

Different Alternative Types

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Abstract

I compare SUCCOTASH to various competitors under the same alternative scenarios in Stephens (2016). SUCCOTASH generally has superior performance in terms of (1) estimating π_0 , (2) having higher AUC, and (3) having lower MSE. LEAPP does better under some scenarios in terms of estimating π_0 but has far worse AUC and MSE. SUCCOTASH accurately estimates the average sign error rate, though not in as conservative manner as one might hope in some scenarios. I also look at limma-shrinking the variances in RUVASH. It doesn't change the results too much. It seems to just make the estimates of π_0 slightly more conservative.

Simulation Setup

I ran through 200 repetitions of generating data from GTEX muscle data under the following parameter conditions:

- $n \in \{10, 20, 40\}$,
- $p = 1000$.
- $\pi_0 \in \{0.5, 0.9\}$,
- The alternative distribution being either spiky, near-normal, flattop, skew, big-normal, or bimodal, where these are the same alternatives defined in Stephens (2016) and the following table. New alternatives are generated every iteration.

Scenario	Alternative Distribution
Spiky	$0.4N(0, 0.25^2) + 0.2N(0, 0.5^2) + 0.2N(0, 1^2), 0.2N(0, 2^2)$
Near Normal	$2/3N(0, 1^2) + 1/3N(0, 2^2)$
Flattop	$(1/7)N(-1.5, .5^2) + N(-1, .5^2) + N(-.5, .5^2) + N(0, .5^2) + N(0.5, .5^2) + N(1.0, .5^2) + N(1.5, .5^2)$
Skew	$(1/4)N(-2, 2^2) + (1/4)N(-1, 1.5^2) + (1/3)N(0, 1^2) + (1/6)N(1, 1^2)$
Big-normal	$N(0, 4^2)$
Bimodal	$0.5N(-2, 1^2) + 0.5N(2, 1^2)$

I extracted the most expressed p genes from the GTEX muscle data and n samples are chosen at random. Half of these samples are randomly given the “treatment” label 1, the other half given the “control” label 0. Of the p genes, $\pi_0 p$ were chosen to be non-null. Signal was added by a Poisson-thinning approach, where the log-2 fold change was sampled from one of five the alternative models above. That is

$$A_1, \dots, A_{p/2} \sim f \tag{1}$$

$$B_i = 2^{A_i} \text{ for } i = 1, \dots, p/2, \tag{2}$$

where f is from the table above. If $A_i > 0$ then we replace $Y_{[1:(n/2), i]}$ with $\text{Binom}(Y_{[j, i]}, 1/B_i)$ for $j = 1, \dots, n/2$. If $A_i < 0$ then we replace $Y_{[(n/2+1):n, i]}$ with $\text{Binom}(Y_{[j, i]}, B_i)$ for $j = n/2 + 1, \dots, n$.

I now describe the justification for this. Suppose that

$$Y_{ij} \sim \text{Poisson}(\lambda_j). \quad (3)$$

Let x_i be the indicator of treatment vs control for individual i . Let Ω be the set of non-null genes. Let Z be the new dataset derived via the steps above. That is

$$Z_{ij}|Y_{ij} = \begin{cases} \text{Binom}(Y_{ij}, 2^{A_j x_i}) & \text{if } A_j < 0 \text{ and } j \in \Omega \\ \text{Binom}(Y_{ij}, 2^{-A_j(1-x_i)}) & \text{if } A_j > 0 \text{ and } j \in \Omega \\ Y_{ij} & \text{if } j \notin \Omega. \end{cases} \quad (4)$$

Then

$$Z_{ij}|A_j, A_j < 0, j \in \Omega \sim \text{Poisson}(2^{A_j x_i} \lambda_j) \quad (5)$$

$$Z_{ij}|A_j, A_j > 0, j \in \Omega \sim \text{Poisson}(2^{-A_j(1-x_i)} \lambda_j), \quad (6)$$

and

$$E[\log_2(Z_{ij}) - \log_2(Z_{kj})|A_j, A_j < 0, j \in \Omega] \approx A_j x_i - A_j x_k, \text{ and} \quad (7)$$

$$E[\log_2(Z_{ij}) - \log_2(Z_{kj})|A_j, A_j > 0, j \in \Omega] \approx -A_j(1-x_i) + A_j(1-x_k). \quad (8)$$

if individual i is in the treatment group and individual k is in the control group, then this just equals A_j . I treat the A_j 's as the true coefficient values when calculating the MSE below.

Methods

I first normalized the counts by $\log_2(COUNTS + 1)$. The number of hidden confounders was estimated using the methods of Buja and Eyuboglu (1992) implemented in the `num.sv()` function in the `sva` package in R.

The confounder adjustment methods I look at in this write-up are:

- OLS + qvalue.
- RUVols + estimate variance inflation using controls + ASH
- RUVgls + estimate variance inflation using controls + ASH
- RUVgls + estimate variance inflation using controls + MLE to UMVUE motivated scaling + ASH
- RUVgls + limma shrink variances + estimate variance inflation using controls + MLE to UMVUE motivated scaling + ASH
- SUCCOTASH using normal mixtures and heteroscedastic PCA as the factor-analysis method.
- The robust regression version of CATE using PCA as the factor analysis method + qvalue.
- SVA + qvalue.
- Negative control version of CATE using PCA as the factor analysis method + qvalue.
- RUV2 + qvalue.
- RUV4 + qvalue.
- Sparse version of LEAPP. Since this is a sparsity-inducing procedure, I used the proportion of zeros as the estimate of π_0 .
- Ridge version of LEAPP + qvalue.

RUVASH

The model for the second step in RUV4, LEAPP, CATE, and SUCCOTASH is

$$Y = \beta_{p \times 1} + \hat{\alpha}_{p \times q} Z_{q \times 1} + E_{p \times 1} \quad (9)$$

$$E \sim N_p(0, \lambda \hat{\Sigma}) \quad (10)$$

$$\hat{\Sigma} = \text{diag}(\hat{\sigma}_1^2, \dots, \hat{\sigma}_p^2), \quad (11)$$

with $\hat{\Sigma}$ and $\hat{\alpha}$ known. λ is a variance inflation parameter that is novel to SUCCOTASH.

Let $\mathcal{C} \subseteq \{1, \dots, p\}$ denote the indices for the negative controls. I.e. $\beta_{\mathcal{C}} = 0$. Then the model for the negative controls is

$$Y_{\mathcal{C}} = \alpha_{\mathcal{C}} Z + E_{\mathcal{C}}. \quad (12)$$

The maximum likelihood estimates of Z and λ are easy:

$$\hat{Z} = (\alpha_{\mathcal{C}}^T \hat{\Sigma}^{-1} \alpha_{\mathcal{C}})^{-1} \alpha_{\mathcal{C}}^T \hat{\Sigma}^{-1} Y_{\mathcal{C}}. \quad (13)$$

Let $R_{\mathcal{C}} = Y_{\mathcal{C}} - \hat{\alpha}_{\mathcal{C}} \hat{Z}$. Then the MLE of λ is

$$\hat{\lambda} = \frac{1}{|\mathcal{C}|} \sum_{i \in \mathcal{C}} R_i^2 / \hat{\sigma}_i^2. \quad (14)$$

RUVASH runs ASH with means $\hat{\beta} = Y - \hat{\alpha} \hat{Z}$ and variances $\hat{\lambda} \hat{\sigma}_i^2$.

In the same way that I made an ad-hoc correction to the SUCCOTASH variance inflation parameter, I also explore using this for RUVASH. Specifically, I set

$$\tilde{\lambda} = \frac{n}{n - k - q} \hat{\lambda} \quad (15)$$

and run ASH with variances $\tilde{\lambda} \hat{\sigma}_i^2$. These are called SRUVASH in the plots below.

Results

Note that in the plots below, n refers to the size of each group, not the total size.

Estimates of π_0

- RUV + ASH Works much better here than vanilla ASH at estimating π_0 , but it is slightly anti-conservative.
- RUV + ad-hoc scaling + ASH has conservative FDR with high probability in almost every scenario. The only one where it isn't conservative is the Bimodal alternative scenario.

- SUCCOTASH has slightly anti-conservative estimates of π_0 in the Flat-top and bimodal Scenarios. It does well for every other scenario for larger n .
- LEAPP does amazingly well in the Big-normal and bimodal scenarios, even for $n = 10$. However, it is far too conservative in every other non-null scenario. This seems to indicate that LEAPP functions best if there is a separation of the alternative signal from zero.
- No method using q -value ever performed as well as succotash. Indeed, none exhibited this “conservative with high-probability” behavior that is desirable.
- Limma-shrinking the variances results in slightly more conservative estimates of π_0 .

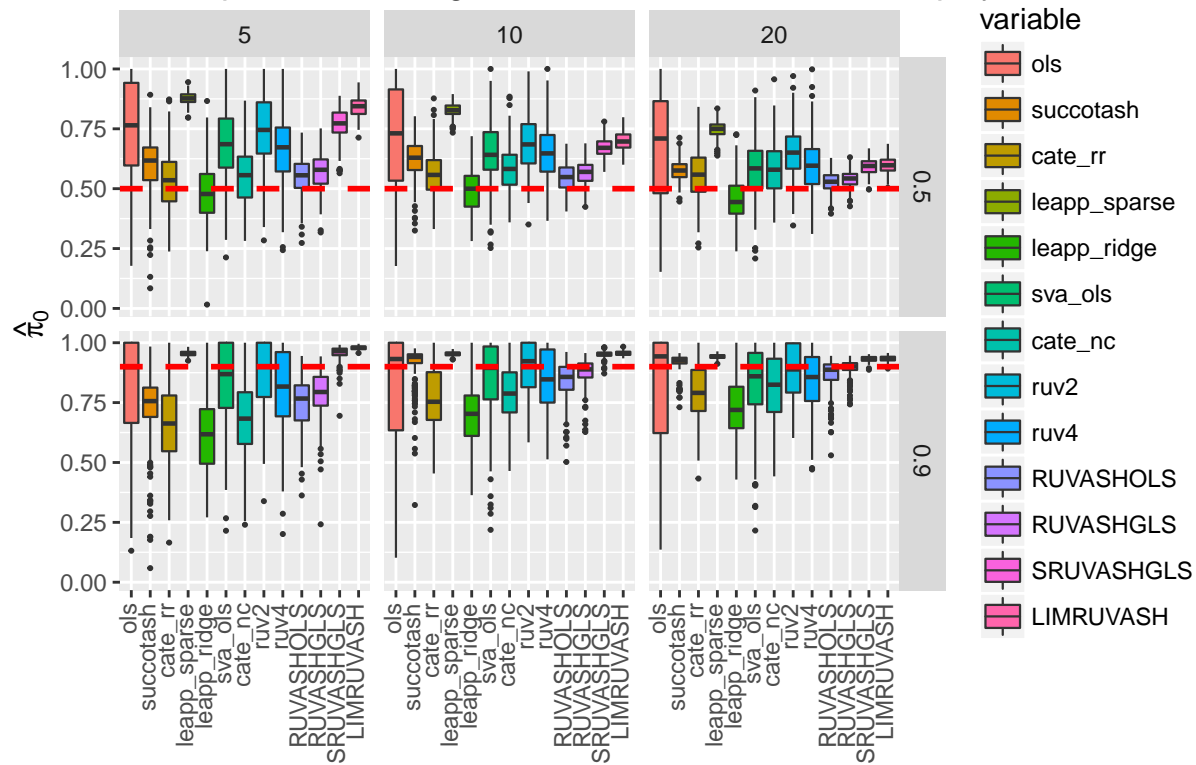
AUC performance.

- SUCCOTASH always has higher AUC, even in the bimodal and big-normal scenarios where LEAPP estimated π_0 more accurately.
- All of the RUV + ASH methods have about the same AUC as SUCCOTASH — including the one with the ad-hoc scaling value that works so well with estimating π_0 .

