homework 04

April 21, 2020

0.0.1 Exercise 1

Let θ_1 and θ_2 be real valued parameters of the model

$$y = \frac{\theta_1 x}{\theta_2 + x}.$$

a. Choose two suitable prior distributions for θ_1 and θ_2 and use HMC algorithm to find their posterior distributions, conditioning on the observations

```
x = (28, 55, 110, 138, 225, 375)y = (0.053, 0.060, 0.112, 0.105, 0.099, 0.122).
```

```
import numpy as np
import matplotlib.pyplot as plt
import matplotlib.style as style
style.use('ggplot')
import seaborn as sns
import torch
import pyro
import pyro.distributions as dist
from pyro.infer.mcmc import MCMC, HMC, NUTS
import pyro.poutine as poutine
np.set_printoptions(precision=4, suppress=True)
```

```
[164]: def model(x, d: torch.distributions.Distribution):

    pyro.clear_param_store()
    # Suitable prior distribution with mean 0 and variance 1
    theta1 = pyro.sample("theta1", d)
    theta2 = pyro.sample("theta2", d)

# Calculated y variable
    yhat = (theta1*obs_x) / (theta2+obs_x)
    y = pyro.sample("y", dist.Normal(yhat,1))
    print("theta1: {}, theta2: {}\nx: {}\ny: {}\".format(theta1,theta2,x,y))

    return y
```

```
def conditioned_model(y, x, d: torch.distributions.Distribution):
           pyro.clear_param_store()
           # Suitable prior distribution with mean 0 and variance 1
           theta1 = pyro.sample("theta1", d)
           theta2 = pyro.sample("theta2", d)
           # Calculated y variable
           yhat = (theta1*obs_x) / (theta2+obs_x)
           y = pyro.sample("y", dist.Normal(yhat,1), obs=y)
           return y
       #model(obs_x, dist.Normal(0,1))
       #conditioned_model(obs_y,obs_x,dist.Normal(0,1))
[165]: # @param model conditioned pyro model
       # @param obs_y observations of y value
       # @param obs_x observation of x value
       # @param dist prior dist of theta1 and theta2
       # Oparam num_sample number of sample from stationary distribution
       # @param warmup burn-in period, number of discarded samples before performing_
       → the actual sampling
       # Oparam chain number of independent MCMC runs
       # Oreturn Progress bars of chains, summary of mcmc, lineplots of chains, dist_\sqcup
       \rightarrow plot of posterior
       # P.S. This functions is written not to have code duplicates
       # for trying different parameters for mcmc algorithm and prior dist, that is_{\sqcup}
        →why it is not that much generalize
       def analyze_hmc(model,obs_y,obs_x, d:torch.distributions.Distribution,__
        →num_sample, warmup, chain):
           '''This function runs HMC algorithm with given parameters
           and plot the chain results and posterior distribution'''
           hmc_kernel = HMC(model=model) # transition kernel
           mcmc = MCMC(hmc_kernel, num_samples=num_sample, warmup_steps=warmup,_
        →num_chains=chain)
           # posteriors
           posterior = mcmc.run(y = obs_y,x = obs_x, d=d)
           # dictionary of sampled values
           print("Keys:",mcmc.get_samples().keys(), "\n")
```

