

hw_4

October 24, 2022

```
[ ]: import pandas as pd
import numpy as np
import matplotlib.pyplot as plt
import statsmodels.api as sm
from sklearn import linear_model
import scipy.stats as stats
import seaborn as sns
import warnings
import matplotlib as mpl

warnings.filterwarnings("ignore")
pd.set_option("display.precision", 4)
```

```
[ ]: plt.style.use("seaborn")
mpl.rcParams['font.family'] = 'serif'
%matplotlib inline
```

```
[ ]: file_name = "C:/Users/dcste/OneDrive/Portfolio_Theory/Homework_Jupyter/
↳portfolio_theory/dfa_analysis_data.xlsx"
descriptions = pd.read_excel(file_name, sheet_name = "descriptions")
descriptions = descriptions.rename(columns = {"Unnamed: 0": "Ticker"})
descriptions.head()
```

```
[ ]: 
```

	Ticker	Name	Unit	Construction \
0	Mkt-RF	Market	Excess Return	Market-cap-weighted
1	SMB	Size	Excess Return	Small Minus Big
2	HML	Value	Excess Return	High Minus Low
3	RF	Risk-free rate	Total Return	Tbills

	Description
0	US Equities
1	Long small stocks and short big stocks
2	Long value (high book-to-market) stocks and sh...
3	NaN

0.1 Homework 4

Dimensional Fund Advisors

DFA believes certain stocks have higher excess returns. In addition to the overall market equity premium, DFA believes that there is a premium attached to size and the value factor. Calculate the univariate statistics on the Market Excess Return, **SMB** (Long small stocks and short big stocks), and **HML** (long value (high book-to-market) stocks and short growth stocks (low book-to-market ratios).

Report the mean, volatility, Sharpe, and VaR (.05)

```
[ ]: df = pd.read_excel(file_name, sheet_name="factors", parse_dates=True)
df = df.set_index("Date")
factors = df.drop("RF", axis = 1)
```

```
[ ]: def summary_stats(df, annual_frac):
    ss_df = (df.mean()*annual_frac).to_frame("Mean")
    ss_df["Volatility"] = df.std()*np.sqrt(annual_frac)
    ss_df["Sharpe"] = ss_df["Mean"]/ss_df["Volatility"]
    ss_df["VaR"] = df.quantile(.05)
    return ss_df.T
```

```
[ ]: print("Period 1926-2022")
summary_stats(factors,12)
```

Period 1926-2022

```
[ ]:
      Mkt-RF      SMB      HML
Mean      0.0808  0.0233  0.0423
Volatility 0.1852  0.1099  0.1234
Sharpe     0.4361  0.2119  0.3430
VaR        -0.0796 -0.0420 -0.0419
```

```
[ ]: print("From 1926 to 1980")
summary_stats(factors[:"1980"],12)
```

From 1926 to 1980

```
[ ]:
      Mkt-RF      SMB      HML
Mean      0.0811  0.0340  0.0495
Volatility 0.2051  0.1146  0.1338
Sharpe     0.3957  0.2966  0.3697
VaR        -0.0840 -0.0434 -0.0429
```

```
[ ]: print("1981-2001")
summary_stats(factors["1981":"2001"],12)
```

1981-2001

```
[ ]:
      Mkt-RF      SMB      HML
Mean      0.0773 -0.0009  0.0637
Volatility 0.1574  0.1184  0.1113
Sharpe     0.4908 -0.0074  0.5727
```

VaR -0.0645 -0.0466 -0.0418

```
[ ]: print("2002-2022")
      summary_stats(factors["2002":],12)
```

2002-2022

```
[ ]:          Mkt-RF      SMB      HML
Mean          0.0833   0.0196   0.0017
Volatility    0.1540   0.0858   0.1045
Sharpe        0.5409   0.2288   0.0161
VaR           -0.0788 -0.0378 -0.0410
```

Answer the following questions:

1. Does each factor have a premium (positive excess return)?
2. Does the premium to the size factor get smaller after 1980?
3. Does the premium to the value factor get smaller during the 1990's?
4. How have the factors performed since the time of the case (2002-present)?

Answers:

1. No, in the period of 1981 - 2001, SMB (**Size Factors**) has negative excess return during this period.
2. No, the size factor premium actually increases after 1980.
3. The value factor gets higher during the 1990s, but then declines during the periods of 2002-to present day.
4. The factors have earned a positive premium to the risk-free rate since the time of the case, but they have not performed that well. The market factor has earned the best premium to the market.

0.2 Question 3

Report the correlation matrix across all three factors and in each subsample period.

```
[ ]: def corr_mat(df_):
      corr_mat = df_.corr()
      return corr_mat
```

```
[ ]: corr_mat(factors)
```

```
[ ]:          Mkt-RF      SMB      HML
Mkt-RF    1.0000   0.3163   0.2297
SMB        0.3163   1.0000   0.1147
HML        0.2297   0.1147   1.0000
```

```
[ ]: corr_mat(factors[:"1980"])
```

```
[ ]:      Mkt-RF      SMB      HML
Mkt-RF  1.0000  0.3663  0.4378
SMB      0.3663  1.0000  0.3091
HML      0.4378  0.3091  1.0000
```

```
[ ]: corr_mat(factors["1981":"2001"])
```

```
[ ]:      Mkt-RF      SMB      HML
Mkt-RF  1.0000  0.1643 -0.5268
SMB      0.1643  1.0000 -0.4548
HML     -0.5268 -0.4548  1.0000
```

```
[ ]: corr_mat(factors["2002":])
```

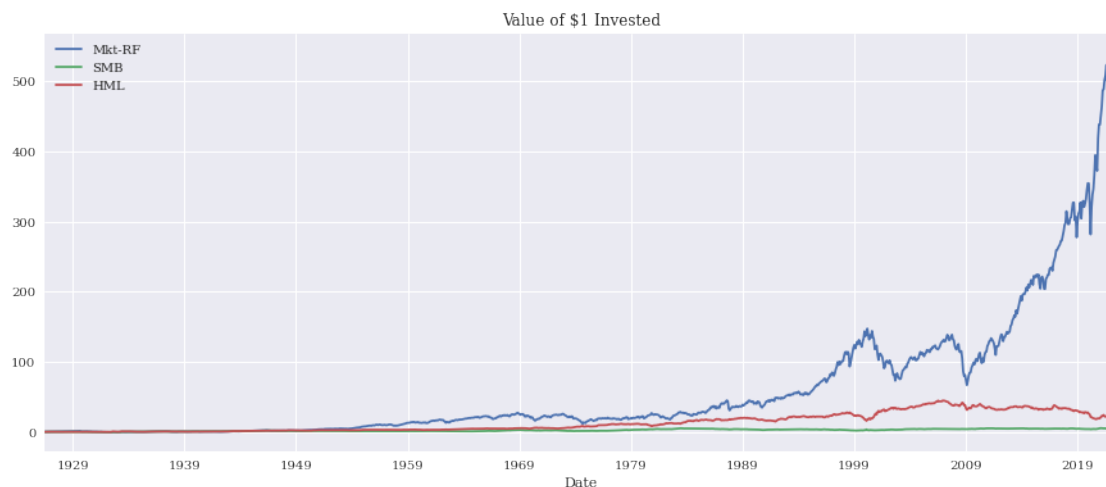
```
[ ]:      Mkt-RF      SMB      HML
Mkt-RF  1.0000  0.3162  0.1317
SMB      0.3162  1.0000  0.0929
HML      0.1317  0.0929  1.0000
```

Yes, I would say the correlations are quite small from sample to sample and even in the full sample suggesting the factors returns are not driven by the the overall market.

0.3 4 Plot Cumulative Returns of the Three Factors

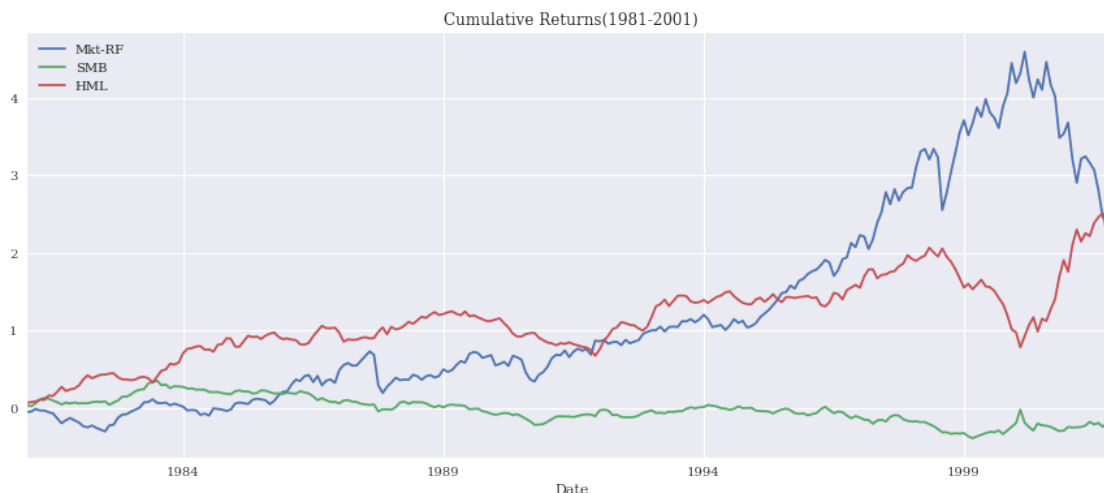
```
[ ]: ((1+factors).cumprod()-1).plot(figsize=(15,6), title = "Value of $1 Invested")
```

```
[ ]: <AxesSubplot:title={'center':'Value of $1 Invested'}, xlabel='Date'>
```



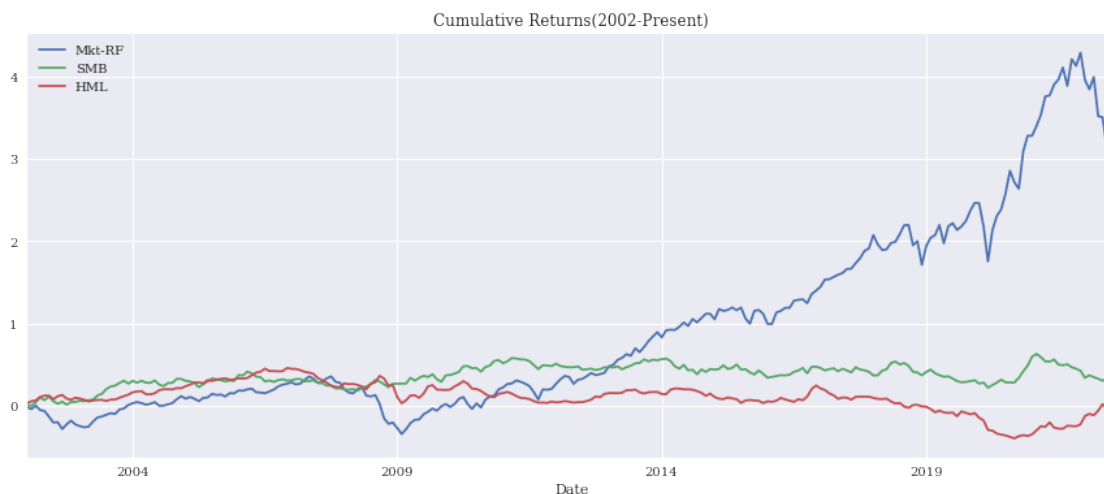
```
[ ]: ((1+factors["1981":"2001"]).cumprod()-1).plot(figsize = (15,6), title = "
↳Cumulative Returns(1981-2001)")
```

```
[ ]: <AxesSubplot:title={'center':'Cumulative Returns(1981-2001)'}, xlabel='Date'>
```



```
[ ]: ((1+factors["2002:"]).cumprod()-1).plot(figsize = (15,6), title = "Cumulative Returns(2002-Present)")
```

```
[ ]: <AxesSubplot:title={'center':'Cumulative Returns(2002-Present)'}, xlabel='Date'>
```



Does it appear that all three factors were valuable in 1981-2001? And post-2001? Would you advise DFA to continue emphasizing all three factors?

