Chile_prev_rmd

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```
#source("Chile_prev.R", local = knitr::knit_global())
library(nleqslv) # Only needed for robince bayesian prevalence
library(janitor)
library(gridExtra)
library(readxl)
library(psych)
library(Hmisc)
## Warning: package 'ggplot2' was built under R version 4.2.3
library(poolr)
## Warning: package 'poolr' was built under R version 4.2.3
library(epitools)
library(corrplot)
library(caret)
library(mltools)
## Warning: package 'mltools' was built under R version 4.2.3
library(ggrepel)
## Warning: package 'ggrepel' was built under R version 4.2.3
library(rjags)
library(rstan)
library(posterior)
library(tidybayes)
library(bayesplot)
library(tidyverse)
```

Bayesian prevalence analysis of autism prevalence in Chile

Load data

```
school_name = nom_rbd,
         school_region_code = cod_reg_rbd,
         school_region_name_abr = nom_reg_rbd_a,
         school_province_code = cod_pro_rbd,
         school commune code = cod com rbd,
         school_commune_name = nom_com_rbd,
         school_dept_code = cod_deprov_rbd,
         school_dept_name = nom_deprov_rbd,
         school dependency code = cod depe, # has categories 1-6, no1 and no2 here are no1 in grouped
         school_dependency_code_grouped = cod_depe2, # has categories 1-5
         school_rurality_code = rural_rbd,
         school_operation_status = estado_estab,
         teaching_code1 = cod_ense, # min = 10, max = 910, eg preschool, special education hearing impa
         teaching_code2 = cod_ense2, # subject matter coding, 1-8
         teaching_code3 = cod_ense3, # age based coding, 1-7
         grade_code1 = cod_grado, # grade of schooling, 1-10, 21-25, 31-34, nests in teaching_code1
         grade_code2 = cod_grado2, # equivalent grade of schooling for adult special education, 1-8, 99
         grade_letter = let_cur, # refers to the class within the grade, close to start of alphabet is
         course_timing = cod_jor, # time of day, morning, afternoon, both, night, no info
         course_type = cod_tip_cur, # 0 = simple course, 1-4 = combined course, 99 = no info
         course_descr = cod_des_cur, # Description of course (TP secondary education only). O: Does not
         student id = mrun,
         sex = gen_alu, # 0 = no info, 1 = male, 2 = female
         dob = fec_nac_alu,
         age_june30 = edad_alu, # age at 30th June 2021
         special needs status = int alu, # integrated student indicator, O = no, 1 = yes. Mostly no
         special_needs_code = cod_int_alu, # ADHD, blindness, etc. 0 = none. 105 = autism, 203 = ADHD.
         student_region_code = cod_reg_alu,
         student_commune_code = cod_com_alu,
         student_commune_name = nom_com_alu,
         economic_sector_code = cod_sec,
         economic_specialty_code = cod_espe,
         economic_branch_code = cod_rama,
         economic_profspec_code = cod_men,
         teaching code new = ens)
chile_stdpop_raw <- read_excel("04_Data/pop_chile_2021_single_age.xlsx") %>%
  clean names()
chile_stdpop <- chile_stdpop_raw %>%
  filter(sex != 9) %>%
  rename("std_pop" = "pop_2021") %>%
  mutate(pop_prop = std_pop / sum(std_pop))
```

Try Bayesian analysis of autism prevalence and specificity and sensitivity of school assessment "Bayesian Estimation of Disease Prevalence and the Parameters of Diagnostic Tests in the Absence of a Gold Standard" Lawrence Joseph, Theresa W. Gyorkos, Louis Coupal https://www.cambridge.org/core/journals/epidemiology-and-psychiatric-sciences/article/bayesian-approach-to-estimating-the-population-prevalence-of-mood-and-anxiety-disorders-using-multiple-measures/DB1D2CA6C27C7E8C85C60B62B969BB72

Use sensitivity and specificity of Social Attention and Communication Surveillance—Revised (SACS-R) tool "Diagnostic Accuracy of the Social Attention and Communication Surveillance—Revised With Preschool Tool for Early Autism Detection in Very Young Children" Josephine Barbaro, Nancy Sadka, Melissa Gilbert, et al https://jamanetwork.com/journals/jamanetworkopen/fullarticle/2789926

```
chile_bayes_aut <- chile_merged %>%
  filter(age_june30 >= 6 & age_june30 <= 18,
         #special_needs_status == 1,
         sex != 0) %>%
  mutate(autism = ifelse(special_needs_code == 105, 1, 0),
         age_cat = ifelse(age_june30 <= 8, 1, ifelse(age_june30 <= 11, 2, ifelse(age_june30 <= 14, 3, 4</pre>
          # 1 = 6-8, 2 = 9-11, 3 = 12-14, 4 = 15-18
         ethnic_2_group = ifelse(ethnic_3_group == "Aymara", "Other ethnic group", ethnic_3_group),
         school_fee = ifelse(school_fee == "", "SIN INFORMACION", school_fee),
         school_fee_group = ifelse(school_fee == "GRATUITO", "Free",
                                  ifelse(school_fee %in% c("$1.000 A $10.000", "$10.001 A $25.000", "$2
                                          ifelse(school_fee == "MAS DE $100.000", "High", "No information
  select(school_region_name_abr,
   sex,
   sex_desc,
   age_june30,
    #edad_alu_2, # equal to age_june30
   age_cat,
   school_rurality_code,
    #rural_rbd_2, # not quite equal to school_rurality_code as it has NA's
   pago_matricula,
   pago_mensual,
   school_fee,
   school_fee_group,
   ethnicity,
   mapuche,
   nationality,
   ethnic_3_group,
   ethnic_2_group,
    #asd_chile, # equal to autism
    autism
  )
# Prevalence of autism in Chile dataset
sum(chile_bayes_aut$autism) / nrow(chile_bayes_aut) # 0.00476 = 0.476%, very low
## [1] 0.004760322
# Is prevalence the same across geographic regions, age, sex?
n std pop <- sum(chile stdpop$std pop)</pre>
```

Define some functions to keep code clean

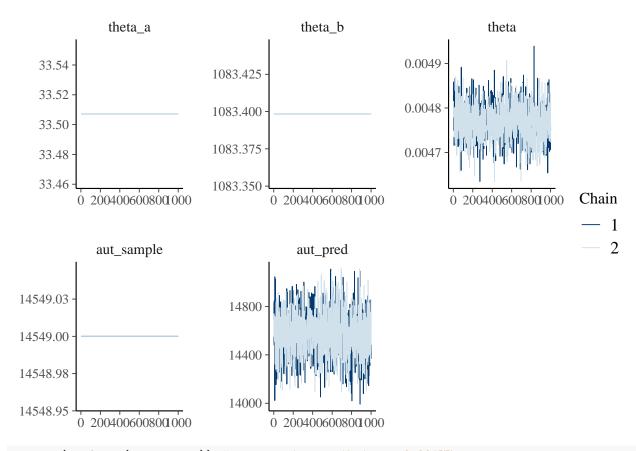
```
get_grouped_prev <- function(x, stdpop, grouping_vars) {
    # Calculates sample prevalence, age- and sex-standardised prevalence and group weighting for supplied
    # x = chile_bayes_aut
    # stdpop = standard population with age and sex counts
    # grouping_vars = variables in x to group by
    x_grouped <- x %>%
        group_by(across(all_of(grouping_vars))) %>%
        summarise(count = n()) %>%
        pivot_wider(names_from = autism, values_from = count) %>%
        rename("n_noautism" = "0", "n_autism" = "1", "age" = "age_june30") %>%
```

```
mutate(n_autism = ifelse(is.na(n_autism), 0, n_autism), # If there are no cases of autism in the gr
           sample_pop_size = n_noautism + n_autism, # Total sample population is autism cases + not cas
           sample_prevalence = n_autism / sample_pop_size) %% # Prevalence of autism in the group
    left_join(stdpop, by = c("age", "sex")) %>%
    mutate(aut_prev_std = n_autism / sample_pop_size * pop_prop, # Prevalence of autism in the group, s
           w = std_pop / (sample_pop_size * n_std_pop), # Weight of the group using standard population
           \#w2 = pop\_prop / sample\_pop\_size,
           \#sum\ std\ pop = sum(std\ pop)
           ) %>%
    ungroup()
  return(x_grouped)
}
get_adjusted_prev <- function(x, grouping_vars) {</pre>
  # Turns grouped prevalences into age- and sex- adjusted prevalences with Fay and Feuer Gamma confiden
  # x = output from get_grouped_prev
  x_adj <- x %>%
  group_by(across(all_of(grouping_vars))) %>%
  summarise(sum_sample_pop_size = sum(sample_pop_size),
            crude_rate = sum(n_autism) / sum(sample_pop_size),
            crude_count = sum(n_autism),
            adjusted_rate = sum(n_autism / sample_pop_size * pop_prop),
            adjusted_count = round(adjusted_rate * sum_sample_pop_size, 0), # had to fudge this to get .
            #adjusted_count = adjusted_rate * sum_sample_pop_size,
            var = sum(pop_prop^2 * n_autism / sample_pop_size^2),
            #se2 = sqrt(sum((std_pop/sum(std_pop))^2 * n_autism/sample_pop_size^2)),
            w M = max(w),
            ci_lower = var / (2*adjusted_rate) * qchisq(p = 0.05/2, df = 2*adjusted_rate^2 / var),
            ci_upper = (var + w_M^2) / (2*(adjusted_rate + w_M)) * qchisq(p = 1-0.05/2, df = 2*(adjusted_rate + w_M))
  arrange(across(all_of(grouping_vars)))
}
Set global parameters
nObs <- nrow(chile_bayes_aut)
nIter <- 1000
nBurn <- 1000
pars <- c("theta_a", "theta_b", "theta", "aut_sample", "aut_pred")</pre>
theta_mu_prior \leftarrow 0.0046
theta_sigma_prior \leftarrow (0.0047-0.0045) / (2*1.96)
theta_mu_sens <- c(0.001, 0.005, 0.01, 0.02, # 0.1%, 0.5%, 1%, 2% prevalence
              rep(0.0046, 4)) # Same as chosen prior
theta_sigma_sens <- c(rep(0.001/1.96, 4), # Same as chosen prior
                 0.0001, 0.001, 0.05, 0.01) # +/- 0.1%, 0.5%, 1%, 5%
theta_mu <- c(theta_mu_prior, theta_mu_sens)</pre>
theta_sigma <- c(theta_sigma_prior, theta_sigma_sens)</pre>
theta_a <- theta_mu * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
theta b <- (1 - theta mu) * (theta mu * (1-theta mu) / theta sigma^2 - 1)
```

Common effects model with sample prevalence

```
# Uniform prior
theta_a_common <- 1</pre>
```

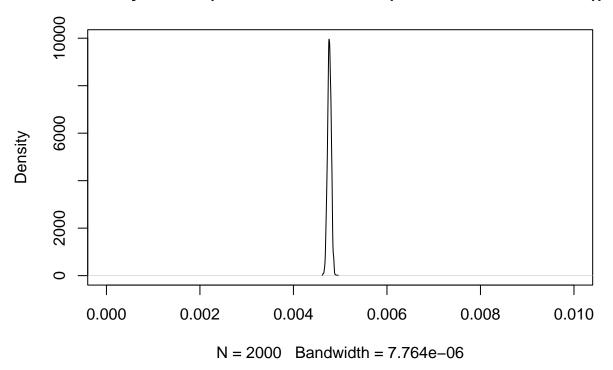
```
theta_b_common <- 1
# This corresponds to a mean of 0.5
# OR Informative prior (global population prevalence)
# Say autism has mean prevalence of 3% and we are 95% confidence that the prevalence is between 2% and .
# Then mu = 0.03, sigma = (0.04-0.02) / (2*1.96)
theta_mu_common <- 0.03
theta_sigma_common <- (0.04-0.02) / (2*1.96)
theta_a_common <- theta_mu_common * (1-theta_mu_common) / theta_sigma_common^2 - 1)
theta_b_common <- (1 - theta_mu_common) * (theta_mu_common * (1-theta_mu_common) / theta_sigma_common^2
common_model <- "model {</pre>
  theta ~ dbeta(theta_a, theta_b) # Prior
  aut_sample ~ dbin(theta, nObs) # Prevalence in sample data
 aut_pred ~ dbin(theta, nObs) # Predicted prevalence in new sample of same size
 #spec ~ dnorm(spec_mu, 1/spec_sd) # dnorm requires prevalence not sd or var
 #sens ~ dnorm(sens mu, 1/sens sd)
 #aut_post <- aut_sample/nObs * sens + (1 - aut_sample/nObs) * spec</pre>
common_data <- list(theta_a = theta_a_common,</pre>
                    theta_b = theta_b_common,
                    n0bs = n0bs,
                    aut_sample = sum(chile_bayes_aut$autism) #,
                    \#spec_mu = 0.996,
                    #spec_sd = (1.00-0.99) / (2*1.96),
                    \#sens_mu = 0.62,
                    \#sens_sd = (0.66-0.57) / (2*1.96)
common_ini \leftarrow list(list(theta = 0.001), #, spec = 0.5, sens = 0.5),
                   list(theta = 0.01)) #, spec = 0.9, sens = 0.9)
common_pars <- c("theta_a", "theta_b", "theta",</pre>
                 #"spec", "sens",
                 "aut_sample", "aut_pred")
# Run JAGS model and discard burn-in samples
common_jag <- jags.model(textConnection(common_model),</pre>
                         data = common_data,
                         inits = common_ini,
                         n.chains = 2,
                         quiet = TRUE)
update(common_jag, n.iter = nBurn)
common_sam <- coda.samples(model = common_jag,</pre>
                           variable.names = common_pars,
                           n.iter = nIter)
# Check for convergence in parameters of interest
```



summary(as_draws(common_sam)) # mean posterior theta is 0.00477

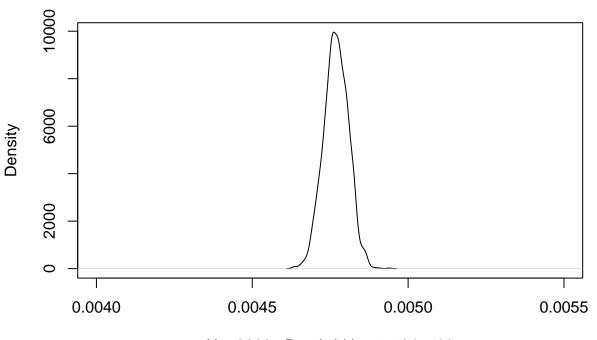
```
## # A tibble: 5 x 10
                                                           q95 rhat ess_b~1 ess_t~2
##
     variable
                 mean median
                                   sd
                                          mad
                                                    q5
                                                 <dbl>
##
     <chr>
                        <dbl>
                                         <dbl>
                                                                       <dbl>
                                                                               <dbl>
                <dbl>
                                <dbl>
                                                         <dbl> <dbl>
## 1 aut_pred 1.46e+4 1.46e+4 1.74e+2 1.65e+2 1.43e+4 1.48e+4 1.00
                                                                       1500.
                                                                               1761.
## 2 aut_sam~ 1.45e+4 1.45e+4 0
                                      0
                                               1.45e+4 1.45e+4 NA
                                                                         NA
                                                                                 NA
## 3 theta
              4.77e-3 4.77e-3 3.94e-5 3.99e-5 4.70e-3 4.83e-3 1.00
                                                                       1216.
                                                                               1125.
## 4 theta_a 3.35e+1 3.35e+1 0
                                               3.35e+1 3.35e+1 NA
                                                                         NA
                                                                                 NA
                                      0
## 5 theta_b 1.08e+3 1.08e+3 0
                                               1.08e+3 1.08e+3 NA
                                                                         NA
                                                                                 NA
                                      0
## # ... with abbreviated variable names 1: ess_bulk, 2: ess_tail
plot(density(extract_variable(common_sam, "theta")), xlim = c(0,0.01))
```

density.default(x = extract_variable(common_sam, "theta"))

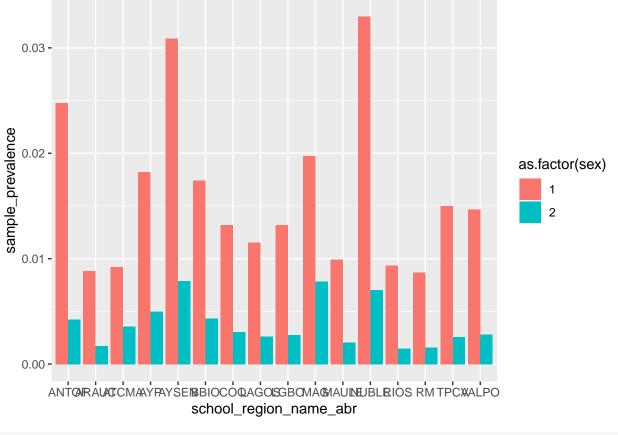


plot(density(extract_variable(common_sam, "theta")), xlim = c(0.004,0.0055))

density.default(x = extract_variable(common_sam, "theta"))



N = 2000 Bandwidth = 7.764e-06



1 is male, 2 is female

Bayesian prevalence analysis

 $Standardise\ prevalence\ by\ Chile's\ age\ and\ sex\ based\ population\ sizes\ using\ https://seer.cancer.gov/seerstat/WebHelp/Rate_Algorithms.htm\ and\ https://wonder.cdc.gov/wonder/help/cancer/fayfeuerconfidenceintervals.pdf$

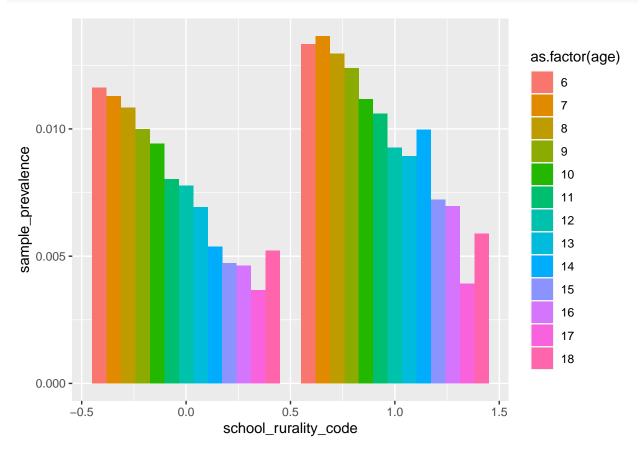
See https://github.com/Dpananos/bayes_multiple_measures/blob/master/analysis/sensitivity_analysis.R for more sensitivity analysis ideas

By region, with sensitivity analysis (alter prior mean and sd)

```
aut_pred[i] ~ dbin(theta[i], nObs[i])
 }
}"
for(j in 1:length(theta_mu)) {
  rand_region_data <- list(theta_a = theta_a[j],</pre>
                           theta_b = theta_b[j],
                           nObs = aut_prev_region_adj$sum_sample_pop_size,
                           aut_sample = aut_prev_region_adj$adjusted_count,
                           nRegion = nRegion)
  rand_region_jag <- jags.model(textConnection(rand_region_model),</pre>
                                 data = rand_region_data,
                                 inits = rand_region_ini,
                                n.chains = 2,
                                quiet = TRUE)
  update(rand_region_jag, n.iter = nBurn)
  rand_region_sam <- coda.samples(model = rand_region_jag,</pre>
                                  variable.names = pars,
                                   n.iter = nIter)
  mcmc_trace(rand_region_sam, paste0("theta[", 1:nRegion, "]")) # Convergence looks fine and rhats <= 1
  mcmc_trace(rand_region_sam, paste0("aut_pred[", 1:nRegion, "]"))# Convergence looks fine and rhats <=</pre>
  rand_region_summ <- summary(subset_draws(as_draws(rand_region_sam), pars),</pre>
                       ~quantile(.x, probs=c(0.025, 0.5, 0.975)),
                       ~mcse_quantile(.x, probs=c(0.025, 0.5, 0.975)),
                       "rhat") %>%
    arrange(desc(mcse_q50))
  rand_region_summ
  # Plot
  aut_prev_region_post <- as_tibble(as_draws_matrix(rand_region_sam), rownames = "Iteration") %>%
    select(c("Iteration", contains("theta["))) %>%
    pivot_longer(cols = contains("theta["),
                 names_to = "Region",
                 values_to = "predicted_prev") %>%
    mutate(school_region_name_abr = factor(Region,
                                            levels = c(paste0("theta[",1:nRegion,"]")),
                                            labels = RegionNames))
  aut_prev_region_ci <- aut_prev_region_post %>%
    group_by(school_region_name_abr) %>%
    summarise(post_lower = quantile(predicted_prev, 0.025),
              post_upper = quantile(predicted_prev, 0.975))
  (ggplot() +
    geom_density(data = aut_prev_region_post, aes(x = predicted_prev)) +
    geom_vline(data = aut_prev_region_ci, aes(xintercept = post_lower), color = "blue", linetype = "dot
    geom_vline(data = aut_prev_region_ci, aes(xintercept = post_upper), color = "blue", linetype = "dot
    geom_vline(data = aut_prev_region_adj, aes(xintercept = ci_lower), color = "red", linetype = "dashe
    geom_vline(data = aut_prev_region_adj, aes(xintercept = ci_upper), color = "red", linetype = "dashe
    facet_wrap(~school_region_name_abr))
}
```

Bayesian prevalence by rurality

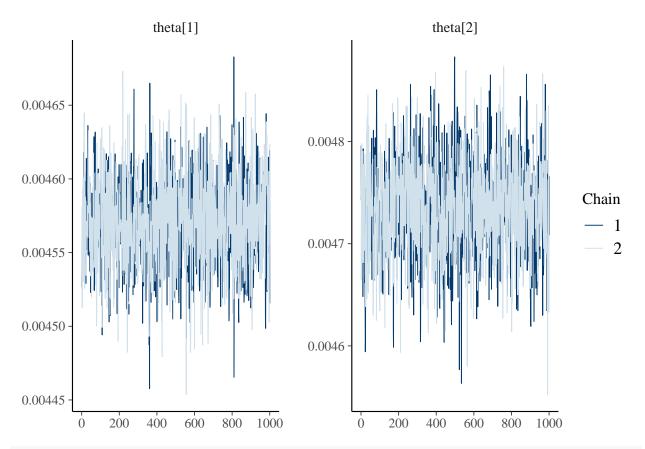
```
## `summarise()` has grouped output by 'school_rurality_code', 'age_june30',
## 'sex'. You can override using the `.groups` argument.
ggplot(data = aut_prev_rural) +
   geom_col(aes(x = school_rurality_code, y = sample_prevalence, group = age, fill = as.factor(age)), po
```



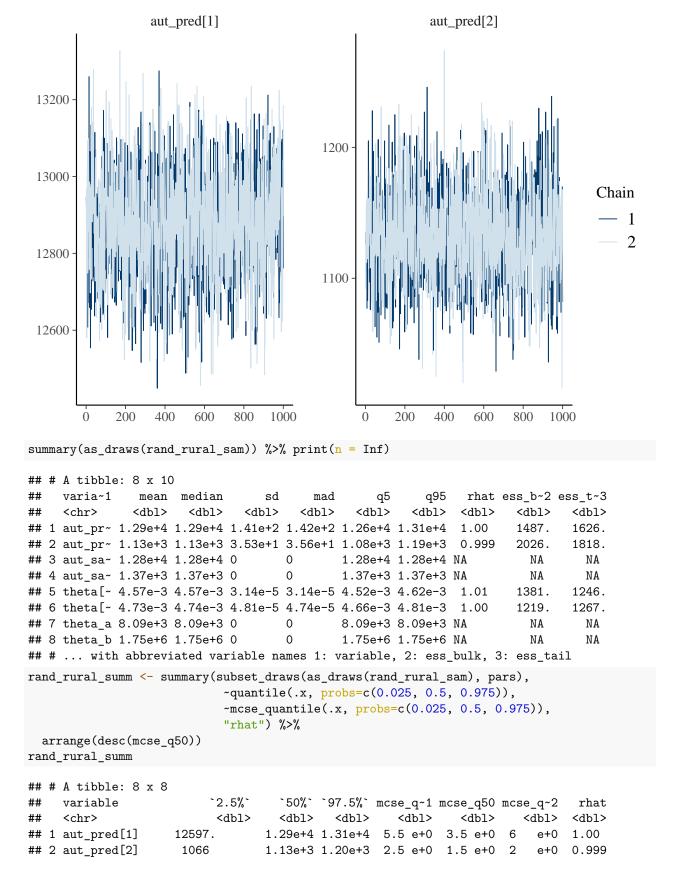
 $\#geom_col(aes(x = school_region_name_abr, y = prevalence, group = sex, fill = as.factor(sex)), position \# 1 is male, 2 is female$