print("Summary of MCMC:\n")

```
mcmc.summary()
   mcmc_samples = mcmc.get_samples(group_by_chain=True)
   chains_1 = mcmc_samples["theta1"]
   chains_2 = mcmc_samples["theta2"]
   n_chains_1, n_samples_1 = chains_1.shape
   n_chains_2, n_samples_2 = chains_2.shape
   print("Shape of chain_1 {} and chain_2 {}\n".format(chains_1.shape,__
⇔chains 2.shape))
   fig, ax = plt.subplots(1, n_chains_1, figsize=(12,3))
   for i, chain in enumerate(chains_1):
       sns.lineplot(x=range(n_samples_1), y=chain, ax=ax[i])
       ax[i].set_title("theta 1_chain "+str(i+1))
   fig, ax = plt.subplots(1, n_chains_2, figsize=(12,3))
   for i, chain in enumerate(chains 2):
       sns.lineplot(x=range(n_samples_2), y=chain, ax=ax[i])
       ax[i].set_title(" theta 2_chain "+str(i+1))
   print("expected theta1 =", mcmc_samples['theta1'].mean().item())
   print("expected theta2 =", mcmc_samples['theta2'].mean().item())
   fig,a = plt.subplots(1,2, figsize=(14,4))
   sns.distplot(mcmc_samples['theta1'], ax = a[0])
   a[0].set_title("P(theta1 | x = obs_x, y = obs_y)")
   a[0].set_xlabel("Theta1")
   sns.distplot(mcmc_samples['theta2'], ax = a[1])
   a[1].set_title("P(theta2 | x = obs_x, y = obs_y)")
   a[1].set_xlabel("Theta2")
```

• Chosen prior distributions for θ_1 and θ_2 are standard normal distribution. Chosen parameters for HMC are num_samples = 500,warmup_steps = 1000 and num_chains = 3.

HBox(children=(FloatProgress(value=0.0, description='Warmup [1]', max=1500.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [2]', max=1500.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [3]', max=1500.0, style=ProgressStyle=

Keys: dict_keys(['theta1', 'theta2'])

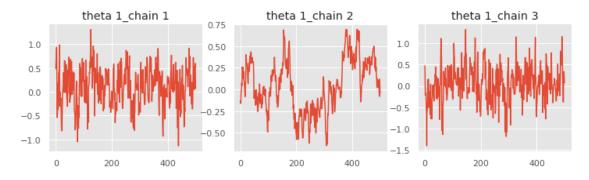
Summary of MCMC:

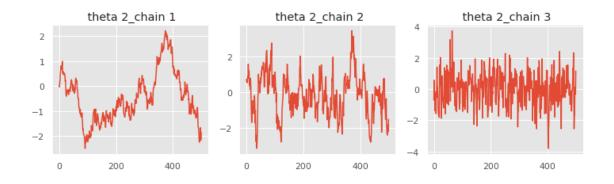
	mean	std	median	5.0%	95.0%	n_eff	r_hat
theta1	0.07	0.37	0.06	-0.50	0.69	104.04	1.01
theta2	-0.19	1.08	-0.20	-1.95	1.58	55.27	1.06

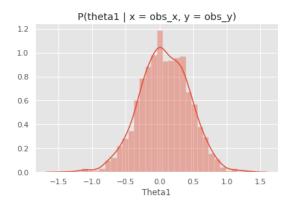
Number of divergences: 0

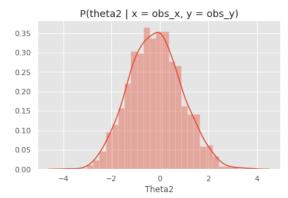
Shape of chain_1 torch.Size([3, 500]) and chain_2 torch.Size([3, 500])

expected theta1 = 0.07132413983345032 expected theta2 = -0.19076305627822876









- There is no divergence in our hmc summary results but n_eff values are low and r_hat values are above 1 so we can say that some chains have not fully converged and above plots showed us that comment that is made before for mcmc.summary is true. Most of the chains are not consistent apart from theta1_chain 3. Also you can find the distribution of $P(\theta_1|x,y)$ and $P(\theta_2|x,y)$ as a result of MCMC algorithm
- b. Discuss how different parameters for both priors and the HMC algorithm lead to different estimates.
- Changing the parameters HMC algorithm may lead better estimates for posterior distribution. Especially increasing number of samples and number of chains may lead better estimations.
- Therefore parameters of mcmc algorithm are increased such as: num_samples = 1200,warmup_steps = 400 and num_chains = 5.

