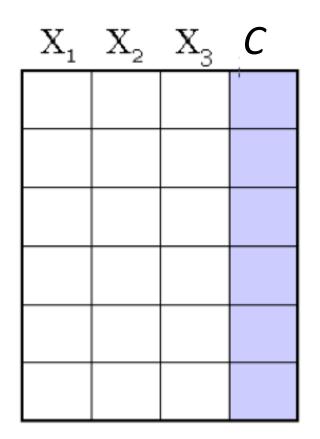
# **UVA CS 4501: Machine Learning**

# Lecture 13 Extra: More about Logistic Regression

Dr. Yanjun Qi

University of Virginia

Department of Computer Science



## A Dataset for classification

$$f:[X]\longrightarrow [c]$$

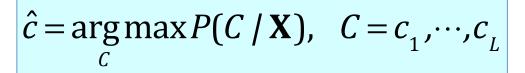
Output as Discrete
Class Label
C<sub>1</sub>, C<sub>2</sub>, ..., C<sub>1</sub>

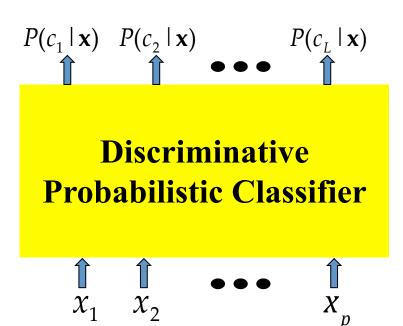
$$\underset{C}{\operatorname{arg\,max}} P(C \mid \mathbf{X}) \quad C = c_{1}, \dots, c_{L}$$

- Data/points/instances/examples/samples/records: [rows]
- Features/attributes/dimensions/independent variables/covariates/predictors/regressors: [ columns, except the last]
- Target/outcome/response/label/dependent variable: special column to be predicted [ last column ]

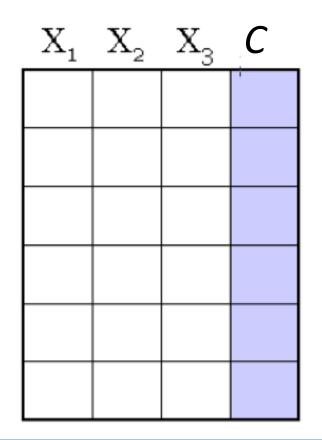
# Establishing a probabilistic model for classification

#### Discriminative model





$$\mathbf{X} = (X_1, X_2, \dots, X_p)$$



## A Dataset for classification

$$f:[X]\longrightarrow [c]$$

Output as Discrete
Class Label
C<sub>1</sub>, C<sub>2</sub>, ..., C<sub>1</sub>

Discriminative

$$\underset{c}{\operatorname{arg\,max}} P(C \mid \mathbf{X}) \quad C = c_{1}, \dots, c_{L}$$

Generative

$$\underset{C}{\operatorname{argmax}} P(C \mid X) = \underset{C}{\operatorname{argmax}} P(X, C) = \underset{C}{\operatorname{argmax}} P(X \mid C) P(C)$$

- Data/points/instances/examples/samples/records: [rows]
- Features/attributes/dimensions/independent variables/covariates/predictors/regressors: [ columns, except the last]
- Target/outcome/response/label/dependent variable: special column to be predicted [ last column ]

Later!

## **Today: Extra**

- Bayes Classifier and MAP Rule?
  - Bayes Classifier
  - Empirical Prediction Error
  - 0-1 Loss function for Bayes Classifier
- ✓ Logistic regression

Parameter Estimation for LR

e BX 1+eBX 13 B

### Bayes classifiers

Treat each feature attribute and the class label as random variables.

### Bayes classifiers

 Treat each feature attribute and the class label as random variables.

- Given a sample **x** with attributes  $(x_1, x_2, ..., x_p)$ :
  - Goal is to predict its class C.
  - Specifically, we want to find the value of  $C_i$  that maximizes  $p(C_i \mid x_1, x_2, ..., x_p)$ .

### Bayes classifiers

 Treat each feature attribute and the class label as random variables.

- Given a sample **x** with attributes  $(x_1, x_2, ..., x_p)$ :
  - Goal is to predict its class  $C. \longrightarrow \{C_1, C_1, ..., C_L\}$
  - Specifically, we want to find the value of  $C_i$  that maximizes  $p(C_i \mid x_1, x_2, ..., x_p)$ .
- Can we estimate  $p(C_i | \mathbf{x}) = p(C_i | x_1, x_2, ..., x_p)$  directly from data?

