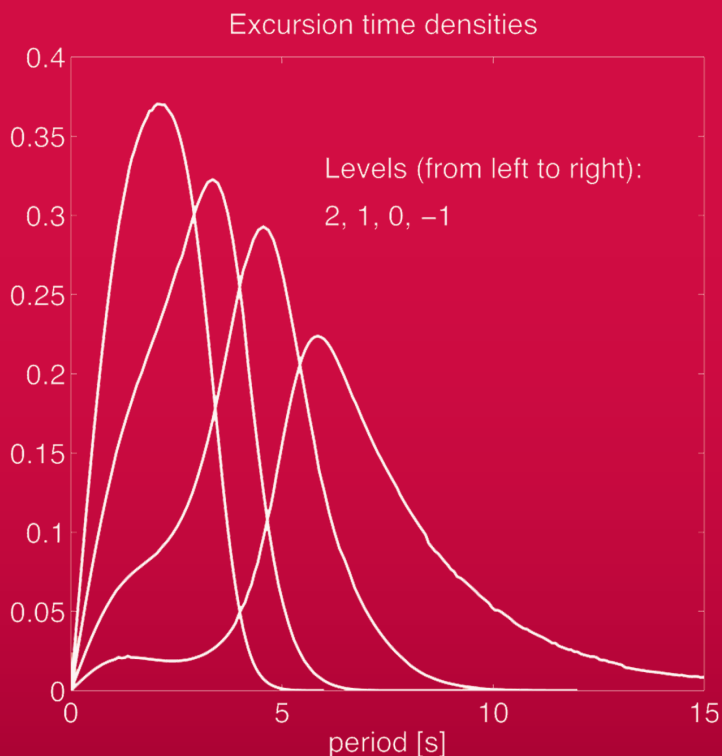


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Stationary Stochastic Processes

Theory and Applications



Georg Lindgren



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Statistics is a science in my opinion, and it is no more a branch of mathematics than are physics, chemistry and economics; for if its methods fail the test of experience – not the test of logic – they are discarded.

John W. Tukey (1953)

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Preface

This book has grown out of my own experiences as teacher and researcher at a department in mathematical statistics with responsibilities both to an engineering and a science community. The spirit of the text reflects those double responsibilities.

The background

The book *Stationary and Related Stochastic Processes* [35] appeared in 1967. Written by Harald Cramér and M.R. Leadbetter, it drastically changed the life of PhD students in mathematical statistics with an interest in stochastic processes and their applications, as well as that of students in many other fields of science and engineering. Through that book, they got access to tools and results for stationary stochastic processes that until then had been available only in rather advanced mathematical textbooks, or through specialized statistical journals. The impact of the book can be judged from the fact that still, after almost fifty years, it is a standard reference to stationary processes in PhD theses and research articles.

Even if many of the more specialized results in the Cramér-Leadbetter book have been superseded by more general results, and simpler proofs have been found for some of the statements, the general attitude in the book makes it enjoyable reading both for the student and for the teacher. It will remain a definite source of reference for many standard results on sample function and crossings properties of continuous time processes, in particular in the Gaussian case. Unfortunately, the book only appeared in a first edition, and it was out of print for many years. The Dover reprint from 2004 filled that gap.

Another book, that at its time created a link between the mathematical theory of stationary processes and their engineering use, is the book *An Introduction to the Theory of Stationary Random Functions*, by A.M. Yaglom from 1962, [124]. The idea of bridging the gap between a strict mathematical treatment and the many practically useful results in the theory of stationary processes is even more explicitly expressed in the later, two-volume work by the same author, *Correlation Theory of Stationary and Related Random*

Functions from 1987, [125]. That work, with its extensive literature survey, presents a rather complete overview of correlation and spectrum theory for stationary processes at its time, with lots of extensions, references, and historical notices.

The third influential book is *Probability*, by Leo Breiman [19], also long out of print until it was reprinted in 1992. In the preface of that book, the author mentions the two-armed character of probability theory. The right arm is rigorous mathematics, based on measure theory; the left arm sets probability in action with gambling, motions of physical particles, and, now, almost fifty years later, remote sensing interpretation and systems control.

This book is for the left handed student, who wants to learn what the right handed side of stationary process theory has to offer, without going very far into the mathematical details! It is my hope that also the right handed probabilist could appreciate some of the fascination that lies in the potential of probabilistic thinking in engineering and science. The Preface to Yaglom's [125], Volume II, expresses these thoughts in a judicious way.

This text has grown out from a series of PhD courses on *Stationary stochastic processes*, which have been given at the Department of Mathematical Statistics at Lund University. Previous versions of the text have also been used in Stockholm, Trondheim, and Umeå. The audiences have consisted of Masters and PhD students in Mathematics and Statistics, but also to a large extent of PhD students in Automatic control, Electrical engineering and Signal processing, Image analysis, and Economics. As far as possible, I have tried to incorporate the links between these concrete topics and the theoretical tools developed.

I am grateful to three decades of PhD students who have followed the courses in Lund and elsewhere. Their comments and suggestions have removed many unclear and obscure points and statements. Their demands have also had the result that the text has grown, with new topics along with the increasing interest in advanced statistical tools in applications.

The text

The text is intended for a "second course" in stationary processes, and the material has been chosen to give a fairly broad overview of the theory behind widely scattered applications in engineering and science. The reader I have in mind has typically some experience with stochastic processes and has felt an urge to know more about "what it really is" and "why."

The emphasis is on second order properties and Fourier methods. Neither time series analysis nor Markov assumptions play any prominent role.

Since the days of the three texts mentioned in the introduction, applications of second order stationary processes have expanded enormously. The basic theory remains, but the applications, in particular new measurement techniques, have inspired much new theory. This is certainly true for spatial and spatio-temporal models, counting processes in space and time, extremes and crossings, and, not least, on computational aspects.

The aim is to provide a reasonably condensed presentation of sample function properties, spectral representations for stationary processes, and fields, including a portion of stationary point processes. One of the most important applications of second order process theory is found in linear filtering theory. The material on linear filters is divided into two chapters, one with a general technique of how to use the spectral representations and one on three important applications.

One of the most difficult topics to explain in elementary courses on stationary processes is ergodicity. I have tried to give both a mathematical and a statistical view with the main aim to couple mathematics and probability to statistics and measurements.

A chapter on crossing problems and extremes with an introduction to some of its applications has grown out from my own research interests.

Some knowledge of the mathematical foundations of probability helps while reading the text; I have included some of it in a probability appendix, together with the existence and basic convergence properties, as well as some Hilbert space concepts. There is also an appendix on how to simulate stationary stochastic processes by spectral methods and the FFT algorithm.

Some suggestions for reading

The reader who wants to design a course on parts of the text may make use of the following selections:

- For a general and concrete level:
 - Section 1.1 (Section 1.2), Sections 1.3–1.5
 - Section 2.1
 - Sections 3.2–3.4
 - Chapter 4 except Section 4.4
 - Section 5.2
 - Sections 7.1.1, 7.2.1
 - Section 8.1.1
- For specialized and concrete levels:
 - Section 3.5

- Section 4.4
- Sections 5.1 and 5.3
- Chapter 7
- Chapter 8
- Appendix B
- For a more abstract level one could add:
 - Section 1.2, Appendix A, plus a book on general probability theory
 - Sections 2.2–2.5
 - Section 3.1
 - Chapter 6

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During the preparation of the final text I have had fruitful discussions with David Bolin, Finn Lindgren, Johan Lindström, Jimmy Olsson, Krzysztof Podgórski, Magnus Wiktorsson; thank you!

But it all goes back to Harald Cramér and Ross Leadbetter!

Georg Lindgren
Lund

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List of notations

Roman letters

A, A_k	random amplitudes
$A_s(\mathbf{v})$	set of functions $y(s)$ with $y(s_j) \leq v_j$
$A_{xy}(\omega) = f_{xy}(\omega) $	cross-amplitude spectrum
$A(z)$	generating polynomial
$A \Delta B$	$(A - B) \cup (B - A)$, difference between sets
a_T, b_T	inverse scale, and location, parameter in asymptotic extreme value distribution
$BB_T(t)$	Brownian bridge
$\mathcal{B}, \mathcal{B}_n, \mathcal{B}_\infty, \mathcal{B}_T$	Borel fields in $\mathbb{R}, \mathbb{R}^n, \mathbb{R}^\infty, \mathbb{R}^T$
C	generic constant
$C(x_1, x_2)$	$E((x_1 - m_1)\overline{(x_2 - m_2)})$, covariance of complex variables
$C(z)$	generating polynomial
\mathbb{C}	space of continuous functions
$c_{xy}(\omega)$	co-spectrum density
$D(u_n), D'(u_n)$	mixing and non-clustering conditions in asymptotic extreme value theory
$d(\mathbf{s}, \mathbf{t})$	$\sqrt{E(x(\mathbf{s}) - x(\mathbf{t}) ^2)}$, distance measure
$dZ(\omega)$	differential spectral component
$\tilde{dZ}(\omega) = \frac{dZ(\omega)}{f(\omega)}$	normalized differential spectral component
$E(x y), E(x \mathcal{A})$	conditional expectation given y or given \mathcal{A}
$e(\mathbf{s}, \mathbf{t})$	Euclidian distance
$F_{t_1 \dots t_n}(b_1, \dots, b_n)$	n -dimensional distribution function for $x(t_1), \dots, x(t_n)$
$\mathbf{F} = \{F_{\mathbf{t}^n}\}$	family of finite-dimensional distribution functions
$F(\omega), F_x(\omega)$	spectral distribution function
F^{ac}, F^d, F^s	absolutely continuous, discrete, and singular part of distribution function
$\tilde{F}(\omega)$	$(F(\omega) + F(\omega - 0))/2$, midpoint at spectral jump
$F_{xy}(\omega), f_{xy}(\omega)$	cross-spectral distribution and density functions
$\mathcal{F}_0, \mathcal{F}$	field and σ -field of events

f	friction coefficient
$f_{x y=v}(u)$	conditional density function
$f(\omega), f_x(\omega)$	spectral density functions for stationary process
$f_x^{(d)}(\omega)$	spectral density after sampling with sampling distance d
$f_x^{\text{time}}(\omega)$	time wave spectrum
$f_x^{\text{space}}(\kappa)$	space wave spectrum
$f_{PM}(\omega), f_J(\omega)$	Pierson-Moskowitz and JONSWAP spectral densities
$f(\omega, \theta)$	$f_x^{\text{time}}(\omega)g(\omega, \theta)$, spectral density for random field with directional spreading
$f_c(\omega)$	$\frac{\lambda}{2\pi} + f(\omega)$, Bartlett spectrum for counting process
$f_n^{\text{per}}(\omega)$	periodogram computed from data
g	constant of gravity
$(g, h)_{\mathcal{H}(F)}$	$\int g(\omega)\overline{h(\omega)}dF(\omega)$, inner product in Hilbert space $\mathcal{H}(F)$
$g(\omega)$	$\int e^{-i\omega u}h(u)du$ frequency response function
$g(\omega, \theta)$	directional spreading function
H	Hurst parameter, index of fractional Brownian motion
$H(\varepsilon)$	entropy
h	water depth
$\mathcal{H}, \mathcal{H}(\Omega)$	Hilbert space of random variable with finite variance
$\mathcal{H}(F)$	Hilbert space of functions with $\int g(\omega) ^2 dF(\omega) < \infty$
$\mathcal{H}(x)$	$\mathcal{S}(x(s); s \in T)$, Hilbert space generated by $x(t), t \in T$
$\mathcal{H}(x, t)$	$\mathcal{S}(x(s); s \leq t)$, Hilbert space generated by $x(s), s \leq t$
$h(u)$	impulse response function
\mathcal{I}	field of unions of a finite number of intervals, or σ -field of invariant sets
$J_m(t)$	Bessel function of the first kind of order m
$K_\nu(t)$	Bessel function of the second kind of order ν
\mathcal{M}_a^b	σ -field generated by $x(t), a \leq t \leq b$
N	Avogadro's number
$N(t)$	Poisson process
$\tilde{N}(I)$	normalized point process of crossings
$N_I(x, u), N_I^+(x, u)$	number of u -crossings/upcrossings by $x(t), t \in I$
\mathbb{N}	the set of natural numbers
P	general probability measure
$P^{hw}(A)$	horizontal window conditioned probability after level crossing

