Lecture 20: Multiple Hypothesis Testing

GENOME 560

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Outline

Why multiple hypothesis testing matters?



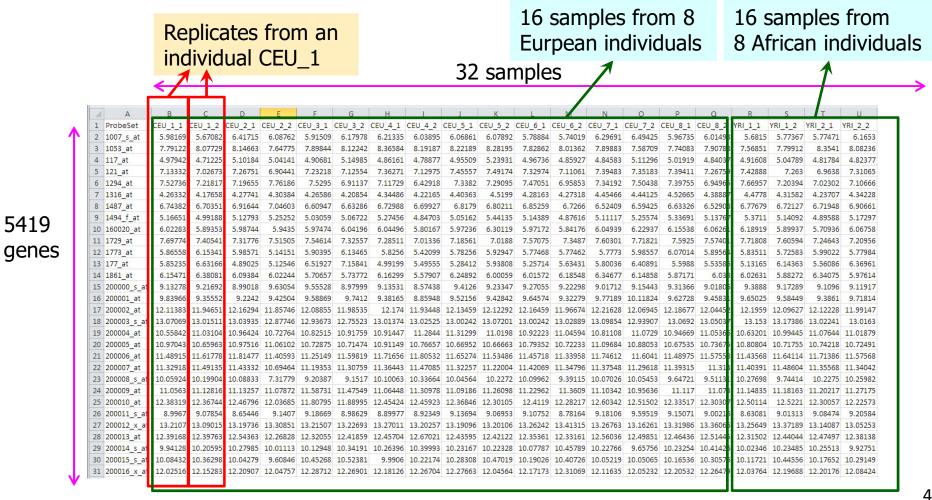
- A motivating example
- R-session
- Basic concepts on multiple hypothesis testing
 - Type I error and Type II error
 - Define the multiple testing problem and related concepts
- Methods for multiple hypothesis testing correction
 - 2 methods that control the family-wise error rate (FWER)
 - 1 method that controls the false discovery rate (FDR)
 - Our goal is to understand what type of error is controlled in each method, such that we know which one to use in our research.

Motivating example

- Storey and Akey (2007). Gene-expression variation
 within and among human populations. AJHG
 - Understand patterns of gene-expression variation within and among human populations.
 - Gain insights into the molecular basis of phenotypic diversity.
- The authors measured expression levels of 5194 genes in 16 human individuals
 - 8 European + 8 African individuals
- Apply t-test to each gene. This results in 5194 p-values!

Expression Data

Here is the expression data from Storey & Akey (2007):



Questions of interest

How many genes show a significant difference in expression levels between European and African individuals?

Which genes show a significant difference?

How would you answer these questions?

Let's try...

Load the data

```
a <- read.table(header = T,
file="http://www.cs.washington.edu/homes/suinlee/genome5
60/RMA_Filtered.txt")
b <- a[,2:33]</pre>
```

- Define a function of performing the t-test fun <- function(d){return(t.test(d[1:16],d[17:32])\$p.value)}</p>
- Obtain the p-values

```
p <- apply(b, 1, fun)
```

See the distribution

```
hist(p, breaks=20)
```

We got many p-values!

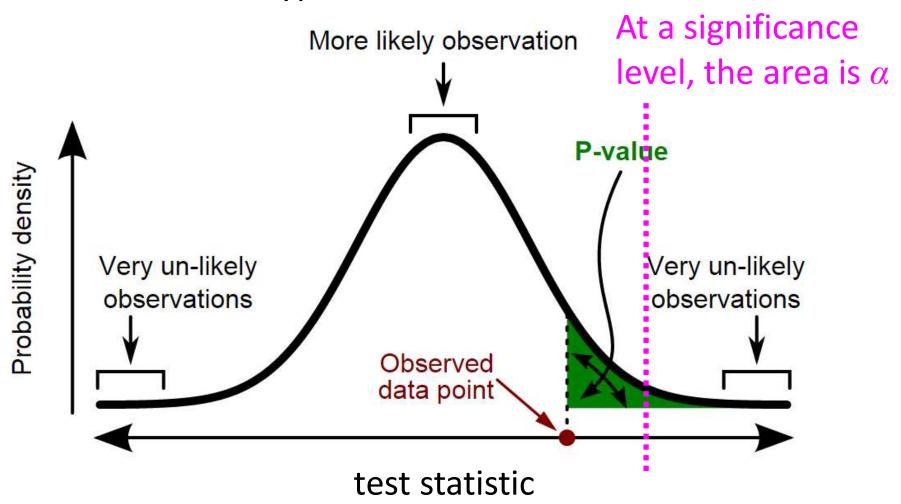
- Which genes show significant difference in expression levels between European and African individuals?
- How many p-values < 0.05?</p>

```
alpha = 0.05
tabulate(as.numeric(p < alpha))</pre>
```

