Calibrated Fairness in Bandits

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ABSTRACT

We study fairness within the stochastic, *multi-armed bandit* (MAB) decision making framework. We adapt the fairness framework of "treating similar individuals similarly" [5] to this setting. Here, an 'individual' corresponds to an arm and two arms are 'similar' if they have a similar quality distribution. First, we adopt a *smoothness constraint* that if two arms have a similar quality distribution then the probability of selecting each arm should be similar. In addition, we define the *fairness regret*, which corresponds to the degree to which an algorithm is not calibrated, where perfect calibration requires that the probability of selecting an arm is equal to the probability with which the arm has the best quality realization. We show that a variation on Thompson sampling satisfies smooth fairness for total variation distance, and give an $\tilde{O}((kT)^{2/3})$ bound on fairness regret. This complements prior work [12], which protects an on-average better arm from being less favored. We also explain how to extend our algorithm to the dueling bandit setting.

1 INTRODUCTION

Consider a sequential decision making problem where, at each time-step, a decision maker needs to select one candidate to hire from a set of k groups (these may be a different ethnic groups, culture, and so forth), whose true qualities are unknown *a priori.* The decision maker would like to make fair decisions with respect to each group's underlying quality distribution and to learn such a rule through interactions. This naturally leads to a stochastic *multi-armed bandit* framework, where each arm corresponds to a group, and quality corresponds to reward.

Earlier studies of fairness in bandit problems have emphasized the need, over all rounds t, and for any pair of arms, to weakly favor an arm that is weakly better in expectation [12]. This notion of meritocratic fairness has provided interesting results, for example a separation between the dependence on the number of arms in the

regret bound between fair and standard non-fair learning. But this is a somewhat weak requirement in that it (i) it allows a group that is slightly better than all other groups to be selected all the time and even if any single sample from the group may be worse than any single sample from another group, and (ii) it allows a random choice to be made even in the case when one group is much better than another group.[1]

In this work, we adopt the framework of "treating similar individuals similarly" of Dwork et al. [5]. In the current context, it is arms that are the objects about which decisions are made, and thus the 'individual' in Dwork et al. corresponds to an 'arm'. We study the classic stochastic bandit problem, and insist that over all rounds t, and for any pair of arms, that if the two arms have a similar quality distribution then the probability with each arm is selected should be similar. This *smooth fairness* requirement addresses concern (i), in that if one group is best in expectation by only a small margin, but has a similar distribution of rewards to other groups, then it cannot be selected all the time.

By itself we don't consider smooth fairness to be enough because it does not also provide a notion of meritocratic fairness— it does not constrain a decision maker in the case that one group is much stronger than another (in particular, a decision maker could choose the weaker group). For this reason, we also care about *calibrated fairness* and introduce the concept of *fairness regret*, which corresponds to the degree to which an algorithm is not calibrated. Perfect calibration requires that the probability of selecting a group is equal to the probability that a group has the best quality realization. Informally, this is a strengthening of "treating similar individuals similarly" because it further requires that dissimilar individuals be treated dissimilarly (and in the right direction.) In the motivating setting of making decisions about who to hire, groups correspond to divisions within society and each activation of an

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[1]Joseph et al. [11] also extend the results to contextual bandits and infinite bandits. Here, there is additional context associated with an arm in a given time period, this context providing information about a specific individual. Weak meritocratic fairness requires, for any pair of arms, to weakly favor an arm that is weakly better in expectation conditioned on context. When this context removes all uncertainty about quality, then this extension addresses critique (i). But in the more general case we think it remains interesting for future work to generalize our definitions to the case of contextual bandits.

arm to a particular candidate. An algorithm with low fairness regret will give individuals a chance proportionally to their probability of being the best candidate rather than protect an entire group based on a higher average quality.

1.1 Our Results

In regard to smooth fairness, we say that a bandit algorithm is $(\epsilon_1, \epsilon_2, \delta)$-*fair* with respect to a divergence function D (for $\epsilon_1, \epsilon_2 \geq 0$, and $0 \leq \delta \leq 1$) if, with probability $1 - \delta$, in every round t and for every pair of arms i and j,

$$D(\pi_t(i)||\pi_t(j)) \leq \epsilon_1 D(r_i||r_j) + \epsilon_2,$$

where $D(\pi_t(i)||\pi_t(j))$ denotes the divergence between the Bernoulli distributions corresponding to activating arms i and j, and $D(r_i||r_j)$ denotes the divergence between the reward distributions of arms i and j.

