A COMPARATIVE STUDY OF TECHNIQUES FOR ESTIMATION AND INFERENCE OF NONLINEAR STOCHASTIC TIME SERIES

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- I. Framing
- 2. Stochastic SIR Model
- 3. Hamiltonian HMCMC
- 4. Iterated Filtering 2
- 5. Forecasting Frameworks
- 6. S-maps & Seasonal Outbreaks
- 7. Spatiotemporal Epidemics
- 8. Parallelism & Future Directions

Framing

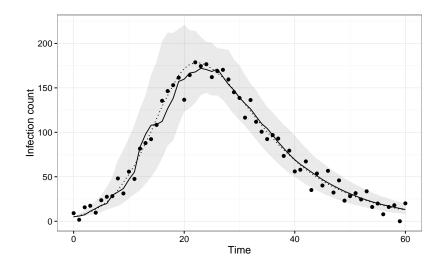


Stochastic SIR Model

Stochastic SIR model

$$\begin{aligned} \frac{\mathrm{dS}}{\mathrm{dt}} &= -\beta \mathrm{SI} \\ \frac{\mathrm{dI}}{\mathrm{dt}} &= \beta \mathrm{SI} - \gamma \mathrm{I} \\ \frac{\mathrm{dR}}{\mathrm{dt}} &= \gamma \mathrm{I} \\ \end{aligned}$$

$$eta_{\mathsf{t+I}} = \mathsf{exp}\left[eta_{\mathsf{t}} + \eta\left(ar{eta} - eta_{\mathsf{t}}
ight) + \mathcal{N}(\mathsf{0}, \sigma_{\mathsf{proc}})\right]$$



Hamiltonian MCMC

MCMC

Iteratively construct Markov chain to approximate posterior

- I. Choose starting parameter set
- 2. Generate N samples by
 - 2.1 Propose new sample
 - 2.2 Compute acceptance ratio
 - 2.3 Accept/reject sample

```
/* Select a starting point | Input : Initialize \theta^{(l)} | for i=2:N do | /* Sample | */ Sample | */ 3 | u \sim \mathcal{U}(0, l) | /* Evaluate acceptance ratio | */ L(\theta) = 0 | /* Step acceptance criterion | */ L(\theta) = 0 | /* Step acceptance criterion | */ L(\theta) = 0 | L(\theta) = 0
```

/* Samples from approximated posterior distribution */ Output: Chain of samples $(\theta^{(1)}, \theta^{(2)}, ..., \theta^{(N)})$

Hamiltonian Dynamics

Energy

Potential

$$K(r) = \frac{1}{2}r^{T}M^{-1}r$$

$$U(\theta) = -\log(\mathcal{L}(\theta)p(\theta))$$

$$H(\theta, r) = U(\theta) + K(r)$$

$$\frac{d\theta}{dt} = M^{-1}r$$

$$\frac{dr}{dt} = -\nabla U(\theta)$$

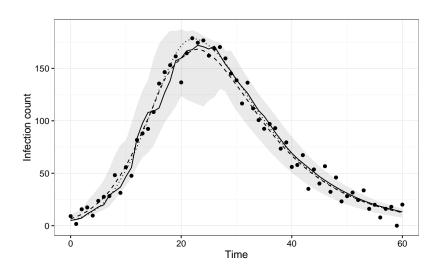
HMCMC algorithm

- I. Choose starting parameter set
- 2. Generate N samples by
 - 2.1 Resampling moments
 - Simulate Hamiltonian dynamics using Leapfrog integration
 - 2.3 Compute acceptance ratio
 - 2.4 Accept/reject sample

