

DSE ECONOMETRICS PROJECT

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LIFE EXPECTANCY AROUND THE WORLD

Answer 1:

According to The World Bank, “*Life expectancy at birth indicates the number of years a newborn infant would live if prevailing patterns of mortality at the time of its birth were to stay the same throughout its life*”.

According to World Health Organization (WHO), “*The average number of years that a newborn could expect to live, if he or she were to pass through life exposed to the sex- and age-specific death rates prevailing at the time of his or her birth, for a specific year, in a given country, territory, or geographic area*”.

For 2019:-

```
.  
. ***FOR 2019:  
.   
. *QUESTION 1:  
. *describing all the variables used in the data  
. desc
```

Contains data

```
obs:      208  
vars:       8  
size:    17,888
```

variable name	storage type	display format	value label	variable label
CountryName	str30	%30s		Country Name
lexp	double	%14.2f		lexp
co2	double	%14.2f		co2
pcrate	double	%14.2f		pcrate
pun	double	%14.2f		pun
hexp	double	%14.2f		hexp
gdppc	double	%14.2f		gdppc
stat	double	%14.2f		stat

Sorted by:

Note: Dataset has changed since last saved.

```
. *regressing lexp with all other explanatory variables to get their relation
. ///with each other
> reg lexp co2 pcrate pun hexp gdppc stat
```

Source	SS	df	MS	Number of obs	=	74
Model	1604.27303	6	267.378838	F(6, 67)	=	18.78
Residual	954.088994	67	14.2401342	Prob > F	=	0.0000
				R-squared	=	0.6271
				Adj R-squared	=	0.5937
Total	2558.36202	73	35.0460551	Root MSE	=	3.7736

lexp	Coeff.	Std. Err.	t	P> t	[95% Conf. Interval]
co2	-.4834918	.2793186	-1.73	0.088	-1.041014 .0740304
pcrate	.1161309	.0380442	3.05	0.003	.0401944 .1920674
pun	-.0895493	.0694143	-1.29	0.201	-.2281008 .0490021
hexp	-.0398276	.2026607	-0.20	0.845	-.4443401 .3646848
gdppc	.0004022	.0000884	4.55	0.000	.0002257 .0005786
stat	.0521342	.04474	1.17	0.248	-.0371672 .1414356
_cons	53.7987	4.457649	12.07	0.000	44.90119 62.6962

```
.
end of do-file
```

Here, *lexp* is the dependent (or explained) variable. *co2*, *pcrate*, *pun*, *hexp*, *gdppc* and *stat* are the independent (or explanatory) variables. All the variables mentioned above are continuous in nature. If we regress *lexp* with all other explanatory variables, then,

- co2* is statistically insignificant at 5% level of significance; which means as *co2* emission increases in each country, *lexp* decreases by 0.4833 years.
- pcrate* is also statistically significant at 5% level of significance; as *pcrate* increases by 1 unit, *lexp* increases by 0.116 years for each country.
- pun* is statistically insignificant at 5% level of significance, thus, as *pun* increases by 1%, *lexp* decreases by 0.0895 years for each country.
- hexp* is statistically insignificant at 5% level of significance; as *hexp* increases by 1% of GDP, *lexp* decreases by 0.0398 years.
- gdppc* is statistically significant at 5% level of significance which indicates that, as *gdppc* increases by 1 unit, *lexp* increases by 0.0004 years.
- stat* is statistically insignificant at 5% level of significance; thus, as *stat* increases by 1 unit, *lexp* increases by 0.052 years for each country.
- the constant term (= 53.79) is statistically significant at 5% level of significance.

Answer 2:

```
. *QUESTION 2:
. *generating a new variable deve which stores the value 0 if developing and
. /// 1 if developed
> *the development threshold for GDP(PPP) per capita of a developed country is
. ///atleast US $22,000
> *comparing gdppc of our data with the given value, we get
. gen deve = 0 if gdppc <= 22000
(90 missing values generated)

. replace deve = 1 if gdppc > 22000
(90 real changes made)

. replace deve = . if(missing(gdppc))
(20 real changes made, 20 to missing)
```

Here, we generated a new variable named *deve* which took the value ‘0’ if the country is developing and ‘1’ if the country is developed. We also replaced the value of *deve* with ‘.’ when it satisfies neither of the conditions mentioned above.

We had compared the value of *gdppc* with the given value (= 22000). This is the development threshold for GDP(PPP) per capita for a developed country (= US \$22,000).

¹ Source for the Development Threshold for GDP (PPP) per capita
[[https://en.wikipedia.org/wiki/Developed_country#:~:text=Another%20commonly%20used%20measure%20of,fit%20three%20out%20of%20four.&text=World%20map%20showing%20country%20classifications,\(last%20updated%20April%202023\).](https://en.wikipedia.org/wiki/Developed_country#:~:text=Another%20commonly%20used%20measure%20of,fit%20three%20out%20of%20four.&text=World%20map%20showing%20country%20classifications,(last%20updated%20April%202023).)]

Answer 3:

```
. *QUESTION 3:  
. *Finding the mean of all the variables except stat score and testing its  
. ///significance for both developing and developed countries  
> mean lexp co2 pcrate pun hexp gdppc if deve == 0
```

Mean estimation Number of obs = 66

	Mean	Std. Err.	[95% Conf. Interval]	
lexp	69.97071	.7210385	68.53069	71.41072
co2	2.057725	.2319248	1.594539	2.520911
pcrate	90.38655	1.785383	86.8209	93.95221
pun	11.26515	1.142248	8.983926	13.54638
hexp	5.741879	.2935322	5.155655	6.328103
gdppc	9999.214	819.077	8363.404	11635.02

```
. mean lexp co2 pcrate pun hexp gdppc if deve == 1
```

Mean estimation Number of obs = 40

	Mean	Std. Err.	[95% Conf. Interval]	
lexp	79.76691	.4994316	78.75672	80.77711
co2	7.869315	.7295501	6.393661	9.34497
pcrate	98.70799	.9928123	96.69984	100.7161
pun	2.875	.1943348	2.481921	3.268079
hexp	7.668487	.432519	6.793635	8.543339
gdppc	47506.83	3048.122	41341.42	53672.24

Here, we are finding the mean of all the variables except stat score for both developing and developed countries.

```
. ttest lexp, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	125	69.21329	.5656498	6.324158	68.09371	70.33287
1	63	79.63789	.449573	3.568375	78.7392	80.53657
combined	188	72.70664	.5413007	7.421942	71.6388	73.77448
diff		-10.4246	.8589884		-12.11921	-8.729986

```
diff = mean(0) - mean(1)                                t = -12.1359
Ho: diff = 0                                           degrees of freedom = 186
```

```
Ha: diff < 0                Ha: diff != 0                Ha: diff > 0
Pr(T < t) = 0.0000          Pr(|T| > |t|) = 0.0000          Pr(T > t) = 1.0000
```

Here, Life expectancy at birth (*lexp*) varies between -12.11 and -8.72

Standard error for the difference in life expectancy at birth is very less (= 0.858)

The null hypothesis is given by,

Ho: diff = 0

From the t-test, we find that for,

- Ha: diff < 0: Ho is rejected at 1% level of significance
- Ha: diff != 0: Ho is rejected at 1% level of significance
- Ha: diff > 0: Ho is accepted at 1% level of significance

```
. ttest co2, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	124	2.057643	.1831721	2.039718	1.695066	2.420221
1	56	8.757779	.7606969	5.692534	7.233308	10.28225
combined	180	4.14213	.3534802	4.742435	3.444606	4.839655
diff		-6.700135	.577988		-7.840726	-5.559545

```
diff = mean(0) - mean(1)                                t = -11.5922
Ho: diff = 0                                           degrees of freedom = 178
```

```
Ha: diff < 0                Ha: diff != 0                Ha: diff > 0
Pr(T < t) = 0.0000          Pr(|T| > |t|) = 0.0000          Pr(T > t) = 1.0000
```

- Ha: $\text{diff} < 0$: H_0 is rejected at 1% level of significance
- Ha: $\text{diff} \neq 0$: H_0 is rejected at 1% level of significance
- Ha: $\text{diff} > 0$: H_0 is accepted at 1% level of significance

```
. ttest pun, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	102	12.31765	1.093659	11.04542	10.14812	14.48717
1	52	3.009615	.1831363	1.320615	2.641954	3.377277
combined	154	9.174675	.8083227	10.03102	7.577761	10.77159
diff		9.308032	1.539719		6.266018	12.35005

