lab05

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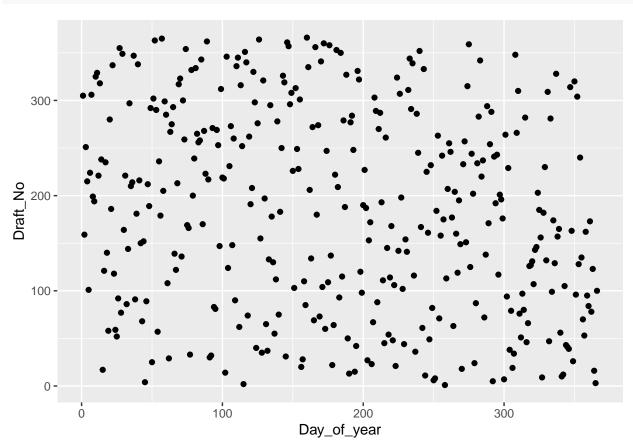
2022-12-13

Question 1: Hypothesis testing

```
set.seed(12345)
lottery=read.csv("lottery.csv", header =TRUE, sep =";")
```

Task 1:Make a scatter plot

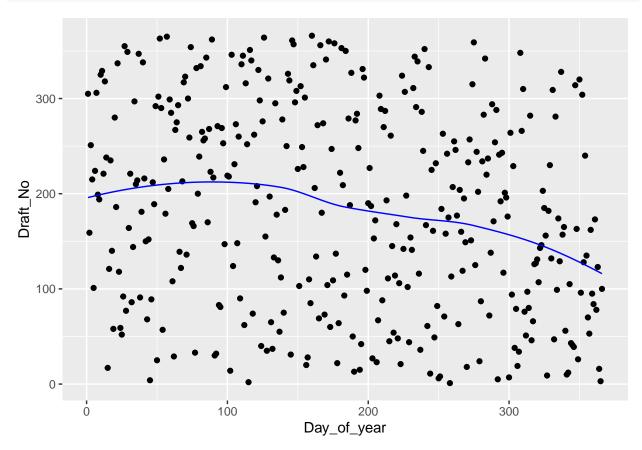
ggplot(data=lottery)+geom_point(aes(x=Day_of_year,y=Draft_No))



The lottery looks random,X and Y seems independent.

Task 2:Plot a loess smoother curve

```
fit=loess(lottery$Draft_No~lottery$Day_of_year,data=lottery)
Y_predict=predict(fit,data=lottery)
lottery1=lottery%>%mutate(predict_Y=Y_predict)
ggplot(data=lottery1)+geom_point(aes(x=Day_of_year,y=Draft_No))+
    geom_line(aes(x=Day_of_year,y=predict_Y),color="blue")
```



We use loess smoother to fit a curve and plot it in the previous graph,now it seems there are slight correlation between X and Y.In the first half year,more Draft No may be larger,but in the second half,more Draft No may be smaller.

Task 3:Estimate the distribution of T using non-parametric bootstrap

```
stat=function(data,vn){
  data=as.data.frame(data[vn,])
  fit_bootstarp=loess(Draft_No~Day_of_year,data=data)
  y_predict_bs=predict(fit_bootstarp,data=data)

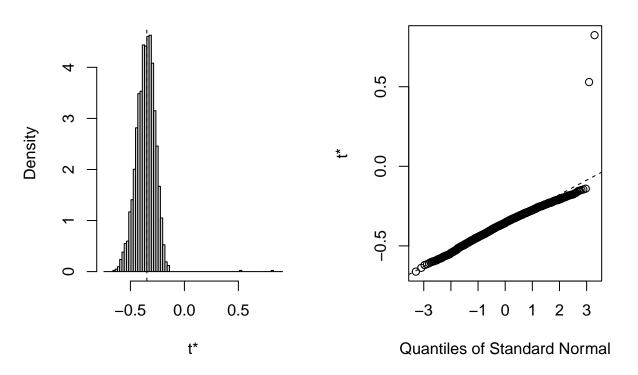
y_predict_max=max(y_predict_bs)
  y_predict_min=min(y_predict_bs)

x_max=data$Day_of_year[which.max(y_predict_bs)]
  x_min=data$Day_of_year[which.min(y_predict_bs)]

t=((y_predict_max-y_predict_min)/(x_max-x_min))
```

```
return(t)
}
res=boot(lottery,stat,R=2000)
boot.ci(res)
## BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
## Based on 2000 bootstrap replicates
##
## CALL :
## boot.ci(boot.out = res)
##
## Intervals :
## Level
              Normal
                                   Basic
         (-0.5146, -0.1600)
                                (-0.4860, -0.1553)
## 95%
##
             Percentile
## Level
                                    BCa
         (-0.5405, -0.2099)
## 95%
                                (-0.5249, -0.1977)
## Calculations and Intervals on Original Scale
plot(res)
```

