Properties of Expected values and Variance

Christopher Croke

University of Pennsylvania

Math 115 UPenn, Fall 2011

Consider a random variable Y = r(X) for some function r, e.g. $Y = X^2 + 3$ so in this case $r(x) = x^2 + 3$.

Consider a random variable Y = r(X) for some function r, e.g. $Y = X^2 + 3$ so in this case $r(x) = x^2 + 3$. It turns out (and we have already used) that

$$E(r(X)) = \int_{-\infty}^{\infty} r(x)f(x)dx.$$

Consider a random variable Y = r(X) for some function r, e.g. $Y = X^2 + 3$ so in this case $r(x) = x^2 + 3$. It turns out (and we have already used) that

$$E(r(X)) = \int_{-\infty}^{\infty} r(x)f(x)dx.$$

This is not obvious since by definition $E(r(X)) = \int_{-\infty}^{\infty} x f_Y(x) dx$ where $f_Y(x)$ is the probability density function of Y = r(X).

Consider a random variable Y = r(X) for some function r, e.g. $Y = X^2 + 3$ so in this case $r(x) = x^2 + 3$. It turns out (and we have already used) that

$$E(r(X)) = \int_{-\infty}^{\infty} r(x)f(x)dx.$$

This is not obvious since by definition $E(r(X)) = \int_{-\infty}^{\infty} x f_Y(x) dx$ where $f_Y(x)$ is the probability density function of Y = r(X). You get from one integral to the other by careful uses of u substitution.

Consider a random variable Y = r(X) for some function r, e.g. $Y = X^2 + 3$ so in this case $r(x) = x^2 + 3$. It turns out (and we have already used) that

$$E(r(X)) = \int_{-\infty}^{\infty} r(x)f(x)dx.$$

This is not obvious since by definition $E(r(X)) = \int_{-\infty}^{\infty} x f_Y(x) dx$ where $f_Y(x)$ is the probability density function of Y = r(X).

You get from one integral to the other by careful uses of u substitution.

One consequence is

$$E(aX + b) = \int_{-\infty}^{\infty} (ax + b)f(x)dx = aE(X) + b.$$



Consider a random variable Y = r(X) for some function r, e.g. $Y = X^2 + 3$ so in this case $r(x) = x^2 + 3$. It turns out (and we have already used) that

$$E(r(X)) = \int_{-\infty}^{\infty} r(x)f(x)dx.$$

This is not obvious since by definition $E(r(X)) = \int_{-\infty}^{\infty} x f_Y(x) dx$ where $f_Y(x)$ is the probability density function of Y = r(X). You get from one integral to the other by careful uses of u

substitution.

One consequence is

$$E(aX + b) = \int_{-\infty}^{\infty} (ax + b)f(x)dx = aE(X) + b.$$

(It is not usually the case that E(r(X)) = r(E(X)).)



Consider a random variable Y = r(X) for some function r, e.g. $Y = X^2 + 3$ so in this case $r(x) = x^2 + 3$. It turns out (and we have already used) that

$$E(r(X)) = \int_{-\infty}^{\infty} r(x)f(x)dx.$$

This is not obvious since by definition $E(r(X)) = \int_{-\infty}^{\infty} x f_Y(x) dx$ where $f_Y(x)$ is the probability density function of Y = r(X). You get from one integral to the other by careful uses of u

substitution.

One consequence is

$$E(aX + b) = \int_{-\infty}^{\infty} (ax + b)f(x)dx = aE(X) + b.$$

(It is not usually the case that E(r(X)) = r(E(X)).) Similar facts old for discrete random variables.

If $X_1, X_2, X_3, ... X_n$ are random variables and $Y = r(X_1, X_2, X_3, ... X_n)$ then

$$E(Y) = \int \int \int ... \int r(x_1, x_2, x_3, ..., x_n) f(x_1, x_2, x_3, ..., x_n) dx_1 dx_2 dx_3 ... dx_n$$

where $f(x_1, x_2, x_3, ..., x_n)$ is the joint probability density function.

If $X_1, X_2, X_3, ... X_n$ are random variables and $Y = r(X_1, X_2, X_3, ... X_n)$ then

$$E(Y) = \int \int \int ... \int r(x_1, x_2, x_3, ..., x_n) f(x_1, x_2, x_3, ..., x_n) dx_1 dx_2 dx_3 ... dx_n$$

where $f(x_1, x_2, x_3, ..., x_n)$ is the joint probability density function. **Problem** Consider again our example of randomly choosing a point in $[0,1] \times [0,1]$.

