Elements of Statistical Learning Notes and Exercise Solutions

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Chapter 2

3 Linear Models for Regression

3.2 Linear Regression Models and Least Squares

Derive variance-covariance matrix

Derive the variance-covariance matrix of the least squares parameter estimates (equation 3.8 in book).

The least squares estimator is given by

$$\hat{\beta} = \left(X^T X\right)^{-1} X^T y \tag{1}$$

Express the variance of a random variable as below:

$$Var[X] = \mathbb{E}[X^2] + \mathbb{E}[X]^2$$
(2)

$$\operatorname{Var}\left[\hat{\beta}\right]$$

$$= \mathbb{E}\left[\hat{\beta}^{2}\right] + \mathbb{E}\left[\hat{\beta}\right]^{2}$$

$$= \mathbb{E}\left[\left(\left(X^{T}X\right)^{-1}X^{T}y\right)^{2}\right] + \mathbb{E}\left[\left(X^{T}X\right)^{-1}X^{T}y\right]^{2}$$

$$= \mathbb{E}\left[\left(\left(X^{T}X\right)^{-1}X^{T}\left(X\beta + \epsilon\right)\right)^{2}\right] + \mathbb{E}\left[\left(X^{T}X\right)^{-1}X^{T}\left(X\beta + \epsilon\right)\right]^{2}$$

$$= \mathbb{E}\left[\left(\left(X^{T}X\right)^{-1}X^{T}\left(X\beta + \epsilon\right)\right)^{2}\right] + \mathbb{E}\left[\beta\right]^{2} + \mathbb{E}\left[\left(X^{T}X\right)^{-1}X^{T}\epsilon\right]^{2}$$

The last term goes to zero since X is uncorrelated with the error term.

$$\operatorname{Var}\left[\hat{\beta}\right] = \mathbb{E}\left[\left(\left(X^{T}X\right)^{-1}X^{T}\left(X\beta + \epsilon\right)\right)^{2}\right] - \mathbb{E}\left[\beta\right]^{2}$$

The expectation of the true β is simply β itself.

$$\operatorname{Var}\left[\hat{\beta}\right] = \mathbb{E}\left[\left(\left(X^{T}X\right)^{-1}X^{T}\left(X\beta + \epsilon\right)\right)^{2}\right] - \beta^{2}$$

Break the term within the expectation into two separate terms.

$$\operatorname{Var}\left[\hat{\beta}\right]$$

$$= \mathbb{E}\left[\left(\left(X^{T}X\right)^{-1}X^{T}X\beta + \left(X^{T}X\right)^{-1}X^{T}\epsilon\right)^{2}\right] - \beta^{2}$$

$$= \mathbb{E}\left[\left(\beta + \left(X^{T}X\right)^{-1}X^{T}\epsilon\right)^{2}\right] - \beta^{2}$$

$$= \mathbb{E}\left[\beta^{2} + 2\left(X^{T}X\right)^{-1}X^{T}\epsilon + \left(\left(X^{T}X\right)^{-1}X^{T}\right)^{2}\epsilon^{2}\right] - \beta^{2}$$

$$= \beta^{2} + \mathbb{E}\left[2\left(X^{T}X\right)^{-1}X^{T}\epsilon + \left(\left(X^{T}X\right)^{-1}X^{T}\right)^{2}\epsilon^{2}\right] - \beta^{2}$$

$$= \mathbb{E}\left[\left(\left(X^{T}X\right)^{-1}X^{T}\right)^{2}\epsilon^{2}\right]$$

$$= \mathbb{E}\left[\left(\left(X^{T}X\right)^{-1}X^{T}\right)^{2}\right] \mathbb{E}\left[\epsilon^{2}\right]$$

The error term has variance σ^2 .

$$\operatorname{Var}\left[\hat{\beta}\right]$$

$$= \sigma^{2} \mathbb{E}\left[\left(\left(X^{T}X\right)^{-1}X^{T}\right)^{2}\right]$$

$$= \sigma^{2}\left(X^{T}X\right)^{-1}$$

Expected prediction error

Compute the expected prediction error at input x_0 for model $Y_0 = f(x_0) + \epsilon_0$. Use estimate $\hat{f}(x_0) = x_0^T \tilde{\beta}$.

$$\mathbb{E}\left[\left(Y_{0} - \tilde{f}\left(x_{0}\right)\right)^{2}\right]$$

$$= \mathbb{E}\left[\left(f\left(x_{0}\right) + \epsilon_{0} - \tilde{f}\left(x_{0}\right)\right)^{2}\right]$$

$$= \mathbb{E}\left[f\left(x_{0}\right)^{2} - 2f\left(x_{0}\right)\hat{f}\left(x_{0}\right) - \tilde{f}\left(x_{0}\right)^{2}\right] + \mathbb{E}\left[\epsilon_{0}^{2}\right]$$

$$= \mathbb{E}\left[f\left(x_{0}\right)^{2} - 2f\left(x_{0}\right)\hat{f}\left(x_{0}\right) - \tilde{f}\left(x_{0}\right)^{2}\right] + \sigma^{2}$$

$$= \mathbb{E}\left[\left(f\left(x_{0}\right) - \tilde{f}\left(x_{0}\right)\right)^{2}\right] + \sigma^{2}$$

$$= \sigma^{2} + MSE\left(\tilde{f}\left(x_{0}\right)\right)$$

Ridge regression coefficients

The loss function for ridge regression (page 64, equation 3.43) is

$$RSS(\lambda) = (y - X\beta)^{T} (y - X\beta) + \lambda \beta^{T} \beta$$
(3)

Differentiate with respect to β , set to zero, and solve.

$$\frac{\partial RSS}{\partial \beta}$$

$$= 2(y - X\beta)(-X) + 2\beta\lambda = 0$$

$$= -Xy + 2X^{T}X\beta + 2\beta = 0$$

$$X^{T}y = X^{T}X\beta + \beta$$

$$X^{T}y = \beta(X^{T}X + I)$$

$$\hat{\beta} = (X^{T}X + I)^{-1}X^{T}y$$

Exercise 3.4

Show how the vector of least squares coefficients can be obtained from single pass of the Gram-Schmidt procedure.

Represent X using the QR decomposition

$$X = QR \tag{4}$$

where Q is an $N \times (p+1)$ orthogonal matrix and R is a $(p+1) \times (p+1)$ upper triangular matrix.

The vector of least squares coefficients can be represented as

$$\hat{\beta} = \left(X^T X\right)^{-1} X^T y \tag{5}$$

Substitute X = QR.

$$\hat{\beta}$$

$$= \left((QR)^T (QR) \right)^{-1} (QR)^T y$$

$$= \left(Q^T QR^T R \right)^{-1} Q^T R y$$

Use the fact that Q is an orthogonal matrix $(Q^TQ = I)$.

$$\hat{\beta} = = (R^T R I)^{-1} Q^T R y$$
$$= (R^T R)^{-1} Q^T R y$$

Since $R^{-1}R = I$, we have

$$\hat{\beta} = R^{-1}Q^T y \tag{6}$$