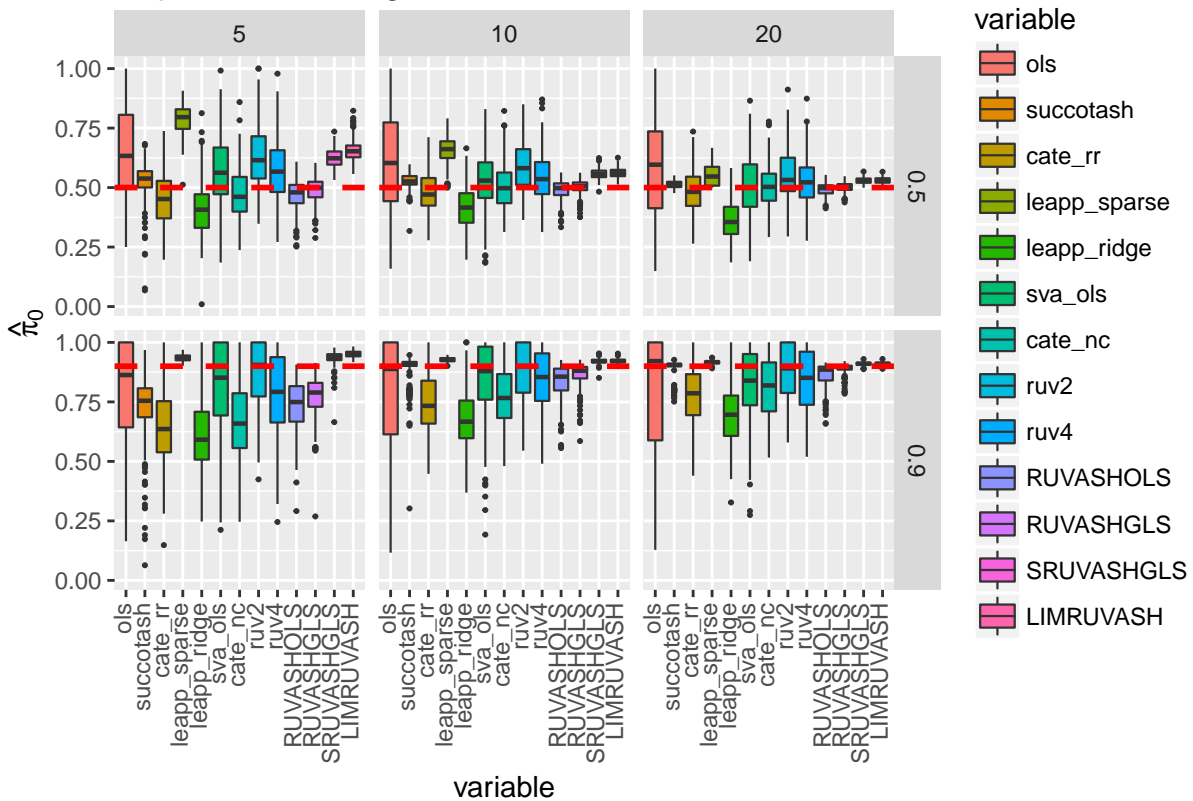
MSE

- SUCCOTASH has superior performance in term of MSE compared to other methods.
- LEAPP has terrible MSE performance in all cases except the all-null setting.
- Even then, SUCCOTASH has better performance.
- RUV + ASH has pretty good MSE performance.
- The ad-hoc scaling version of RUV + ASH seems to suffer in MSE a little bit, especially at low sample sizes.

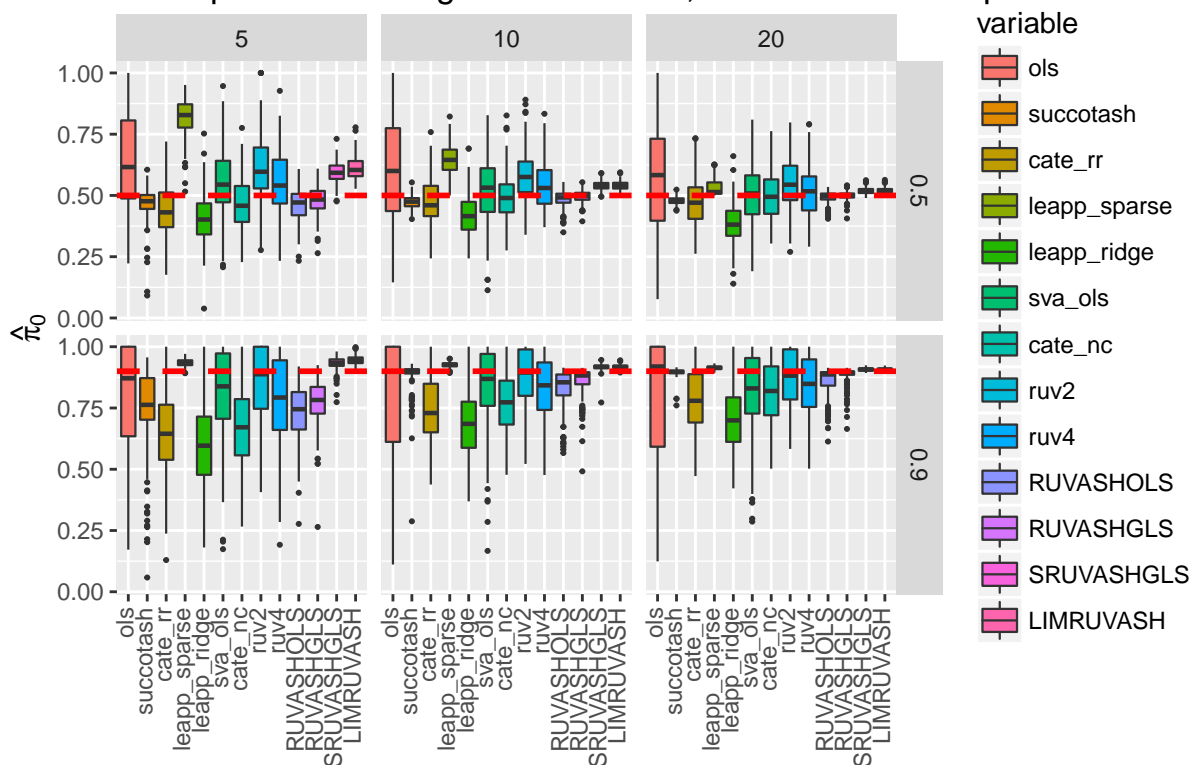
Estimates of π_0 When Using Muscle Tissue, Alternative = spiky



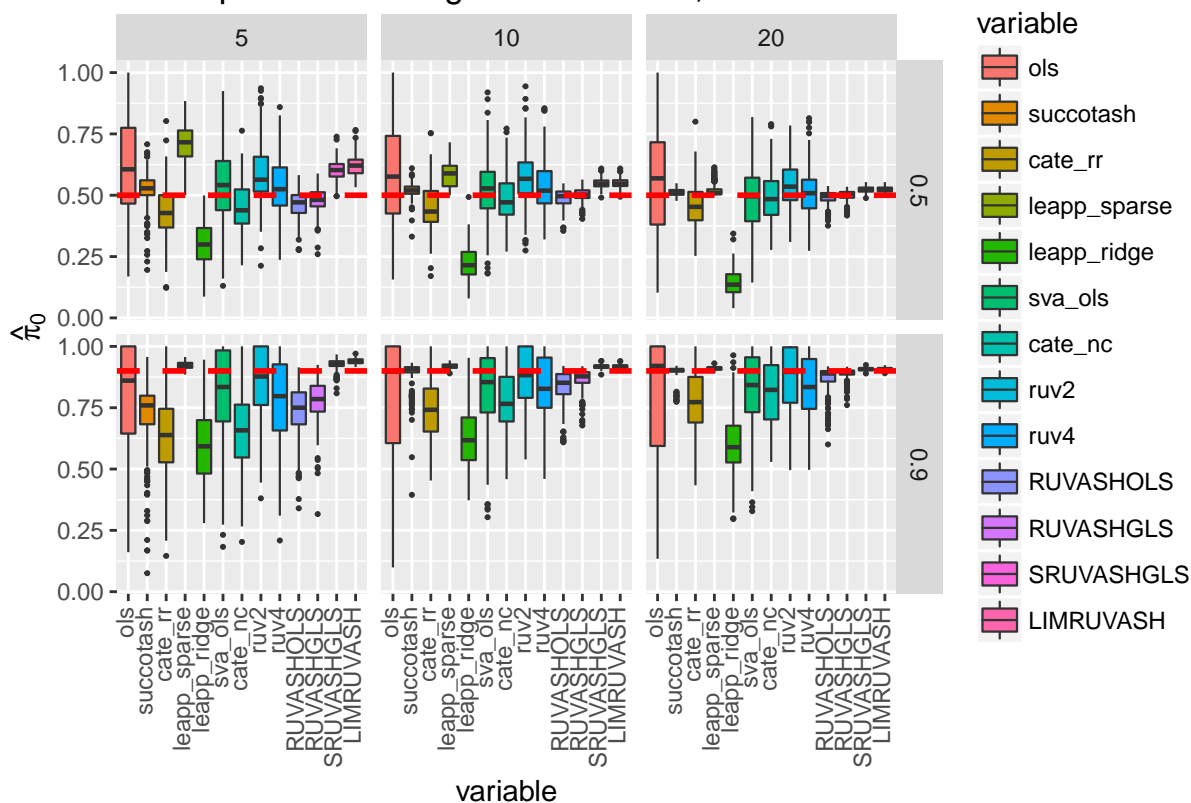
Estimates of π_0 When Using Muscle Tissue, Alternative = near_normal



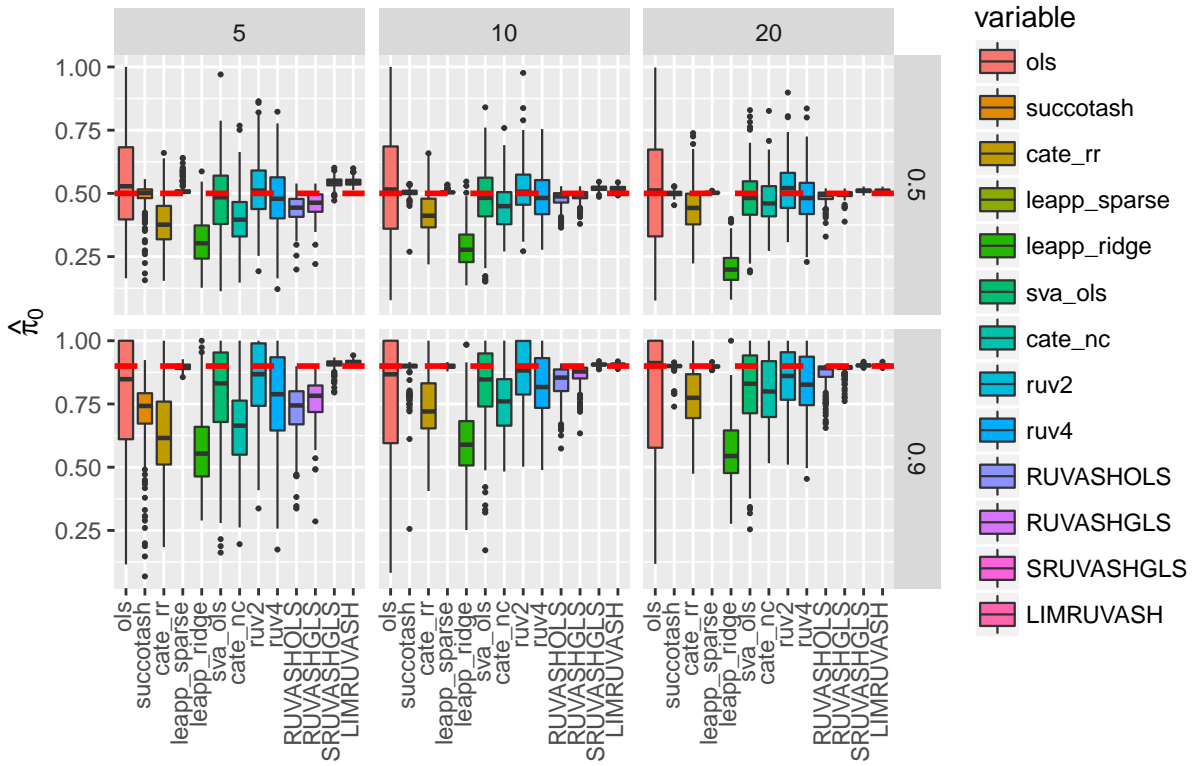
Estimates of π_0 When Using Muscle Tissue, Alternative = flattop