```
diff = mean(0) - mean(1)                                t = 6.0453
Ho: diff = 0                                             degrees of freedom = 152
```

```
Ha: diff < 0                                Ha: diff != 0                                Ha: diff > 0
Pr(T < t) = 1.0000                        Pr(|T| > |t|) = 0.0000                        Pr(T > t) = 0.0000
```

Here, percentage of undernourished population (*pun*) varies between 6.266 and 12.35

Standard error for the difference in percentage of undernourished population is 1.53

The null hypothesis is given by,

Ho: diff = 0

From the t-test, we find that for,

- Ha: diff < 0: Ho is accepted at 1% level of significance
- Ha: diff != 0: Ho is rejected at 1% level of significance
- Ha: diff > 0: Ho is rejected at 1% level of significance

Two-sample t test with equal variances

```
diff = mean(0) - mean(1)                                t = -3.8285
Ho: diff = 0                                           degrees of freedom = 174

Ha: diff < 0                Ha: diff != 0                Ha: diff > 0
Pr(T < t) = 0.0001          Pr(|T| > |t|) = 0.0002          Pr(T > t) = 0.9999
```

- a) $H_a: \text{diff} < 0$: H_0 is rejected at 1% level of significance
- b) $H_a: \text{diff} \neq 0$: H_0 is rejected at 1% level of significance
- c) $H_a: \text{diff} > 0$: H_0 is accepted at 10% level of significance

```
. ttest gdppc, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	125	9509.119	590.3338	6600.133	8340.683	10677.55
1	63	49745.26	2762.249	21924.68	44223.6	55266.92
combined	188	22992.51	1711.827	23471.39	19615.53	26369.48
diff		-40236.14	2125.673		-44429.67	-36042.61

```
diff = mean(0) - mean(1)                                t = -18.9287
Ho: diff = 0                                             degrees of freedom = 186
```

```
Ha: diff < 0                Ha: diff != 0                Ha: diff > 0
Pr(T < t) = 0.0000          Pr(|T| > |t|) = 0.0000          Pr(T > t) = 1.0000
```

.

Here, GDP per capita (*gdppc*) varies between -44429.67 and -36042.61

Standard error for the difference in GDP per capita is very high (=2125.673)

The null hypothesis is given by,

Ho: diff = 0

From the t-test, we find that for,

- Ha: diff < 0: Ho is rejected at 1% level of significance
- Ha: diff != 0: Ho is rejected at 1% level of significance
- Ha: diff > 0: Ho is accepted at 1% level of significance

Answer 4:

```
. *QUESTION 4:
. *Taking log of GDP per capita
. gen lgdppc = ln(gdppc)
(20 missing values generated)

. *creating a dummy variable from stat score
. gen ss = 1 if stat <= 25
(207 missing values generated)

. replace ss = 2 if stat > 25 & stat <= 50
(29 real changes made)

. replace ss = 3 if stat > 50 & stat <= 75
(70 real changes made)

. replace ss = 4 if stat > 75 & stat <= 100
(42 real changes made)

. replace ss = . if (missing(stat))
(0 real changes made)

.
end of do-file
```

Here, we took log of *gdppc* and stored it in *lgdppc*. Next, we are generating a new dummy variable from stat score named 'ss' which stores the values '1', '2', '3', '4' and '.' under the following conditions given.

Answer 5:

```
. *QUESTION 5:  
. *running a regression on lexp and all other explanatory variables including  
. ///the dummy variable  
> reg lexp co2 pcrate pun hexp lgdppc stat i.ss
```

Source	SS	df	MS	Number of obs	=	74
				F(8, 65)	=	16.68
Model	1720.50292	8	215.062865	Prob > F	=	0.0000
Residual	837.859097	65	12.89014	R-squared	=	0.6725
				Adj R-squared	=	0.6322
Total	2558.36202	73	35.0460551	Root MSE	=	3.5903

lexp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
co2	-.4530016	.2487928	-1.82	0.073	-.9498751	.0438719
pcrate	.0826651	.0374112	2.21	0.031	.0079497	.1573805
pun	.0115544	.074892	0.15	0.878	-.1380152	.161124
hexp	-.0621712	.1938881	-0.32	0.750	-.4493924	.3250501
lgdppc	4.367693	.9802781	4.46	0.000	2.409942	6.325443
stat	.0072442	.0777578	0.09	0.926	-.1480489	.1625374
ss						
3	-1.564096	2.296125	-0.68	0.498	-6.149775	3.021583
4	1.400875	3.568512	0.39	0.696	-5.725936	8.527685
_cons	24.51965	10.02219	2.45	0.017	4.503957	44.53533

Here, we are regressing life expectancy at birth (*lexp*) with *co2*, *pcrate*, *pun*, *hexp*, *lgdppc*, *stat*; which are the explanatory variables in our data along with the newly created dummy variable named '*ss*'. Here, while regressing, we considered '*i.ss*' for the dummy variable. This is because *i.* stands for the intercept dummy of the categorical variable *ss* in our data. All other variables are continuous in nature.

Now, from the regression, we find that,

- *co2*, *pun*, *hexp*, *stat* are statistically insignificant at 5% level of significance
- *pcrate* and *lgdppc* are statistically significant at 5% level of significance
- if *co2* increases by 1 kiloton (kt), *lexp* decreases by 0.453 years
- if *pcrate* increases by 1 unit, *lexp* increases by 0.082 years

- if *pun* increases by 1 member, *lexp* increases by 0.115 years
- if *hexp* increases by 1 percent, *lexp* decreases by 0.062 years
- if *lgdppc* increases by 1 US \$, *lexp* increases by 4.367 years
- if *stat* increases by 1 score, *lexp* increases by 0.0072 years

For the dummy variable *ss*,

- compared to $s = 1$ and $s = 2$, if *stat* score lies between 50 and 75, *lexp* decreases by 1.564 years
- compared to $s = 1$ and $s = 2$, if *stat* score lies between 75 and 100, *lexp* increases by 1.4 years

However, the dummy variable *s* is statistically insignificant at 5% level of significance.

The constant term of the regression is 24.51 which is statistically significant at 5% level of significance.

Answer 6:

Regression Diagnostics:-

```
. predict lexp1
(option xrb assumed; fitted values)
(134 missing values generated)

. predict res, residuals
(134 missing values generated)

. *checking for normality
. *(jarque-bera test)
. jbr res
Jarque-Bera normality test:  1.353 Chi(2)  .5084
Jarque-Bera test for Ho: normality:
```

Jarque-Bera test is a no (no parameter) test where the null hypothesis is presented as,

Ho: normality

The value of Jarque-Bera normality test is 1.353

Here, the value of $\chi^2(2) = 0.5084$ with 2 degrees of freedom

This is statistically insignificant at 5% level of significance and we reject the null hypothesis of normality.

```
. histogram res, normal
[bin=8, start=-10.254249, width=2.1857467]

.
end of do-file

. graph save Graph "F:\Stata MP 14.2\MD PROJECT\histogram 2019.gph"
[file F:\Stata MP 14.2\MD PROJECT\histogram 2019.gph saved]

. do "C:\Users\student\AppData\Local\Temp\STD000000000.tmp"

. rvfplot, yline(0)

.
end of do-file

. graph save Graph "F:\Stata MP 14.2\MD PROJECT\rvfplot 2019.gph"
[file F:\Stata MP 14.2\MD PROJECT\rvfplot 2019.gph saved]
```

We had plotted a histogram where we found, the histogram is very mean-centric (i.e. it is leptokurtic in nature).

We had also plotted a scatterplot of residuals versus fitted values where we found there is no such relation between the two axes. It is not depicting any particular shape (or a pattern).

Thus, there is no presence of heteroscedasticity in our data.

```

. *checking for homoscedasticity
. *# graphical method
. rvfplot, yline(0)

.
. *# formal test
. hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
    Ho: Constant variance
    Variables: fitted values of lexp

    chi2(1)      =      6.50
    Prob > chi2   =     0.0108

```

After performing the graphical method for detecting heteroscedasticity, we then performed the formal method for detecting heteroscedasticity.