$P^u(A)$	distribution of Slepian model for process after u -upcrossing
$P_j^{\max}(A)$	distributions for Slepian models after local maximum, $j = 1, 2$
P_x, P_y, \dots	probability measure for processes $x(t), y(t), \dots$
$P_x^{(n)}, P_y^{(n)}, \dots$	n -dimensional probability measure for $x(t), y(t), \dots$
$p^u(z)$	Rayleigh density for derivative at level upcrossing in Gaussian process
$q_{xy}(\omega)$	quadrature spectrum
R	Boltzmann's constant
$R(t)$	envelope process
$\mathbb{R}, \mathbb{R}^n, \mathbb{R}^\infty, \mathbb{R}^T$	real line, n -dimensional real space, space of infinite real sequences, space of real function defined on T
$r(s, t)$	$C(x(s), x(t))$ covariance function
$r(t)$	$C(x(s+t), x(s))$ covariance function for stationary process or field
$r_{xy}(t)$	cross-covariance between $x(s+t)$ and $y(s)$
$r_N^*(t)$	covariance estimate based on the periodogram
$r_{\Delta_1}(s, t), r_{\Delta_2}(s, t)$	covariance functions for Gaussian residual in Slepian models after local maximum
$r_\kappa(s, t)$	covariance function for Gaussian residual in Slepian model after level upcrossing
$\mathcal{S}(L)$	space spanned by random variables in L
$\mathcal{S}(x, \varepsilon)$	ε -surrounding around point or function x
$\text{sinc}(x) = \frac{\sin x}{x}$	sinc function
T	parameter space, often \mathbb{R} or \mathbb{R}^+ , or absolute temperature
$T(\omega)$	transformation of a probability space
$T_2 = 2\pi\sqrt{\omega_0/\omega_2}$	mean period
$t_n^{(k)} = k/2^n$	dyadic number
$(u, v)_{\mathcal{H}(x)}$	$E(u\bar{v})$, inner product in Hilbert space $\mathcal{H}(x)$
$V_f(I)$	total variation of $f(t), t \in I$
$w(t)$	Wiener process
$w'(t)$	Gaussian white noise
$w(\tau)$	covariance intensity function
$w_c(\tau) = \lambda \delta_0(\tau) + w(\tau)$	complete covariance intensity function
$\hat{x}(t)$	Hilbert transform of $\{x(t), t \in \mathbb{R}\}$
\mathbb{Z}	the set of integer numbers
$Z(\omega), Z_j(\omega)$	spectral processes

Greek letters

$\alpha = \omega_2 / \sqrt{\omega_0 \omega_4}$	spectral width parameter
γ	peakedness parameter in JONSWAP spectrum
$\gamma(\mathbf{u}, \mathbf{v})$	semi-variogram
$\Delta F_\omega, \Delta F_k$	spectral distribution jump at ω and ω_k
$\Delta Z_\omega, \Delta Z_k$	spectral process jump at ω and ω_k
δ_{jk}	Kronecker delta
$\delta(t), \delta_0(t), \delta_\tau(t)$	Dirac delta functions
$\varepsilon = \sqrt{1 - \alpha^2}$	spectral width parameter
$\zeta = c / (2\sqrt{mk})$	relative damping in linear oscillator
θ	wave direction in random field
κ	space frequency, wave number
κ_k	spectral moment in space
$\kappa_{xy}^2(\omega)$	squared coherence spectrum
λ	intensity in point process
λ_k	eigenvalue to a covariance operator
$\mu^+ = \frac{1}{\pi} \sqrt{\omega_2 / \omega_0}$	mean frequency
$\mu(u)$	mean number of u -crossings per time unit
$\mu^+(u)$	mean number of u -upcrossings per time unit
$\mu_{\min} = \mu_{\max}$	mean number of local extremes per time unit
$\xi_u(t)$	Slepian model for process after u -upcrossing
$\xi_1^{\max}(t), \xi_2^{\max}(t)$	Slepian models for process after local maximum
$\rho^2 = \omega_1^2 / \omega_0 \omega_2$	spectral width parameter
$\rho_{xy}(\omega)$	in-phase correlation in vector processes
$\tilde{\rho}_{xy}(\omega)$	out-of-phase correlation in vector processes
Σ	covariance matrix
$\sigma^2 / 2\pi$	constant spectral density of white noise
$\sigma(x)$	σ -field generated by random variable
$\Phi_{xy}(\omega)$	phase spectrum
ϕ, ϕ_k	random phases
$\phi_k(t)$	eigenfunction to a covariance operator
ϕ, Φ	standard normal density and distribution functions
$\varphi(s), \varphi(x y)$	characteristic function and conditional expectation
χ_A	indicator function for event A
ψ	phase parameter in linear oscillator
$\psi_\varepsilon(t)$	characteristic function
Ω	sample space
(Ω, \mathcal{F})	measurable space, sample space with σ -field
(Ω, \mathcal{F}, P)	probability space

ω	frequency or element in sample space
$\omega_0 = \sqrt{k/m}$	resonance frequency in linear oscillator
ω_2	$-r''(0) = V(x(t))$, second spectral moment
ω_k	discrete frequency, or spectral moment $\int \omega ^k dF(\omega)$
$\omega_{\alpha} = \prod_k \omega_k^{\alpha_k}$	multiple spectral moment
$\omega_x(h)$	continuity modulus of function x
$\omega_{x,d}(h)$	continuity modulus of function x in d -distance

Special symbols

$\xrightarrow{a.s.}$	almost sure convergence, with probability one
$\xrightarrow{q.m.}$	quadratic mean convergence
\xrightarrow{P}	convergence in probability
$\xrightarrow{\mathcal{L}}$	convergence in distribution
$x \stackrel{\mathcal{L}}{=} y$	x and y have the same distribution

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Some probability and process background

This introductory chapter gives examples of some of the main concepts in stationary processes theory that will be dealt with in detail in later chapters. The intention here is not to dwell on the precise mathematical properties, but to point to the statistical content and interpretation of the mathematical constructions, and on the driving ideas. The mathematical details are instead presented in Appendix A, and refreshed when needed.

First of all, we deal with the finite-dimensional distribution functions (d.f.), which uniquely define probabilities for a sufficiently rich family of events, namely events that can be identified through process values at discrete sets of times. In particular, they allow us to find conditions for sample function continuity and differentiability and other properties that are essential in statistical model building. This is the topic of Chapter 2.

Another central topic in the book is the covariance properties of stationary processes and their spectral representations, dealt with in Chapter 3. These also form the tools in the chapters on filter manipulations of stationary processes, Chapters 4 and 5. Prediction and limiting behavior are the topics of Chapter 6. Extensions to multivariate models and stochastic fields are treated in Chapter 7, and the final Chapter 8 deals with techniques and important applications of extremes and crossings.

This chapter ends with four examples of path breaking innovative use of stochastic models in physics and engineering, innovations which helped to shape the presented tools and techniques. The examples are the Wiener process as a model for Brownian motion, stationary processes as models for noise in electronic communication systems, stochastic wave models used in ocean engineering, and finally the link between statistical inference and stochastic processes as a basis for signal processing.

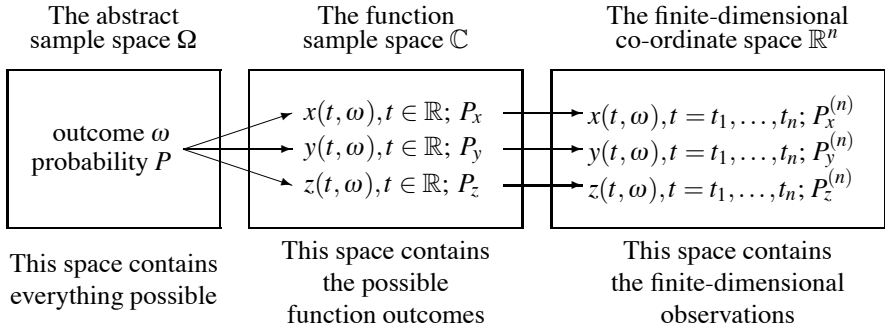


Figure 1.1: *Overview of the three types of worlds in which our processes live.*

1.1 Sample space, sample function, and observables

Stochastic processes are often called random functions; these two notions put emphasis on two different aspects of the theory, namely,

- stochastic processes as families of infinitely many random variables on the same sample space, usually equipped with a fixed probability measure,
- stochastic processes as a means to assign probabilities to sets of functions, for example some specified sets of continuous functions, or sets of piecewise constant functions with unit jumps.

These two aspects of stochastic processes can be illustrated as in Figure 1.1, corresponding to an experiment where the outcomes are continuous functions.

The figure illustrates the three levels of abstraction and observability for a random experiment. To be concrete, think of an experiment controlling the steering of a ship. The sample space Ω can be envisioned to describe the conditions that the ship will meet on the ocean, the wind and wave climate over the lifetime of the ship, the actual weather at a certain trip, etc. The function sample space \mathbb{C} contains the different characteristics that physically characterize the movements of the ship and the steering process, while the co-ordinate space \mathbb{R}^n describes the finite number of measurements that can be taken on these signals.

The general sample space Ω is an abstract set that contains all the possible outcomes of the experiment that can conceivably happen – and it may contain more. A probability measure P is defined on Ω that assigns probabilities to all interesting subsets – we need only one single probability measure to describe our whole world. We may choose to measure many different characteristics, with different statistical distributions, but they are all uniquely defined by the single underlying “world” probability measure P .