- Would a standard p-value cutoff α = 0.05 (or 0.01) be useful when there are many hypotheses?
- Definition of p-value:
 - The estimated percentage of observations more extreme than the one observed under the assumption that the null hypothesis is true

Review: p-value

Distributin of the test statistic (e.g., f-ratio in ANOVA) when the null hypothesis is true:



Why Multiple Testing Matters I

- We have 5194 hypothesis tests in this problem
 - A typical microarray experiment might result in performing 10,000 separate hypothesis tests.
 - Genomics: Lots of data, Lots of hypothesis tests
- We would expect ~260 (5194 x 0.05) genes to be deemed "significant" by chance.
- A standard p-value cutoff α = 0.05 needs to be adjusted.

Why Multiple Testing Matters II

- In general, if we perform m hypothesis tests, what is the probability of at least 1 false positive?
 - Assume that all the null hypotheses are true

Rejecting the null hypothesis when that hypothesis is true.

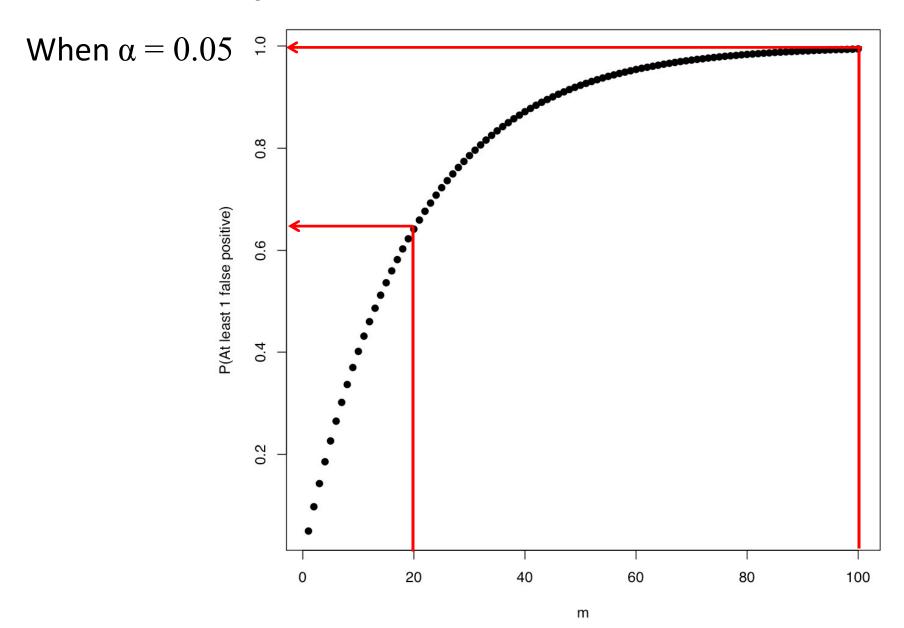
P(Making an error) = α

P(Not making an error) = $1 - \alpha$

P(Not making an error in *m* tests) = $(1 - \alpha)^{m}$

P(Making at least 1 error in m tests) = 1 - $(1 - \alpha)^m$

Probability of At Least 1 False Positive



"Correcting" for Multiple Testing

- We need to adjust p-values for the number of hypothesis tests performed"
- Very active area of statistics many different methods have been described
- Although these varied approaches have the same goal, they go about it in fundamentally different ways
 - Our goal is to understand what errors each method controls, such that we know which method to use in research.
- Many of them control the Type I error

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- Basic concepts on multiple hypothesis testing



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Type I and II Errors

Actual Situation "Truth"