[168]: analyze_hmc(conditioned_model,obs_y,obs_x,dist.Normal(0,1),1200, 400, 5)

HBox(children=(FloatProgress(value=0.0, description='Warmup [1]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [2]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [3]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [4]', max=1600.0, style=ProgressStyle=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [5]', max=1600.0, style=ProgressStyle=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [5]', max=1600.0, style=ProgressStyle=Prog

Keys: dict_keys(['theta1', 'theta2'])

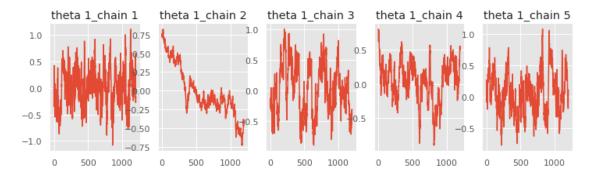
Summary of MCMC:

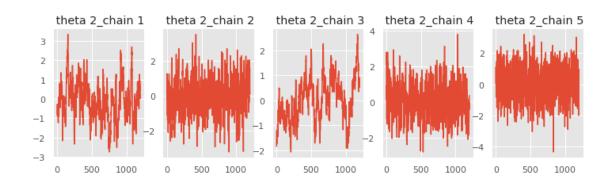
	mean	std	median	5.0%	95.0%	n_eff	r_hat
theta1	0.03	0.35	0.01	-0.58	0.57	106.30	1.06
theta2	-0.02	0.98	-0.02	-1.61	1.61	398.26	1.01

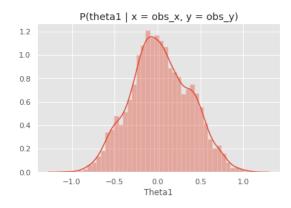
Number of divergences: 0

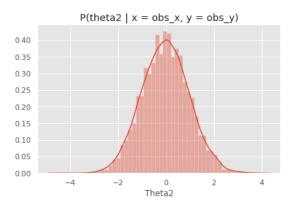
Shape of chain_1 torch.Size([5, 1200]) and chain_2 torch.Size([5, 1200])

expected theta1 = 0.030372297391295433 expected theta2 = -0.019049106165766716









- For θ_1 , we have high n_eff and close value of r_hat to 1. And most of the chains are consistent and converged just the second one seems diverged. So we can say that we have good results for θ_1 . However for θ_2 even though we can say that result are quite better than previous ones, it is not as good as θ_1 .
- Changing the parameters of prior distribution may lead better results. Therefore same analyze will be done for $\theta_k \sim \mathcal{N}(5, 10)$ where k = 1, 2

[169]: analyze_hmc(conditioned_model,obs_y,obs_x,dist.Normal(5,10),1200, 400, 5)

HBox(children=(FloatProgress(value=0.0, description='Warmup [1]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [2]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [3]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [4]', max=1600.0, style=ProgressStyle=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [5]', max=1600.0, style=ProgressStyle=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [5]', max=1600.0, style=ProgressStyle=Prog

Keys: dict_keys(['theta1', 'theta2'])

Summary of MCMC:

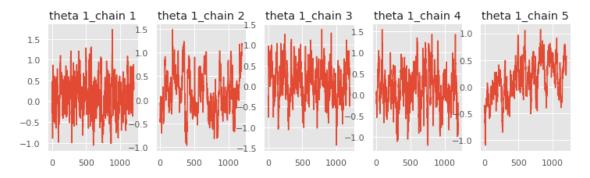
mean std median 5.0% 95.0% n_eff r_hat theta1 0.13 0.43 0.11 -0.56 0.85 329.08 1.02

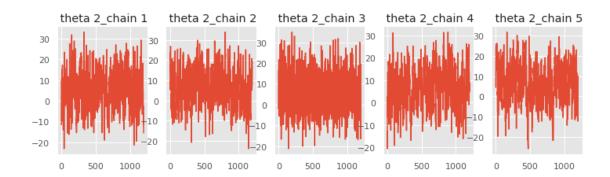
theta2 6.20 9.65 6.15 -10.74 21.00 1036.69 1.00

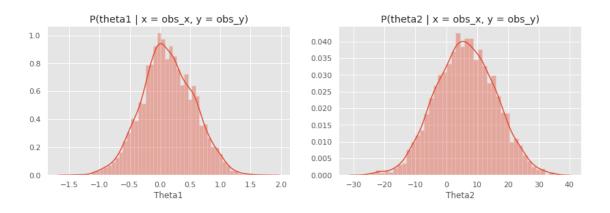
Number of divergences: 20

Shape of chain_1 torch.Size([5, 1200]) and chain_2 torch.Size([5, 1200])

expected theta1 = 0.12991181015968323 expected theta2 = 6.203857898712158







According to results (particularly checking n_eff, r_hat and plots), we have more consistent result

for theta2 this time. But it is worth to point out that theta values have higher standard deviation this time.