During the experiment one can record the time evolution of a number of things, such as rudder angle, which we call $\{x(t), t \in \mathbb{R}\}$, ship head angle, $\{y(t), t \in \mathbb{R}\}$, and roll angle $\{z(t), t \in \mathbb{R}\}$. Each observed function is an observation of a continuous random process. In the figure, the randomness is indicated by the dependence of the experiment outcome ω . The distributions of the different processes are P_x, P_y, P_z – we need one probability measure for each of the phenomena we have chosen to observe.¹

In practice, the continuous functions are sampled in discrete time steps, $t = t_1, \dots, t_n$, resulting in an observation vector, (x_1, \dots, x_n) , with an n -dimensional distribution, $P_x^{(n)}$, etc. This familiar type is illustrated in the third box in Figure 1.1.

Since we do not always want to specify a finite value for n , the natural mathematical model for the practical situation is to replace the middle box, the sample space \mathbb{C} of continuous functions, by the sample space \mathbb{R}^∞ of infinite sequences of real numbers (x_0, x_1, \dots) . This is close, as we shall see later, really very close, to the finite-dimensional space \mathbb{R}^n , and mathematically not much more complicated.

Taking the set \mathbb{C} of continuous functions as a sample space and assigning probabilities P_x , etc., on it, is not as innocent as it may sound from the description above. Chapter 2 deals with conditions that guarantee that a stochastic process is continuous, i.e., has continuous sample functions. In fact, these conditions are all on the finite-dimensional distributions.

SUMMARY: *The abstract sample space Ω for an experiment contains everything that can happen and is therefore very complex and detailed. Each outcome $\omega \in \Omega$ is unique, and we need only one comprehensive probability measure P to describe every outcome of the experiment.*

An experiment is a way to “observe the world.” The function (sequence) sample space $\mathbb{C}(\mathbb{R}^\infty)$ is simple. It can be used as sample space for a specified experiment for which the result is a function or a sequence of numbers. We have to define a specific probability measure for each experiment.

¹The symbol ω is here used to represent the elementary experimental outcome, a practice that is standard in probability theory. In most parts of this book, ω will stand for (angular) frequency; no confusion should arise from this.

1.2 Random variables and stochastic processes

1.2.1 Probability space and random variables

A *probability space* (Ω, \mathcal{F}, P) is a triple consisting of: a sample space Ω of outcomes, a family \mathcal{F} of events, i.e., subsets of Ω , and a probability measure P , that assigns probabilities to the events in \mathcal{F} . The family \mathcal{F} may be chosen in many different ways; it may contain just a few subsets, or it may contain many. It is even possible to let \mathcal{F} contain *all* subsets of Ω .² The obligatory properties of the family \mathcal{F} are that it contains Ω and is closed under formation of complements, unions, and intersections of countably many of its elements. A family of subsets with these properties is called a σ -field.

A probability measure assigns probabilities to the events in \mathcal{F} in such a way that the basic rules for adding probabilities, usually called the *Kolmogorov's axioms of probabilities*, are fulfilled. The exact definitions of σ -field and the summation rules are formulated in Definition A.1 in Appendix A.1.

The link between the sample space Ω and the observable world is the random variable, a real valued function $\Omega \ni \omega \mapsto x(\omega) \in \mathbb{R}$. In order that x be called random variable it is essential that the *distribution* of x is well defined, i.e., that the *event* " $x \leq b$ " has a probability for each value of $b \in \mathbb{R}$. This means that the set of outcomes $\{\omega \in \Omega; x(\omega) \leq b\}$ must be one of the sets in \mathcal{F} , those which have been assigned probability by P . Functions that fulfill this requirement are called *measurable*, or, more precisely, *Borel measurable*.

It is of course possible to define many random variables, x_1, \dots, x_n , on the same probability space (Ω, \mathcal{F}, P) . Since P assigns a probability to each of the individual events

$$"x_1 \leq b_1", \dots, "x_n \leq b_n"$$

it also assigns a probability to the joint occurrence of them all; this is by the property of the σ -field \mathcal{F} that it contains intersections. The result is a multivariate random variable, $\mathbf{x} = \{x_k\}_{k=1}^n$ with an n -dimensional distribution.

In Appendix A.1.1 we explain what more events are given probabilities in this way. These events are the *Borel sets* in \mathbb{R}^n and the family is called the *Borel σ -field*, denoted \mathcal{B}_n . Examples of such events are those defined by two-sided inequalities, like the half-open rectangle $a_1 < x_1 \leq b_1, a_2 < x_2 \leq b_2$, unions of rectangles, complements of unions, and many more. The Borel σ -

²If \mathcal{F} contains all subsets of Ω every probability measure has to be discrete.

field is a very rich family of events, which all get probabilities defined by the finite-dimensional distribution functions.

1.2.2 Stochastic processes and their finite-dimensional distributions

We are now ready to define stochastic processes in general. There is no difficulty in considering an infinite sequence of random variables $\mathbf{x} = \{x_k\}_{k=1}^{\infty}$, on a single probability space (Ω, \mathcal{F}, P) . One can even consider more than countably many random variables at the same time. Then it is common to index the variables by the letter t instead of k , and let the *parameter* t take values in a general *parameter space* T . Usually one thinks of t as “time” and T as the whole or part of the real line.

Thus, we can consider a family of functions, $\{x(t, \omega) \in \mathbb{R}\}_{t \in T}$, where each $x(t) = x(t, \cdot)$ is a random variable, i.e., a measurable function from Ω to \mathbb{R} . Hence, it has a distribution with a distribution function on \mathbb{R} , which we denote $F(\cdot; t)$, i.e.,

$$F(b; t) = \text{Prob}(x(t) \leq b) = P(\{\omega; x(t, \omega) \leq b\}).$$

Taking several variables, observed at times t_1, \dots, t_n , we get an n -variate random variable $(x(t_1), \dots, x(t_n))$ with an n -variate distribution in \mathbb{R}^n ,

$$\begin{aligned} F_{t_1 \dots t_n}(b_1, \dots, b_n) &= \text{Prob}(x(t_1) \leq b_1, \dots, x(t_n) \leq b_n) \\ &= P(\cap_1^n \{\omega; x(t_k, \omega) \leq b_k\}). \end{aligned} \quad (1.1)$$

Write $F_{\mathbf{t}^n}$ for the n -dimensional distribution function of $(x(t_1), \dots, x(t_n))$. We summarize the terminology in a formal, but simple, definition.

Definition 1.1. *Let T be a discrete or continuous real parameter set. A stochastic process $\{x(t), t \in T\}$ indexed by the parameter $t \in T$ is a family of random variables $x(t)$ defined on one and the same probability space (Ω, \mathcal{F}, P) . In other words, a stochastic process is a function*

$$T \times \Omega \ni (t, \omega) \mapsto x(t, \omega) \in \mathbb{R},$$

such that, for fixed $t = t_0$, $x(t_0, \cdot)$ is a random variable, i.e., a Borel measurable function, $\Omega \ni \omega \mapsto x(t_0, \omega) \in \mathbb{R}$, and, for fixed $\omega = \omega_0$, $x(\cdot, \omega_0)$ is a function $T \ni t \mapsto x(t, \omega_0) \in \mathbb{R}$, called a sample path or sample function.

We now have the following concepts at our disposal in the three scenes in Figure 1.1:

sample space	events	probability
abstract space: Ω	σ -field \mathcal{F}	P
space of continuous functions: \mathcal{C}	$\boxed{?}$	$\boxed{?}$
real sequences: \mathbb{R}^∞	$\boxed{?}$	$\boxed{?}$
real vectors: \mathbb{R}^n	Borel sets: \mathcal{B}_n	$P_{\mathbf{t}^n}$ from n -dimensional d.f.'s $F_{\mathbf{t}^n}$
real line: \mathbb{R}	Borel sets: \mathcal{B}	P from a d.f. F

In the table, the $\boxed{?}$ indicate what we yet have to define – or even show existence of – to reach beyond the elementary probability theory, and into the world of stochastic processes.

1.2.3 The distribution of random sequences and functions

In the previous section we discussed finite-dimensional distributions for one, two, and (finitely) many random variables, how these distributions are defined by the finite-dimensional distribution functions, and what finite-dimensional events get their probabilities defined.

Our aim now is to find the events in \mathbb{R}^∞ (= all real sequences) and \mathbb{R}^T (= all real functions defined on T), and see how one can define a probability measure for these. When this is done, it is legitimate to talk about *the distribution of an infinite random sequence or function*, and of the probability that an experiment results in a sample function that satisfies some rather general requirement, e.g., that its maximum is less than a given constant.

It was this step, from probabilities for one-dimensional or n -dimensional real sets and events, to probabilities in \mathbb{R}^∞ and \mathbb{R}^T , that was made axiomatic in A.N. Kolmogorov's celebrated *Grundbegriffe der Wahrscheinlichkeitsrechnung* from 1933 [71].

Generalized rectangles, intervals, and the Borel field

The basic requirement on the events in \mathbb{R}^∞ or \mathbb{R}^T is that they should not be simpler than the events in the finite-dimensional spaces \mathbb{R}^n , which means that if an event $B_n \in \mathcal{B}_n$ in \mathbb{R}^∞ is expressed by means of a finite set of random variables x_1, \dots, x_n , then it should be an event also in the space \mathbb{R}^∞ . Now, such an event can be written

$$\begin{aligned} \{\mathbf{x} = (x_1, x_2, \dots) \in \mathbb{R}^\infty; (x_1, x_2, \dots, x_n) \in B_n\} &= B_n \times \mathbb{R} \times \mathbb{R} \times \dots \\ &= B_n \times \mathbb{R}^\infty. \end{aligned}$$

A set of this form is called a *generalized rectangle* in \mathbb{R}^∞ .

Similarly, we define a generalized rectangle in \mathbb{R}^T as a set defined by a condition on any finite set of $x(t)$ -values,

$$\{\mathbf{x} \in \mathbb{R}^T; (x(t_1), x(t_2), \dots, x(t_n)) \in B_n\}.$$

Hence, we have to require that the σ -field of events contains at least all generalized rectangles. The natural event field is exactly the smallest σ -field which contains all such sets; cf. Example A.2 in Appendix A.1. This σ -field is denoted \mathcal{B}_∞ and \mathcal{B}_T , respectively, and is, naturally, called the *Borel field* in \mathbb{R}^∞ or \mathbb{R}^T . Symbolically, we can write, for sequences,

$$\mathcal{B}_\infty = \sigma(\cup_{n=1}^\infty (\mathcal{B}_n \times \mathbb{R}^\infty)).$$

A particularly simple form of rectangles is the *intervals*. These are, in \mathbb{R}^∞ , the sets of form

$$I = (a_1, b_1] \times (a_2, b_2] \times \dots \times (a_n, b_n] \times \mathbb{R}^\infty,$$

where each $(a_j, b_j]$ is a real half-open interval. Thus, the half-open intervals in \mathbb{R}^∞ and \mathbb{R}^T are defined by inequalities like

$$a_1 < x_1 \leq b_1, a_2 < x_2 \leq b_2, \dots, a_n < x_n \leq b_n. \quad (1.2)$$

$$a_1 < x(t_1) \leq b_1, a_2 < x(t_2) \leq b_2, \dots, a_n < x(t_n) \leq b_n. \quad (1.3)$$

Sets which are unions of a finite number of intervals are important; they form a field, which we denote \mathcal{I} . The σ -field generated by \mathcal{I} is exactly \mathcal{B}_∞ , and \mathcal{B}_T , respectively, in \mathbb{R}^∞ and \mathbb{R}^T . The reader should draw a picture of some sets in the field \mathcal{I} in \mathbb{R}^2 .