Decision

Do Not Reject H_o

Reject H₀

H_0	Tr	ue

Correct Decision (True Negative) $1-\alpha$

Incorrect Decision (False Positive)

Type I Error α

H₀ False

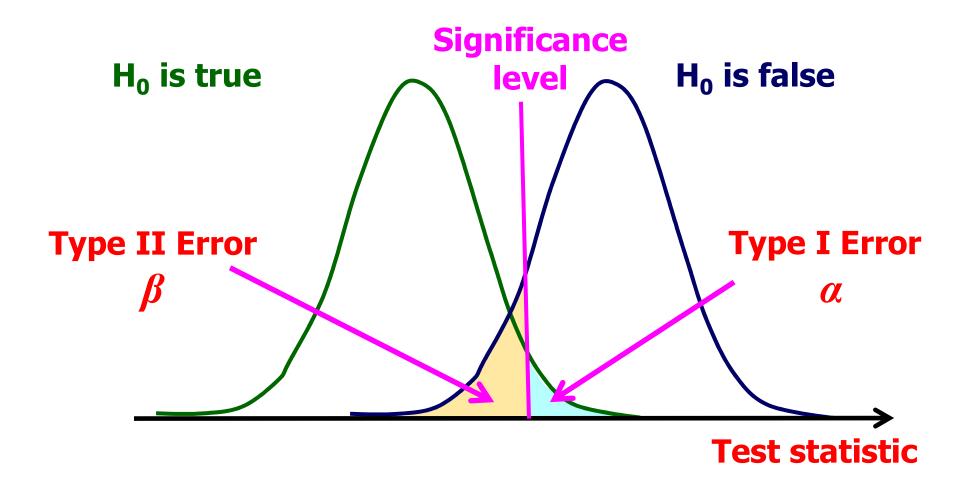
Incorrect Decision (False Negative)
Type II Error β

Correct Decision (True Positive)
1-β

$$\alpha$$
 = P(Type I Error) β = P(Type II Error)
Power = 1 - β

Type I and Type II Errors

Consider the distribution of your test statistic



Counting Errors

- Assume that we are testing m hypotheses: H^1 , ..., H^m
 - m_0 = # of **true** null hypotheses
 - \blacksquare R = # of rejected null hypotheses

	Null True	Alternative True	Total
Not Called Significant	U	7	m-R
Called Significant	V	5	R
	m_0	<i>m</i> - <i>m</i> ₀	m

V = # Type I errors [false positives]

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Family-Wise Error Rate (FWER)

- Assume that we are testing m hypotheses: $H^1, ..., H^m$
- Family-Wise Error Rate (FWER)
 - Here, the term "family" refers to the collection of hypotheses H^1 , ..., H^m
 - The probability of making one type I error among all the hypotheses, P(V≥1)
- Two general types of FWER corrections:
 - Single step: equivalent adjustments made to each p-value
 - Sequential: adaptive adjustment made to each p-value

Single Step Approach: Bonferroni

- By assuring FWER $\leq \alpha$, the probability of making even one Type I error in the family is controlled at level α .
- Adjust p-values as following:

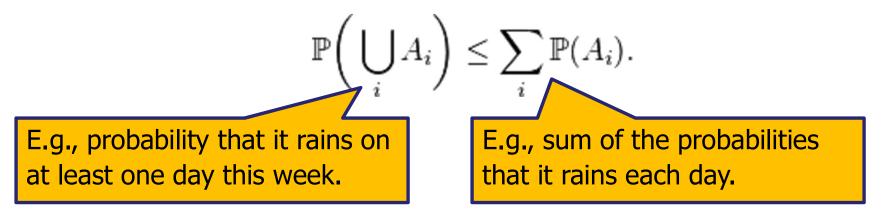
$$\widetilde{p}_j = \min[mp_j, 1]$$

- Rejects any hypothesis with p-value $\leq \alpha/m$.
- Example
 - Say that we want the probability of making at least one Type I error to be <0.05 when we perform 10,000 hypothesis tests.
 - Then, we need a p-value of $0.05/10,000 = 5 \times 10^{-6}$ to declare significance