The *fairness regret* $R_{f,T}$ of a bandit algorithm over T rounds is the total deviation from calibrated fairness:

$$R_{f,T} = \sum_{t=1}^{T} \mathbb{E}\left[\sum_{i=1}^{k} \max(\mathbb{P}^*(i) - \pi_t(i), 0)\right] \quad (1)$$

where $\mathbb{P}^*(i)$ is the probability that the realized quality of arm i is highest and $\pi_t(i)$ is the probability that arm i is activated by the algorithm in round t.

Our main result is stated for the case of Bernoulli bandits. We show that a Thompson-sampling based algorithm, modified to include an initial uniform exploration phase, satisfies:

(1) $(2, \epsilon_2, \delta)$-fair with regard to total variation distance for any $\epsilon_2 > 0, \delta > 0$, where the amount of initial exploration on each arm scales as $1/\epsilon_2^2$ and $\log(1/\delta)$, and

(2) fairness regret that is bounded by $\tilde{O}((kT)^{2/3})$, where k is the number of arms and T the number of rounds.

We also show that a simpler version of Thompson sampling can immediately satisfy a *subjective version* of smooth fairness. Here, the relevant reward distributions are defined with respect to the posterior reward distribution under the belief of a Bayesian decision maker, this decision maker having an initially uninformed prior. In addition, we draw a connection between calibrated fairness and proper scoring functions: there exists a loss function on reward whose maximization in expectation would result in a calibrated-fair policy. In Section 5 we also extend our results to the *dueling bandit* setting in which the decision maker receives only pairwise comparisons between arms.

1.2 Related work

Joseph et al. [12] were the first to introduce fairness concepts in the bandits setting. These authors adopt the notion of weak meritocratic fairness, and study it within the classic and contextual bandit setting. Their main results establish a separation between the regret for a fair and an un-fair learning algorithm, and an asymptotically regret-optimal, fair algorithm that uses an approach of chained confidence intervals. While their definition promotes meritocracy in regard to expected quality, this present paper emphasizes instead the distribution on rewards, and in this way connects with the smoothness definitions and "similar people be treated similarly" of Dwork et al. [5].

Joseph et al. [11] study a more general problem in which there is no group structure; rather, a number of individuals are available to select in each period, each with individual context (they also consider an infinite bandits setting.) Jabbari et al. [10] also extend the notion of weakly meritocratic fairness to Markovian environments, whereby fairness requires the algorithm to be more likely to play actions that have a higher utility under the optimal policy.

In the context of fair statistical classification, a number of papers have asked what does it mean for a method of scoring individuals (e.g., for the purpose of car insurance, or release on bail) to be fair. In this setting it is useful to think about each individual as having a latent outcome, either positive or negative (no car accident, car accident.) One suggestion is that of *statistical parity*, which requires the average score of all members of each group be equal. For bandits we might interpret the activation probability as the score, and thus statistical parity would relate to always selecting each arm with equal probability. Another suggestion is *calibration within groups* [14], which requires for any score $s \in [0, 1]$ and any group, the approximate fraction of individuals with a positive outcome should be s; see also Chouldechova [2] for a related property. Considering also that there is competition between arms in our setting, this relates to our notion of calibrated fairness, where an arm is activated according to the probability that its realized reward is highest. Other definitions first condition on the latent truth; e.g., *balance* [14] requires that the expected score for an individual should be independent of group when conditioned on a positive outcome; see also Hardt et al. [9] for a related property. These concepts are harder to interpret in the present context of bandits problems. Interestingly, these different notions of fair classification are inherently in conflict with each other [2, 14].

This statistical learning framework has also been extended to decision problems by Corbett-Davies et al. [3], who analyze the tradeoff between utility maximization and the satisfaction of fairness constraints. Another direction is to consider *subjective fairness*, where the beliefs of the decision maker or external observer are also taken into account [4]. The present paper also briefly considers a specific notion of subjective fairness for bandits, where the similarity of arms is defined with respect to their marginal reward distribution.

2 THE SETTING

We consider the stochastic bandits problem, in which at each time step, a decision maker chooses one of k possible arms (possibly in a randomized fashion), upon which the decision maker receives a reward. We are interested in decision rules that are fair in regard to the decisions made about which arms to activate while achieving high total reward.

