

Homework_1

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Math 742

Casella / Berger

1.2d : Prove $A \cup B = A \cup (B \cap A^c)$

I'll start with the Right-Hand Side and simplify it using the distributive law, which is stated in theorem 1.1.4.

In set theory proofs, equality of sets is proven by double containment

$$\textcircled{i} \quad A \cup B \subseteq A \cup (B \cap A^c) \quad \wedge \quad A \cup (B \cap A^c) \subseteq A \cup B \quad \textcircled{ii}$$

\textcircled{i} Taking an arbitrary element $x \in A \cup B$. By definition of union:

$$x \in A \cup B \rightarrow x \in A \text{ or } x \in B$$

Now, let's consider the two cases.

$$\text{case 1: } x \in A \rightarrow x \in A \subseteq A \cup (B \cap A^c)$$

$$\text{case 2: } x \in B \wedge x \notin A \quad \begin{cases} x \in B \\ x \in A^c \end{cases}$$

$$\text{So, } x \in B \cap A^c \rightarrow x \in A \cup (B \cap A^c)$$

In both cases, $x \in A \cup (B \cap A^c)$ thus: $A \cup B \subseteq (B \cap A^c)$

\textcircled{ii} Taking an arbitrary $x \in A \cup (B \cap A^c) \rightarrow x \in A \vee x \in (B \cap A^c)$

$$\text{case 1: } x \in A \text{ then } x \in A \cup B \quad \text{case 2: } x \in B \cap A^c \text{ then } x \in B \rightarrow x \in A \cup B$$

Again in both cases $x \in A \cup B$ Thus $A \cup (B \cap A^c) \subseteq A \cup B$

After proving the above result, prove: $P(A \cup B) = P(A) + P(B) - P(A \cap B)$

from above $A \cup B = A \cup (B \cap A^c) \rightarrow A \text{ and } B \cap A^c \text{ are disjoint } A \cap (B \cap A^c) = \emptyset$

why? $A \cap (B \cap A^c) \rightarrow$ using associative of intersection $A \cap (B \cap A^c) = (A \cap B) \cap A^c$. so any element x in this set must satisfy all 3 conditions: $x \in A$, $x \in B$, $x \in A^c$. Then by definition $(A^c) \cap A^c = \emptyset$

Then $(A \cap B) \cap A^c \subseteq A \cap A^c = \emptyset \rightarrow A \cap (B \cap A^c) = \emptyset$

from above, if $C \cap D = \emptyset$ then $P(C \cup D) = P(C) + P(D) \rightarrow P(A \cup B) = P(A) + P(B \cap A^c)$

Rewriting $P(B \cap A^c)$, $B = (B \cap A) \cup (B \cap A^c)$ (decomposition)

$$P(B) = P(B \cap A) + P(B \cap A^c) \rightarrow P(B \cap A^c) = P(B) - P(B \cap A)$$

$$\text{Then } P(A \cup B) = P(A) + P(B \cap A^c)$$

$$= P(A) + P(B) - P(A \cap B)$$

1.3c De Morgan's Laws. Prove this without Venn diagrams.

$$\textcircled{i} \quad (A \cup B)^c = \underline{A^c \cap B^c} \quad \wedge \quad \textcircled{ii} \quad (A \cap B)^c = \underline{A^c \cup B^c}$$

\textcircled{i} If an $x \in (A \cup B)^c \leftrightarrow x \notin A \vee x \notin B \leftrightarrow x \in A^c \wedge x \in B^c \leftrightarrow x \in A^c \cup B^c$

\textcircled{ii} If an $x \in (A \cap B)^c \leftrightarrow x \notin A \cap B \leftrightarrow x \notin A \wedge x \notin B \leftrightarrow x \in A^c \vee x \in B^c \leftrightarrow x \in A^c \cup B^c$

1.4/10: Prove that $F_X(x) = \frac{1}{1+e^{-x}}$, $x \in (-\infty, \infty)$ is a CDF.

