

August 22: Exploring causal assumptions with string diagrams

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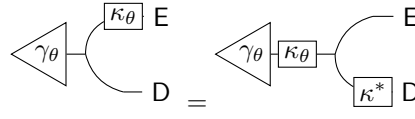
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1 Recoverability

A natural assumption suggested by the notion of a CSDP is that of *recoverability* - that a causal theory $\mathcal{T} : E \times D \rightarrow E$ permits some decision function that reproduces the distribution of the observed data. That is, we assume that for every $(\kappa_\theta, \mu_\theta) := \theta \in \mathcal{T}$ there exists $\gamma_\theta \in \Delta(\mathcal{D})$ such that

$$\gamma_\theta \kappa_\theta = \mu_\theta \quad (1)$$

Suppose also that we have some κ^* that, for all $\theta \in \mathcal{T}$, is a Bayesian inversion of γ_θ and κ_θ ; that is:



$$\text{String diagram (2)} \quad (2)$$

A sufficient condition for the existence of such a κ^* is the assumption that decisions correspond to *variable setting* - that is, there is some variable $X : E \rightarrow X$ such that for all $a \in D$, $\theta \in \mathcal{T}$ we have $\delta_a \kappa_\theta F_X = \delta_a$ (such an assumption arises in graphical models as hard interventions, and in potential outcomes as “potential-outcome identifiers”). Indeed F_X is in this case a candidate for κ^* . It is not necessary that κ^* be deterministic, however - suppose every κ ignores D . Then choose $\gamma_\theta = \gamma$ for arbitrary $\gamma \in \Delta(\mathcal{D})$ and it can be verified that $\kappa^* : b \mapsto \gamma$ satisfies 2.

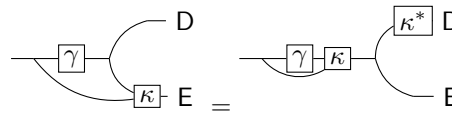
I believe a weaker sufficient condition for the existence of a universal κ^* is that every κ_θ factorises as $\kappa_\theta = h \vee (\text{Id}_F \otimes j_\theta)$ for some fixed $h : D \rightarrow \Delta(\mathcal{F})$, but I have not yet shown this.

We will proceed somewhat rashly: suppose that by defining $\gamma : \mathcal{T} \rightarrow \Delta(\mathcal{D})$, $\mu : \mathcal{T} \rightarrow \Delta(\mathcal{E})$ and $\kappa : \mathcal{T} \times D \rightarrow \Delta(\mathcal{E})$ by $\gamma : \theta \rightarrow \gamma_\theta$, $\mu : \theta \rightarrow \mu_\theta$ and $\kappa : (\theta, d) \rightarrow \kappa_\theta(d; \cdot)$ that all resulting objects are Markov kernels, and that \mathcal{T} is a standard measurable space.

By previous assumptions, we have the following properties:



$$\text{String diagram (3)} \quad (3)$$



$$\text{String diagram (4)} \quad (4)$$



$$\text{String diagram (5)} \quad (5)$$

18 From 4 we also have

$$\begin{array}{c} \text{---} \boxed{\gamma} \text{---} \end{array} \begin{array}{c} \text{---} \text{D} \\ \text{---} \boxed{\kappa^*} \text{---} \end{array} \begin{array}{c} \text{---} \end{array} = \begin{array}{c} \text{---} \boxed{\gamma} \text{---} \boxed{\kappa} \text{---} \end{array} \begin{array}{c} \text{---} \boxed{\kappa^*} \text{---} \text{D} \\ \text{---} \text{---} \end{array} \begin{array}{c} \text{---} \end{array} \quad (6)$$

$$\text{---} \boxed{\gamma} \text{---} \text{D} = \text{---} \boxed{\mu} \boxed{\kappa^*} \text{---} \text{D} \quad (7)$$

19 Where 7 follows from 1.

20 Suppose we have that μ is deterministic, relieving us of having to deal with any issues regarding
 21 inferring the observed distribution from a finite sample (e.g. we are in the world of classical physics).
 22 μ therefore has a left inverse ${}^\dagger\mu$.