At each time step t, the decision maker chooses a distribution π_t over the available arms, which we refer to as the *decision rule*. Then nature draws an action $a_t \sim \pi_t$, and draws rewards:

$$r_i(t)|a_t = i \sim P(r_i|\theta_i),$$

where θ_i is the unknown parameter of the selected arm $a_t = i$, and where we denote the realized reward for arm i at time t by $r_i(t)$.

We denote the reward distribution $P(r_i|\theta_i)$ of arm i under some parameter θ_i as $r_i(\theta_i)$, with r_i denote the true reward distribution. Denote the vector form as $\mathbf{r} = (r_1, ..., r_k)$, while $\mathbf{r}_{-i,j}$ removes r_i

and r_j from **r**. If the decision maker has prior knowledge of the parameters $\theta = (\theta_1, ..., \theta_k)$, we denote this by $\beta(\theta)$.

2.1 Smooth Fairness

For divergence function D, let $D(\pi_t(i)\|\pi_t(j))$ to denote the divergence between the Bernoulli distributions with parameters $\pi_t(i)$ and $\pi_t(j)$, and use $D(r_i\|r_j)$ as a short-hand for the divergence between the reward distributions of arm i and j with true parameters θ_i and θ_j.

We define $(\epsilon_1, \epsilon_2, \delta)$-fair w.r.t. a divergence function D for an algorithm with an associated sequence of decision rules $\{\pi_t\}_t$ as:

Definition 2.1 (Smooth fairness). A bandit process is $(\epsilon_1, \epsilon_2, \delta)$-fair w.r.t. divergence function D, and $\epsilon_1 \geq 0, \epsilon_2 \geq 0, 0 \leq \delta \leq 1$, if with probability at least $1 - \delta$, in every round t, and for every pair of arms i and j:

$$D(\pi_t(i)\|\pi_t(j)) \leq \epsilon_1 D(r_i\|r_j) + \epsilon_2. \qquad (2)$$

Interpretation. This adapts the concept of "treating similar individuals similarly" [5] to the banditrs setting. If two arms have a similar reward distribution, then we can only be fair by ensuring that our decision rule has similar probabilities. The choice of D is crucial. For the KL divergence, if r_i, r_j do not have common support, our action distributions may be arbitrarily different. A Wasserstein distance, requires to treat two arms with a very close mean but different support similarly to each other. Most of the technical development will assume the total variation divergence.

As a preliminary, we also consider a variation on smooth fairness where we would like to be fair with regard to a posterior belief of the decision maker about the distribution on rewards associated with each arm.

For this, let the *posterior distribution on the parameter* θ_i of arm i be $\beta(\theta_i|h^t)$, where $h^t = (a_1, r_{a_1}(1), \ldots, a_t, r_{a(t)}(t))$, is the history of observations until time t. The *marginal reward distribution under the posterior belief* is

$$r_i(h^t) \triangleq \int_\Theta P(r_i \mid \theta_i) \, d\beta(\theta_i \mid h^t).$$

Definition 2.2 (Subjective smooth fairness). A bandit process is $(\epsilon_1, \epsilon_2, \delta)$-subjective fair w.r.t. divergence function D, and $\epsilon_1 \geq 0, \epsilon_2 \geq 0$, and $0 \leq \delta \leq 1$, if, with probability at least $1-\delta$, for every period t, and every pair of arms i and j,

$$D(\pi_t(i)\|\pi_t(j)) \leq \epsilon_1 D(r_i(h^t)\|r_j(h^t)) + \epsilon_2, \qquad (3)$$

where the initial belief of the decision maker is an uninformed prior for each arm.

2.2 Calibrated Fairness

Smooth fairness by itself does not seem strong enough for fair bandits algorithms. In particular, it does not require meritocracy: if two arms have quite different reward distributions then the weaker arm can be selected with higher probability than the stronger arm. This seems unfair to individuals in the group associated with the stronger arm.

For this reason we also care about *calibrated fairness*: an algorithm should sample each arm with probability equal to its reward being the greatest. This would ensure that even very weak arms

will be pulled sometimes, and that better arms will be pulled significantly more often.

Definition 2.3 (Calibrated fair policy). A policy π_t is *calibrated-fair* when it selects actions a with probability

$$\pi_t(a) = \mathbb{P}^*(a), \qquad \mathbb{P}^*(a) \triangleq P(a = \arg\max_{j \in [k]}\{r_j\}), \qquad (4)$$

equal to the probability that the reward realization of arm a is the highest, and we break ties at random in the case that two arms have the same realized reward.