The sample spaces we are mostly going to work with are \mathbb{R}^∞ and \mathbb{R}^T , infinite real sequences, and real functions defined on T , and the Borel σ -fields of events to be given probabilities are \mathcal{B}_∞ and \mathcal{B}_T , respectively, generated by the finite-dimensional intervals.

These spaces are important since they provide us with a natural world of objects for stochastic modeling.

Definition 1.2 (The co-ordinate process). *The simplest process on $(\mathbb{R}^\infty, \mathcal{B}_\infty)$ is the co-ordinate process, also known as “the canonical process,” $\{x_n, n \in \mathbb{N}\}$ defined by $x_k(\omega) = \omega_k$, on the outcome $\omega = (\omega_1, \omega_2, \dots)$.*

Similarly, the simplest process on $(\mathbb{R}^T, \mathcal{B}_T)$ is the co-ordinate process $x(t, \omega) = \omega(t)$ for $\omega \in \mathbb{R}^T$.

Finding a probability measure

The next big step is to assign probabilities to the Borel sets in \mathcal{B}_∞ and \mathcal{B}_T , and this can be done in either of two ways, from the abstract side or from the observable, finite-dimensional side. Going back to Figure 1.1 we shall fill in the two remaining question marks in the following scheme:

$$(\Omega, \mathcal{F}, P) \xrightarrow{x} \left\{ \begin{array}{l} (\mathbb{R}^\infty, \mathcal{B}_\infty, ?) \\ (\mathbb{R}^T, \mathcal{B}_T, ?) \end{array} \right\} \leftarrow (\mathbb{R}^n, \mathcal{B}_n, P_x^{(n)}), n = 1, 2, \dots$$

The mapping \xrightarrow{x} symbolizes that we have already defined a probability P on some measurable space (Ω, \mathcal{F}) and want to define probabilities on a sample function space for a stochastic process x with discrete or continuous parameter. The mapping \leftarrow symbolizes that we shall define the probabilities from a family of finite-dimensional distributions on $(\mathbb{R}^n, \mathcal{B}_n), n = 1, 2, \dots$ so that the co-ordinate process gets exactly these finite-dimensional distributions.

Definition from a random process

Suppose the probability space (Ω, \mathcal{F}, P) is given a priori, and a process x is defined on (Ω, \mathcal{F}) , either as $\{x_n, n \in \mathbb{N}\}$ or as $\{x(t), t \in T\}$. We shall use P on (Ω, \mathcal{F}) to define the distribution of x on the sample function spaces $(\mathbb{R}^\infty, \mathcal{B}_\infty)$ or $(\mathbb{R}^T, \mathcal{B}_T)$. The following lemma may look complicated, but its content is simple. It says that there is a direct connection between the events in \mathcal{B}_∞ or \mathcal{B}_T and certain events in \mathcal{F} on Ω .

Lemma 1.1. *For every set $A \in \mathcal{B}_\infty$ or $A \in \mathcal{B}_T$, the inverse set $x^{-1}(A) \in \mathcal{F}$:*

$$x^{-1}(A) = \{\omega \in \Omega; \mathbf{x}(\cdot, \omega) \in A\} \in \mathcal{F}.$$

Proof. Let \mathcal{C} be the class of sets C in \mathcal{B}_∞ or \mathcal{B}_T such that $x^{-1} \in \mathcal{F}$. By the definition of a stochastic process, \mathcal{C} contains all rectangles, and it is easy to show (do that as Exercise 1:3) that it is a σ -field. Since the Borel field is the smallest σ -field that contains the rectangles, the claim is shown. \square

It is now easy to define probabilities on the sample function spaces. If x is a stochastic process, i.e., a function from Ω to \mathbb{R}^∞ or \mathbb{R}^T , then a probability measure P_x is defined on $(\mathbb{R}^\infty, \mathcal{B}_\infty)$ and $(\mathbb{R}^T, \mathcal{B}_T)$, respectively, by

$$P_x(B) = P(x^{-1}(B)), \quad \text{for } B \in \mathcal{B}_\infty \text{ or } \mathcal{B}_T. \quad (1.4)$$

Thus, each stochastic process produces a probability measure P_x on $(\mathbb{R}^\infty, \mathcal{B}_\infty)$ or $(\mathbb{R}^T, \mathcal{B}_T)$.

Example 1.1 (“Random amplitude and phase”). *Here is a simple example of a stochastic process that is often used as a building block in more complicated models. Take two random variables A, ϕ , with $A \geq 0$ and ϕ uniformly distributed between 0 and 2π , and define a stochastic process $\{x(t), t \in \mathbb{R}\}$ as $x(t) = A \cos(t + \phi)$. This procedure defines the process x as a function from the sample space of (A, ϕ) -outcomes, $\Omega = [0, \infty) \times [0, 2\pi] = \{(A, \phi)\}$, to the function space \mathbb{R}^T , with $T = \mathbb{R}$. Of course, only cosine functions with different amplitudes and phases can appear as sample functions, and the probability measure on \mathbb{R}^T gives probability one to the subset of such functions. The finite-dimensional distributions of $x(t)$ can be calculated from the joint distribution of A and ϕ .*

Example 1.2 (“Poisson process”). *The sample functions of a common Poisson process are piecewise constant and increase in unit steps. It can be explicitly constructed from a sequence of independent exponential variables defining the time that elapses between two jumps, and the finite-dimensional distributions of the Poisson process can be calculated from the distribution of these intervals in $(\mathbb{R}^\infty, \mathcal{B}_\infty)$.*

It is essential, in the two examples, that we already have a probability space that allows us to construct the amplitude and phase variables, and the inter-event times, respectively. It is also worth to notice, that the more complex a process we wanted to generate, the more complex did we choose the basic sample space. This causes no problem, since one can always expand the sample space to include the desired variability. There is no need to look for the *simplest* probability space for a particular application.

If we start with a huge sample space (Ω, \mathcal{F}) with a single but complex probability measure P we can construct a rich variety of processes on the single probability space (Ω, \mathcal{F}, P) , and, by (1.4), obtain the probabilities of

sets B in \mathcal{B}_∞ or \mathcal{B}_T . We can also derive the finite-dimensional distributions. The negative side of this approach is that we don't know how to construct the measure P from experiments or from simple process models like Poisson or Gaussian models. We therefore usually rely on the other alternative to construct process distributions.

Definition from a family of finite-dimensional distributions:

The family $\{F_{\mathbf{t}^n}\}_{n=1}^\infty$ of finite-dimensional distributions is the family of functions

$$F_{t_1 \dots t_n}(b_1, \dots, b_n) = \text{Prob}(x(t_1) \leq b_1, \dots, x(t_n) \leq b_n),$$

for $t_j \in T, t_1 < t_2 < \dots < t_n, n = 1, 2, \dots$. A family of distribution functions for a stochastic process has to be *consistent* to satisfy certain, trivial requirements.

Definition 1.3. A family $\{F_{\mathbf{t}^n}\}_{n=1}^\infty$ of finite-dimensional distributions is called consistent if, for every n , and $t_1, b_1, \dots, t_n, b_n$,

$$\lim_{b_k \uparrow \infty} F_{t_1, \dots, t_n}(b_1, \dots, b_n) = F_{t_1, \dots, \tilde{t}_k, \dots, t_n}(b_1, \dots, \tilde{b}_k, \dots, b_n),$$

$$F_{t_1, t_2}(b_1, b_2) = F_{t_2, t_1}(b_2, b_1), \text{ etc.}$$

where the notation $\tilde{\cdot}$ indicates that the variable is missing.

Suppose a family of finite-dimensional distribution functions,

$$\mathbf{F} = \{F_{\mathbf{t}^n}\}_{n=1}^\infty, t_k \in T,$$

$$F_{t_1 \dots t_n}(b_1, \dots, b_n) = \text{Prob}(x(t_1) \leq b_1, \dots, x(t_n) \leq b_n),$$

is given a priori, with one distribution for each n and for each set (t_1, \dots, t_n) of times. If the family is consistent (Definition 1.3), one can use it to define probabilities for all half-open n -dimensional intervals in \mathbb{R}^∞ or \mathbb{R}^T , for example by, for $n = 1, 2, \dots$, taking (cf. (1.1))

$$P_{\mathbf{x}}((a_1, b_1] \times (a_2, b_2] \times \dots \times (a_n, b_n] \times \mathbb{R}^\infty) = P_n((a_1, b_1] \times (a_2, b_2] \times \dots \times (a_n, b_n]).$$

This will give us a unique probability measure for each n and each n -dimensional selection of time points (t_1, \dots, t_n) .

Theorem 1.1 (Existence of stochastic process; Kolmogorov, 1933).

To every consistent family of finite-dimensional distribution functions, $\mathbf{F} = \{F_{\mathbf{t}^n}\}_{n=1}^{\infty}$, $t_k \in T$, there exists one and only one probability measure P_x on $(\mathbb{R}^T, \mathcal{B}_T)$ with

$$P_x(x(t_1) \leq b_1, \dots, x(t_n) \leq b_n) = F_{t_1 \dots t_n}(b_1, \dots, b_n), \text{ for } n = 1, 2, \dots$$

In Appendix A.2 we present an outline of the proof of the existence theorem. The main steps are to show that the finite dimensional distributions give a countably additive probability measure on the field \mathcal{J} of finite unions of intervals, defined for different times (t_1, \dots, t_n) . By the extension property of probability measures on fields, Theorem A.1, one can then conclude that this defines a unique probability measure P_x on the σ -field generated by the intervals, i.e., on the Borel field \mathcal{B}_{∞} or \mathcal{B}_T , respectively. The finite-dimensional distributions are the distributions in \mathbf{F} . The proof of the countable additivity is a significant part of Kolmogorov's existence theorem for stochastic processes; see Theorem A.3.

By this, we have defined events in \mathbb{R}^{∞} and \mathbb{R}^T and know how to define probability measures in $(\mathbb{R}^{\infty}, \mathcal{B}_{\infty})$ and $(\mathbb{R}^T, \mathcal{B}_T)$.