```
aut_prev_rural_adj <- aut_prev_rural %>%
  group_by(school_rurality_code) %>%
  summarise(sum_sample_pop_size = sum(sample_pop_size),
            crude_rate = sum(n_autism) / sum(sample_pop_size),
            crude_count = sum(n_autism),
            adjusted_rate = sum(n_autism / sample_pop_size * pop_prop),
            adjusted_count = round(adjusted_rate * sum_sample_pop_size, 0), # had to fudge this to get .
            #adjusted count = adjusted rate * sum sample pop size,
            var = sum(pop_prop^2 * n_autism / sample_pop_size^2),
            #se2 = sqrt(sum((std_pop/sum(std_pop))^2 * n_autism/sample_pop_size^2)),
            w M = max(w),
            ci_lower = var / (2*adjusted_rate) * qchisq(p = 0.05/2, df = 2*adjusted_rate^2 / var),
            ci_upper = (var + w_M^2) / (2*(adjusted_rate + w_M)) * qchisq(p = 1-0.05/2, df = 2*(adjusted_rate + w_M))
  arrange(school_rurality_code)
# Prior: age and sex standardised prevalence in the whole Chile dataset
theta_mu \leftarrow 0.0046
theta_sigma \leftarrow (0.0047-0.0045) / (2*1.96)
theta_a <- theta_mu * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
theta_b <- (1 - theta_mu) * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
nRural <- length(unique(aut_prev_rural$school_rurality_code))</pre>
rand_rural_model <- "model {</pre>
 for(i in 1:nRural) { # For each rurality
    theta[i] ~ dbeta(theta_a, theta_b)
    aut_sample[i] ~ dbin(theta[i], nObs[i])
    aut_pred[i] ~ dbin(theta[i], nObs[i])
  }
}"
rand_rural_data <- list(theta_a = theta_a,</pre>
                          theta_b = theta_b,
                          nObs = aut_prev_rural_adj$sum_sample_pop_size,
                          aut_sample = aut_prev_rural_adj$adjusted_count,
                          nRural = nRural)
\#rand\_rural\_ini \leftarrow list(list(theta = 0.001), \#, spec = 0.5, sens = 0.5),
                     list(theta = 0.01)) #, spec = 0.9, sens = 0.9))
# Run JAGS model and discard burn-in samples
rand_rural_jag <- jags.model(textConnection(rand_rural_model),</pre>
                               data = rand_rural_data,
                               #inits = rand_region_ini,
                               n.chains = 2,
                               quiet = TRUE)
update(rand_rural_jag, n.iter = nBurn)
rand_rural_sam <- coda.samples(model = rand_rural_jag,</pre>
                                 variable.names = pars,
                                 n.iter = nIter)
```

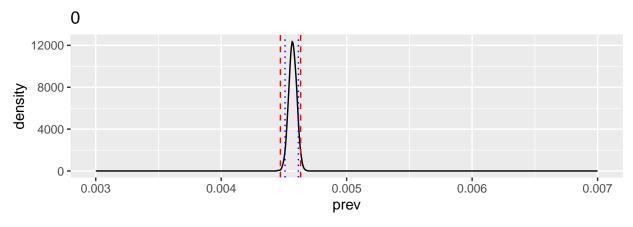
mcmc_trace(rand_rural_sam, paste0("theta[", 1:nRural, "]")) # Convergence looks fine and rhats <= 1.1</pre>

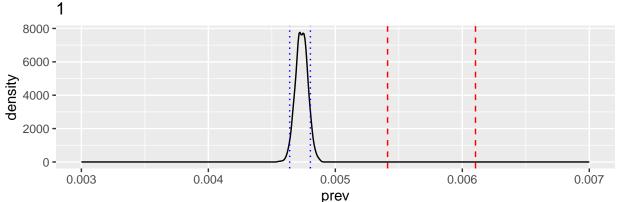


mcmc_trace(rand_rural_sam, paste0("aut_pred[", 1:nRural, "]"))# Convergence looks fine and rhats <= 1.1</pre>