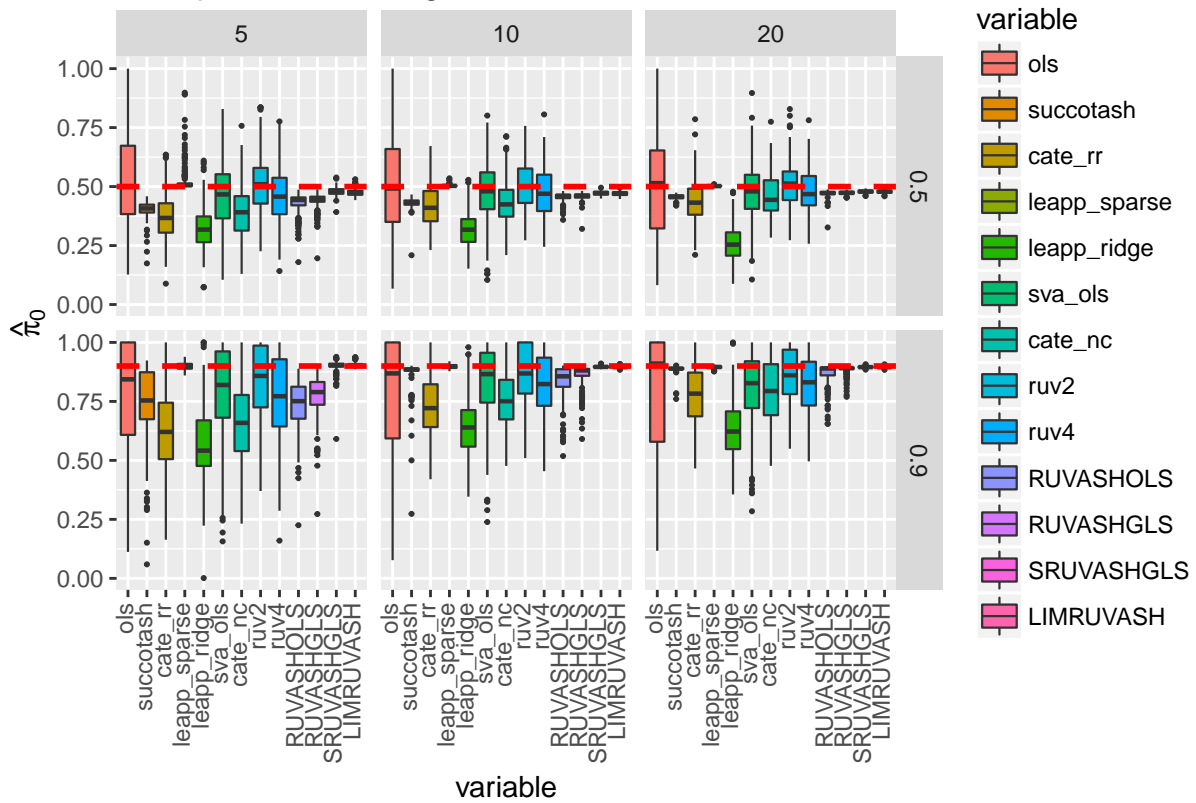
Estimates of π_0 When Using Muscle Tissue, Alternative = skew



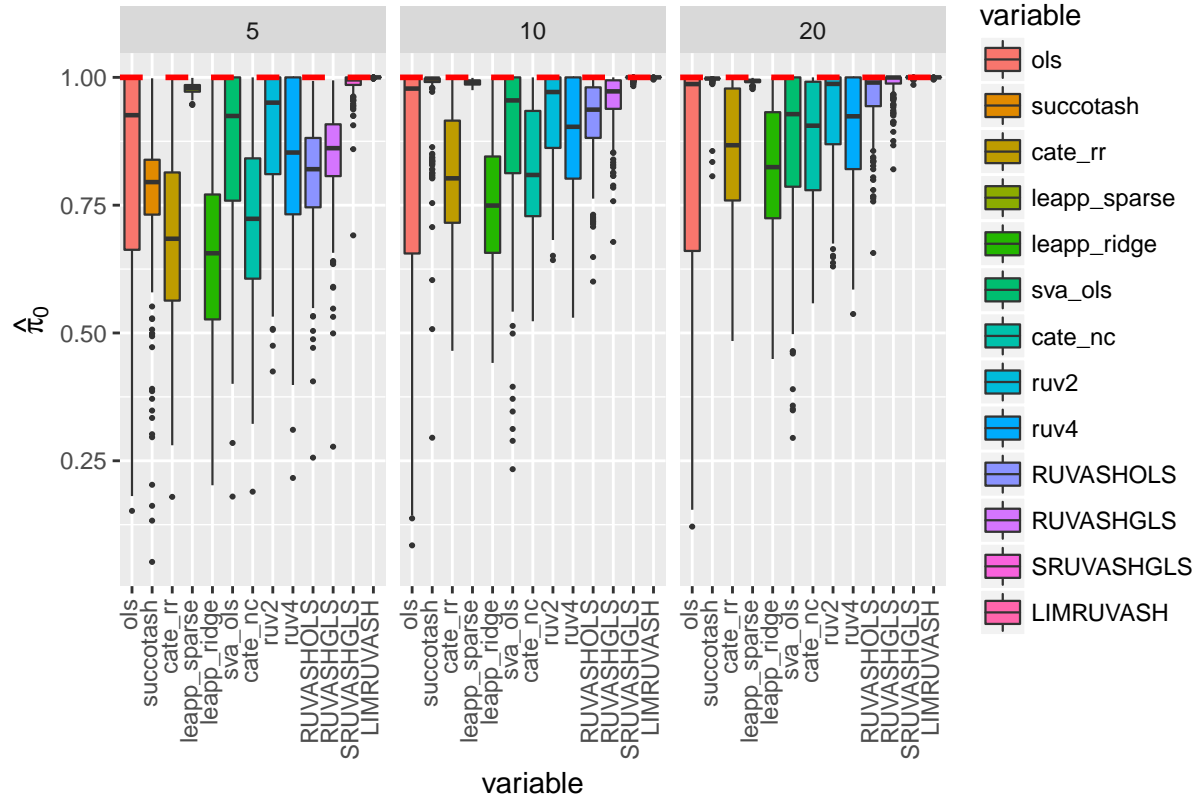
Estimates of π_0 When Using Muscle Tissue, Alternative = big_normal



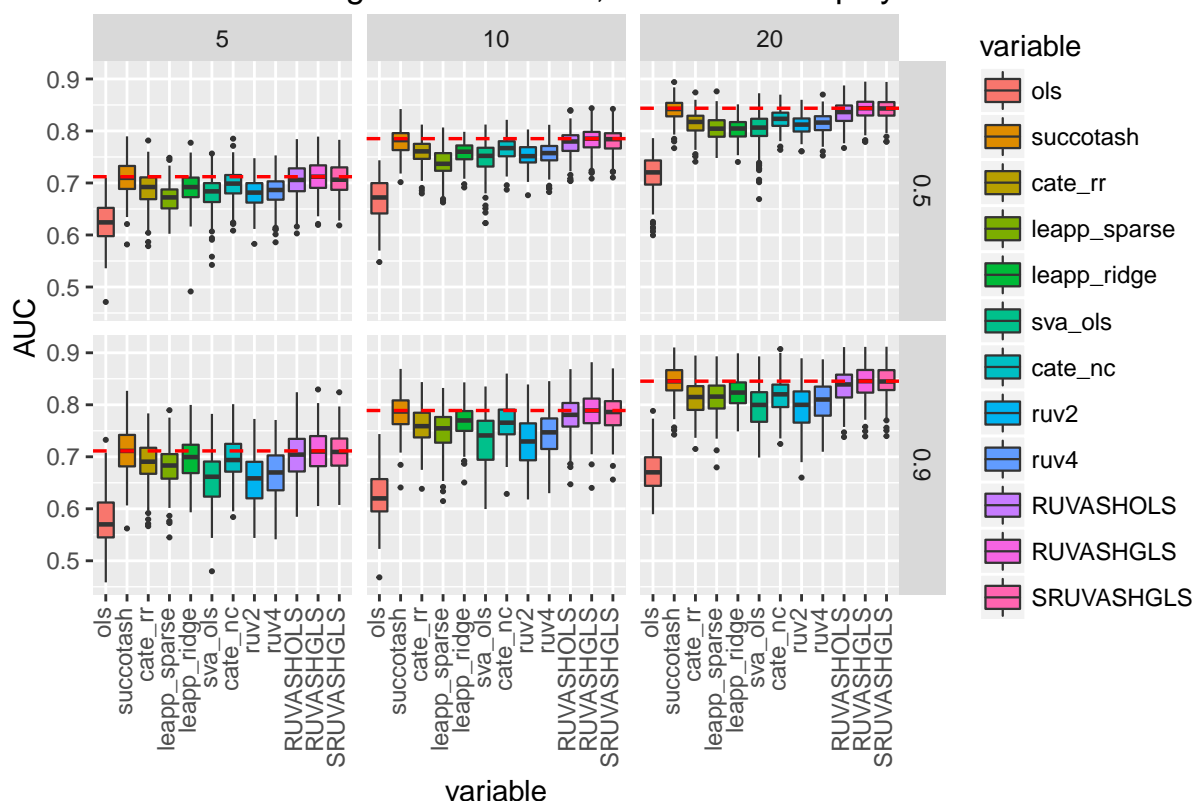
Estimates of π_0 When Using Muscle Tissue, Alternative = bimodal



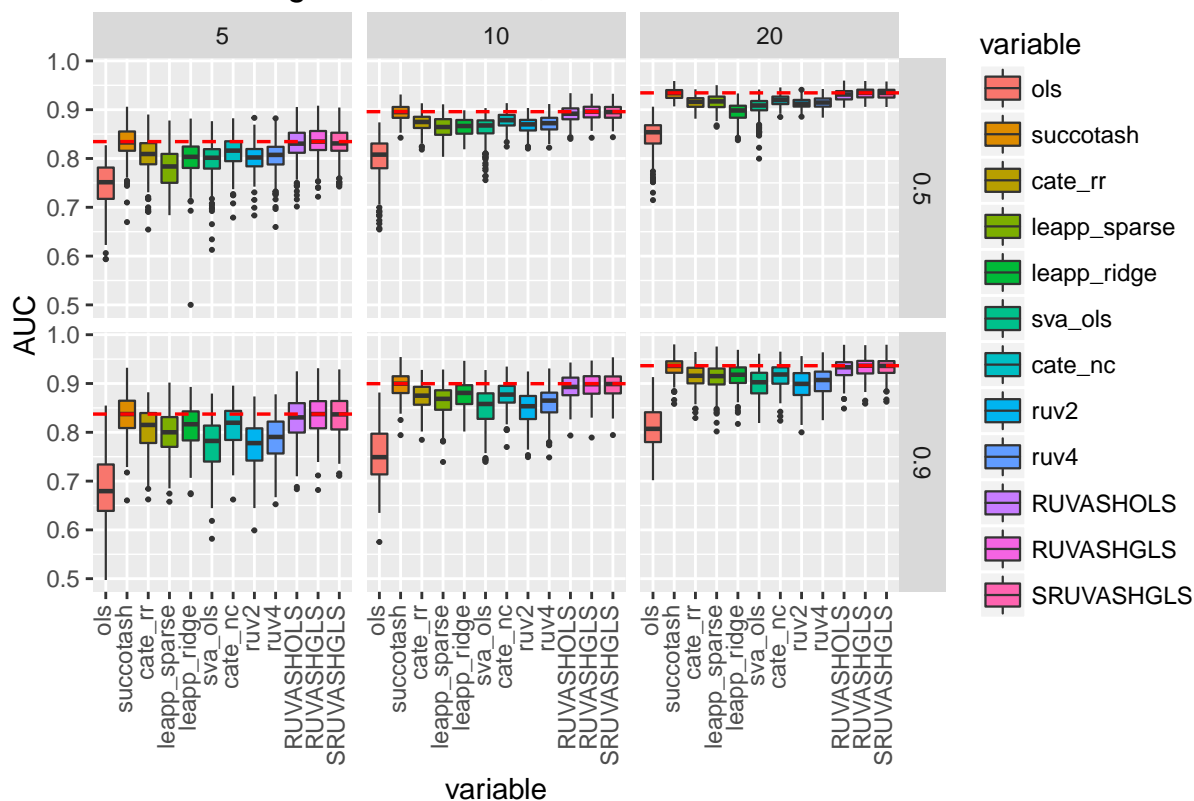
Estimates of π_0 When Using Muscle Tissue and All Null