Remark 1.1. *It is worth noting that all the well-known processes in elementary probability theory can be introduced by finite dimensional distributions. Going back to Example 1.2, we saw there an explicit way to construct a Poisson process. On the other hand, following the route via the finite-dimensional distribution functions the Poisson process on the real line is characterized by the fact that the increments over a finite number of disjoint intervals are independent Poisson variables with expectation proportional to the interval length. This property can easily be formulated as a property of the finite-dimensional distributions at the interval end points, even if one never does so explicitly.*

Both approaches lead to a probability measure P_x , defined for all $B \in \mathcal{B}_T$, either by (1.4) or by the Kolmogorov existence theorem. There is an important difference between the two approaches. With the explicit construction all sample functions are piecewise constant with unit jumps, by construction, and the probability measure P_x has the potential to take that into account and it can easily be extended to more events (subsets) than just those in \mathcal{B}_T . The construction by the Kolmogorov theorem does not allow any strong statement about sample function properties. In the next section and in Chapter 2 we discuss some consequences of this fact in more detail.

Finally, also the explicit construction relies on the Kolmogorov existence theorem, since it requires the existence of an infinite sequence of independent exponential random variables.

1.2.4 An important comment on probabilities on \mathbb{R}^T

Processes with continuous time need some extra comments. Even if the existence of such processes can be deduced from Kolmogorov's theorem, the practical applications need some care. The sample space Ω is an abstract space and a mathematical construction, and the link to reality is provided by the random variables. In an experiment, one can observe the values of one or more random variables, x_1, x_2 , etc. and also find their distribution, by some statistical procedure. There is no serious difficulty to allow the outcome to be any real number, and to define probability distributions on \mathbb{R} or \mathbb{R}^n .

When the result of an experiment is a whole function with continuous parameter, the situation is more complicated. In principle, *all functions* of $t \in T$ are potential outcomes, and the sample space of all functions on T is simply so big that one needs to be restrictive on which sets of outcomes that should be given probability. There are too many realizations that ask for probability.

Here, practice comes to our assistance. In an experiment one can only observe the values of $x(t)$ at a finite number of times, t_1, t_2, \dots, t_n ; with $n = \infty$ we allow an unlimited series of observations. The construction of processes with continuous time was built on exactly this fact: the observable events are those that can be defined by countably many $x(t_j)$, $j \in \mathbb{N}$, and the probability measure we constructed assigns probabilities to only such events.

A set of functions that is defined by restrictions on countably many $x(t_j)$ is said to have a *countable basis*; see more on this in Appendix A.1. There are many important function sets that do not have a countable basis and are not Borel sets in \mathbb{R}^T . These sets are not given a unique probability by the finite-dimensional distribution functions.

Example 1.3. Here are some examples of function sets with and without countable basis, when $T = [0, 1]$:

- $\{x \in \mathbb{R}^T; \lim_{n \rightarrow \infty} x(1/n) \text{ exists}\} \in \mathcal{B}_T$,
- $\{x \in \mathbb{R}^T; \lim_{t \rightarrow 0} x(t) \text{ exists}\} \notin \mathcal{B}_T$,
- $\{x \in \mathbb{R}^T; x \text{ is a continuous function}\} \notin \mathcal{B}_T$,
- $\{x \in \mathbb{R}^T; x(t) \leq 2 \text{ for all rational } t\} \in \mathcal{B}_T$.

Approximation by finite-dimensional events

The events in the σ -field \mathcal{B}_∞ in \mathbb{R}^∞ can be approximated in probability by finite-dimensional sets. If $(\mathbb{R}^\infty, \mathcal{B}_\infty, P)$ is a probability space, and $B \in \mathcal{B}_\infty$, then, for every $\varepsilon > 0$, there is a finite n and an event $B_n \in \mathcal{B}_n$ such that

$$P(B \Delta \tilde{B}_n) \leq \varepsilon,$$

where $\tilde{B}_n = \{x \in \mathbb{R}^\infty; (x_1, \dots, x_n) \in B_n\}$ and $A \Delta B = (A - B) \cup (B - A)$.

Similarly, events in \mathcal{B}_T in \mathbb{R}^T can be approximated arbitrarily close by events defined by the values of $x(t)$ for a finite number of t -values: $P(B \Delta \tilde{B}_n) \leq \varepsilon$, with

$$\tilde{B}_n = \{x \in \mathbb{R}^T; (x(t_1), \dots, x(t_n)) \in B_n\}.$$

Remember that every probability measure on $(\mathbb{R}^\infty, \mathcal{B}_\infty)$ is uniquely determined by its finite-dimensional distributions, which implies that also every probability measure P on $(\mathbb{R}^T, \mathcal{B}_T)$ is determined by the finite-dimensional distributions, $\{F_t^n\}_{n=1}^\infty$. In particular, the probability

$$P\left(\lim_{n \rightarrow \infty} x(t_0 + 1/n) \text{ exists and is equal to } x(t_0)\right)$$

is determined by the finite-dimensional distributions. But $x(t_0 + 1/n) \rightarrow x(t_0)$ as $n \rightarrow \infty$ is almost, but not quite, the same as $x(t) \rightarrow x(t_0)$ as $t \rightarrow t_0$. To deal with sample function continuity we need a more refined construction of the probability measure from the finite-dimensional distributions. We deal with these matters in Chapter 2.

1.3 Stationary processes and fields

This section summarizes some elementary notation for stationary processes. More details, properties, and proofs of the most important facts will be given in Chapter 3. To avoid too cumbersome notation we sometimes allow ourselves to write “the process $x(t)$,” when we should have used the notation $\{x(t), t \in \mathbb{R}\}$ for the process (or any other relevant parameter set T instead of \mathbb{R}). The simple notation x will also be frequently used, when the parameter space is clear from the context. If we mean “the random variable $x(t)$,” we will say so explicitly. At this stage, we deal with real valued processes, only. Later, we will need to consider also complex valued processes.

We write $m(t) = E(x(t))$ for the *mean value function* and $r(s, t) = C(x(s), x(t))$ for the *covariance function* of the process $\{x(t), t \in \mathbb{R}\}$. If $\{x(t), t \in \mathbb{R}\}$ and $\{y(t), t \in \mathbb{R}\}$ are two processes, $r_{x,y}(s, t) = C(x(s), y(t))$ is the *cross-covariance function*.

1.3.1 Stationary processes, covariance, and spectrum

Definition 1.4 (Stationarity). A stochastic process $\{x(t), t \in \mathbb{R}\}$ is strictly stationary if all n -dimensional distributions of $x(t_1 + \tau), \dots, x(t_n + \tau)$ are independent of τ . It is called weakly stationary (or second order stationary) if its mean is constant, $E(x(t)) = m$, and if its covariance function,

$$r(t) = C(x(s+t), x(s)),$$

is a function only of the time lag t .

Here are some simple properties of a covariance function of a stationary process.

Theorem 1.2. If $r(t)$ is the covariance function of a (real) stationary process, then,

- a) $V(x(t+h) \pm x(t)) = 2(r(0) \pm r(h))$, so $|r(h)| \leq r(0)$,
- b) if $|r(\tau)| = r(0)$ for some $\tau \neq 0$, then $r(t)$ is periodic,
- c) if $r(t)$ is continuous at $t = 0$, then it is continuous everywhere.

Proof. Property (a) is clear, and it implies $|r(t)| \leq r(0)$. From (a) it also follows that if $r(\tau) = \pm r(0)$, then $x(t + \tau) = \pm x(t)$ for all t , so $x(t)$ is periodic with period τ and 2τ , respectively, which is (b). For (c), use the inequality, $C(x, y)^2 \leq V(x)V(y)$, on

$$\begin{aligned} (r(t+h) - r(t))^2 &= (C(x(0), x(t+h)) - C(x(0), x(t)))^2 \\ &= (C(x(0), x(t+h) - x(t)))^2 \\ &\leq V(x(0))V(x(t+h) - x(t)) = 2r(0) \cdot (r(0) - r(h)), \end{aligned}$$

which goes to 0 with h if $r(t)$ is continuous for $t = 0$. □

The covariance function is said to define the *time domain* properties of the process. Equally important for the theory is the *frequency domain* properties, expressed by the *spectral distribution*, studied in detail in Chapter 3.

Every continuous covariance function is a Fourier integral,

$$r(t) = \int_{-\infty}^{\infty} e^{i\omega t} dF(\omega), \quad (1.5)$$

where the function $F(\omega)$ is the spectral distribution function.

The spectral distribution is characterized by the properties:³

- symmetry: $dF(-\omega) = dF(\omega)$,
- monotonicity: $\omega \leq \omega'$ implies $F(\omega) \leq F(\omega')$,
- boundedness: $F(+\infty) - F(-\infty) = r(0) = V(x(t)) < \infty$.

As indicated by the way we write the three properties, $F(\omega)$ is defined only up to an additive constant, and we usually take $F(-\infty) = 0$. The spectral distribution function is then equal to a cumulative distribution function multiplied by a positive constant, equal to the variance of the process. As we shall see in later chapters, ω is in a natural way interpreted as an *angular frequency*, not to be confused with the elementary event ω in basic probability theory.

In this work we will mainly use the symmetric form (1.5) of the spectral representation when we deal with theoretical aspects of the covariance function and of the spectrum. However, in many examples we will use a one-sided real form with $\omega \geq 0$; for details on this, see Section 3.2.3. In Section 3.3.5 we will meet examples where negative frequencies play an important role in their own right.

If $F(\omega)$ is absolutely continuous with $F(\omega) = \int_{s=-\infty}^{\omega} f(s) ds$, then the spectrum is said to be (absolutely) continuous, and

$$f(\omega) = \frac{d}{d\omega} F(\omega)$$

is the *spectral density function*.

The *spectral moments* are defined as

$$\omega_k = \int_{-\infty}^{\infty} |\omega|^k dF(\omega).$$

Note that the odd spectral moments are defined as absolute moments. Since F is symmetric around 0 the signed odd moments are always 0, when they exist.

³The differential notation in $dF(-\omega) = dF(\omega)$ will be used throughout the book as a convenient way to simplify increment notation.

Spectral moments may be finite or infinite. As we shall see in the next chapter, the finiteness of the spectral moments are coupled to the smoothness properties of the process $x(t)$. For example, the process is differentiable (in quadratic mean), see Section 2.1, if $\omega_2 = -r''(0) < \infty$, and the spectral density of the derivative is

$$f_{x'}(\omega) = \omega^2 f_x(\omega).$$

To ensure that sample functions are differentiable functions one needs just a little bit more than $\omega_2 = \int_{-\infty}^{\infty} \omega^2 f_x(\omega) d\omega < \infty$, namely $-r''(t) = \omega_2 - C|t|^\alpha + o(|t|^\alpha)$ as $t \rightarrow 0$, for some $\alpha \in (1, 2]$; see Theorem 2.10, Section 2.3.1.

