

- Consider a single-server system in which the arrivals occur with a Poisson distribution and the service time has an exponential distribution.
- Suppose the study objective is to measure the mean waiting time, defined as the time entities spend waiting to receive service and excluding the service time itself.
- This system is commonly denoted by M/M/1 which indicates; first, that the inter-arrival time is distributed exponentially; second that the service time is distributed exponentially; and, third, that there is one server. The M stands for Markovian, which implies an exponential distribution.



- In a simulation run, the simplest approach is to estimate the mean waiting time by accumulating the waiting time of n successive entities and dividing by n.
- This measure, the sample mean, is denoted by x(n) to emphasize the fact that its value depends upon the number of observations taken.
- If x_i (i=1,2,....,n) are the individual waiting times (including the value 0 for those entities that do not have to wait), then

$$\overline{x}(n) = \frac{1}{n} \sum_{i=1}^{n} x_i$$



- Whenever a waiting line forms, the waiting time of each entity on the line clearly depends upon the waiting time of its predecessors.
- Any series of data that has this property of having one value affect other values is said to be autocorrelated.
- The sample mean of autocorrelated data can be shown to approximate a normal distribution as the sample size increases.



- The equation $\bar{x}(n) = \frac{1}{n} \sum_{i=1}^{n} x_i$ remains a satisfactory

estimate for the mean of autocorrelated data.

- A simulation run is started with the system in some initial state, frequently the idle state, in which no service is being given and no entities are waiting.
- The early arrivals then have a more than normal probability of obtaining service quickly, so a sample mean that includes the early arrivals will be biased.



- For a given sample size starting from a given initial condition, the sample mean distribution is stationary; but, if the distributions could be compared for different sample sizes, the distribution would be slightly different.
- The following figure is based on theoretical results, which shows how the expected value of sample mean depends upon the sample length, for the M/M/1 system, starting from an initial empty state, with a server utilization of 0.9.



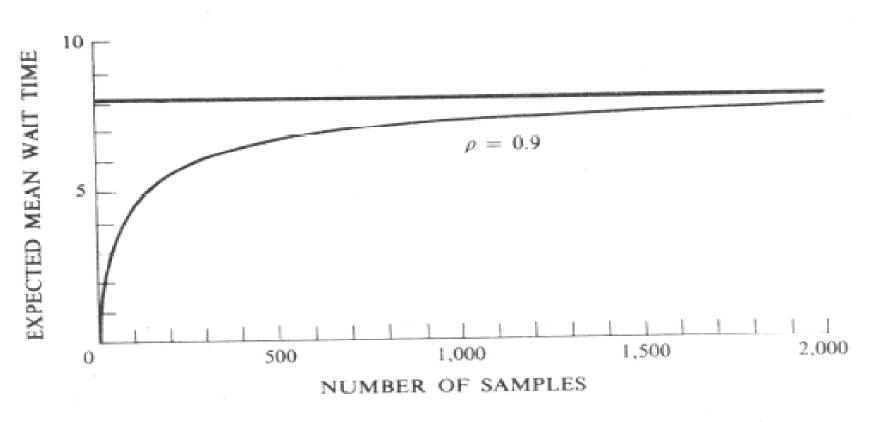


Figure 14-2. Mean wait time in M/M/1 system for different sample sizes.



- The precision of results of a dynamic stochastic can be increased by repeating the experiment with different random numbers strings.
- For each replication of a small sample size, the sample mean is determined.
- The sample means of the independent runs can be further used to estimate the variance of distribution. Let X ij be the ith observation in jth run, then the sample mean and variance for the jth run are:



$$\overline{x_j}(n) = \frac{1}{n} \sum_{i=1}^n \chi_{ij}$$

$$s_j^2 = \frac{1}{n-1} \sum_{i=1}^n \left[x_{ij} - \overline{x}_j(n) \right]^2$$



 When we have similar means and variances for m independent measurements, then by combining them, the mean and variance for the population can be obtained as:

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$$\frac{1}{x} = \frac{1}{p} \sum_{j=1}^{p} x_j(n)$$

$$S^{2} = \frac{1}{p} \sum_{j=1}^{p} s_{j}^{2}(n)$$



 The following figure shows the result of applying the procedure to experimental results for the M/M/1 system.

