

# Gaussian processes for classification

Manel Martínez-Ramón

ECE, UNM

October 2019

- In classification, we assume that a joint probability distribution  $p(\mathbf{x}, y)$  exists.
- Then, by the Bayes rule we can decompose this probability in

$$p(\mathbf{x}, y) = p(\mathbf{x}|y)p(y) = p(y|\mathbf{x})p(\mathbf{x})$$

The first decomposition is the generative approach, where  $p(y|\mathbf{x})$  is the class conditional distribution and

$$p(y|\mathbf{x}) = \frac{p(\mathbf{x}|y)p(y)}{p(\mathbf{x})}$$

where  $p(\mathbf{x}) = \sum_c p(\mathbf{x}|(C)_c)p(\mathcal{C}_c)$  and  $\mathcal{C}_c$  is each one of the possible class distributions.

- The discriminative approach is intended to model  $p(y|\mathbf{x})$  in a direct way without making inference on  $p(\mathbf{x}|y)$ .
- Both methods, nevertheless, need probabilistic models. For the generative model, one can use a Gaussian class conditional model

$$p(\mathbf{x}|\mathcal{C}_c) = \mathcal{N}(\boldsymbol{\mu}_c, \Sigma_c).$$

- For the discriminative case, we can establish an activation or response function bounded between 0 and 1. The linear Logistic Regression is often used

$$p(y = 1|\mathbf{x}) = \frac{1}{1 + \exp(-\mathbf{x}^\top \mathbf{w})}$$

- For the logistic function, if  $\mathbf{x}^\top \mathbf{w}$  is positive, then the probability model tends to 1 when its absolute value increases, and when it is negative, the probability tends to zero if the absolute value increases.
- The Probit function is also used to model a discriminative distribution as

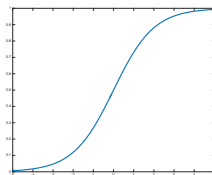
$$p(y = 1|\mathbf{x}) = \int_{-\infty}^{\mathbf{x}^\top \mathbf{w}} \mathcal{N}(\tau, 1) d\tau$$

- Discriminative approaches are easy to implement. Generative approaches are hard, and the complexity increases with the number of data
- Generative approaches are able to deal with situations where some labels are missing (semisupervised) or where there are no labelling (unsupervised), while discriminative approaches cannot.
- The approach taken by Gaussian processes are nevertheless discriminative.

- The likelihood of a label  $y$  is defined as

$$p(y = 1|\mathbf{x}, \mathbf{w}) = \sigma(\mathbf{x}^\top \mathbf{w})$$

where  $\sigma$  is any sigmoidal function.



- It is easy to see that for the logistic function we have

$$\log \left( \frac{p(y = 1|\mathbf{x}, \mathbf{w})}{p(y = -1|\mathbf{x}, \mathbf{w})} \right) = \mathbf{x}^\top \mathbf{w}$$

The quotient is the so called logit function.

- Assume a dataset  $\{\mathbf{x}_1, y_1 \cdots \mathbf{x}_N, y_N\}$  and use a Gaussian prior  $\mathcal{N}(\mathbf{0}, \Sigma_p)$  for the parameters  $\mathbf{w}$ . Then, we can compute a posterior for the parameters of the form

$$\log p(\mathbf{w}|\mathbf{X}, \mathbf{y}) = -\frac{1}{2}\mathbf{w}^\top \Sigma_p^{-1} \mathbf{w} + \sum_{n=1}^N \log \sigma(y_n f_n)$$

where  $f_n = \mathbf{x}_n^\top \mathbf{w}$  and where the constant (normalization) terms have been omitted.

- Note that the first term is a quadratic form, and the second one is monotonic, so the posterior has a single maximum.

- The posterior allows us to compute the predictive distribution for a test data

$$p(y = 1|\mathbf{x}^*, \mathbf{X}, \mathbf{y}) = \int p(y = 1|\mathbf{x}^*, \mathbf{w})p(\mathbf{w}|\mathbf{X}, \mathbf{y})d\mathbf{w}$$



- Assume a prior over the latent function  $f(\mathbf{x})$  and then a logistic function over this prior.

$$\pi(\mathbf{x}) = p(y = 1|\mathbf{x}) = \sigma(f(\mathbf{x}))$$

- A predictive distribution can then be computed as it is done in regression as

$$p(f(\mathbf{x}^*)|\mathbf{X}, \mathbf{y}, \mathbf{x}^*) = \int p(f(\mathbf{x}^*)|\mathbf{X}, \mathbf{x}^*, \mathbf{f})p(\mathbf{f}|\mathbf{X}, \mathbf{y})d\mathbf{f}$$

where  $p(\mathbf{f}|\mathbf{X}, \mathbf{y})$  is the posterior of the latent function over the data.

