

Introducing the OccupancyModels package

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The use of occupancy models is quite common, but the expertise to write a Bayesian model to do the analysis is a substantial burden. The package `unmarked` has helped, but it lacks some functionality, particularly for analyses that have an availability component. While the primary focus of the `OccupancyModels` package is to provide a convenient R interface to Bayesian models that include multiple detection devices at a site, we also include this missing functionality.

This vignette is primarily designed to introduce the relevant models with appropriate references, and demonstrate fitting them in the `OccupancyModel` package. If package `unmarked` contains the functionality to fit a particular model, that process will also be demonstrated.

Software issues

The package was designed to use `stan`. To learn more about `stan` and install it and the package `rstan` go to <https://github.com/stan-dev/rstan/wiki/RStan-Getting-Started>.

Introduction stuff

- Ψ = probability that a site is occupied
 - Occupied means species use on at least one day
- $y_{ij} = 1$ denotes a detection at site i on day j
- $Z_i = 1$ denotes if a site is actually occupied
 - Latent variable and not necessarily known (unless a detection is made)
- π is the probability of detection given the the site is occupied
 - $P(y_{ij} | Z_i = 1) = \pi$

Occurrence Models

We will use the terminology of “Occurrence Models” to address a class of models introduced by Mackenzie et al 2002.

- $\Psi_i = \text{logit}^{-1}(X_i^T \beta)$ where the X_i covariates represent site level information that increase or decrease the occurrence probability and the estimation of the β values is the primary interest of the study.
- $Z_i \sim \text{Bern}(\Psi_i)$
- $y_{ij} \sim \text{Bern}(Z_i \cdot \pi)$
- Assumes the detection probability is constant across days.

- Inappropriate in many cases
- Can aggregate multiple days into multiple sampling periods
 - * What is an appropriate number of days to aggregate?
- Can be fit using **unmarked**

Occupancy / Availability models

- Similar to models in Nichols et al, 2008
- Can have a second latent variable that indicates daily availability
 - $\theta_j = 1$ probability an animal is available on day j
 - $W_{ij} = 1$ (latent variable denoting if an animal is available at site i on day j)
- $\Psi_i = \text{logit}^{-1}(X_i^T \beta)$
- $Z_i \sim \text{Bern}(\Psi_i)$
- $W_{ij} \sim \text{Bern}(Z_i \cdot \theta_j)$
- $y_{ij} \sim \text{Bern}(W_{ij} \cdot \pi)$
- Assumptions
 - Availability on day j is the same across sites.
 - Constant value of availability better?

To fit these models, users have had to write their own code in WinBUGS, JAGS, or stan. Fortunately, we have a program to do most of this automatically.

No covariates

We first introduce the produce using a model with no covariates. That is to say that Ψ_i is constant across sites. In our model, this can be coded as $X_i = 1$ for all i and $\Psi_i = \text{logit}^{-1}(\beta)$.

```

# tools that allow me to download from GitHub
library(devtools)
install_github('dereksonderegger/OccupancyModels')
library(OccupancyModels)

# Use the make.data function to generate some simulated data
# Check out the help file!
# ?make.data
temp <- make.data(
  Occupancy.p = .4, # thus beta = logit(.4) = -.405
  Available.p = .5,
  Detection.p = .7,
  n.sites = 40,
  n.days = 5,
  n.cameras = 3)

# What have we made...
head(temp)

##   site day camera Detections
## 1    1  1      1           0
## 2    1  1      2           0
## 3    1  1      3           0
## 4    1  2      1           0
## 5    1  2      2           0
## 6    1  2      3           0