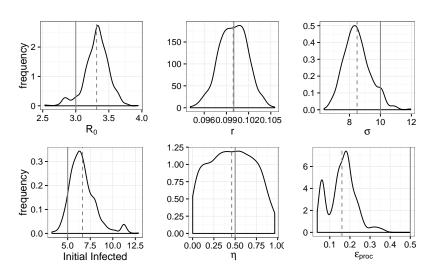
```
/* Select a starting point
     Input : Initialize \theta^{(l)}
\int for i = 2 \cdot N do
               /* Resample moments
               for i = 1 : n do
                         r(i) \leftarrow \mathcal{N}(0, 1)
7
               /* Leapfrog initialization
               \theta_0 \leftarrow \theta^{(i-1)}
5
               r_0 \leftarrow r - \nabla U(\theta_0) \cdot \varepsilon/2
               /* Leapfrog intermediate steps
               for j = l : L - l do
                         \theta_i \leftarrow \theta_{i-1} + M^{-1} r_{i-1} \cdot \varepsilon
                         r_i \leftarrow r_{i-1} - \nabla \cup (\theta_i) \cdot \varepsilon
               /* Leapfrog last steps
               \theta^* \leftarrow \theta_{1-1} + M^{-1}r_{1-1} \cdot \varepsilon
               r^* \leftarrow \nabla U(\theta_1) \cdot \varepsilon/2 - r_{1-1}
               /* Evaluate acceptance ratio
               r = \exp\left[H(\theta^{(i-1)}, r) - H(\theta^*, r^*)\right]
Ш
               I* Sample
               u \sim \mathcal{U}(0, I)
12
               /* Step acceptance criterion
               if u < min \{l, r\} then
13
                         \theta^{(i)} = \theta^*
14
15
               else
                         \theta^{(i)} = \theta^{(i-1)}
     /* Samples from approximated posterior
```

Output: Chain of samples $(\theta^{(1)}, \theta^{(2)}, \dots, \theta^{(N)})$

HMCMC state reconstruction sample



HMCMC parameter estimation kernels sample



Iterated Filtering 2

Basic Particle Filter

Iterative prediction-update cycle prunes particle cohort of poor parameter estimates

- I. Initialize particles with parameter sets
- 2. For each data point
 - 2.1 Evolve particle states
 - 2.2 Weight via likelihood
 - 2.3 Resample proportional to weights

```
/* Select a starting point
   Input : Observations D = y_1, y_2, ..., y_T, initial particle
               distribution Pn of size J
   /* Setup
Initialize particle cohort by sampling
      (p^{(1)}, p^{(2)}, \dots, p^{(J)}) from P_0
2 for t = 1 : T do
            /* Evolve
            for j = I:J do
                   X_{t}^{(j)} \leftarrow f_{I}(X_{t-1}^{(j)}, \theta^{(j)})
           /* Weight
            for j = I:J do
                  w^{(j)} \leftarrow P(y_t|X_t^{(j)}, \theta^{(j)}) = f_2(X_t^{(j)}, \theta^{(j)})
           /* Normalize
            for j = I:J do
               w^{(j)} \leftarrow w^{(j)} / \sum_{i}^{J} w^{(j)}
           I* Resample
            p^{(I:J)} \leftarrow sample(p^{(I:J)}, prob = w, replace =
   /* Samples from approximated posterior distribution */
    Output: Cohort of posterior samples
```

 $(\theta^{(1)}, \theta^{(2)}, \dots, \theta^{(J)})$

Iterated Filtering 2

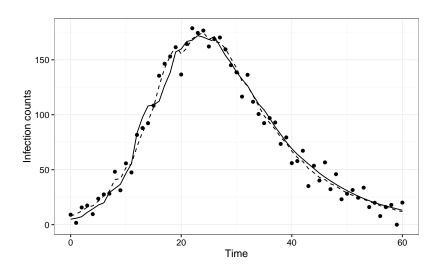
Treat parameter estimates as stochastic processes

Multiple passes through data (a la Data cloning)