Histogram of t



The distribution of T-value is around -0.35, which is significantly different from 0, so the lottery is not random.

p_value=mean(res\$t>=0)

p_value

[1] 0.001

Task 4:implement permutation test

```
B=2000
permu_test=function(data,B){
  Ms=stat(data,1:nrow(data))
  sta=numeric(B)
  n=dim(data)[1]
  for(b in 1:B){
    GB=sample(data$Day_of_year,n)
    data=data%>%mutate(Day_of_year=GB)
    sta[b]=stat(data,1:n)
  }
  #Ms=stat(lottery,1:nrow(lottery))
  #calculate p value
  #test is two-sided
  p_value_permut=mean(abs(sta)>=abs(Ms))
  return(p_value_permut)
res_permu_test(lottery,B)
res_permu
```

[1] 0.154

The p-value is around 0.15, so We can't reject H0, that lottery is random.

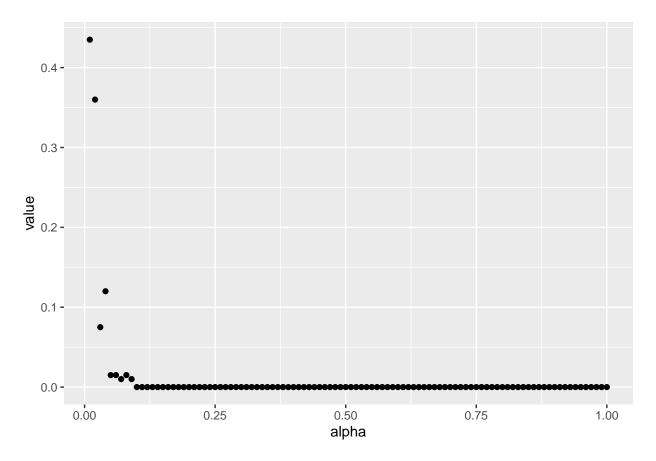
Task 5:Study the power

```
repeat_power=function(alpha){
    new_dataset=as.data.frame(lottery$Day_of_year)
    colnames(new_dataset)=c("Day_of_year")
    new_dataset=new_dataset%>%
        mutate(Draft_No=NA)
    for(i in 1:nrow(new_dataset)){
        new_dataset$Draft_No[i]=max(0,min(alpha*new_dataset$Day_of_year[i]+rnorm(1,183,sd=10),366))
}

res_permu=permu_test(new_dataset,200)
    return(res_permu)
}

power_data=sapply(seq(0.01,1,0.01),FUN=repeat_power)
res_power=data.frame("value"=power_data,"alpha"=seq(0.01,1,0.01))

ggplot(data=res_power)+geom_point(aes(x=alpha,y=value))
```



In this case,

H0:Lottery is random

H1:Lottery is non-random

When the alpha is 0.07, the p-value is 0.01, There is only 1% chance that the data distribution is random, so we reject the null hypothesis, hence the power is 0.99.

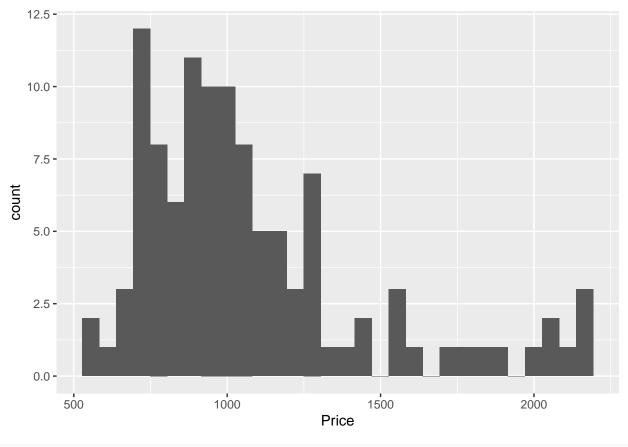
Consider when α become larger, The distribution of data will gradually become like non-random, hence the power will become larger. This is consistent with the visualization we observed, so the quality of the test statistics is good.

