CSCI5525 Assignment 1 Answer

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1 Problem 1

(1) Let $\frac{\partial E[l(f(x),y)]}{\partial f(x)} = 0$, then

$$\int_{y} (f(x) - y)p(y|x)p(x)dy = 0$$
$$\int_{y} f(x)p(y|x)p(x)dy = \int_{y} yp(y|x)p(x)dy$$

Since f(x) and p(x) is independent of y, and $\int_{\mathcal{U}} p(y|x) dy = 1$,

$$f(x)p(x) \int_{y} p(y|x) dy = p(x) \int_{y} yp(y|x) dy$$
$$f(x) = \int_{y} yp(y|x) dy$$
$$= E[y|x]$$

Thus the optimal f(x) is E[y|x], where $E[y|x] = \int_y y p(y|x) \mathrm{d}y$.

(2) Let $\frac{\partial E[l(f(x),y)]}{\partial f(x)} = 0$, then

$$\int \operatorname{sgn}(f(x) - y)p(y|x)p(x)dy = 0$$

$$\int_{-\infty}^{f(x)} p(y|x)p(x) = \int_{f(x)}^{\infty} p(y|x)p(x)$$

which means the optimal f(x) is the median of the distribution of y, i.e., $p(y \le f(x)|x) = 0.5$.

2 Problem 2

The expectation for y is given by:

$$E[y] = \sum_{i=1:M} \int_{x \in \mathcal{R}_i} p(y_i, x) dx$$

The error rate for class $y = C_j$ equals to the rate that x is in \mathcal{R}_k where $k \neq j$. Thus

$$err[y = C_j] = \sum_{i=1:M, i \neq j} \int_{x \in \mathcal{R}_i} p(y_i, x) dx$$

$$= \sum_{i=1:M, i \neq j} \int_{x \in \mathcal{R}_i} p(C_j | x) p(x) dx$$

$$= \sum_{i=1:M} \int_{x \in \mathcal{R}_i} p(C_j | x) p(x) dx - \int_{x \in \mathcal{R}_j} p(C_j | x) p(x) dx$$

Proved.

3 Problem 3

3.1 (a)

First modify the target T of boston dataset so that p(y=0)=0.5 and p(y=1)=0.5. And then Fisher's linear discriminant analysis (LDA) is applied to project boston dataset to 1 dimension. Within-class covariance S_w is given by:

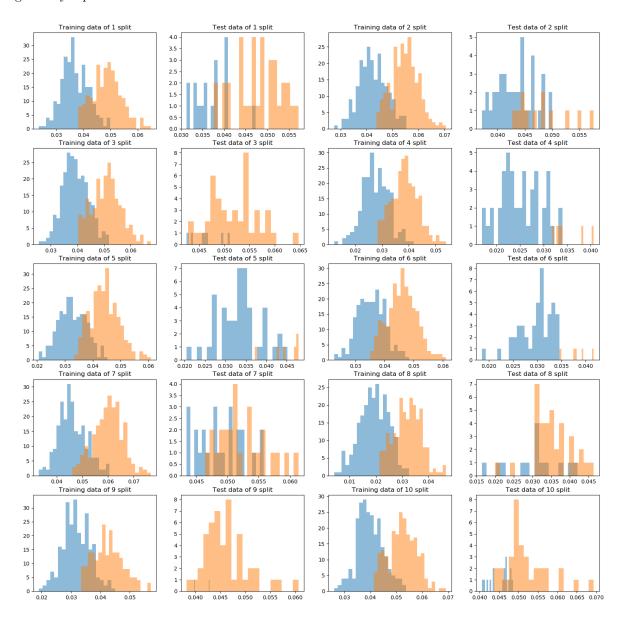
$$S_w = \sum_{\mathbf{x}_n \in C_1} (\mathbf{x}_n - \mathbf{m}_1)(x_n - \mathbf{m}_1)^T + \sum_{\mathbf{x}_n \in C_2} (\mathbf{x}_n - \mathbf{m}_2)(\mathbf{x}_n - \mathbf{m}_2)^T$$

where m_i is mean of class i, and weight vector \mathbf{w} is given by:

$$\mathbf{w} \propto S_w^{-1} (\mathbf{m}_2 - \mathbf{m}_1)$$

and the projected data $f(x) = \mathbf{w}^T \mathbf{x}$.

The results of projected training data and test data of 10 cross-validation are below. Noted that LDA generally separated the two classes.



3.2 (b)

No, since boston dataset has two classes after modification and LDA can project datasets to a maximum of K-1 dimensions, where K is the number of classes.