- Notations
 - H^1 , ..., H^m : a family of hypotheses
 - $p_1, ..., p_m$: the corresponding p-values.
 - I_0 : set of (unknown)true null hypotheses, having m_0 members
- The FWER is the probability of rejecting at least one of the members in I_0
- The Bonferroni correction states that rejecting all $p_i \le \alpha/m$ will control the FWER $\le \alpha$

The proof is based on the Boole's inequality

- In probability theory, Boole's inequality says that for any finite set of events, the probability that at least one of the events happens is no greater than the sum of the probabilities of the individual events.
- Boole's inequality is named after George Boole.
- Formally, for a set of events, A_1 , A_2 ,..., we have:



- Notations
 - H^1 , ..., H^m : a family of hypotheses
 - $p_1, ..., p_m$: the corresponding p-values.
 - I_0 : set of the true null hypotheses, having m_0 members
- The FWER is bounded above by α :

$$FWER = Pr\left\{\bigcup_{I_o} \left(p_i \leq \frac{\alpha}{m}\right)\right\}$$
 The probability that at least one of the true null hypotheses has a p-value $\leq \alpha/m$

- Notations
 - $H^1, ..., H^m$: a family of hypotheses
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 - I_0 : set of the true null hypotheses, having m_0 members
- The FWER is bounded above by α :

$$FWER = Pr\left\{\bigcup_{I_o} \left(p_i \le \frac{\alpha}{m}\right)\right\} \le \sum_{I_o} \left\{Pr\left(p_i \le \frac{\alpha}{m}\right)\right\}$$

The probability that at least one of the true null hypotheses has a p-value $\leq \alpha/m$

Sum of the probabilities that each of the true null hypotheses has a p-value $\leq \alpha/m$

- Notations
 - H^1 , ..., H^m : a family of hypotheses
 - $p_1, ..., p_m$: the corresponding p-values.
 - I_0 : set of the true null hypotheses, having m_0 members
- The FWER is bounded above by α :

$$FWER = Pr\left\{\bigcup_{I_o} \left(p_i \le \frac{\alpha}{m}\right)\right\} \le \sum_{I_o} \left\{Pr\left(p_i \le \frac{\alpha}{m}\right)\right\} \le m_0 \frac{\alpha}{m}$$

The probability that a true null hypothesis test gets a p-value $\leq \alpha/m = \alpha/m$ (see the definition of p-value on slide 8)

- Notations
 - H^1 , ..., H^m : a family of hypotheses
 - $p_1, ..., p_m$: the corresponding p-values.
 - I_0 : set of the true null hypotheses, having m_0 members
- The FWER is bounded above by α :

$$FWER = Pr\left\{\bigcup_{I_o} \left(p_i \leq \frac{\alpha}{m}\right)\right\} \leq \sum_{I_o} \left\{Pr\left(p_i \leq \frac{\alpha}{m}\right)\right\} \leq m_0 \frac{\alpha}{m} \leq m \frac{\alpha}{m} = \alpha$$

An alternative method

- Drawback of the Boferroni correction
 - It often leaves very few hypotheses that are deemed significant
- "Holm method" a.k.a "Holm-Bonferroni method"
- It is known to be more powerful than the Bonferroni method
 - More "lenient" correction than Bonferroni method

Basic idea:

- The Bonferroni correction p-value cut-off is α/m .
- This could have been α/m_0 , where m_0 is the number of true null hypotheses.
- We do not know m_0 , but could estimate it.

FWER: Sequential Adjustments

- Simplest sequential method is Holm's Method
 - Order the unadjusted p-values such that $p_1 \le p_2 \le ... \le p_m$
 - For control of the FWER at level α , the step-down Holm adjusted p-values are

$$\widetilde{p}_j = \min[(m-j+1) \mid p_j,1]$$

- The point here is that we do not multiply every p_i by the same factor m
- For example, when m = 10000:

$$\widetilde{p}_1 = 10000 \cdot p_1, \quad \widetilde{p}_2 = 9999 \cdot p_2, \dots, \widetilde{p}_m = 1 \cdot p_m$$

Philosophical Objections to Bonferroni Corrections

- "Bonferroni adjustments are, at best, unnecessary and, at worst, deleterious to sound statistical inference" Perneger (1998)
- Counter-intuitive: interpretation of finding depends on the number of other tests performed (shared by all methods)
- The general null hypothesis (that all the null hypotheses are true) is rarely of interest
- High probability of Type II errors, i.e., of not rejecting the general null hypothesis when important effects exist

Who Cares About Not Making ANY Type I Errors?