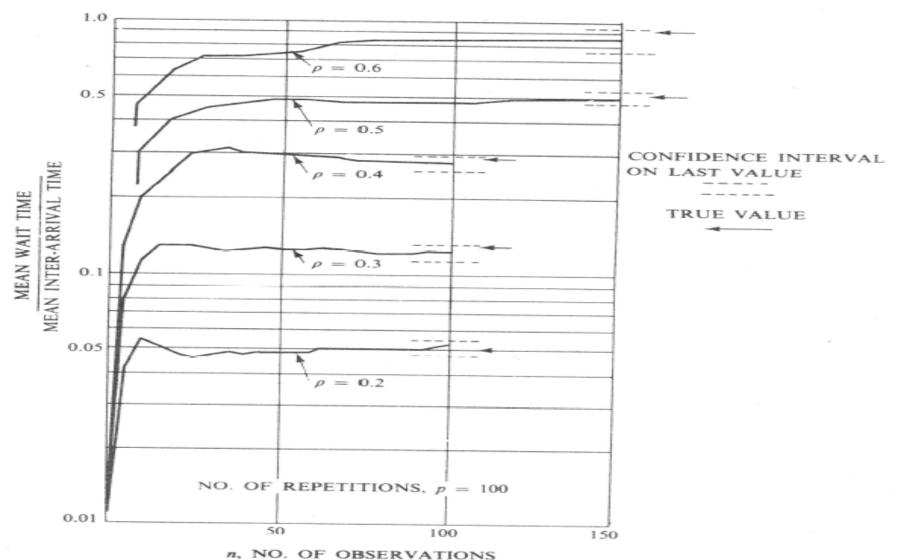


Figure 14-3. Experimentally measured wait time in M/M/1 system for different sample sizes.

- This variance can further be used to establish the confidence interval for p-1 degrees of freedom.
- The length of run of replications is so selected that all combined it comes to the sample size N. i.e. p.n=N.
- By increasing the number of replications and shortening their length of run, the confidence interval can be narrowed.
- But due to shortening of length of replication the effect of starting conditions will increase.
- The results obtained will not be accurate, especially when the initiazation of the runs is not proper.
- Thus, a compromise has to be made.
- There is no established procedure of dividing the sample size N into replications.
- However, it is suggested that the number of replications should not be very large, and that the sample means should approximate a normal distribution.



- Two general approaches can be taken to remove the bias: the system can be started in a more representative state that the empty state, or the first part of the simulation can be ignored.
- The ideal situation is to know the steady state distribution for the system, and select the initial condition from that distribution.
- In the study previously discussed, repeated the experiments on the M/M/1 system, supplying an initial waiting line for each run, selected at random from the known steady state distribution of waiting line.



- The case of 40 repetitions of 320 samples, which previously resulted in a coverage of only 9% was improved to coverage of 88%.
- The more common approach to removing the initial bias is to eliminate an initial section of the run.
- The run is started from an idle state and stopped after a certain period of time.



- The run is then restarted with statistics being gathered from the point of restart.
- It is usual to program the simulation so that statistics are gathered from the beginning, and simply wipe out the statistics gathered up to the point of restart.
- No simple rules can be given to decide how long an interval should be eliminated.



- The disadvantage of eliminating the first part of a simulation run is that the estimate of the variance, needed to establish a confidence limit, must be based on less information.
- The reduction in bias, therefore, is obtained at the price of increasing the confidence interval size.



Reference

Geoffrey Gordon, System Simulation,
Chapter 14, analysis of simulation output