

# BIOS526\_CourseSummary

Monday, November 27, 2023 13:54



BIOS526\_C  
ourseSum...

## BIOS 526 Course Summary

1. Ordinary least squares (M1):

$$\begin{aligned} Y &\in \mathbb{R}^N \\ \text{Cov } Y &= \sigma^2 \mathbf{I} \\ \hat{\beta}^{OLS} &= (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'Y \end{aligned}$$

2. Generalized least squares (M4, slide 16) is a formulation for any given covariance matrix:

$$\begin{aligned} \text{Cov } Y &= \Sigma \\ \hat{\beta}^{GLS} &= (\mathbf{X}'\Sigma^{-1}\mathbf{X})^{-1}\mathbf{X}'\Sigma^{-1}Y \end{aligned}$$

- Need to know  $\Sigma$ .
- In practice, we impose structure on  $\Sigma$  and estimate a statistical model, e.g., mixed models with Gaussian assumptions, GEE with Gaussian assumptions.

3. Linear mixed models (M2)

- Random intercept:

$$\begin{aligned} y_{ij} &= x_{ij}'\beta + \theta_i + \epsilon_{ij} \\ \theta_i &\stackrel{iid}{\sim} N(0, \tau^2) \\ \epsilon_{ij} &\stackrel{iid}{\sim} N(0, \sigma^2) \\ \epsilon_{ij} &\perp\!\!\!\perp \theta_i \end{aligned}$$

- Also see random slopes and hierarchical models

4. Generalized linear models (M3 part I):

$$\begin{aligned} y_i &\stackrel{ind}{\sim} P(Y_i | x_i'\beta) \\ g(E(Y_i)) &= x_i'\beta \end{aligned}$$

- logistic regression for 0, 1 response:

– Model:

$$y_i \overset{\text{ind}}{\sim} \text{Bernoulli}(E[Y_i])$$

$$\log \left\{ \frac{E[Y_i]}{1 - E[Y_i]} \right\} = \mathbf{x}_i' \boldsymbol{\beta}$$

NO  $\varepsilon_i$  !!!

logit link

–  $\text{Var}(Y_i) = E(Y_i) \{1 - E(Y_i)\}$ .

- Poisson regression for count data. i.e.  $p_i(1-p_i)$

– Model:

$$y_i \overset{\text{ind}}{\sim} \text{Poisson}(E[Y_i]),$$

$$\log \{E[Y_i]\} = \mathbf{x}_i' \boldsymbol{\beta}.$$

–  $\text{Var}(Y_i) = E(Y_i) \rightarrow$  assumption rarely satisfied in practice

– Watch out for overdispersion, i.e.,  $\text{Var}(Y_i) > E(Y_i)$ .

can fit quasipoisson (ok approach), or a GEE (robust variance)

## 5. Generalized linear mixed models (M2, M3 part II):

Random intercept model:

$$y_{ij} \sim P(Y_{ij} | \mathbf{x}_{ij}' \boldsymbol{\beta} + \theta_i)$$

$$g \{E(y_{ij} | \theta_i)\} = \mathbf{x}_{ij}' \boldsymbol{\beta} + \theta_i$$

$$\theta_i \overset{\text{iid}}{\sim} N(0, \tau^2)$$

- Handle repeated measurements / longitudinal / clustered data. ↓ observations that are not independent
- For Gaussian, interpretation of  $\boldsymbol{\beta}$  not impacted by conditional versus marginal (the estimates of  $\boldsymbol{\beta}$  from LMM and GEE are different but usually similar, in some cases GEE with exchangeable correlation structure and random intercept LMM have equivalent  $\hat{\boldsymbol{\beta}}$ ).
- For logistic, interpretation of  $\boldsymbol{\beta}$  in GLMM (conditional model) is different from the interpretation in a GEE (the estimates of  $\boldsymbol{\beta}$  from GLMM and GEE are different).
- For Poisson, the interpretation of  $\beta_1, \dots, \beta_p$  in GLMM (conditional model) is the same as in the GEE due to a special property of the log link. The intercept changes, as the intercept in the marginal models includes  $\tau^2/2$ ; see the R code.
- Use mixed models if interested in subject-specific predictions (shrinkage towards population effects).
- Can use if no overdispersion in logistic or Poisson, no heteroscedasticity in Gaussian.

