

MoHBF Analytical Sheet 3

Stephan Gabler

June 14, 2010

I cooperated with Duncan Blythe and Rafael Schultze-Kraft.

1 Exercise 1 (Linear Regression and LMSE)

The error function is given by $E = \sum_{i=1}^N (y_i - \alpha x_i - \beta)^2$. Differentiate E with respect to α and β to get:-

$$\frac{\partial E}{\partial \beta} = -2 \sum_{i=1}^N (y_i - \alpha x_i - \beta) = 0 \quad (1)$$

Which is zero at a minimum.
so for a minimum w.r.t. beta:

$$\beta = \frac{1}{N} \left(\sum_{i=1}^N y_i - \alpha \sum_{i=1}^N x_i \right) \quad (2)$$

But also if we substitute this value for β into the error function and differentiate w.r.t. α , then we get:-

$$\begin{aligned} \frac{\partial E}{\partial \alpha} &= \frac{\partial}{\partial \alpha} \sum_{i=1}^N \left(y_i - \alpha x_i - \frac{1}{N} \sum_{j=1}^N y_j + \frac{\alpha}{N} \sum_{j=1}^N x_j \right)^2 \\ &= \frac{\partial}{\partial \alpha} \sum_{i=1}^N \left(y_i - \alpha x_i - \frac{1}{N} \sum_{j=1}^N y_j + \frac{\alpha}{N} \sum_{j=1}^N x_j \right) \times \left(-x_i + \frac{1}{N} \sum_{j=1}^N x_j \right) \\ &= 2 \sum_{i=1}^N \left(-y_i x_i + \alpha x_i^2 + \frac{x_i}{N} \sum_{j=1}^N y_j - \frac{\alpha}{N} x_i \sum_{j=1}^N x_j \right) + \sum_{i=1}^N \left(y_i \frac{1}{N} \sum_{j=1}^N x_j - \frac{1}{N^2} \sum_{j=1}^N y_j \sum_{j=1}^N x_j + \frac{\alpha}{N^2} \left(\sum_{j=1}^N x_j \right)^2 \right) \end{aligned}$$

But this last expression is zero whenever

$$\begin{aligned} &\alpha \left(\sum_{i=1}^N x_i^2 - \frac{1}{N} \left(\sum_{i=1}^N x_i \right)^2 - \frac{1}{N} \left(\sum_{i=1}^N x_i \right)^2 + \frac{N}{N^2} \left(\sum_{i=1}^N x_i \right)^2 \right) \\ &= \sum_{i=1}^N y_i x_i - \frac{1}{N} \sum_{i=1}^N x_i \sum_{i=1}^N y_i - \frac{1}{N} \sum_{i=1}^N y_i \sum_{i=1}^N x_i + \frac{N}{N^2} \sum_{i=1}^N y_i \sum_{i=1}^N x_i \end{aligned}$$

And this is the result we wanted:-

$$\alpha = \frac{N \sum_{i=1}^N x_i y_i - \sum_{i=1}^N x_i \sum_{i=1}^N y_i}{N \sum_{i=1}^N x_i^2 - \left(\sum_{i=1}^N x_i \right)^2} \quad (3)$$

2 Exercise 2 (LMSE and Maximum Likelihood)

All data-points $y_1 \cdots y_i$ are assumed to be gaussian distributed around the true model $f(\vec{x}, \vec{\alpha})$. The standard distributions σ are the same for all data points. The joint density for all data-points is then:

$$\prod_{i=1}^N \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y_i - \mu)^2}{2\sigma^2}} \quad (4)$$

with μ being equal to the value of the true model $f(\vec{x}, \vec{\alpha})$. The negative log-likelihood of the model is

$$L(\vec{\alpha}, \vec{x}) = \sum_{i=1}^N \ln\left(\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y_i - \mu)^2}{2\sigma^2}}\right) \quad (5)$$

which we then want to maximize.

$$\begin{aligned} \max_{\vec{\alpha}} L(\vec{\alpha}, \vec{x}) &= \sum_{i=1}^N \ln\left(\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y_i - f(\vec{x}, \vec{\alpha}))^2}{2\sigma^2}}\right) \\ &= -n \ln\left(\frac{1}{\sqrt{2\pi\sigma^2}}\right) - \sum_{i=1}^N \ln\left(e^{-\frac{(y_i - f(\vec{x}, \vec{\alpha}))^2}{2\sigma^2}}\right) \\ &= \underbrace{-\frac{n}{2\sigma^2} \ln\left(\frac{1}{\sqrt{2\pi\sigma^2}}\right)}_{\text{constant}} + \sum_{i=1}^N (y_i - f(\vec{x}, \vec{\alpha}))^2 \end{aligned} \quad (6)$$

Because we can neglect the constant term, only the last part has to be maximized which is the same as the LMSE.

$$\sum_{i=1}^N (y_i - f(\vec{x}, \vec{\alpha}))^2 \quad (7)$$

I think that for both behavioral and psychophysical experiments assuming σ to be constant for all \vec{x} might not be the best assumption. For example could the std of a effect increase with time because the subjects get tired.

3 Exercise 3 (Maximum Likelihood and Binomial Data)

The binomial distribution has the form $Pr(K = k) = \binom{N}{k} p^k (1-p)^{N-k}$. Here we set N to the no. of trials, K to the number of successes on the part of the observer. and p to the actual probability of detection. Then our likelihood function is $\binom{N}{k} \tilde{p}^k (1-\tilde{p})^{N-k}$. Differentiating this w.r.t \tilde{p} gives:- $\binom{N}{k} k \tilde{p}^{k-1} (1-\tilde{p})^{N-k} - \binom{N}{k} (N-k) \tilde{p}^k (1-\tilde{p})^{N-k-1}$. This is zero if and only if:-

$$k \tilde{p}^{k-1} (1-\tilde{p})^{N-k} = (N-k) \tilde{p}^k (1-\tilde{p})^{N-k-1} \quad (8)$$

$$k(1-\tilde{p}) = (N-k)\tilde{p} \quad (9)$$

$$k = N\tilde{p} \quad (10)$$

$$\tilde{p} = \frac{k}{N} \quad (11)$$

Now try the log likelihood, which is given by:- $\log\left(\binom{N}{k}\right) + \log(\tilde{p}^k) + \log((1 - \tilde{p})^{N-k})$. Differentiating this w.r.t. \tilde{p} and setting to zero gives:-

$$\frac{k\tilde{p}^{k-1}}{(\tilde{p}^k)} = \frac{(N-k)(1-\tilde{p})^{N-k-1}}{(1-\tilde{p})^{N-k}} \quad (12)$$

This clearly gives the same result as before.

4 Exercise 4 (Binomial Data and Adaptive Procedures)

First we need to define what we mean here by convergence since we know that at the least after every 3 presentations our stimulus intensity will be changed by a discrete amount according to how the 3 down 1 up procedure is defined, so the sequence of stimulus values themselves don't converge in a mathematical sense.

We could do this in terms of average of sum of the preceding terms. So we might say (for example) that the sequence of stimulus presentations converges to a if for any $\epsilon > 0$ there exists a time point T such that if $t > T$ $|\frac{1}{M} \sum_{i=1}^t s(i) - a| < \epsilon$.

But this won't work since it's conceivable that these terms $s(i)$ could alternate between two distinct values, and the *convergence* point will lie somewhere in between.