If $X_1, X_2, X_3, ... X_n$ are random variables and $Y = r(X_1, X_2, X_3, ... X_n)$ then

$$E(Y) = \int \int \int ... \int r(x_1, x_2, x_3, ..., x_n) f(x_1, x_2, x_3, ..., x_n) dx_1 dx_2 dx_3 ... dx_n$$

where $f(x_1, x_2, x_3, ..., x_n)$ is the joint probability density function. **Problem** Consider again our example of randomly choosing a point in $[0,1] \times [0,1]$. We could let X be the random variable of choosing the first coordinate and Y the second. What is E(X+Y)? (note that f(x,y)=1.)

If $X_1, X_2, X_3, ... X_n$ are random variables and $Y = r(X_1, X_2, X_3, ... X_n)$ then

$$E(Y) = \int \int \int ... \int r(x_1, x_2, x_3, ..., x_n) f(x_1, x_2, x_3, ..., x_n) dx_1 dx_2 dx_3 ... dx_n$$

where $f(x_1, x_2, x_3, ..., x_n)$ is the joint probability density function. **Problem** Consider again our example of randomly choosing a point in $[0,1] \times [0,1]$. We could let X be the random variable of choosing the first coordinate and Y the second. What is E(X+Y)? (note that f(x,y)=1.)

Easy properties of expected values:

If
$$Pr(X \ge a) = 1$$
 then $E(X) \ge a$.
If $Pr(X \le b) = 1$ then $E(X) \le b$.



A little more surprising (but not hard and we have already used):

$$E(X_1 + X_2 + X_3 + ... + X_n) = E(X_1) + E(X_2) + E(X_3) + ... + E(X_n).$$

A little more surprising (but not hard and we have already used):

$$E(X_1 + X_2 + X_3 + ... + X_n) = E(X_1) + E(X_2) + E(X_3) + ... + E(X_n).$$

Another way to look at binomial random variables; Let X_i be 1 if the i^{th} trial is a success and 0 if a failure.

A little more surprising (but not hard and we have already used):

$$E(X_1 + X_2 + X_3 + ... + X_n) = E(X_1) + E(X_2) + E(X_3) + ... + E(X_n).$$

Another way to look at binomial random variables; Let X_i be 1 if the i^{th} trial is a success and 0 if a failure. Note that $E(X_i) = 0 \cdot q + 1 \cdot p = p$.

A little more surprising (but not hard and we have already used):

$$E(X_1 + X_2 + X_3 + ... + X_n) = E(X_1) + E(X_2) + E(X_3) + ... + E(X_n).$$

Another way to look at binomial random variables;

Let X_i be 1 if the i^{th} trial is a success and 0 if a failure. Note that $E(X_i) = 0 \cdot q + 1 \cdot p = p$.

Our binomial variable (the number of successes) is

$$X = X_1 + X_2 + X_3 + ... + X_n$$
 so

$$E(X) = E(X_1) + E(X_2) + E(X_3) + ... + E(X_n) = np.$$

A little more surprising (but not hard and we have already used):

$$E(X_1 + X_2 + X_3 + ... + X_n) = E(X_1) + E(X_2) + E(X_3) + ... + E(X_n).$$

Another way to look at binomial random variables;

Let X_i be 1 if the i^{th} trial is a success and 0 if a failure. Note that $E(X_i) = 0 \cdot q + 1 \cdot p = p$.

Our binomial variable (the number of successes) is

$$X = X_1 + X_2 + X_3 + ... + X_n$$
 so

$$E(X) = E(X_1) + E(X_2) + E(X_3) + ... + E(X_n) = np.$$

What about products?



A little more surprising (but not hard and we have already used):

$$E(X_1 + X_2 + X_3 + ... + X_n) = E(X_1) + E(X_2) + E(X_3) + ... + E(X_n).$$

Another way to look at binomial random variables;

Let X_i be 1 if the i^{th} trial is a success and 0 if a failure. Note that $E(X_i) = 0 \cdot q + 1 \cdot p = p$.