- If we use this distribution to construct the posterior over the latent  $f(\mathbf{x}^*)$  we produce a probabilistic prediction

$$p(y_* = 1 | \mathbf{X}, \mathbf{y}, \mathbf{x}^*) = \int \sigma(f(\mathbf{x}^*)) p(f(\mathbf{x}^*) | \mathbf{X}, \mathbf{y}, \mathbf{x}^*) df(\mathbf{x}^*)$$

- These two integrals are analytically intractable, so we need analytical approximations of integrals or Monte Carlo sampling.
- In GP two approaches are used: The Laplace approximation and the Expectation Propagation (EP) method.

- The method is based on a Gaussian approximation of the posterior probability  $p(\mathbf{f}|\mathbf{X}, \mathbf{y})$ .
- The approximation is obtained based on a second order Taylor expansion of the posterior:

$$q(\mathbf{f}|\mathbf{X}, \mathbf{y}) = \mathcal{N}(\mathbf{f}|\hat{\mathbf{f}}, \mathbf{A}^{-1}) \propto \exp\left(-\frac{1}{2}(\mathbf{f} - \hat{\mathbf{f}})^\top \mathbf{A}(\mathbf{f} - \hat{\mathbf{f}})\right)$$

where

$$\hat{\mathbf{f}} = \arg \max_{\mathbf{f}} p(\mathbf{f}|\mathbf{X}, \mathbf{y})$$

$$\mathbf{A} = -\nabla \nabla \log p(\mathbf{f}|\mathbf{X}, \mathbf{y})|_{\mathbf{f}=\hat{\mathbf{f}}}$$

- By the Bayes' rule

$$p(\mathbf{f}|\mathbf{X}, \mathbf{y}) = \frac{p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|\mathbf{X})}{p(\mathbf{y}|\mathbf{X})} \propto p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|\mathbf{X})$$

Taking logarithms

$$\log p(\mathbf{f}|\mathbf{X}, \mathbf{y}) \propto \Psi(\mathbf{f}) = \log p(\mathbf{y}|\mathbf{f}) + \log p(\mathbf{f}|\mathbf{X})$$

- The second term has been computed for regression:

$$\Psi(\mathbf{f}) = \log p(\mathbf{y}|\mathbf{f}) - \frac{1}{2}\mathbf{f}^\top \mathbf{K}^{-1}\mathbf{f} - \frac{1}{2}\log |\mathbf{K}| - \frac{n}{2}\log 2\pi$$

- The first goal here is to find the maximum of this posterior. We use the Newton's method, which needs the gradient and the Hessian:

$$\nabla \Psi(\mathbf{f}) = \nabla \log p(\mathbf{y}|\mathbf{f}) + \mathbf{K}^{-1}\mathbf{f}$$

$$\nabla \nabla \Psi(\mathbf{f}) = \nabla \nabla \log p(\mathbf{y}|\mathbf{f}) + \mathbf{K}^{-1}$$

where the gradient and the Hessian of  $\log p(\mathbf{y}|\mathbf{f})$  depend on the model that we use for this distribution (see textbook, table pag. 43).

- The gradient has to be nulled:

$$\nabla \log p(\mathbf{y}|\mathbf{f}) + \mathbf{K}^{-1}\mathbf{f} = \mathbf{0}$$

with which

$$\hat{\mathbf{f}} = \mathbf{K} \nabla \log p(\mathbf{y}|\hat{\mathbf{f}})$$

Since this equation is transcendental, we can only find  $\hat{\mathbf{f}}$  iteratively as

$$\mathbf{f}^{k+1} = \mathbf{f}^k - (\nabla \nabla \Psi)^{-1} \nabla \Psi$$

- We define  $\mathbf{W} = -\nabla \nabla \log p(\mathbf{y}|\mathbf{f})$ , and then the posterior can be approximated by a Gaussian

$$p(\mathbf{f}|\mathbf{X}, \mathbf{y}) \approx q(\mathbf{f}|\mathbf{X}, \mathbf{y}) = \mathcal{N}(\hat{\mathbf{f}}, (\mathbf{K}^{-1} + \mathbf{W})^{-1})$$

- The goal of the prediction is to find the expectation of the predictive distribution. Since the approximation is Gaussian, we can use the expression obtained for regression

$$\mathbb{E}[f_*|\mathbf{X}, \mathbf{y}, \mathbf{x}_*] = \mathbf{k}(\mathbf{x}_*)\mathbf{K}^{-1}\hat{\mathbf{f}}$$

and since

$$\hat{\mathbf{f}} = \mathbf{K}\nabla \log p(\mathbf{y}|\hat{\mathbf{f}})$$

then

$$\mathbb{E}[f_*|\mathbf{X}, \mathbf{y}, \mathbf{x}_*] = \mathbf{k}(\mathbf{x}_*)\nabla \log p(\mathbf{y}|\hat{\mathbf{f}})$$