```
## 3 theta[2]
                          0.00464 4.74e-3 4.83e-3 3.59e-6 1.79e-6 4.38e-6 1.00
## 4 theta[1]
                          0.00451 4.57e-3 4.63e-3 3.52e-6 9.47e-7 2.40e-6 1.01
## 5 theta a
                                   8.09e+3 8.09e+3 NA
                      8091.
                                                             NA
                                                                      NA
                                                                                NA
                                   1.75e+6 1.75e+6 NA
                                                                                NΑ
## 6 theta_b
                    1750915.
                                                             NA
                                                                      NA
                                   1.28e+4 1.28e+4 NA
## 7 aut_sample[1]
                     12823
                                                             NA
                                                                      NA
                                                                                NA
## 8 aut sample[2]
                       1370
                                   1.37e+3 1.37e+3 NA
                                                             NA
                                                                      NA
                                                                                NA
## # ... with abbreviated variable names 1: mcse_q2.5, 2: mcse_q97.5
aut prev rural plots <- list()</pre>
rural_post_ci_lower <- list()</pre>
rural_post_ci_upper <- list()</pre>
for(i in 1:nRural) {
  prevs <- data.frame(prev = extract_variable(rand_rural_sam, paste0("theta[", i, "]")))</pre>
  rural_post_ci_lower[[i]] <- quantile(prevs$prev, 0.025)</pre>
  rural_post_ci_upper[[i]] <- quantile(prevs$prev, 0.925)</pre>
  density_plot <- ggplot(prevs, aes(x = prev), color = "blue") +</pre>
    geom_density() +
    xlim(c(0.003, 0.007)) +
    geom_vline(xintercept = rural_post_ci_lower[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = rural_post_ci_upper[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = aut_prev_rural_adj$ci_lower[i], color = "red", linetype = "dashed") +
    geom_vline(xintercept = aut_prev_rural_adj$ci_upper[i], color = "red", linetype = "dashed") +
    labs(title = aut_prev_rural_adj$school_rurality_code[i])
  aut_prev_rural_plots[[i]] <- density_plot</pre>
do.call(grid.arrange, aut_prev_rural_plots)
```





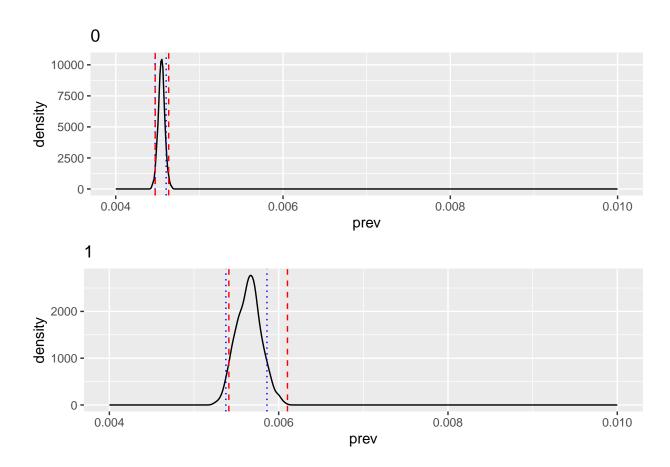
```
#autism_prev_rural_plots <- do.call(grid.arrange, aut_prev_rural_plots)
#ggsave("autism_prev_rural_plots.png", autism_prev_rural_plots, height = 10, width = 15)</pre>
```

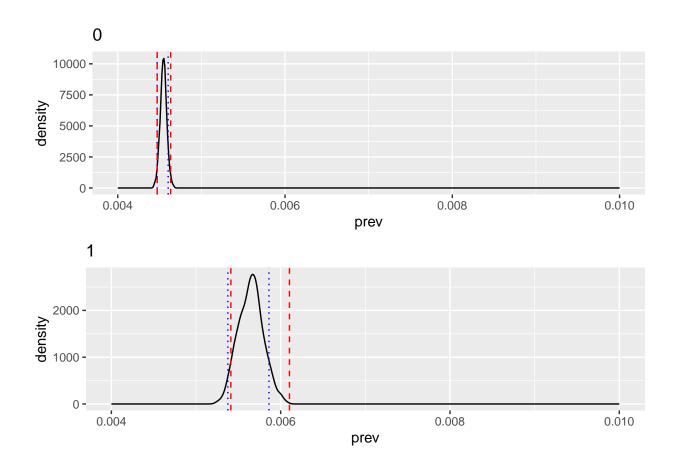
Assuming 0 = city, 1 = rural. Narrower CI for city because sample size is bigger

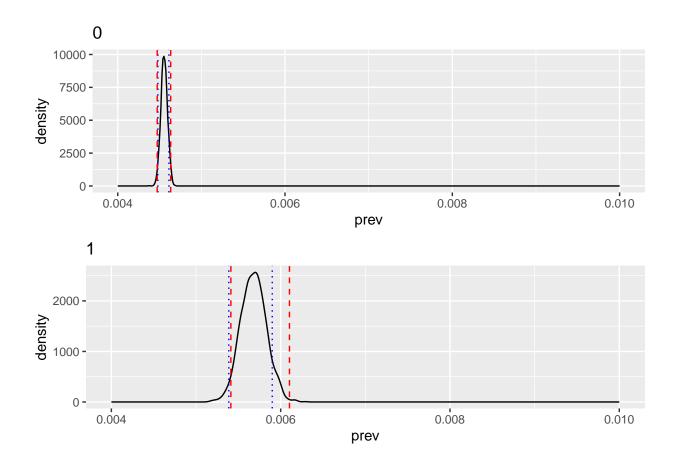
Sensitivity analysis - alter prior mean and sd

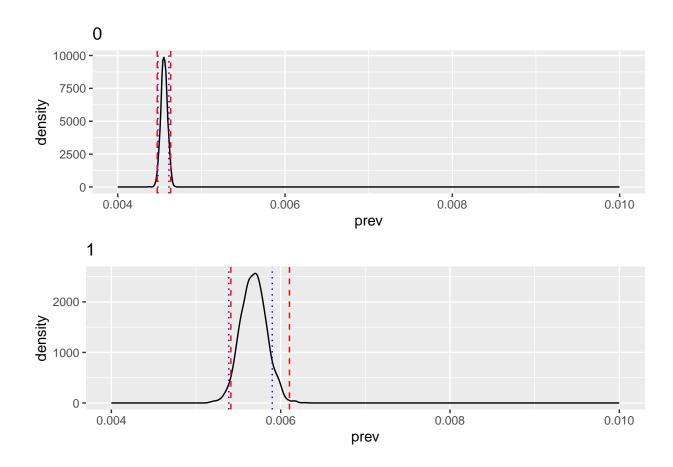
```
theta_mu <- c(0.001, 0.005, 0.01, 0.02, # 0.1\%, 0.5\%, 1\%, 2\% prevalence
              rep(0.0046, 4)) # Same as chosen prior
theta_sigma <- c(rep(0.001/1.96, 4), # Same as chosen prior
                 0.0001, 0.001, 0.05, 0.01) # +/- 0.1%, 0.5%, 1%, 5%
theta_a <- theta_mu * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
theta b <- (1 - theta mu) * (theta mu * (1-theta mu) / theta sigma^2 - 1)
for(j in 1:length(theta_mu)) {
  rand_rural_data <- list(theta_a = theta_a[j],</pre>
                          theta_b = theta_b[j],
                           nObs = aut_prev_rural_adj$sum_sample_pop_size,
                           aut_sample = aut_prev_rural_adj$adjusted_count,
                          nRural = nRural)
  rand_rural_jag <- jags.model(textConnection(rand_rural_model),</pre>
                                data = rand_rural_data,
                                #inits = rand_region_ini,
                                n.chains = 2,
                                quiet = TRUE)
  update(rand_rural_jag, n.iter = nBurn)
```

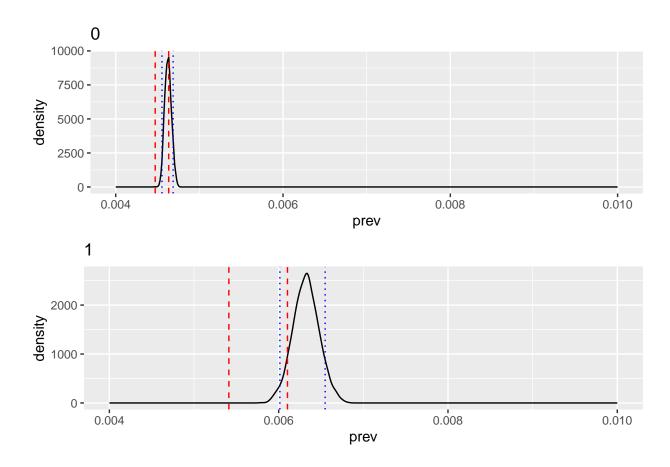
```
rand_rural_sam <- coda.samples(model = rand_rural_jag,</pre>
                                variable.names = pars,
                                n.iter = nIter)
# Plots
aut_prev_rural_plots <- list()</pre>
rural_post_ci_lower <- list()</pre>
rural post ci upper <- list()</pre>
for(i in 1:nRural) {
  prevs <- data.frame(prev = extract_variable(rand_rural_sam, paste0("theta[", i, "]")))</pre>
  rural_post_ci_lower[[i]] <- quantile(prevs$prev, 0.025)</pre>
  rural_post_ci_upper[[i]] <- quantile(prevs$prev, 0.925)</pre>
  density_plot <- ggplot(prevs, aes(x = prev), color = "blue") +</pre>
    geom_density() +
    xlim(c(0.004, 0.01)) +
    geom_vline(xintercept = rural_post_ci_lower[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = rural_post_ci_upper[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = aut_prev_rural_adj$ci_lower[i], color = "red", linetype = "dashed") +
    geom_vline(xintercept = aut_prev_rural_adj$ci_upper[i], color = "red", linetype = "dashed") +
    labs(title = aut_prev_rural_adj$school_rurality_code[i])
  aut_prev_rural_plots[[i]] <- density_plot</pre>
do.call(grid.arrange, aut_prev_rural_plots)
autism_prev_rural_plots <- do.call(grid.arrange, aut_prev_rural_plots)</pre>
ggsave(paste0("autism_prev_rural_plots_", j, ".png"), autism_prev_rural_plots, height = 10, width = 1
```

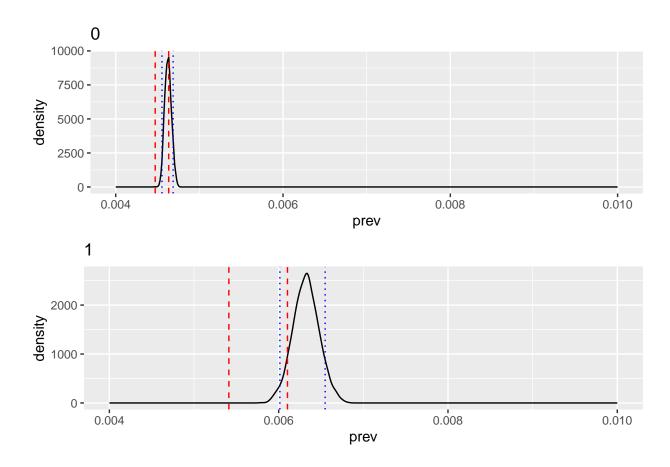


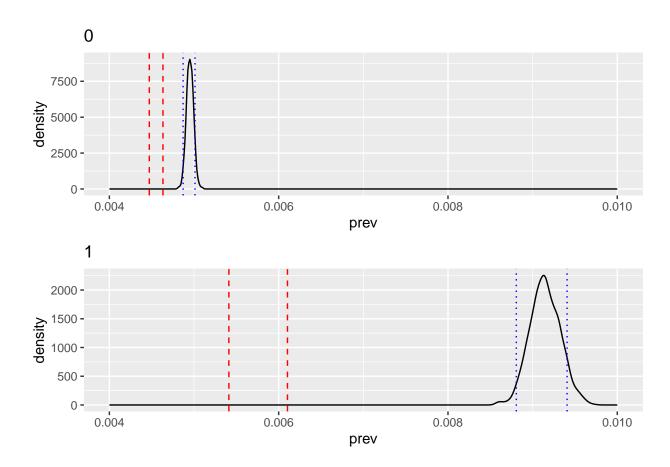


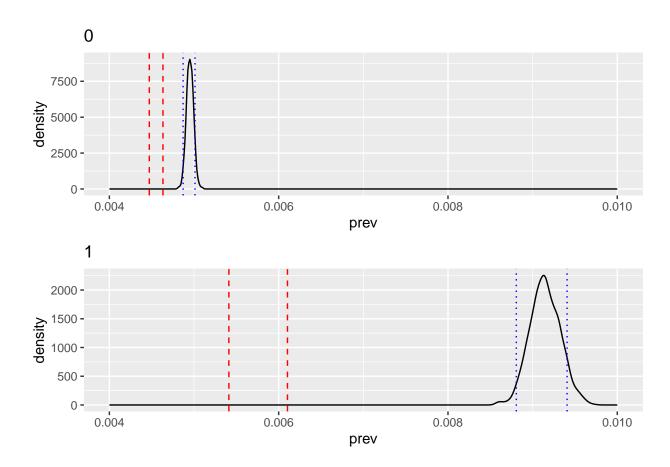


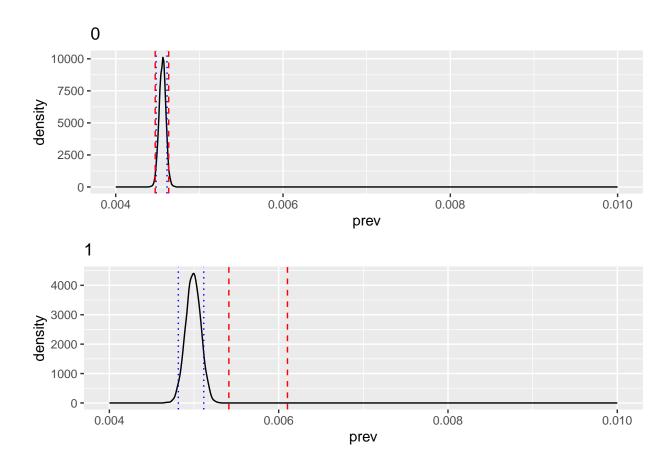


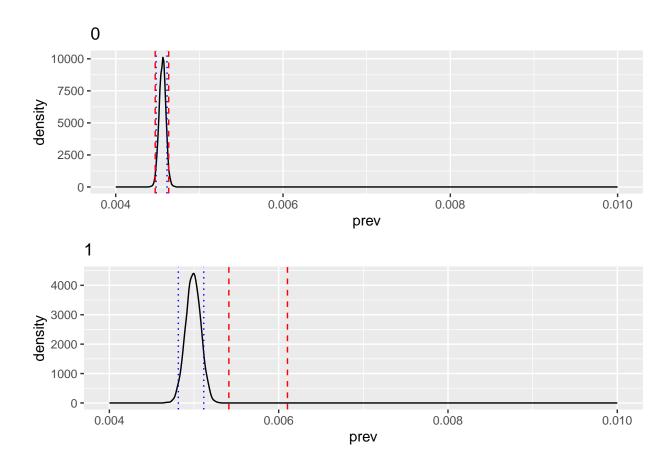


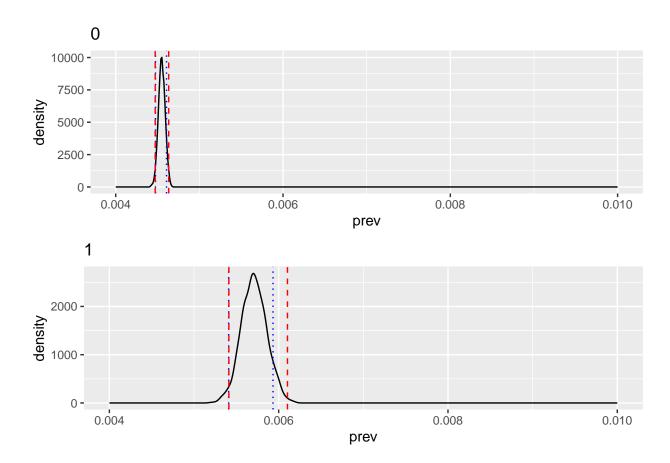


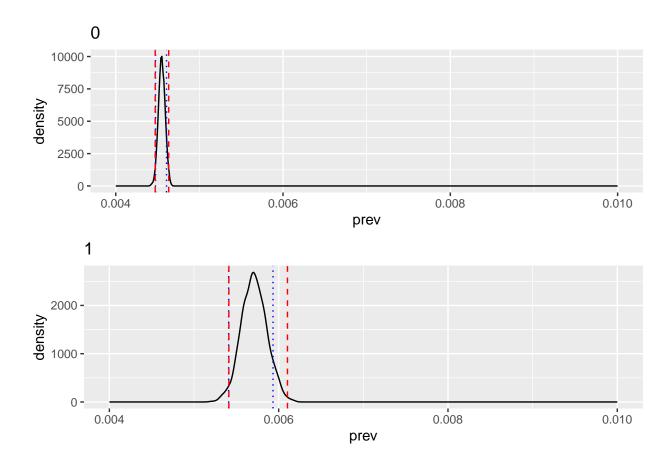


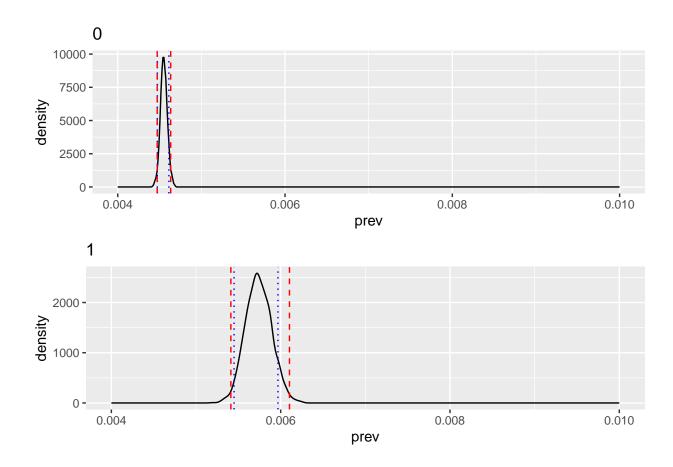


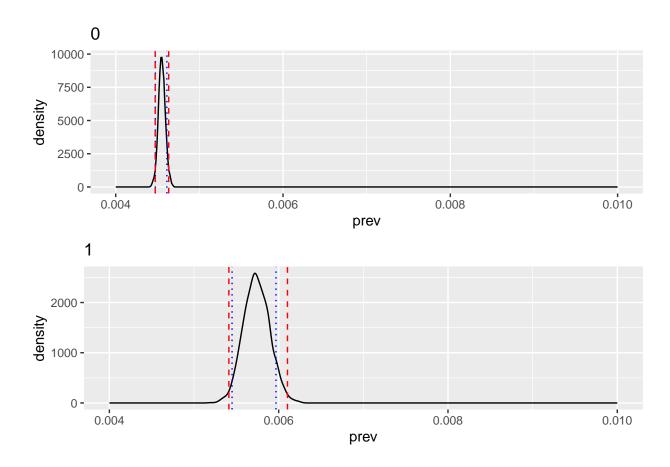


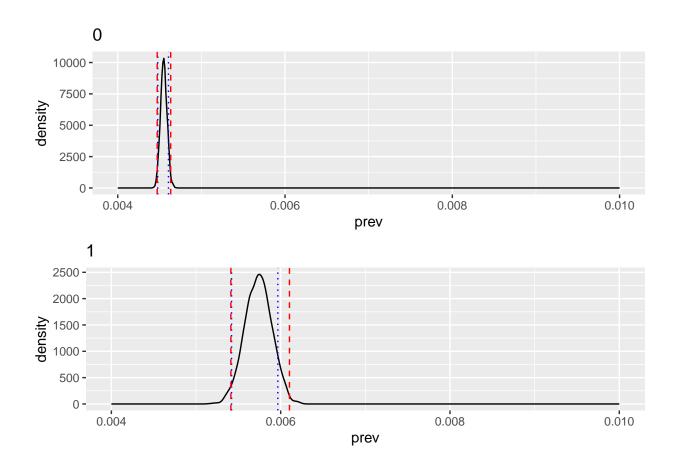


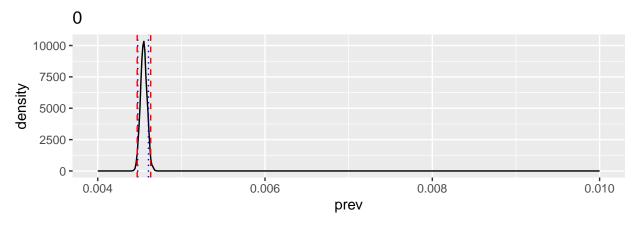


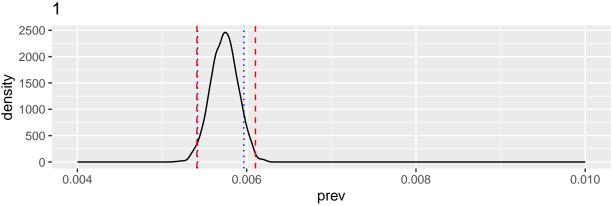






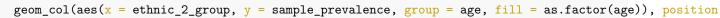


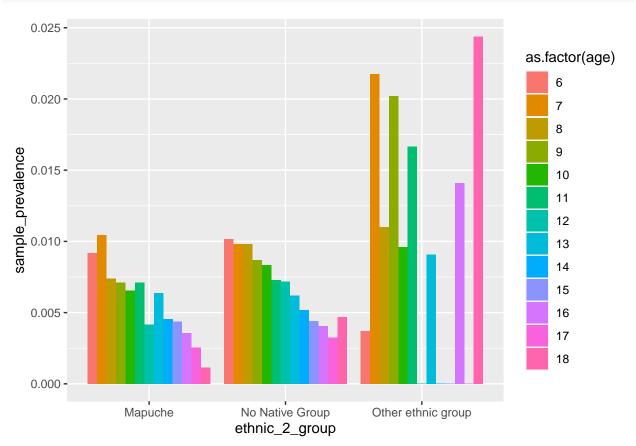