This is because the maximum rank of between-class covariance S_b is K-1, and thus for $S_w^{-1}S_b$ there are at most K-1 eigenvalues and corresponding K-1 eigenvectors. The subspace dimension of the projected data is given by the eigenvectors, which is at most K-1.

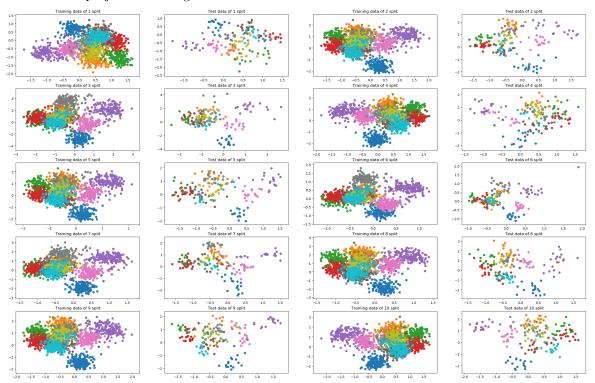
3.3 (c)

Between-class and within-class covariance S_w is given by:

$$S_b = \sum_{i=1}^K N_i (\mathbf{m}_i - \mathbf{m}) (\mathbf{m}_i - \mathbf{m})^T$$
$$S_w = \sum_{i=1}^K \sum_{x_n \in C_i} (\mathbf{x}_n - \mathbf{m}_i) (x_n - \mathbf{m}_i)^T$$

where K is the number of classes, N_i is the number of instances in class i, m_i is mean of class i. The weight vector \mathbf{w} is given by the top 2 eigenvectors of $S_w^{-1}S_b$ by the largest magnitude, and the projected data $f(x) = \mathbf{w}^T \mathbf{x}$.

The results of projected training data and test data of 10 cross-validation are below.



The testing error followed by Gaussian generative modeling is bellow

```
1 split:
                        0.31666667
Test error of
                split:
                        0.27222222
                split:
                        0.46666667
              4 split:
                        0.40555556]
              5 split:
                        0.35]
                         0.4555556
              6 split:
              8 split:
Test error of 9 split:
                       [0.31284916
Test error of 10 split: [0.31284916]
Test error mean: 0.36332712600869027
Test error std: 0.06076858471325458
```

4 Problem 4

4.1 (a) Logistic regression (LR)

Logistic regression is a discriminative model, and the class posteriors are given by:

$$p(C_k|\mathbf{x}) = \frac{\exp(a_k)}{\sum_j \exp(a_j)}$$

where $a_k = \mathbf{w}_k^T \mathbf{x}$. Then the likelihood is given by:

$$p(\mathbf{y}|\mathbf{w}_1,...,\mathbf{w}_k) = \prod_{n=1}^{N} \prod_{k=1}^{K} p(C_k|\mathbf{x}_n)^{y_{nk}} = \prod_{n=1}^{N} \prod_{k=1}^{K} \pi_{nk}^{y_{nk}}$$

The optimization problem of \mathbf{w} is solved by Iteratively Reweighted Least Squares (IRLS). Here the multi-class problem is treated as one-versus-the-rest problem. The update of \mathbf{w} is then given by:

$$\mathbf{w}^{new} = \mathbf{w}^{old} - H^{-1}(\mathbf{w}^{old}) \nabla E(\mathbf{w}^{old})$$
$$= (X^T R X)^{-1} X^T R \mathbf{z}$$

where $\mathbf{z} = X\mathbf{w}^{old} - R^{-1}(\pi - \mathbf{y})$, and R is a diagonal matrix initialized with $r^i = 1$, which is updated by

$$r_i = \frac{1}{\max(\delta, |y_i - X^i \mathbf{w}|)}$$

where δ is some small value.

4.2 (b) Naive-Bayes with marginal Gaussian distributions (GNB)

For K-class problem of Naive-Bayes, the posterior probability for class k:

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{p(\mathbf{x})} = \frac{\exp(a_k)}{\sum_{j=1}^{K} \exp(a_j)}$$

were a_k is given by:

$$a_k = \log p(\mathbf{x}|C_k) + \log p(C_k)$$

And for marginal Gaussian distribution, $p(x_i|C_k)$ is given by:

$$p(x_i|C_k) = \mathcal{N}(\mu_{ik}, \sigma_{ik}^2)$$
$$p(\mathbf{x}|C_k) = \prod_{n=1}^{D} p(x_i|C_k) = \frac{1}{(2\pi)^{D/2}(\prod_{i=1}^{D} \sigma_{ik})} \exp\left(-\sum_{i=1}^{D} \frac{(x_i - \mu_{ik})^2}{2\sigma_{ik}^2}\right)$$

where μ_k and σ_k^2 are means and stds of class k of training data.

The results of LR and GNB are printed below:

