hessian.

Arguments:

- **X:** Design matrix for fixed effects ($N \times p$)
- **Z:** Design matrix for random effects ($N \times q$)
- **y:** Response vector ($N \times 1$)
- **p:** Number of fixed effects
- **q:** Number of random effects
- **n:** Number of observations
- **alpha_e, gamma_e:** Gamma prior parameters for τ_e
- **alpha_u, gamma_u:** Gamma prior parameters for τ_u
- **model_type:** "M1" (linear) or "M3" (hierarchical)

Calculations:

For M1: Optimise $\boldsymbol{\theta} = (\boldsymbol{\beta}, \log \tau_e)$

Negative log-posterior:

$$-\log p(\boldsymbol{\theta} \mid \mathbf{y}) = - \left[\underbrace{-\frac{N}{2} \log(2\pi) + \frac{N}{2} \log \tau_e - \frac{\tau_e}{2} \sum r_i^2}_{\text{likelihood}} + \underbrace{(\alpha_e - 1) \log \tau_e - \gamma_e \tau_e}_{\text{prior for } \tau_e} + \underbrace{-\frac{1}{200} \|\boldsymbol{\beta}\|^2}_{\text{prior for } \boldsymbol{\beta}} \right]$$

Where $r_i = y_i - \mathbf{X}_i \boldsymbol{\beta}$

For M3: Optimise $\boldsymbol{\theta} = (\boldsymbol{\beta}, \mathbf{u}, \log \tau_e, \log \tau_u)$

Additional prior term for random effects:

$$-\frac{q}{2} \log(2\pi) + \frac{q}{2} \log \tau_u - \frac{\tau_u}{2} \mathbf{u}^\top \mathbf{K}^{-1} \mathbf{u}$$

Plus prior: $(\alpha_u - 1) \log \tau_u - \gamma_u \tau_u$

Use BFGS optimisation to find $\hat{\boldsymbol{\theta}}_{\text{MAP}}$, compute Hessian \mathbf{H} , approximate:

$$q(\boldsymbol{\theta}) \approx N(\hat{\boldsymbol{\theta}}_{\text{MAP}}, \mathbf{H}^{-1})$$

Returns:

- **theta_map:** MAP estimate $\hat{\boldsymbol{\theta}}_{\text{MAP}}$
- **cov_matrix:** Inverse Hessian \mathbf{H}^{-1}
- **model_type, p, q:** Metadata

0.4 run_gibbs_sampler()

Purpose: Full Gibbs sampler (MCMC) for linear or hierarchical model.

Arguments:

- `X, Z, y, p, q, n`: As in `run_laplace()`
- `alpha_e, gamma_e, alpha_u, gamma_u`: Prior hyperparameters
- `model_type`: "M1" or "M3"
- `n_iter`: Total MCMC iterations (default: 5000)
- `n_burnin`: Burn-in iterations (default: 1000)

Calculations:

For M1: Gibbs sampling for (β, τ_e)

Full conditionals:

$$\begin{aligned}\beta \mid \tau_e, \mathbf{y} &\sim N(\mu_\beta, \Sigma_\beta) \\ \Sigma_\beta &= (\tau_e \mathbf{X}^\top \mathbf{X} + \mathbf{I}_p/100)^{-1} \\ \mu_\beta &= \tau_e \Sigma_\beta \mathbf{X}^\top \mathbf{y} \\ \tau_e \mid \beta, \mathbf{y} &\sim \text{Gamma}(a_{\text{post}}, b_{\text{post}}) \\ a_{\text{post}} &= \alpha_e + N/2 \\ b_{\text{post}} &= \gamma_e + \frac{1}{2} \sum (y_i - \mathbf{X}_i \beta)^2\end{aligned}$$

For M3: Gibbs sampling for $(\beta, \mathbf{u}, \tau_e, \tau_u)$

Joint update for (β, \mathbf{u}) :

$$\begin{aligned}\begin{bmatrix} \beta \\ \mathbf{u} \end{bmatrix} \mid \tau_e, \tau_u, \mathbf{y} &\sim N(\mu_{\beta u}, \Sigma_{\beta u}) \\ \Sigma_{\beta u} &= \left(\tau_e [\mathbf{X}|\mathbf{Z}]^\top [\mathbf{X}|\mathbf{Z}] + \tau_u \begin{bmatrix} \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{K}^{-1} \end{bmatrix} + \mathbf{I}_{p+q}/100 \right)^{-1} \\ \mu_{\beta u} &= \tau_e \Sigma_{\beta u} [\mathbf{X}|\mathbf{Z}]^\top \mathbf{y}\end{aligned}$$

Update τ_e :

$$\tau_e \mid \beta, \mathbf{u}, \mathbf{y} \sim \text{Gamma}(\alpha_e + N/2, \gamma_e + \frac{1}{2} \|\mathbf{y} - \mathbf{X}\beta - \mathbf{Z}\mathbf{u}\|^2)$$