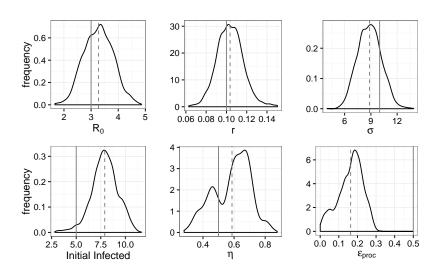
```
/* Setup
 Initialize particle cohort by sampling
       (p^{(1)}, p^{(2)}, \dots, p^{(J)}) from P_0
    /* Particle seeding distribution
    \Theta \leftarrow P_0
3 \text{ for } m = 1 \cdot M \text{ do}
              /* Pass perturbation
             for j = I:J do
                       p^{(j)} \sim h(\Theta^{(j)}, \sigma_m)
             for t = I : T do
                       for j = I:J do
                                 /* Iteration perturbation
                                 p^{(j)} \sim h(p^{(j)}, \sigma_m)
                               X_{t}^{(j)} \leftarrow f_{I}(X_{t-1}^{(j)}, \theta^{(j)})
                                w^{(j)} \leftarrow P(y_t|X_t^{(j)},\theta^{(j)}) =
IO
                                   f_2(X_h^{(j)}, \theta^{(j)})
                       /* Normalize
                       for j = I:J do
                          w^{(j)} \leftarrow w^{(j)} / \sum_{i=1}^{J} w^{(j)}
                       /* Resample
                       p^{(I:J)} \leftarrow sample(p^{(I:J)}, prob =
13
                          w. replace = true)
              /* Collect particles for next pass
14
              for i = I : J do
                       \Theta^{(j)} \leftarrow p^{(j)}
15
```

/* Samples from approximated posterior distribution */
Output: Cohort of posterior samples $(\theta^{(1)}, \theta^{(2)}, \dots, \theta^{(J)})$

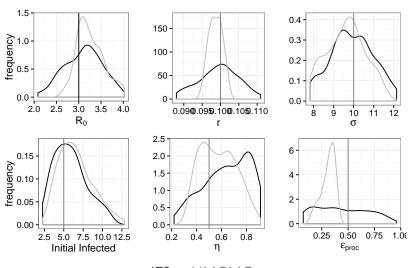
IF2 state reconstruction sample



IF2 parameter estimation kernels sample

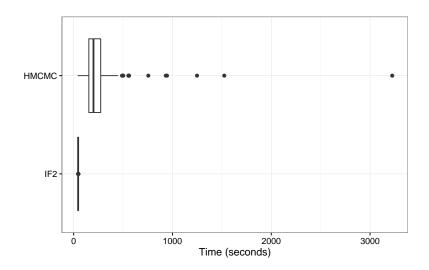


Mean parameter estimation distributions



IF2 HMCMC

Running times



Forecasting Frameworks

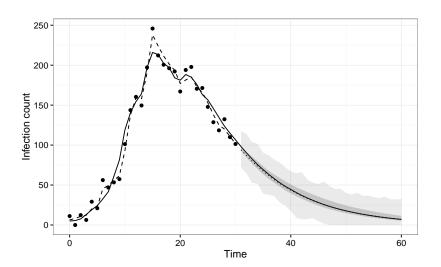


IF2 + Parametric bootstrapping + Forward simulation

- Provides additional samples from posterior distribution
- Forward simulation using states point estimate, posterior samples