- c. Plot the most reliable posterior distributions, according to convergence checks on the traces.
- After trying different prior distributions like Normal, Exponential, Half-Cauchy and Inverse-Gamma, Exponential distribution gave the best result according to n_eff, r_hat and plots. you can find the results belove. Of course theta1 and theta2 may have different prior distribution but for simplicity I assume that they have same distribution with same parameters. But one can easily set a new parameter to assign them different distributions and parameters

```
[171]: analyze_hmc(conditioned_model,obs_y,obs_x,dist.Exponential(2),1200, 400, 5)
```

```
HBox(children=(FloatProgress(value=0.0, description='Warmup [1]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [2]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [3]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [4]', max=1600.0, style=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [5]', max=1600.0, style=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=Progress(value=0.0, description='Warmup [5]', max=1600.0, style=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=ProgressStyle=Prog
```

Keys: dict_keys(['theta1', 'theta2'])

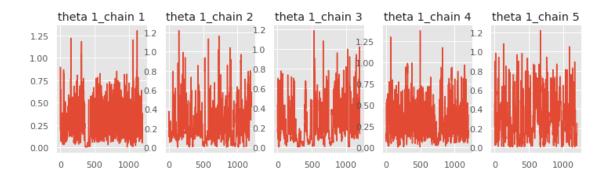
Summary of MCMC:

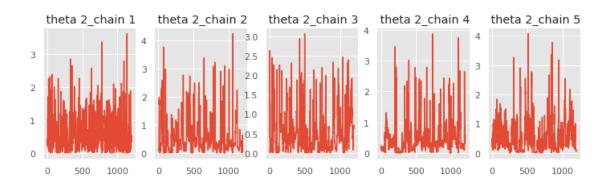
	mean	std	median	5.0%	95.0%	n_eff	r_hat
theta1	0.26	0.21	0.21	0.00	0.55	610.93	1.01
theta2	0.49	0.50	0.34	0.00	1.12	730.47	1.01

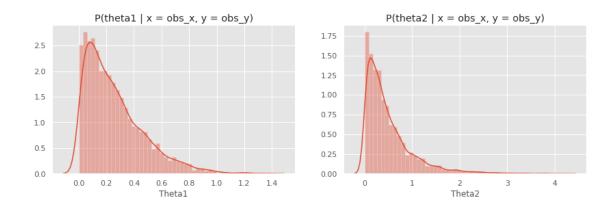
Number of divergences: 0

Shape of chain_1 torch.Size([5, 1200]) and chain_2 torch.Size([5, 1200])

expected theta1 = 0.2577589452266693 expected theta2 = 0.4914016127586365







0.0.2 Exercise 2

A bivariate Gibbs sampler for a vector $x=(x_1,x_2)$ draws iteratively from the posterior conditional distributions in the following way: - choose a starting value $p(x_1|x_2^{(0)})$ - for each iteration i: - draw $x_2(i)$ from $p(x_2|x_1^{(i-1)})$ - draw $x_1(i)$ from $p(x_1|x_2^{(i)})$

```
[172]: def gibbs_sampler(iters, warmup ,mu1, mu2, rho):
            '''This is a bivarite implementation of gibbs sampling for normal _{\sqcup}
        \hookrightarrow distribution'''
           # Starting points and ro
           rho = rho
           x0_1 = torch.tensor([0.])
           x0_2 = torch.tensor([0.])
           x1 = []
           x2 = []
           for i in pyro.plate("samples", iters):
                x_1 = pyro.sample("x_1", dist.Normal(mu1 + rho*(x0_2-mu2), (1-rho**2)))_{i}
           \# P(A \mid B_{t-1})
               x_2 = pyro.sample("x_2", dist.Normal(mu2 + rho*(x_1-mu1), (1-rho**2)))
          \# P(B \mid A_{t})
               x1.append(x_1.item())
               x2.append(x_2.item())
                #update B O with B-t
               x0_2 = x_2
           x1 = torch.tensor(x1)
           x2 = torch.tensor(x2)
           return x1[warmup:], x2[warmup:]
```