```
### Bayesian prevalence by ethnicity
aut_prev_ethnic <- chile_bayes_aut %>%
 filter(school_region_name_abr %in% c("ARAUC", "BBIO", "LAGOS", "RIOS", "RM")) %>%
 group_by(ethnic_3_group, age_june30, sex, autism) %>%
 summarise(count = n()) %>%
 pivot_wider(names_from = autism, values_from = count) %>%
 rename("n_noautism" = "0", "n_autism" = "1", "age" = "age_june30") %>%
 mutate(ethnic_2_group = ifelse(ethnic_3_group == "Aymara", "Other ethnic group", ethnic_3_group),
        n_autism = ifelse(is.na(n_autism), 0, n_autism),
        sample_pop_size = n_noautism + n_autism,
        sample_prevalence = n_autism / sample_pop_size) %>%
 left_join(chile_stdpop, by = c("age", "sex")) %>%
 mutate(aut_prev_std = n_autism / sample_pop_size * pop_prop,
        w = std_pop / (sample_pop_size * n_std_pop),
        w2 = pop_prop / sample_pop_size,
        \#sum\_std\_pop = sum(std\_pop)
 ) %>%
 ungroup()
```

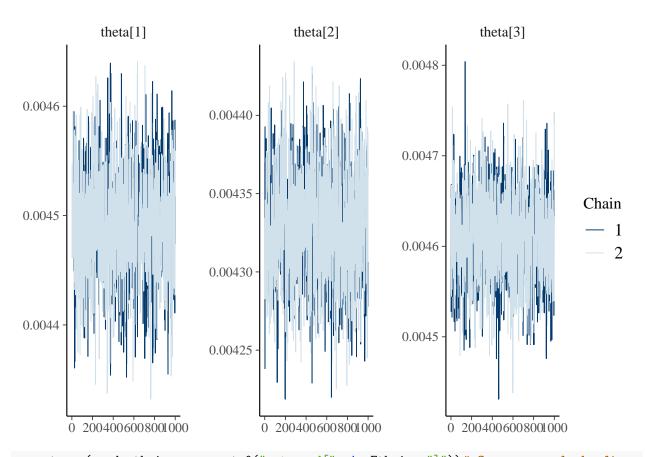
`summarise()` has grouped output by 'ethnic_3_group', 'age_june30', 'sex'. You



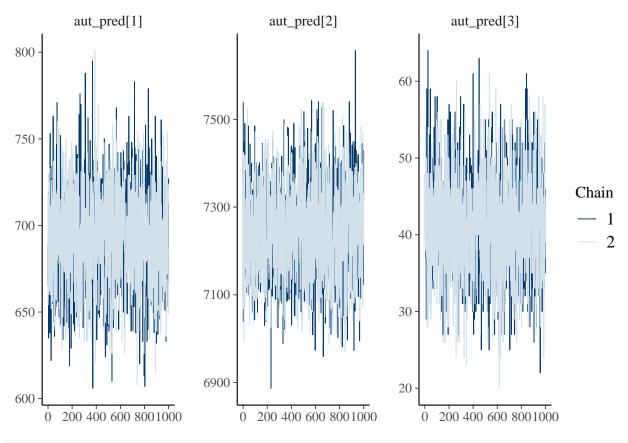


```
\#geom\_col(aes(x = ethnic\_3\_group, y = sample\_prevalence, group = sex, fill = as.factor(sex)), positio
# 1 is male, 2 is female
aut_prev_ethnic_adj <- aut_prev_ethnic %>%
  #group_by(ethnic_3_group) %>%
  group_by(ethnic_2_group) %>%
  summarise(sum_sample_pop_size = sum(sample_pop_size),
            crude_rate = sum(n_autism) / sum(sample_pop_size),
            crude_count = sum(n_autism),
            adjusted_rate = sum(n_autism / sample_pop_size * pop_prop),
            adjusted_count = round(adjusted_rate * sum_sample_pop_size, 0), # had to fudge this to get
            #adjusted_count = adjusted_rate * sum_sample_pop_size,
            var = sum(pop_prop^2 * n_autism / sample_pop_size^2),
            #se2 = sqrt(sum((std_pop/sum(std_pop))^2 * n_autism/sample_pop_size^2)),
            w M = max(w),
            ci_lower = var / (2*adjusted_rate) * qchisq(p = 0.05/2, df = 2*adjusted_rate^2 / var),
            ci_upper = (var + w_M^2) / (2*(adjusted_rate + w_M)) * qchisq(p = 1-0.05/2, df = 2*(adjusted_rate + w_M))
  #arrange(ethnic_3_group)
  arrange(ethnic_2_group)
# Prior: age and sex standardised prevalence in the whole Chile dataset
theta_mu \leftarrow 0.0046
theta_sigma \leftarrow (0.0047-0.0045) / (2*1.96)
theta_a <- theta_mu * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
```

```
theta_b <- (1 - theta_mu) * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
#nEthnic <- length(unique(aut_prev_ethnic$ethnic_3_group))</pre>
nEthnic <- length(unique(aut_prev_ethnic$ethnic_2_group))</pre>
rand_ethnic_model <- "model {</pre>
  for(i in 1:nEthnic) { # For each ethnic group
    theta[i] ~ dbeta(theta a, theta b)
    aut_sample[i] ~ dbin(theta[i], nObs[i])
    aut_pred[i] ~ dbin(theta[i], nObs[i])
  }
}"
rand_ethnic_data <- list(theta_a = theta_a,</pre>
                         theta_b = theta_b,
                         n0bs = aut_prev_ethnic_adj$sum_sample_pop_size,
                         aut_sample = aut_prev_ethnic_adj$adjusted_count,
                         nEthnic = nEthnic)
\#rand\_rural\_ini \leftarrow list(list(theta = 0.001), \#, spec = 0.5, sens = 0.5),
                     list(theta = 0.01)) #, spec = 0.9, sens = 0.9))
\# Run JAGS model and discard burn-in samples
rand_ethnic_jag <- jags.model(textConnection(rand_ethnic_model),</pre>
                              data = rand_ethnic_data,
                              #inits = rand region ini,
                              n.chains = 2,
                              quiet = TRUE)
update(rand_ethnic_jag, n.iter = nBurn)
rand_ethnic_sam <- coda.samples(model = rand_ethnic_jag,</pre>
                                variable.names = pars,
                                n.iter = nIter)
# Check for convergence in parameters of interest
#mcmc_trace(rand_region_sam, pars)
mcmc_trace(rand_ethnic_sam, paste0("theta[", 1:nEthnic, "]")) # Convergence looks fine and rhats <= 1.1
```



mcmc_trace(rand_ethnic_sam, paste0("aut_pred[", 1:nEthnic, "]"))# Convergence looks fine and rhats <= 1</pre>

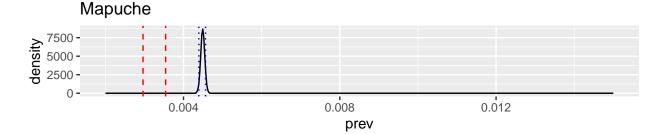


summary(as_draws(rand_ethnic_sam)) %>% print(n = Inf)