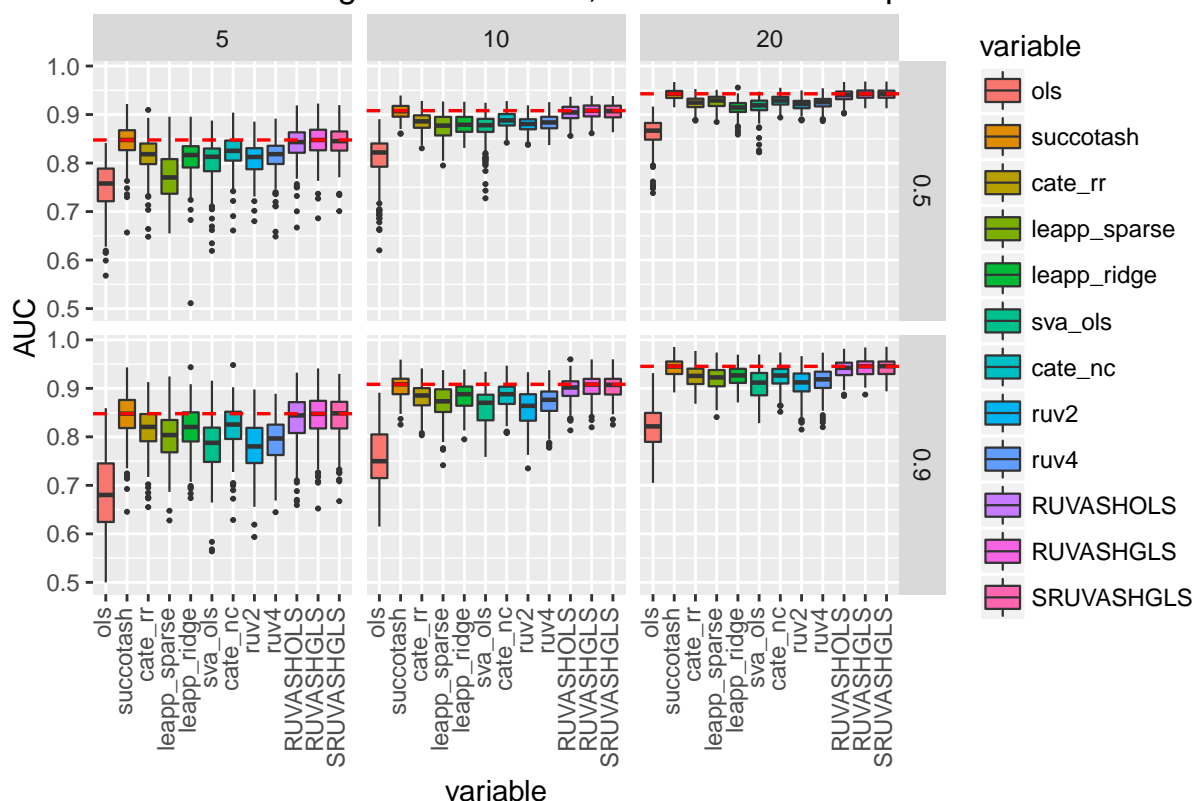
AUC When Using Muscle Tissue, Alternative = spiky



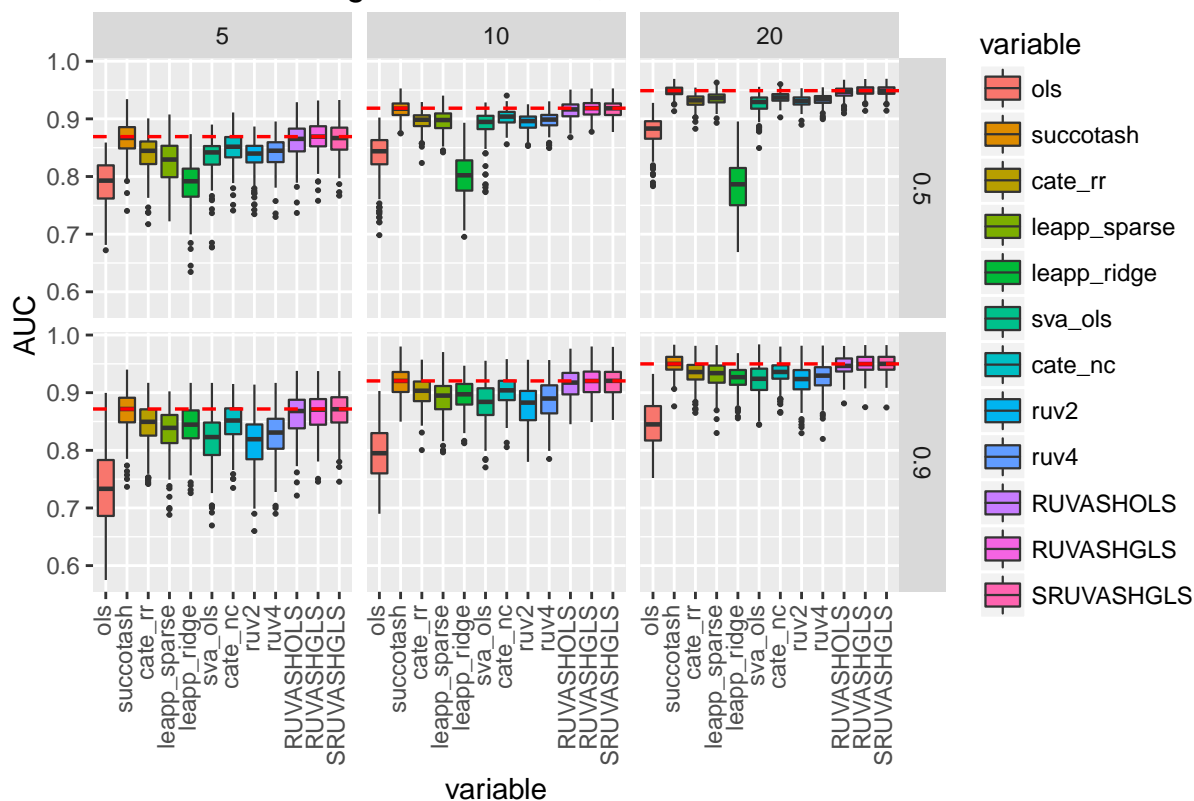
AUC When Using Muscle Tissue, Alternative = near_normal



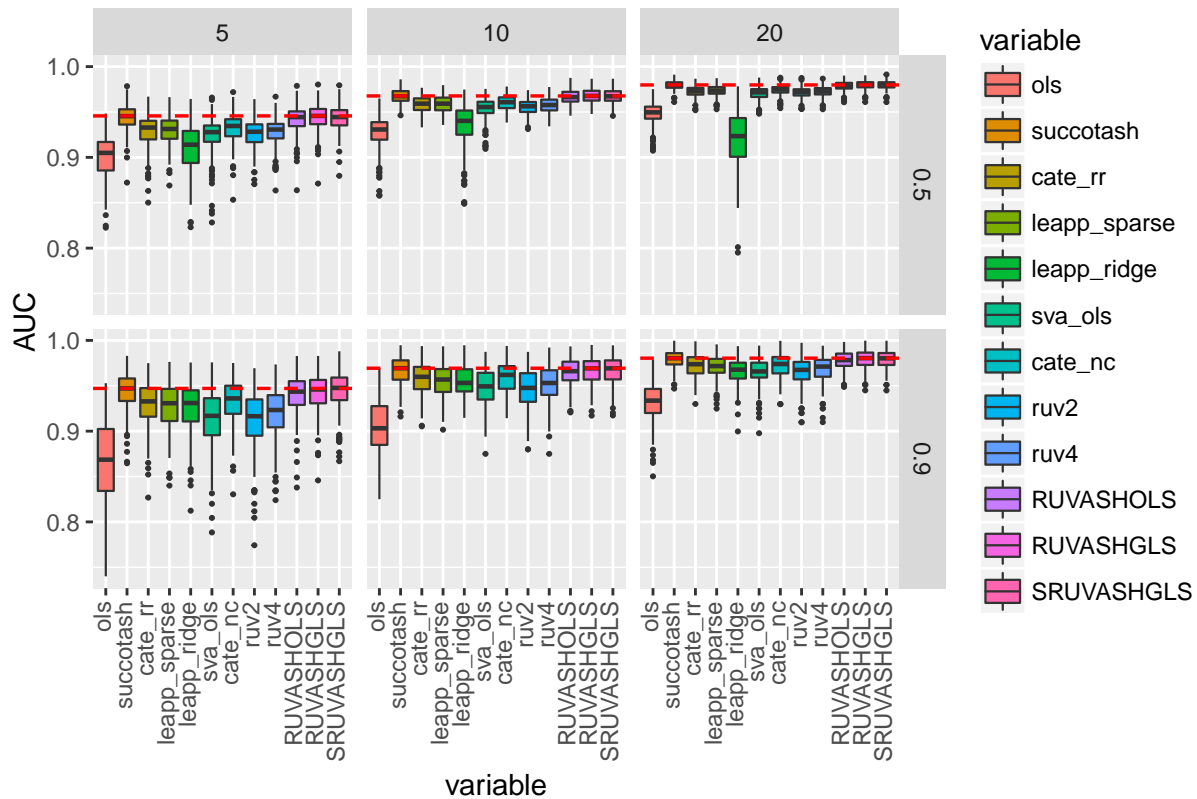
AUC When Using Muscle Tissue, Alternative = flattop