The *mean frequency* of a stationary process,

$$\mu^+ = \frac{1}{2\pi} \sqrt{\omega_2/\omega_0}, \quad (1.6)$$

and its inverse, $1/\mu^+$, the *mean period*, characterize the main feature of the spectrum. The width of the spectrum around the mean frequency is measured by the *spectral width* parameter $1/\alpha = (\omega_0\omega_4)^{1/2}/\omega_2$.

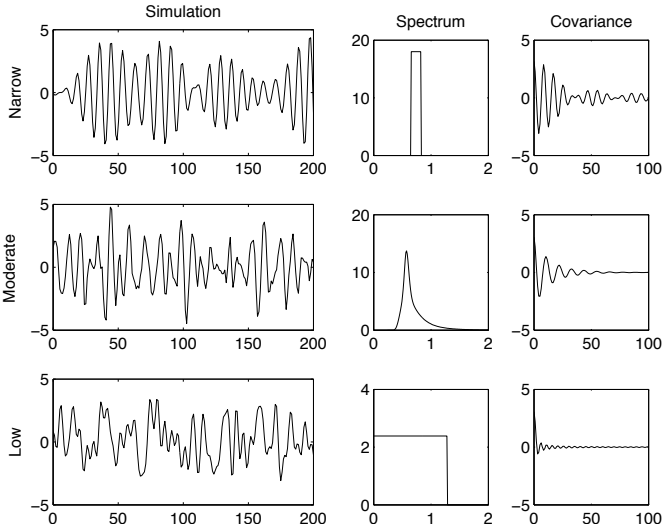


Figure 1.2 Top: Processes with very narrow band spectrum, oscillating covariance function, and “fading” sample function; Middle: Realistic ocean waves with moderate band width wave spectrum; Bottom: low frequency white noise spectrum. The spectra are illustrated by their one-sided spectral densities.

Example 1.4. Here is a first example on the visual characteristics of spectrum, covariance function, and simulated sample function. Figure 1.2 illustrates one very narrow spectrum with oscillating covariance function and

regular “fading” sample path, one realistic water wave spectrum with less regular sample paths, and one irregular process with “low frequency white noise” spectrum.

1.3.2 Stationary streams of events

The Poisson process is a *counting process* that counts the number of events that happen at discrete time points. Its characteristic property is that the number of events in any interval of length h has a Poisson distribution with mean λh , with parameter λ called the *intensity*. Further, the number of events in disjoint intervals are statistically independent. The events counted by the Poisson process is an example of a *stationary point process*. This is a natural generalization of the notion “stationary process,” since the distribution of the number of events in a set of intervals is the same, regardless of their absolute location.

If one relaxes the independence condition, one obtains a more general class of point processes, where there may be causal or statistical dependence between events. For example, the intensity in a Poisson process may be allowed to vary like a stationary process. It is then possible to describe this dependence by means of a type of covariance function that describes the conditional intensity of events a distance t apart. We shall deal with this covariance density in Section 3.5, where we also introduce the corresponding spectral distribution. For general accounts of points processes, see [37, 39].

1.3.3 Random fields

A random field is a stochastic process $\{x(\mathbf{t})\}$ with multi-dimensional parameter $\mathbf{t} = (t_1, \dots, t_p) \in \mathbf{T}$, which can be discrete or continuous. For example, if $\mathbf{t} = (t_1, t_2)$ is two-dimensional we can think of $(t_1, t_2, x(\mathbf{t}))$, $(t_1, t_2) \in \mathbb{R}^2$, as a random surface. The mean value and covariance functions are defined in the natural way, $m(\mathbf{t}) = E(x(\mathbf{t}))$ and $r(\mathbf{t}, \mathbf{u}) = C(x(\mathbf{t}), x(\mathbf{u}))$.

A random field is called *homogeneous* if it has constant mean value $m(\mathbf{t}) = m$ and the covariance function $r(\mathbf{t}, \mathbf{u})$ depends only on the vector $\mathbf{t} - \mathbf{u}$ between the two observation points, i.e., assuming $m = 0$,

$$r(\mathbf{t}) = r(\mathbf{u} + \mathbf{t}, \mathbf{u}) = E(x(\mathbf{u} + \mathbf{t}) \cdot x(\mathbf{u})).$$

The covariance of the process values at two parameter points depends on *distance* as well as on *direction* of the vector between the two points. This is the equivalent of a stationary process on the real line.

If the covariance between $x(\mathbf{u})$ and $x(\mathbf{v})$ depends only on the distance $\tau = \|\mathbf{u} - \mathbf{v}\|$ between the observation points and not on the direction, the field is

called *isotropic*. This requirement poses strong restrictions on the covariance function, as we shall see in Chapter 7, where random fields are treated in more detail.

A digital image can be thought of as a random field with discrete parameter $\mathbf{t} \in \mathbf{T} \subset \mathbb{N}^2$. A time-dependent random surface is a field $(s_1, s_2, x(t, s_1, s_2))$ with $\mathbf{t} = (t, s_1, s_2)$, where t represents time and $(s_1, s_2) \in \mathbb{R}^2$ is location.

1.4 Gaussian processes

1.4.1 Normal distributions and Gaussian processes

The Gaussian, or normal, univariate distribution with mean m and variance σ^2 has probability density function

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-(x-m)^2/(2\sigma^2)} = \frac{1}{\sigma} \phi\left(\frac{x-m}{\sigma}\right),$$

where $\phi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$ is the standardized normal density function. A distribution concentrated at one point m is regarded as Gaussian with $\sigma^2 = 0$.

Definition 1.5. A vector $\boldsymbol{\xi} = (\xi_1, \dots, \xi_p)$ of p random variables is said to have a p -variate Gaussian (normal) distribution if every linear combination of its components $\mathbf{a}\boldsymbol{\xi}' = \sum_k a_k \xi_k$ has a normal distribution. The variables ξ_1, \dots, ξ_p are said to be “jointly normal.”

With mean vector $\mathbf{m} = E(\boldsymbol{\xi})$ and covariance matrix

$$\boldsymbol{\Sigma} = C(\boldsymbol{\xi}; \boldsymbol{\xi}) = E((\boldsymbol{\xi} - \mathbf{m})'(\boldsymbol{\xi} - \mathbf{m})),$$

the variance of $\mathbf{a}\boldsymbol{\xi}'$ is

$$V(\mathbf{a}\boldsymbol{\xi}') = \mathbf{a}\boldsymbol{\Sigma}\mathbf{a}' \geq 0.$$

The *characteristic function* of a normal variable ξ is (see Appendix A.5)

$$\varphi(s) = E(e^{is\xi}) = e^{ims - \sigma^2 s^2/2},$$

and that of a multi-variate normal $\boldsymbol{\xi}$ is

$$\varphi(\mathbf{s}) = E(e^{i\mathbf{s}\boldsymbol{\xi}'}) = e^{i\mathbf{m}\mathbf{s}' - \mathbf{s}\boldsymbol{\Sigma}\mathbf{s}'/2}. \quad (1.7)$$

It follows from the form of the characteristic function that if two jointly normal variables are uncorrelated, i.e., have zero covariance, then they are also independent. To see this, one can first note that if Σ is diagonal, then the characteristic function factorizes into a product of characteristic functions for independent normals. The uniqueness of characteristic functions does the rest.

If the determinant of Σ is positive, the distribution of ξ is non-singular and has a density

$$f_{\xi}(\mathbf{x}) = \frac{1}{(2\pi)^{p/2} \sqrt{\det \Sigma}} e^{-\frac{1}{2}(\mathbf{x}-\mathbf{m})\Sigma^{-1}(\mathbf{x}-\mathbf{m})'}.$$

If the determinant is zero, the distribution of ξ is concentrated to a linear subspace of \mathbb{R}^n and there exists at least one linear relationship between the components, i.e., there is at least one \mathbf{a} for which $\mathbf{a}\xi'$ is a constant.

Definition 1.6. A stochastic process $\{x(t), t \in \mathbb{R}\}$ is a Gaussian process if every linear combination $S = \sum_k a_k x(t_k)$ for real a_k and $t_k \in \mathbb{R}$ has a Gaussian distribution.

Example 1.5 (Existence of a Gaussian process). What knowledge is needed before we can ascertain the existence of a specific Gaussian process? In view of Kolmogorov's existence Theorem 1.1, we must present a consistent family of finite-dimensional distribution functions, such that every linear combination $\sum_k a_k x(t_k)$ of variables is Gaussian, i.e., that for every selection, the variables $x(t_1), \dots, x(t_n)$ are jointly normal.

Since the multi-variate normal distribution is defined by its characteristic function, it is clear from (1.7) that it suffices to specify a function $m(t)$ and a function $r(s, t)$ such that

$$\sum_{j,k} a_j a_k r(t_j - t_k) \geq 0, \quad (1.8)$$

for all real a_k and t_k . Then there exists a Gaussian process with $m(t)$ and $r(s, t)$ as mean value and covariance function, respectively. We will meet condition (1.8) in Chapter 8.

It is an easy consequence of the definition that the derivative of a Gaussian

process is also Gaussian (when it exists), since it is the limit of the Gaussian variable $z_h = (x(t+h) - x(t))/h$ as $h \rightarrow 0$. For a stationary Gaussian process $\{x(t), t \in \mathbb{R}\}$ the mean of z_h is 0 and it has variance $V(z_h) = 2(r(0) - r(h))/h^2$. As we shall prove in Section 2.1.2 this converges to $\omega_2 = \int \omega^2 dF(\omega \leq \infty)$. The derivative exists only if this limit is finite.

Also the integral of a Gaussian process is a Gaussian variable; conditions for the existence will be given in Section 2.4.

1.4.2 Conditional normal distributions

The multivariate normal distribution has the very useful property that, conditional on observations of a subset of variables, the unobserved variables are also normal.⁴ Further, the conditional expectation is a linear combination of the observations while variances and covariances are independent of the observed values.

Let $\xi = (\xi_1, \dots, \xi_n)$ and $\eta = (\eta_1, \dots, \eta_m)$ be two jointly Gaussian vectors with mean values

$$E(\xi) = m_\xi, \quad E(\eta) = m_\eta,$$

and with covariance matrix (with $\Sigma_{\xi\eta} = \Sigma'_{\eta\xi}$),

$$\Sigma = C((\xi, \eta); (\xi, \eta)) = \begin{pmatrix} \Sigma_{\xi\xi} & \Sigma_{\xi\eta} \\ \Sigma_{\eta\xi} & \Sigma_{\eta\eta} \end{pmatrix}.$$

If the determinant of the covariance matrix Σ is positive, then the distribution of (ξ, η) has a non-singular density

$$f_{\xi\eta}(\mathbf{x}, \mathbf{y}) = \frac{1}{(2\pi)^{(m+n)/2} \sqrt{\det \Sigma}} e^{-\frac{1}{2}(\mathbf{x} - m_\xi, \mathbf{y} - m_\eta) \Sigma^{-1} (\mathbf{x} - m_\xi, \mathbf{y} - m_\eta)' }.$$

The density of η is

$$f_\eta(\mathbf{y}) = \frac{1}{(2\pi)^{m/2} \sqrt{\det \Sigma_{\eta\eta}}} e^{-\frac{1}{2}(\mathbf{y} - m_\eta) \Sigma_{\eta\eta}^{-1} (\mathbf{y} - m_\eta)' },$$

and the conditional density $f_{\xi|\eta}(\mathbf{x}|\mathbf{y})$, defined as

$$f_{\xi|\eta}(\mathbf{x} | \mathbf{y}) = \frac{f_{\xi\eta}(\mathbf{y}, \mathbf{x})}{f_\eta(\mathbf{y})}.$$

⁴A summary of conditional distributions is found in Section A.3.3 in Appendix A.

