

CS 5291: Stochastic Processes for Networking

HW1 Solution

1. (10%) $E[Y] = E[X^2] = \int_0^\infty x^2 \lambda e^{-\lambda x} dx$

$$= x^2(-e^{-\lambda x}) \Big|_0^\infty - \int_0^\infty 2x(-e^{-\lambda x}) dx = \frac{2}{\lambda} \int_0^\infty x \lambda e^{-\lambda x} dx = \frac{2}{\lambda^2}$$

2. (10%) $E[X] = \int_0^\infty x f(x) dx$

$$= \int_0^\infty (1 - F_X(x)) dx = x(1 - F_X(x)) \Big|_0^\infty + \int_0^\infty x f(x) dx$$

$$= \int_0^\infty x f(x) dx = E[X]$$

3. (10%) Let N be the number of tosses required;

$$\text{Then } E[X] = E[E[N|X]]$$

$$= E[E[N|X=1]]P(X=1) + E[E[N|X=0]]P(X=0)$$

$$= (1 + E[N]) \times \frac{1}{2} + \left(2 \times \frac{1}{2} + (2 + E[N]) \times \frac{1}{2} \right) \times \frac{1}{2}$$

$$= \frac{3}{4}E[N] + \frac{3}{2}$$

$$E[N] = 6$$

4. (10%) $P\{E_1, E_2, E_3, E_4\} = \frac{C_1^4 C_{12}^{48}}{C_{13}^{52}} \times \frac{C_1^3 C_{12}^{36}}{C_{13}^{39}} \times \frac{C_1^2 C_{12}^{24}}{C_{13}^{26}} \times 1$

5.

(a) (10%) If x is a real number between 0 and 1, then $X < x$ if and only if $Y_k < x$ for all k , and $P(Y_k < x) = x$ since x is uniform on $(0,1)$. (The pdf is equal to 1, so the cdf is $\int_0^\infty 1 dx = x$)

Because the Y_k are independent,

$$P(X < x) = P(Y_1 < x) \times P(Y_2 < x) \times \cdots \times P(Y_n < x) = x^n$$

(b) (5%) The PDF of X is obtained by differentiating,

$$f_X(x) = \frac{F_X(x)}{dx} = nx^{n-1}, \quad 0 \leq x \leq 1$$

6. (a) (10%)

Uniform distribution: $f_X(x) = 1/(b-a)$

$$\varphi(t) = E[e^{tX}]$$

$$\varphi(t) = \int_a^b e^{tx} * \left(\frac{1}{(b-a)}\right) dx = \left(\frac{1}{(b-a)}\right) * (1/t) \int_a^b te^{tx} dx$$

$$= \left(\frac{1}{t(b-a)}\right) \left[e^{tx} \Big|_a^b \right] = \frac{(e^{tb} - e^{ta})}{t(b-a)}$$

$$\varphi'(t) = \left(\frac{1}{(b-a)}\right) * \left(\frac{(be^{tb} - ae^{ta})}{t} - \frac{(e^{tb} - e^{ta})}{t^2} \right)$$

$$\varphi''(t) = \left(\frac{1}{(b-a)}\right) * \left(\frac{2(e^{tb} - e^{ta})}{t^3} - \frac{2(be^{tb} - ae^{ta})}{t^2} + \frac{(b^2e^{tb} - a^2e^{ta})}{t} \right)$$

$$E[X] = \varphi'(0) = \lim_{s \rightarrow 0} \left(\frac{1}{(b-a)}\right) * \left(\frac{(be^{tb} - ae^{ta})}{t} - \frac{(e^{tb} - e^{ta})}{t^2} \right)$$

$$= \frac{(a+b)}{2} \text{ (by L'Hospital's rule)}$$

$$E[X^2] = \varphi''(0) = \frac{a^2 + ab + b^2}{3}$$

$$\text{Var}(X) = E[X^2] - E[X]^2 = \frac{(b-a)^2}{12}$$

(b) (10%)