6. Generalized estimating equations (M4):

$$y_{ij} \sim P(Y_{ij} | x'_{ij} \beta)$$

$$g(E[Y_{ij}]) = x'_{ij} \beta$$

$$\text{Cov}(\mathbf{Y}_i) = \mathbf{D}_i^{1/2} \mathbf{R}(\alpha) \mathbf{D}_i^{1/2}$$

where  $\mathbf{R}(\alpha)$  is the working correlation and  $\mathbf{D}_i$  is a diagonal matrix with diagonal elements equal to the variance determined by the likelihood.

- Handle repeated measurements / longitudinal / clustered data.
- Use robust standard errors.
- Use if heteroscedasticity and/or overdispersion (valid inference, unlike GLMM).
- Marginal inference (no random effects).

$$Y_i \in \mathbb{R}^{P_i}$$

$$\begin{bmatrix} 1 & \alpha & \alpha \\ 2 & 1 & \alpha \\ 2 & 1 & 1 \end{bmatrix}$$

$$\begin{bmatrix} [p_{11}(1-p_{11})]^{1/2} & 0 & 0 \\ 0 & [p_{12}(1-p_{12})]^{1/2} & 0 \\ 0 & 0 & [p_{13}(1-p_{13})]^{1/2} \end{bmatrix}$$

$$\text{working correlation: } \text{Cor}(y_{ij}, y_{i'j}) = 0$$

7. Generalized additive models (M5):

$$y_i \sim P(Y_i | \beta_0, s_1(x_{i1}), \dots, s_p(x_{ip}))$$

$$g(E(Y_i)) = \beta_0 + s_1(x_{i1}) + \dots + s_p(x_{ip})$$

- Handle non-linear effects.
- Can incorporate random effects for longitudinal / repeated measures / clustered data.
- Can generalize interactions from linear models to bivariate splines, e.g.,  $s(x_{i1}, x_{i2})$ , i.e., 2D surfaces.
- Estimate  $s(x_{ik})$  using either cross-validation or mixed model formulation of spline coefficients.

8. Bias-Variance Tradeoff (M5, part I, slides 33-43, M6, part II, slides 5-6)

- $MSE(\hat{f}(x)) = Var(\hat{f}(x)) + Bias(\hat{f}(x))^2$
- Fewer parameter: more bias, less variance
- More parameters: less bias, more variance
- Use cross-validation or generalized cross-validation to approximately minimize the MSE

9. Principal component analysis (M6 I): uses the singular value decomposition on standardized  $N \times p$  data:

$$\mathbf{X}_{scaled} = \mathbf{U} \mathbf{D} \mathbf{V}'$$

- Lower dimensional representation using first  $q$  left eigenvectors.
- Principal component scores:  $\mathbf{U}_{1:q} \mathbf{D}_{1:q}$ ,

3

$$\text{loading } V_{1:q} : \quad \mathbf{X}_{scaled} V_{1:q} = \mathbf{U} \mathbf{D} \mathbf{V}' \mathbf{V}_{(1:q)} \\ = u_1 \dots u_q$$

- Can use in principal component regression when have issues with multicollinearity.

10. Ridge Regression (L2-norm regularization) (M5 part II, M6 part II):

$$\hat{\beta}^{Ridge} = (\mathbf{X}'\mathbf{X} + \lambda\mathbf{I})^{-1}\mathbf{X}'\mathbf{Y} \quad \text{— solution for gaussian case, i.e., when}$$

- Regularization method with a nice closed form.
- Also extends to likelihoods (M6 part II):

loss fun = SSE

$$\underset{\beta}{\operatorname{argmin}} \quad - \sum_{i=1}^n \ell(y_i; \mathbf{x}_i' \beta) + \lambda \|\beta\|_2^2.$$

- Can use when lots of covariates,  $p > n$ .
- Use for shrinking spline coefficients in GAMs (used in MGCV).

11. Lasso (L1-norm regularization) (M6 part II):

$$\hat{\beta}^{Lasso} = \underset{\beta}{\operatorname{argmin}} \quad - \sum_{i=1}^n \ell(y_i; \mathbf{x}_i' \beta) + \lambda \|\beta\|_1$$

- Regularization that results in variable selection by setting many coefficients equal to 0.

often a go-to

12. Elastic net (L1-norm and L2-norm regularization) (M6 part II):

$$\hat{\beta}^{ENet} = \underset{\beta}{\operatorname{argmin}} \quad - \sum_{i=1}^n \ell(y_i; \mathbf{x}_i' \beta) + \lambda \sum_{j=1}^p \left( \alpha |\beta_j| + \frac{(1-\alpha)}{2} \beta_j^2 \right).$$

- A good choice when predictors are correlated.
- Use for variable selection.