This is again a Gaussian density and its mean and covariance properties can easily be found by some matrix algebra. We summarize the important facts as a theorem.

Theorem 1.3. *The conditional distribution of ξ given $\eta = \mathbf{y}$ is Gaussian with conditional mean matrix*

$$\begin{aligned} E(\xi \mid \eta = \mathbf{y}) &= \widehat{\xi}(\mathbf{y}) = E(\xi) + C(\xi; \eta) \Sigma_{\eta\eta}^{-1} (\mathbf{y} - E(\eta))' \\ &= m_\xi + \Sigma_{\xi\eta} \Sigma_{\eta\eta}^{-1} (\mathbf{y} - m_\eta)'. \end{aligned} \quad (1.9)$$

The conditional covariance is

$$\Sigma_{\xi\xi|\eta} = E((\xi - \widehat{\xi}(\eta))'(\xi - \widehat{\xi}(\eta))) = \Sigma_{\xi\xi} - \Sigma_{\xi\eta} \Sigma_{\eta\eta}^{-1} \Sigma_{\eta\xi}. \quad (1.10)$$

In two dimensions the formulas read

$$m_{x|y} = m_x + \sigma_x \rho_{xy} \cdot \frac{y - m_y}{\sigma_y}, \quad \sigma_{x|y}^2 = \sigma_x^2 (1 - \rho_{xy}^2),$$

with $\rho_{xy} = C(x, y) / \sqrt{V(x)V(y)}$; thus the squared correlation ρ_{xy}^2 gives the relative reduction of the variability (uncertainty) in the random variable x gained by observation of y .

Observe the mnemotechnical friendliness of the formulas for conditional mean and covariance. For example, the covariance matrix $\Sigma_{\xi\xi|\eta}$ has dimension $n \times n$ and the configuration on the right hand side of (1.10) is the only way to combine the matrices involved that matches their dimensions – of course, you have to remember the general structure.

1.4.3 Linear prediction and reconstruction

Prediction and reconstruction are two of the most important applications of stationary process theory. Even though these problems are not main topics in this work, we present one of the basic concepts here; in Section 6.6 we will deal with the more philosophical sides of the prediction problem.

Suppose we have observed the outcomes of a set of random variables, $\eta = (\eta_1, \dots, \eta_m)$, and that we want to give a statement $\widehat{\xi}$ about the outcome of some other variable ξ , either to be observed sometime in the future, or perhaps a missing observation in a time series. These two cases constitute the framework of *prediction* and *reconstruction*, respectively. Also suppose that

we want to make the statement in the best possible way in the mean square sense, i.e., we want $E((\xi - \hat{\xi})^2)$ to be as small as possible.

It is easy to see (Theorem A.4 in Appendix A) that the best solution in mean square sense is given by the conditional expectation, $\hat{\xi} = E(\xi | \boldsymbol{\eta}) = \varphi(\boldsymbol{\eta})$. On the other hand, if the variables are jointly Gaussian, then we know from formula (1.9) that the conditional expectation of ξ given $\boldsymbol{\eta}$ is linear in $\boldsymbol{\eta}$, so that for Gaussian variables the optimal solution is

$$\hat{\xi} = E(\xi | \boldsymbol{\eta}) = m_{\xi} + \boldsymbol{\Sigma}_{\xi\boldsymbol{\eta}} \boldsymbol{\Sigma}_{\boldsymbol{\eta}\boldsymbol{\eta}}^{-1} (\boldsymbol{\eta} - m_{\boldsymbol{\eta}})', \quad (1.11)$$

an expression that depends only on the mean values and the second order moments, i.e., variances and covariances.

Looking at the general case, without assuming normality, we restrict ourselves to solutions that are linear functions of the observed variables. It is clear that the solution that is optimal in the mean square sense can only depend on the mean values and variances/covariances of the variables. It therefore has the same form for all variables with the same first and second order moments. Thus, (1.11) gives the best linear predictor in mean square sense, regardless of the distribution.

1.5 Four historical landmarks

1.5.1 Brownian motion and the Wiener process

There is no other Gaussian process with as wide applicability as the Wiener process. Even if it is a non-stationary process it appears repeatedly in the theory of stationary processes, and we spend this section to describe some of its properties and applications.

Definition 1.7 (Wiener process). *The Wiener process $\{w(t); t \geq 0\}$ is a Gaussian process with $w(0) = 0$ such that $E(w(t)) = 0$, and the variance of the increment $w(t+h) - w(t)$ over any interval $[t, t+h]$, $h > 0$, is proportional to the interval length,*

$$V(w(t+h) - w(t)) = h\sigma^2.$$

A Wiener process $\{w(t), t \in \mathbb{R}\}$ over the whole real line is a combination of two independent Wiener processes w_1 and w_2 , so that $w(t) = w_1(t)$ for $t \geq 0$, and $w(t) = w_2(-t)$ for $t < 0$.

It is an easy consequence of the definition that the increment $w(t) - w(s)$ is uncorrelated with $w(s)$ for $s < t$; just expand the variance of $x(t)$,

$$\begin{aligned} t\sigma^2 &= V(w(t)) = V(w(s)) + V(w(t) - w(s)) + 2C(w(s), w(t) - w(s)) \\ &= s\sigma^2 + (t - s)\sigma^2 + 2C(w(s), w(t) - w(s)). \end{aligned}$$

It also follows that $C(w(s), w(t)) = \sigma^2 \min(s, t)$. Since the increments over disjoint intervals are jointly normal, by the definition of a normal process, they are also independent.

A characteristic feature of the Wiener process is that its future changes are statistically independent of its actual and previous values. It is intuitively clear that a process with this property cannot be differentiable. The increment over a small time interval from t to $t + h$ is of the order \sqrt{h} , which is small enough to make the process continuous, but it is too large to give a differentiable process. In Chapter 2 we will give conditions for continuity and differentiability of sample functions of Gaussian and general processes.

The sample functions of a Wiener process are in fact objects that have *fractal dimension*, and the process is *self similar* in the sense that when magnified with proper scales it retains its statistical geometrical properties. More precisely, for each $a > 0$, the process $\sqrt{a}w(t/a)$ has the same distributions as the original process $w(t)$.

The Wiener process is commonly used to model phenomena where the local changes are virtually independent. Symbolically, one can write $dw(t)$ for the infinitesimal independent increments, or simply use “the derivative” $w'(t)$. In this form, the Wiener process is used as a driving force in many stochastic systems, appearing in Chapter 4. The Brownian motion is a good example of how one can use the Wiener process to get models with more or less physical realism.

The Brownian motion, observed 1827 by the Scottish biologist Robert Brown, is an erratic movement by small particles immersed in a fluid, for example pollen particles in water as in Brown’s original experiment. Albert Einstein presented in 1905 a quantitative model for the Brownian movement in his paper *On the movements of small particles in a stationary liquid demanded by the molecular-kinetic theory of heat*, reprinted in [46], based on the assumption that the movements are caused by independent impacts on the particle by the molecules of the surrounding fluid medium.

In Einstein’s model the changes in location due to collisions over separate time intervals are supposed to be independent. This requires however that the particles have no mass, which is physically wrong, but sufficiently accurate for microscopic purposes on a human time scale. According to Einstein, the

change in location in any of the three directions (x, y, z) over a time interval of length t is random and normal with mean zero, which is not surprising, since the changes are the results of a very large number of independent collisions. What made Einstein's contribution conclusive was that he derived an expression for the variance in terms of other physical parameters, namely

$$V(x(t)) = V(y(t)) = V(z(t)) = t \frac{4RT}{Nf} = t \sigma^2, \quad (1.12)$$

where T is the absolute temperature, R is the Boltzmann constant, and N is Avogadro's number, i.e., the number of molecules per mole, and the friction coefficient f depends on the shape and size of the particle and on the viscosity of the fluid. Here, the coordinates are independent Wiener processes.

Observations of the Brownian movement and estimation of its variance makes it possible to calculate any of the factors in σ^2 , for example N , from the other ones.

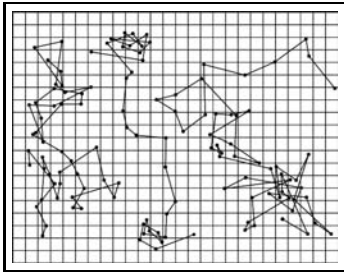


Figure 1.3 *Three particle paths in Perrin's experiment.*

The French physicist J.B. Perrin estimated in a series of experiments 1908-1911 Avogadro's number in this way by observing suspended rubber particles and found an estimate that is correct within about 10%. The figure, reproduced from the book *Les Atomes* by Perrin [96], shows the paths of the motions of three colloidal particles as seen under the microscope. Dots show the positions every 30 seconds. For his experiments and conclusions, Perrin was awarded the Nobel prize in physics, 1926, for his work on the "discontinuous structure of matter."

In a more realistic model, one takes also particle mass and velocity into account. If $v(t)$ denotes the velocity at time t , the fluid offers a resistance from the friction force, which is equal to $f v(t)$, with the friction coefficient as in (1.12). Further, the particle offers a resistance to changes in velocity proportional to its mass m . Finally, one needs to model the independent collisions from the surrounding fluid molecules, and here the Wiener process can again be used, more precisely its increments $dw(t)$. This gives the *Langevin equation*, named after the French physicist Paul Langevin, for the particle velocity,

$$dv(t) + \alpha v(t) dt = \frac{1}{m} dw(t), \quad (1.13)$$

where $\alpha = f/m$, and $w(t)$ is a standardized Wiener process. It is usually written

$$\frac{dv(t)}{dt} + \alpha v(t) = \frac{1}{m} w'(t). \quad (1.14)$$

We will meet the Langevin equation in Example 3.5 on the Ornstein-Uhlenbeck process in Section 3.4.2.

1.5.2 *S.O. Rice and electronic noise*

The two papers *Mathematical analysis of random noise*, by Steve O. Rice, appeared in Bell System Technical Journal, 1944–1945 [100]. They represent a landmark in the history of stochastic processes in that they bring together and exhibit the wide applicability in engineering of the spectral formulation of a stationary process as a sum, or asymptotically an integral, of harmonic cosine functions with random amplitudes and phases. Correlation functions and their Fourier transforms had been studied at least since the early 1900s, and Rice's work brought together these results in a systematic way. But it also contained many new results, in particular pertaining to crossing related properties, and on the statistical properties of stationary processes “obtained by passing random noise through physical devices.”

