

## STATISTICAL AVERAGES

(1)

In some cases we cannot obtain the pdf  $\Rightarrow$  So for the description of random variables we use statistical averages or mean values.

### • Average of a Discrete Random Variable

- $\Rightarrow$  Random discrete variable  $X$
- $\Rightarrow$  Possible values of the  $X$ :  $X_1, X_2, \dots, X_m$ .
- $\Rightarrow$  Respective Probabilities:  $P_1, P_2, \dots, P_m$ .

### Statistical Average/Expectation

$$\bar{X} = E[X] = \sum_{j=1}^m X_j * P_j$$

$$X_1 = 100$$

$$X_2 = 30.$$

$$P_1 = 0.8.$$

$$P_2 = 0.2.$$

$$\Rightarrow \bar{X} = \sum_{j=1}^2 X_j * P_j$$

$$= X_1 * P_1 + X_2 * P_2.$$

$$= \underline{100 * 0.8 + 30 * 0.2.}$$

### Proof.

We will use the relative frequency of observations. Let's say that the experiment is repeated a large number of times  $N$ , and  $X_1$  is observed  $n_1$  times,  $X_2$  is observed  $n_2$  times etc. In order to get the mathematical average:

$$\frac{n_1 X_1 + n_2 X_2 + \dots + n_m X_m}{N} = \sum_{j=1}^m X_j \frac{n_j}{N}$$

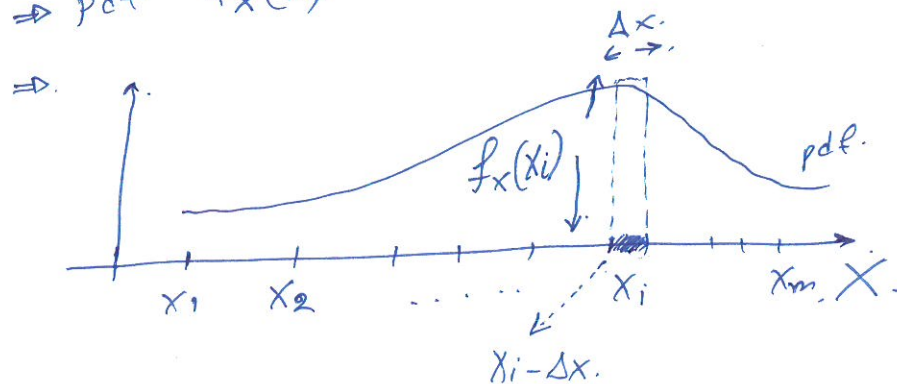
By the relative frequency interpretation when  $N \rightarrow \infty$   
then  $\lim_{N \rightarrow \infty} \frac{n_j}{N} = P_j.$

# • Average of a Continuous Random Variable.

(2)

⇒ Random Continuous Variable  $X$

⇒ pdf  $f_X(x)$



We consider the range of value that  $X$  may take on, say  $x_0$  to  $x_m$ , to be broken up into a large number of small subintervals of length  $\Delta x$ .

Probability that  $X$  lies between  $x_i - \Delta x$  and  $x_i$ :

$$P(x_i - \Delta x \leq X \leq x_i) \approx f_X(x_i) \cdot \Delta x.$$

[ We have approximated  $X$  by a discrete random variable that takes on the values  $x_0, x_1, \dots, x_m$  with probabilities  $f_X(x_0)\Delta x, \dots, f_X(x_m)\Delta x$ .

Expectation

$$E[X] \approx \sum_{i=0}^m x_i f_X(x_i) \Delta x.$$

As  $\Delta x \rightarrow 0$  the  $E[X]$  becomes a better approximation

So as  $\Delta x \rightarrow 0 \Rightarrow \Delta x = dx$

$$\text{So } E[X] = \int_{-\infty}^{+\infty} x f_X(x) dx.$$

Note:  $E[X]$   $\begin{cases} \rightarrow \text{expectation} \\ \rightarrow \text{mean.} \\ \rightarrow \text{first moment} \end{cases}$  of  $X$ .