a. Supposing that samples are drawn from a bivariate normal distribution

$$\begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \sim \mathcal{N} \left[\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix} \right],$$

implement a Gibbs sampler for x which takes as inputs the number of iterations iters and the number of warmup draws warmup.

$$If: \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \sim \mathcal{N} \left[\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix} \right],$$

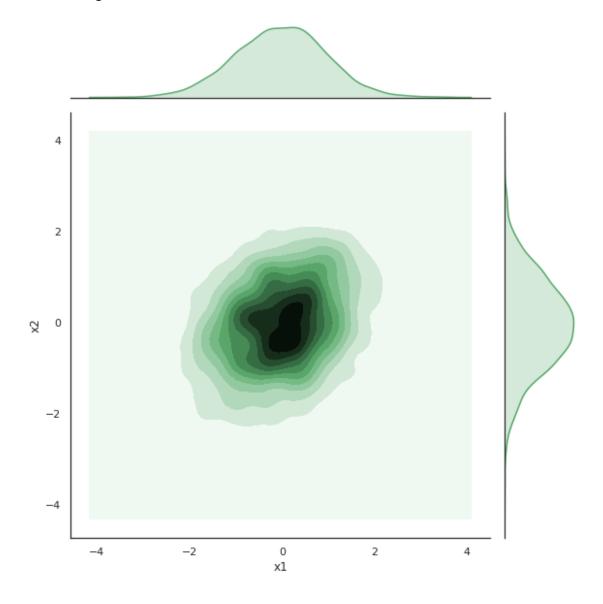
Then:
$$x_1|x_2, y \sim \mathcal{N}(\mu_1 + \rho(x_2 - \mu_2), [1 - \rho^2])x_2|x_1, y \sim \mathcal{N}(\mu_2 + \rho(x_1 - \mu_1), [1 - \rho^2])$$

reference 1, reference 2

```
[173]: x1,x2 = gibbs_sampler(3000,500,0,0.25) torch.set_printoptions(precision=5) print("Mean of x1 {0:.4f}, std of x1 {1:.4f}, mean of x2 {2:.4f}, std of x2 {3:. \hookrightarrow4f}".format(x1.mean(),x1.std(),x2.mean(),x2.std())) sns.set(style="white", color_codes=True) x = sns.jointplot(x1,x2,kind="kde",color="g",height=8) x.set_axis_labels('x1','x2')
```

Mean of x1 -0.0293, std of x1 0.9694, mean of x2 -0.0240, std of x2 0.9841

[173]: <seaborn.axisgrid.JointGrid at 0x7fa814502d90>



b. Use your implementation of Gibbs sampler to infer the parameters $\theta = (\theta_1, \theta_2)$ from **Exercise**

1.

By using the information from previous exercise:

$$\begin{pmatrix} \theta_1 \\ \theta_1 \end{pmatrix} \sim \mathcal{N} \left[\begin{pmatrix} 0.13 \\ 6.20 \end{pmatrix}, \begin{pmatrix} 1 & 0.0336 \\ 0.0336 & 1 \end{pmatrix} \right],$$

P.S: I have used the best normal prior that I get from exercise 1 not the exponential one because of simplicity. 0.13 and 6.20 was found from mcmc.summary table and ρ was found by np.corrcoef by using all samples from chains.

Mean of theta1 is 0.0958 and std of theta1 is 0.9821 Mean of theta2 is 6.2100 and std of theta2 is 0.9986

[174]: <seaborn.axisgrid.JointGrid at 0x7fa814f3b290>