- The variance can also be computed by using the expression obtained for regression

$$\mathbb{V}(f_*|\mathbf{X}, \mathbf{y}, \mathbf{x}_*) = k(\mathbf{x}_*, \mathbf{x}_*) - \mathbf{k}_*^\top (\mathbf{K} + \mathbf{W}^{-1})^{-1} \mathbf{k}_*$$

- Approximations of multiclass classifier exist by using the Softmax activation

$$p(y_i^c | \mathbf{f}_i) = \frac{\exp(f_i^c)}{\sum_c \exp(f_i^c)}$$



- It is useful to compute an approximation for the marginal likelihood:

$$\log q(\mathbf{y}|\mathbf{X}, \boldsymbol{\theta}) = -\frac{1}{2}\hat{\mathbf{f}}^\top \mathbf{K}^{-1}\hat{\mathbf{f}} + \log p(\mathbf{y}|\hat{\mathbf{f}}) - \frac{1}{2} \log |\mathbf{B}|$$

with  $\mathbf{B} = |\mathbf{I} + \mathbf{W}^{\frac{1}{2}}\mathbf{K}\mathbf{W}^{\frac{1}{2}}|$

- It is necessary in order to find optimal values for the hyperparameters (see Rasmussen, Ch. 5)

- The EP procedure is used as an alternative way of estimating the parameters of the GP for classification.
- It is based on the factorization of the posterior distribution over the latent variables  $\mathbf{f}$ :

$$p(\mathbf{f}|\mathbf{X}, \mathbf{y}) = \frac{p(\mathbf{f}|\mathbf{X})p(\mathbf{y}|\mathbf{f})}{p(\mathbf{y}|\mathbf{X})} = \frac{1}{\mathbf{Z}}p(\mathbf{f}|\mathbf{X}) \prod_{i=1}^N p(y_i|f_i)$$

where  $\mathbf{Z} = p(\mathbf{y}|\mathbf{X}) = \int p(\mathbf{f}|\mathbf{X}) \prod_{i=1}^N p(y_i|f_i) d\mathbf{f}$

- Here, the likelihood over labels  $y_i$  is modelled using the probit function

$$p(y_i|f_i) = \Phi(f_i y_i)$$

But this makes the posterior intractable, so we approximate it by a Gaussian

$$p(y_i|f_i) \approx t_i(f_i|\tilde{\mathbf{Z}}_i, \tilde{\boldsymbol{\mu}}_i, \tilde{\sigma}_i^2) = \tilde{\mathbf{Z}}_i \mathcal{N}(f_i, \tilde{\boldsymbol{\mu}}_i, \tilde{\sigma}_i^2)$$

- With this approximation, the product of (local) likelihoods becomes

$$\prod_{i=1}^N t_i(f_i | \tilde{\mathbf{Z}}_i, \tilde{\boldsymbol{\mu}}_i, \tilde{\sigma}_i^2) = \mathcal{N}(\tilde{\boldsymbol{\mu}}, \tilde{\boldsymbol{\Sigma}}) \prod_i \tilde{\mathbf{Z}}_i$$

Where the mean is a vector of all means and the covariance matrix is a diagonal containing all variances  $\tilde{\sigma}_i^2$

- With that, the posterior  $p(\mathbf{f}|\mathbf{X}, \mathbf{y})$  is approximated as

$$q(\mathbf{f}|\mathbf{X}, \mathbf{y}) = \frac{1}{\mathbf{Z}_{EP}} p(\mathbf{f}|\mathbf{X}) \prod_{i=1}^N t_i(f_i|\tilde{\mathbf{Z}}_i, \tilde{\boldsymbol{\mu}}_i, \tilde{\sigma}_i^2) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

which is a Gaussian with parameters  $\boldsymbol{\mu} = \boldsymbol{\Sigma} \tilde{\boldsymbol{\Sigma}}^{-1} \tilde{\boldsymbol{\mu}}$  and  $\boldsymbol{\Sigma} = (\mathbf{K}^{-1} + \tilde{\boldsymbol{\Sigma}}^{-1})^{-1}$

- The EP chooses then the parameters of the local likelihoods in a sequential way.