Unlike smooth fairness, which can always be achieved exactly (e.g., through selecting each arm with equal probability), this notion of calibrated fairness is not possible to achieve exactly in a bandits setting while the algorithm is learning the quality of each arm. For this reason, we define the cumulative violation of calibration across all rounds T:

Definition 2.4 (Fairness regret). The fairness regret R_f of a policy π at time t is:

$$R_f(t) \triangleq \mathbb{E}\left[\sum_{i=1}^k \max(\mathbb{P}^*(i) - \pi_t(i), 0) \,\middle|\, \theta \right].$$

The cumulative fairness regret is defined as $R_{f,T} \triangleq \sum_{t=1}^T R_f(t)$.

Example 2.5. Consider a bandits problem with two arms, whose respective reward functions are random variables with realization probabilities:

- $P(r_1 = 1) = 1.0$;
- $P(r_2 = 0) = 0.6$ and $P(r_2 = 2) = 0.4$.

Since $\mathbb{E}(r_1) = 1.0$ and $\mathbb{E}(r_2) = 0.8$, a decision maker who optimizes expected payoff (and knows the distributions) would prefer to always select arm 1 over arm 2. Indeed, this satisfies weakly meritocratic fairness [12].

In contrast, calibrated fairness requires that arm 1 be selected 60% of the time and arm 2 40% of the time, since this matches the frequency with which arm 2 has the higher realized reward. In a learning context, we would not expect an algorithm to be calibrated in every period. Fairness regret measures the cumulative amount by which an algorithm is miscalibrated across rounds.

Smooth fairness by itself does not require calibration. Rather, smooth fairness requires, in every round, that the probability of selecting arm 1 be close to that of arm 2, where "close" depends on the particular divergence function. In particular, smooth fairness would not insist on arm 1 being selected with higher probability than arm 2, without an additional constraint such as maximising expected reward.

In Section 3, we introduce a simple Thompson-sampling based algorithm, and show that it satisfies smooth-subjective fairness. This algorithm provides a building block towards our main result, which is developed in Section 4, and provides smooth fairness and low fairness regret. Section 5 extends this algorithm to the dueling bandits setting.

3 SUBJECTIVE FAIRNESS

Subjective fairness is a conceptual departure from current approaches to fair bandits algorithms, which empasize fairness in every period t with respect to the true reward distributions for each arm. Rather, subjective fairness adopts the *interim* perspective of a Bayesian decision maker, who is fair with respect to his or her current beliefs. Subjective smooth fairness is useful as a building block towards our main result, which reverts to smooth fairness with regard to the true, objective reward distribution for each arm.

3.1 Stochastic-Dominance Thompson sampling

In *Thompson sampling* (TS), the probability of selecting an arm is equal to its probability of being the best arm under the subjective belief (posterior). This draws an immediate parallel with the Rawlsian notion of equality of opportunity, while taking into account informational constraints.

In this section we adopt a simple, multi-level sampling variation, which we refer to as *stochastic-dominance Thompson sampling*, SD_TS. This first samples parameters θ from the posterior, and then samples rewards for each arm, picking the arm with the highest reward realization.

The version of this algorithm for Bernoulli bandits with a Beta prior, where each arm's reward is generated according to a Bernoulli random variable, is detailed in Algorithm 1, which considers the marginal probability of an individual arm's reward realization being the greatest, and immediately provides subjective smooth fairness.

Algorithm 1 (SD_TS): Stoch.-Dom. Thompson sampling

For each action $a \in \{1, 2, ..., k\}$, set $S_a = F_a = 1/2$ (parameters for priors of Beta distributions).

for $t = 1, 2, ...,$ **do**

    For each action, sample $\theta_a(t)$ from Beta(S_a, F_a).

    Draw $\tilde{r}_a(t) \sim$ Bernoulli$(\theta_a(t)), \forall a$.

    Play arm $a_t := \text{argmax}_a \tilde{r}_a(t)$ (with random tie-breaking).