No, not all the factors were valuable in 1981-2001. The **size factor** earns a negative premium. However, even though these factors do not earn a better returns to the market, nor have better sharpe ratios, this does not mean DFA should disregard these factors. Since their correlations are small, these factors could have have great diversification benefits for investors.

0.4 CAPM

DFA believes that premia in stocks and stock portfolios is related to the three factors. Let's test 25 equity portfolios that span a wide range of size and value measures.

- For this entire problem focus on the 1981-Present subsample.
1. Calculate summary statistics for each portfolio,
 1. Use the risk-free rate column in the factors tab to convert these total returns to excess returns.
 2. Calculate the annualized univariate statistics from 1.1
 3. Can the difference in mean excess returns of the portfolios be explained by the difference in their volatilities.

```
[ ]: portfolio = pd.read_excel(file_name,sheet_name="portfolios (total returns)")
portfolio = portfolio.set_index("Date")
portfolios_ex = portfolio.subtract(df["RF"], axis = 'rows')
```

```
[ ]: subsample = portfolios_ex["1981":]
port_summary = summary_stats(subsample,12).T
port_summary["portfolio_num"] = range(1,26)
port_summary
```

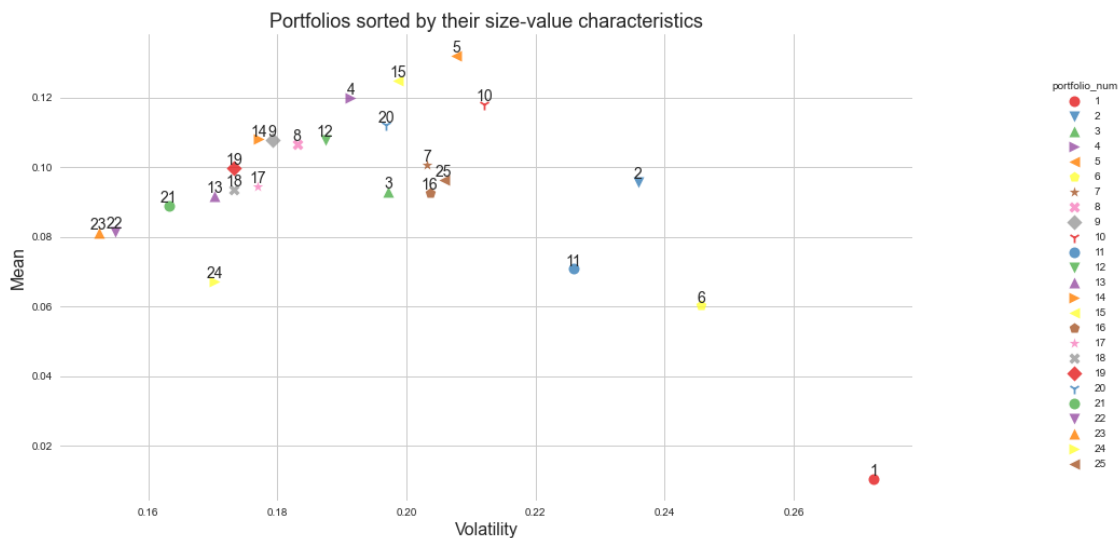
```
[ ]:
```

	Mean	Volatility	Sharpe	VaR	portfolio_num
SMALL LoBM	0.0105	0.2724	0.0384	-0.1165	1
ME1 BM2	0.0957	0.2359	0.4058	-0.0976	2
ME1 BM3	0.0929	0.1972	0.4711	-0.0864	3
ME1 BM4	0.1200	0.1914	0.6270	-0.0752	4
SMALL HiBM	0.1321	0.2077	0.6361	-0.0883	5
ME2 BM1	0.0603	0.2457	0.2453	-0.1046	6
ME2 BM2	0.1006	0.2032	0.4953	-0.0840	7
ME2 BM3	0.1066	0.1831	0.5824	-0.0787	8
ME2 BM4	0.1077	0.1792	0.6012	-0.0764	9
ME2 BM5	0.1181	0.2120	0.5571	-0.0942	10
ME3 BM1	0.0708	0.2259	0.3135	-0.1023	11
ME3 BM2	0.1078	0.1874	0.5753	-0.0786	12
ME3 BM3	0.0918	0.1702	0.5392	-0.0738	13
ME3 BM4	0.1080	0.1771	0.6097	-0.0720	14
ME3 BM5	0.1249	0.1987	0.6282	-0.0824	15
ME4 BM1	0.0927	0.2036	0.4551	-0.0839	16
ME4 BM2	0.0945	0.1770	0.5341	-0.0734	17
ME4 BM3	0.0936	0.1733	0.5402	-0.0756	18
ME4 BM4	0.0998	0.1733	0.5756	-0.0693	19
ME4 BM5	0.1120	0.1968	0.5693	-0.0833	20
BIG LoBM	0.0889	0.1632	0.5447	-0.0753	21
ME5 BM2	0.0816	0.1549	0.5271	-0.0656	22
ME5 BM3	0.0811	0.1523	0.5324	-0.0704	23
ME5 BM4	0.0671	0.1703	0.3943	-0.0756	24
BIG HiBM	0.0963	0.2058	0.4681	-0.0985	25

```
[ ]: markers=['o', 'v', '^', '>', '<', 'p', '*', 'X', 'D', '1', 'o', \
            'v', '^', '>', '<', 'p', '*', 'X', 'D', '1', 'o', 'v', '^', '>', '<']
sns.set_style("whitegrid")
ax = sns.lmplot('Volatility', # Horizontal axis
               'Mean', # Vertical axis
               hue='portfolio_num', palette="Set1", scatter_kws={"s": 100},
               data=port_summary, # Data source
               fit_reg=False, # Don't fix a regression line
               markers=markers,
               aspect =1) # size and dimension
ax.fig.set_figwidth(16)
ax.fig.set_figheight(7)
plt.title('Portfolios sorted by their size-value characteristics', fontsize=18)
plt.xlabel('Volatility', fontsize=16)
plt.ylabel('Mean', fontsize=16)

def label_point(x, y, val, ax):
    a = pd.concat({'x': x, 'y': y, 'val': val}, axis=1)
    for i, point in a.iterrows():
        ax.text(point['x'], point['y'], str(int(point['val'])), fontsize=14, \
                horizontalalignment='center', verticalalignment='bottom')

label_point(port_summary.Volatility, port_summary.Mean, port_summary.
            portfolio_num, plt.gca())
```



```
[ ]: markers=['o', 'v', '^', '>', '<', 'p', '*', 'X', 'D', '1', 'o', \
            'v', '^', '>', '<', 'p', '*', 'X', 'D', '1', 'o', 'v', '^', '>', '<']
sns.set_style("whitegrid")
ax = sns.lmplot('VaR', # Horizontal axis
```