Here, we found that for the Breusch-Pagan/ Cook-Weisberg test for heteroscedasticity, the null hypothesis is presented as,

Ho: Constant variance

The $\chi^2(1)$ value is 6.50 which is greater than the critical value ($= 3.84$), thus we reject the null hypothesis of constant variance. The $\chi^2(1)$ value is statistically insignificant at 5% level of significance.

Thus, there is a presence of heteroscedasticity in our data.

```
. reg lexp co2 pcrate pun hexp lgdppc stat i.ss, robust
```

```
Linear regression               Number of obs   =          74
                               F(8, 65)         =         24.83
                               Prob > F          =         0.0000
                               R-squared          =         0.6725
                               Root MSE       =         3.5903
```

lexp	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
co2	-.4530016	.1957965	-2.31	0.024	-.8440343	-.0619689
pcrate	.0826651	.0315637	2.62	0.011	.0196281	.1457021
pun	.0115544	.0899919	0.13	0.898	-.1681718	.1912806
hexp	-.0621712	.2131474	-0.29	0.771	-.4878559	.3635135
lgdppc	4.367693	.9961031	4.38	0.000	2.378337	6.357048
stat	.0072442	.0852536	0.08	0.933	-.163019	.1775075
ss						
3	-1.564096	2.731703	-0.57	0.569	-7.019684	3.891492
4	1.400875	3.853447	0.36	0.717	-6.294991	9.09674
_cons	24.51965	9.840201	2.49	0.015	4.867408	44.17188

Next, we performed regression with robust standard errors. "Robust" standard errors is a technique to obtain unbiased standard errors of OLS coefficients under heteroscedasticity.

Now, some of the coefficients are statistically significant at 5% level of significance.


```
*checking for multicollinearity
vif
```

Variable	VIF	1/VIF
co2	2.29	0.437030
pcrate	1.56	0.640425
pun	2.68	0.373562
hexp	1.17	0.857127
lgdppc	4.39	0.227771
stat	5.48	0.182362
ss		
3	7.57	0.132158
4	18.06	0.055363
Mean VIF	5.40	

We can use vif command after the regression to check for multicollinearity. As a rule of thumb, a variable whose vif value is > 10 may merit further investigation.

Here, the vif value = 5.40 which is less than 10, thus there is no multicollinearity issue in our data and it is a full column rank matrix.

```
. *descriptive statistics
. desc lexp - stat
```

variable name	storage type	display format	value label	variable label
lexp	double	%14.2f		lexp
co2	double	%14.2f		co2
pcrate	double	%14.2f		pcrate
pun	double	%14.2f		pun
hexp	double	%14.2f		hexp
gdppc	double	%14.2f		gdppc
stat	double	%14.2f		stat

```
. sum lexp - stat
```

Variable	Obs	Mean	Std. Dev.	Min	Max
lexp	206	72.96525	7.449954	52.91	85.18049
co2	188	4.040135	4.674155	.0337149	31.8772
pcrate	132	92.64288	13.5823	54.72869	120.4473
pun	160	9.60625	10.62906	2.5	54.8
hexp	180	6.458008	2.94654	2.017115	20.88979
gdppc	188	22992.51	23471.39	760.6041	128031.2
stat	142	64.5227	16.66675	16.66667	96.66667

```
corr lexp - stat
```

```
(obs=74)
```

	lexp	co2	pcrate	pun	hexp	gdppc	stat
lexp	1.0000						
co2	0.5126	1.0000					
pcrate	0.5587	0.4250	1.0000				
pun	-0.5989	-0.5379	-0.4863	1.0000			
hexp	0.0901	0.0810	0.1475	0.1192	1.0000		
gdppc	0.7161	0.7761	0.4275	-0.6208	0.1387	1.0000	
stat	0.5399	0.5004	0.4057	-0.5066	0.1236	0.5763	1.0000

There is high correlation among the variables as the values are greater than or equal to 0.5

```
. *principle component analysis (PCA)
. factor lexp - stat, pcf
(obs=74)
```

```
Factor analysis/correlation      Number of obs   =      74
Method: principal-component factors Retained factors =       2
Rotation: (unrotated)           Number of params =     13
```

Factor	Eigenvalue	Difference	Proportion	Cumulative
Factor1	3.76368	2.69533	0.5377	0.5377
Factor2	1.06835	0.37921	0.1526	0.6903
Factor3	0.68915	0.15493	0.0984	0.7887
Factor4	0.53422	0.10624	0.0763	0.8651
Factor5	0.42798	0.06027	0.0611	0.9262
Factor6	0.36770	0.21877	0.0525	0.9787
Factor7	0.14893	.	0.0213	1.0000

```
LR test: independent vs. saturated:  chi2(21) = 238.08 Prob>chi2 = 0.0000
```

```
Factor loadings (pattern matrix) and unique variances
```

Variable	Factor1	Factor2	Uniqueness
lexp	0.8343	-0.0054	0.3040
co2	0.7987	-0.0157	0.3618
pcrate	0.6822	0.1208	0.5201
pun	-0.7851	0.3296	0.2749
hexp	0.1306	0.9697	0.0426
gdppc	0.8836	0.0229	0.2188
stat	0.7418	0.0627	0.4459

- There are two factors (retained factors = 2) which satisfies the kaiser criterion. The first 2 eigenvalues are 3.76 and 1.06.
- We performed the factor analysis, where we considered rotation (unrotated) to understand the factor loadings.
- We perform LR test: independent vs saturated (correlated to each other)
- The null hypothesis is H_0 : independence (i.e. no correlation) (or, zero correlation)
- The value of $\chi^2(21) = 238.08$ and $\text{Prob} > \chi^2 = 0.0000$
- The null hypothesis, H_0 is rejected at 1% level of significance; thus, running factor analysis is a good idea.

Now,

1. *lexp* has as much as 83.43% commonness with factor 1
2. *co2* has as much as 79.87% commonness with factor 1
3. *pcrate* has as much as 68.22% commonness with factor 1
4. *hexp* has as much as 13.06% commonness with factor 1
5. *gdppc* has as much as 88.36% commonness with factor 1
6. *stat* has as much as 74.18% commonness with factor 1

```
. estat kmo
```

Kaiser-Meyer-Olkin measure of sampling adequacy

Variable	kmo
lexp	0.7967
co2	0.7646
pcrate	0.7933
pun	0.8593
hexp	0.3286
gdppc	0.7294
stat	0.9377
Overall	0.7907

```
. factortest lexp - stat
```

Determinant of the correlation matrix

Det = 0.035

Bartlett test of sphericity

Chi-square = 234.723

Degrees of freedom = 21

p-value = 0.000

H0: variables are not intercorrelated

Kaiser-Meyer-Olkin Measure of Sampling Adequacy

KMO = 0.791

Here, kmo value is 0.7907 which is close to 1, thus we have adequate data to run factor analysis.

For the Bartlett's test of Sphericity,

1. the null hypothesis is H_0 : variables are not intercorrelated
2. the value of chi-square is 234.723 with 21 degrees of freedom and the p-value is 0.0000
3. we reject the null hypothesis that variables are not intercorrelated at 5% level of significance.
4. Therefore, factor analysis is a good idea.

```
. *scree plot
. screeplot

. screeplot, yline(1)

.
end of do-file

. graph save Graph "D:\Stata MP 14.2\MD PROJECT\screeplot 2019.gph"
(file D:\Stata MP 14.2\MD PROJECT\screeplot 2019.gph saved)
```

We had then plotted the screeplot which is a graphical method of detecting and dropping the eigenvalues which are less than 1 since these provide less information than is provided by a single variable.

```
. *rotation
. *# orthogonal rotation
. rotate, varimax
```

```
Factor analysis/correlation      Number of obs   =      74
Method: principal-component factors    Retained factors =      2
Rotation: orthogonal varimax (Kaiser off)  Number of params =     13
```

Factor	Variance	Difference	Proportion	Cumulative
Factor1	3.73652	2.64102	0.5338	0.5338
Factor2	1.09550	.	0.1565	0.6903

```
LR test: independent vs. saturated:  chi2(21) = 238.08 Prob>chi2 = 0.0000
```

Rotated factor loadings (pattern matrix) and unique variances

Variable	Factor1	Factor2	Uniqueness
lexp	0.8306	0.0784	0.3040
co2	0.7963	0.0646	0.3618
pcrate	0.6666	0.1887	0.5201
pun	-0.8142	0.2492	0.2749
hexp	0.0326	0.9779	0.0426
gdppc	0.8768	0.1115	0.2188
stat	0.7317	0.1368	0.4459

Factor rotation matrix

	Factor1	Factor2
Factor1	0.9950	0.1004
Factor2	-0.1004	0.9950

This is an orthogonal rotation.