### \* Average of a Function of a random Variable

(3)

→ We are interested in statistical averages of functions of  $X$ .

$$\rightarrow Y = g(X) \quad [Y = X^2 + 1]$$

$$\text{So: } E[Y] = \int_{-\infty}^{\infty} y f_Y(y) dy$$

where  $f_Y(y)$  is the pdf of  $Y$ .

↓  
We do not have it!

So:

$$Y = E[Y] = E[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx.$$

EXAMPLE.

Suppose that the random variable  $\Theta$  has the pdf:

$$f_{\Theta}(\theta) = \begin{cases} \frac{1}{2\pi} & , \quad |\theta| \leq \pi \\ 0 & , \quad \text{otherwise.} \end{cases}$$

( $\Theta$ : continuous)

We have to compute:

$$E[\Theta^n] = \int_{-\infty}^{+\infty} \theta^n * f_{\Theta}(\theta) d\theta = \int_{-\pi}^{\pi} \theta^n * \frac{1}{2\pi} d\theta$$

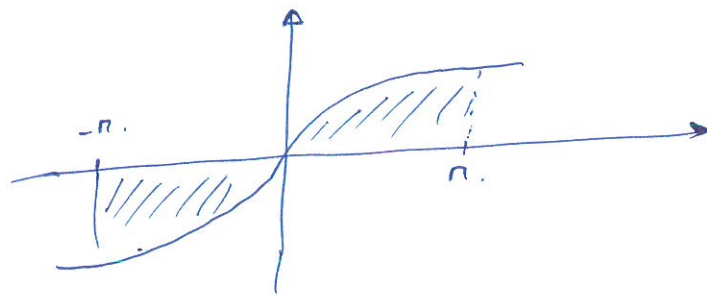
↓  
 $g(\theta) = \theta^n$

$$\Rightarrow E[\Theta^n] = \frac{1}{2\pi} \int_{-\pi}^{\pi} \theta^n d\theta.$$

• Let's suppose that  $n=3$ .

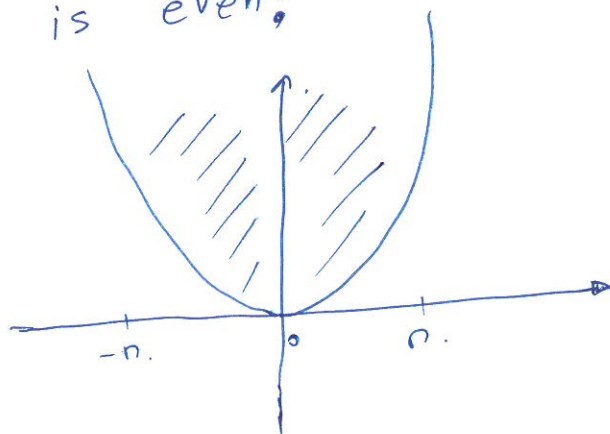
(4)

$$\int_{-n}^n \theta^3 d\theta = \frac{\theta^4}{4} \Big|_{-n}^n = \frac{n^4}{4} - \frac{(-n)^4}{4} = 0.$$



$\Rightarrow$  If  $n$  is odd, then  $E[\Theta^n] = 0$ .

$\Rightarrow$  If  $n$  is even;



$$E[\Theta^n] = 2 * \frac{1}{2\pi} \int_0^n \theta^n d\theta = \frac{1}{\pi} \cdot \frac{\theta^{n+1}}{n+1} \Big|_0^n = \frac{n^{n+1}}{n+1} - 0 = \frac{n^{n+1}}{n+1}.$$

$$\Rightarrow \text{If } n=1 \Rightarrow E[\Theta] = \frac{1}{2\pi} \int_{-n}^n \theta d\theta = \frac{1}{2\pi} \cdot \frac{\theta^2}{2} \Big|_{-n}^n.$$

$$= \frac{1}{2\pi} \cdot \left( \frac{n^2}{2} - \frac{(-n)^2}{2} \right) = 0.$$



## Average of a Function of More than One Random Variables.

(5)

- Function  $g(X, Y)$  of two random variables  $X$  and  $Y$
- The expectation of  $g$  is defined in a manner analogous to the case of a single random variable
- $f_{XY}(x, y)$ : joint pdf of  $X$  and  $Y$

$$\rightarrow E[g(X, Y)] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} g(x, y) f_{XY}(x, y) dx dy.$$

↑  
It can be generalized for more than 2 variables.

