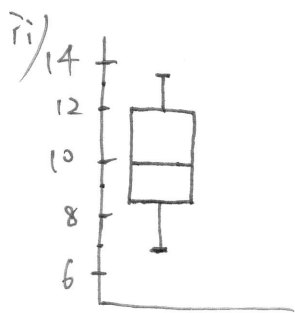


Stats Sample 3.

Q1. a) i) 6.9 7.9 8.8 9.3 10.6 11.3 12.7 13.2

{ 6.9, 8.35, 9.95, 12, 13.2 }



fairly symmetric

No outlier

$$1.5 \text{ iqr} = 3.65 \times 1.5 = 5.475$$

$$8.35 - 5.475 = 2.875$$

$$12 + 5.475 = 17.475$$

median 10, range 6.9 ~ 13.2

iii) D. small standard deviation without any outlier.

iv) Normally distributed.

$$\begin{aligned} v) \mu &= \bar{x} \pm t_{7; 0.975} \frac{s}{\sqrt{n}} \\ &= 10.0875 \pm 2.365 \times \frac{2.2510}{\sqrt{8}} \\ &= 10.0875 \pm 1.8822 \\ &= [8.2053, 11.9697] \end{aligned}$$

vi) $H_0: \mu = 9$ against $H_a: \mu > 9$

$$\begin{aligned} t_0 &= \frac{\sqrt{n}(\bar{x} - \mu_0)}{s} \sim t_{n-1} \\ &= \frac{\sqrt{8}(10.0875 - 9)}{2.2510} \sim t_7 \\ &= 1.366. \end{aligned}$$

$$t_{7; 0.95} = 1.895.$$

$$1.366 < 1.895$$

\therefore Do not reject H_0 .

Therefore we cannot contradict that the mean ozone concentration in this city does not exceed 9 ppm.

vii) P-value = $P(T > 1.366)$

from table $0.10 < \text{p-value} < 0.15$, means we can reject H_0 at any significance level greater than or equal to 0.15.

This is bigger than α , 0.05, So we do not reject H_0 .

If we reject H_0 , we would have 10% ~ 15% chance of being wrong.

b). See sample 1 Q2 b).

Q2. a) i) independence assumption
normality assumption.
equal variance assumption

ii) quantile plot — dots lies around a straight line shows normality.
plotting ~~residual~~ residual against ~~time~~ ~~order~~ ^{fitted value}. No pattern should be observed (independence)
plotting residuals against fitted value. No pattern shows equal variance.

b). 14.8182. ~~14.35~~

c) 99.6%

d) 0.998

e) 3.646

f) $H_0: \beta_1 = 0$ against $H_a: \beta_1 \neq 0$.

~~Reject H_0 if $\hat{\beta}_1 \notin [-t_{10, 0.975} \frac{s}{\sqrt{S_{xx}}}, t_{10, 0.975} \frac{s}{\sqrt{S_{xx}}}]$~~

$t_0 = 48.6 \sim t_{10}$

Reject H_0 if p-value $< \alpha^{0.05}$.

p-value = $2 \times P(T > 48.6) < 2 \times 0.0005 = 0.001 < 0.05$

\therefore Reject H_0 .

Therefore, ~~there~~ ^X has significant impacts to Y.

g) $\beta_1 = \hat{\beta}_1 \pm t_{10, 0.975} \times SE \text{ coef}$
 $= 14.8182 \pm 2.228 \times 0.3049$
 $= 14.8182 \pm 0.6793$
 $= [14.1289, 15.4975]$

h) We have 0.081 chance of being wrong if we reject the hypothesis that there is an ^X intercept ~~at X=0~~.

i). $\hat{y} = 4.348 + 14.8182 \times 7.5 = 115.3845$.

$[\hat{y} \pm \overset{3.646}{t_{10, 0.975}} \sqrt{1 + \frac{1}{12} + \frac{(7.5 - 6.5)^2}{143}}] = [115.3845 \pm 0.9805]$
 $= [114.404, 116.365]$

5. a). Independence

Normality

Equal variance.

b).

Source	df	SS	MS	F
Factor	2	165.778	82.889	13.04
Error	15	95.333	6.356	
Total	17	261.111		

c). μ_1, μ_2, μ_3 be ^{true} mean tensile strength for different concentration

$H_0: \mu_1 = \mu_2 = \mu_3$ against H_a : Not all ~~the~~ means equal.

observed $f_0 = 13.04$, following $F_{2;15}$

$$f_{2;15;0.95} = 3.68 < 13.04.$$

Reject ~~if~~ ^{H_0} ~~$f_0 > f_{2;15;0.95}$~~ since $f_0 > f_{2;15;0.95}$.

p-value ~~is~~ = $P(X > \overset{13.04}{f_0})$ $X \sim F_{2;15}$.

$$p\text{-value} < 0.005 < 0.05.$$

\therefore There is significant difference in ~~con~~ tensile strength at different concentration.

d)

$$\begin{aligned} CI &= \left[\bar{x}_2 - \bar{x}_3 \pm t_{15;0.975} \sqrt{MS_{ER} \left(\frac{1}{n_2} + \frac{1}{n_3} \right)} \right] \\ &= \left[-1.33 \pm 2.131 \times \sqrt{6.356 \times \left(\frac{1}{6} + \frac{1}{6} \right)} \right] \\ &= [-1.33 \pm 3.102] \\ &= [-4.432, 1.772] \end{aligned}$$

No. since $0 \notin [-4.432, 1.772]$

e).

$$\frac{0.05}{3} = 0.0167$$

$0.0004 < 0.0167$. \therefore Reject $H_0: \mu_1 = \mu_3$.

\therefore Not same conclusion.