

# Maximum Likelihood Estimation

Given “data”  $X_1, X_2, \dots, X_n$ , a random sample (iid) from a distribution with unknown parameter  $\theta$ , we want to find the value of  $\theta$  in the parameter space that maximizes our “probability” of observing that data.

- If  $X_1, X_2, \dots, X_n$  are discrete, we can look at

$$P(X_1 = x_1, X_2 = x_2, \dots, X_n = x_n)$$

as a function of  $\theta$ , and find the  $\theta$  that maximizes it.

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as a function of  $\theta$ , and find the  $\theta$  that maximizes it.

This is the joint pmf for  $X_1, X_2, \dots, X_n$ .

- The analogue for continuous  $X_1, X_2, \dots, X_n$  is to maximize the joint pdf with respect to  $\theta$ .

The pmf/pdf for any one of  $X_1, X_2, \dots, X_n$  is denoted by  $f(x)$ .

No 1,2, ..., n here.

We will emphasize the dependence of  $f$  on a parameter  $\theta$  by writing it as

$$f(x; \theta)$$

The joint pmf/pdf for all  $n$  of them is

$$\underbrace{f(x_1, x_2, \dots, x_n; \theta)}_{f(\vec{x}; \theta)} = \prod_{i=1}^n f(x_i; \theta)$$

The joint pmf/pdf for all  $n$  of them is

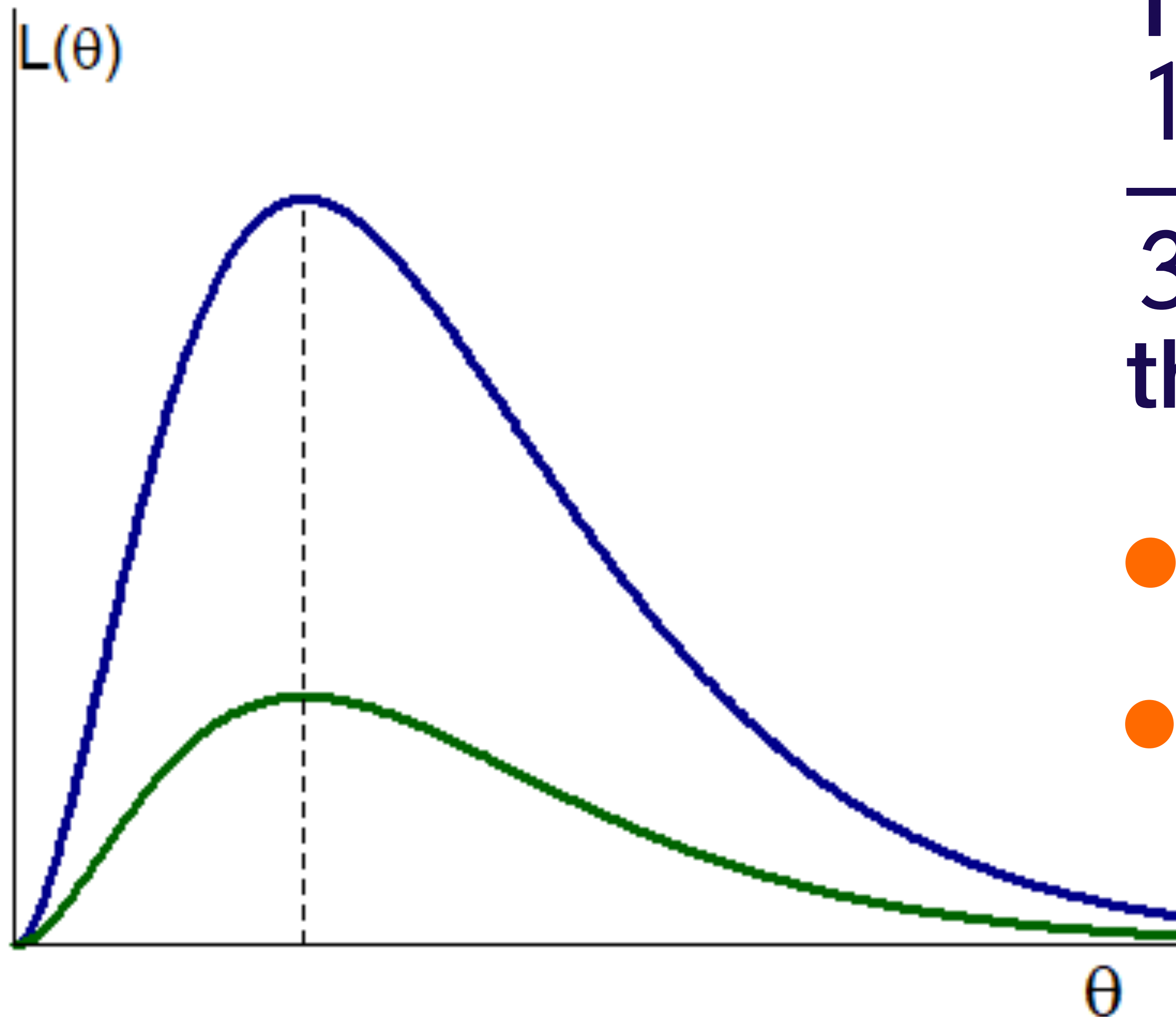
$$f(\vec{x}; \theta) = \prod_{i=1}^n f(x_i; \theta)$$

- The data (the  $x$ 's) are fixed.
- Think of the  $x$ 's as fixed and the joint pdf as a function of  $\theta$ .