 $F_1, F_2, ..., F_M$

IF2 parametric bootstrapping forecast

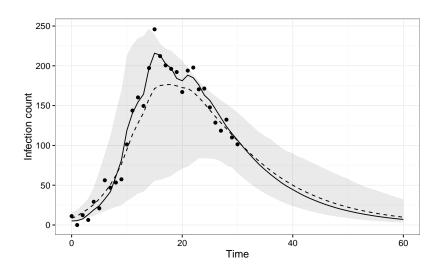


HMCMC + State reconstructions + Forward simulation

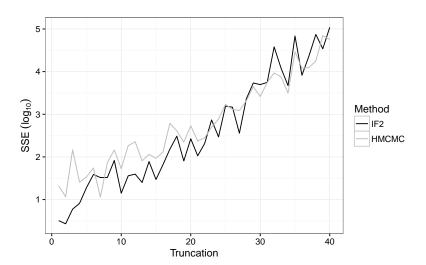
- Reconstruct states using latent process noise samples
- Simulate forward

```
/* Sample from the posterior | */ Input : Posterior samples \theta_1, \theta_2, ..., \theta_M | /* Reconstruct state estimates | */ 2 | for i=1:M do 3 | D_i \leftarrow S(\theta_i) | /* Simulate forward to produce forecasts */ 4 | for i=1:M do 5 | F_i \leftarrow S(D_i, \theta_i) | Output: Forecast trajectories F_1, F_2, ..., F_M
```

HMCMC parametric bootstrapping forecast



Forecast accuracy comparison

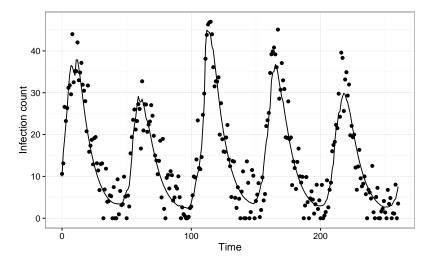


S-maps & Seasonal Outbreaks

Stochastic SIRS model

$$\begin{aligned} \frac{\mathrm{dS}}{\mathrm{dt}} &= -\beta \mathrm{SI} + \alpha \mathrm{R} \\ \frac{\mathrm{dI}}{\mathrm{dt}} &= \beta \mathrm{SI} - \gamma \mathrm{I} \\ \frac{\mathrm{dR}}{\mathrm{dt}} &= \gamma \mathrm{I} - \alpha \mathrm{R} \\ &+ \end{aligned}$$

$$eta_{\mathsf{t+I}} = \mathsf{exp}\left[eta_{\mathsf{t}} + \eta\left(ar{eta} - eta_{\mathsf{t}}
ight) + \mathcal{N}(\mathsf{0}, \sigma_{\mathsf{proc}})\right]$$

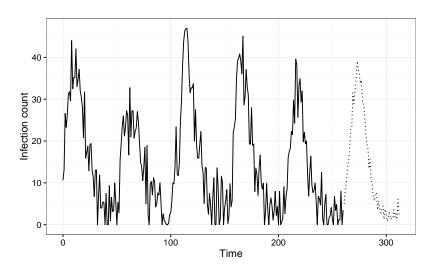


S-map

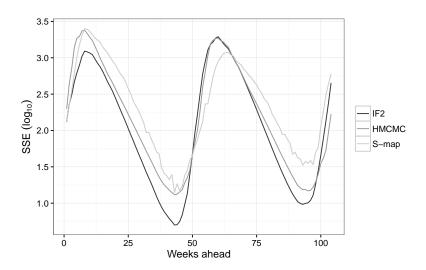
- Construct global mapping of time-lagged vectors (the library) to future states
- Weightings are used to penalize poorly matching library vectors

```
/* Select a starting point
   Input: Time series x_1, x_2, ..., x_T,
             embedding dimension E, distance
             penalization \theta, forecast length L,
             predictor vector xt
   /* Construct library {x<sub>i</sub>}
i \text{ for } i = E : T do
     | x_i = (x_i, x_{i-1}, ..., x_{i-F-1})
   /* Construct mapping from library vectors
3 for i = I : (T_E + I) do
          for j = I : E do
                 A(i,j) = w(||x_i - x_t||)x_i(j) \\
6 for i = I : (T_F + I) do
   b(i) = w(||x_i - x_t||)y_i
   /* Use SVD to solve the mapping system, Ac
8 SVD(Ac = b)
   I* Compute forecast
9 \hat{y_t} = \sum_{i=0}^{E} c_t(j) x_t(j)
   /* Forecasted value in time series
   Output: Forecast v+
```