    Observe the true $r_{a_t}(t)$:

    • If $r_{a_t}(t) = 1$, $S_{a_t} := S_{a_t} + 1$;

    • else $F_{a_t} := F_{a_t} + 1$.

end for

THEOREM 3.1. *With (*SD_TS*), we can achieve* $(2, 0, 0)$*-subjective fairness under the total variation distance.*

PROOF. Define:

$$X_j(r_i) = \begin{cases} 1 & \text{if } r_i(h^t) > \max\{r'_j, \mathbf{r}'_{-i,j}\} \\ 0 & \text{if } r'_j > \max\{r_i(h^t), \mathbf{r}'_{-i,j}\} \\ Bin(1, \frac{1}{2}) & \text{otherwise} \end{cases}$$

where $r'_j \sim r_j(h^t)$ (similarly for $\mathbf{r}'_{-i,j}$) and Bin is a binomial random variable. First, we have for Thompson sampling:

$$D(r_i(h^t)\|r_j(h^t)) = \frac{1}{2} \cdot D(r_i(h^t)\|r_j(h^t)) + \frac{1}{2} \cdot D(r_i(h^t)\|r_j(h^t))$$

$$\overset{(a)}{\geq} \frac{1}{2} \cdot D(X_i(r_i(h_t))\|X_i(r_j(h_t))) + \frac{1}{2} \cdot D(X_j(r_i(h_t))\|X_j(r_j(h_t)))$$

$$= \frac{1}{2} \cdot D(\frac{1}{2}\|\pi_t(j) + \frac{1}{2} \cdot \pi_t(l \neq i, j)) + \frac{1}{2} \cdot D(\pi_t(i) + \frac{1}{2} \cdot \pi_t(l \neq i, j)\|\frac{1}{2})$$

(l denotes an arbitrary other agent than i, j.)

$$\overset{(b)}{\geq} D(\frac{1}{2} \cdot \frac{1}{2} + \frac{1}{2} \cdot \pi_t(i) + \frac{1}{2} \cdot \frac{1}{2} \cdot \pi_t(l \neq i, j)$$

$$\|\frac{1}{2} \cdot \frac{1}{2} + \frac{1}{2} \cdot \pi_t(j) + \frac{1}{2} \cdot \frac{1}{2} \cdot \pi_t(l \neq i, j))$$

$$= \frac{1}{2}|\pi_t(i) - \pi_t(j)| = \frac{1}{2} \cdot D(\pi_t(i)\|\pi_t(j))$$

where step (a) is by monotonicty and step (b) is by convexity of divergence function D. Therefore, ϵ_1 is equal to 2, and $\epsilon_2 = \delta = 0$. □

To further reduce the value of ϵ_1, we can randomize between the selection of the arms in the following manner:

- With probability $\epsilon/2$ select an arm selected by (SD_TS);
- Otherwise select uniformly at random another arm.

In that case, we have:

$$D(\pi_t(i)\|\pi_t(j)) = D(\frac{\epsilon}{2}\pi_{t,ts}(i) + \frac{1-\epsilon}{2}\frac{1}{2}\|\frac{\epsilon}{2}\pi_{t,ts}(j) + \frac{1-\epsilon}{2}\frac{1}{2})$$

$$\overset{\text{monotonicity}}{\leq} \frac{\epsilon}{2}D(\pi_{t,ts}(i)\|\pi_{t,ts}(j)) + \frac{1-\epsilon}{2}D(\frac{1}{2}\|\frac{1}{2})$$

$$\leq \epsilon D(r_i(h^t)\|r_j(h^t)).$$

Also see Sason and Verdú [15] for how to to bound $D(r_i(h_t)\|r_j(h_t))$ using another f-divergence (e.g. through Pinsker's inequality).

While, SD_TS algorithm is defined in a subjective setting, we can develop a minor variant of it in the objective setting. Even though the original algorithm already uses an uninformative prior, [2] to ensure that the algorithm output is more data than prior-driven, in the following section we describe an algorithm, based on SD_TS, which can achieve fairness with respect to the actual reward distribution of the arms.

4 OBJECTIVE FAIRNESS

In this section, we introduce a variant of SD_TS, which includes an initial phase of uniform exploration. We then prove the modified algorithm satisfies (objective) smooth fairness.