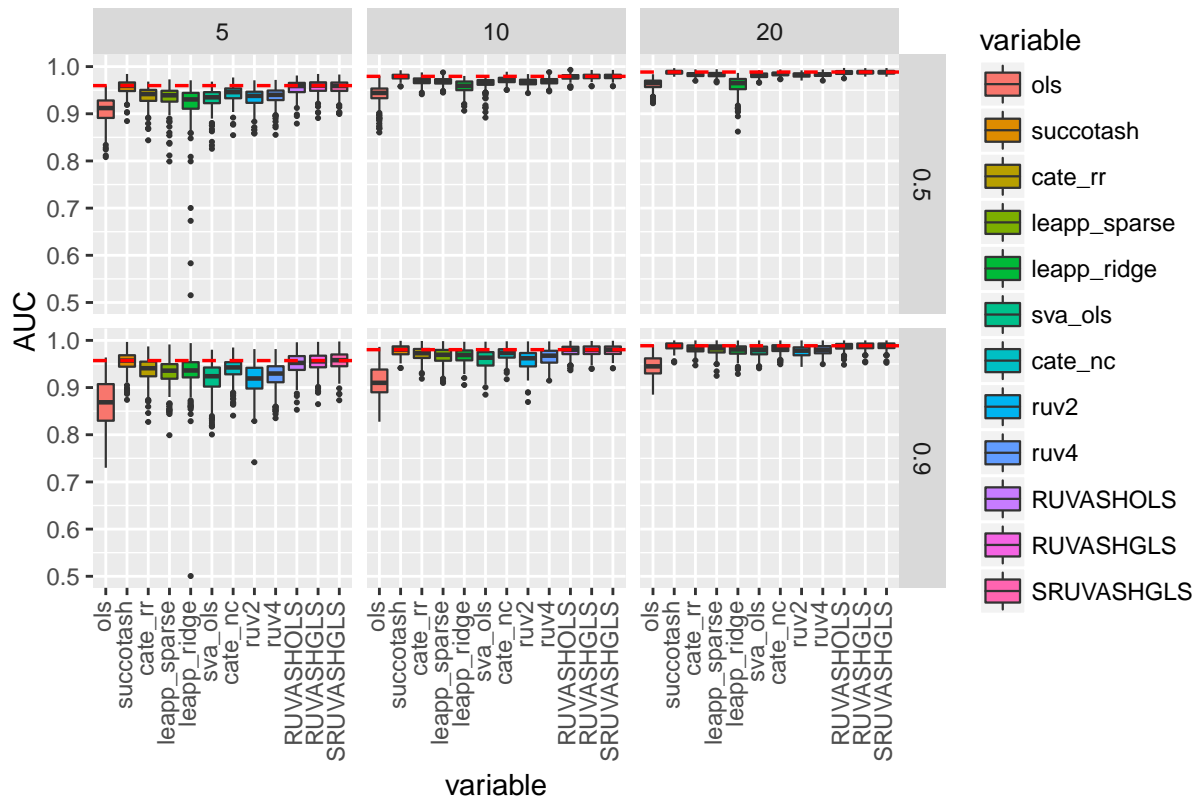
AUC When Using Muscle Tissue, Alternative = skew



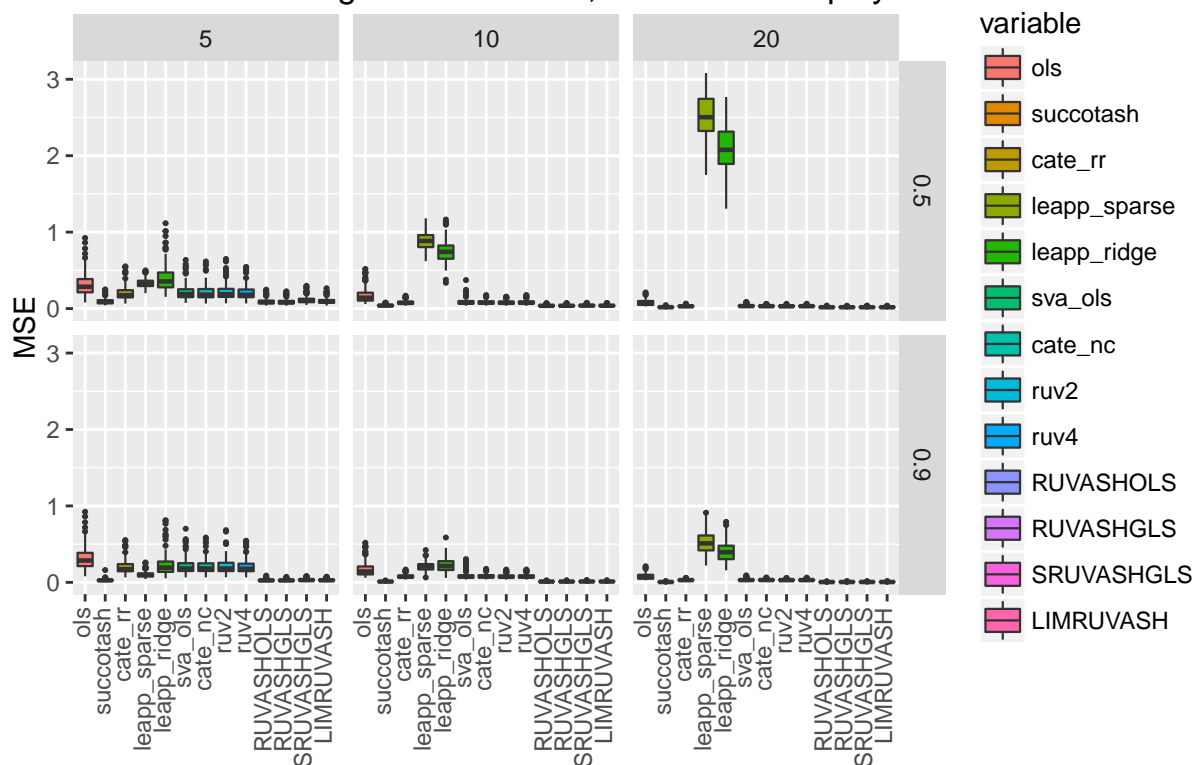
AUC When Using Muscle Tissue, Alternative = big_normal



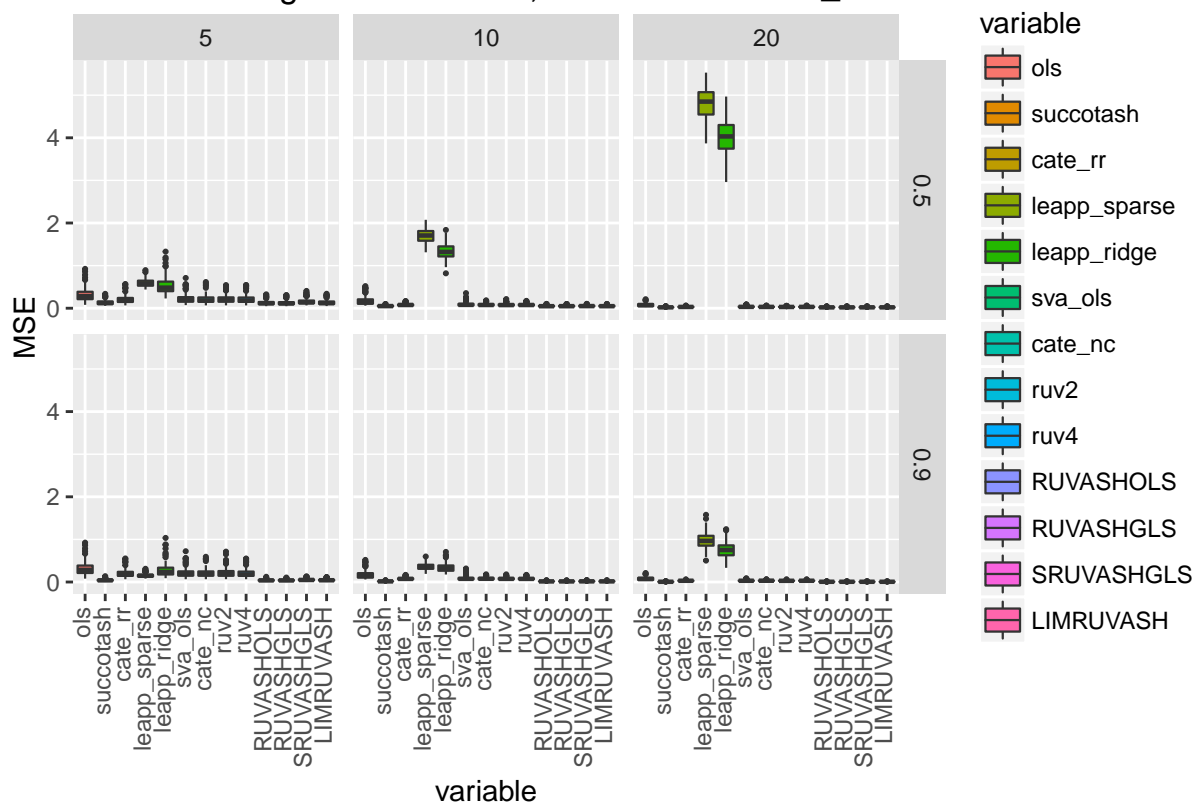
AUC When Using Muscle Tissue, Alternative = bimodal



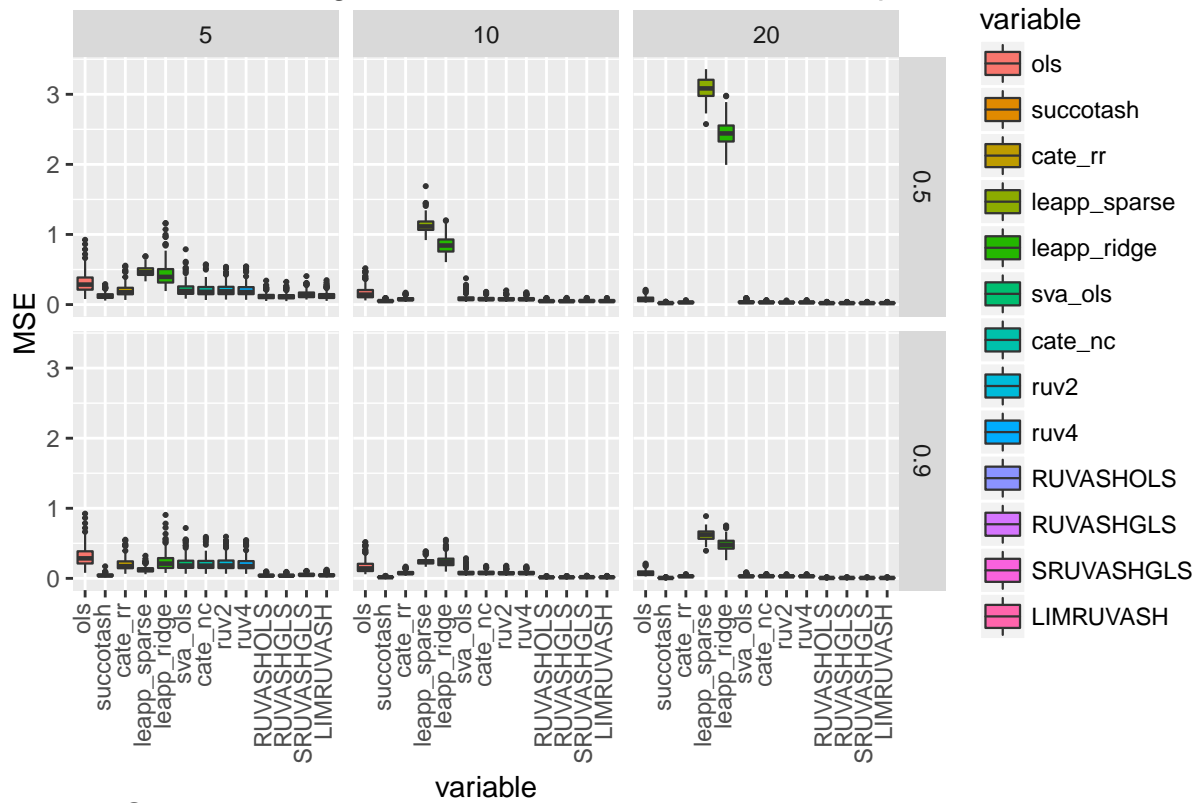
MSE When Using Muscle Tissue, Alternative = spiky