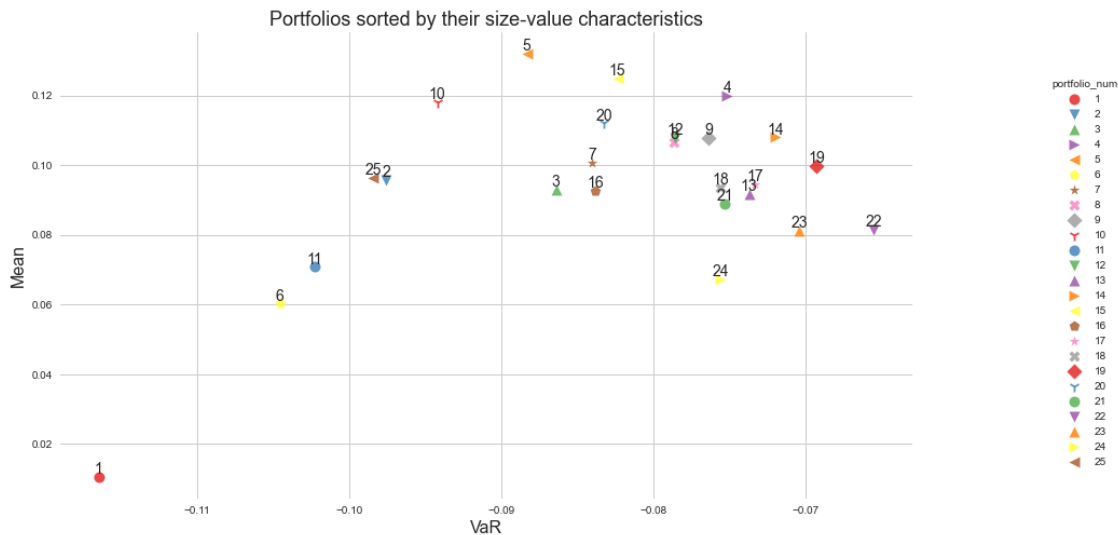
```

'Mean', # Vertical axis
hue='portfolio_num',palette="Set1",scatter_kws={"s": 100},
data=port_summary, # Data source
fit_reg=False, # Don't fix a regression line
markers=markers,
aspect =1) # size and dimension
ax.fig.set_figwidth(16)
ax.fig.set_figheight(7)
plt.title('Portfolios sorted by their size-value characteristics', fontsize=18)
plt.xlabel('VaR', fontsize=16)
plt.ylabel('Mean', fontsize=16)

def label_point(x, y, val, ax):
    a = pd.concat({'x': x, 'y': y, 'val': val}, axis=1)
    for i, point in a.iterrows():
        ax.text(point['x'], point['y'], str(int(point['val'])), fontsize=14,\
                horizontalalignment='center', verticalalignment='bottom')

label_point(port_summary.VaR, port_summary.Mean,port_summary.portfolio_num, plt.
    ↳ gca())

```



From the two graphs above, the differences in mean excess returns cannot be explained by their volatilities or Value-at-Risk measures. This is interesting, because holding everything else constant, higher risk should mean higher excess returns. As we can see, some portfolios with higher volatilities actually have lower mean returns.

The Capital Asset Pricing Model (CAPM) asserts that an asset (or portfolio's) expected excess return is completely a function of its beta to the equity market index (SPY, or in this case, MKT.)

For each of the $n = 25$ test portfolios, run the CAPM time-series regression. Report the estimated

$\beta^{i,m}$, Treynor ratio, α^i , and Information Ratio for each of the $n = 25$ regressions.

```
[ ]: capm_data = portfolios_ex.join(df['Mkt-RF'])["1981":]

[ ]: capm_report = pd.DataFrame(index=portfolios_ex.columns)
rhs = sm.add_constant(capm_data['Mkt-RF'])
bm_residuals = pd.DataFrame(columns=portfolios_ex.columns)
t_p_values = pd.DataFrame()

for portf in portfolios_ex.columns:
    lhs = capm_data[portf]
    res = sm.OLS(lhs, rhs, missing='drop').fit()
    capm_report.loc[portf, 'alpha_hat'] = res.params['const'] * 12
    capm_report.loc[portf, 'beta_hat'] = res.params['Mkt-RF']
    capm_report.loc[portf, 'info_ratio'] = np.sqrt(12) * res.params['const'] / \
    ↪ res.resid.std()
    capm_report.loc[portf, 'treynor_ratio'] = 12 * capm_data[portf].mean() / \
    ↪ res.params['Mkt-RF']
    capm_report.loc[portf, "Mean Return"] = 12*capm_data[portf].mean()
    bm_residuals[portf] = res.resid
    t_p_values.loc[portf, 't-value'] = res.params['const']
    t_p_values.loc[portf, 't-value'] = res.tvalues['const']
    t_p_values.loc[portf, 'p-value'] = round(res.pvalues['const'], 4)
```

```
[ ]: capm_report
```

```
[ ]:
alpha_hat  beta_hat  info_ratio  treynor_ratio  Mean Return
SMALL LoBM    -0.0987    1.3600    -0.5750         0.0077         0.0105
ME1 BM2         0.0018    1.1702     0.0122         0.0818         0.0957
ME1 BM3         0.0097    1.0365     0.0857         0.0896         0.0929
ME1 BM4         0.0425    0.9652     0.3585         0.1243         0.1200
SMALL HiBM     0.0526    0.9907     0.3778         0.1333         0.1321
ME2 BM1     -0.0475    1.3433    -0.3677         0.0449         0.0603
ME2 BM2         0.0101    1.1286     0.0984         0.0892         0.1006
ME2 BM3         0.0248    1.0193     0.2712         0.1046         0.1066
ME2 BM4         0.0303    0.9654     0.3094         0.1116         0.1077
ME2 BM5         0.0292    1.1086     0.2363         0.1066         0.1181
ME3 BM1     -0.0316    1.2767    -0.2936         0.0555         0.0708
ME3 BM2         0.0204    1.0896     0.2549         0.0990         0.1078
ME3 BM3         0.0136    0.9743     0.1754         0.0942         0.0918
ME3 BM4         0.0294    0.9797     0.3253         0.1102         0.1080
ME3 BM5         0.0413    1.0407     0.3586         0.1200         0.1249
ME4 BM1     -0.0030    1.1919    -0.0353         0.0778         0.0927
ME4 BM2         0.0093    1.0615     0.1465         0.0890         0.0945
ME4 BM3         0.0135    0.9989     0.1753         0.0937         0.0936
ME4 BM4         0.0215    0.9749     0.2566         0.1023         0.0998
ME4 BM5         0.0285    1.0410     0.2548         0.1076         0.1120
```