Our binomial variable (the number of successes) is

$$X = X_1 + X_2 + X_3 + ... + X_n$$
 so

$$E(X) = E(X_1) + E(X_2) + E(X_3) + ... + E(X_n) = np.$$

What about products? Only works out well if the random variables are **independent**. If $X_1, X_2, X_3, ... X_n$ are independent random variables then:

$$E(\prod_{i=1}^{n} X_i) = \prod_{i=1}^{n} E(X_i).$$

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information!

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information! Also assume $E(X_1^2)=3$, $E(X_2^2)=1$, and $E(X_3^2)=2$.

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information! Also assume $E(X_1^2)=3$, $E(X_2^2)=1$, and $E(X_3^2)=2$.

Facts about Var(X):

• Var(X) = 0 means the same as: there is a c such that Pr(X = c) = 1.

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information! Also assume $E(X_1^2)=3$, $E(X_2^2)=1$, and $E(X_3^2)=2$.

- Var(X) = 0 means the same as: there is a c such that Pr(X = c) = 1.
- $\sigma^2(X) = E(X^2) E(X)^2$ (alternative definition)

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information! Also assume $E(X_1^2)=3$, $E(X_2^2)=1$, and $E(X_3^2)=2$.

- Var(X) = 0 means the same as: there is a c such that Pr(X = c) = 1.
- $\sigma^2(X) = E(X^2) E(X)^2$ (alternative definition)
- $\sigma^2(aX + b) = a^2\sigma^2(X)$.

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information! Also assume $E(X_1^2)=3$, $E(X_2^2)=1$, and $E(X_3^2)=2$.

- Var(X) = 0 means the same as: there is a c such that Pr(X = c) = 1.
- $\sigma^2(X) = E(X^2) E(X)^2$ (alternative definition)
- $\sigma^2(aX + b) = a^2\sigma^2(X)$. Proof: $\sigma^2(aX + b) = E[(aX + b (a\mu + b))^2]$

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information! Also assume $E(X_1^2)=3$, $E(X_2^2)=1$, and $E(X_3^2)=2$.

- Var(X) = 0 means the same as: there is a c such that Pr(X = c) = 1.
- $\sigma^2(X) = E(X^2) E(X)^2$ (alternative definition)
- $\sigma^2(aX + b) = a^2\sigma^2(X)$. Proof: $\sigma^2(aX + b) =$

$$E[(aX+b-(a\mu+b))^2] = E[(aX-a\mu)^2]$$

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information! Also assume $E(X_1^2)=3$, $E(X_2^2)=1$, and $E(X_3^2)=2$.

- Var(X) = 0 means the same as: there is a c such that Pr(X = c) = 1.
- $\sigma^2(X) = E(X^2) E(X)^2$ (alternative definition)
- $\sigma^2(aX + b) = a^2\sigma^2(X)$. Proof: $\sigma^2(aX + b) =$

$$E[(aX+b-(a\mu+b))^2] = E[(aX-a\mu)^2] = a^2E[(X-\mu)^2]$$

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information! Also assume $E(X_1^2) = 3$, $E(X_2^2) = 1$, and $E(X_3^2) = 2$.

- Var(X) = 0 means the same as: there is a c such that Pr(X = c) = 1.
- $\sigma^2(X) = E(X^2) E(X)^2$ (alternative definition)
- $\sigma^2(aX + b) = a^2\sigma^2(X)$. Proof: $\sigma^2(aX + b) =$

$$E[(aX+b-(a\mu+b))^2] = E[(aX-a\mu)^2] = a^2 E[(X-\mu)^2] = a^2 \sigma^2(X).$$

Problem: Consider **independent** random variables X_1 , X_2 , and X_3 , where $E(X_1) = 2$, $E(X_2) = -1$, and $E(X_3) = 0$. Compute $E(X_1^2(X_2 + 3X_3)^2)$.

There is not enough information! Also assume $E(X_1^2)=3$, $E(X_2^2)=1$, and $E(X_3^2)=2$.

