

Summary of Short-term Research Objectives

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1 Model Specification

Let L be a K -dimensional Bernoulli random variable denoting the true state. Consider the general log linear model:

$$f(l; \Theta) = \exp\{\Theta_1^T l + \Theta_2^T u_2 + \dots + \Theta_K^T u_K - A^*(\Theta)\}$$

where U_k is a $\binom{K}{k} \times 1$ vector of k -way cross-products, $k = 1, \dots, K$, and $\Theta = (\Theta_1, \dots, \Theta_K)$ contains the the natural parameters, which is a $(2^K - 1) \times 1$ vector.

Model restrictions, let $\tilde{l} = (l, u_2, \dots, u_K)^T$, and $S = \sum_{j=1}^K L_j = s$ has some fixed pmf

$$\begin{aligned} \pi(s) &:= P(S = s) \\ &= \frac{1}{A(\Theta)} \sum_{\tilde{l}: S=s} \exp\{\Theta^T \tilde{l}\}, \quad s = 0, 1, \dots, K \end{aligned} \tag{1}$$

$$A(\Theta) = \sum_{\tilde{l}: l \in \{0,1\}^K} \exp\{\Theta^T \tilde{l}\} \tag{2}$$

1.1 Additional Definition of Parameters

Using previous notations, with some properly defined γ , we have:

$$\begin{aligned} P(L_i; \pi, \gamma) &= P(L_i, S_i; \pi, \gamma) \\ &= P(L_i | S_i; \pi, \gamma) P(S_i; \pi, \gamma) \end{aligned}$$

where $P(L_i | S_i; \pi, \gamma) = P(L_i | S_i; \gamma)$ because $1(S_i = s)$ is the sufficient statistic for π_s , furthermore S_i is the sufficient statistic for π . Also $P(S_i; \pi, \gamma) = \pi_{S_i}$ by definition.

Therefore, we have

$$P(L_i; \pi, \gamma) = P(L_i | S_i; \gamma) \pi_{S_i}$$

Then we define the following parameters:

$$\begin{aligned} \gamma_{j_1, \dots, j_s} &= P(L_{ij_1} = \dots = L_{ij_s} = 1 | S_i = s) \\ \gamma &= (\gamma_1, \gamma_2, \dots, \gamma_{12}, \dots, \gamma_{1\dots K})^T, \text{ where } \sum_j \gamma_j = \sum_{j \neq j'} \gamma_{jj'} = \sum_{j \neq j' \neq j''} \gamma_{jj'j''} = \dots = \gamma_{1\dots K} = 1 \end{aligned} \quad (3)$$

Therefore $(\pi, \gamma)^T$ is a vector of length $2^K + K$ with degrees of freedom $2^K - 1$.

Let $J_i = \{j : L_{ij} = 1\}$. We have,

$$P(L_i; \pi, \gamma) = \gamma_{J_i} \pi_{S_i} \quad (4)$$

The relation between (π, γ) and Θ is defined by equation (2) together with:

$$\gamma_{J_i} = \frac{\exp(\Theta^T \tilde{l}_i)}{\sum_{l: l^T 1 = S_i} \exp(\Theta^T \tilde{l})} \quad (5)$$

with $2^K - 1 - K$ degrees of freedom.

Therefore (1), (3) and (5) together define $2^K - 1$ non-linear equations for $2^K - 1$ unknowns. If there exists a unique root for the above non-linear system, then there is a one-to-one mapping between (π, γ) and Θ , which provides the re-parameterization.

1.2 Find the Re-parameterization

1.2.1 Quasi-Newton Method

Numerically solve the system defined by (1), (3) and (5). As the dimension of L grows ($K > 6$), multiple sets of starting values are needed to reach the solution. Also, solutions to high order Θ are subject to larger error.

See code `GammaToTheta()` in the appendix.

1.2.2 Restrict Θ to QE model

Setting all high order interaction parameter to 0, using only the equations defined by $pi_0, pi_1, \gamma_1, \dots, \gamma_{K-1}$ and $\gamma_{11}, \dots, \gamma_{K-2, K}$, which are in total $\frac{K(K+1)}{2}$ equations, we can solve for Θ for larger value of K .

See code `GammaToTheta.QE()` in the appendix.