## Bayes classifiers→ MAP classification rule

- Establishing a probabilistic model for classification
- → MAP classification rule
  - MAP: Maximum A Posterior

## Bayes classifiers→ MAP classification rule

- Establishing a probabilistic model for classification
- → MAP classification rule
  - MAP: Maximum A Posterior
  - Assign x to  $c^*$  if

$$P(C = c^* | \mathbf{X} = \mathbf{x}) > P(C = c | \mathbf{X} = \mathbf{x})$$
  
for  $c \neq c^*$ ,  $c = c_1, \dots, c_L$ 

## Bayes classifiers→ MAP classification rule

- Establishing a probabilistic model for classification
- → MAP classification rule
  - MAP: Maximum A Posterior
  - Assign x to  $c^*$  if

$$P(C=c^*|X=x) > P(C=c|X=x) \quad (c \neq c^*, c = c_1, \dots, c_L)$$

$$P(C=c^*|X=x) > P(C=c|X=x) \quad (c \neq c^*, c = c_1, \dots, c_L)$$

$$P(C=c^*|X=x) > P(C=c|X=x) \quad (c \neq c^*, c = c_1, \dots, c_L)$$

$$P(C=c^*|X=x) > P(C=c|X=x) \quad (c \neq c^*, c = c_1, \dots, c_L)$$

$$P(C=c^*|X=x) > P(C=c|X=x) \quad (c \neq c^*, c = c_1, \dots, c_L)$$

### Bayes Classifiers – MAP Rule

*Task*: Classify a new instance X based on a tuple of attribute values  $X = \left\langle X_1, X_2, \dots, X_p \right\rangle$  into one of the classes

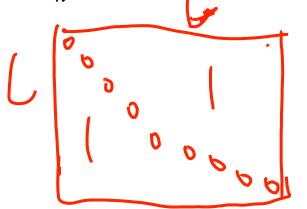
$$c_{MAP} = \underset{c_j \in C}{\operatorname{argmax}} P(c_j \mid x_1, x_2, \dots, x_p)$$

WHY?

MAP = Maximum Aposteriori Probability

#### 0-1 LOSS for Classification

- Procedure for categorical output variable C  $\forall k=1, k(k,k)=0$
- Frequently, 0-1 loss function used:  $L(k, \ell)$ if  $k \neq \ell$ ,  $L(k, \ell) = \ell$
- $L(k, \ell)$  is the price paid for misclassifying an element from class  $C_k$  as belonging to class  $C_{\ell}$ 
  - $\rightarrow L*L matrix$



C1, 63, ..., C/

### Expected prediction error (EPE)

• Expected prediction error (EPE), with expectation taken w.r.t. the joint distribution Pr(C,X)

$$-\Pr(C,X) = \Pr(C \mid X) \Pr(X)$$

$$\nearrow \text{?.9.} \text{o-(l.s)}$$

$$\text{EPE}(f) = E_{X,C}(L(C,f(X)))$$

$$= E_X \sum_{k=1}^{L} L[C_k, f(X)] Pr(C_k | X)$$

Consider sample population distribution

Dr. Yanjun Qi / UVA CS / s18  $EPE(f) = E_{X,C}(L(C,f(X)))$ = Ez Eciz [L(c,f(z))] z Disarte RV's Expetention E [ [Ck, f(x)] Pr(Ck | X) point wise minization  $f(X=x) = argmin \sum_{k=1}^{\infty} L((k,f(x))) Pr((k|X=x))$   $f(x) \in C$  $\Rightarrow$  f(x) = argmax Pr(Ck|X=x)3/15/18

### Expected prediction error (EPE)

$$EPE(f) = E_{X,C}(L(C, f(X))) = E_X \sum_{k=1}^{K} L[C_k, f(X)] Pr(C_k | X)$$

Consider sample population distribution

- Pointwise minimization suffices
- $\rightarrow$  simply  $\hat{f}(X) = \operatorname{argmin}_{g \in C} \sum_{k=1}^{K} L(C_k, g) \Pr(C_k | X = x)$

$$\hat{f}(X) = C_k$$
 if  

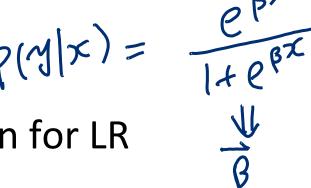
$$\Pr(C_k | X = x) = \max_{g \in C} \Pr(g | X = x)$$