```
# A tibble: 11 x 10
##
##
      variable
                         mean median
                                            sd
                                                   mad
                                                             q5
                                                                    q95
                                                                          rhat ess_b~1
                                                 <dbl>
##
      <chr>
                        <dbl>
                                <dbl>
                                         <dbl>
                                                          <dbl>
                                                                  <dbl>
                                                                         <dbl>
                                                                                  <dbl>
                      6.93e+2 6.93e+2 2.77e+1 2.82e+1 6.48e+2 7.38e+2
                                                                                  1863.
##
    1 aut_pred[1]
                                                                         1.00
                      7.25e+3 7.25e+3 1.04e+2 1.07e+2 7.08e+3 7.42e+3
                                                                         0.999
                                                                                  1653.
##
    2 aut pred[2]
                      4.12e+1 4.1 e+1 6.61e+0 7.41e+0 3.1 e+1 5.3 e+1
                                                                                  1849.
##
    3 aut_pred[3]
                                                                         1.00
                      4.99e+2 4.99e+2 0
                                               0
                                                       4.99e+2 4.99e+2 NA
                                                                                    NA
##
    4 aut_sample[1]
##
    5 aut_sample[2]
                      6.77e+3 6.77e+3 0
                                               0
                                                       6.77e+3 6.77e+3 NA
                                                                                    NΑ
                                                                                    NA
##
    6 aut_sample[3]
                      5.9 e+1 5.9 e+1 0
                                               0
                                                       5.9 e+1 5.9 e+1 NA
##
    7 theta[1]
                      4.49e-3 4.49e-3 5.00e-5 4.90e-5 4.41e-3 4.57e-3
                                                                         1.00
                                                                                  1252.
##
    8 theta[2]
                      4.33e-3 4.33e-3 3.60e-5 3.72e-5 4.27e-3 4.39e-3
                                                                                  1106.
    9 theta[3]
                                                                                  1320.
                      4.61e-3 4.61e-3 4.93e-5 5.09e-5 4.53e-3 4.69e-3
                                                                        1.00
##
## 10 theta_a
                      8.09e+3 8.09e+3 0
                                               0
                                                       8.09e+3 8.09e+3 NA
                                                                                    NA
##
  11 theta_b
                      1.75e+6 1.75e+6 0
                                               0
                                                       1.75e+6 1.75e+6 NA
                                                                                    NA
     ... with 1 more variable: ess_tail <dbl>, and abbreviated variable name
## #
       1: ess_bulk
rand_ethnic_summ <- summary(subset_draws(as_draws(rand_ethnic_sam), pars),</pre>
                            -quantile(.x, probs=c(0.025, 0.5, 0.975)),
                            ~mcse_quantile(.x, probs=c(0.025, 0.5, 0.975)),
                            "rhat") %>%
  arrange(desc(mcse_q50))
rand_ethnic_summ
```

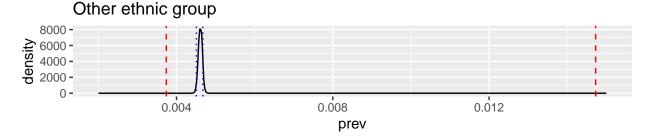
A tibble: 11 x 8