Rice used the *spectral representation* of a stationary Gaussian process as a sum over discrete positive frequencies $\omega_k > 0$,

$$x(t) = \sum_k a_k \cos \omega_k t + b_k \sin \omega_k t = \sum_k c_k \cos(\omega_k t + \phi_k), \quad (1.15)$$

where the amplitudes a_k, b_k are normal and independent random variables with mean 0 and $E(a_k^2) = E(b_k^2) = \sigma_k^2$, and ϕ_k uniformly distributed over $(0, 2\pi)$, independent of the amplitudes. As we shall see in Chapter 3, such a process has covariance

$$r(t) = \sum_k \sigma_k^2 \cos \omega_k t.$$

The spectral distributions function is a discrete distribution with point mass $\sigma_k^2/2$ at the symmetrically located frequencies $\pm\omega_k$.

The continuous, integral form of the spectral representation is presented as limiting cases in Rice's work. At about the same time, Harald Cramér and Michel Loève gave a probabilistic formulation of the continuous spectral representation, in a mathematically impeccable way [31, 33]. In Chapter 3 we shall deal in more depth with spectral representation, in different forms.

Besides the previously known “Rice’s formula” for the expected number of level crossings, Rice’s 1945 paper also analyzed crossing and excursion distributions and investigated the joint occurrence of crossings of a fixed level at two distinct points, necessary for calculation of the variance of the number of crossings. The simple Rice’s formula can be generalized to vector processes and to random fields, and has developed into one of the most powerful tools in the statistical analysis of random systems; see [6]. Crossing problems are the topic of Chapter 8.

The flexibility and generality of Rice’s methods and examples made correlation and spectral theory fundamental ingredients in communication theory and signal processing for decades to come. An example by Rice himself is the ingenious explanation of the intriguing click noise in analogue FM radio [101].

1.5.3 Gaussian random wave models

Steve Rice’s analysis of time dependent stationary processes had, as mentioned, great influence on signal processing in the information sciences. Less well known in the statistical world is the effect his work had in oceanography and naval architecture.

To give the reader a feeling for the dramatic effect Rice’s work on random noise had in a quite different field, we cite in extenso (references deleted) the first two paragraphs in Manley St Denis and Willard J. Pierson’s paper, *On the motion of ships in confused seas*, which came out 1954 [115]:

HISTORY

Three years ago the first co-author of the present work collaborated with Weinblum in the writing of a paper entitled “On the motion of ships at sea.” In that paper Lord Rayleigh was quoted saying: “The basic law of the seaway is the apparent lack of any law.” Having made this quotation, however, the authors then proceed to consider the seaway as being composed of “a regular train of waves defined by simple equations.” This artificial substitution of pattern for chaos was dictated by the necessity of reducing the utterly confused reality to a simple form amenable to mathematical treatment.

Yet at the same time and in other fields the challenging study of confusion was being actively pursued. Thus in 1945 Rice was writing on the mathematical analysis of random noise and in 1949 Tukey and Hamming were writing on the properties of stationary time series and their power spectra in connection with colored noise. In the same year Wiener published his now famous book on time series. These works

were written as contributions to the theory of communication. Nevertheless the fundamental mathematical discipline expounded therein can readily be extended to other fields of scientific endeavor. Thus in 1952 the second co-author, inspired by a contribution of Tukey, was able to apply the foregoing theories to the study of actual ocean waves. As the result of analyses of actual wave records, he succeeded in giving not only a logical explanation as to why waves are irregular, but a statement as well of the laws underlying the behavior of a seaway. There is indeed a basic law of the seaway. Contrary to the obvious inference from the quotation of Lord Rayleigh, the seaway can be described mathematically and precisely, albeit in a statistical way.

If Rice's work had been in the vein of generally accepted ideas in communication theory, the St Denis and Pierson paper represented a complete revolution in common naval practice. Nevertheless, its treatment of irregular water waves as, what is now called, a random field was almost immediately accepted, and set a standard for much of naval architecture. A parallel theory was developed in England where M.S. Longuet-Higgins studied the geometric properties of random wave fields. His long paper [85] from 1957 is still much cited in oceanography, optics, and statistics.

One possible reason for the success of the St Denis-Pierson approach can be that the authors succeeded to formulate and analyze the motions of a ship that moved with constant speed through the random field. The random sea could directly be used as input to a linear (later also non-linear) filter representing the ship. Linear filters are the topic of Chapters 4 and 5.

St Denis and Pierson extended the one-dimensional description of a time dependent process $\{x(t), t \in \mathbb{R}\}$, useful for example to model the waves measured at a single point, to a random field $x(t, (s_1, s_2))$ with time parameter t and location parameter (s_1, s_2) . They generalized the sum (1.15) to be a sum of directed waves, with $\boldsymbol{\omega} = (\omega, \kappa_1, \kappa_2)$,

$$\sum_{\boldsymbol{\omega}} A_{\boldsymbol{\omega}} \cos(\omega t - \kappa_1 s_1 - \kappa_2 s_2 + \phi_{\boldsymbol{\omega}}), \quad (1.16)$$

with random amplitudes and phases.

For fixed t each element in (1.16) is a cosine-function in the plane, which is zero along lines $\omega t - \kappa_1 s_1 - \kappa_2 s_2 + \phi_{\boldsymbol{\omega}} = \pi/2 + k\pi$, k integer. The parameters κ_1 and κ_2 , called the *wave numbers*, determine the direction and wave length of the waves and ω is the wave (*angular*) *frequency*, when the sea is observed at a fixed location (s_1, s_2) . More on this will follow in Chapter 7.

1.5.4 Detection theory and statistical inference

The first three landmarks illustrated the relation between stochastic model building and physical knowledge, in particular how the concepts of statistical independence and dependence between signal and functions relate to the physical world. About the same time as Rice and St Denis & Pierson advanced physically based stochastic modeling, the statistical inference methodology was placed firmly into a theoretical mathematical framework, as it was by then documented by the classical book by Harald Cramér, *Mathematical Methods of Statistics*, 1945 [32].

A few years later, the connection between the theoretical basis for statistical inference and important engineering questions related to signal detection was elegantly exploited by Ulf Grenander in his PhD thesis from Stockholm, *Stochastic processes and statistical inference* [56]. The classical problem in signal processing of deciding whether a deterministic signal of known shape $s(t)$ is present in an environment of Gaussian dependent, colored as opposed to white, random noise $x(t)$ can be treated as an infinite dimensional decision problem, testing an infinite dimensional statistical hypothesis; see, e.g., the standard work [117] on detection theory.

Suppose one observes a Gaussian stochastic process $x(t)$, $a \leq t \leq b$, with known correlation structure, but with unknown mean value function $m(t)$. If no signal is present, the mean value is 0, but with signal, the mean is equal to a function $s(t)$ of known shape. In statistical terms one has to test the following two hypotheses against each other:

$$H_0 : m(t) = 0,$$

$$H_1 : m(t) = s(t).$$

Grenander introduced a series of independent Gaussian *observables*,

$$y_k = \int h_k(t)x(t) dt,$$

by choosing the filter functions h_k as solutions to the integral equation

$$\int r(s,t)h_k(t) dt = c_k h_k(s),$$

with $c_1 \geq c_2 \geq \dots, c_k \rightarrow 0$ as $k \rightarrow \infty$. Under H_1 the observables will have mean $a_k = \int h_k(t)m(t) dt$ and variance c_k , while under H_0 they will have mean 0, and the same variances. So instead of a continuous problem, we have gotten a denumerable problem, in which one can make a Likelihood-Ratio test of the two alternatives. We will return to this problem in Section 5.3.4.

Exercises

- 1:1. Consider the sample space $\Omega = [0, 1]$ with uniform probability P , $P([a, b]) = b - a$, $0 \leq a \leq b \leq 1$. Construct a stochastic process $\mathbf{y} = (x_1, x_2, \dots)$ on Ω such that the components are independent zero-one variables, with $P(x_k = 0) = P(x_k = 1)$.
- 1:2. Show that $\mathcal{B}(\mathbb{R})$ = the class of Borel sets in \mathbb{R} is generated by
- the open intervals,
 - the closed intervals.
- 1:3. Prove the claim in Lemma 1.1 that the inverse $x^{-1}(\mathcal{B}) \in \mathcal{F}$.
- 1:4. A set $A \subset \{1, 2, \dots\}$ is said to have asymptotic density θ if

$$\lim_{n \rightarrow \infty} n^{-1} |A \cap \{1, 2, \dots, n\}| = \theta.$$

(Note, $|B|$ denotes the number of elements in B .) Let \mathcal{A} be the family of sets for which the asymptotic density exists. Is \mathcal{A} a field?

- 1:5. Take \mathbb{R}^n and show that the family \mathcal{F}_0 whose elements are unions of finitely many rectangles $(a_i, b_j]$ (with possibly infinite end points) is a field. Then, let T be an interval and convince yourself that the finite dimensional rectangles in \mathbb{R}^T and unions of finitely many such rectangles, is a field.
- 1:6. Show that the co-ordinate process $x_k(\omega) = \omega_k$ is a stochastic process on $(\mathbb{R}^\infty, \mathcal{B}^\infty)$.
- 1:7. Let \mathbb{Z} be the integers, and \mathcal{A} the family of subsets A , such that either A or its complement A^c is finite. Let $P(A) = 0$ in the first case and $P(A) = 1$ in the second case. Show that P can not be extended to a probability to $\sigma(\mathcal{A})$, the smallest σ -field that contains \mathcal{A} .
- 1:8. Take $T = [0, 1]$, and consider the set of functions which are continuous on the rational numbers, i.e.,

$$C_Q = \{x \in \mathbb{R}^T; x(q) \rightarrow x(q_0) \text{ for all rational numbers } q_0\},$$

where the limit is taken as q tends to q_0 through the rational numbers. Show that $C_Q \in \mathcal{B}_T$.

- 1:9. Prove that the increments of a Wiener process, as defined as in Definition 1.7, are independent and normal.
- 1:10. Prove the following useful inequality valid for any non-negative, integer-valued random variable N ,

$$E(N) - \frac{1}{2}E(N(N-1)) \leq P(N > 0) \leq E(N).$$

Generalize it to the following inequalities, where

$$\alpha_i = E(N(N-1) \cdot \dots \cdot (N-i+1))$$

is the i^{th} factorial moment,

$$\frac{1}{k!} \sum_{i=0}^{2n-1} (-1)^i \frac{1}{i!} \alpha_{(k+i)} \leq P(N=k) \leq \frac{1}{k!} \sum_{i=0}^{2n} (-1)^i \frac{1}{i!} \alpha_{(k+i)}.$$

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