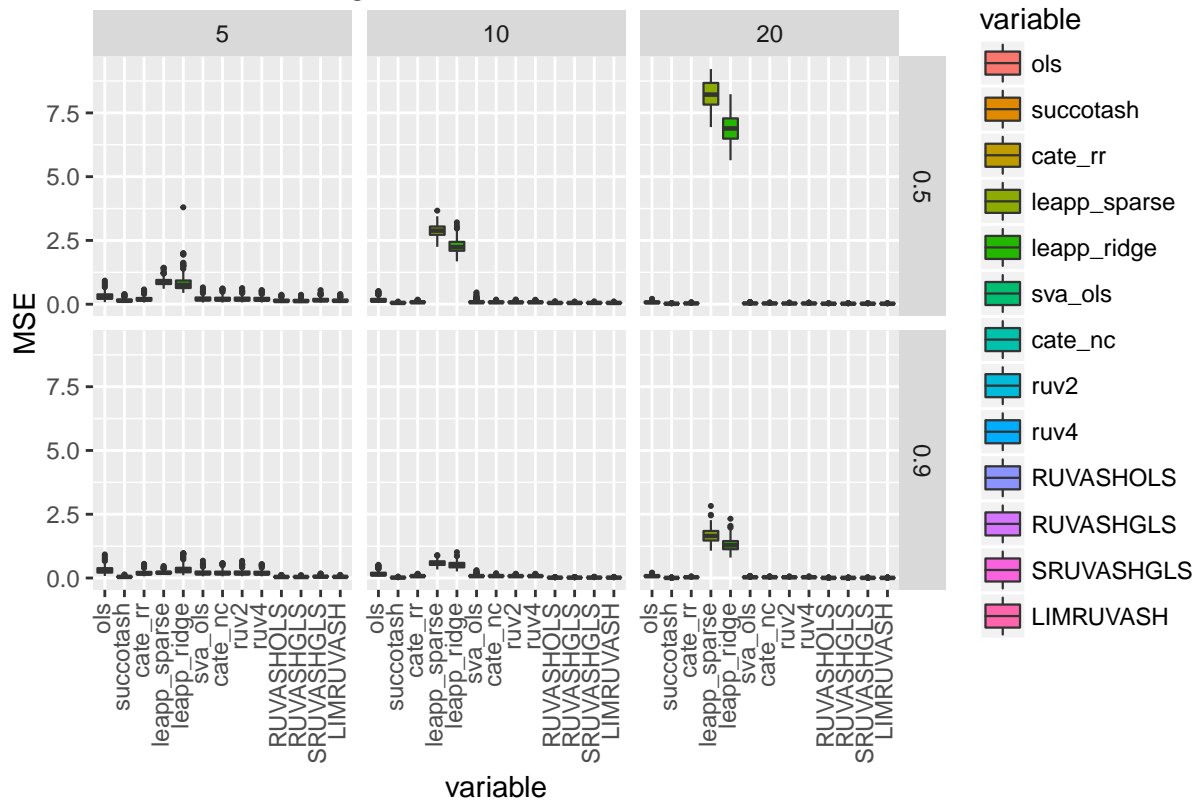
MSE When Using Muscle Tissue, Alternative = near_normal



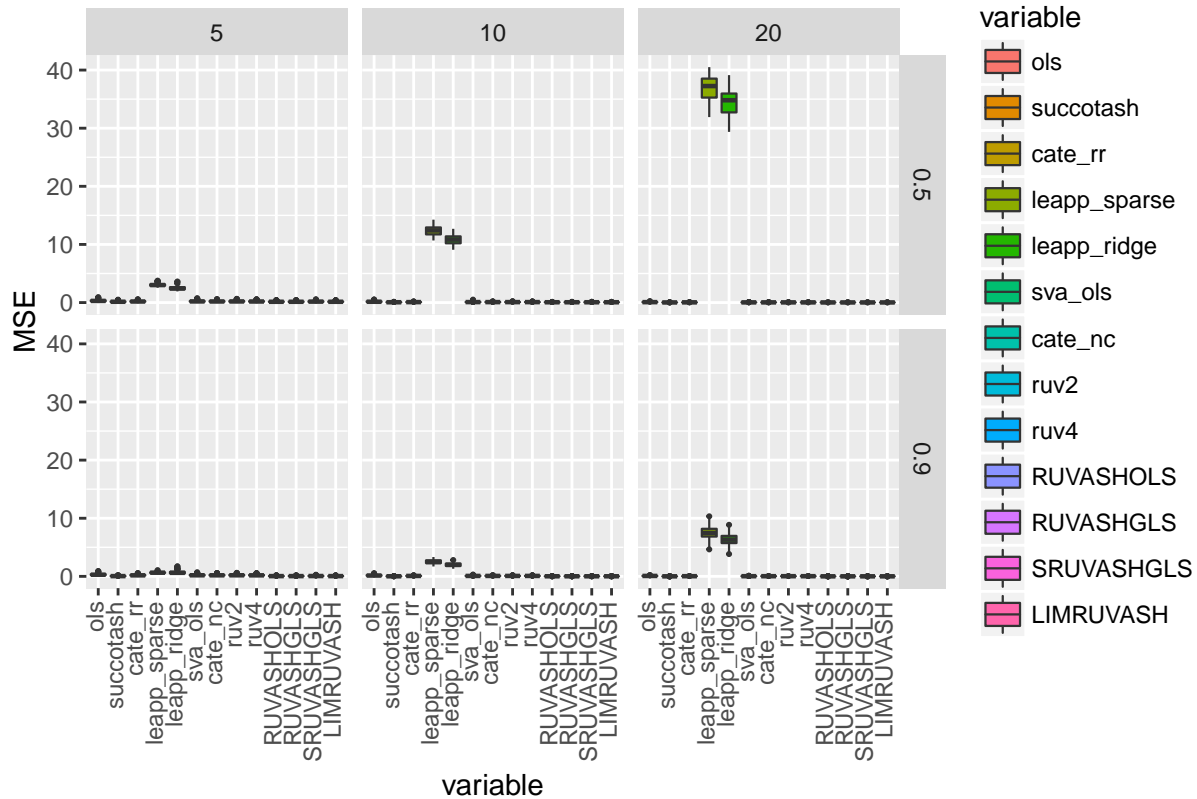
MSE When Using Muscle Tissue, Alternative = flattop



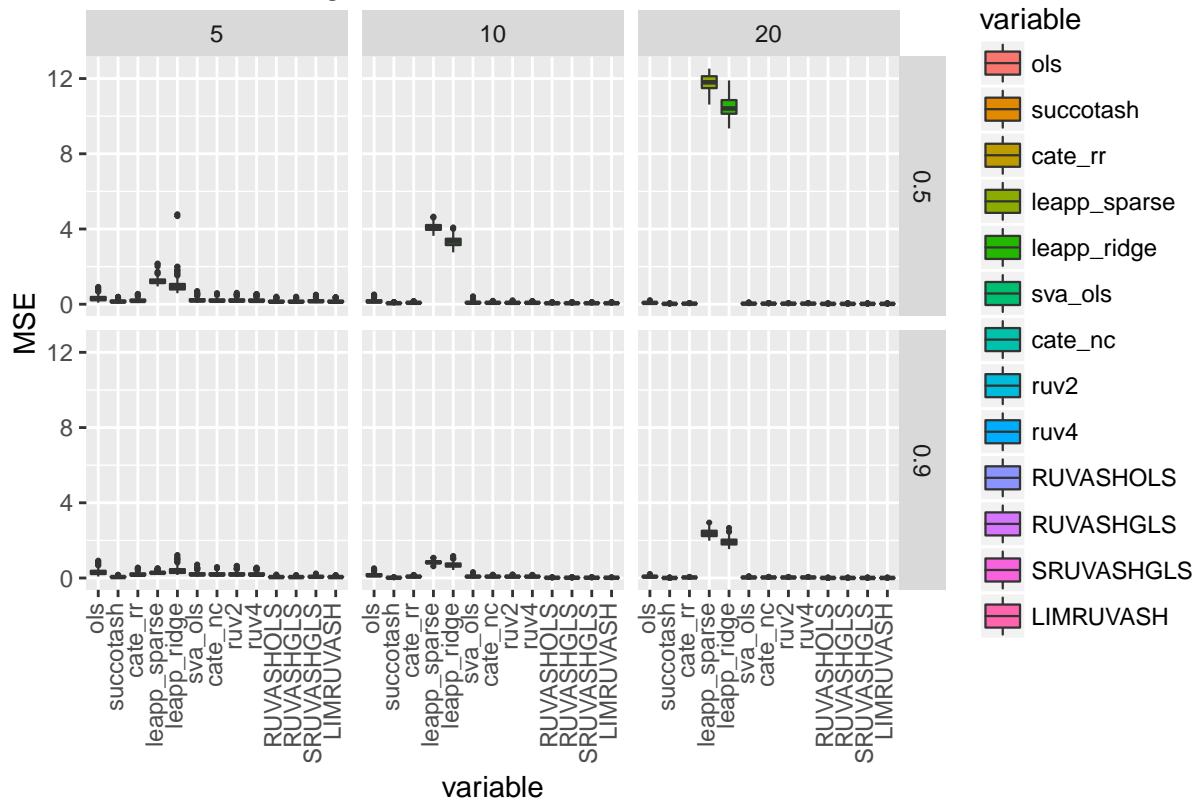
MSE When Using Muscle Tissue, Alternative = skew



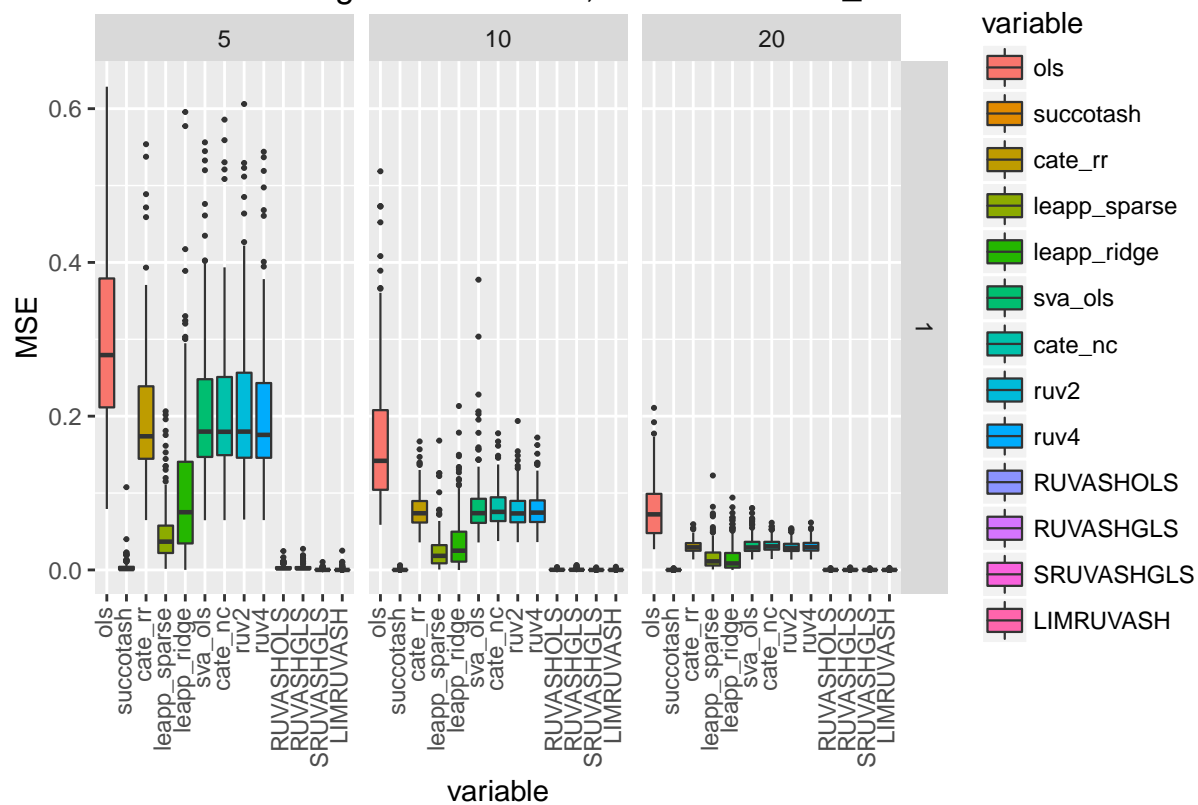
MSE When Using Muscle Tissue, Alternative = big_normal