Update τ_u :

$$\tau_u \mid \mathbf{u} \sim \text{Gamma}(\alpha_u + q/2, \gamma_u + \frac{1}{2} \mathbf{u}^\top \mathbf{K}^{-1} \mathbf{u})$$

Returns:

- Matrix of post-burn-in samples $(n_{\text{iter}} - n_{\text{burnin}}) \times d$
- M1: Columns $[\beta_0, \dots, \beta_{p-1}, \tau_e]$
- M3: Columns $[\beta_0, \dots, \beta_{p-1}, u_1, \dots, u_q, \tau_e, \tau_u]$

0.5 plot_vb_posteriors()

Purpose: Creates 8-panel comparison plot showing VB, Exact, Laplace, and Gibbs posteriors.

Arguments:

- `mu_beta`: VB posterior mean for (β, \mathbf{u})
- `Sigma_beta`: VB posterior covariance
- `mu_beta_exact`: Exact posterior mean for β
- `Sigma_beta_exact`: Exact posterior covariance
- `laplace_result`: Output from `run_laplace()`
- `gibbs_samples`: Output from `run_gibbs_sampler()`
- `p, q`: Dimensions
- `beta_true, u_true`: True parameter values
- `tau_e_true, tau_u_true`: True precision values
- `E_tau_e, E_tau_u`: VB expectations
- `a_e_new, b_e_new`: VB Gamma parameters for τ_e
- `a_u_new, b_u_new`: VB Gamma parameters for τ_u
- `gibbs_tau_e, gibbs_tau_u`: Gibbs samples for precisions
- `run_gibbs`: Logical, whether Gibbs was run
- `model_type`: "M1" or "M3"

Calculations:

For each parameter $\theta_i \in \{\beta_0, \dots, \beta_{p-1}, u_1, u_2, \tau_e, \tau_u\}$:

1. Construct marginal densities:

- **VB**: $q(\theta_i) = N(\mu_i, \Sigma_{ii})$ for β, \mathbf{u} ; Gamma for τ_e, τ_u
- **Exact**: $p(\theta_i | \mathbf{y})$ from analytical posterior
- **Laplace**: $q(\theta_i) = N(\hat{\theta}_i, H_{ii}^{-1})$
- **Gibbs**: Kernel density estimate from MCMC samples

2. Compute SD ratios:

$$\frac{SD_{VB}}{SD_{Exact}}, \quad \frac{SD_{VB}}{SD_{Laplace}}, \quad \frac{SD_{VB}}{SD_{Gibbs}}$$

3. Overlay densities with colour-coded lines and true value as vertical line

Returns:

- ggplot2 patchwork object with 8 panels (M3) or 5 panels (M1)
- Each panel shows 4 density overlays with SD ratio annotations

0.6 run_vb_algorithm()

Purpose: Mean-field variational Bayes with coordinate ascent for hierarchical models.

Arguments:

- **X, Z, y:** Data matrices
- **K:** Covariance structure for random effects (typically \mathbf{I}_q)
- **p, q, n:** Dimensions
- **alpha_e, gamma_e, alpha_u, gamma_u:** Prior hyperparameters
- **model_type:** "M1" or "M3"
- **max_iter:** Maximum iterations (default: 100)
- **tol:** Convergence tolerance (default: 10^{-5})

Calculations:

Mean-field factorisation:

$$q(\boldsymbol{\beta}, \mathbf{u}, \tau_e, \tau_u) = q(\boldsymbol{\beta}, \mathbf{u}) q(\tau_e) q(\tau_u)$$

Coordinate ascent updates:

1. Update $q(\boldsymbol{\beta}, \mathbf{u})$:

$$\begin{aligned} q(\boldsymbol{\beta}, \mathbf{u}) &\sim N(\boldsymbol{\mu}_{\beta u}, \boldsymbol{\Sigma}_{\beta u}) \\ \boldsymbol{\Sigma}_{\beta u} &= \left(E[\tau_e] [\mathbf{X}|\mathbf{Z}]^\top [\mathbf{X}|\mathbf{Z}] + E[\tau_u] \begin{bmatrix} \mathbf{0} & \mathbf{0} \\ \mathbf{0} & \mathbf{K}^{-1} \end{bmatrix} + \mathbf{I}_{p+q}/100 \right)^{-1} \\ \boldsymbol{\mu}_{\beta u} &= E[\tau_e] \boldsymbol{\Sigma}_{\beta u} [\mathbf{X}|\mathbf{Z}]^\top \mathbf{y} \end{aligned}$$

Where **mu_betau** stores $\boldsymbol{\mu}_{\beta u}$, **Sigma_betau** stores $\boldsymbol{\Sigma}_{\beta u}$, and **E_tau_e**, **E_tau_u** store the expectations.