Many phased reinforcement learning algorithms [13], such as those based on successive elimination [6], explicitly separate time into exploration and exploitation phases. In the exploration phase, arms are prioritized that haven't been selected enough times. In the exploitation phase, arms are selected in order to target the chosen objective as best as possible given the available information. The algorithm maintains statistics on the arms, so that $O(t)$ is the set which we have not selected sufficiently to determine their

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[2]The use of Beta parameters equal to 1/2, corresponds to a Jeffrey's prior for Bernoulli distributions.

value. Following the structure of the deterministic exploration algorithm [17], we exploit whenever this set is empty, and uniformly choosing among all arms otherwise.[3]

Algorithm 2 Fair_SD_TS

At any t, denote by $n_i(t)$ the number of times arm i is selected up to time t. Check the following set:

$$O(t) = \{i : n_i(t) \leq C(\epsilon_2, \delta)\},$$

where $C(\epsilon_2, \delta)$ depends on ϵ_2 and δ.
- If $O(t) = \emptyset$, follow (SD_TS), using the collected statistics. (*exploitation*)
- If $O(t) \neq \emptyset$, select all arms equally likely. (*exploration*)

THEOREM 4.1. *For any $\epsilon_2, \delta > 0$, setting*

$$C(\epsilon_2, \delta) := \frac{(2 \max D(r_i || r_j) + 1)^2}{2\epsilon_2^2} \log \frac{2}{\delta},$$

we have that (Fair_SD_TS) *is* $(2, 2\epsilon_2, \delta)$-*fair w.r.t. total variation; and further it has fairness regret bounded as* $R_{f,T} \leq \tilde{O}((kT)^{2/3})$.

The proof of Theorem 4.1 is given in the following sketch.

PROOF. (sketch) We begin by proving the first part of Theorem 4.1: that for any $\epsilon_2, \delta > 0$, and setting $C(\epsilon_2, \delta)$ appropriately, we will have that Fair_SD_TS is $(2, 2\epsilon_2, \delta)$-fair w.r.t. total variation divergence. In the exploration phase, $D(\pi_t(i) || \pi_t(j)) = 0$, so the fairness definition is satisfied. For other steps, using Chernoff bounds we have that with probability at least $1 - \delta$

$$|\tilde{\theta}_i - \theta_i| \leq \frac{\epsilon_2}{2 \max D(r_i || r_j) + 1}, \forall i$$

Let the error term for θ_i be $\epsilon(i)$. Note that for a Bernoulli random variable, we have the following for the mixture distribution:

$$r_i(\tilde{\theta}_i) = (1 - \epsilon(i)/2)r(\theta_i) + \epsilon(i)/2r(1 - \theta_i)$$

with $\epsilon(i) \leq \frac{\epsilon_2}{2 \max D(r_i || r_j) + 1}$. Furthermore, using the convexity of D we can show that:

$$D(r_i(\tilde{\theta}_i) || r_j(\tilde{\theta}_j)) \leq D(r_i || r_j) + \epsilon_2 \quad (5)$$

Following the proof for Theorem 3.1, we then obtain that

$$D(\pi_t(i) || \pi_t(j)) \leq 2D(r_i(\tilde{\theta}_i) || r_j(\tilde{\theta}_j)),$$

which proves our statement.

We now establish the fairness regret. The regret incurred during the exploration phase can be bounded as $O(k^2 C(\epsilon_2, \delta))$.[4] For the exploitation phase, the regret is bounded by $O((\epsilon_2 + \delta)T)$. Setting

$$O((\epsilon_2 + \delta)T) = O(k^2 C(\epsilon_2, \delta))$$

we have the optimal ϵ is $\epsilon := k^{2/3}T^{-1/3}$. Further setting $\delta = O(T^{-1/2})$, we can show the regret is at the order of $\tilde{O}((kT)^{2/3})$. □

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[3]However, in our case, the actual drawing of the arms is stochastic to ensure fairness.
[4]This is different from standard deterministic bandit algorithms, where the exploration regret is often at the order of $kC(\epsilon_2, \delta)$. The additional k factor is due to the uniform selection in the exploration phase, while in standard deterministic explorations, the arm with the least number of selections will be selected.

It is possible to modify the sampling of the exploitation phase, alternating between sampling according to SD_TS and sampling uniformly randomly. This can be used to bring the factor 2 down to any $\epsilon_1 > 0$, at the expense of reduced utility.

4.1 Connection with proper scoring rules

There is a connection between calibrated fairness and proper scoring rules. Suppose we define a *fairness loss function* $\mathcal{L}_f$ for decision policy π, such that $L_f(\pi) = \mathcal{L}(\pi, a_{t,best})$, where arm $a_{t,best}$ is the arm with the highest realized reward at time t. The expected loss for policy π is

$$\mathbb{E}(L_f(\pi)) = \sum_{i=1}^{k} \mathbb{P}^*(i) \cdot \mathcal{L}(\pi, i).$$

If $\mathcal{L}$ is strictly proper [7], then the optimal decision rule π in terms of L_f is calibrated.