Here, variance = 3.736 has been extracted by factor 1 and variance = 1.095 has been extracted by factor 2

The cumulative variance is 0.6903

However, all the variables are loaded on both the factors

```
. rotate, varimax blanks(.49)
```

```
Factor analysis/correlation      Number of obs   =      74
Method: principal-component factors    Retained factors =      2
Rotation: orthogonal varimax (Kaiser off)  Number of params =     13
```

Factor	Variance	Difference	Proportion	Cumulative
Factor1	3.73652	2.64102	0.5338	0.5338
Factor2	1.09550	.	0.1565	0.6903

```
LR test: independent vs. saturated:  chi2(21) = 238.08 Prob>chi2 = 0.0000
```

```
Rotated factor loadings (pattern matrix) and unique variances
```

Variable	Factor1	Factor2	Uniqueness
lexp	0.8306		0.3040
co2	0.7963		0.3618
pcrate	0.6666		0.5201
pun	-0.8142		0.2749
hexp		0.9779	0.0426
gdppc	0.8768		0.2188
stat	0.7317		0.4459

```
(blanks represent abs(loading)<.49)
```

```
Factor rotation matrix
```

	Factor1	Factor2
Factor1	0.9950	0.1004
Factor2	-0.1004	0.9950

We had considered 0.49 as the rule of thumb. Thus, values which are less than 0.49 are represented by blanks in the pattern matrix.

Therefore, *lexp* to *pun* and *gdppc* and *stat* are now loaded on factor 1

But, *hexp* is now loaded on factor 2

```
. *saving factor score
. predict pc1 pc2
(regression scoring assumed)
```

Scoring coefficients (method = regression; based on varimax rotated factors)

Variable	Factor1	Factor2
lexp	0.22105	0.01723
co2	0.21262	0.00669
pcrate	0.16898	0.13073
pun	-0.23852	0.28605
hexp	-0.05657	0.90658
gdppc	0.23143	0.04489
stat	0.19020	0.07815

```
. estat common
```

Correlation matrix of the varimax rotated common factors

Factors	Factor1	Factor2
Factor1	1	
Factor2	0	1

```
. corr pc1 pc2
(obs=74)
```

	pc1	pc2
pc1	1.0000	
pc2	0.0000	1.0000

We find that, there is no correlation in orthogonal rotation.


```
. *# oblique rotation
. rotate, promax
```

```
Factor analysis/correlation      Number of obs    =      74
Method: principal-component factors    Retained factors =      2
Rotation: oblique promax (Kaiser off)  Number of params =     13
```

Factor	Variance	Proportion	Rotated factors are correlated
Factor1	3.76003	0.5371	
Factor2	1.14107	0.1630	

```
LR test: independent vs. saturated:  chi2(21) = 238.08 Prob>chi2 = 0.0000
```

Rotated factor loadings (pattern matrix) and unique variances

Variable	Factor1	Factor2	Uniqueness
lexp	0.8306	0.0255	0.3040
co2	0.7970	0.0138	0.3618
pcrate	0.6585	0.1470	0.5201
pun	-0.8355	0.3030	0.2749
hexp	-0.0307	0.9819	0.0426
gdppc	0.8750	0.0558	0.2188
stat	0.7274	0.0906	0.4459

Factor rotation matrix

	Factor1	Factor2
Factor1	0.9993	0.1643
Factor2	-0.0368	0.9864

This is oblique rotation.

Here, variance = 3.76 has been extracted by factor 1 and variance = 1.141 has been extracted by factor 2

There is correlation in oblique rotation as the rotated factors are correlated.

However, all the variables are loaded on both the factors

```
. rotate, promax blanks(.44)
```

```
Factor analysis/correlation      Number of obs   =      74
Method: principal-component factors  Retained factors =      2
Rotation: oblique promax (Kaiser off) Number of params =     13
```

Factor	Variance	Proportion	Rotated factors are correlated
Factor1	3.76003	0.5371	
Factor2	1.14107	0.1630	

```
LR test: independent vs. saturated:  chi2(21) = 238.08 Prob>chi2 = 0.0000
```

```
Rotated factor loadings (pattern matrix) and unique variances
```

Variable	Factor1	Factor2	Uniqueness
lexp	0.8306		0.3040
co2	0.7970		0.3618
pcrate	0.6585		0.5201
pun	-0.8355		0.2749
hexp		0.9819	0.0426
gdppc	0.8750		0.2188
stat	0.7274		0.4459

```
(blanks represent abs(loading)<.44)
```

```
Factor rotation matrix
```

	Factor1	Factor2
Factor1	0.9993	0.1643
Factor2	-0.0368	0.9864

We had considered 0.44 as the rule of thumb. Thus, values which are less than 0.44 are represented by blanks in the pattern matrix.

Therefore, *lexp* to *pun* and *gdppc* and *stat* are now loaded on factor 1

But, *hexp* is now loaded on factor 2

```
. *saving factor score
. predict pc3 pc4
(regression scoring assumed)
```

Scoring coefficients (method = regression; based on promax(3) rotated factors)

Variable	Factor1	Factor2
lexp	0.22170	0.03144
co2	0.21261	0.02037
pcrate	0.17697	0.14134
pun	-0.21981	0.27009
hexp	0.00131	0.90105
gdppc	0.23382	0.05970
stat	0.19479	0.09024

```
. estat common
```

Correlation matrix of the promax(3) rotated common factors

Factors	Factor1	Factor2
Factor1	1	
Factor2	.1279	1

```
. corr pc3 pc4
(obs=74)
```

	pc3	pc4
pc3	1.0000	
pc4	0.1279	1.0000

There is some amount of correlation in oblique rotation, however no presence of multicollinearity in the data.

`. linktest`

Source	SS	df	MS	Number of obs	=	74
Model	1722.0889	2	861.044448	F(2, 71)	=	73.10
Residual	836.273124	71	11.7784947	Prob > F	=	0.0000
				R-squared	=	0.6731
				Adj R-squared	=	0.6639
Total	2558.36202	73	35.0460551	Root MSE	=	3.432

lexp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
_hat	.0641803	2.551769	0.03	0.980	-5.023902	5.152263
_hatsq	.006723	.0183225	0.37	0.715	-.029811	.043257
_cons	32.40201	88.50066	0.37	0.715	-144.0633	208.8673

Here, we have performed model misspecification tests. For the linktest, we get,

- lexp* has been regressed on *hat* and *hatsq*
- here, *hat* is significant as it should be at 1% level of significance and *hatsq* is insignificant at 1% level of significance.

Thus, model is highly misspecified.

`. ovtest`

```
Ramsey RESET test using powers of the fitted values of lexp
Ho: model has no omitted variables
      F(3, 62) =      0.69
      Prob > F =      0.5627
```

For the ovtest, we get,

- the null hypothesis is given by, H_0 : model has no omitted variables
- $F(3,62) = 0.69 < 2.68$ (the critical value of $F(3,62)$ at 5% level of significance); thus we reject the null hypothesis that the model has no omitted variables at 5% level of significance.
- $\text{Prob} > F = 0.5627 > 0.05$; thus it is insignificant at 5% level of significance.
- Thus, the model has a lot of explanatory variables which are omitted variables.