#### SUMMARY: WHEN EPE USES DIFFERENT LOSS

Loss Function	Estimator $\hat{f}(x)$
$L_2 \qquad \stackrel{L(\epsilon)}{\longleftarrow} \epsilon$	$FRE = E_{X,Y} (Y - f(X))^{2}$ $\widehat{f}(x) = E[Y X = x]$
$L_1 \qquad \qquad \bullet \epsilon$	$\widehat{f}(x) = \text{median}(Y X=x)$
$\begin{array}{c c}  & L(\epsilon) \\ \hline  & -\delta & \delta \\ \end{array}$	$\widehat{f}(x) = \arg\max_{Y} P(Y X=x)$ (Bayes classifier / MAP)

## **Today: Extra**

- ✓ Why Bayes Classification MAP Rule?
  - Empirical Prediction Error
  - 0-1 Loss function for Bayes Classifier
- ✓ Logistic regression
  - Parameter Estimation for LR



### Newton's method for optimization

- The most basic second-order optimization algorithm
- Updating parameter with  $\neg D: \Theta_{K1} = O_K V G_K$

#### Review: Hessian Matrix / n==2 case

Singlevariate

→ multivariate

1st derivative to gradient,

$$g = \nabla f = \begin{pmatrix} \frac{\partial f}{\partial x} \\ \frac{\partial f}{\partial y} \end{pmatrix}$$

2<sup>nd</sup> derivative to Hessian

$$H = \begin{pmatrix} \frac{\partial^2 f}{\partial x^2} & \frac{\partial^2 f}{\partial x \partial y} \\ \frac{\partial^2 f}{\partial x \partial y} & \frac{\partial^2 f}{\partial y^2} \end{pmatrix}$$

#### Review: Hessian Matrix

Suppose that  $f: \mathbb{R}^n \to \mathbb{R}$  is a function that takes a vector in  $\mathbb{R}^n$  and returns a real number. Then the **Hessian** matrix with respect to x, written  $\nabla_x^2 f(x)$  or simply as H is the  $n \times n$  matrix of partial derivatives,

$$\nabla_x^2 f(x) \in \mathbb{R}^{n \times n} = \begin{bmatrix} \frac{\partial^2 f(x)}{\partial x_1^2} & \frac{\partial^2 f(x)}{\partial x_1 \partial x_2} & \cdots & \frac{\partial^2 f(x)}{\partial x_1 \partial x_n} \\ \frac{\partial^2 f(x)}{\partial x_2 \partial x_1} & \frac{\partial^2 f(x)}{\partial x_2^2} & \cdots & \frac{\partial^2 f(x)}{\partial x_2 \partial x_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial^2 f(x)}{\partial x_n \partial x_1} & \frac{\partial^2 f(x)}{\partial x_n \partial x_2} & \cdots & \frac{\partial^2 f(x)}{\partial x_n^2} \end{bmatrix}.$$

#### Newton's method for optimization

Making a quadratic/second-order Taylor series approximation

$$oldsymbol{f}_{quad}(oldsymbol{ heta}) = f(oldsymbol{ heta}_k) + \mathbf{g}_k^T(oldsymbol{ heta} - oldsymbol{ heta}_k) + rac{1}{2}(oldsymbol{ heta} - oldsymbol{ heta}_k)^T \mathbf{H}_k(oldsymbol{ heta} - oldsymbol{ heta}_k)$$

Finding the minimum solution of the above right quadratic approximation (quadratic function minimization is easy!)

$$\widehat{S(\theta)} = \widehat{S(0\kappa)} + \widehat{J_{K}}(0 - 0\kappa) + \frac{1}{2}(0 - 0\kappa) + \frac{1}{2}(0 - 0\kappa) + \frac{1}{2}(0 - 0\kappa)^{T} + \frac{1}{2}(0 - 0\kappa)$$