PROPOSITION 4.2. *Consider a fairness loss function L_f defined as:*

$$L_f(\pi) = \mathcal{L}(\pi, a_{t,best}),$$

where $\mathcal{L}$ is a strictly proper loss function. Then a decision rule $\bar{\pi}$ that minimizes expected loss is calibrated fair.

PROOF. We have:

$$\bar{\pi} \in \arg\min_\pi \mathbb{E}(L_f(\pi)) = \arg\min_\pi \sum_{i=1}^{k} \mathbb{P}^*(i) \cdot \mathcal{L}(\pi, i) = \{\mathbb{P}^*(i)\},$$

where the last equality comes from the strict properness of $\mathcal{L}$. □

This connection between calibration and proper scoring rules suggests an approach to the design of bandits algorithms with low fairness regret, by considering different proper scoring rules along with online algorithms to minimize loss.

5 DUELING BANDIT FEEDBACK

After an initial exploration phase, Fair_SD_TS selects an arm according to how likely its sample realization will dominate those of other arms. This suggests that we are mostly interested in the stochastic dominance probability, rather than the joint reward distribution. Recognizing this, we now move to the *dueling bandits* framework [18], which examines pairwise stochastic dominance.

In a dueling bandit setting, at each time step t, the decision maker chooses two arms $a_t(1), a_t(2)$ to "duel" with each other. The decision maker doesn't observe the actual rewards of $r_{a_t(1)}(t), r_{a_t(2)}(t)$, but rather the outcome $\mathbb{1}(r_{a_t(1)}(t) > r_{a_t(2)}(t))$. In this section, we extend our fairness results to the dueling bandits setting.

5.1 A Plackett-Luce model

Consider the following model. Denote the probability of arm i's reward being greater than arm j's reward by:

$$p_{i,j} := \mathbb{P}(i > j) := P(r_i > r_j), \quad (6)$$

where we assume a stationary reward distribution over time t. To be concrete, we adopt the *Plackett-Luce* (PL) model [1, 8], where every arm i is parameterized by a *quality parameter*, $v_i \in \mathbb{R}_+$, such that

$$p_{i,j} = \frac{v_i}{v_i + v_j}. \quad (7)$$

Furthermore, let $\mathcal{M} = [p_{i,j}]$ denote the matrix of pairwise probabilities $p_{i,j}$. This is a standard setting to consider in the dueling bandit literature [16, 18].

With knowledge of $\mathcal{M}$, we can efficiently simulate the best arm realization. In particular, for the rank over arms rank $\sim \mathcal{M}$ generated according to the PL model (by selecting pairwise comparisons one by one, each time selecting one from the remaining set with probability proportional to v_i), we have [8]:

$$P(\text{rank}|v) = \prod_{i=1}^{k} \frac{v_{o_i}}{\sum_{j=i}^{k} v_{o_j}},$$

where $\mathbf{o} = \text{rank}^{-1}$. In particular, the marginal probability in the PL model that an arm is rank 1 (and the best arm) is just:

$$P(\text{rank}(1) = i) = \frac{v_i}{\sum_j v_j} = \frac{1}{1 + \sum_{j \neq i} v_j / v_i}.$$

Finally, knowledge of $\mathcal{M}$ allows us to directly qcalculate each arm's quality from (7).

$$v_j / v_i := \frac{p_{j,i}}{1 - p_{j,i}}.$$

Thus, with estimates of the quality parameters ($\mathcal{M}$) we can estimate $P(\text{rank}(1) = i)$ and directly sample from the best arm distribution and simulate stochastic-dominance Thompson sampling.

We will use dueling bandit feedback to estimate pairwise probabilities, denoted $\tilde{p}_{i,j}$, along with the corresponding comparison matrix denoted by $\tilde{\mathcal{M}}$. In particular, let $n_{i,j}(t)$ denote the number of times arms i and j are selected up to time t. Then we estimate the pairwise probabilities as:

$$\tilde{p}_{i,j}(t) = \frac{\sum_{n=1}^{n_{i,j}(t)} \mathbb{1}(r_i(n) > r_j(n))}{n_{i,j}(t)}, \; n_{i,j}(t) \geq 1. \tag{8}$$