23 A corollary of Lemma 2.5 is that left inverses have the following property:

$$\begin{array}{c} \text{---} \boxed{{}^\dagger A} \text{---} \end{array} \begin{array}{c} \text{---} \\ \text{---} \boxed{A} \text{---} \end{array} = \begin{array}{c} \text{---} \boxed{{}^\dagger A} \text{---} \end{array} \begin{array}{c} \text{---} \end{array} \quad (8)$$

24 And Fong [2013] has shown that for deterministic A we have

$$\begin{array}{c} \text{---} \boxed{A} \text{---} \end{array} \begin{array}{c} \text{---} \\ \text{---} \end{array} = \begin{array}{c} \text{---} \boxed{A} \text{---} \\ \text{---} \boxed{A} \text{---} \end{array} \quad (9)$$

25 We then have

$$\begin{array}{c} \text{---} \boxed{\mu} \text{---} \boxed{\kappa^*} \text{---} \boxed{\kappa} \text{---} \end{array} \begin{array}{c} \text{---} \\ \text{---} \end{array} \stackrel{4}{=} \begin{array}{c} \text{---} \boxed{\mu} \text{---} \boxed{{}^\dagger\mu\mu} \text{---} \end{array} \begin{array}{c} \text{---} \\ \text{---} \boxed{\kappa^*} \text{---} \end{array} \quad (10)$$

$$\stackrel{4}{=} \begin{array}{c} \text{---} \boxed{\mu} \text{---} \boxed{{}^\dagger\mu} \text{---} \boxed{\mu} \text{---} \boxed{\mu} \text{---} \boxed{\kappa^*} \text{---} \boxed{\kappa} \text{---} \end{array} \begin{array}{c} \text{---} \\ \text{---} \boxed{\kappa^*} \text{---} \end{array} \quad (11)$$

$$\stackrel{4}{=} \begin{array}{c} \text{---} \boxed{\mu} \text{---} \boxed{{}^\dagger\mu} \text{---} \boxed{\mu} \text{---} \boxed{\kappa^*} \text{---} \boxed{\kappa} \text{---} \end{array} \begin{array}{c} \text{---} \\ \text{---} \end{array} \quad (12)$$

$$\stackrel{29}{=} \begin{array}{c} \text{---} \boxed{\mu} \text{---} \boxed{{}^\dagger\mu} \text{---} \boxed{\kappa^*} \text{---} \boxed{\kappa} \text{---} \end{array} \begin{array}{c} \text{---} \\ \text{---} \end{array} \quad (13)$$

$$\stackrel{9}{=} \begin{array}{c} \text{---} \boxed{\mu} \text{---} \boxed{{}^\dagger\mu} \text{---} \boxed{\mu\kappa^*} \text{---} \boxed{\kappa} \text{---} \end{array} \begin{array}{c} \text{---} \\ \text{---} \end{array} \quad (14)$$

26 A key question is does 14 imply anything non-trivial regarding the following “identifiability” condition
 27 for arbitrary $J : E \rightarrow \Delta(\mathcal{D})$:

$$(15)$$

28 I call this identifiability because the left hand side is the kernel $\mathcal{T} \rightarrow \Delta(\mathcal{E})$ that computes the “result”
 29 of a given state and decision function, while the right hand side implies it is possible to find a J that
 30 minimises the expected utility κu independent of the causal state (this is because we see one of the
 31 input wires and control the other).

32 We may be able to get some insight into this by asking, given matrices A, B, C, D of appropriate
 33 shapes, if $BA = CA$ when does $BDA = CDA$?

34 2 Notes on category theoretic probability and string diagrams

35 Category theoretic treatments of probability theory often start with *probability monads* (for a good
 36 overview, see [Jacobs, 2018]). A monad on some category C is a functor $T : C \rightarrow C$ along with
 37 natural transformations called the unit $\eta : 1_C \rightarrow T$ and multiplication $\mu : T^2 \rightarrow T$. Roughly,
 38 functors are maps between categories that preserve identity and composition structure and natural
 39 transformations are “maps” between functors that also preserve composition structure. The monad
 40 unit is similar to the identity element of a monoid in that application of the identity followed by
 41 multiplication yields the identity transformation. The multiplication transformation is also (roughly
 42 speaking) associative.