Answer 1:

For 2020:

```
. ***FOR 2020:
.
. *QUESTION 1:
. *describing all the variables used in the data
. desc
```

Contains data

```
obs:      208
vars:      8
size:     17,888
```

variable name	storage type	display format	value label	variable label
CountryName	str30	%30s		Country Name
lexp	double	%14.2f		lexp
co3	double	%14.2f		co3
pcrate	double	%14.2f		pcrate
pun	double	%14.2f		pun
hexp	double	%14.2f		hexp
gdppc	double	%14.2f		gdppc
stat	double	%14.2f		stat

Sorted by:

Note: Dataset has changed since last saved.

```
. *regressing lexp with all other explanatory variables to get their relation
. ///with each other
> reg lexp co3 pcrate pun hexp gdppc stat
```

Source	SS	df	MS	Number of obs	=	60
				F(6, 53)	=	18.95
Model	1247.01478	6	207.835797	Prob > F	=	0.0000
Residual	581.312867	53	10.9681673	R-squared	=	0.6821
				Adj R-squared	=	0.6461
Total	1828.32765	59	30.9886042	Root MSE	=	3.3118

lexp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
co3	-.3202865	.2918929	-1.10	0.277	-.9057496	.2651766
pcrate	.1247831	.0398583	3.13	0.003	.0448375	.2047288
pun	-.1328368	.0717487	-1.85	0.070	-.2767466	.0110729
hexp	.1481946	.194948	0.76	0.451	-.2428216	.5392108
gdppc	.0003295	.0000836	3.94	0.000	.0001618	.0004972
stat	.0342244	.0388422	0.88	0.382	-.0436831	.1121319
_cons	53.60646	4.744244	11.30	0.000	44.09071	63.12221

Here, *lexp1* is the dependent (or explained) variable. *co3*, *pcrate1*, *pun1*, *hexp1*, *gdppc1* and *stat1* are the independent (or explanatory) variables. All the variables mentioned above are continuous in nature. If we regress *lexp1* with all other explanatory variables, then,

- co3* is statistically insignificant at 5% level of significance; which means as *co3* emission increases in each country, *lexp1* decreases by 0.3208 years.
- pcrate1* is also statistically significant at 5% level of significance; as *pcrate1* increases by 1 unit, *lexp1* increases by 0.124 years for each country.
- pun1* is statistically insignificant at 5% level of significance, thus, as *pun1* increases by 1%, *lexp1* decreases by 0.132 years for each country.
- hexp1* is statistically insignificant at 5% level of significance; as *hexp1* increases by 1% of GDP, *lexp1* decreases by 0.148 years.
- gdppc1* is statistically significant at 5% level of significance which indicates that, as *gdppc1* increases by 1 unit, *lexp1* increases by 0.0003 years.
- stat1* is statistically insignificant at 5% level of significance; thus; as *stat1* increases by 1 unit, *lexp1* increases by 0.034 years for each country.
- the constant term (=53.60) is statistically significant at 5% level of significance.

Answer 2:

```
. *QUESTION 2:
. *generating a new variable deve which stores the value 0 if developing and
. /// 1 if developed
> *the development threshold for GDP(PPP) per capita of a developed country is
. ///atleast US $22,000
> *comparing gdppc of our data with the given value, we get
. gen deve = 0 if gdppc <= 22000
(84 missing values generated)

. replace deve = 1 if gdppc > 22000
(84 real changes made)

. replace deve = . if(missing(gdppc))
(21 real changes made, 21 to missing)
```

Here, we generated a new variable named *deve* which took the value ‘0’ if the country is developing and ‘1’ if the country is developed. We also replaced the value of *deve* with ‘.’ when it satisfies neither of the conditions mentioned above.

We had compared the value of *gdppc* with the given value (= 22000). This is the development threshold for GDP(PPP) per capita for a developed country (= US \$22,000).

Answer 3:

```
. *QUESTION 3:  
. *Finding the mean of all the variables except stat score and testing its  
. ///significance for both developing and developed countries  
> mean lexp co3 pcrate pun hexp gdppc if deve == 0
```

Mean estimation Number of obs = 52

	Mean	Std. Err.	[95% Conf. Interval]	
lexp	69.99173	.7648911	68.45615	71.52732
co3	2.061425	.2467412	1.566071	2.556778
pcrate	92.79357	1.94169	88.89547	96.69168
pun	10.20577	1.207201	7.782213	12.62933
hexp	6.289619	.3329459	5.621202	6.958035
gdppc	10138.76	877.9201	8376.259	11901.26

```
. mean lexp co3 pcrate pun hexp gdppc if deve == 1
```

Mean estimation Number of obs = 40

	Mean	Std. Err.	[95% Conf. Interval]	
lexp	79.11657	.5190952	78.0666	80.16654
co3	7.09856	.7912014	5.498204	8.698916
pcrate	98.7742	.9477576	96.85718	100.6912
pun	2.98	.2237845	2.527353	3.432647
hexp	8.688491	.4654086	7.747113	9.629869
gdppc	46002.95	3122.254	39687.6	52318.31

Here, we are finding the mean of all the variables except stat score for both developing and developed countries.


```
. *executing t-test among two group of countries
. ttest lexp, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	124	68.50131	.5489023	6.112317	67.41479	69.58783
1	63	79.10329	.4643617	3.685756	78.17505	80.03154
combined	187	72.0731	.539755	7.381038	71.00827	73.13793
diff		-10.60198	.8387973		-12.25682	-8.947144

```
diff = mean(0) - mean(1)                                t = -12.6395
Ho: diff = 0                                           degrees of freedom = 185

Ha: diff < 0                                Ha: diff != 0                                Ha: diff > 0
Pr(T < t) = 0.0000                Pr(|T| > |t|) = 0.0000                Pr(T > t) = 1.0000
```

Here, Life expectancy at birth (*lexp*) varies between -12.25 and -8.94

Standard error for the difference in life expectancy at birth is very less (= 0.838)

The null hypothesis is given by,

Ho: diff = 0

From the t-test, we find that for,

- Ha: diff < 0: Ho is rejected at 1% level of significance
- Ha: diff != 0: Ho is rejected at 1% level of significance
- Ha: diff > 0: Ho is accepted at 1% level of significance

```
. ttest co3, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	122	1.807898	.1608685	1.77685	1.489416	2.126379
1	57	8.187992	.7786513	5.878689	6.628165	9.747818
combined	179	3.839548	.349803	4.680045	3.149253	4.529842
diff		-6.380094	.5805183		-7.525722	-5.234466

```
diff = mean(0) - mean(1)                                t = -10.9903
Ho: diff = 0                                           degrees of freedom = 177

Ha: diff < 0                Ha: diff != 0                Ha: diff > 0
Pr(T < t) = 0.0000          Pr(|T| > |t|) = 0.0000          Pr(T > t) = 1.0000
```

Here, CO3 emission (*co3*) varies between -7.525 and -5.234

Standard error for the difference in CO3 emission is very less (= 0.58)

The null hypothesis is given by,

Ho: diff = 0

From the t-test, we find that for,

- Ha: diff < 0: Ho is rejected at 1% level of significance
- Ha: diff != 0: Ho is rejected at 1% level of significance
- Ha: diff > 0: Ho is accepted at 1% level of significance

```
. ttest pcrate, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	64	90.85485	1.956955	15.65564	86.94419	94.76551
1	46	99.10757	.848447	5.754448	97.39871	100.8164
combined	110	94.30599	1.25066	13.11703	91.82722	96.78476
diff		-8.252725	2.420257		-13.05009	-3.455356

diff = mean(0) - mean(1)

t = -3.4099

Ho: diff = 0

degrees of freedom = 108

Ha: diff < 0

Ha: diff != 0

Ha: diff > 0

Pr(T < t) = 0.0005

Pr(|T| > |t|) = 0.0009

Pr(T > t) = 0.9995

Here, primary completion rate (*pcrate*) varies between -13.05 and -3.455

Standard error for the difference in primary completion rate is 2.42

The null hypothesis is given by,

Ho: diff = 0

From the t-test, we find that for,

- Ha: diff < 0: Ho is rejected at 1% level of significance
- Ha: diff != 0: Ho is rejected at 1% level of significance
- Ha: diff > 0: Ho is accepted at 1% level of significance

```
. ttest hexp, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	119	6.370955	.2724702	2.972299	5.83139	6.910521
1	56	8.610337	.3858615	2.887523	7.837053	9.383621
combined	175	7.087557	.235727	3.118375	6.622305	7.55281
diff		-2.239382	.4773391		-3.18154	-1.297223

diff = mean(0) - mean(1)

t = -4.6914

Ho: diff = 0

degrees of freedom = 173

Ha: diff < 0

Ha: diff != 0

Ha: diff > 0

Pr(T < t) = 0.0000

Pr(|T| > |t|) = 0.0000

Pr(T > t) = 1.0000

Here, health expenditure (% of GDP) (*hexp*) varies between -3.181 and -1.297

Standard error for the difference in health expenditure (% of GDP) is very less (=0.477)