- FWER is appropriate when you want to guard against ANY false positives
- However, in many cases (particularly in genomics) we can live with a certain number of false positives
- In these cases, the more relevant quantity to control is the false discovery rate (FDR)

```
# falsely rejected # rejected in total
```

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False Discovery Rate

	Null True	Alternative True	Total
Not Called Significant	U	7	m-R
Called Significant	V	<i>S</i>	R
	m_o	m - m ₀	m

- V = # Type I errors [false positives]
- False discovery rate (FDR) is designed to control the proportion of false positives among the set of rejected hypotheses (R) -- V/R

Benjamini and Hochberg FDR

- To control FDR at level δ:
 - 1. Order the unadjusted p-values: $p_1 \le p_2 \le ... \le p_m$
 - 2. Then find the test with the highest rank, j, for which the p-value, p_i , is less than or equal to (j/m) x δ
 - 3. Declare the tests of rank 1, 2, ..., j as significant

$$p(j) \le \delta \frac{j}{m}$$

Adjust p-values as following (now called q-values):

$$\widetilde{p}_j = \min[\frac{m}{j}p_j,1]$$

B&H FDR Example

• Controlling the FDR at $\delta = 0.05$

Rank (j)	P-value	(j/m) x δ	Reject H ₀ ?
1	0.0008	0.005	1
2	0.009	0.010	1
3	0.165	0.015	0
4	0.205	0.020	0
5	0.396	0.025	0
6	0.450	0.030	0
7	0.641	0.035	0
8	0.781	0.040	0
9	0.900	0.045	0
10	0.993	0.050	0

Summary

- Why care about multiple hypoithesis testing?
- Controling the family-wise error rate (FWER)
 - Bonferroni correction
 - Holm method
- Controling the false discovery rate (FDR)
 - B-F FDR correction

Let's quickly do the following

Load the data

```
a <- read.table(header = T,
file="http://www.cs.washington.edu/homes/suinlee/genome5
60/RMA_Filtered.txt")
b <- a[,2:33]</pre>
```

Define a function of performing the t-test

```
fun \leftarrow function(d) \{ return(t.test(d[1:16],d[17:32]) \} p.value) \}
```

Obtain the p-values

p <- apply(b, 1, fun)

See the distribution

hist(p, breaks=20)

Check how many are < 0.05</p>

tabulate(as.numeric(p<0.05))

R commands

- tabulate(as.numeric(p<0.05))</p>
 - tabulate & as.numeric
 - Counting how many elements are < 0.05
- as.numeric
 - What is the output of "p<0.05"
 - What is the output of "as.numeric(p < 0.05)"?
 - as.numeric (which is identical to as.double) coerces to the class
- tabulate
 - tabulate takes the integer-valued vector bin and counts the number of times each integer occurs in it.
- Alternatively, we can use a different command "???"

Let's apply the correction methods

What methods are available?

```
p.adjust.methods
```

Let's apply each of them

```
padj1 <- p.adjust(p, "bonferroni")
padj2 <- p.adjust(p, "holm")
padj3 <- p.adjust(p, "BH")</pre>
```

How many of the adjusted p-values are less than 0.05?

```
alpha = 0.05
tabulate(as.numeric(padj1 < alpha) )
tabulate(as.numeric(padj2 < alpha) )
tabulate(as.numeric(padj3 < alpha) )</pre>
```

Let's compare the adjusted p-values!

Plot the histograms of the original and adjusted p-values.

```
par( mfrow = c(1,2) )
hist( p, breaks = 50)
hist( padj1, breaks = 50)
```

Let's plot the histogram of 4 sets of p-values.

```
par( mfrow = c(1,4) )
hist( p, breaks = 50)
hist( padj1, breaks = 50)
hist( padj2, breaks = 50)
hist( padj3, breaks = 50)
```

Which method is the most lenient and which one is the most harsh adjustment?