BIG LoBM	0.0092	0.9934	0.1748	0.0895	0.0889
ME5 BM2	0.0071	0.9289	0.1270	0.0879	0.0816
ME5 BM3	0.0112	0.8712	0.1607	0.0931	0.0811
ME5 BM4	-0.0052	0.9010	-0.0533	0.0745	0.0671
BIG HiBM	0.0149	1.0150	0.1127	0.0949	0.0963

0.5 Testing the CAPM Model

- Conduct a cross-sectional regression on the individual regressions we just performed above.

```
[ ]: y = subsample.mean()
X = sm.add_constant(capm_report["beta_hat"])
capm_reg = sm.OLS(y,X, missing = 'drop').fit()

capm_reg.summary()
```

```
[ ]: <class 'statsmodels.iolib.summary.Summary'>
"""
                                OLS Regression Results
=====
Dep. Variable:                  y      R-squared:                0.258
Model:                        OLS      Adj. R-squared:           0.226
Method:                    Least Squares  F-statistic:                8.008
Date:                Mon, 24 Oct 2022  Prob (F-statistic):       0.00950
Time:                        11:44:05  Log-Likelihood:            123.51
No. Observations:                25    AIC:                       -243.0
Df Residuals:                    23    BIC:                       -240.6
Df Model:                        1
Covariance Type:                nonrobust
=====
               coef      std err          t      P>|t|      [0.025      0.975]
-----
const           0.0165      0.003      5.362      0.000      0.010      0.023
beta_hat       -0.0082      0.003     -2.830      0.009     -0.014     -0.002
=====
Omnibus:                 4.531   Durbin-Watson:           1.311
Prob(Omnibus):            0.104   Jarque-Bera (JB):         2.975
Skew:                    -0.820   Prob(JB):                 0.226
Kurtosis:                 3.408   Cond. No.                 17.0
=====

Notes:
[1] Standard Errors assume that the covariance matrix of the errors is correctly
specified.
"""
```

```
[ ]: print("Alpha: {}; Beta {}: ; Rsquared: {}".format(capm_reg.
    ↪params[0]*12,capm_reg.params[1]*12,capm_reg.rsquared))
```

Alpha: 0.19798077049490984; Beta -0.09815760379979337: ; Rsquared:
0.2582451192136339

The **CAPM** says the beta is the only risk associated with excess returns. In this case, the r-squared should be 100% and alpha should be zero, however, this is not what we see in the regression data. The **Betas** of the portfolios only explain 25% of the variation of the excess returns of the portfolios.

- There is no evidence suggesting size and value portfolios contain a premium unrelated to the market. However, we can say that risk matters beyond just beta since α is clearly not close to 0 from the regression.

0.6 Extensions

Re-do the analysis of the CAPM and use all three factors *MKT, SMB, HML*. Thus you will be testing the **FAMA-French 3-factor Model**.

```
[ ]: ff_data = portfolios_ex.join(df)["1981:"]
ff_data
```

```
[ ]:
      SMALL LoBM  ME1 BM2  ME1 BM3  ME1 BM4  SMALL HiBM  ME2 BM1  \
Date
1981-01-31    -0.0549 -0.0149  0.0099  0.0059      0.0211 -0.0597
1981-02-28    -0.0447 -0.0105  0.0014  0.0092      0.0144 -0.0145
1981-03-31     0.0607  0.0824  0.0736  0.0777      0.0610  0.0681
1981-04-30     0.0181  0.0359  0.0246  0.0365      0.0441  0.0060
1981-05-31     0.0382  0.0378  0.0324  0.0134      0.0168  0.0505
...
2022-04-30    -0.1616 -0.1325 -0.1000 -0.0635     -0.0411 -0.1820
2022-05-31    -0.0586 -0.0359 -0.0101 -0.0018      0.0293 -0.0332
2022-06-30    -0.0671 -0.0433 -0.0431 -0.0663     -0.0994 -0.0211
2022-07-31     0.1530  0.1287  0.0818  0.0955      0.0563  0.1272
2022-08-31     0.0362 -0.0033  0.0034 -0.0021     -0.0045  0.0348
```

```
      ME2 BM2  ME2 BM3  ME2 BM4  ME2 BM5  ...  ME4 BM5  BIG LoBM  \
Date
1981-01-31 -0.0107 -0.0127  0.0076  0.0206  ...  0.0039 -0.0717
1981-02-28  0.0051  0.0174  0.0244  0.0279  ...  0.0272  0.0219
1981-03-31  0.0715  0.0663  0.0663  0.0632  ...  0.0649  0.0174
1981-04-30  0.0247  0.0266  0.0268  0.0294  ...  0.0063 -0.0383
1981-05-31  0.0103  0.0169 -0.0086  0.0136  ...  0.0220  0.0043
...
2022-04-30 -0.1208 -0.0855 -0.0630 -0.0598  ... -0.0706 -0.1087
2022-05-31 -0.0285  0.0198  0.0401  0.0572  ...  0.0826 -0.0304
2022-06-30 -0.0630 -0.0685 -0.0628 -0.1303  ... -0.1330 -0.0787
2022-07-31  0.1357  0.1155  0.0992  0.0914  ...  0.0608  0.1264
2022-08-31 -0.0101 -0.0468 -0.0403 -0.0266  ... -0.0056 -0.0552
```

```
      ME5 BM2  ME5 BM3  ME5 BM4  BIG HiBM  Mkt-RF      SMB      HML  \
Date
```