```

The first step in the analysis is to take our detection data and turn it into a 3-dimensional array, where the dimensions correspond to **site**, **day**, and **camera**.

```

Y <- acast(temp, site~day~camera, value.var='Detections')
str(Y)

##   int [1:40, 1:5, 1:3] 0 0 1 0 0 0 0 0 0 0 ...
##   - attr(*, "dimnames")=List of 3
##   ..$ : chr [1:40] "1" "2" "3" "4" ...
##   ..$ : chr [1:5] "1" "2" "3" "4" ...
##   ..$ : chr [1:3] "1" "2" "3"

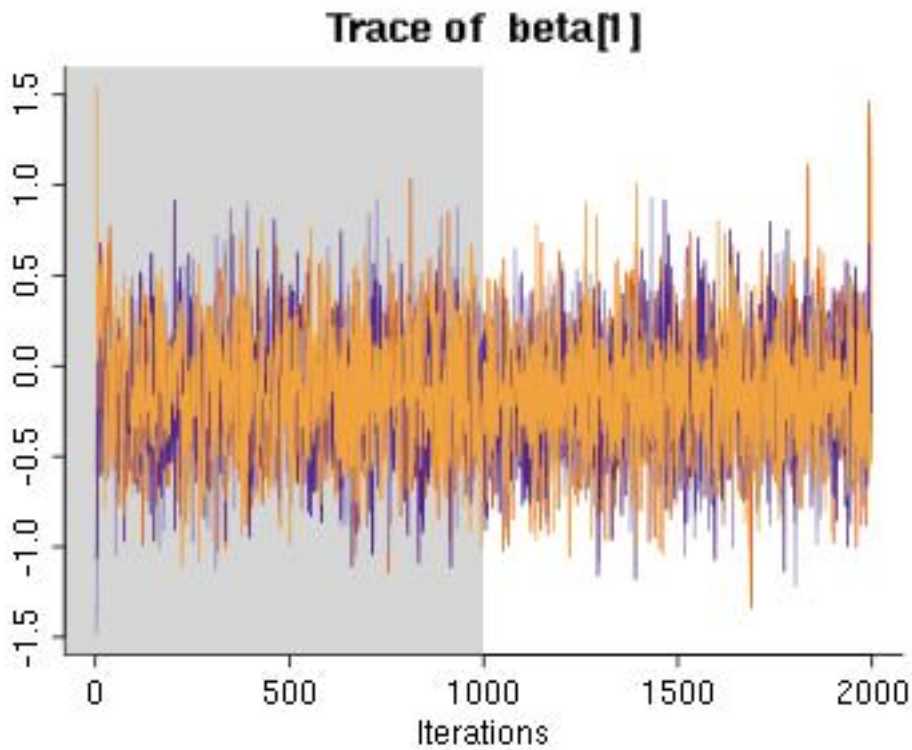
```

Next we can call the multiple detection occupancy model `mdOcc`.

```
model <- mdOcc(Y) # No covariates
```

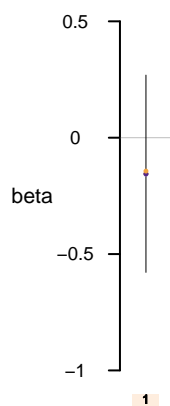
We have now run the model through `stan` and can inspect the output. In particular I want to look at the values for β which is the logit transformed occupancy probability.

```
traceplot(model, pars='beta')
```



```
plot(model, pars='beta')
```

an model 'md_Covariate' (4 chains: iter=2000; warmup=1000; thin=1) fitted at Wed Oct 1 00:34:06 20
medians and 80% intervals



Rhat: 1 < 1.1 2 < 1.2 3 < 1.5 4 < 2 5 >= 2 6 NaN/Inf

We can do our own summary statistics by extracting the chains for the β value (which in this simple case is just the inverse logit of the occupancy probability).

```

beta.posterior <- extract(model, pars='beta')[[1]]
output <- apply( inv.logit( beta.posterior ),
  MARGIN=2,
  quantile, c(.025, .25, .50, .75, .975) )
output

##
##           [,1]
## 2.5%  0.3104
## 25%   0.4073
## 50%   0.4617
## 75%   0.5163
## 97.5% 0.6181

```

This shows us that the posterior distribution of Ψ , the occupancy probability, is between 0.32 and 0.64, and this interval happily includes the true value of $\Psi = 0.4$.

With Covariates

We next allow there to be covariates in the model. In this case, we'll create data with one covariate and set a positive relationship between the covariate and Ψ .

```

# Number of plots
n <- 50

# make up some site level covariates
site.data <- data.frame(x=rnorm(n, 0, 2))

# relationship between x and Psi
psi <- inv.logit( 0 + (1/2)*site.data$x )

# Make data
sim.data <- make.data(n, n.days=20, n.cameras=3,
  Occupancy.p=psi, Available.p=.5, Detection.p=.4 )

# reshape the data frame into a 3D array
Y <- acast(sim.data, site~day~camera, value.var='Detections')

# Fit the model using the site level covariates
model <- mdOcc(Y, site.data, ~ x)

```

Using the `extract` function as we did before, we obtain the posterior distribution of the β values and also exam the logit^{-1} transformed values as well.

```

beta.posterior <- extract(model, pars='beta')[[1]]
beta.summary <- apply( beta.posterior, MARGIN=2,
                        quantile, c(.025, .25, .50, .75, .975) )
inv.logit.beta.summary <- apply( inv.logit( beta.posterior ), MARGIN = 2,
                                quantile, c(.025, .25, .50, .75, .975) )

beta.summary

##
##           [,1]  [,2]
##  2.5%  -0.50178  0.2876
##  25%   -0.07687  0.4901
##  50%    0.15689  0.6277
##  75%    0.39556  0.7732
##  97.5%  0.87446  1.1005

inv.logit.beta.summary

##
##           [,1]  [,2]
##  2.5%   0.3771  0.5714
##  25%   0.4808  0.6201
##  50%   0.5391  0.6520
##  75%   0.5976  0.6842
##  97.5% 0.7057  0.7504

```