43 An example of a probability monad is the discrete probability monad given by the functor $\mathcal{D} : \mathbf{Set} \rightarrow$
 44 \mathbf{Set} which maps a countable set X to the set of functions from $X \rightarrow [0, 1]$ that are probability
 45 measures on X , denoted $\mathcal{D}(X)$. \mathcal{D} maps a measurable function f to $\mathcal{D}f : X \rightarrow \mathcal{D}(X)$ given by
 46 $\mathcal{D}f : x \mapsto \delta_{f(x)}$. The unit of this monad is the map $\eta_X : X \rightarrow \mathcal{D}(X)$ given by $\eta_X : x \mapsto \delta_x$ (which
 47 is equivalent to $\mathcal{D}1_X$) and multiplication is $\mu_X : \mathcal{D}^2(X) \rightarrow \mathcal{D}(X)$ where $\mu_X : \Omega \mapsto \sum_{\phi} \Omega(\phi)\phi$.

48 For continuous distributions we have the Giry monad on the category \mathbf{Meas} of measurable spaces
 49 given by the functor \mathcal{G} which maps a measurable space X to the set of probability measures on X ,
 50 denoted $\mathcal{G}(X)$. Other elements of the monad (unit, multiplication and map between morphisms) are
 51 the “continuous” version of the above.

52 Of particular interest is the Kleisli category of the monads above. The Kleisli C_T category of a
 53 monad T on category C is the category with the same objects and the morphisms $X \rightarrow Y$ in C_T is
 54 the set of morphisms $X \rightarrow TY$ in C . Thus the morphisms $X \rightarrow Y$ in the Kleisli category $\mathbf{Set}_{\mathcal{D}}$ are
 55 morphisms $X \rightarrow \mathcal{D}(Y)$ in \mathbf{Set} , i.e. stochastic matrices, and in the Kleisli category $\mathbf{Meas}_{\mathcal{G}}$ we have
 56 Markov kernels. Composition of arrows in the Kleisli categories correspond to Matrix products and
 57 “kernel products” respectively.

58 Both \mathcal{D} and \mathcal{G} are known to be *commutative* monads, and the Kleisli category of a commutative
 59 monad is a symmetric monoidal category.

60 Diagrams for symmetric monoidal categories consist of wires with arrows, boxes and a couple of
 61 special symbols. The identity object (which we identify with the set $\{*\}$) is drawn as nothing at all
 62 $\{*\} := \square$ and identity maps are drawn as bare wires:

$$\text{Id}_X := \uparrow X \quad (16)$$

63 We draw Kleisli arrows from the unit (i.e. probability distributions) $\mu : \{*\} \rightarrow X$ as triangles and
 64 Kleisli arrows $\kappa : X \rightarrow Y$ (i.e. Markov kernels $X \rightarrow \Delta(\mathcal{Y})$) as boxes. We draw the Kleisli arrow

65 $\mathbb{1}_X : X \rightarrow \{*\}$ (which is unique for each X) as below

$$\mu := \begin{array}{c} \uparrow X \\ \triangleleft \mu \end{array} \quad \kappa := \begin{array}{c} \uparrow Y \\ \boxed{\kappa} \end{array} \quad (17)$$

66 The product of objects in **Meas** is given by $(X, \mathcal{X}) \cdot (Y, \mathcal{Y}) = (X \times Y, \mathcal{X} \otimes \mathcal{Y})$, which we will
 67 often write as just $X \times Y$. Horizontal juxtaposition of wires indicates this product, and horizontal
 68 juxtaposition also indicates the tensor product of Kleisli arrows. Let $\kappa_1 : X \rightarrow W$ and $\kappa_2 : Y \rightarrow Z$:

$$(X \times Y, \mathcal{X} \otimes \mathcal{Y}) := \begin{array}{c} \uparrow X \quad \uparrow Y \end{array} \quad \kappa_1 \otimes \kappa_2 := \begin{array}{c} \uparrow W \quad \uparrow Z \\ \boxed{\kappa_1} \quad \boxed{\kappa_2} \\ \downarrow X \quad \downarrow Y \end{array} \quad (18)$$

69 Composition of arrows is achieved by “wiring” boxes together. For $\kappa_1 : X \rightarrow Y$ and $\kappa_2 : Y \rightarrow Z$
 70 we have

$$\kappa_1 \kappa_2(x; A) = \int_Y \kappa_2(y; A) \kappa_1(x; dy) := \begin{array}{c} \uparrow Z \\ \boxed{\kappa_2} \\ \downarrow Y \\ \boxed{\kappa_1} \\ \downarrow X \end{array} \quad (19)$$

71 Symmetric monoidal categories have the following coherence theorem[Selinger, 2010]:

72 **Theorem 2.1** (Coherence (symmetric monoidal)). *A well-formed equation between morphisms in*
 73 *the language of symmetric monoidal categories follows from the axioms of symmetric monoidal*
 74 *categories if and only if it holds, up to isomorphism of diagrams, in the graphical language.*

75 Isomorphism of diagrams for symmetric monoidal categories (somewhat informally) is any planar
 76 deformation of a diagram including deformations that cause wires to cross. We consider a diagram
 77 for a symmetric monoidal category to be well formed only if all wires point upwards.