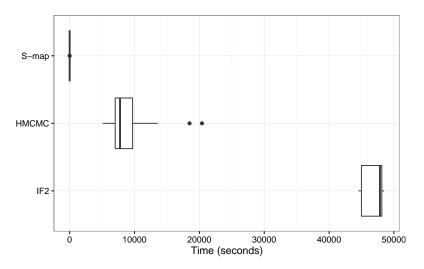
S-map forecast



SIRS model forecasting error



SIRS model forecasting runtimes



S-map: 316,000x faster than IF2, 61,800x faster than HMCMC

Spatiotemporal Epidemics

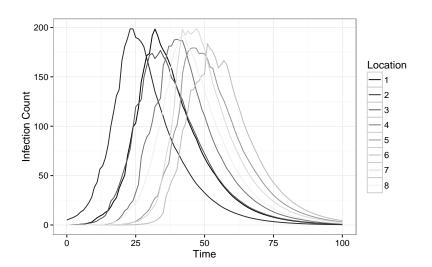
Stochastic Spatial SIR model

$$\begin{split} \frac{dS_i}{dt} &= -\left(I - \phi \frac{M}{M+I}\right) \beta_i S_i I_i - \frac{\phi}{M+I} S_i \sum_{j=1}^M \beta_{ij} I_j \\ \frac{dI_i}{dt} &= \left(I - \phi \frac{M}{M+I}\right) \beta_i S_i I_i + \frac{\phi}{M+I} S_i \sum_{j=1}^M \beta_{ij} I_j - \gamma I_i \\ \frac{dR_i}{dt} &= \gamma I_i \end{split}$$

$$+$$

$$eta_{i,t+l} = \exp\left[eta_{i,t} + \eta\left(ar{eta} - eta_{i,t}
ight) + \mathcal{N}(0,\sigma_{\mathsf{proc}})
ight]$$

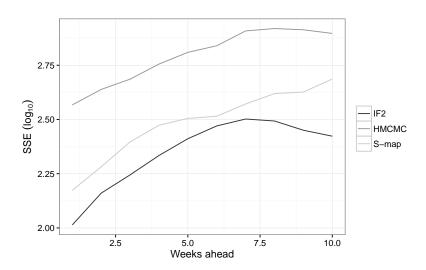
Stochastic spatial SIR model simulation (ring topology)



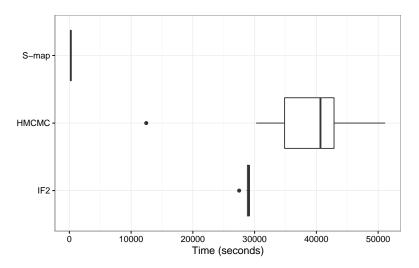
Dewdrop Regression

- "Stitching" together multiple short time series into single library
 - Requires scaling

Stochastic SIR model forecasting error



Spatial SIR model forecasting runtimes



Parallelism & Future Directions



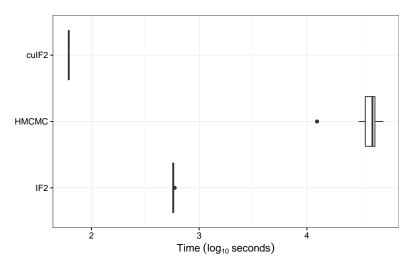
Conclusions

- IF2 produces superior forecasts in all scenarios
- S-mapping runs orders of magnitude faster than other methods

More than Moore

- Moore's Law is ceasing to hold
- Focus now on distributed computing
- MCMC-based methods are resistant to parallelization
 - Chain construction requires iterative dependence
- IF2 exhibits high parallel potential
 - Preliminary CUDA (GPU-accelerated) implementation - cuIF2

Spatial SIR model fitting times



cuIF2: 9.33x faster than IF2, 6I7x faster than HMCMC



THE BEST THESIS DEFENSE IS A GOOD THESIS OFFENSE.