Exponential distribution: $f_X(x) = \lambda e^{-\lambda x}$

$$\varphi(t) = E[e^{tX}] = \int_0^\infty e^{tx} \lambda e^{-\lambda x} dx = \lambda \int_0^\infty e^{-(\lambda-t)x} dx$$

$$= \frac{\lambda}{(\lambda-t)} \int_0^\infty (\lambda-t) e^{-(\lambda-t)x} dx = \frac{\lambda}{(\lambda-t)}$$

$$\varphi'(t) = \frac{\lambda}{(\lambda-t)^2}$$

$$\varphi''(t) = \frac{2\lambda}{(\lambda - t)^3}$$

$$E[X] = \varphi'(0) = \frac{1}{\lambda}$$

$$E[X^2] = \varphi''(0) = \frac{2}{\lambda^2}$$

$$\text{Var}(X) = E[X^2] - E[X]^2 = \frac{1}{\lambda^2}$$

7.

(a) (10%)

$$f_X(x) = \frac{(e^{-\lambda} \lambda^x)}{x!}$$

$$M_X(t) = \sum_{x=0}^{\infty} \frac{e^{tx} (e^{-\lambda} \lambda^x)}{x!} = e^{-\lambda} \sum_{x=0}^{\infty} \frac{(e^t \lambda)^x}{x!} = e^{-\lambda} e^{e^t \lambda} = e^{-\lambda(1-e^t)}$$

(b) (5%)

$$P\{X \geq a\} \leq \min_t (e^{-ta} M_X(t)) = \min_t (e^{-ta - \lambda(1-e^t)})$$

$$\text{Let } g(t) = e^{-ta - \lambda(1-e^t)}$$

To find the minimum of $g(t)$, we need to solve $\frac{dg(t)}{dt} = 0$

$$\frac{dg(t)}{dt} = 0 \Rightarrow (\lambda e^t - a) e^{-ta - \lambda(1-e^t)} = 0 \Rightarrow t = \ln\left(\frac{a}{\lambda}\right)$$

$$P\{X \geq a\} \leq g\left(\ln\left(\frac{a}{\lambda}\right)\right) = e^{-a \ln\left(\frac{a}{\lambda}\right) - \lambda + a} = e^{-\lambda \left(\frac{\lambda e}{a}\right)^a}$$

8. (10%)

$$e^K(t) = E[e^{tX}] = \int_{-\infty}^{\infty} e^{tX} f_X(x) dx$$

differentiating both sides w.r.t. t .

$$\begin{aligned} \frac{d}{dt} e^K(t) &= e^K(t) K'(t) = \frac{d}{dt} \left(\int_{-\infty}^{\infty} e^{tX} f_X(x) dx \right) = \int_{-\infty}^{\infty} X e^{tX} f_X(x) dx \\ &= E[X e^{tX}] \end{aligned}$$

$$e^{K(t)}K'(t) = E[e^{tX}]K'(t) = E[Xe^{tX}] \Rightarrow K'(t) = \frac{E[Xe^{tX}]}{E[e^{tX}]}$$

$$K'(0) = \frac{E[X]}{1} = E[X]$$

differentiating again w.r.t. t

$$\frac{d}{dt}\left(e^{K(t)}K'(t)\right) = e^{K(t)}K'(t)^2 + e^{K(t)}K''(t) = \frac{d}{dt}\left(\int_{-\infty}^{\infty} X e^{tX} f_X(x) dx\right)$$

$$= \int_{-\infty}^{\infty} X^2 e^{tX} f_X(x) dx = E[X^2 e^{tX}]$$

$$e^{K(t)}K'(t)^2 + e^{K(t)}K''(t) = E[e^{tX}] \left(\frac{E[Xe^{tX}]^2}{E[e^{tX}]^2} + K''(t) \right) = E[X^2 e^{tX}]$$

$$K''(t) = \frac{E[X^2 e^{tX}]}{E[e^{tX}]} - \frac{E[Xe^{tX}]^2}{E[e^{tX}]^2} = \frac{E[X^2 e^{tX}]E[e^{tX}] - E[Xe^{tX}]^2}{E[e^{tX}]^2}$$

$$K''(0) = E[X^2] - E[X]^2 = \text{Var}(X)$$