- The algorithm sequentially approximates each  $t_i$  by using the current approximations for all other variables. Each  $t_i$  contains the following terms:
  - 1 The prior  $p(\mathbf{f}|\mathbf{X})$
  - 2 The rest of likelihoods  $t_j$
  - 3 The exact likelihood  $p(y_i|f_i) = \Phi(y_i f_i)$
- To this end, we construct the *cavity function*

$$q_{-i}(f_i) \propto \int p(\mathbf{f}|\mathbf{X}) \prod_{j \neq i} t_j(f_j | \tilde{\mathbf{Z}}_i, \tilde{\boldsymbol{\mu}}_i, \tilde{\sigma}_i^2) df_j$$

- This cavity function is another Gaussian

$$q_i(f_i) = \mathcal{N}(f_i | \mu_{-i}, \sigma_{-i}^2) \quad (1)$$

with  $\mu_{-i} = \sigma_{-i}^2(\sigma_i^{-2}\mu_i - \tilde{\sigma}_i^{-2}\tilde{\mu}_i)$  and  $\sigma_{-i}^2 = (\sigma_i^2 - \tilde{\sigma}_i^2)^{-1}$

- The EP algorithm can then be summarized as
  - 1 computing the marginal for sample  $i$

$$\hat{q}(f_i) \triangleq \hat{Z}_i \mathcal{N}(\hat{\mu}_i, \hat{\sigma}_i^2) \approx q_{-i}(f_i) p(y_i | f_i)$$

- 2 Find the parameters  $\tilde{\mu}_i, \tilde{\sigma}_i$  of the approximate likelihood that minimize the KL divergence with respect the true likelihood.

We must iterate this procedure for all samples.

The values of the parameters for each iteration are:

$$\hat{Z}_i = \Phi(z_i) \qquad \hat{\mu}_i = \mu_{-i} + \frac{y_i \sigma_{-i}^2 \mathcal{N}(z_i)}{\Phi(z_i) \sqrt{1 + \sigma_{-1}^2}}$$

$$\hat{\sigma}_i^2 = \sigma_{-1}^2 - \frac{\sigma_{-i}^4 \mathcal{N}(z_i)}{(1 + \sigma_{-i}^2) \Phi(z_i)} \left( z_i + \frac{\mathcal{N}(z_i)}{\Phi(z_i)} \right) \qquad \text{with} \quad z_i = \frac{y_i \mu_{-i}}{\sqrt{1 + \sigma_{-1}^2}}$$



Then, the moments of the approximate posterior are computed as

$$\tilde{\mu}_i = \tilde{\sigma}_i^2(\hat{\sigma}_i^{-2}\hat{\mu}_i - \sigma_{-i}^{-2}\mu_{-i}), \quad \tilde{\sigma}_i^2 = (\hat{\sigma}_i^{-2} - \sigma_{-i}^{-2})^{-1}$$

$$\tilde{Z}_i = \hat{Z}_i \sqrt{2\pi} \sqrt{\sigma_{-i}^2 + \tilde{\sigma}_i^2} \exp\left(\frac{1}{2}(\mu_{-i} - \tilde{\mu}_i)^2 / (\sigma_{-i}^2 + \tilde{\sigma}_i^2)\right)$$

The predictive mean and variance are:

$$\mathbb{E}[f_*|\mathbf{X}, \mathbf{y}, \mathbf{x}_*] = \mathbf{k}_* \left( \mathbf{K} + \tilde{\Sigma} \right)^{-1} \tilde{\boldsymbol{\mu}}$$

$$\mathbb{V}[f_*|\mathbf{X}, \mathbf{y}, \mathbf{x}_*] = k(\mathbf{x}_*, \mathbf{x}_*) - \mathbf{k}_*^\top \left( \mathbf{K} + \tilde{\Sigma} \right)^{-1} \mathbf{k}_*$$

And the predictive distribution for the target is

$$q(y_* = 1|\mathbf{X}, \mathbf{y}, \mathbf{x}_*) = \mathbb{E}[\pi_*|\mathbf{X}, \mathbf{y}, \mathbf{x}_*] = \int \Phi(f_*) q(f_*|\mathbf{X}, \mathbf{y}, \mathbf{x}_*) df_*$$

that becomes

$$q(y_* = 1|\mathbf{X}, \mathbf{y}, \mathbf{x}_*) = \Phi \left( \frac{\mathbb{E}[f_*|\mathbf{X}, \mathbf{y}, \mathbf{x}_*]}{\sqrt{1 + \mathbb{V}[f_*|\mathbf{X}, \mathbf{y}, \mathbf{x}_*]}} \right)$$

- Input:  $\mathbf{K}$  and  $y$
- Initialize parameters of slide 21 to zero.
- Repeat until convergence
  - ➊ Compute parameters of eq. (1)
  - ➋ Compute moments of slide 24
  - ➌ Update parameters of slide 25
  - ➍ Update  $\Sigma$  and  $\mu$