78 In fact the Kleisli categories of the probability monads above have (for each object) unique *copy*:
 79 $X \rightarrow X \times X$ and *erase*: $X \rightarrow \{*\}$ maps that satisfy the *commutative comonoid axioms* that (thanks
 80 to the coherence theorem above) can be stated graphically. These differ from the copy and erase
 81 maps of *finite product* or *cartesian* categories in that they do not necessarily respect composition of
 82 morphisms.

$$\text{Erase} = \mathbb{1}_X := \begin{array}{c} * \\ \uparrow \end{array} \quad \text{Copy} = x \mapsto \delta_{x,x} := \begin{array}{c} \swarrow \quad \searrow \\ \uparrow \end{array} \quad (20)$$

$$\begin{array}{c} \swarrow \quad \searrow \\ \uparrow \end{array} = \begin{array}{c} \swarrow \quad \searrow \\ \uparrow \end{array} := \begin{array}{c} \swarrow \quad \searrow \\ \uparrow \end{array} \quad (21)$$

$$\begin{array}{c} * \\ \uparrow \end{array} = \begin{array}{c} * \\ \uparrow \end{array} = \begin{array}{c} \uparrow \end{array} \quad (22)$$

$$\begin{array}{c} \swarrow \quad \searrow \\ \uparrow \end{array} = \begin{array}{c} \swarrow \quad \searrow \\ \uparrow \end{array} \quad (23)$$

83 Finally, $\{*\}$ is a terminal object in the Kleisli categories of either probability monad. This means
 84 that the map $X \rightarrow \{*\}$ is unique for all objects X , and as a consequence for all objects X, Y and all
 85 $\kappa : X \rightarrow Y$ we have

$$\begin{array}{c} * \\ \boxed{\kappa} \\ \downarrow \\ X \end{array} = \begin{array}{c} * \\ \downarrow \\ X \end{array} \quad (24)$$

86 This is equivalent to requiring for all $x \in X$ $\int_Y \kappa(x; dy) = 1$. In the case of $\mathbf{Set}_{\mathcal{D}}$, this condition is
 87 what differentiates a stochastic matrix from a general positive matrix (which live in a larger category
 88 than $\mathbf{Set}_{\mathcal{D}}$).

89 Thus when manipulating diagrams representing Markov kernels in particular (and, importantly, not
 90 more general symmetric monoidal categories) diagram isomorphism also includes applications of 21,
 91 22, 23 and 24.

92 A particular property of the copy map in $\mathbf{Meas}_{\mathcal{G}}$ (and probably $\mathbf{Set}_{\mathcal{D}}$ as well) is that it commutes with
 93 Markov kernels iff the markov kernels are deterministic [Fong, 2013].

94 2.1 Disintegration and Bayesian inversion

95 *Disintegration* is a key operation on probability distributions (equivalently arrows $\{*\} \rightarrow X$) in
 96 the categories under discussion. It corresponds to “finding the conditional probability” (though
 97 conditional probability is usually formalised in a slightly different way).

98 Given a distribution $\mu : \{*\} \rightarrow X \otimes Y$, a disintegration $c : X \rightarrow Y$ is a Markov kernel that satisfies

$$\begin{array}{c} X \quad Y \\ \downarrow \quad \downarrow \\ \triangleleft \mu \end{array} = \begin{array}{c} X \quad Y \\ \downarrow \quad \downarrow \\ \boxed{c} \\ \downarrow \\ \triangleleft \mu \end{array} \quad (25)$$

99 Disintegrations always exist in $\mathbf{Set}_{\mathcal{D}}$ but not in $\mathbf{Meas}_{\mathcal{G}}$. They do exist in the latter if we restrict
 100 ourselves to standard measurable spaces. If c_1 and c_2 are disintegrations $X \rightarrow Y$ of μ , they are equal
 101 μ -A.S. In fact, this equality can be strengthened somewhat - they are equal almost surely with respect
 102 to any distribution that shares the “ X -marginal” of μ .