1981-01-31	-0.0836	-0.0743	-0.0100	0.0095	-0.0504	0.0292	0.0672
1981-02-28	0.0094	-0.0164	0.0204	-0.0169	0.0057	-0.0034	0.0102
1981-03-31	0.0199	-0.0089	0.0222	0.0696	0.0356	0.0354	0.0064
1981-04-30	-0.0535	-0.0380	0.0104	-0.0404	-0.0211	0.0440	0.0228
1981-05-31	-0.0328	-0.0314	0.0267	-0.0060	0.0011	0.0200	-0.0042
...
2022-04-30	-0.0922	-0.0716	-0.0553	-0.0743	-0.0946	-0.0141	0.0619
2022-05-31	-0.0035	0.0278	0.0644	0.0637	-0.0034	-0.0185	0.0841
2022-06-30	-0.0559	-0.0884	-0.1216	-0.1224	-0.0843	0.0209	-0.0597
2022-07-31	0.0543	0.0410	0.0727	0.0672	0.0957	0.0281	-0.0410
2022-08-31	-0.0330	-0.0210	-0.0014	-0.0324	-0.0378	0.0139	0.0031

RF

Date

1981-01-31	0.0104
1981-02-28	0.0107
1981-03-31	0.0121
1981-04-30	0.0108
1981-05-31	0.0115
...	...
2022-04-30	0.0001
2022-05-31	0.0003
2022-06-30	0.0006
2022-07-31	0.0008
2022-08-31	0.0019

[500 rows x 29 columns]

```
[ ]: ff_report = pd.DataFrame(index=portfolios_ex.columns)
rhs = sm.add_constant(ff_data[['Mkt-RF', 'SMB', 'HML']])

for portf in portfolios_ex.columns:
    lhs = ff_data[portf]
    res = sm.OLS(lhs, rhs, missing='drop').fit()
    ff_report.loc[portf, 'alpha_hat'] = res.params['const'] * 12
    ff_report.loc[portf, 'beta_mkt'] = res.params['Mkt-RF']
    ff_report.loc[portf, 'beta_s'] = res.params['SMB']
    ff_report.loc[portf, 'beta_v'] = res.params['HML']
    ff_report.loc[portf, 'info_ratio'] = np.sqrt(12) * res.params['const'] / \
↳ res.resid.std()
    ff_report.loc[portf, 'treynor_ratio'] = 12 * ff_data[portf].mean() / res.
↳ params['Mkt-RF']
ff_report
```

```
[ ]:          alpha_hat  beta_mkt  beta_s  beta_v  info_ratio  treynor_ratio
SMALL LoBM -8.3200e-02    1.1162  1.3722 -0.2629    -0.9288         0.0094
ME1 BM2     5.9122e-03    0.9722  1.3145 -0.0126     0.0847         0.0985
```

ME1 BM3	8.6004e-05	0.9216	1.0463	0.2774	0.0019	0.1008
ME1 BM4	2.3969e-02	0.8783	1.0580	0.4768	0.5033	0.1366
SMALL HiBM	2.4209e-02	0.9359	1.0610	0.6956	0.3231	0.1411
ME2 BM1	-2.9881e-02	1.1424	1.0150	-0.3323	-0.4994	0.0528
ME2 BM2	7.0806e-03	1.0120	0.9012	0.1201	0.1421	0.0994
ME2 BM3	8.8349e-03	0.9730	0.7078	0.3982	0.1705	0.1096
ME2 BM4	6.2321e-03	0.9409	0.7422	0.5797	0.1390	0.1145
ME2 BM5	-5.2842e-03	1.0933	0.9219	0.8224	-0.1127	0.1080
ME3 BM1	-1.2503e-02	1.1073	0.7554	-0.3810	-0.2291	0.0640
ME3 BM2	1.5043e-02	1.0265	0.5759	0.1534	0.2688	0.1050
ME3 BM3	-3.8841e-03	0.9767	0.3964	0.4131	-0.0688	0.0940
ME3 BM4	2.6522e-03	1.0074	0.4339	0.6212	0.0480	0.1072
ME3 BM5	4.9672e-03	1.0893	0.5127	0.8410	0.0739	0.1146
ME4 BM1	1.5052e-02	1.0708	0.4376	-0.3754	0.2802	0.0866
ME4 BM2	6.3776e-04	1.0609	0.2096	0.2060	0.0111	0.0891
ME4 BM3	-5.1411e-03	1.0385	0.1565	0.4235	-0.0822	0.0901
ME4 BM4	-2.7009e-03	1.0271	0.2003	0.5520	-0.0444	0.0971
ME4 BM5	-8.0351e-03	1.1270	0.2497	0.8289	-0.1129	0.0994
BIG LoBM	2.3672e-02	0.9817	-0.2591	-0.3381	0.6931	0.0906
ME5 BM2	3.5560e-03	0.9703	-0.2118	0.0657	0.0706	0.0841
ME5 BM3	-3.0737e-03	0.9474	-0.2063	0.3051	-0.0552	0.0856
ME5 BM4	-3.6379e-02	1.0336	-0.2075	0.6832	-0.6521	0.0650
BIG HiBM	-2.4203e-02	1.1720	-0.1965	0.8587	-0.2693	0.0822