```
##
      variable
                            `2.5%`
                                     `50%` `97.5%` mcse_q~1 mcse_q50 mcse_q~2
##
      <chr>
                             <dbl>
                                             <dbl>
                                                      <dbl>
                                                                <dbl>
                                                                                <dbl>
                                     <dbl>
                                                                         <dbl>
   1 aut pred[2]
                                   7.25e+3 7.45e+3 4
                                                                                0.999
##
                       7050.
                                                         e+0
                                                             3.5 e+0
                                                                      8.5 e+0
                                   6.93e+2 7.46e+2 1.5 e+0
   2 aut_pred[1]
                        639
                                                             5
                                                                           e+0
                                                                                1.00
                                                                  e-1
##
    3 theta[1]
                           0.00439 4.49e-3 4.59e-3 4.28e-6
                                                              1.86e-6
                                                                       3.19e-6
## 4 theta[2]
                           0.00426 4.33e-3 4.40e-3 2.33e-6 1.52e-6 3.99e-6 1.00
## 5 theta[3]
                          0.00451 4.61e-3 4.70e-3 6.32e-6 1.44e-6 5.14e-6 1.00
## 6 aut_pred[3]
                                   4.1 e+1 5.5 e+1 1
                                                                           e-1 1.00
                         29
                                                         e+0 0
                                                                       5
##
   7 theta a
                       8091.
                                   8.09e+3 8.09e+3 NA
                                                             NA
                                                                      NA
                                                                               NA
                    1750915.
                                                                      NA
                                                                               NA
## 8 theta_b
                                   1.75e+6 1.75e+6 NA
                                                             NA
## 9 aut_sample[1]
                        499
                                   4.99e+2 4.99e+2 NA
                                                             NA
                                                                      NA
                                                                               NA
                       6772
                                   6.77e+3 6.77e+3 NA
## 10 aut_sample[2]
                                                             NA
                                                                      NA
                                                                               NA
                                   5.9 e+1 5.9 e+1 NA
## 11 aut_sample[3]
                         59
                                                             NA
                                                                      NA
                                                                               NA
## # ... with abbreviated variable names 1: mcse_q2.5, 2: mcse_q97.5
aut_prev_ethnic_plots <- list()</pre>
ethnic_post_ci_lower <- list()</pre>
ethnic_post_ci_upper <- list()</pre>
for(i in 1:nEthnic) {
  prevs <- data.frame(prev = extract_variable(rand_ethnic_sam, paste0("theta[", i, "]")))</pre>
  ethnic_post_ci_lower[[i]] <- quantile(prevs$prev, 0.025)</pre>
  ethnic_post_ci_upper[[i]] <- quantile(prevs$prev, 0.925)</pre>
  density_plot <- ggplot(prevs, aes(x = prev)) +
    geom_density() +
    xlim(c(0.002, 0.015)) +
    geom_vline(xintercept = ethnic_post_ci_lower[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = ethnic_post_ci_upper[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = aut_prev_ethnic_adj$ci_lower[i], color = "red", linetype = "dashed") +
    geom_vline(xintercept = aut_prev_ethnic_adj$ci_upper[i], color = "red", linetype = "dashed") +
    #labs(title = aut_prev_ethnic_adj$ethnic_3_group[i])
    labs(title = aut_prev_ethnic_adj$ethnic_2_group[i])
  aut_prev_ethnic_plots[[i]] <- density_plot</pre>
do.call(grid.arrange, aut_prev_ethnic_plots)
```



No Native Group 10000 5000 -0 0.008 0.012

prev



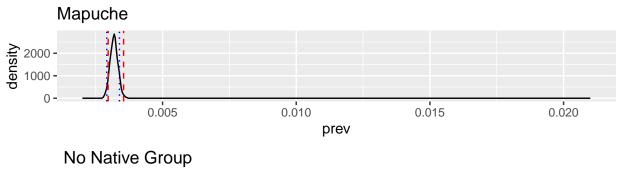
#autism prev ethnic plots <- do.call(grid.arrange, aut prev ethnic plots) #ggsave("autism_prev_ethnicity_plots.png", autism_prev_ethnic_plots, height = 10, width = 15)

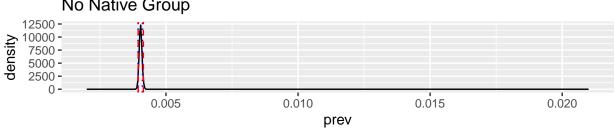
Sensitivity analysis - alter prior mean and sd

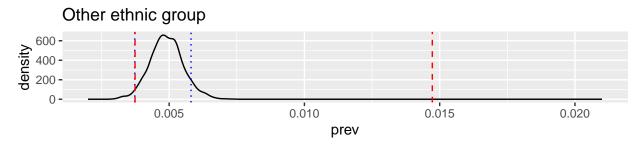
0.004

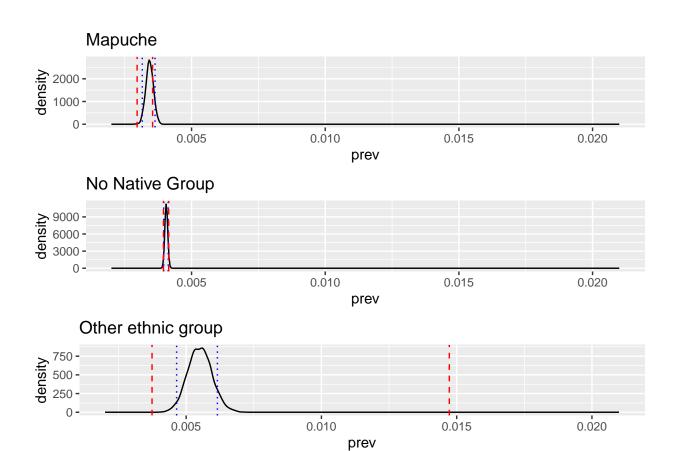
```
theta_mu <- c(0.001, 0.005, 0.01, 0.02, # 0.1%, 0.5%, 1%, 2% prevalence
              rep(0.0046, 4)) # Same as chosen prior
theta_sigma <- c(rep(0.001/1.96, 4), # Same as chosen prior
                 0.0001, 0.001, 0.05, 0.01) # +/- 0.1%, 0.5%, 1%, 5%
theta_a <- theta_mu * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
theta_b <- (1 - theta_mu) * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
for(j in 1:length(theta mu)) {
  rand_ethnic_data <- list(theta_a = theta_a[j],</pre>
                           theta_b = theta_b[j],
                           nObs = aut_prev_ethnic_adj$sum_sample_pop_size,
                           aut_sample = aut_prev_ethnic_adj$adjusted_count,
                           nEthnic = nEthnic)
  rand_ethnic_jag <- jags.model(textConnection(rand_ethnic_model),</pre>
                                data = rand_ethnic_data,
                                 #inits = rand_region_ini,
                                 n.chains = 2,
                                 quiet = TRUE)
  update(rand_ethnic_jag, n.iter = nBurn)
  rand_ethnic_sam <- coda.samples(model = rand_ethnic_jag,</pre>
                                   variable.names = pars,
```

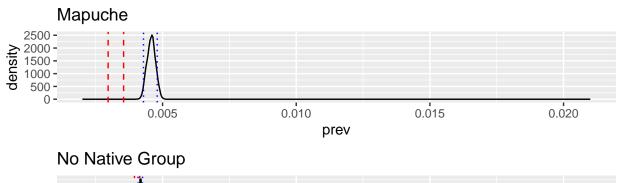
```
n.iter = nIter)
# Plots
aut_prev_ethnic_plots <- list()</pre>
ethnic_post_ci_lower <- list()</pre>
ethnic_post_ci_upper <- list()</pre>
for(i in 1:nEthnic) {
 prevs <- data.frame(prev = extract_variable(rand_ethnic_sam, paste0("theta[", i, "]")))</pre>
  ethnic_post_ci_lower[[i]] <- quantile(prevs$prev, 0.025)</pre>
  ethnic_post_ci_upper[[i]] <- quantile(prevs$prev, 0.925)</pre>
 density_plot <- ggplot(prevs, aes(x = prev)) +
    geom_density() +
    xlim(c(0.002, 0.021)) +
    geom_vline(xintercept = ethnic_post_ci_lower[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = ethnic_post_ci_upper[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = aut_prev_ethnic_adj$ci_lower[i], color = "red", linetype = "dashed") +
    geom_vline(xintercept = aut_prev_ethnic_adj$ci_upper[i], color = "red", linetype = "dashed") +
    #labs(title = aut_prev_ethnic_adj$ethnic_3_group[i])
    labs(title = aut_prev_ethnic_adj$ethnic_2_group[i])
 aut_prev_ethnic_plots[[i]] <- density_plot</pre>
do.call(grid.arrange, aut_prev_ethnic_plots)
#autism_prev_ethnic_plots <- do.call(grid.arrange, aut_prev_ethnic_plots)</pre>
#ggsave(paste0("autism_prev_ethnicity_plots_", j, ".png"), autism_prev_ethnic_plots, height = 10, wid
```

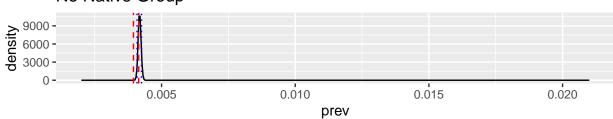


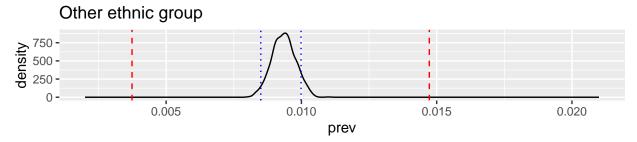


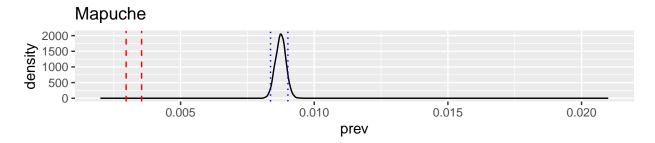


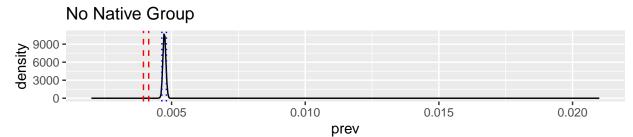


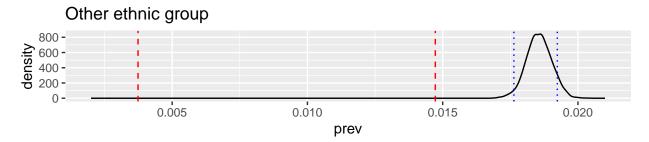


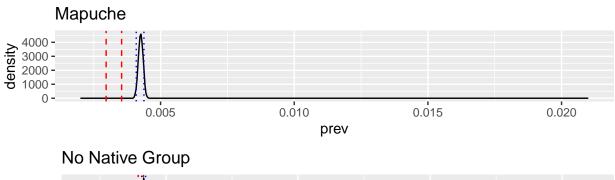


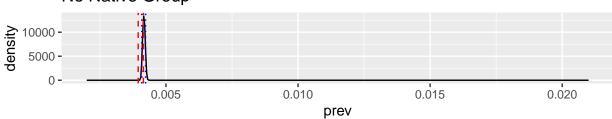


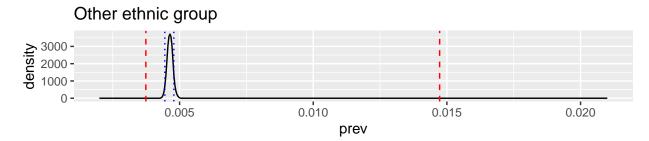


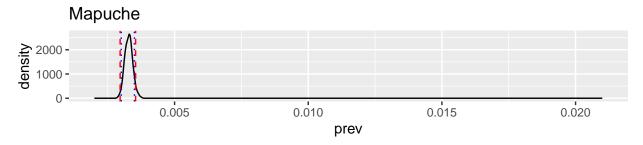


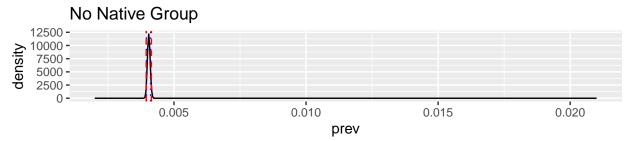


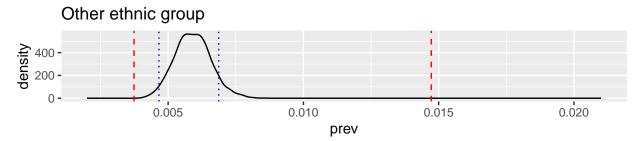


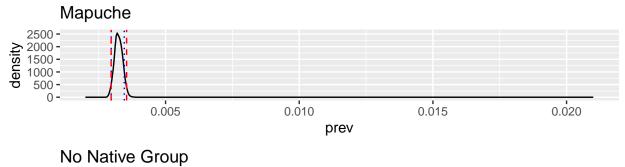


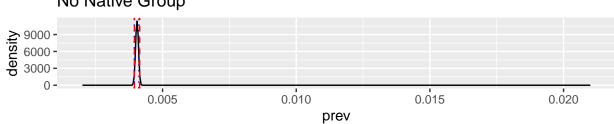


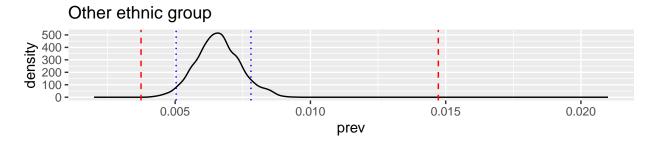


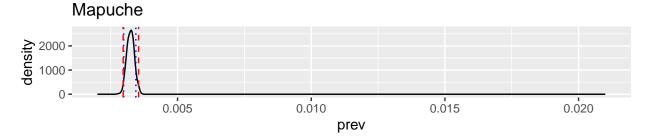


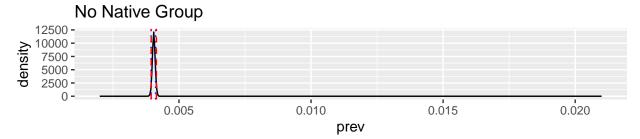


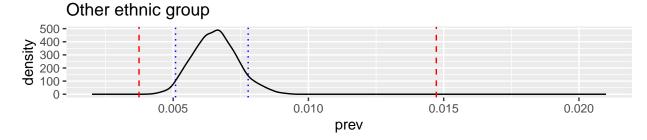












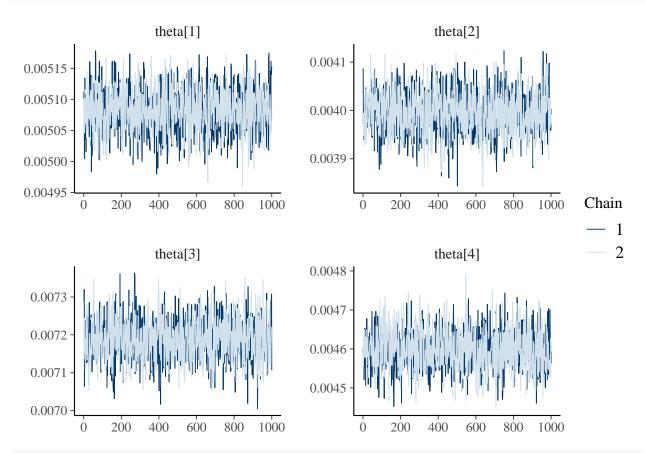
Bayesian prevalence by economic status

