

Supporting Information: Experimental designs for testing the interactive effects of temperature and light in ecology and the problem of periodicity

Estimating the effects of experimental periodicity covariance mathematically

When experimental designs couple thermo- and photo- periodicity and therefore, forcing and photoperiod treatments covary, they may incorrectly estimate the effects of forcing, photoperiod, and (for experiments with crossed designs) their interaction. While robustly testing the effect of this experimental covariation requires additional experiments, we can solve for its potential effect based on the geometry of the experimental design given several assumptions.

If we assume forcing and photoperiod effects are additive and linear (i.e., there is no interaction), then the treatments and their resulting effects occupy a plane, and we can solve algebraically for the separate effects of forcing and photoperiod. If we replace the qualitative factor (high forcing/low forcing) by the quantitative effect of forcing (daily thermal sums) to properly account for the difference in forcing between short and long photoperiods, then the plane is defined on a three dimensional space with thermal sums on the x axis, photoperiod on the y axis, and the response on the z axis. Given the assumption of no interaction, we can solve for the expected response if forcing and photoperiod were fully independent.

Here, we work through an example based on the experiment described in Flynn & Wolkovich (2018). First, we redefine the forcing treatment (20/10°C vs. 15/5°C) as thermal sums (degree hours) to reflect the duration and intensity of temperature to define the x axis. The y axis is defined by the photoperiod treatment levels (8 hours vs. 12 hours), and the z axis is the response, defined as the change in days to budburst (Δ) for this experiment. With these axes, we can use a system of three equations to define a plane (Equations 1-3),

$$ax_1 + by_1 + cz_1 + d = 0 \tag{1}$$

$$ax_2 + by_2 + cz_2 + d = 0 \tag{2}$$

$$ax_3 + by_3 + cz_3 + d = 0 \tag{3}$$

and take three points that define the main effects in the experiment (low forcing/short day, low forcing/long day, high forcing/short day; i.e., the effect of daylength holding forcing constant, and the effect of forcing holding daylength constant) as the three points that define the plane:

$$200a + 8b + 0 + d = 0 \quad (4)$$

$$240a + 12b + 4.5c + d = 0 \quad (5)$$

$$320a + 8b + 9c + d = 0 \quad (6)$$

where Equation 4 is the result of the low forcing, short day treatment where thermal sum (x) is 200 degree hours, photoperiod (y) is 8 hours, and change in days to budburst is 0; Equation 5 is the result of the low forcing, long day treatment ($x = 240$ degree hours, $y = 12$ hours, $z = 4.5$ days); and Equation 6 is the result of the high forcing, short day treatment ($x = 320$ degree hours, $y = 8$ hours, $z = 9$ days) (see Flynn & Wolkovich (2018), Table S5). Algebraically, we simply solve for each unknown in turn.

Solving Equation 4 for d in terms of a and b

$$200a + 8b + 0 + d = 0 \text{ and solving for } d \text{ yields:} \quad (7)$$

$$-200a - 8b = d \quad (8)$$

Now, we can solve for c (in terms of a and b) using our solution for d :

$$240a + 12b + 4.5c + (-200a - 8b) = 0 \quad (9)$$

$$-40/4.5a + 4/4.5b = c \quad (10)$$

And similarly then solve for a given our last set of coordinates and solutions for d and c (not shown). This yields:

$$a = \frac{1}{5}b \quad (11)$$

$$c = -\frac{8}{3}b \quad (12)$$

$$d = -48b \quad (13)$$

Putting these back into the equation for a plane yields an overall equation to estimate the response z (days) for any combination of x (thermal sum) and y (photoperiod):

$$z = \frac{3}{40}x + \frac{3}{8}y - 18 \quad (14)$$

Using this equation, we can calculate the response for an existing treatment, e.g., low forcing/short day where thermal sum (x) is 200 degree hours and photoperiod (y) results in the measured days to budburst of (z) as zero. We can also calculate the expected response for a treatment that *was not* included in the experiment, e.g., a short day thermal sum with a long day photoperiod, $x =$

200 degree hours and $y = 12$ hours, to calculate the expected change in days to budburst z ,

$$\frac{3}{40} * (200) + \frac{3}{8} * (8) - 18 = 0 \quad (15)$$

$$\frac{3}{40} * (200) + \frac{3}{8} * (12) - 18 = 3 \quad (16)$$

showing that the effect of changing photoperiod from 8 to 12 hours *without* a change in forcing — given our assumptions — is a 3 day advance in budburst.

DB says: Does this mean that 1.5 days of the original 4.5 photoperiod effect is due to forcing? Previously we talked about it the other way around (ie 3 of the estimated 4.5 photo effect due to forcing)

Finally, we can calculate the response for high forcing/long day, given the assumption of no interaction:

$$\frac{3}{40} * (360) + \frac{3}{8} * (12) - 18 = 13.5 \quad (17)$$

Modeling methods

To investigate the effects of experimental photo- and thermo- period covariance on the estimated effects of photoperiod and forcing cues on spring phenology, we obtained data files from two published phenology studies Flynn & Wolkovich (2018) and Buonaiuto & Wolkovich (2021). Both studies used dormant twigs collected from Harvard Forest in Petersham, MA, USA (42.5314°N, 72.1900°W) and exposed them to comparable photoperiod and forcing treatments levels (photoperiod: 12 vs. 8 hours; forcing: 20/10°C vs. 15/5°C day/night (Flynn & Wolkovich, 2018); 24/18°C vs. 18/12°C day/night (Buonaiuto & Wolkovich, 2021)) in full factorial growth chamber treatments. We subset each date set to include only matching phenological observations (“leaf expansion” or BBCH 11 (Finn *et al.*, 2007)), species. This data subset included 713 observations ($n_{\text{Flynn\&Wolkovich (2018)}} = 384, n_{\text{Buonaiuto\&Wolkovich (2021)}} = 329, n_{\text{species}} = 7$).

We estimated the effect of study design on parameter estimates with a Bayesian hierarchical mixed effect model using weakly informative priors. We included forcing, photoperiod, study and their interactions as main effects, and species as a random effect. We ran the model using the R package “brms” (Bürkner, 2018) on 4 chains with 2000 iterations and warmup of 1000 iterations for a total of 4000 posterior per parameter. The model is written below.

$$\text{leafexpansion}_{[i]} \sim N(\alpha_{sp[i]} + \beta_{study} \times \beta_{forcing} \times \beta_{photoperiod}, \sigma_y^2)$$

We modeled the intercept (α) at the species level using the formula:

$$\alpha_{x_{sp}} \sim N(\mu_x, \sigma_x^2)$$

We note that the because the subset of overlapping data between the two studies is considerably smaller than it is not surprising that the relative effect sizes of photoperiod and forcing estimated here differ from those reported in the full, published studies.

References

- Buonaiuto, D.M. & Wolkovich, E.M. (2021) Differences between flower and leaf phenological responses to environmental variation drive shifts in spring phenological sequences of temperate woody plants. *Journal of Ecology* **109**, 2922–2933.
- Bürkner, P.C. (2018) Advanced bayesian multilevel modeling with the r package brms. *R Journal* **10**, 395–411.
- Finn, G.A., Straszewski, A.E. & Peterson, V. (2007) A general growth stage key for describing trees and woody plants. *Annals of Applied Biology* **151**, 127–131.
- Flynn, D.F.B. & Wolkovich, E.M. (2018) Temperature and photoperiod drive spring phenology across all species in a temperate forest community. *New Phytologist* **219**, 1353–1362.