The null hypothesis is given by,

Ho: diff = 0

From the t-test, we find that for,

- Ha: diff < 0: Ho is rejected at 1% level of significance
- Ha: diff != 0: Ho is rejected at 1% level of significance
- Ha: diff > 0: Ho is accepted at 10% level of significance

```
. ttest pun, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	100	13.061	1.132858	11.32858	10.81316	15.30884
1	53	3.035849	.1945762	1.416536	2.645403	3.426295
combined	153	9.588235	.8369678	10.35272	7.934643	11.24183
diff		10.02515	1.564907		6.93321	13.11709

```
diff = mean(0) - mean(1)                                t = 6.4062
Ho: diff = 0                                             degrees of freedom = 151

Ha: diff < 0                Ha: diff != 0                Ha: diff > 0
Pr(T < t) = 1.0000          Pr(|T| > |t|) = 0.0000          Pr(T > t) = 0.0000
```

Here, percentage of undernourished population (*pun*) varies between 6.933 and 13.11

Standard error for the difference in percentage of undernourished population is 1.564

The null hypothesis is given by,

Ho: diff = 0

From the t-test, we find that for,

- Ha: diff < 0: Ho is accepted at 1% level of significance
- Ha: diff != 0: Ho is rejected at 1% level of significance
- Ha: diff > 0: Ho is rejected at 1% level of significance

```
. ttest gdppc, by(deve)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	124	8843.649	530.7116	5909.755	7793.138	9894.16
1	63	46908.57	2505.692	19888.31	41899.76	51917.38
combined	187	21667.66	1602.762	21917.44	18505.73	24829.59
diff		-38064.92	1931.069		-41874.67	-34255.17

```
diff = mean(0) - mean(1)                                t = -19.7118
Ho: diff = 0                                           degrees of freedom = 185

Ha: diff < 0                Ha: diff != 0                Ha: diff > 0
Pr(T < t) = 0.0000          Pr(|T| > |t|) = 0.0000          Pr(T > t) = 1.0000
```

Here, GDP per capita (*gdppc*) varies between -41874.67 and -34255.17

Standard error for the difference in GDP per capita is very high (=1931.069)

The null hypothesis is given by,

Ho: diff = 0

From the t-test, we find that for,

- Ha: diff < 0: Ho is rejected at 1% level of significance
- Ha: diff != 0: Ho is rejected at 1% level of significance
- Ha: diff > 0: Ho is accepted at 1% level of significance

Answer 4:

```
. *QUESTION 4:
. *Taking log of GDP per capita
. gen lgdppc = ln(gdppc)
(21 missing values generated)

. *creating a dummy variable from stat score
. gen ss = 1 if stat <= 25
(207 missing values generated)

. replace ss = 2 if stat > 25 & stat <= 50
(33 real changes made)

. replace ss = 3 if stat > 50 & stat <= 75
(70 real changes made)

. replace ss = 4 if stat > 75 & stat <= 100
(38 real changes made)

. replace ss = . if (missing(stat))
(0 real changes made)
```

Here, we took log of *gdppc* and stored it in *lgdppc*. Next, we are generating a new dummy variable from stat score named 'ss' which stores the values '1', '2', '3', '4' and '.' under the following conditions given.

Answer 5:

```
. *QUESTION 5:  
. *running a regression on lexp and all other explanatory variables including  
. ///the dummy variable  
> reg lexp co3 pcrate pun hexp lgdppc stat i.ss
```

Source	SS	df	MS	Number of obs	=	60
Model	1327.55074	8	165.943843	F(8, 51)	=	16.90
Residual	500.776904	51	9.81915498	Prob > F	=	0.0000
				R-squared	=	0.7261
				Adj R-squared	=	0.6831
Total	1828.32765	59	30.9886042	Root MSE	=	3.1336

lexp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
co3	-.3639859	.2703337	-1.35	0.184	-.9067036 .1787317
pcrate	.0627868	.0395569	1.59	0.119	-.0166271 .1422007
pun	-.0655449	.0744067	-0.88	0.383	-.2149226 .0838328
hexp	.2051085	.1828109	1.12	0.267	-.1618997 .5721168
lgdppc	4.121248	.9174177	4.49	0.000	2.279455 5.963041
stat	.0064151	.0785306	0.08	0.935	-.1512417 .1640719
ss					
3	-2.546789	2.1	-1.21	0.231	-6.762716 1.669137
4	-.6976435	3.460376	-0.20	0.841	-7.644638 6.249351
_cons	28.46513	8.894389	3.20	0.002	10.6089 46.32137

Here, we are regressing life expectancy at birth (*lexp*) with *co3*, *pcrate*, *pun*, *hexp*, *lgdppc*, *stat*; which are the explanatory variables in our data along with the newly created dummy variable named 'ss'. Here, while regressing, we considered '*i.ss*' for the dummy variable. This is because *i.* stands for the intercept dummy of the categorical variable *ss* in our data. All other variables are continuous in nature.

Now, from the regression, we find that,

- *co3*, *pcrate*, *pun*, *hexp*, *stat* are statistically insignificant at 5% level of significance
- *lgdppc* is statistically significant at 5% level of significance
- if *co3* increases by 1 kiloton (kt), *lexp* decreases by 0.363 years
- if *pcrate* increases by 1 unit, *lexp* increases by 0.062 years

- if *pun* increases by 1 member, *lexp* decreases by 0.065 years
- if *hexp* increases by 1 percent, *lexp* increases by 0.205 years
- if *lgdppc* increases by 1 US \$, *lexp* increases by 4.12 years
- if *stat* increases by 1 score, *lexp* increases by 0.0064 years

For the dummy variable *ss*,

- compared to $s = 1$ and $s = 2$, if *stat* score lies between 50 and 75, *lexp* decreases by 2.54 years
- compared to $s = 1$ and $s = 2$, if *stat* score lies between 75 and 100, *lexp* decreases by 0.69 years

However, the dummy variable *s* is statistically insignificant at 5% level of significance.

The constant term of the regression is 28.46 which is statistically significant at 5% level of significance.

Answer 6:

Regression Diagnostics:

```
. predict lexp1
(option xb assumed; fitted values)
(148 missing values generated)

. predict res, residuals
(148 missing values generated)

.
end of do-file

. do "C:\Users\DILIP\AppData\Local\Temp\STD000000000.tmp"

. *checking for normality
. *(jarque-bera test)
. jbr res
Jarque-Bera normality test:  5.176 Chi(2)  .0752
Jarque-Bera test for Ho: normality:

. histogram res, normal
(bin=7, start=-8.1841908, width=1.9835954)

.
end of do-file

. graph save Graph "D:\Stata MP 14.2\MD PROJECT\histogram 2020.gph"
(file D:\Stata MP 14.2\MD PROJECT\histogram 2020.gph saved)
```

Jarque-Bera test is a no (no parameter) test where the null hypothesis is presented as,

Ho: normality

The value of Jarque-Bera normality test is 5.176

Here, the value of $\text{Chi}(2) = 0.0752$ with 2 degrees of freedom

This is statistically insignificant at 5% level of significance and we reject the null hypothesis of normality.

```
. *checking for homoscedasticity
. *# graphical method
. rvfplot, yline(0)

.
end of do-file

. graph save Graph "D:\Stata MP 14.2\MD PROJECT\rvfplot 2020.gph"
(file D:\Stata MP 14.2\MD PROJECT\rvfplot 2020.gph saved)
```

We had plotted a histogram where we found, the histogram is very mean-centric (i.e. it is leptokurtic in nature).

We had also plotted a scatterplot of residuals versus fitted values where we found there is no such relation between the two axes. It is not depicting any particular shape (or a pattern).