⇒ The generalization include also the single random variable case.

Suppose that  $g(X, Y)$  is replaced by  $h(X)$ .

So:

$$\begin{aligned} E[h(X)] &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} h(x) f_{X,Y}(x, y) dx dy \\ &= \int_{-\infty}^{+\infty} h(x) f_X(x) dx. \end{aligned}$$

where:

$$\int_{-\infty}^{+\infty} f_{X,Y}(x, y) dy = f_X(x)$$

## Average of a Function of More than One Random Variable

(5)

→ Function  $g(X, Y)$  of two random variables  $X$  and  $Y$ .

→ The expectation of  $g$  is defined in a manner analogous to the case of a single random variable.

→  $f_{XY}(x, y)$ : joint pdf of  $X$  and  $Y$ .

So, for the expectation we have:

$$E[g(X, Y)] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} g(x, y) f_{XY}(x, y) dx dy$$

↖  
This equation can be generalized for more than 2 random variables.

• The generalization includes also the single random variable case. (Proof).  
Suppose that  $g(X, Y)$  is replaced by  $h(X)$ .

So:

$$E[g(X, Y)] = E[h(X)] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} h(x) f_{XY}(x, y) dx dy.$$

$$= \int_{-\infty}^{+\infty} h(x) f_X(x) dx$$

where:

$$\int_{-\infty}^{+\infty} f_{XY}(x, y) dy = f_X(x).$$

EXAMPLE.

⑥

•  $g(x, y) = XY$

• joint pdf:  $f_{XY}(x, y) = \begin{cases} 2 \cdot e^{-(2x+y)} & x, y \geq 0 \\ 0, & \text{otherwise.} \end{cases}$

$$E[g(x, y)] = E[XY] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f_{XY}(x, y) dx dy$$

$$= \int_0^{+\infty} \int_0^{+\infty} 2xy e^{-(2x+y)} dx dy$$

$$= 2 \int_0^{+\infty} x \cdot e^{-2x} dx \int_0^{+\infty} y \cdot e^{-y} dy.$$

Integration by parts.

$$\int f dg = fg - \int g df.$$

•  $\int_0^{+\infty} x \cdot e^{-2x} dx$

let  $f = x$   
 $df = dx$

and  $dg = e^{-2x}$

and  $g = -\frac{e^{-2x}}{2}$

$$\left[ \sim \left( -\frac{e^{-2x}}{2} \right)' = +\frac{2e^{-2x}}{2} = e^{-2x} \right]$$

$\Rightarrow$

$$\int_0^{+\infty} x \cdot e^{-2x} dx = -x \cdot \left( \frac{e^{-2x}}{2} \right) \Big|_0^{\infty} + \int_0^{\infty} \frac{e^{-2x}}{2} dx \quad (7)$$

$$= \left[ -x \cdot \frac{e^{-2x}}{2} \right]_0^{+\infty} + \frac{1}{2} \left[ \frac{e^{-2x}}{2} \right]_0^{+\infty}$$

$$\bullet \lim_{x \rightarrow \infty} \left[ -x \cdot \frac{e^{-2x}}{2} \right] = 0.$$

$$\bullet \lim_{x \rightarrow +\infty} \left[ \frac{e^{-2x}}{2} \right]_0^{+\infty} = 0.$$

$$\bullet \frac{1}{2} \left[ \frac{e^0}{2} \right] = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}.$$

$$\int_0^{+\infty} x \cdot e^{-2x} dx = \frac{1}{4}.$$

$$\text{Similarly, } \int_0^{+\infty} y \cdot e^{-y} dy = 1.$$

$$E[g(x, y)] = 2 * \frac{1}{4} * 1 = \frac{1}{2}.$$