$$\begin{array}{c}
1 \\
\lim - 0 \\
n \to \infty
\end{array}$$

$$\lim_{n\to\infty} \left(\frac{1}{2}\right)^n = 0$$

$$\lim_{n \to \infty} \frac{2n^2 + 3}{3n^3 - 7} = 0$$

$$\lim_{n\to\infty} \frac{2n^2 + 3}{3n^2 - 7} = \frac{2}{3}$$

$$a_{n} = \frac{n-1}{n} \Rightarrow \lim_{n \to \infty} a_{n} = 1$$

$$a_{n} = \frac{n-1}{n} \Rightarrow \lim_{n \to \infty} a_{n} = 1$$

Consider a sequence of random variables: X₁, X₂, X₃, ...

$$\lim_{n\to\infty} X_n = ?$$

This is meaningless!

Convergence in Probability

The sequence of random variables:

converges in probability to a random variable X if, for any $\varepsilon > 0$

$$\lim_{n\to\infty} P(|X_n - X| > \varepsilon) = 0$$

We write $X_n \stackrel{P}{\rightarrow} X$.

An Integral Notation:

Rewrite
$$\int_{0}^{2} f(x) dx = \int_{A} f(x) dx$$

where $A=\{x: 0\leq x\leq 2\}$.

Suppose that X has the exponential distribution with rate λ.

How can we find $P(|\sin(X)|>1/2)$?

How can we find P(|sin(X)|>1/2)?

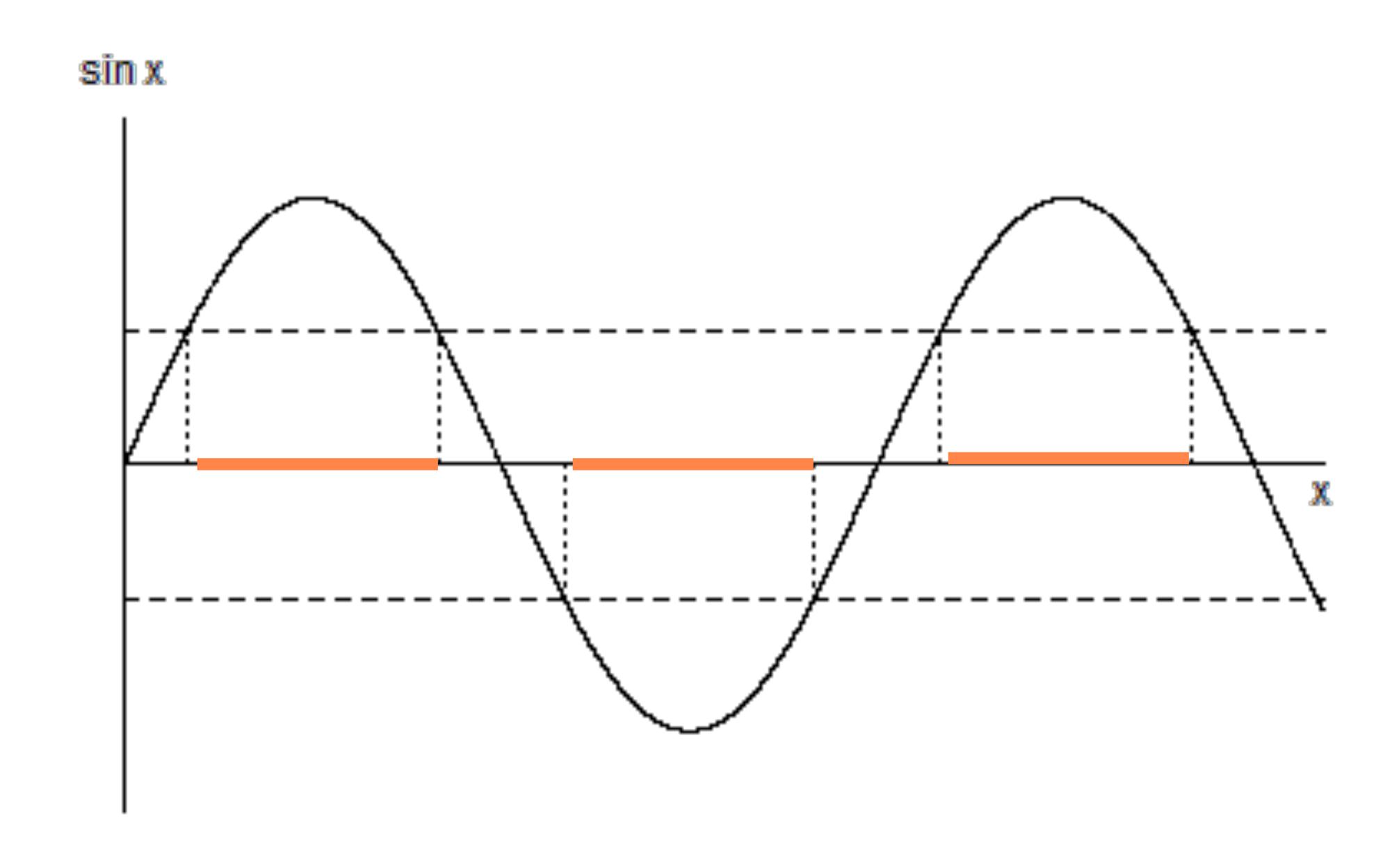
We can define a new random variable Y=|sin(X)|, try to find its pdf and then

$$P(|\sin(X)| > 1/2) = P(Y > 1/2)$$

$$= \int_{1/2}^{\infty} f_{Y}(y) \, dy$$

How can we find $P(|\sin(X)|>1/2)$?