2. Update $q(\tau_e)$:

$$\begin{aligned} q(\tau_e) &\sim \text{Gamma}(a_e^{\text{new}}, b_e^{\text{new}}) \\ a_e^{\text{new}} &= \alpha_e + N/2 \\ b_e^{\text{new}} &= \gamma_e + \frac{1}{2} \left[\|\mathbf{y}\|^2 - 2\mathbf{y}^\top [\mathbf{X}|\mathbf{Z}] \boldsymbol{\mu}_{\beta u} + \text{tr}([\mathbf{X}|\mathbf{Z}]^\top [\mathbf{X}|\mathbf{Z}] (\boldsymbol{\Sigma}_{\beta u} + \boldsymbol{\mu}_{\beta u} \boldsymbol{\mu}_{\beta u}^\top)) \right] \end{aligned}$$

Where **a_e_new** stores a_e^{new} , **b_e_new** stores b_e^{new} , and **E_tau_e** = $a_e^{\text{new}}/b_e^{\text{new}}$.

3. Update $q(\tau_u)$:

$$\begin{aligned} q(\tau_u) &\sim \text{Gamma}(a_u^{\text{new}}, b_u^{\text{new}}) \\ a_u^{\text{new}} &= \alpha_u + q/2 \\ b_u^{\text{new}} &= \gamma_u + \frac{1}{2} \left[\underbrace{\boldsymbol{\mu}_u^\top \mathbf{K}^{-1} \boldsymbol{\mu}_u}_{\text{quad_form}} + \underbrace{\text{tr}(\mathbf{K}^{-1} \boldsymbol{\Sigma}_{uu})}_{\text{trace_u}} \right] \end{aligned}$$

Where **quad_form** = $\boldsymbol{\mu}_u^\top \mathbf{K}^{-1} \boldsymbol{\mu}_u$ measures the squared magnitude of $\boldsymbol{\mu}_u$ (the quadratic form), and **trace_u** = $\text{tr}(\mathbf{K}^{-1} \boldsymbol{\Sigma}_{uu})$ accounts for uncertainty in $\boldsymbol{\Sigma}_{uu}$ (the trace term).

The R code stores: $\mathbf{a_u_new} = a_u^{\text{new}}$, $\mathbf{b_u_new} = b_u^{\text{new}}$, $\mathbf{E_tau_u} = a_u^{\text{new}}/b_u^{\text{new}}$.

Extracting random effects: $\boldsymbol{\mu}_u = \boldsymbol{\mu}_{\beta u}[(p+1):(p+q)]$ and $\boldsymbol{\Sigma}_{uu} = \boldsymbol{\Sigma}_{\beta u}[(p+1):(p+q), (p+1):(p+q)]$

ELBO:

$$\mathcal{L} = E_q[\log p(\mathbf{y}, \boldsymbol{\beta}, \mathbf{u}, \tau_e, \tau_u)] - E_q[\log q(\boldsymbol{\beta}, \mathbf{u}, \tau_e, \tau_u)]$$

Iterate until $|\mathcal{L}^{(t)} - \mathcal{L}^{(t-1)}| < \text{tol}$ or `max_iter` reached.

Returns:

- `mu_betau`: Posterior mean $\boldsymbol{\mu}_{\beta u}$
- `Sigma_betau`: Posterior covariance $\boldsymbol{\Sigma}_{\beta u}$
- `E_tau_e`, `E_tau_u`: a/b from Gamma posteriors
- `a_e_new`, `b_e_new`, `a_u_new`, `b_u_new`: Gamma parameters
- `elbo_history`: ELBO trajectory
- `E_tau_e_history`, `E_tau_u_history`: Parameter trajectories

0.7 plot_convergence()

Purpose: Diagnostic plots showing VB parameter convergence over iterations.

Arguments:

- `results`: Output from `run_vb_algorithm()`
- `scenario_name`: String label for plot title
- `tau_e_true, tau_u_true`: True parameter values
- `model_type`: "M1" or "M3"

Calculations:

Extract iteration histories:

- $E[\tau_e]^{(t)}$ for $t = 1, \dots, T$
- $E[\tau_u]^{(t)}$ for $t = 1, \dots, T$ (M3 only)

Create line plots with true values as horizontal reference lines.

Returns:

- ggplot2 patchwork object:
 - M1: Single plot for $E[\tau_e]$
 - M3: Two-panel plot for $E[\tau_e]$ and $E[\tau_u]$

0.8 plot_elbo()

Purpose: Plot Evidence Lower BOund (ELBO) trajectory for convergence diagnostics.

Arguments:

- **results:** Output from `run_vb_algorithm()`
- **scenario_name:** String label for plot title

Calculations:

Extract ELBO history: $\mathcal{L}^{(1)}, \mathcal{L}^{(2)}, \dots, \mathcal{L}^{(T)}$

The ELBO should be:

- **Monotonically increasing:** Each coordinate ascent update improves (or maintains) the ELBO
- **Flattening:** Convergence indicated when slope approaches zero

Returns:

- ggplot2 line plot showing ELBO vs iteration number
- Helps identify: premature convergence, oscillations, or divergence