Assignment 2:Bootstrap, jackknife and confidence intervals

```
price=read.csv("prices1.csv",header =TRUE,sep =";")
```

Task 1:Histogram and mean price

```
ggplot(data=price)+geom_histogram(aes(x=Price),bins = 30)
```



```
cat("The mean price is :",mean(price$Price))
```

The mean price is : 1080.473

Does it remind any conventional distribution? The distribution looks likes gamma distribution.

Task 2
Estimate the distribution of the mean price of the house using bootstrap

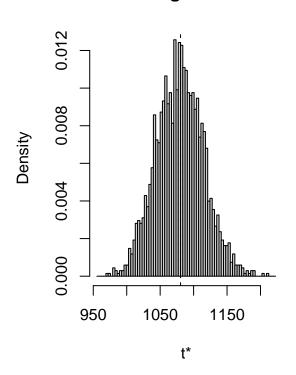
```
stat_mean_price=function(data,vn){
  data=data[vn,]
  t=mean(data$Price)
  return(t)
}

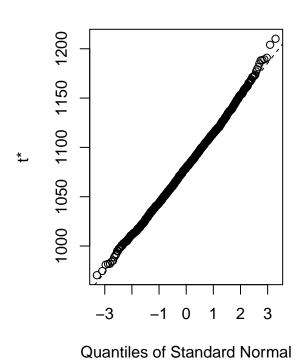
res=boot(price,stat_mean_price,R=2000)
res

##
## ORDINARY NONPARAMETRIC BOOTSTRAP
##
##
## Call:
## boot(data = price, statistic = stat_mean_price, R = 2000)
##
##
##
```

```
## Bootstrap Statistics :
## original bias std. error
## t1* 1080.473 -0.9710045 35.31229
plot(res)
```

Histogram of t





Bootstrap bias-correction and the variance of the mean price

Bias corrected estimator is:

$$T_1 := 2T(D) - \frac{1}{B} \sum_{i=1}^{B} T_i^*$$

```
cat("The variance of the mean price:",35.64^2,"\nThe bias correction:",
    2*mean(price$Price)-mean(res$t))
```

```
## The variance of the mean price: 1270.21
## The bias correction: 1081.444
```

Compute a 95% confidence interval

```
boot.ci(res)
## BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
## Based on 2000 bootstrap replicates
##
## CALL:
## boot.ci(boot.out = res)
```

```
##
## Intervals :
## Level
              Normal
                                   Basic
       (1012, 1151) (1009, 1148)
## 95%
## Level
                                    BCa
             Percentile
         (1013, 1152)
                          (1015, 1157)
## 95%
## Calculations and Intervals on Original Scale
The 95% confidence interval for the mean price using: bootstrap percentile: (1013,1152)
bootstrap BCa:(1015,1157)
first-order normal approximation: (1012,1151)
```

Task 3:Estimate the variance of the mean price using the jackknife and compare it with the bootstrap estimate

Jackknife (n = B):

$$\widehat{Var[T(\mathring{\mathbf{u}})]} = \frac{1}{n(n-1)} \sum_{i=1}^{n} ((T_i^*) - J(T))^2$$

```
where T_i^* = nT(D) - (n-1)T(D_i^*) J(T) = \frac{1}{n} \sum_{i=1}^n T_i^*
```

```
## The variance of the mean price using the jackknife: 1320.911 ## The variance of the mean price using the bootstrap: 1270.21
```

Compare to the bootstrap estimate, the variance of the mean price using the jackknife is larger, because Jackknife overestimate variance.

Task 4:Compare the confidence intervals obtained with respect to their length and the location of the estimated mean in these intervals

```
a=mean(price$Price)-1.96*sqrt(var_jackknife)
b=mean(price$Price)+1.96*sqrt(var_jackknife)
cat("The confidence intervals (95%) of means of price using jackknife is:(",a,",",b,")")
```

The confidence intervals (95%) of means of price using jackknife is:(1009.238 , 1151.708)

##		confidence_intervals	length	estimated_mean
##	Normal	(1009,1153)	144	1081.0
##	Basic	(1009,1151)	142	1080.0
##	Percentile	(1010,1152)	142	1081.0
##	BCa	(1013,1158)	145	1085.5