- Var(X) = 0 means the same as: there is a c such that Pr(X = c) = 1.
- $\sigma^2(X) = E(X^2) E(X)^2$ (alternative definition)
- $\sigma^2(aX + b) = a^2\sigma^2(X)$. Proof: $\sigma^2(aX + b) = E[(aX + b (a\mu + b))^2] = E[(aX a\mu)^2] = a^2E[(X \mu)^2] = a^2\sigma^2(X)$.
- For independent $X_1, X_2, X_3, ..., X_n$

$$\sigma^{2}(X_{1}+X_{2}+X_{3}+...+X_{n})=\sigma^{2}(X_{1})+\sigma^{2}(X_{2})+\sigma^{2}(X_{3})+...+\sigma^{2}(X_{n}).$$

Note that the last statement tells us that

$$\sigma^{2}(a_{1}X_{1} + a_{2}X_{2} + a_{3}X_{3} + \dots + a_{n}X_{n}) =$$

$$= a_{1}^{2}\sigma^{2}(X_{1}) + a_{2}^{2}\sigma^{2}(X_{2}) + a_{3}^{2}\sigma^{2}(X_{3}) + \dots + a_{n}^{2}\sigma^{2}(X_{n}).$$

Note that the last statement tells us that

$$\sigma^{2}(a_{1}X_{1} + a_{2}X_{2} + a_{3}X_{3} + \dots + a_{n}X_{n}) =$$

$$= a_{1}^{2}\sigma^{2}(X_{1}) + a_{2}^{2}\sigma^{2}(X_{2}) + a_{3}^{2}\sigma^{2}(X_{3}) + \dots + a_{n}^{2}\sigma^{2}(X_{n}).$$

Now we can compute the variance of the binomial distribution with parameters n and p.

Note that the last statement tells us that

$$\sigma^{2}(a_{1}X_{1} + a_{2}X_{2} + a_{3}X_{3} + \dots + a_{n}X_{n}) =$$

$$= a_{1}^{2}\sigma^{2}(X_{1}) + a_{2}^{2}\sigma^{2}(X_{2}) + a_{3}^{2}\sigma^{2}(X_{3}) + \dots + a_{n}^{2}\sigma^{2}(X_{n}).$$

Now we can compute the variance of the binomial distribution with parameters n and p. As before $X = X_1 + X_2 + ... + X_n$ where the X_i are independent with $Pr(X_i = 0) = q$ and $Pr(X_i = 1) = p$.

Note that the last statement tells us that

$$\sigma^{2}(a_{1}X_{1} + a_{2}X_{2} + a_{3}X_{3} + \dots + a_{n}X_{n}) =$$

$$= a_{1}^{2}\sigma^{2}(X_{1}) + a_{2}^{2}\sigma^{2}(X_{2}) + a_{3}^{2}\sigma^{2}(X_{3}) + \dots + a_{n}^{2}\sigma^{2}(X_{n}).$$

Now we can compute the variance of the binomial distribution with parameters n and p. As before $X = X_1 + X_2 + ... + X_n$ where the X_i are independent with $Pr(X_i = 0) = q$ and $Pr(X_i = 1) = p$. So $\mu(X_i) = p$

Note that the last statement tells us that

$$\sigma^{2}(a_{1}X_{1} + a_{2}X_{2} + a_{3}X_{3} + \dots + a_{n}X_{n}) =$$

$$= a_{1}^{2}\sigma^{2}(X_{1}) + a_{2}^{2}\sigma^{2}(X_{2}) + a_{3}^{2}\sigma^{2}(X_{3}) + \dots + a_{n}^{2}\sigma^{2}(X_{n}).$$

Now we can compute the variance of the binomial distribution with parameters n and p. As before $X=X_1+X_2+...+X_n$ where the X_i are independent with $Pr(X_i=0)=q$ and $Pr(X_i=1)=p$. So $\mu(X_i)=p$ and $\sigma^2(X_i)=E(X_i^2)-\mu(X_i)^2=$

Note that the last statement tells us that

$$\sigma^{2}(a_{1}X_{1} + a_{2}X_{2} + a_{3}X_{3} + \dots + a_{n}X_{n}) =$$

$$= a_{1}^{2}\sigma^{2}(X_{1}) + a_{2}^{2}\sigma^{2}(X_{2}) + a_{3}^{2}\sigma^{2}(X_{3}) + \dots + a_{n}^{2}\sigma^{2}(X_{n}).$$

Now we can compute the variance of the binomial distribution with parameters n and p. As before $X=X_1+X_2+...+X_n$ where the X_i are independent with $Pr(X_i=0)=q$ and $Pr(X_i=1)=p$. So $\mu(X_i)=p$ and $\sigma^2(X_i)=E(X_i^2)-\mu(X_i)^2=p-p^2=p(1-p)=pq$.