Thus, there is no presence of heteroscedasticity in our data.

```

. *# formal test
. hettest

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
    Ho: Constant variance
    Variables: fitted values of lexp

    chi2(1)      =      0.62
    Prob > chi2   =      0.4293

. reg lexp co3 pcrate pun hexp lgdpcc stat i.ss, robust

Linear regression                               Number of obs   =           60
                                                F(8, 51)         =          30.07
                                                Prob > F          =          0.0000
                                                R-squared        =          0.7261
                                                Root MSE        =          3.1336

```

lexp	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
co3	-.3639859	.1892717	-1.92	0.060	-.7439648	.0159929
pcrate	.0627868	.0363616	1.73	0.090	-.0102121	.1357857
pun	-.0655449	.0769133	-0.85	0.398	-.2199549	.0888651
hexp	.2051085	.1695959	1.21	0.232	-.1353694	.5455865
lgdpcc	4.121248	.7351396	5.61	0.000	2.645393	5.597102
stat	.0064151	.0943056	0.07	0.946	-.1829113	.1957415
ss						
3	-2.546789	2.266373	-1.12	0.266	-7.096723	2.003144
4	-.6976435	3.832307	-0.18	0.856	-8.39132	6.996033
_cons	28.46513	7.376782	3.86	0.000	13.65563	43.27464

After performing the graphical method for detecting heteroscedasticity, we then performed the formal method for detecting heteroscedasticity.

Here, we found that for the Breusch-Pagan/ Cook-Weisberg test for heteroscedasticity, the null hypothesis is presented as,

Ho: Constant variance

The chi2(1) value is 0.62 which is less than the critical value (= 3.84), thus we accept the null hypothesis of constant variance. The chi2(1) value is statistically insignificant at 5% level of significance.

Thus, there is no presence of heteroscedasticity in our data.

```
. *checking for multicollinearity
. vif
```

Variable	VIF	1/VIF
co3	2.15	0.464895
pcrate	1.67	0.597545
pun	2.39	0.417640
hexp	1.17	0.851671
lgdppc	4.02	0.248666
stat	7.21	0.138754
ss		
3	6.67	0.149937
4	18.11	0.055221
Mean VIF	5.43	

We can use vif command after the regression to check for multicollinearity. As a rule of thumb, a variable whose vif value is > 10 may merit further investigation.

Here, the vif value = 5.43 which is less than 10, thus there is no multicollinearity issue in our data and it is a full column rank matrix

```
. *descriptive statistics
. desc lexp - stat
```

variable name	storage type	display format	value label	variable label
lexp	double	%14.2f		lexp
co3	double	%14.2f		co3
pcrate	double	%14.2f		pcrate
pun	double	%14.2f		pun
hexp	double	%14.2f		hexp
gdppc	double	%14.2f		gdppc
stat	double	%14.2f		stat

```
. sum lexp - stat
```

Variable	Obs	Mean	Std. Dev.	Min	Max
lexp	206	72.34202	7.407253	52.777	85.49756
co3	188	3.775203	4.620423	.0325848	31.72684
pcrate	114	94.53861	12.96707	51.19454	115.6307
pun	160	9.935625	10.81931	2.5	53.1
hexp	180	7.064762	3.123236	2.007863	21.53917
gdppc	187	21667.66	21917.44	751.2009	120010.2
stat	142	63.10251	16.4038	22.2222	94.44447

```
. corr lexp - stat
(obs=60)
```

	lexp	co3	pcrate	pun	hexp	gdppc	stat
lexp	1.0000						
co3	0.5105	1.0000					
pcrate	0.5966	0.4166	1.0000				
pun	-0.6510	-0.5468	-0.5315	1.0000			
hexp	0.3070	0.0487	0.2383	-0.1197	1.0000		
gdppc	0.7326	0.7035	0.4039	-0.6133	0.2625	1.0000	
stat	0.5020	0.4686	0.2505	-0.4473	0.2247	0.5798	1.0000

There is high correlation among the variables as the values are greater than or equal to 0.5

```
. *principle component analysis (PCA)
. factor lexp - stat, pcf
(obs=60)
```

```
Factor analysis/correlation          Number of obs   =          60
Method: principal-component factors   Retained factors =           2
Rotation: (unrotated)                Number of params =          13
```

Factor	Eigenvalue	Difference	Proportion	Cumulative
Factor1	3.77218	2.76884	0.5389	0.5389
Factor2	1.00334	0.19348	0.1433	0.6822
Factor3	0.80986	0.32320	0.1157	0.7979
Factor4	0.48665	0.07225	0.0695	0.8674
Factor5	0.41441	0.06945	0.0592	0.9266
Factor6	0.34496	0.17636	0.0493	0.9759
Factor7	0.16860	.	0.0241	1.0000

```
LR test: independent vs. saturated:  chi2(21) = 189.00 Prob>chi2 = 0.0000
```

```
Factor loadings (pattern matrix) and unique variances
```

Variable	Factor1	Factor2	Uniqueness
lexp	0.8657	0.1027	0.2400
co3	0.7620	-0.3576	0.2916
pcrate	0.6728	0.1375	0.5285
pun	-0.8010	0.1736	0.3282
hexp	0.3399	0.9001	0.0743
gdppc	0.8730	-0.0743	0.2324
stat	0.6857	-0.0159	0.5295

- g. There are two factors (retained factors = 2) which satisfies the kaiser criterion. The first 2 eigenvalues are 3.77 and 1.00.
- h. We performed the factor analysis, where we considered rotation (unrotated) to understand the factor loadings.
- i. We perform LR test: independent vs saturated (correlated to each other)
- j. The null hypothesis is H_0 : independence (i.e. no correlation) (or, zero correlation)
- k. The value of $\chi^2(21) = 189.00$ and $\text{Prob}>\chi^2 = 0.0000$
- l. The null hypothesis, H_0 is rejected at 1% level of significance; thus, running factor analysis is a good idea.

Now,

7. *lexp* has as much as 86.57% commonness with factor 1
8. *co2* has as much as 76.20% commonness with factor 1
9. *pcrate* has as much as 67.28% commonness with factor 1
10. *hexp* has as much as 33.99% commonness with factor 1
11. *gdppc* has as much as 87.30% commonness with factor 1
12. *stat* has as much as 68.57% commonness with factor 1

```
. estat kmo
```

Kaiser-Meyer-Olkin measure of sampling adequacy

Variable	kmo
lexp	0.7967
co3	0.7738
pcrate	0.7572
pun	0.8958
hexp	0.6630
gdppc	0.7593
stat	0.9112
Overall	0.8020

```
. factortest lexp - stat
```

Determinant of the correlation matrix

Det = 0.036

Bartlett test of sphericity

Chi-square = 185.677

Degrees of freedom = 21

p-value = 0.000

H0: variables are not intercorrelated

Kaiser-Meyer-Olkin Measure of Sampling Adequacy

KMO = 0.802

Here, kmo value is 0.8020 which is close to 1, thus we have adequate data to run factor analysis.

For the Bartlett's test of Sphericity,

1. the null hypothesis is H_0 : variables are not intercorrelated
2. the value of chi-square is 185.677 with 21 degrees of freedom and the p-value is 0.0000
3. we reject the null hypothesis that variables are not intercorrelated at 5% level of significance.
4. Therefore, factor analysis is a good idea.

```
. *scree plot
. screeplot

. screeplot, yline(1)

.
end of do-file

. graph save Graph "D:\Stata MP 14.2\MD PROJECT\screeplot 2020.gph"
(file D:\Stata MP 14.2\MD PROJECT\screeplot 2020.gph saved)
```

We had then plotted the screeplot which is a graphical method of detecting and dropping the eigenvalues which are less than 1 since these provide less information than is provided by a single variable.

```
. *rotation
. *# orthogonal rotation
. rotate, varimax
```

```
Factor analysis/correlation      Number of obs   =      60
Method: principal-component factors    Retained factors =       2
Rotation: orthogonal varimax (Kaiser off)  Number of params =     13
```

Factor	Variance	Difference	Proportion	Cumulative
Factor1	3.54968	2.32384	0.5071	0.5071
Factor2	1.22584	.	0.1751	0.6822

```
LR test: independent vs. saturated:  chi2(21) = 189.00 Prob>chi2 = 0.0000
```

Rotated factor loadings (pattern matrix) and unique variances

Variable	Factor1	Factor2	Uniqueness
lexp	0.8011	0.3439	0.2400
co3	0.8321	-0.1269	0.2916
pcrate	0.6062	0.3226	0.5285
pun	-0.8174	-0.0606	0.3282
hexp	0.0708	0.9595	0.0743
gdppc	0.8582	0.1762	0.2324
stat	0.6621	0.1791	0.5295

Factor rotation matrix

	Factor1	Factor2
Factor1	0.9590	0.2835
Factor2	-0.2835	0.9590

This is an orthogonal rotation.