$$\frac{1}{2}(0 + 0\kappa)^{T} + \frac{1}{2}(0 + 0\kappa)$$

$$\frac{1}{2}(0 + 0\kappa)^{T} + \frac{1}{2}(0 - 0\kappa)$$

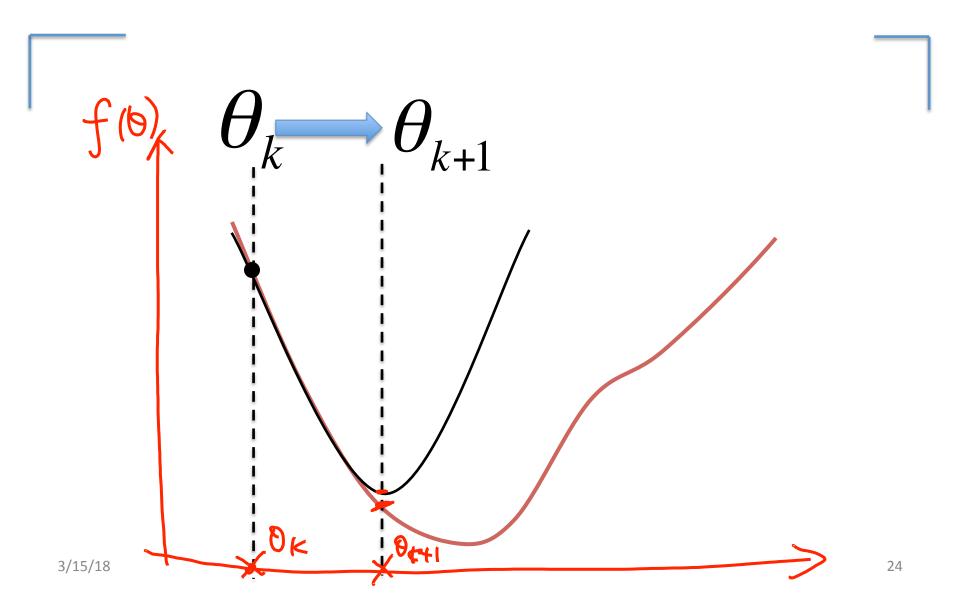
$$\frac{1}{2}(0 + 0\kappa)^{T} + \frac{1}{2}(0 + 0\kappa)$$

$$\frac{1}{2}($$

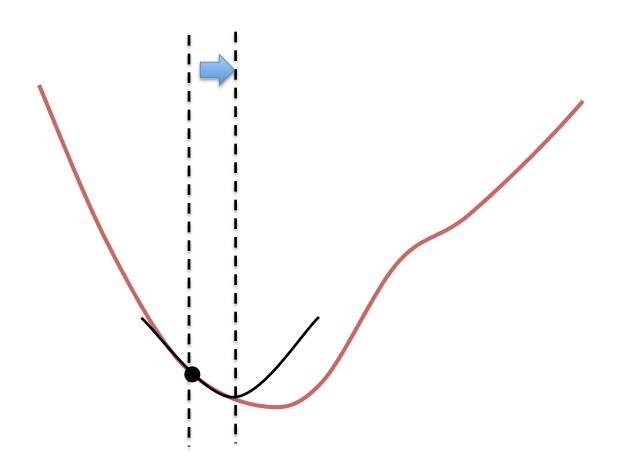
$$3k + Hk(0 - \theta k) = 0$$

$$\Rightarrow 0 = \theta k - Hkgk$$

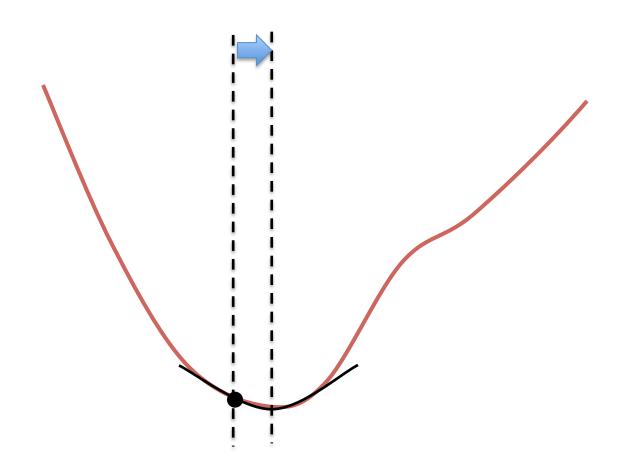
# Newton's Method / second-order cs/s18 Taylor series approximation