Re-do the analysis of 3.3 and 3.3, but instead of using the market return as the factor, use a new factor: the in-sample tangency portfolio of the $n = 25$ portfolios. You will not use the factor data for this problem!

```
[ ]: y = subsample.mean()

# The regressor, (x): the market beta from each of the n = 25 time-series
↳ regressions.
X = sm.add_constant(ff_report[['beta_mkt', 'beta_s', 'beta_v']])

res = sm.OLS(y,X,missing='drop').fit()
res.summary()
```

```
[ ]: <class 'statsmodels.iolib.summary.Summary'>
"""

                        OLS Regression Results
=====
Dep. Variable:          y      R-squared:                0.467
Model:                  OLS      Adj. R-squared:          0.390
Method:                 Least Squares      F-statistic:        6.121
Date:                   Mon, 24 Oct 2022      Prob (F-statistic):    0.00369
Time:                   11:59:42      Log-Likelihood:       127.63
No. Observations:       25      AIC:                 -247.3
Df Residuals:           21      BIC:                 -242.4
```

```

Df Model:                                3
Covariance Type:                        nonrobust
=====
              coef      std err          t      P>|t|      [0.025      0.975]
-----
const          0.0150      0.005      3.311      0.003      0.006      0.024
beta_mkt       -0.0080      0.004     -1.845      0.079     -0.017      0.001
beta_s          0.0002      0.001      0.361      0.721     -0.001      0.002
beta_v          0.0030      0.001      3.705      0.001      0.001      0.005
=====
Omnibus:                15.621    Durbin-Watson:                1.265
Prob(Omnibus):           0.000    Jarque-Bera (JB):           16.758
Skew:                    -1.490    Prob(JB):                   0.000230
Kurtosis:                 5.685    Cond. No.                   30.5
=====

```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

"""

```

[ ]: print(f"R-squared: {res.rsquared}")
      print(f"Intercept: {res.params[0] * 12}")
      print(f"Regression coefficient for market beta: {res.params[1]*12}")
      print(f"Regression coefficient for size beta: {res.params[2]*12}")
      print(f"Regression coefficient for value beta: {res.params[3]*12}")

```

```

R-squared: 0.46651395727533507
Intercept: 0.1797402802352507
Regression coefficient for market beta: -0.09567612643110464
Regression coefficient for size beta: 0.0028937271062240573
Regression coefficient for value beta: 0.036210032456927506

```

```

[ ]: print(f"Expected market excess return: {ff_data['Mkt-RF'].mean()*12}")
      print(f"Expected premium to size factor: {ff_data['SMB'].mean()*12}")
      print(f"Expected premium to value factor: {ff_data['HML'].mean()*12}")

```

```

Expected market excess return: 0.08024879999999998
Expected premium to size factor: 0.0092952
Expected premium to value factor: 0.032952000000000016

```

```

[ ]: # 4.2
      def compute_tangency(df_tilde, diagonalize_Sigma=False):

          """Compute tangency portfolio given a set of excess returns.

          Also, for convenience, this returns the associated vector of average

```

```

returns and the variance-covariance matrix.

Parameters
-----

diagonalize_Sigma: bool

    When `True`, set the off diagonal elements of the variance-covariance
    matrix to zero.

"""
Sigma = df_tilde.cov()

# N is the number of assets
N = Sigma.shape[0]

Sigma_adj = Sigma.copy()

if diagonalize_Sigma:

    Sigma_adj.loc[:, :] = np.diag(np.diag(Sigma_adj))

mu_tilde = df_tilde.mean()

Sigma_inv = np.linalg.inv(Sigma_adj)

weights = Sigma_inv @ mu_tilde / (np.ones(N) @ Sigma_inv @ mu_tilde)

# For convenience, I'll wrap the solution back into a pandas.Series object.
omega_tangency = pd.Series(weights, index=mu_tilde.index)

return omega_tangency, mu_tilde, Sigma_adj

```

```
[ ]: omega_tangency, mu_silde, Sigma_adj = compute_tangency(subsample)
```

```
[ ]: omega_tangency
```

```
[ ]: SMALL LoBM    -2.2576
      ME1 BM2      0.8976
      ME1 BM3     -0.0612
      ME1 BM4      1.3570
      SMALL HiBM   0.9409
      ME2 BM1     -0.3169
      ME2 BM2      0.8624
```

```

ME2 BM3      0.2559
ME2 BM4     -0.3609
ME2 BM5     -0.8624
ME3 BM1     -0.5990
ME3 BM2      0.1254
ME3 BM3     -0.5004
ME3 BM4      0.2890
ME3 BM5      0.2625
ME4 BM1      1.2849
ME4 BM2     -0.5254
ME4 BM3     -0.3362
ME4 BM4      0.0566
ME4 BM5      0.2444
BIG LoBM     1.0422
ME5 BM2     -0.2006
ME5 BM3      0.0697
ME5 BM4     -0.9925
BIG HiBM     0.3244
dtype: float64

```

```
[ ]: ex_return_tan = subsample @ omega_tangency
```

```
[ ]: tan_report = pd.DataFrame(index=portfolios_ex.columns)
rhs = sm.add_constant(ex_return_tan)

for portf in portfolios_ex.columns:
    lhs = ff_data[portf]
    res = sm.OLS(lhs, rhs, missing='drop').fit()
    tan_report.loc[portf, 'alpha_hat'] = res.params['const'] * 12
    tan_report.loc[portf, 'beta_hat'] = res.params[0]
    tan_report.loc[portf, 'info_ratio'] = np.sqrt(12) * res.params['const'] / \
↳ res.resid.std()
tan_report
```

```
[ ]:
alpha_hat  beta_hat  info_ratio
SMALL LoBM 1.7564e-16  0.0259  6.4488e-16
ME1 BM2    3.8641e-16  0.2366  1.6769e-15
ME1 BM3    4.6577e-16  0.2296  2.4385e-15
ME1 BM4   -3.0314e-16  0.2965 -1.6787e-15
SMALL HiBM 3.9942e-16  0.3264  2.0418e-15
ME2 BM1   -3.9031e-16  0.1490 -1.6021e-15
ME2 BM2   -1.3791e-16  0.2487 -7.0331e-16
ME2 BM3    1.0434e-15  0.2635  5.9898e-15
ME2 BM4    2.6021e-16  0.2663  1.5310e-15
ME2 BM5    6.8695e-16  0.2919  3.3895e-15
ME3 BM1    4.9440e-17  0.1750  2.2188e-16
ME3 BM2   -2.6021e-17  0.2665 -1.4572e-16