MSE When Using Muscle Tissue, Alternative = bimodal

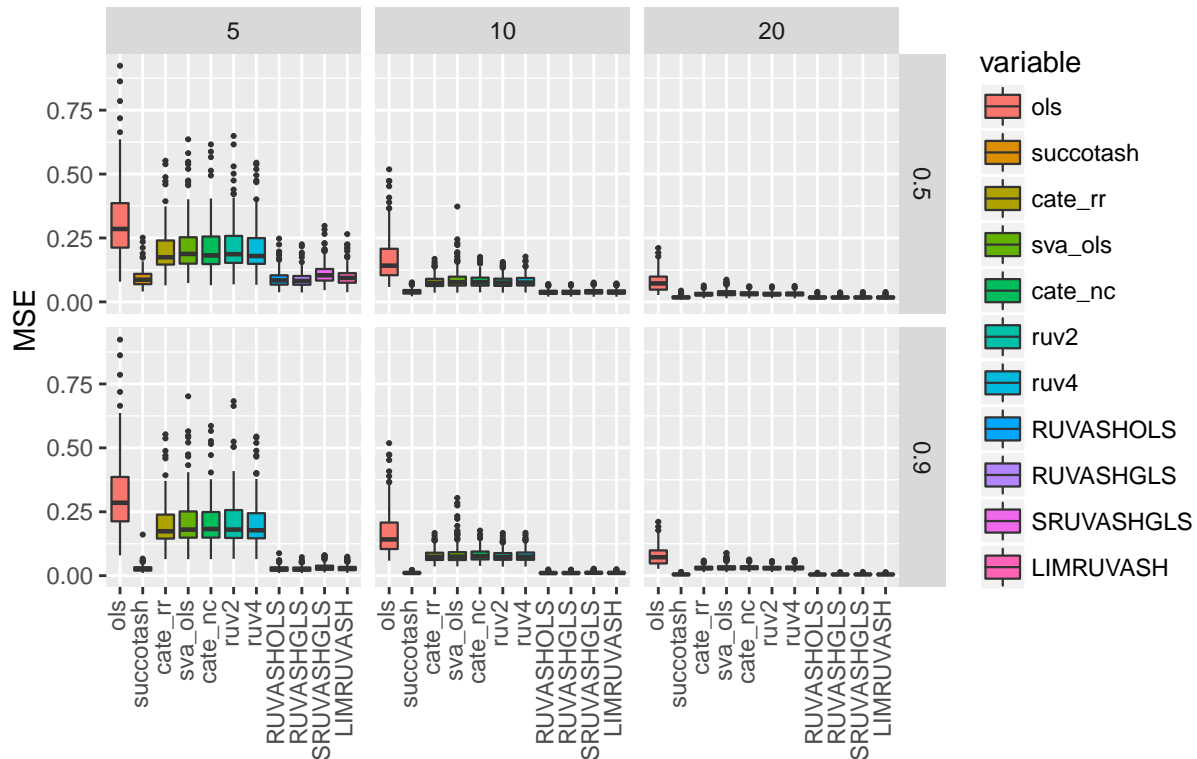


MSE When Using Muscle Tissue, Alternative = all_null

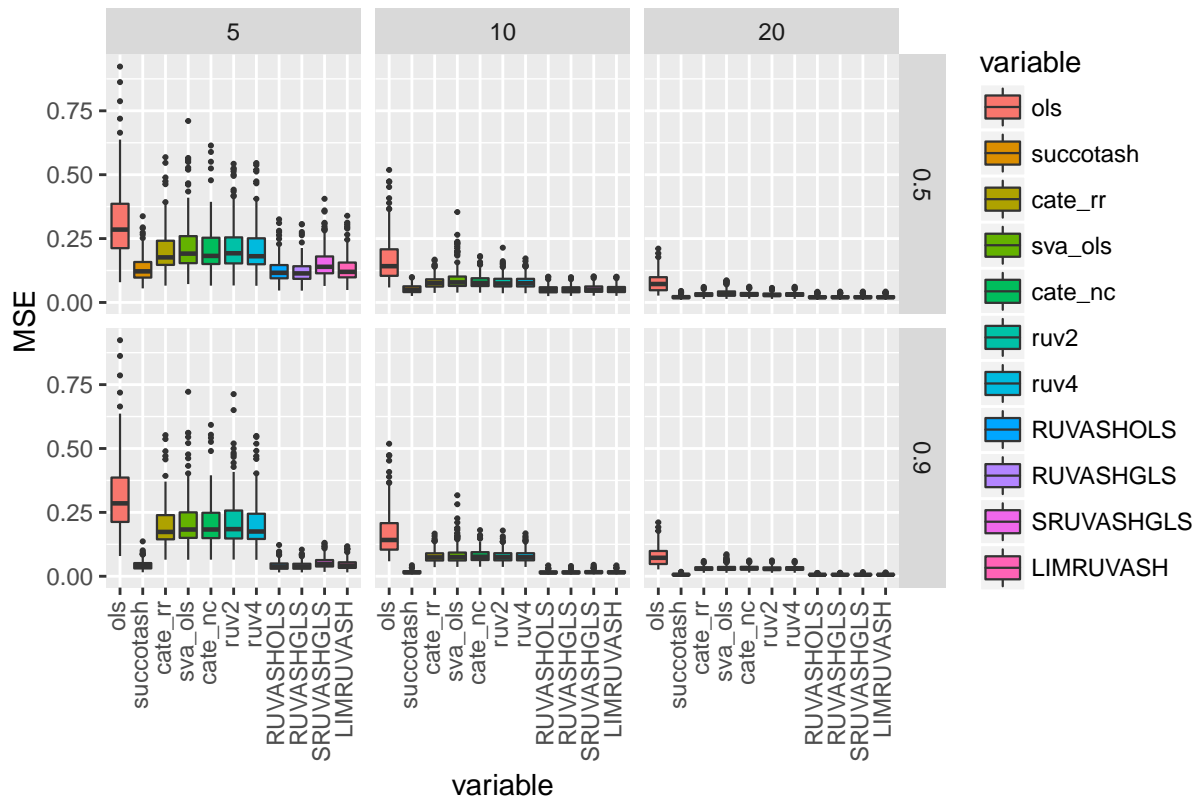


MSE without LEAPP

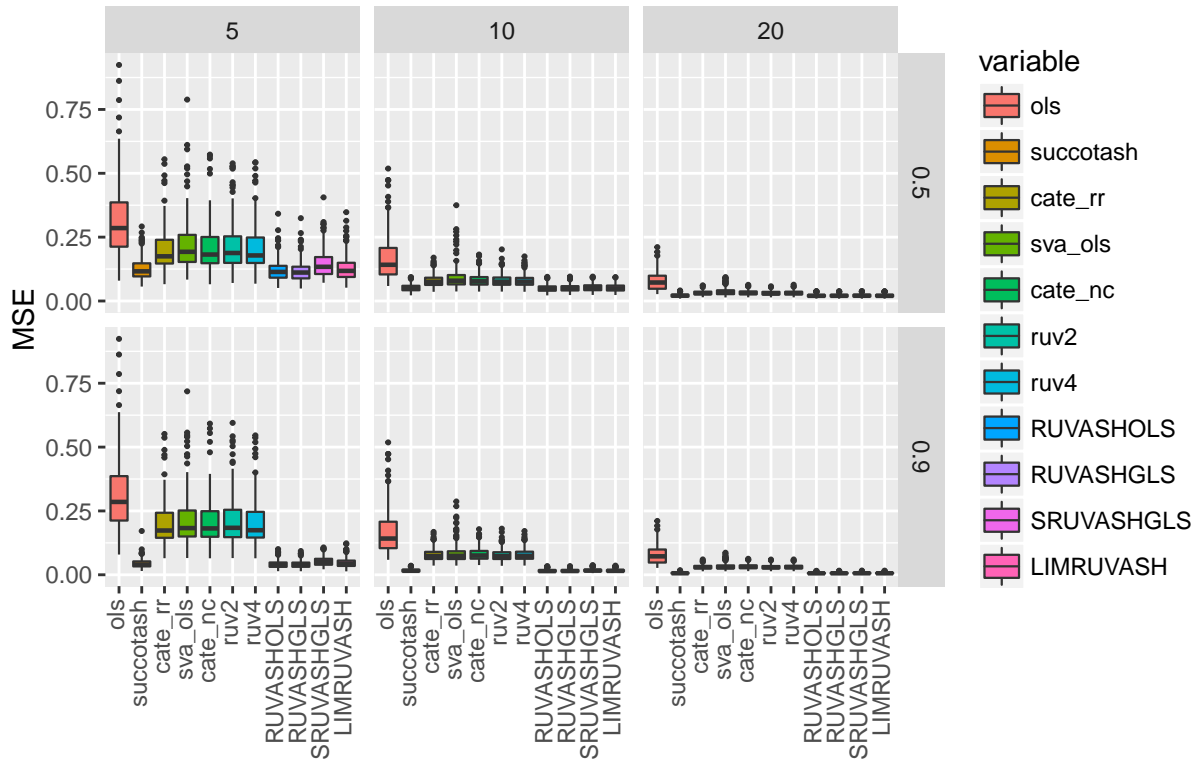
MSE When Using Muscle Tissue, Alternative = spiky



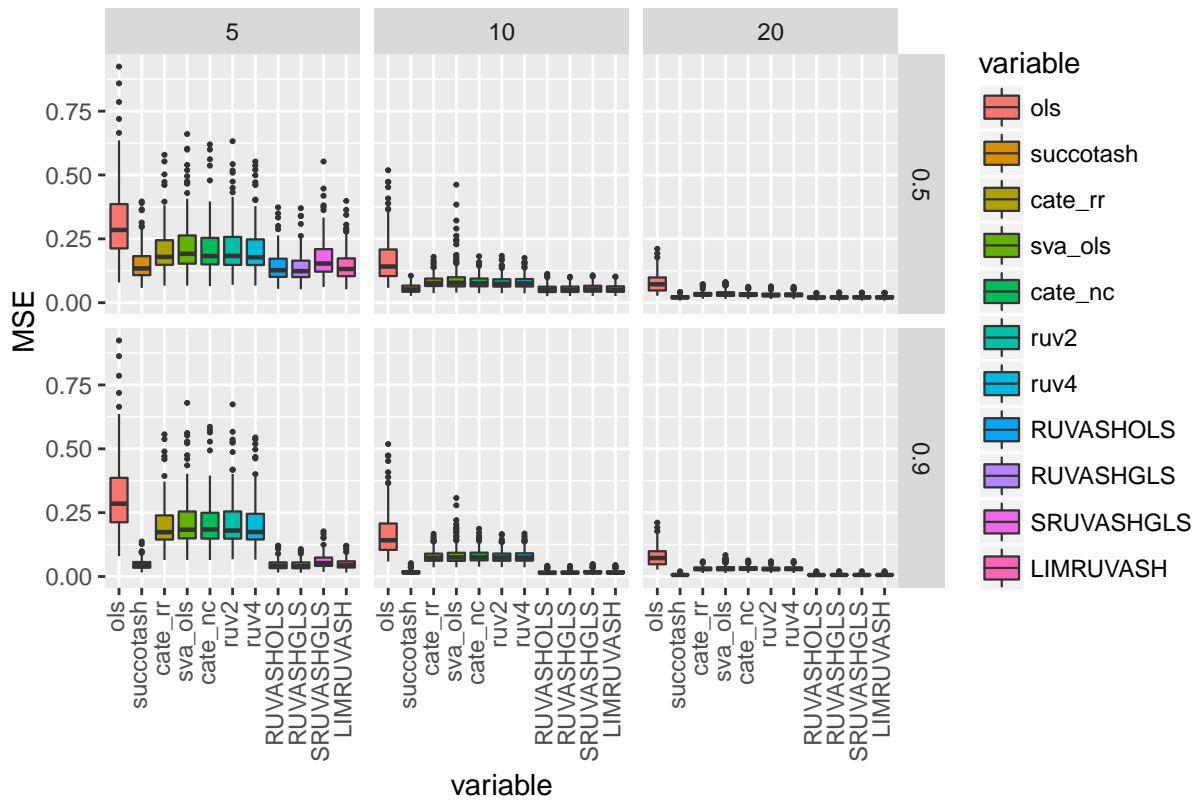
MSE When Using Muscle Tissue, Alternative = near_normal