```
aut_prev_econ <- chile_bayes_aut %>%
  mutate(school_fee = ifelse(school_fee == "", "SIN INFORMACION", school_fee),
         school_fee_group = ifelse(school_fee == "GRATUITO", "Free",
                                  ifelse(school_fee %in% c("$1.000 A $10.000", "$10.001 A $25.000", "$2
                                         ifelse(school_fee == "MAS DE $100.000", "High", "No information
  group_by(school_fee, school_fee_group, age_june30, sex, autism) %>%
  summarise(count = n()) %>%
  pivot_wider(names_from = autism, values_from = count) %>%
  rename("n_noautism" = "0", "n_autism" = "1", "age" = "age_june30") %>%
  mutate(n_autism = ifelse(is.na(n_autism), 0, n_autism),
         sample_pop_size = n_noautism + n_autism,
         sample_prevalence = n_autism / sample_pop_size) %>%
  left_join(chile_stdpop, by = c("age", "sex")) %>%
  mutate(aut_prev_std = n_autism / sample_pop_size * pop_prop,
         w = std_pop / (sample_pop_size * n_std_pop),
         w2 = pop_prop / sample_pop_size,
         sum_std_pop = sum(std_pop)) %>%
  ungroup()
## `summarise()` has grouped output by 'school_fee', 'school_fee_group',
## 'age_june30', 'sex'. You can override using the `.groups` argument.
aut_prev_econ_adj <- aut_prev_econ %>%
  #group_by(school_fee) %>%
```

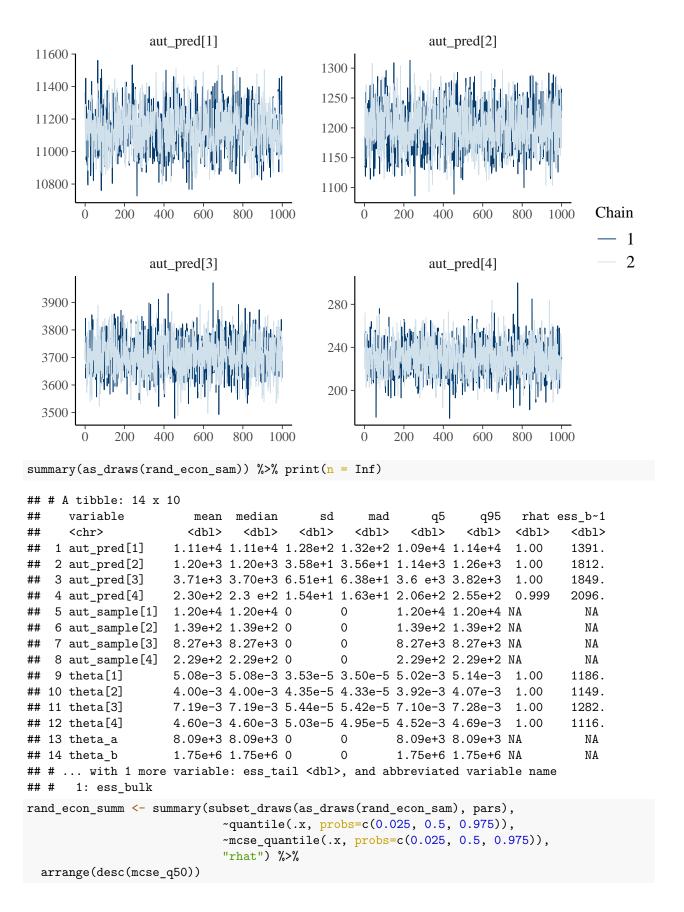
```
group_by(school_fee_group) %>%
  summarise(sum_sample_pop_size = sum(sample_pop_size),
            crude_rate = sum(n_autism) / sum(sample_pop_size),
            crude_count = sum(n_autism),
            adjusted_rate = sum(n_autism / sample_pop_size * pop_prop),
            adjusted_count = round(adjusted_rate * sum_sample_pop_size, 0), # had to fudge this to get .
            #adjusted_count = adjusted_rate * sum_sample_pop_size,
            var = sum(pop_prop^2 * n_autism / sample_pop_size^2),
            #se2 = sqrt(sum((std_pop/sum(std_pop))^2 * n_autism/sample_pop_size^2)),
            w M = max(w),
            ci_lower = var / (2*adjusted_rate) * qchisq(p = 0.05/2, df = 2*adjusted_rate^2 / var),
            ci_upper = (var + w_M^2) / (2*(adjusted_rate + w_M)) * qchisq(p = 1-0.05/2, df = 2*(adjusted_rate + w_M))
  #arrange(school fee)
  arrange(school_fee_group)
# Try informative prior
theta_mu \leftarrow 0.0046
theta_sigma \leftarrow (0.0047-0.0045) / (2*1.96)
theta_a <- theta_mu * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
theta_b <- (1 - theta_mu) * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
nEcon <- length(unique(aut_prev_econ$school_fee))</pre>
nEcon <- length(unique(aut_prev_econ$school_fee_group))</pre>
rand econ model <- "model {</pre>
 for(i in 1:nEcon) { # For each economic status level
    theta[i] ~ dbeta(theta a, theta b)
    aut_sample[i] ~ dbin(theta[i], nObs[i])
    aut_pred[i] ~ dbin(theta[i], n0bs[i])
 }
}"
rand_econ_data <- list(theta_a = theta_a,</pre>
                          theta_b = theta_b,
                          nObs = aut_prev_econ_adj$sum_sample_pop_size,
                          aut_sample = aut_prev_econ_adj$adjusted_count,
                          nEcon = nEcon)
rand_econ_ini <- list(list(theta = rep(0.001, nEcon)), #, spec = 0.5, sens = 0.5),
                         list(theta = rep(0.01, nEcon))) #, spec = 0.9, sens = 0.9))
# Run JAGS model and discard burn-in samples
rand_econ_jag <- jags.model(textConnection(rand_econ_model),</pre>
                               data = rand_econ_data,
                               inits = rand_econ_ini,
                               n.chains = 2,
                               quiet = TRUE)
update(rand_econ_jag, n.iter = nBurn)
rand_econ_sam <- coda.samples(model = rand_econ_jag,</pre>
                                 variable.names = pars,
                                 n.iter = nIter)
```

Check for convergence in parameters of interest #mcmc_trace(rand_region_sam, pars)

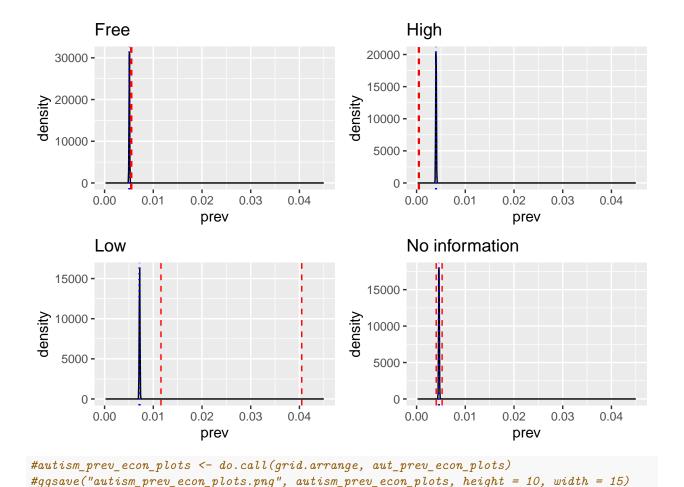
mcmc_trace(rand_econ_sam, paste0("theta[", 1:nEcon, "]")) # Convergence looks fine and rhats <= 1.1</pre>



mcmc_trace(rand_econ_sam, paste0("aut_pred[", 1:nEcon, "]"))# Convergence looks fine and rhats <= 1.1</pre>



```
rand_econ_summ
## # A tibble: 14 x 8
      variable
                            `2.5%`
                                     `50%` `97.5%` mcse_q~1 mcse_q50 mcse_q~2
                                                                                  rhat
##
      <chr>
                             <dbl>
                                             dbl>
                                                       <dbl>
                                                                <dbl>
                                                                          <dbl>
                                                                                 <dbl>
                                     <dbl>
                                   1.11e+4 1.14e+4
                                                                       8.5 e+0
                                                                                 1.00
   1 aut pred[1]
                      10883.
                                                         e+0
                                                                  e+0
    2 aut_pred[3]
                                                         e+0
##
                        3577.
                                   3.70e+3 3.83e+3 5
                                                              2
                                                                  e+0
                                                                       3.5 e+0 1.00
##
   3 aut_pred[2]
                       1133.
                                   1.20e+3 1.27e+3 1.5 e+0
                                                              1
                                                                  e+0
                                                                       1.5 e+0
                                                                                 1.00
## 4 aut_pred[4]
                        201
                                   2.3 e+2 2.6 e+2 1
                                                                            e+0 0.999
                                                         e+0
                                                              5
                                                                  e-1
                                                                       6.40e-6
## 5 theta[3]
                           0.00708 7.19e-3 7.30e-3 4.07e-6
                                                             1.72e-6
## 6 theta[4]
                           0.00450 4.60e-3 4.70e-3 3.61e-6 1.70e-6 4.34e-6
## 7 theta[2]
                           0.00391 4.00e-3 4.08e-3 3.69e-6 1.32e-6 3.74e-6 1.00
## 8 theta[1]
                           0.00501 5.08e-3 5.15e-3 2.04e-6 1.07e-6 2.30e-6 1.00
## 9 theta_a
                        8091.
                                   8.09e+3 8.09e+3 NA
                                                             NA
                                                                      NA
                                                                               NA
                                   1.75e+6 1.75e+6 NA
## 10 theta_b
                     1750915.
                                                             NA
                                                                      NA
                                                                               NA
## 11 aut_sample[1]
                      11980
                                   1.20e+4 1.20e+4 NA
                                                             NA
                                                                      NA
                                                                                NA
## 12 aut sample[2]
                        139
                                   1.39e+2 1.39e+2 NA
                                                             NA
                                                                      NA
                                                                               NΑ
                                   8.27e+3 8.27e+3 NA
## 13 aut_sample[3]
                        8267
                                                             NA
                                                                      NA
                                                                                NA
## 14 aut_sample[4]
                         229
                                   2.29e+2 2.29e+2 NA
                                                             NA
                                                                                NA
## # ... with abbreviated variable names 1: mcse_q2.5, 2: mcse_q97.5
aut_prev_econ_plots <- list()</pre>
econ_post_ci_lower <- list()</pre>
econ_post_ci_upper <- list()</pre>
for(i in 1:nEcon) {
  prevs <- data.frame(prev = extract_variable(rand_econ_sam, paste0("theta[", i, "]")))</pre>
  econ_post_ci_lower[[i]] <- quantile(prevs$prev, 0.025)</pre>
  econ_post_ci_upper[[i]] <- quantile(prevs$prev, 0.925)</pre>
  density_plot \leftarrow ggplot(prevs, aes(x = prev)) +
    geom_density() +
    xlim(c(0.0002, 0.045)) +
    geom_vline(xintercept = econ_post_ci_lower[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = econ_post_ci_upper[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = aut_prev_econ_adj$ci_lower[i], color = "red", linetype = "dashed") +
    geom_vline(xintercept = aut_prev_econ_adj$ci_upper[i], color = "red", linetype = "dashed") +
    #labs(title = aut_prev_econ_adj$school_fee[i])
    labs(title = aut_prev_econ_adj$school_fee_group[i])
  aut_prev_econ_plots[[i]] <- density_plot</pre>
do.call(grid.arrange, aut_prev_econ_plots)
```



Sensitivity analysis - alter prior mean and sd

```
theta_mu <- c(0.001, 0.005, 0.01, 0.02, # 0.1\%, 0.5\%, 1\%, 2\% prevalence
              rep(0.0046, 4)) # Same as chosen prior
theta_sigma <- c(rep(0.001/1.96, 4), # Same as chosen prior
                 0.0001, 0.001, 0.05, 0.01) # +/- 0.1%, 0.5%, 1%, 5%
theta_a <- theta_mu * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
theta_b <- (1 - theta_mu) * (theta_mu * (1-theta_mu) / theta_sigma^2 - 1)
for(j in 1:length(theta mu)) {
  #print(j)
  #print(theta_a[j])
  #print(theta_b[j])
  rand_econ_data <- list(theta_a = theta_a[j],</pre>
                           theta_b = theta_b[j],
                           nObs = aut_prev_econ_adj$sum_sample_pop_size,
                           aut_sample = aut_prev_econ_adj$adjusted_count,
                           nEcon = nEcon)
  rand_econ_jag <- jags.model(textConnection(rand_econ_model),</pre>
                                 data = rand_econ_data,
                                 inits = rand_econ_ini,
                                 n.chains = 2,
                                 quiet = TRUE)
```

```
update(rand_econ_jag, n.iter = nBurn)
rand_econ_sam <- coda.samples(model = rand_econ_jag,</pre>
                                 variable.names = pars,
                                 n.iter = nIter)
mcmc_trace(rand_econ_sam, paste0("theta[", 1:nEcon, "]")) # Convergence looks fine and rhats <= 1.1
mcmc_trace(rand_econ_sam, paste0("aut_pred[", 1:nEcon, "]"))# Convergence looks fine and rhats <= 1.1</pre>
aut_prev_econ_plots <- list()</pre>
econ_post_ci_lower <- list()</pre>
econ_post_ci_upper <- list()</pre>
for(i in 1:nEcon) {
 prevs <- data.frame(prev = extract_variable(rand_econ_sam, paste0("theta[", i, "]")))</pre>
  econ_post_ci_lower[[i]] <- quantile(prevs$prev, 0.025)</pre>
  econ_post_ci_upper[[i]] <- quantile(prevs$prev, 0.925)</pre>
  density_plot <- ggplot(prevs, aes(x = prev)) +
    geom_density() +
    xlim(c(0.0002, 0.05)) +
    geom_vline(xintercept = econ_post_ci_lower[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = econ_post_ci_upper[[i]], color = "blue", linetype = "dotted") +
    geom_vline(xintercept = aut_prev_econ_adj$ci_lower[i], color = "red", linetype = "dashed") +
    geom_vline(xintercept = aut_prev_econ_adj$ci_upper[i], color = "red", linetype = "dashed") +
    #labs(title = aut_prev_econ_adj$school_fee[i])
    labs(title = aut_prev_econ_adj$school_fee_group[i])
 aut_prev_econ_plots[[i]] <- density_plot</pre>
#autism_prev_econ_plots <- do.call(grid.arrange, aut_prev_econ_plots)</pre>
\#ggsave(paste0("autism\_prev\_econ\_plots\_", j, ".png"), autism\_prev\_econ\_plots, height = 10, width = 15
```