```


ME3 BM3	5.9848e-16	0.2268	3.6672e-15
ME3 BM4	6.6093e-16	0.2669	3.9408e-15
ME3 BM5	7.6501e-16	0.3086	4.0801e-15
ME4 BM1	5.8287e-16	0.2290	2.9486e-15
ME4 BM2	4.0593e-16	0.2336	2.3908e-15
ME4 BM3	4.4756e-16	0.2313	2.6944e-15
ME4 BM4	4.8919e-16	0.2466	2.9626e-15
ME4 BM5	0.0000e+00	0.2769	0.0000e+00
BIG LoBM	1.0408e-16	0.2197	6.6584e-16
ME5 BM2	3.6950e-16	0.2017	2.4842e-15
ME5 BM3	4.6317e-16	0.2004	3.1685e-15
ME5 BM4	2.6021e-17	0.1659	1.5621e-16
BIG HiBM	5.0741e-16	0.2380	2.5446e-15

```
[ ]: y = subsample.mean()

# The regressor, (x): the market beta from each of the n = 25 time-series
↳ regressions.
X = sm.add_constant(tan_report['beta_hat'])

res = sm.OLS(y,X,missing='drop').fit()
res.summary()
```

```
[ ]: <class 'statsmodels.iolib.summary.Summary'>
"""
```

```

                                OLS Regression Results
=====
Dep. Variable:                  y      R-squared:                  1.000
Model:                            OLS      Adj. R-squared:              1.000
Method:                    Least Squares      F-statistic:                1.308e+29
Date:                Mon, 24 Oct 2022      Prob (F-statistic):          1.08e-320
Time:                        12:07:26      Log-Likelihood:              918.63
No. Observations:                25      AIC:                        -1833.
Df Residuals:                    23      BIC:                        -1831.
Df Model:                        1
Covariance Type:                nonrobust
=====
               coef      std err          t      P>|t|      [0.025      0.975]
-----
const      -3.253e-18   2.24e-17     -0.145     0.886   -4.96e-17   4.31e-17
beta_hat      0.0337   9.32e-17    3.62e+14     0.000     0.034     0.034
=====
Omnibus:                 1.317      Durbin-Watson:              1.976
Prob(Omnibus):            0.518      Jarque-Bera (JB):            1.170
Skew:                    -0.480      Prob(JB):                    0.557
Kurtosis:                 2.549      Cond. No.                     17.7
=====
```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

"""

```
[ ]: print(f"R-squared: {res.rsquared}")
      print(f"Intercept: {res.params[0] * 12}")
      print(f"Regression coefficient for tangency beta: {res.params[1]}")
```

R-squared: 1.0

Intercept: -3.903127820947816e-17

Regression coefficient for tangency beta: 0.033720564437420955

```
[ ]: print(f"Expected premium to tangency portfolio: {ex_return_tan.mean()}")
```

Expected premium to tangency portfolio: 0.03372056443742081

Solution: The cross-sectional regression coefficient for the tangency beta is exactly the same as the expected premium of the tangency portfolio.

```
[ ]: MAE_alpha = (100 * capm_report['alpha_hat']).abs().mean()
      print('MAE = {:.2f} %'.format(MAE_alpha))
```

MAE = 2.43 %

Under classic statistical assumptions, we can test the null hypothesis that the CAPM works by calculating the following:

```
[ ]: T = subsample.dropna().shape[0]
      SR = capm_data['Mkt-RF'].mean() / capm_data['Mkt-RF'].std()
      Sigma = bm_residuals.cov()
      Sigma_inv = pd.DataFrame(np.linalg.inv(Sigma), index=Sigma.index, columns=Sigma.
                               ↪columns)
      alpha = capm_report['alpha_hat']
      H = T * (1 + SR**2)**(-1) * alpha @ Sigma_inv @ alpha
      print('H = {:.2f}'.format(H))
      pvalue = 1 - stats.chi2.cdf(H, df=25)
      print('p-value = {:.4f}'.format(pvalue))
```

H = 20068.02

p-value = 0.0000

```
[ ]: # (c)
      display(t_p_values.sort_values(by='p-value', ascending=False))
```

	t-value	p-value
ME1 BM2	0.0775	0.9382
ME4 BM1	-0.2250	0.8221
ME5 BM4	-0.3397	0.7342
ME1 BM3	0.5468	0.5847

ME2 BM2	0.6276	0.5306
BIG HiBM	0.7185	0.4728
ME5 BM2	0.8102	0.4182
ME4 BM2	0.9342	0.3506
ME5 BM3	1.0249	0.3059
BIG LoBM	1.1152	0.2653
ME4 BM3	1.1178	0.2642
ME3 BM3	1.1189	0.2637
ME2 BM5	1.5073	0.1324
ME4 BM5	1.6250	0.1048
ME3 BM2	1.6255	0.1047
ME4 BM4	1.6369	0.1023
ME2 BM3	1.7297	0.0843
ME3 BM1	-1.8726	0.0617
ME2 BM4	1.9735	0.0490
ME3 BM4	2.0747	0.0385
ME1 BM4	2.2864	0.0227
ME3 BM5	2.2872	0.0226
ME2 BM1	-2.3451	0.0194
SMALL HiBM	2.4095	0.0163
SMALL LoBM	-3.6673	0.0003

Which is a stricter test: checking if individual alphas are individually significant or if they are all jointly significant?

The stricter test is joint significance because it means that those portfolio's factors are all jointly different from 0.

Conceptually, how does the test-statistic H relate to checking whether \tilde{r}^m spans the tangency portfolio?

Answer:

- We have demonstrated that an LFM is exactly the same as asserting that a set of factors spans the MV frontier, (and thus spans the tangency portfolio.) Thus, this test can be interpreted as checking whether investing in the alphas gets beyond the tangency portfolio. The test statistic is basically comparing the square SR of the alphas to the square SR of the factors.