Here, variance = 3.549 has been extracted by factor 1 and variance = 1.225 has been extracted by factor 2

The cumulative variance is 0.6822

However, all the variables are loaded on both the factors

```
. rotate, varimax blanks(.49)
```

```
Factor analysis/correlation      Number of obs   =      60
Method: principal-component factors    Retained factors =      2
Rotation: orthogonal varimax (Kaiser off)  Number of params =     13
```

Factor	Variance	Difference	Proportion	Cumulative
Factor1	3.54968	2.32384	0.5071	0.5071
Factor2	1.22584	.	0.1751	0.6822

```
LR test: independent vs. saturated:  chi2(21) = 189.00 Prob>chi2 = 0.0000
```

```
Rotated factor loadings (pattern matrix) and unique variances
```

Variable	Factor1	Factor2	Uniqueness
lexp	0.8011		0.2400
co3	0.8321		0.2916
pcrate	0.6062		0.5285
pun	-0.8174		0.3282
hexp		0.9595	0.0743
gdppc	0.8582		0.2324
stat	0.6621		0.5295

```
(blanks represent abs(loading)<.49)
```

```
Factor rotation matrix
```

	Factor1	Factor2
Factor1	0.9590	0.2835
Factor2	-0.2835	0.9590

We had considered 0.49 as the rule of thumb. Thus, values which are less than 0.49 are represented by blanks in the pattern matrix.

Therefore, *lexp* to *pun* and *gdppc* and *stat* are now loaded on factor 1

But, *hexp* is now loaded on factor 2

```
. *saving factor score
. predict pc1 pc2
(regression scoring assumed)
```

Scoring coefficients (method = regression; based on varimax rotated factors)

Variable	Factor1	Factor2
lexp	0.19108	0.16319
co3	0.29474	-0.28450
pcrate	0.13219	0.18198
pun	-0.25269	0.10572
hexp	-0.16789	0.88582
gdppc	0.24293	-0.00543
stat	0.17883	0.03629

```
. estat common
```

Correlation matrix of the varimax rotated common factors

Factors	Factor1	Factor2
Factor1	1	
Factor2	0	1

```
. corr pc1 pc2
(obs=60)
```

	pc1	pc2
pc1	1.0000	
pc2	-0.0000	1.0000

We find that, there is no correlation in orthogonal rotation.

```
. *# oblique rotation
. rotate, promax
```

```
Factor analysis/correlation          Number of obs    =      60
Method: principal-component factors   Retained factors =      2
Rotation: oblique promax (Kaiser off) Number of params =     13
```

Factor	Variance	Proportion	Rotated factors are correlated
Factor1	3.73232	0.5332	
Factor2	1.47741	0.2111	

```
LR test: independent vs. saturated:  chi2(21) =  189.00 Prob>chi2 = 0.0000
```

```
Rotated factor loadings (pattern matrix) and unique variances
```

Variable	Factor1	Factor2	Uniqueness
lexp	0.7820	0.2159	0.2400
co3	0.8827	-0.2764	0.2916
pcrate	0.5827	0.2278	0.5285
pun	-0.8402	0.0799	0.3282
hexp	-0.0661	0.9800	0.0743
gdppc	0.8658	0.0325	0.2324
stat	0.6617	0.0697	0.5295

```
Factor rotation matrix
```

	Factor1	Factor2
Factor1	0.9928	0.4138
Factor2	-0.1200	0.9104

This is oblique rotation.

Here, variance = 3.732 has been extracted by factor 1 and variance = 1.477 has been extracted by factor 2

There is correlation in oblique rotation as the rotated factors are correlated.

However, all the variables are loaded on both the factors

```
. rotate, promax blanks(.44)
```

```
Factor analysis/correlation      Number of obs   =      60
Method: principal-component factors  Retained factors =      2
Rotation: oblique promax (Kaiser off) Number of params =     13
```

Factor	Variance	Proportion	Rotated factors are correlated
Factor1	3.73232	0.5332	
Factor2	1.47741	0.2111	

```
LR test: independent vs. saturated:  chi2(21) = 189.00 Prob>chi2 = 0.0000
```

```
Rotated factor loadings (pattern matrix) and unique variances
```

Variable	Factor1	Factor2	Uniqueness
lexp	0.7820		0.2400
co3	0.8827		0.2916
pcrate	0.5827		0.5285
pun	-0.8402		0.3282
hexp		0.9800	0.0743
gdppc	0.8658		0.2324
stat	0.6617		0.5295

```
(blanks represent abs(loading)< .44)
```

```
Factor rotation matrix
```

	Factor1	Factor2
Factor1	0.9928	0.4138
Factor2	-0.1200	0.9104

We had considered 0.44 as the rule of thumb. Thus, values which are less than 0.44 are represented by blanks in the pattern matrix.

Therefore, *lexp* to *pun* and *gdppc* and *stat* are now loaded on factor 1

But, *hexp* is now loaded on factor 2

```

. *saving factor score
. predict pc3 pc4
(regression scoring assumed)

```

Scoring coefficients (method = regression; based on promax(3) rotated factors)

Variable	Factor1	Factor2
lexp	0.21556	0.18812
co3	0.24330	-0.24086
pcrate	0.16062	0.19856
pun	-0.23158	0.06964
hexp	-0.01818	0.85396
gdppc	0.23864	0.02832
stat	0.18238	0.06075

```

. estat common

```

Correlation matrix of the promax(3) rotated common factors

Factors	Factor1	Factor2
Factor1	1	
Factor2	.3016	1

```

. corr pc3 pc4
(obs=60)

```

	pc3	pc4
pc3	1.0000	
pc4	0.3016	1.0000

There is some amount of correlation in oblique rotation, however no presence of multicollinearity in the data.

```
. *model specification tests
. linktest
```

Source	SS	df	MS	Number of obs	=	60
Model	1247.12113	2	623.560566	F(2, 57)	=	61.15
Residual	581.206515	57	10.1966055	Prob > F	=	0.0000
				R-squared	=	0.6821
				Adj R-squared	=	0.6710
Total	1828.32765	59	30.9886042	Root MSE	=	3.1932

lexp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
_hat	1.262007	2.567952	0.49	0.625	-3.880227	6.404242
_hatsq	-.0018839	.0184524	-0.10	0.919	-.0388342	.0350665
_cons	-9.068041	89.05248	-0.10	0.919	-187.3926	169.2565

Here, we have performed model misspecification tests. For the linktest, we get,

- lexp* has been regressed on *hat* and *hatsq*
- here, *hat* is significant as it should be at 1% level of significance and *hatsq* is insignificant at 1% level of significance.

Thus, model is highly misspecified.

```
. ovtest
```

```
Ramsey RESET test using powers of the fitted values of lexp
Ho: model has no omitted variables
      F(3, 50) =      0.53
      Prob > F =      0.6610
```

For the ovtest, we get,

- the null hypothesis is given by, H_0 : model has no omitted variables
- $F(3,62) = 0.6610 < 2.68$ (the critical value of $F(3,62)$ at 5% level of significance); thus we reject the null hypothesis that the model has no omitted variables at 5% level of significance.
- $\text{Prob} > F = 0.53 > 0.05$; thus it is insignificant at 5% level of significance.
- Thus, the model has a lot of explanatory variables which are omitted variables.

Answer 7:

If we compare the data given for the years 2019 and 2020, we see that,

- *pcrate* and *gdppc* are statistically significant at 5% level of significance for both the years.
- *lexp*, *co2*, *pun*, *hexp*, *stat* are statistically insignificant at 5% level of significance for 2019.
- *lexp*, *co3*, *pun*, *hexp*, *stat* are statistically insignificant at 5% level of significance for 2020.
- There is significant presence of heteroscedasticity in the data for 2019, but no presence of heteroscedasticity in the data for 2020.
- There is no presence of multicollinearity in the data for both the years.
- There is significant presence of omitted variable bias in the data for both the years.