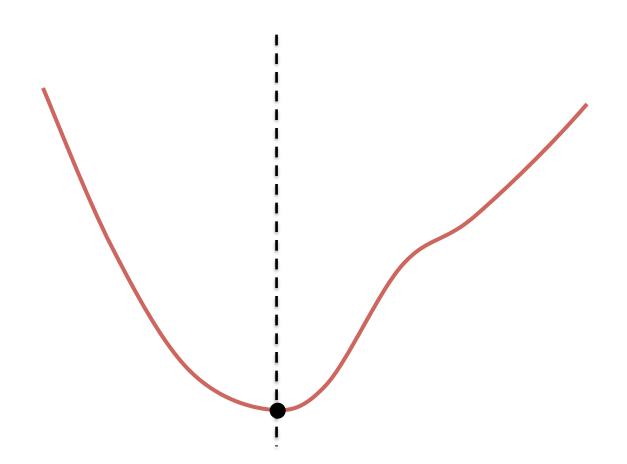
### Newton's Method / second-order Taylor series approximation



### Newton's Method / second-order Taylor series approximation



### Newton's Method / second-order Taylor series approximation



#### Newton's Method

At each step:

$$\theta_{k+1} = \theta_k - \frac{f'(\theta_k)}{f''(\theta_k)}$$

$$\theta_{k+1} = \theta_k - H^{-1}(\theta_k) \nabla f(\theta_k)$$

- Requires 1<sup>st</sup> and 2<sup>nd</sup> derivatives
- Quadratic convergence
- However, finding the inverse of the Hessian matrix is often expensive

### Newton vs. GD for optimization

Newton: a quadratic/second-order Taylor

series approximation 
$$\mathbf{f}_{quad}(\boldsymbol{\theta}) = f(\boldsymbol{\theta}_k) + \mathbf{g}_k^T(\boldsymbol{\theta} - \boldsymbol{\theta}_k) + \frac{1}{2}(\boldsymbol{\theta} - \boldsymbol{\theta}_k)^T\mathbf{H}_k(\boldsymbol{\theta} - \boldsymbol{\theta}_k)$$

GD: a approximation

Finding the minimum solution of the above right quadratic approximation (quadratic function minimization is easy!)

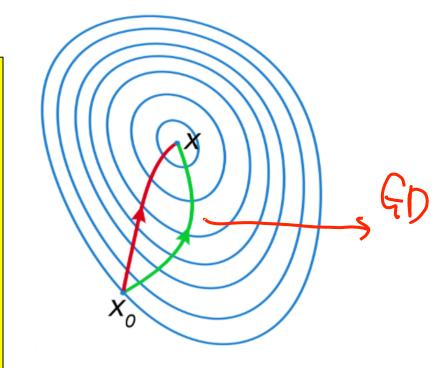
$$\begin{aligned} \widehat{f_{quad}}(\boldsymbol{\theta}) &= f(\boldsymbol{\theta}_k) + \mathbf{g}_k^T (\boldsymbol{\theta} - \boldsymbol{\theta}_k) + \frac{1}{2} (\boldsymbol{\theta} - \boldsymbol{\theta}_k)^T \frac{1}{\alpha} (\boldsymbol{\theta} - \boldsymbol{\theta}_k) \\ & \qquad \qquad \qquad \\ \mathbf{0}_{kkl} &= \mathbf{0}_{kl} - \mathbf{0}_{kl} \mathbf{g}(\mathbf{0}_{kl}) \end{aligned}$$

### Comparison

Newton's method vs. Gradient descent

A comparison of gradient descent (green) and Newton's method (red) for minimizing a function (with small step sizes).

Newton's method uses curvature information to get a more direct route ...



## Parameter Estimation for LR → MLE from the data

RECAP: Linear regression → Least squares

Logistic regression: Maximum likelihood estimation

#### MLE for Logistic Regression Training

Let's fit the logistic regression model for K=2, i.e., number of classes is 2

Training set:  $(x_i, y_i)$ , i=1,...,N

For Bernoulli distribution

$$p(y \mid x)^{y} (1-p)^{1-y}$$

(conditional) Log-likelihood:

$$l(\beta) = \sum_{i=1}^{N} \{ logPr(Y = y_i | X = x_i) \}$$

How? 
$$= \sum_{i=1}^{N} y_i \log(\Pr(Y=1|X=x_i)) + (1-y_i) \log(\Pr(Y=0|X=x_i))$$

$$= \sum_{i=1}^{N} (y_i \log \frac{\exp(\beta^T x_i)}{1 + \exp(\beta^T x_i)}) + (1-y_i) \log \frac{1}{1 + \exp(\beta^T x_i)})$$