Call it a **likelihood function** and denote it by  $L(\theta)$ .

# The likelihood function $L(\theta)$ :

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The maximum of  $\frac{1}{3} L(\theta)$  occurs in the same place.

- Same for  $3L(\theta)$ .
- Same for  $x_1 L(\theta)$ .

The likelihood  $L(\theta)$  is defined to be anything proportional to the joint pmf/pdf.

## Example:

$$X_1, X_2, \dots, X_n \stackrel{\text{iid}}{\sim} \text{Bernoulli}(p)$$

The pmf for one of them is

$$f(x; p) = p^x (1 - p)^{1-x} \mathbf{I}_{\{0,1\}}(x)$$

The joint pmf for all of them is

$$f(\vec{x}; p) = \prod_{i=1}^n f(x_i; p)$$

The parameter space is  $[0,1]$ .

$$= \prod_{i=1}^n p^{x_i} (1 - p)^{1-x_i} \mathbf{I}_{\{0,1\}}(x_i)$$



$$f(\vec{x}; p)$$

$$= p^{\sum_{i=1}^n x_i} (1 - p)^{n - \sum_{i=1}^n x_i} \prod_{i=1}^n \mathbb{I}_{\{0,1\}}(x_i)$$

A likelihood is

$$L(p) = p^{\sum_{i=1}^n x_i} (1 - p)^{n - \sum_{i=1}^n x_i}$$

It is almost always easier to minimize the “log-likelihood”:

$$\ell(p) = \ln L(p) = \sum_{i=1}^n x_i \ln p + (n - \sum_{i=1}^n x_i) \ln(1 - p)$$

$$\ell(p) = \sum_{i=1}^n x_i \ln p + (n - \sum_{i=1}^n x_i) \ln(1 - p)$$

$$\frac{\partial}{\partial p} \ell(p) = \frac{\sum_{i=1}^n x_i}{p} - \frac{n - \sum_{i=1}^n x_i}{1 - p} \stackrel{\text{set}}{=} 0$$

$$p(1 - p) \left[ \frac{\sum_{i=1}^n x_i}{p} - \frac{n - \sum_{i=1}^n x_i}{1 - p} \right] = p(1 - p) \cdot 0$$

$$(1 - p) \sum_{i=1}^n x_i - p \left( n - \sum_{i=1}^n x_i \right) = 0$$



$$\sum_{i=1}^n x_i - p \sum_{i=1}^n x_i - np + p \sum_{i=1}^n x_i = 0$$

$$\Rightarrow p = \frac{\sum_{i=1}^n x_i}{n}$$

The maximum likelihood estimator  
for  $p$  is:

$$\hat{p} = \frac{\sum_{i=1}^n X_i}{n} = \bar{X}$$

**Uppercase!**

## Continuous Example:

$$X_1, X_2, \dots, X_n \stackrel{\text{iid}}{\sim} \exp(\text{rate} = \lambda)$$

The pdf for one of them is

$$f(x; \lambda) = \lambda e^{-\lambda x} \mathbf{I}_{(0, \infty)}(x)$$

The joint pdf for all of them is

$$\begin{aligned} f(\vec{x}; \lambda) &= \prod_{i=1}^n f(x_i; \lambda) \\ &= \prod_{i=1}^n \lambda e^{-\lambda x_i} \mathbf{I}_{(0, \infty)}(x_i) \end{aligned}$$

The parameter space is  $(0, \infty)$ .

$$f(\vec{x}; \mathbf{p}) = \lambda^n e^{-\lambda \sum_{i=1}^n x_i} \prod_{i=1}^n I_{(0, \infty)}(x_i)$$

A likelihood is

$$L(\lambda) = \lambda^n e^{-\lambda \sum_{i=1}^n x_i}$$

The log-likelihood is

$$\ell(\lambda) = n \ln \lambda - \lambda \sum_{i=1}^n x_i$$

$$\ell(\lambda) = n \ln \lambda - \lambda \sum_{i=1}^n x_i$$

$$\frac{\partial}{\partial \lambda} \ell(\lambda) = \frac{n}{\lambda} - \sum_{i=1}^n x_i \stackrel{\text{set}}{=} 0$$

$$\Rightarrow \lambda = \frac{n}{\sum_{i=1}^n x_i}$$

Same as  
method of  
moments.  
Biased! 🤔

The maximum likelihood estimator  
for  $\lambda$  is:

$$\hat{\lambda} = \frac{n}{\sum_{i=1}^n X_i} = \frac{1}{\bar{X}}$$

Uppercase!