2 Posterior Distribution

$$\begin{aligned}
P(\mu, \theta^{(2)}|L) &\propto P(L, \mu, \theta^{(2)}) \\
&\propto P(L, \mu, \theta^{(1)}, \theta^{(2)}, \pi) \\
&\propto P(L|\mu, \theta^{(1)}, \theta^{(2)}, \pi)P(\mu, \theta^{(1)}, \theta^{(2)}, \pi) \\
&\propto P(L|\theta^{(1)}, \theta^{(2)})P(\theta^{(1)}, \theta^{(2)}|\mu, \pi)P(\mu)P(\pi) \\
&\propto \text{QE}(L; \theta^{(1)}, \theta^{(2)})\text{UFR}(\theta^{(1)}, \theta^{(2)}|\mu, \pi)\text{N}(\text{logit}(\mu), \Sigma)\text{tPois}(\pi)
\end{aligned}$$

where QE is the second-order log linear model.

UFR is a Multivariate distribution of $[\theta^{(1)}, \theta^{(2)}|\mu, \pi]$ subject to non-linear constraints: $M(\theta^{(1)}, \theta^{(2)}) = \mu$ and $\Pi(\theta^{(1)}, \theta^{(2)}) = \pi$, which can be sampled by a two-step procedure. tPois is a truncated conjugate Poisson distribution defined as:

$$\begin{aligned}
\pi &\sim \text{Dirichlet}(\text{hist}(\vec{s})) \\
s &\sim \frac{\lambda^s}{s!} e^{-\lambda} / [1 - \sum_{s>K} \frac{\lambda^s}{s!} e^{-\lambda}]
\end{aligned}$$

2.1 On sampling $[\theta^{(1)}, \theta^{(2)}|\mu, \pi]$

2.1.1 QE model

For QE model, we have a $J_1 \times J_2$ design matrix \tilde{L} , where $J_1 = 2^K - 1$, $J_2 = \frac{1}{2}K(K+1)$. Recall that

$$\begin{aligned}
A(\Theta) &= \frac{1}{\pi(0)} \\
\pi(s) &= \frac{1}{A(\Theta)} \sum_{\tilde{l}: S=s} \exp\{\Theta^T \tilde{l}\}, \quad s = 1, \dots, K
\end{aligned} \tag{6}$$

$$\mu_k = \frac{1}{A(\Theta)} \sum_{\tilde{l}: l_k=1} \exp\{\Theta^T \tilde{l}\}, \quad k = 1, \dots, K \tag{7}$$

Define intermediate parameter $\phi_j = \exp(\theta^T \tilde{l}_j) > 0$, $\theta = (\theta^{(1)}, \theta^{(2)})$, $j = 1, \dots, J_1$ and two $K \times J_1$ sub-design matrices B , C , where $B[k, j] = 1(\tilde{l}_j^T \mathbf{1} = k)$, $C[k, j] = 1(\tilde{L}[j, k] = 1)$. Thus (6) and (7) become

$$\begin{aligned}
\vec{\phi} &> 0 \\
B\vec{\phi} &= \vec{\pi}/\pi(0) \\
C\vec{\phi} &= \vec{\mu}/\pi(0)
\end{aligned}$$

Based on [1][2], we can sample $\vec{\phi}$ from Uniform distribution subject to the above linear constraints efficiently and robustly. Now we have a over-determined linear system: (J_1 equations with J_2 unknowns)

$$\tilde{L}\theta = \log \vec{\phi}$$

Then we can use Least Square method to solve for θ .

2.1.2 General model

For General model, \tilde{L} is $J_1 \times J_1$, the intermediate parameters $\vec{\phi}$ and $\pi(0)$ fully specify all cell probabilities, thus the posterior distribution becomes

$$P(\mu, \vec{\phi}|L) \propto \text{LL}(L; \vec{\phi})\text{UFR}(\vec{\phi}|\mu, \pi)\text{N}(\text{logit}(\mu), \Sigma)\text{tPois}(\pi)$$

Reference

1. Smith RL . Efficient Monte-Carlo Procedures for Generating Points Uniformly Distributed over Bounded Regions. *Operations Research*, 32(6), 1296-1308 (1984).
2. Van den Meersche, Karel, K. E. R. Soetaert, and D. J. Van Oevelen. "xsample (): an R function for sampling linear inverse problems." *Journal of Statistical Software* 30 (2009).