103 Given $\sigma : \{*\} \rightarrow X$ and a channel $c : X \rightarrow Y$, a Bayesian inversion of (σ, c) is a channel $d : Y \rightarrow X$
 104 such that

$$\begin{array}{c} X \quad Y \\ \downarrow \quad \downarrow \\ \boxed{c} \\ \downarrow \\ \triangleleft \sigma \end{array} = \begin{array}{c} X \quad Y \\ \downarrow \quad \downarrow \\ \boxed{d} \\ \downarrow \\ \boxed{c} \\ \downarrow \\ \triangleleft \sigma \end{array} \quad (26)$$

105 We can obtain disintegrations from Bayesian inversions and vice-versa.

106 Clerc et al. [2017] offer an alternative view of Bayesian inversion which they claim doesn’t depend
 107 on standard measurability conditions, but there is a step in their proof I didn’t follow.

108 2.2 Generalisations

109 Cho and Jacobs [2019] make use of a larger “CD” category by dropping 24. I’m not completely clear
 110 whether you end up with arrows being “Markov kernels for general measures” or something else (can
 111 we have negative arrows?). This allows for the introduction of “observables” or “effects” of the form

$$\begin{array}{c} \triangle \\ f \\ \downarrow \end{array} .$$

113 Jacobs et al. [2019] make use of an embedding of \mathbf{Set}_D in $\mathbf{Mat}(\mathbb{R}^+)$ with morphisms all positive
 114 matrices (I’m not totally clear on the objects, or how they are self-dual - this doesn’t seem to be
 115 exactly the same as the category of finite dimensional vector spaces). This latter category is compact
 116 closed, which - informally speaking - supports the same diagrams as symmetric monoidal categories
 117 with the addition of “upside down” wires.

118 2.3 Key questions for Causal Theories

119 We will first define *labeled diagrams*. Rather than labelling the wires of our diagrams with *spaces* (as is
 120 typical [Selinger, 2010]), we assign a unique label to each “wire segment” (with some qualifications).
 121 That is, we assign a unique label to each bare wire in the diagram with the following additional
 122 qualifications:

- 123 • If we have a box in the diagram representing the identity map, the incoming and outgoing
 124 wires are given the same label
- 125 • If we have a wire crossing in the diagram, the diagonally opposite wires are given the same
 126 label
- 127 • The input wire and the *two* output wires of the copy map are given the same label

128 Given two diagrams G_1 and G_2 that are isomorphic under transformations licenced by the axioms of
 129 symmetric monoidal categories and commutative comonoid axioms, suppose we have a labelling of
 130 G_1 . We can label G_2 using the following translation rule:

- 131 • For each box in G_2 , we can identify a corresponding box in G_1 via labels on each box. For
 132 each such pair of boxes, we label the incoming wires of the G_2 box with the labels of the
 133 G_1 box preserving the left-right order. We do likewise for outgoing wires.

134 These rules will lead to a unique labelling of G_2 with all wire segments are labelled. We would like
 135 for these rules to yield the following:

- 136 • For any sequence of diagram isomorphisms beginning with G_1 and ending with G_2 , we end
 137 up with the same set of labels
- 138 • If we label G_2 according to the rules above then relabel G_1 from G_2 according to the same
 139 rules we retrieve the original labels of G_1

140 We do not prove these properties here, but motivate them via the following considerations:

- 141 • These properties obviously hold for the wire segments into and out of boxes
- 142 • The only features a diagram may have apart from boxes and wires are wire crossings, copy
 143 maps and erase maps
- 144 • The labeling rule for wire crossings respects the symmetry of the swap map
- 145 • The labeling rule for copy maps respects the symmetry of the copy map and the property
 146 described in Equation 23