Note that the last statement tells us that

$$\sigma^{2}(a_{1}X_{1} + a_{2}X_{2} + a_{3}X_{3} + \dots + a_{n}X_{n}) =$$

$$= a_{1}^{2}\sigma^{2}(X_{1}) + a_{2}^{2}\sigma^{2}(X_{2}) + a_{3}^{2}\sigma^{2}(X_{3}) + \dots + a_{n}^{2}\sigma^{2}(X_{n}).$$

Now we can compute the variance of the binomial distribution with parameters n and p. As before $X=X_1+X_2+...+X_n$ where the X_i are independent with $Pr(X_i=0)=q$ and $Pr(X_i=1)=p$. So $\mu(X_i)=p$ and $\sigma^2(X_i)=E(X_i^2)-\mu(X_i)^2=p-p^2=p(1-p)=pq$. Thus $\sigma^2(X)=\Sigma\sigma^2(X_i)=npq$.

Consider our random variable Z which is the sum of the coordinates of a point randomly chosen from $[0,1] \times [0,1]$.

Consider our random variable Z which is the sum of the coordinates of a point randomly chosen from $[0,1] \times [0,1]$. Z = X + Y where X and Y both represent choosing a point randomly from [0,1].

Consider our random variable Z which is the sum of the coordinates of a point randomly chosen from $[0,1] \times [0,1]$. Z = X + Y where X and Y both represent choosing a point randomly from [0,1]. They are independent and last time we showed $\sigma^2(X) = \frac{1}{12}$.

Consider our random variable Z which is the sum of the coordinates of a point randomly chosen from $[0,1] \times [0,1]$. Z = X + Y where X and Y both represent choosing a point randomly from [0,1]. They are independent and last time we showed $\sigma^2(X) = \frac{1}{12}$. So $\sigma^2(Z) = \frac{1}{6}$.

Consider our random variable Z which is the sum of the coordinates of a point randomly chosen from $[0,1] \times [0,1]$. Z = X + Y where X and Y both represent choosing a point randomly from [0,1]. They are independent and last time we showed $\sigma^2(X) = \frac{1}{12}$. So $\sigma^2(Z) = \frac{1}{6}$.

What is E(XY)?

Consider our random variable Z which is the sum of the coordinates of a point randomly chosen from $[0,1] \times [0,1]$. Z = X + Y where X and Y both represent choosing a point randomly from [0,1]. They are independent and last time we showed $\sigma^2(X) = \frac{1}{12}$. So $\sigma^2(Z) = \frac{1}{6}$.

What is E(XY)? They are independent so $E(XY) = E(X)E(Y) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$.

Consider our random variable Z which is the sum of the coordinates of a point randomly chosen from $[0,1] \times [0,1]$. Z = X + Y where X and Y both represent choosing a point randomly from [0,1]. They are independent and last time we showed $\sigma^2(X) = \frac{1}{12}$. So $\sigma^2(Z) = \frac{1}{6}$.

What is E(XY)? They are independent so $E(XY) = E(X)E(Y) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$. $(\sigma^2(XY))$ is more complicated.)

$$E(\sum_{i=1}^{n} X_i) = \int \int ... \int (x_1 + x_2 + ... + x_n) f(x_1, x_2, ..., x_n) dx_1 dx_2 ... dx_n$$

$$E(\sum_{i=1}^{n} X_i) = \int \int ... \int (x_1 + x_2 + ... + x_n) f(x_1, x_2, ..., x_n) dx_1 dx_2 ... dx_n$$
$$= \sum_{i=1}^{n} \int \int ... \int x_i f(x_1, x_2, ..., x_n) dx_1 dx_2 ... dx_n.$$

$$E(\sum_{i=1}^{n} X_i) = \int \int ... \int (x_1 + x_2 + ... + x_n) f(x_1, x_2, ..., x_n) dx_1 dx_2 ... dx_n$$

$$= \sum_{i=1}^{n} \int \int ... \int x_i f(x_1, x_2, ..., x_n) dx_1 dx_2 ... dx_n.$$

$$= \sum_{i=1}^{n} \int x_i f_i(x_i) dx_i = \sum_{i=1}^{n} E(X_i).$$