$$= \sum_{i=1}^{N} (y_i \beta^T x_i - \log(1 + \exp(\beta^T x_i)))$$

 $x_i$  are (p+1)-dimensional input vector with leading entry 1 \beta is a (p+1)-dimensional vector

$$I(\beta) = \sum_{i=1}^{N} \{ \log \Pr(Y = y_i | X = x_i) \}$$

$$\int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \} \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{0}^{N} \{ Y = y_i | X = x_i \}$$

$$= \int_{$$

#### Newton-Raphson for LR (optional)

$$\frac{\partial l(\beta)}{\partial \beta} = \sum_{i=1}^{N} (y_i - \frac{\exp(\beta^T x)}{1 + \exp(\beta^T x)}) x_i = 0$$

(p+1) Non-linear equations to solve for (p+1) unknowns

Solve by Newton-Raphson method:

$$\beta^{new} \leftarrow \beta^{old} - \left[ \left( \frac{\partial^2 l(\beta)}{\partial \beta \partial \beta^T} \right) \right]^{-1} \frac{\partial l(\beta)}{\partial \beta},$$

where, 
$$\left(\frac{\partial^2 l(\beta)}{\partial \beta \partial \beta^T}\right) = -\sum_{i=1}^N x_i x_i^T \left(\frac{\exp(\beta^T x_i)}{1 + \exp(\beta^T x_i)}\right) \left(\frac{1}{1 + \exp(\beta^T x_i)}\right)$$

minimizes a quadratic approximation to the function we are really interested in.

$$oldsymbol{ heta}_{k+1} = oldsymbol{ heta}_k - \mathbf{H}_K^{-1} \mathbf{g}_k$$

 $p(x_i; \beta)$ 

 $1 - p(x_i; \beta)$ 

#### Newton-Raphson for LR...

$$\frac{\partial l(\beta)}{\partial \beta} = \sum_{i=1}^{N} (y_i - \frac{\exp(\beta^T x)}{1 + \exp(\beta^T x)}) x_i = X^T (y - p)$$

$$(\frac{\partial^2 l(\beta)}{\partial \beta \partial \beta^T}) = -X^T W X$$

So, NR rule becomes: 
$$\beta^{new} \leftarrow \beta^{old} + (X^T W X)^{-1} X^T (y-p),$$

$$X = \begin{bmatrix} x_1^T \\ x_2^T \\ \vdots \\ x_N^T \end{bmatrix}_{N-by-(p+1)}, y = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_N \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \exp(\beta^T x_2)/(1 + \exp(\beta^T x_2)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \exp(\beta^T x_2)/(1 + \exp(\beta^T x_2)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp($$

 $X: N \times (p+1)$  matrix of  $x_i$ 

 $y: N \times 1$  matrix of  $y_i$ 

 $p: N \times 1$  matrix of  $p(x_i; \beta^{old})$ 

 $W: N \times N$  diagonal matrix of  $p(x_i; \beta^{old})(1 - p(x_i; \beta^{old}))$ 

$$\left(\frac{\exp(\boldsymbol{\beta}^T \boldsymbol{x}_i)}{(1+\exp(\boldsymbol{\beta}^T \boldsymbol{x}_i))}\right)\left(1-\frac{1}{(1+\exp(\boldsymbol{\beta}^T \boldsymbol{x}_i))}\right)$$

#### Newton-Raphson for LR...

#### Newton-Raphson

$$-\beta^{new} = \beta^{old} + (X^T W X)^{-1} X^T (y - p) 
= (X^T W X)^{-1} X^T W (X \beta^{old} + W^{-1} (y - p)) 
= (X^T W X)^{-1} X^T W z$$

Re expressing Newton step as weighted least square step

Adjusted response

$$z = X\beta^{old} + W^{-1}(y-p)$$

Iteratively reweighted least squares (IRLS)

$$\beta^{new} \leftarrow \arg\min_{\beta} (z - X\beta^{T})^{T} W (z - X\beta^{T})$$

$$\leftarrow \arg\min_{\beta} (y - p)^{T} W^{-1} (y - p)$$

#### References

- Prof. Tan, Steinbach, Kumar's "Introduction to Data Mining" slide
  - ☐ Prof. Andrew Moore's slides
  - ☐ Prof. Eric Xing's slides
  - ☐ Prof. Ke Chen NB slides
  - ☐ Hastie, Trevor, et al. *The elements of statistical learning*. Vol. 2. No. 1. New York: Springer, 2009.