Our $F_X(x)$ satisfies the conditions of theorem 1.5.3

Theorem 1.5.3 The function $F(x)$ is a CDF \Leftrightarrow The following 3 conditions hold:

a. $\lim_{x \rightarrow -\infty} F(x) = 0 \wedge \lim_{x \rightarrow \infty} F(x) = 1$ (Monotonicity)

b. $F(x)$ is a NONDECREASING function of x .

c. $F(x)$ is RIGHT continuos; that is, for every number x_0 , $\lim_{x \downarrow x_0} F(x) = F(x_0)$

outline of proof: To prove necessity, the three properties can be verified by writting F in terms of the probability function.

If $F_X(\cdot)$ is a CDF, $F_X(x) = P(X \leq x)$. Hence $\lim_{x \rightarrow -\infty} P(X \leq x) = 0 \wedge \lim_{x \rightarrow \infty} P(X \leq x) = 1$

$F(x)$ is nondecreasing since the set $\{x : X \leq x\}$ is nondecreasing in x .

Lastly, as $x \downarrow x_0$, $P(X \leq x) \rightarrow P(X \leq x_0)$. Then $F_X(\cdot)$ is RIGHT continuos.

4.4 A PDF is defined by,

$$f(x, y) = \begin{cases} c(x+2y) & : 0 < y < 1 \wedge 0 < x < 2 \\ 0 & : \text{Otherwise} \end{cases}$$

a) find the value of c .

$$\int_0^2 \int_0^1 c(x+2y) dy dx = 1 \quad (\text{the total probability})$$

$$(y) \rightarrow \int_0^1 (x+2y) dy = (xy + y^2) \Big|_0^1 = (x+1) - (0) = x+1$$

$$(x) \rightarrow c \int_0^2 (x+1) dx = c \left(\frac{x^2}{2} + x \right) \Big|_0^2 = c((2+2) - (0)) = 4c$$

$$4c = 1 \rightarrow c = 1/4$$

b) Find the marginal distribution of X .

To find $f_X(x)$, I should integrate the joint PDF over the range of y

$$f_X(x) = \begin{cases} \int_0^1 1/4 (x+2y) dy, \text{ we did it previously} & = \frac{1}{4}(x+1), \quad 0 < x < 2 \\ 0 & \text{otherwise} \end{cases}$$

c) Find the joint CDF of X and Y . $= F_{XY}(x, y) = P(X \leq x, Y \leq y)$

$$P(X \leq x, Y \leq y) = \int_{-\infty}^x \int_0^y f(v, u) dv du = \int_0^y \int_0^x \frac{1}{4}(u+2v) dv du$$

$$\text{Inner integral (w.r.t. } v) \quad \frac{1}{4} [uv + v^2]_0^y = \frac{1}{4}(uy + y^2)$$

$$\text{Outer integral (w.r.t. } u) \quad \int_0^x \frac{1}{4}(uy + y^2) du = \frac{1}{4} [\frac{1}{2}u^2y + uy^2]_0^x$$

$$F_{XY}(x, y) = \frac{x^2y}{8} + \frac{xy^2}{4} \quad \text{for } 0 < x < 2 \wedge 0 < y$$

$$f_{XY}(x, y) = \int_{-\infty}^x \int_0^y f(u, v) dv du = \int_0^y \int_0^x \frac{1}{4}(u+2v) dv du = \frac{x^2}{8} + \frac{x}{4}$$

Then, the full definition of F_{XY} is

$$F_{XY}(x, y) = \begin{cases} 0 & x \leq 0 \vee y \leq 0 \\ \frac{x^2y}{8} + \frac{xy^2}{4} & 0 < x < 2 \wedge 0 < y < 1 \\ \frac{y}{2} + \frac{y^2}{2} & 2 \leq x \wedge 0 < y < 1 \\ \frac{x^2}{8} + \frac{x}{4} & 0 < x < 2 \wedge 1 \leq y \\ 1 & 2 \leq x \wedge 1 \leq y \end{cases}$$

d) Find the PDF of the random variable $Z = g(X+1)^2$

Apply Theorem 2.1.5 (Monotonicity) computing derivative:

$$g'(x) = -\frac{18}{(x+1)^3} < 0 \quad \text{for } x > 0. \quad \text{so } g(x) \text{ is strictly monotone decreasing on } 0 < x < 2$$