With accurate estimation of the pairwise probabilities, we are able to accurately approximate the probability that each arm will be rank 1. Denote $\widetilde{\text{rank}} \sim \tilde{M}$ as the rank generated according to the PL model that corresponds with matrix $\tilde{M}$. We estimate the ratio of quality parameters $\widetilde{(\frac{v_i}{v_j})}$ using $\tilde{p}_{i,j}$, as

$$\widetilde{\left(\frac{v_i}{v_j}\right)} = \frac{\tilde{p}_{i,j}}{1 - \tilde{p}_{i,j}}.$$

Given this, we can then estimate the probability than arm i has the best reward realization:

$$P(\widetilde{\text{rank}}(1) = i) = \frac{1}{1 + \sum_{j \neq i} \widetilde{\left(\frac{v_j}{v_i}\right)}}. \tag{9}$$

LEMMA 5.1. *When* $|\tilde{p}_{i,j} - p_{i,j}| \leq \epsilon$, *and ϵ is small enough, we have that in the Plackett-Luce model,* $|P(\widetilde{\text{rank}}(1) = i) - P(\text{rank}(1) = i)| \leq O(k\epsilon)$.

This lemma can be established by establishing a concentration bound on $\widetilde{(\frac{v_i}{v_j})}$. We defer the details to a long version of this paper.

Given this, we can derive an algorithm similar to Fair_SD_TS that can achieve calibrated fairness in this setting, by appropriately setting the length of the exploration phase, and by simulating the probability that a given arm has the highest reward realization. This dueling version of Fair_SD_TS is the algorithm Fair_SD_DTS, and detailed in Algorithm 3.

Algorithm 3 (Fair_SD_DTS)

At any t, select two arms $a_1(t), a_2(t)$, and receive a realization of the following comparison: $\mathbb{1}(r_{a_t(1)}(t) > r_{a_t(2)}(t))$. Check the following set:

$$O(t) = \{i : n_{i,j}(t) \leq C(\epsilon_2, \delta)\},$$

where $C(\epsilon_2, \delta)$ depends on ϵ_2 and δ.
- If $O(t) = \emptyset$, follow (SD_TS), using the collected statistics. (*exploitation*)
- If $O(t) \neq \emptyset$, select all pairs of arms equally likely. (*exploration*)

Update $\tilde{p}_{a_t(1), a_t(2)}$, for the pair of selected arms (Eqn. (8)).
Update $P(\text{rank}(1) = i)$ using $\tilde{M}$ (Eqn. (9)).

Due to the need to explore all pairs of arms, a larger number of exploration rounds $C(\epsilon_2, \delta)$ is needed, and thus the fairness regret scales as $R_f(T) \leq \tilde{O}(k^{4/3} T^{2/3})$:

THEOREM 5.2. *For any* $\epsilon_2, \delta > 0$, *setting*

$$C(\epsilon_2, \delta) \triangleq O\left(\frac{(2 \max D(r_i || r_j) + 1)^2 k^2}{2\epsilon_2^2} \log \frac{2}{\delta}\right),$$

we have that (Fair_SD_DTS) *is* $(2, 2\epsilon_2, \delta)$-*fair w.r.t. total variation; and further it has fairness regret bounded as* $R_{f,T} \leq \tilde{O}(k^{4/3} T^{2/3})$.

This proof is similar to the fairness regret proof for Theorem 4.1, once we established Lemma 5.1. We defer the details to the full version of the paper.

6 CONCLUSION

In this paper we adapt the notion of "treating similar individuals similarly"[5] to the bandits problem, with similarity based on the distribtion on rewards, and this property of smooth fairness required to hold along with (approximate) calibrated fairness. Calibrated fairness requires that arms that are worse in expectation still be played if they have a chance of being the best, and that better arms be played significantly more often than weaker arms.

We analyzed Thompson-sampling based algorithms, and showed that a variation with an initial uniform exploration phase can achieve a low regret bound with regard to calibration as well as smooth fairness. We further discussed how to adopt this algorithm to a dueling bandit setting together with Plackett-Luce.

In future work, it will be interesting to consider contextual bandits (in the case in which the context still leaves residual uncertainty about quality), to establish lower bounds for fairness regret, to consider ways to achieve good calibrated fairness uniformly across rounds, and to study the utility of fair bandits algorithms (e.g., with respect to standard notions of regret) and while allowing for a tradeoff against smooth fairness for different divergence functions and fairness regret. In addition, it will be interesting to explore the connection between strictly-proper scoring rules and calibrated fairness, as well as to extend Lemma 5.1 to more general ranking models.

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