Chernoff Bound

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We will prove a fairly general form of the Chernoff bound. This proof is due to Van Vu, who gave it at a lecture at the University of California, San Diego.

Theorem 1. Let $X_1, ..., X_n$ be discrete, independent random variables such that $E[X_i] = 0$ and $|X_i| \le 1$ for all i. Let $X = \sum_{i=1}^n X_i$ and σ^2 be the variance of X. Then

$$\Pr[|X| \ge \lambda \sigma] \le 2e^{-\lambda^2/4}$$

for any $0 \le \lambda \le 2\sigma$.

Proof. We will prove

$$\Pr[X \ge \lambda \sigma] \le e^{-\lambda^2/4}$$

which will suffice, by the symmetry of argument. Let $0 \le t \le 1$, to be determined later. Note that

$$\begin{array}{lcl} \Pr[X \geq \lambda \sigma] & = & \Pr[tX \geq t\lambda \sigma] \\ & = & \Pr[e^{tX} \geq e^{t\lambda \sigma}] \\ & \leq & \frac{\mathrm{E}[e^{tX}]}{e^{t\lambda \sigma}}, \end{array}$$

the last by the Markov Inequality.

Before going any further, we establish a bound on $E[e^{tZ}]$, where $-1 \le Z \le 1$ and E[Z] = 0. Additionally, let $t \le 1$. By the definition of expectation,

$$E[e^{tZ}] = \sum_{j=1}^{m} p_{j}e^{tz_{j}}$$

$$= \sum_{j=1}^{m} p_{j} (1 + tz_{j} + \frac{1}{2!}(tz_{j})^{2} + \frac{1}{3!}(tz_{j})^{3} + \cdots)$$

$$= \sum_{j=1}^{m} p_{j} + t \sum_{j=1}^{m} p_{j}z_{j} + \sum_{j=1}^{m} p_{j} (\frac{1}{2!}(tz_{j})^{2} + \frac{1}{3!}(tz_{j})^{3} + \cdots).$$

Summation *A* is the sum of all probabilities, so A = 1. Summation *B* is the expectation of *Z*, so B = 0. Since $|ta_i| \le 1$, we can upper-bound *C* by

$$\sum_{j=1}^m p_j(ta_j)^2 \left(\frac{1}{2!} + \frac{1}{3!} + \cdots\right) \le t^2 \sum_{j=1}^m p_j a_j^2.$$

But the above summation is just Var[Z], giving

$$E[e^{tZ}] \le 1 + t^2 Var[Z].$$

Returning to our main computation,

$$\begin{split} \mathbf{E}[e^{tX}] &=& \mathbf{E}[e^{t(X_1+X_2+\cdots+X_n)}] \\ &=& \mathbf{E}\big[\prod_{i=1}^n e^{tX_i}\big] \\ &=& \prod_{i=1}^n \mathbf{E}[e^{tX_i}] \qquad \text{by the independence of } X_i \\ &\leq& \prod_{i=1}^n \left(1+t^2 \mathrm{Var}[X_i]\right) \\ &\leq& \prod_{i=1}^n e^{t^2 \mathrm{Var}[X_i]} \qquad \text{since } 1+\alpha \leq e^\alpha \text{ for } \alpha \geq 0 \\ &=& e^{t^2\sigma^2}. \qquad \text{by the independence of } X_i \end{split}$$

Thus,

$$\Pr[X \ge \lambda \sigma] \le \frac{e^{t^2 \sigma^2}}{e^{t\lambda \sigma}}$$
$$= e^{t\sigma(t\sigma - \lambda)}.$$

Optimizing t we get $t = \lambda/2\sigma$, which gives

$$\Pr[X \ge \lambda \sigma] \le e^{-\lambda^2/4}.$$