- find the support of Z :

$$\text{Evaluate endpoints at } x=0: \quad Z = \frac{g}{(1)^2} = g \quad \text{at } x=2: \quad Z = \frac{g}{(3)^2} = 1$$

since g is decreasing

$$1 < Z < 9$$

- Invert the transformation

Start from

$$Z = \frac{g}{(x+1)^2} \quad \text{solve for } x: \quad (x+1)^2 = \frac{g}{Z} \rightarrow x+1 = \frac{\sqrt{g}}{\sqrt{Z}}$$

$$x = \frac{\sqrt{g}}{\sqrt{Z}} - 1 \quad (\text{only positive root is valid since } x+1 > 0)$$

$$\sqrt{z}$$

- Compute the Jacobian. $\frac{\partial x}{\partial z} = -\frac{3}{2} z^{-3/2} \rightarrow \frac{\partial x}{\partial z} = \frac{3}{2 z^{3/2}}$

- For a monotone transformation:

$$f_z(z) = f_x(x(z)) \left| \frac{\partial x}{\partial z} \right|$$

- substitute

$$\text{compute } f_x(x(z)) \quad x+1 = \frac{3}{\sqrt{z}}$$

$$f_x(x) = \frac{1}{4}(x+1) = \frac{3}{4\sqrt{z}}$$

$$\text{- Multiply by Jacobian} \quad f_z(z) = \frac{3}{4\sqrt{z}} \cdot \frac{3}{2z^{3/2}} = \frac{9}{8z^2} \rightarrow f_z(z) = \begin{cases} \frac{9}{8z^2}, & 1 < z < 9 \\ 0 & \text{otherwise} \end{cases}$$

4.10 The random pair (X, Y) has the distribution

	$X=1$	$X=2$	$X=3$
$Y=2$	$\frac{1}{12}$	$\frac{1}{6}$	$\frac{1}{12}$
$Y=3$	$\frac{1}{6}$	0	$\frac{1}{6}$
$Y=4$	0	$\frac{1}{3}$	0

a) Show that X and Y are dependent. Two discrete random variables X, Y are independent if: $P(X=x, Y=y) = P(X=x)P(Y=y) \forall x, y$

One counterexample

$$P(X=2, Y=3) = 0$$

But if X and Y were independent:

$$P(X=2)P(Y=3) = \frac{1}{2} \cdot \frac{1}{3} = \frac{1}{6}$$

$$0 \neq \frac{1}{6} \rightarrow X \text{ and } Y \text{ are dependent}$$

b) Give a probability table for random variables U and V that have the same marginals as X as Y but are independent.

The distribution that satisfies $P(U=x, V=y) = P(U=x)P(V=y)$ where $U \sim X \wedge V \sim Y$ is

	$U=1$	$U=2$	$U=3$
$V=2$	$1/12$	$1/6$	$1/12$
$V=3$	$1/12$	$1/6$	$1/12$
$V=4$	$1/12$	$1/6$	$1/12$

Wasserman

3.10 $X \sim N(0,1)$, $Y \sim e^X$. Determine $E(Y)$ and $V(Y)$

$$E(Y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\sigma} e^x e^{-x^2/2} dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\sigma} e^{\frac{x(2-x)}{2}} dx \quad \text{let } u=x-1$$

$$\rightarrow E(Y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\sigma} e^{\frac{1-u}{2}} du = e^{1/2} \int_{-\infty}^{\sigma} \frac{e^{-u^2/2}}{\sqrt{2\pi}} du = e^{1/2}$$

$$E(Y^2) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\sigma} e^{2x - \frac{x^2}{2}} dx \quad \text{lets } u=x-2$$

$$\rightarrow E(Y^2) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\sigma} e^2 e^{-\frac{u^2}{2}} du = e^2$$

$$V(Y) = E(Y^2) - E(Y)^2 = e^2 - (e^{1/2})^2 = e^2 - e$$

NO TIME TO WORK ON IT

3.19

Extra Problem:

Let X have a gamma distribution with parameters α, β .

Prove that $E(X) = \alpha\beta$ and $V(X) = \alpha\beta^2$

The Gamma PDF is defined as:

- $\alpha > 0$

- $\beta > 0$

$$f_X(x) = \frac{\beta^\alpha}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} \quad \Gamma(\alpha) = \int_0^\infty t^{\alpha-1} e^{-t} dt$$

Show that: $E(X) = \alpha\beta$ and $V(X) = \alpha\beta^2$