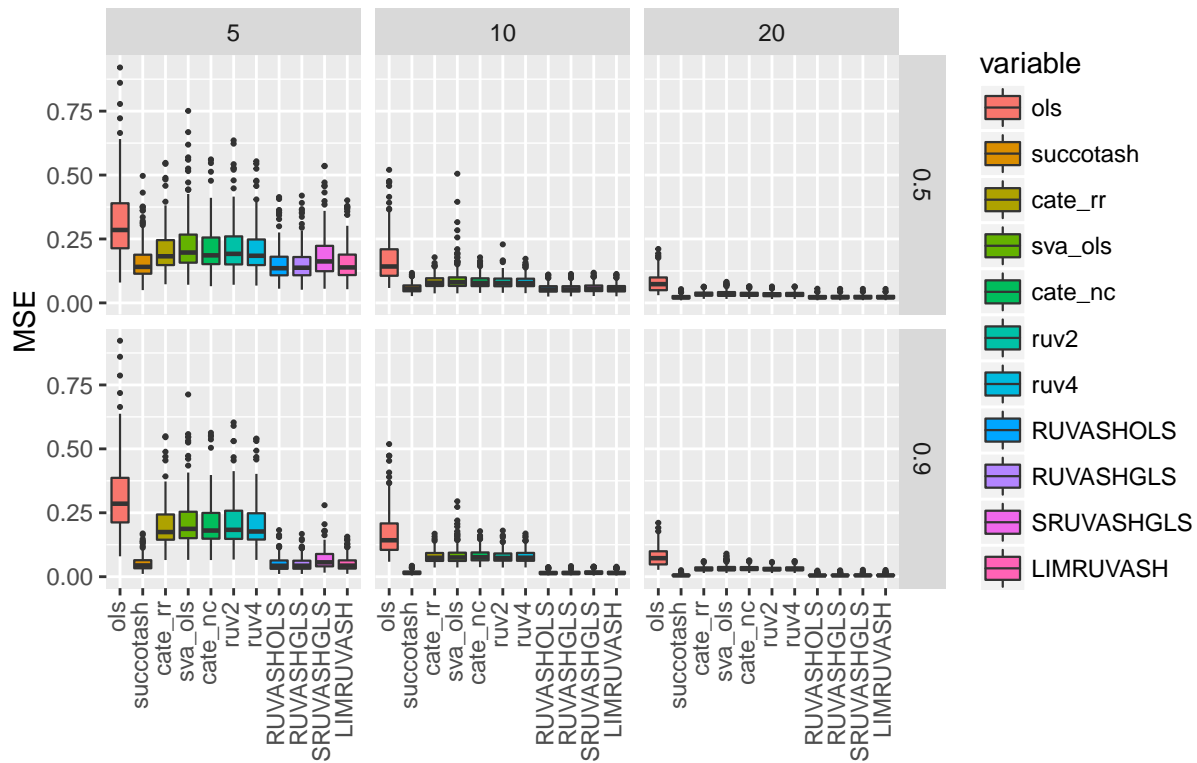
MSE When Using Muscle Tissue, Alternative = flattop



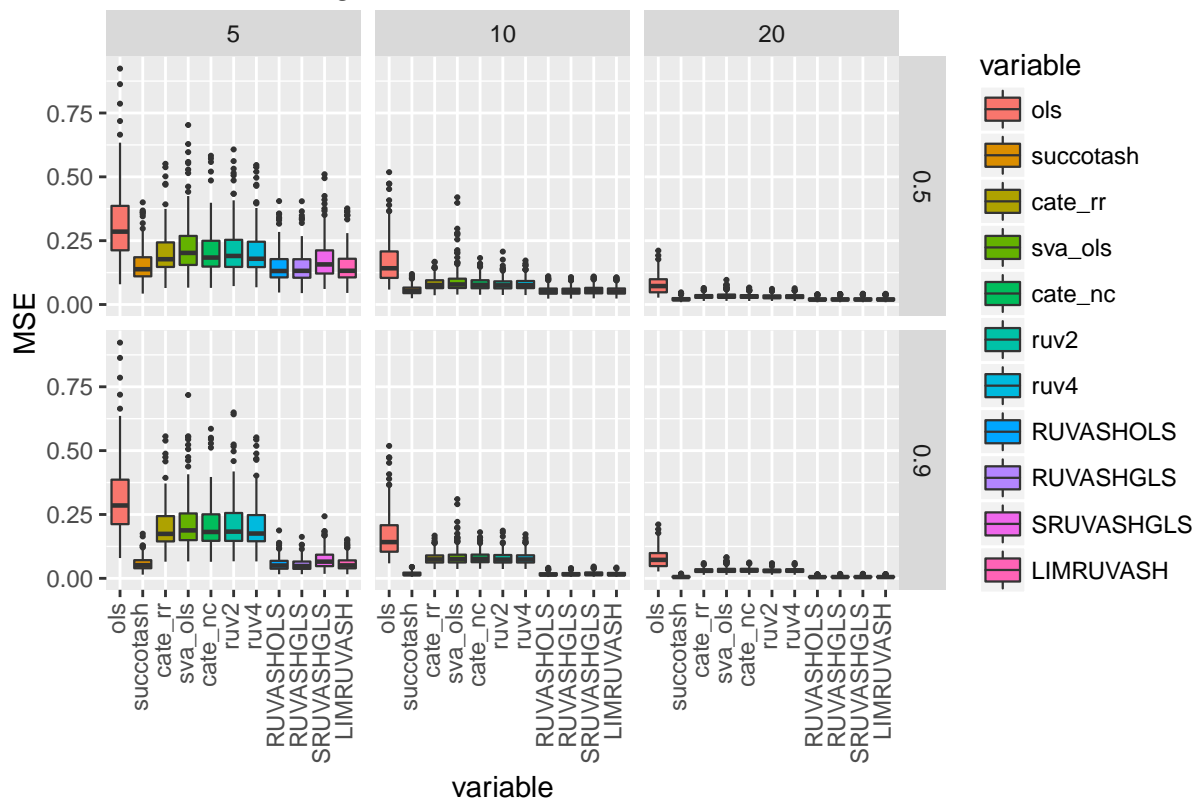
MSE When Using Muscle Tissue, Alternative = skew



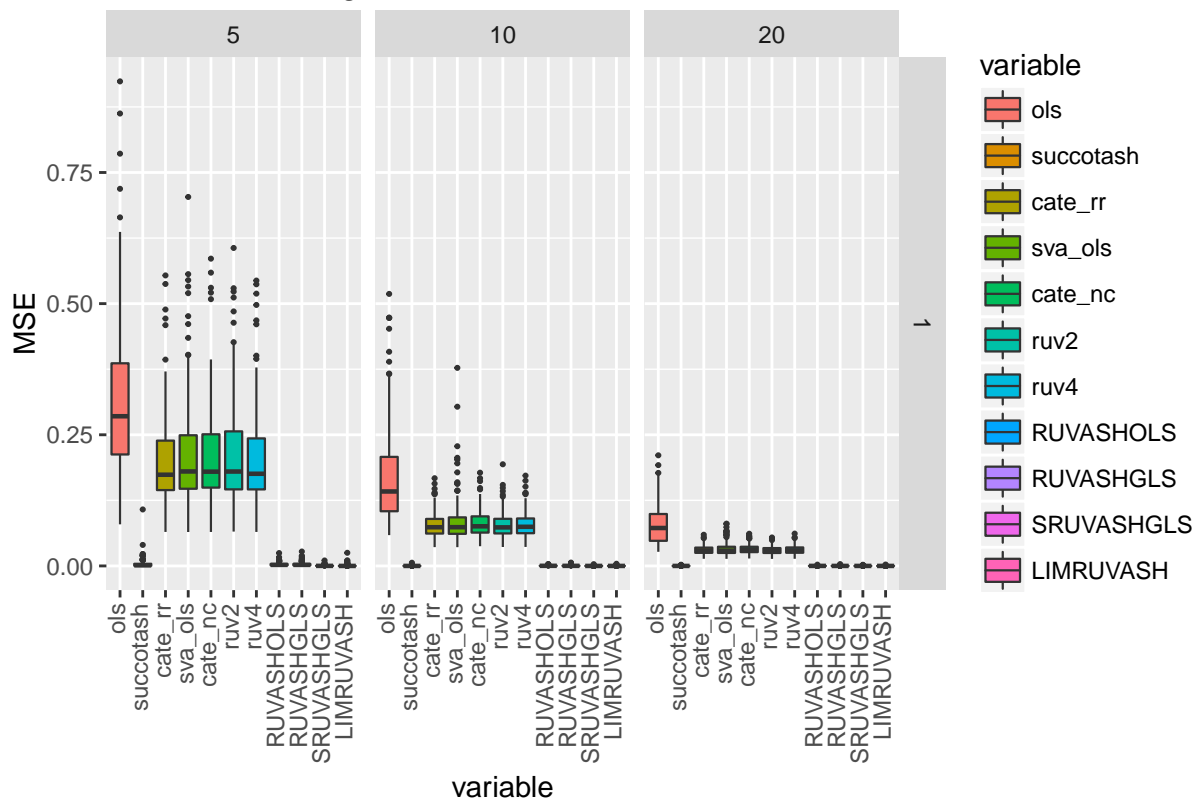
MSE When Using Muscle Tissue, Alternative = big_normal



MSE When Using Muscle Tissue, Alternative = bimodal



MSE When Using Muscle Tissue, Alternative = all_null



`sessionInfo()`

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## Platform: x86_64-pc-linux-gnu (64-bit)
## Running under: Ubuntu 14.04.4 LTS
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## locale:
##  [1] LC_CTYPE=en_US.UTF-8      LC_NUMERIC=C
##  [3] LC_TIME=en_US.UTF-8      LC_COLLATE=en_US.UTF-8
##  [5] LC_MONETARY=en_US.UTF-8  LC_MESSAGES=en_US.UTF-8
##  [7] LC_PAPER=en_US.UTF-8     LC_NAME=C
##  [9] LC_ADDRESS=C             LC_TELEPHONE=C
## [11] LC_MEASUREMENT=en_US.UTF-8 LC_IDENTIFICATION=C
##
## attached base packages:
## [1] stats      graphics  grDevices  utils      datasets  methods    base
##
## other attached packages:
## [1] pROC_1.8      dplyr_0.4.3  reshape2_1.4.1 ggplot2_2.1.0
##
## loaded via a namespace (and not attached):
##  [1] Rcpp_0.12.5      knitr_1.12.28  magrittr_1.5    munsell_0.4.3
##  [5] colorspace_1.2-6 R6_2.1.2       stringr_1.0.0   plyr_1.8.3
##  [9] tools_3.3.0      parallel_3.3.0 grid_3.3.0      gtable_0.2.0
## [13] DBI_0.4          htmltools_0.3.5 yaml_2.1.13     lazyeval_0.1.10
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Buja, Andreas, and Nermin Eyuboglu. 1992. “Remarks on Parallel Analysis.” *Multivariate Behavioral Research* 27 (4). Taylor & Francis: 509–40.

Stephens, Matthew. 2016. “False Discovery Rates: A New Deal.” *BioRxiv*. Cold Spring Harbor Labs Journals, 038216.