Appendix

Contribution statement

```
part 1: Shipeng Liu, Dongwei Ni
part 2: Shipeng Liu
library(ggplot2)
library(dplyr)
library(boot)
# part 1
set.seed(12345)
lottery=read.csv("lottery.csv",header =TRUE,sep =";")
ggplot(data=lottery)+geom_point(aes(x=Day_of_year,y=Draft_No))
fit=loess(lottery$Draft_No~lottery$Day_of_year,data=lottery)
Y_predict=predict(fit,data=lottery)
lottery1=lottery%>%mutate(predict_Y=Y_predict)
ggplot(data=lottery1)+geom_point(aes(x=Day_of_year,y=Draft_No))+
  geom_line(aes(x=Day_of_year,y=predict_Y),color="blue")
stat=function(data,vn){
  data=as.data.frame(data[vn,])
  fit_bootstarp=loess(Draft_No~Day_of_year, data=data)
  y_predict_bs=predict(fit_bootstarp,data=data)
 y_predict_max=max(y_predict_bs)
  y_predict_min=min(y_predict_bs)
  x_max=data$Day_of_year[which.max(y_predict_bs)]
  x_min=data$Day_of_year[which.min(y_predict_bs)]
 t=((y_predict_max-y_predict_min)/(x_max-x_min))
 return(t)
res=boot(lottery,stat,R=2000)
```

```
boot.ci(res)
plot(res)
p_value=mean(res$t>=0)
p_value
B=2000
permu test=function(data,B){
  Ms=stat(data,1:nrow(data))
  sta=numeric(B)
  n=dim(data)[1]
  for(b in 1:B){
    GB=sample(data$Day_of_year,n)
    data=data%>%mutate(Day_of_year=GB)
    sta[b]=stat(data,1:n)
  }
  #Ms=stat(lottery,1:nrow(lottery))
  #calculate p value
  #test is two-sided
  p_value_permut=mean(abs(sta)>=abs(Ms))
  return(p_value_permut)
res_permu=permu_test(lottery,B)
res permu
repeat_power=function(alpha){
  new_dataset=as.data.frame(lottery$Day_of_year)
  colnames(new_dataset)=c("Day_of_year")
  new_dataset=new_dataset%>%
    mutate(Draft_No=NA)
  for(i in 1:nrow(new_dataset)){
    new_dataset$Draft_No[i]=max(0,min(alpha*new_dataset$Day_of_year[i]+rnorm(1,183,sd=10),366))
  res_permu=permu_test(new_dataset,200)
  return(res_permu)
power_data=sapply(seq(0.01,1,0.01),FUN=repeat_power)
res_power=data.frame("value"=power_data, "alpha"=seq(0.01,1,0.01))
ggplot(data=res_power)+geom_point(aes(x=alpha,y=value))
## part 2
price=read.csv("prices1.csv",header =TRUE,sep =";")
ggplot(data=price)+geom_histogram(aes(x=Price),bins = 30)
cat("The mean price is :",mean(price$Price))
stat_mean_price=function(data,vn){
  data=data[vn,]
  t=mean(data$Price)
  return(t)
```

```
res=boot(price,stat_mean_price,R=2000)
plot(res)
cat("The variance of the mean price:",35.64^2,"\nThe bias correction:",
    2*mean(price$Price)-mean(res$t))
boot.ci(res)
jackknife=function(data,B){
  res=c()
  for(i in 1:B){
    res=c(res,mean(data$Price[-i]))
 return(res)
}
res_jackknife=jackknife(price,nrow(price))
n=nrow(price)
Ti_star=n*mean(price$Price)-(n-1)*res_jackknife
var_jackknife=sum(((Ti_star)-mean(Ti_star))^2)/(n*(n-1))
cat("The variance of the mean price using the jackknife: ", var jackknife,
    "\nThe variance of the mean price using the bootstrap:",35.64^2)
a=mean(price$Price)-1.96*sqrt(var_jackknife)
b=mean(price$Price)+1.96*sqrt(var jackknife)
cat("The confidence intervals (95%) of means of price using jackknife is:(",a,",",b,")")
ic_res=data.frame("confidence_intervals"=c("(1009,1153)","(1009,1151)",
                                    "(1010,1152)","(1013,1158)"),
           "length"=c(1153-1009,1151-1009,1152-1010,1158-1013),
           "estimated_mean"=c(((1153+1009)/2),((1151+1009)/2),
                              ((1152+1010)/2),((1158+1013)/2)))
rownames(ic_res)=c("Normal", "Basic", "Percentile", "BCa")
ic_res
```