147 We will follow the convention whereby “internal” wire labels are omitted from diagrams.

148 Note also that each wire that terminates in a free end can be associated with a random variable.
 149 Suppose for $N \in \mathbb{N}$ we have a kernel $\kappa : A \rightarrow \Delta(\times_{i \in N} X_i)$. Define by p_j ($j \in [N]$) the projection
 150 map $p_j : \times_{i \in N} X_i \rightarrow X_j$ defined by $p_j : (x_0, \dots, x_N) \mapsto x_j$. p_j is a measurable function, hence
 151 a random variable. Define by π_j the projection kernel $\mathcal{G}(\pi_j)$ (that is, $\pi_j : \mathbf{x} \mapsto \delta_{p_j(\mathbf{x})}$). Note that
 152 $\kappa(y; p_j^{-1}(A)) = \int_{X_j} \delta_{p_j(\mathbf{x})}(A) \kappa(y; d\mathbf{x}) = \kappa \pi_j$. Diagrammatically, π_j is the identity map on the j -th
 153 wire tensored with the erase map on every other wire. Thus the j -th wire carries the distribution
 154 associated with the random variable p_j . We will therefore consider the labels of the “outgoing” wires
 155 of a diagram to denote random variables (though there are obviously many random variables not
 156 represented by such wires). We will additionally distinguish wire labels from spaces by font - wire
 157 labels are sans serif A, B, C, X, Y, Z while spaces are serif A, B, C, X, Y, Z .

Wire labels appear to have a key advantage over random variables: they allow us to “forget”
 the sample space as the correct typing is handled automatically by composition and erasure of
 wires

I’m sure one of the papers I read mentioned labeled diagrams, I just couldn’t find it when I looked for it

Since writing this, I found Kissinger [2014] as an example of a diagrammatic system with labeled wires, I will follow it up

159 **generalised disintegrations** : Of key importance to our work is generalising the notion of disinte-
 160 gration (and possibly Bayesian inversion) to general kernels $X \rightarrow Y$ rather than restricting ourselves
 161 to probability distributions $\{\ast\} \rightarrow Y$. We will define generalised disintegrations as a straightforward
 162 analogy regular disintegrations, but the conditions under which such disintegrations exist are more
 163 restrictive than for regular disintegrations.

164 **Definition 2.2** (Label signatures). If a kernel $\kappa : X \rightarrow \Delta(Y)$ can be represented by a diagram
 165 G with incoming wires X_1, \dots, X_n and outgoing wires Y_1, \dots, Y_m , we can assign the kernel a “label
 166 signature” $\kappa : X_1 \otimes \dots \otimes X_n \dashrightarrow Y_1 \otimes \dots \otimes Y_m$ or, for short, $\kappa : X_{[n]} \dashrightarrow Y_{[m]}$. Note that this
 167 signature associates each label with a unique space - the space of X_1 is the space associated with the
 168 left-most wire of G and so forth. We will implicitly leverage this correspondence and write with X_1
 169 the space associated with X_1 and so forth. Note that while X_1 is by construction always different from
 170 X_2 (or any other label), the space X_1 may coincide with X_2 - the fact that labels always maintain
 171 distinctions between wires is the fundamental reason for introducing them in the first place.

172 There might actually be some sensible way to consider κ to be transforming the measurable
 functions of a type similar to $\otimes_{i \in [n]} X_i$ to functions of a type similar to $\otimes_{i \in [m]} Y_i$ (or vise
 versa - perhaps related to Clerc et al. [2017]), but wire labels are all we need at this point

173 **Definition 2.3** (Generalised disintegration). Given a kernel $\kappa : X \rightarrow \Delta(Y)$ with label signature
 174 $\kappa : X_{[n]} \dashrightarrow Y_{[m]}$ and disjoint subsets $S, T \subset [m]$ such that $S \cup T = [m]$, a kernel c is a *g-*
 175 *disintegration from S to T* if it’s type is compatible with the label signature $c : Y_S \dashrightarrow Y_T$ and we
 176 have the identity (omitting incoming wire labels):

$$(27)$$

177 I have introduced without definition additional labeling operations here: first, each label has
 a particular space associated with it (in order to license the notion of “type compatible with
 label signature”), and we have supposed labels can be “bundled”.

178 In contrast to regular disintegrations, generalised disintegrations “usually” do not exist. Consider
 179 $X = \{0, 1\}$, $Y = \{0, 1\}^2$ and κ has label signature $X_1 \dashrightarrow Y_{\{1,2\}}$ with

$$\kappa : \begin{cases} 1 \mapsto \delta_1 \otimes \delta_1 \\ 0 \mapsto \delta_1 \otimes \delta_0 \end{cases} \quad (28)$$

180 κ imposes contradictory requirements for any disintegration $c : \{0, 1\} \rightarrow \{0, 1\}$ from $\{1\}$ to $\{2\}$:
 181 equality for $X_1 = 1$ requires $c(1; \cdot) = \delta_1$ while equality for $X_1 = 0$ requires $c(1; \cdot) = \delta_0$. Subject
 182 to some regularity conditions (similar to standard Borel conditions for regular disintegrations),
 183 we can define g-disintegrations of a canonically related kernel that do generally exist; intuitively,
 184 g-disintegrations exist if they take the “input wires” of κ as input wires themselves.

185 **Lemma 2.4.** Given $\kappa : X \rightarrow \Delta(Y)$, a kernel κ^\dagger is a right inverse iff we have for all $x \in X$
 186 $\kappa^\dagger(y; A) = \delta_x(A)$, $\kappa(x; \cdot)$ -almost surely.

187 *Proof.* Suppose κ^\dagger satisfies the almost sure equality for all $x \in X$. Then for all $x \in X$, $A \in \mathcal{Y}$ we
 188 have $\kappa \kappa^\dagger(x; A) = \int_Y \kappa^\dagger(y; A) \kappa(x; dy) = \int_Y \delta_x(A) \kappa(x; dy) = \delta_x(A)$; that is, $\kappa \kappa^\dagger = \text{Id}_X$, so κ^\dagger is
 189 a right inverse of κ .

190 Suppose we have a right inverse κ^\dagger . By definition, for all $x \in X$ and $A \in \mathcal{Y}$ we have
 191 $\int_Y \kappa^\dagger(y; A) \kappa(x; dy) = \delta_x(A)$. Suppose $x \notin A$ and let $B_\epsilon = \kappa_A^{\dagger-1}((\epsilon, 1])$ for some $\epsilon > 0$. We
 192 have $\int_Y \kappa^\dagger(y; A) \kappa(x; dy) \geq \epsilon \kappa(x; B)$. For any $\epsilon > 0$ we have $\kappa(x; B_\epsilon) = 0$. Consider the set
 193 $B_0 = \kappa_A^{\dagger-1}((0, 1])$. For some sequence $\{\epsilon_i\}_{i \in \mathbb{N}}$ such that $\lim_{i \rightarrow \infty} \epsilon_i = 0$ we have $B_0 = \cup_{i \in \mathbb{N}} B_{\epsilon_i}$.

194 By countable additivity, $\kappa(x; B_0) = 0$. Suppose $x \in A$ and let $B^{1-\epsilon} = \kappa_A^{\dagger-1}([0, 1 - \epsilon])$. By
 195 an argument analogous to the above, we have $\kappa(x; B^1) = 0$. Thus the $\kappa(x; \cdot)$ measure of the set
 196 on which $\kappa^\dagger(y; A)$ disagrees with $\delta_x(A)$ is $\kappa(x; B_0) + \kappa(x; B^1) = 0$ and hence $\kappa^\dagger(y; A) = \delta_x(A)$
 197 $\kappa(x; \cdot)$ -almost surely. \square

I haven't shown that any map inverting κ implies the existence of a Markov kernel that does so

I am using countable sets below to get my general argument in order without getting too hung up on measurability; I will try to lift it to standard measurable once it's all there

200 **Lemma 2.5.** Given $\kappa : X \rightarrow \Delta(Y)$ and a right inverse κ^\dagger , we have

$$(29)$$

201 *Proof.* Let the diagram on the left hand side be L and the diagram on the right hand side be R .

$$L(x; A \times B) = \int_Y \int_{Y \times Y} \text{Id}_Y \otimes \kappa_S^\dagger(y, y'; A \times B) \delta_{(z, z)}(dy \times dy') \kappa \pi_S(x; dz) \quad (30)$$

$$= \int \text{Id}_Y \otimes \kappa^\dagger(z, z; A \times B) \kappa \pi_S(x; dz) \quad (31)$$

$$= \int \delta_z(A) \kappa_S^\dagger(z; B) \kappa \pi_S(x; dz) \quad (32)$$

$$= \int_A \kappa_S^\dagger(z; B) \kappa \pi_S(x; dz) \quad (33)$$

$$= \delta_x(B) \kappa \pi_S(x; A) \quad (34)$$

202 Where 34 follows from Lemma 2.4.

$$R(x; A \times B) = \int \delta_{(x, x)}(dy \times dy') \kappa \pi_S \otimes \text{Id}_X(y, y'; A \times B) \quad (35)$$

$$= \kappa \pi_S(x; A) \delta_x(B) = L \quad (36)$$

203

\square

204 **Theorem 2.6.** Given countable X and standard measurable Y , $n, m \in \mathbb{N}$, $S, T \subset [m]$, κ with label
 205 signature $X_{[n]} \dashrightarrow Y_{[m]}$ a g -disintegration exists from S to T if $\kappa \pi_S$ is right-invertible

via a Markov kernel

207 *Proof.* In addition, as R is a composition of Markov kernels, and hence a Markov kernel itself, L
 208 must also be a Markov kernel even if κ^\dagger is not.

209 For all $x \in X$ we have a (regular) disintegration $c_x : Y_S \rightarrow \Delta(Y_T)$ of $\kappa(x; \cdot)$ by standard mea-
 210 surability of Y . Define $c : X \otimes Y_S \rightarrow \Delta(Y_T)$ by $c : (x, y_S) \mapsto c_x(y_S)$. Clearly, $c(x, y_S)$ is a
 211 probability distribution on Y_T for all $(x, y_S) \in X \otimes Y_S$. It remains to show $c(\cdot)^{-1}(B)$ is measurable
 212 for all $B \in \mathcal{B}([0, 1])$. But $c(\cdot)^{-1}(B) = \cap_{x \in X} c_y(\cdot)^{-1}(B)$. The right hand side is measurable by
 213 measurability of $c_y(\cdot)^{-1}(B)$ countability of X , so c is a Markov kernel.

214 By the definition of c_x , we have for all $x \in X$

$$\begin{array}{c}
Y_S \quad Y_T \\
\downarrow \quad \downarrow \\
\boxed{\kappa} \\
\downarrow \\
\triangle \delta_x
\end{array}
=
\begin{array}{c}
Y_S \quad Y_T \\
\downarrow \quad \downarrow \\
\boxed{C_x} \\
\downarrow \quad \downarrow \\
\boxed{\kappa}^* \\
\downarrow \\
\triangle \delta_x
\end{array}
\quad (37)$$

$$=
\begin{array}{c}
Y_S \quad Y_T \\
\downarrow \quad \downarrow \\
\boxed{C} \\
\downarrow \quad \downarrow \\
\boxed{\kappa}^* \\
\downarrow \\
\triangle \delta_x
\end{array}
\quad (38)$$

215 Which implies

$$\begin{array}{c}
Y_S \quad Y_T \\
\downarrow \quad \downarrow \\
\boxed{\kappa}
\end{array}
=
\begin{array}{c}
Y_S \quad Y_T \\
\downarrow \quad \downarrow \\
\boxed{C} \\
\downarrow \quad \downarrow \\
\boxed{\kappa}^* \\
\downarrow \\
\triangle \delta_x
\end{array}
\quad (39)$$

216 Finally, we have

$$\begin{array}{c}
Y_S \quad Y_T \\
\downarrow \quad \downarrow \\
\boxed{C} \\
\downarrow \quad \downarrow \\
\boxed{\kappa_S^\dagger}^* \\
\downarrow \\
\boxed{\kappa}
\end{array}
=
\begin{array}{c}
Y_S \quad Y_T \\
\downarrow \quad \downarrow \\
\boxed{C} \\
\downarrow \quad \downarrow \\
\boxed{\kappa_S^\dagger}^* \\
\downarrow \\
\boxed{\kappa}
\end{array}
\quad (40)$$

$$=
\begin{array}{c}
Y_S \quad Y_T \\
\downarrow \quad \downarrow \\
\boxed{C} \\
\downarrow \quad \downarrow \\
\boxed{\kappa}^* \\
\downarrow \\
\triangle \delta_x
\end{array}
\quad (41)$$

217 Where the first line follows from 22 and the second line from 29. If κ_S^\dagger is a Markov kernel, then
218 $\forall (\text{Id}_{Y_S} \otimes \kappa_S^\dagger) c$ is a g-disintegration. \square

219 In the reverse direction, suppose κ is such that $\kappa \pi_T = \text{Id}_X$; that is, π_T is a right inverse of κ . If
220 $\kappa \pi_S$ is not right invertible then, by definition, there is no d such that $\kappa \pi_S d \pi_T = \text{Id}_X$. However, if a
221 g-disintegration of κ exists then there is a d such that $\kappa \pi_S d = \kappa$, a contradiction. Thus if $\kappa \pi_S$ is not
222 right invertible then there is *in general* no g-disintegration from S to T .

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