2. We can integrate the <u>pdf for</u> X over the relevant region.



How can we find $P(|\sin(X)|>1/2)$?

2. We can integrate the pdf for X over the relevant region.

$$P(|\sin(X)| > 1/2)$$

$$= \int_{\sin^{-1}(1/2)}^{\pi - \sin^{-1}(1/2)} f_X(x) dx + \int_{\pi + \sin^{-1}(1/2)}^{2\pi - \sin^{-1}(1/2)} f_X(x) dx + \cdots$$

Notation
$$=\int f_X(x) dx$$

 $\{x: |\sin(x)| > 1/2\}$

Let X be a random variable. Let g be a non-negative function and let c>0.

Then

$$P(g(X) \ge c) \le \frac{E[g(X)]}{c}$$

When g(x)=|x|, this is known as Markov's inequality.

$$P(g(X) \ge c) \le \frac{E[g(X)]}{c}$$

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f_X(x) dx$$

$$= \int g(x)f_X(x) dx + \int g(x)f_X(x) dx$$

$$\{x : g(x) \ge c\} \qquad \{x : g(x) < c\}$$

$$P(g(X) \ge c) \le \frac{E[g(X)]}{c}$$

$$\begin{split} & E[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) \, dx \\ & = \int g(x) f_X(x) \, dx + \int g(x) f_X(x) \, dx \\ & \{x: g(x) \geq c\} \qquad \{x: g(x) < c\} \\ & \geq \int g(x) f_X(x) \, dx \geq \int c \, f_X(x) \, dx \\ & \{x: g(x) \geq c\} \qquad \{x: g(x) \geq c\} \end{split}$$

$$P(g(X) \ge c) \le \frac{E[g(X)]}{c}$$

$$E[g(X)] \ge \int c f_X(x) dx$$

$$\{x : g(x) \ge c\}$$

$$= c \int f(x) dx = c P(g(X) \ge c)$$

{x: g(x) \ge c}

Chebyshev's Inequality:

Let X be a random variable with mean μ and variance $\sigma^2 < \infty$. Let k > 0.

Then

$$P(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}$$

or, equivalently,

$$P(|X - \mu| < k\sigma) > 1 - \frac{1}{k^2}$$

"the probability that X is within k standard deviations of its mean"

Chebyshev's Inequality:

$$P(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}$$

Proof:

$$P(|X - \mu| \ge k\sigma) = P((X - \mu)^2 \ge k^2\sigma^2)$$

$$g(x)$$
c

$$\leq \frac{\mathsf{E}[\mathsf{g}(\mathsf{X})]}{\mathsf{c}} = \frac{\mathsf{E}[(\mathsf{X} - \mu)^2]}{\mathsf{k}^2 \sigma^2} = \frac{\sigma^2}{\mathsf{k}^2 \sigma^2}$$

$$= 1/k^2$$

Convergence in Probability $(X_n \xrightarrow{P} X)$

The sequence of random variables:

converges in probability to a random variable X if, for any $\varepsilon > 0$

$$\lim_{n\to\infty} P(|X_n - X| > \varepsilon) = 0$$

could be a constant!

Ex:
$$X = 3 \text{ w.p.1}$$

$$P(X = x) = \begin{cases} 1, & \text{if } x=3\\ 0, & \text{otherwise} \end{cases}$$

Suppose that X_1, X_2, X_3, \ldots is a sequence of iid random variables from any distribution with mean μ and variance $\sigma^2 < \infty$.

Then

- $\mathbf{E}[\mathbf{X}] = \mu$
- $Var[\overline{X}] = \sigma^2/n$

Proof: Let ε>0.

Choose K so that this

Chebyshev:
$$P(|X - \mu| \ge k\sigma) \le \frac{1}{k^2}$$

Here:

$$P(|\overline{X} - \mu_{\overline{X}}| \ge k\sigma_{\overline{X}}) \le \frac{1}{k^2}$$

Which is:

$$P(|\overline{X} - \mu| \ge k\sigma/\sqrt{n}) \le \frac{1}{k^2}$$

$$P(|\overline{X} - \mu| \ge \varepsilon) \le \frac{1}{(\varepsilon \sqrt{n/\sigma})^2} = \frac{\sigma^2}{\varepsilon^2 n}$$

$$\lim_{n\to\infty} P(|\overline{X} - \mu| \ge \varepsilon) \le \lim_{n\to\infty} \frac{\sigma^2}{\varepsilon^2 n} = 0$$

$$\lim_{n\to\infty} P(|\overline{X} - \mu| \ge \varepsilon) = 0$$

$$\lim_{n\to\infty} P(|\overline{X} - \mu| \ge \varepsilon) = 0$$

$$\downarrow \downarrow$$

Example:

$$X_1, X_2, X_3, \dots \stackrel{iid}{\sim} \exp(\text{rate} = \lambda)$$

$$\frac{P}{X} \rightarrow 1/\lambda$$

$$X_1, X_2, X_3, \dots \stackrel{\text{iid}}{\sim} \Gamma(\alpha, \beta)$$

$$X \rightarrow \alpha/\beta$$