

Flipping coin large no. of times or head = $\frac{1}{2}$ → limiting value

- set element has some properties :-
- 1) ordering does not matter in sets
- 2) equal $A = B \Leftrightarrow A \subset B \text{ and } B \subset A$
- 3) S or Ω (universal sets), sample space
mutually exclusive / disjoint

Partition : disjoint + union gives back original cardinality of set = no. of elements in set

$$A = \{2, 3\} \text{ cardinality } |A| = 2$$

$$|A \cup B| = |A| + |B| - |A \cap B|$$

$$|A \cup B \cup C| = |A| + |B| + |C| - |A \cap B| - |A \cap C| - |B \cap C| + |A \cap B \cap C|$$

general →

$$\begin{aligned} & |A| + |B| - \\ & - |A \cap B| - \\ & + |A \cap B \cap C| - \\ & - |A \cap B \cap C \cap D| + |A \cap B \cap C \cap D \cap E| - \end{aligned}$$

countable & uncountable sets :

A set S is said to be countable

- 1) if it is finite
- 2) if there is a one-one and onto correspondence with the set of natural numbers N .

$$f: N \rightarrow S$$

e.g. 1 set of all even positive integers →

$$f: N \rightarrow S$$

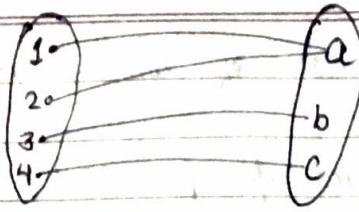
$$f(n) = 2n, \quad \forall n \in N$$

one-one for two diff o/p we should have
2 diff diff i/p

for one-one
for $f(a) = f(b)$
then $a = b$

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$$f(1) = f(2)$$

but $1 \neq 2$

so it is not one-one

counter part

$1 \neq 2$, then prove $f(1) \neq f(2)$

continuing eq \rightarrow let $f(n) = f(m)$

$$2n = 2m$$

one-one

$$n = m$$

so thing is that we should make a right function in this it is $f(n) = 2n$

each & every entry should be mapped

For Onto :- if m is even then

$$\boxed{f\left(\frac{m}{2}\right) = f(m)}$$

now this is countable

eg. 2 set of all integers :-

$$f: N \rightarrow \mathbb{Z}$$

$$f(n) = \begin{cases} n/2 & \text{if } n \text{ is even} \\ -\frac{(n-1)}{2} & \text{if } n \text{ is odd.} \end{cases}$$

let $n \neq m$

$$\frac{n}{2} \neq \frac{m}{2}$$

$$-\frac{(n-1)}{2} \neq -\frac{(m-1)}{2}$$

for one-one

$$\boxed{f(n) \neq f(m)}$$

we used counter-part.

If we take $\rightarrow f(n) = f(m)$

1) when $m \& n$ both are odd

2) when $m \& n$ both are even

2nd Argument

In exam just prove this by one argument
only

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- 3) m is even, n is odd
4) m is odd, n is even

Taking case - m is odd, n is even

$$-\frac{(m-1)}{2} = \frac{n}{2} \quad m+n=1$$

m and n should be natural nos. Min possible values are 1 and 2 $m+n=3$ \times so this case is not possible

→ Taking case n odd, m even

$$-\frac{(n-1)}{2} = \frac{m}{2} \quad m+n=1$$

same as above

→ If both are odd

$$f(m) = f(n)$$

$$-\frac{(m-1)}{2} = -\frac{(n-1)}{2} \quad m=n$$

→ if both are even $m=n$

so in this we need to consider all pairs

onto :-

If $m=0$
let $m \in \mathbb{Z}$ if $m > 0$
if $m < 0$

This is basically about
the o/p

$$\begin{aligned} f(\underline{m}) &= 0 && \text{In this we} \\ f(2m) &= m && \text{covered all the} \\ f(1-2m) &= -\frac{1-2m-1}{2} && \text{Integers} \\ i/p &= m \end{aligned}$$

Total value is going to be positive
value

so this is
onto

Q3 Set of positive rational numbers is
countable?

set of

what kind of function from natural no:
to set of rational numbers.

solution

$N \rightarrow$ Numerator

$N+D = 2$ in $\frac{1}{1}$

$D \rightarrow$ Denominator

$N+D$ is 3 in $\frac{1}{2}, \frac{2}{1}$

$N+D$ is 4 in $\frac{1}{3}, \frac{3}{1}, \frac{2}{2}$

$N+D$ is 5 in $\frac{1}{4}, \frac{4}{1}, \frac{2}{3}, \frac{3}{2}$

$N+D$ is 6 in $\frac{1}{5}, \frac{5}{1}, \frac{3}{3}, \frac{2}{4}, \frac{4}{2}$

$\frac{1}{1}, \frac{1}{2}, \frac{2}{1}, \frac{1}{3}, \frac{2}{2}, \frac{3}{1}, \frac{1}{4}, \frac{2}{3}, \frac{3}{2}, \frac{4}{1}$

Arranging them

$\frac{1}{1}$	$\frac{1}{2}$	$\frac{1}{3}$	$\frac{1}{4}$	$\frac{1}{5}$	$\frac{1}{6}$...
$\frac{2}{1}$	$\frac{2}{2}$	$\frac{2}{3}$	$\frac{2}{4}$	$\frac{2}{5}$	$\frac{2}{6}$...
$\frac{3}{1}$	$\frac{3}{2}$	$\frac{3}{3}$	$\frac{3}{4}$	$\frac{3}{5}$	$\frac{3}{6}$...
$\frac{4}{1}$	$\frac{4}{2}$	$\frac{4}{3}$	$\frac{4}{4}$	$\frac{4}{5}$	$\frac{4}{6}$...

In this manner we are able to cover all Rational numbers

eg 4. set of all rational numbers countable?

eg 5. set of Real numbers countable?

Real = union of rational and irrational

sample space :- set of all possible outcomes of a statistical experiment.

each outcome is called a sample point.

3 ways to describe a sample space :-

list (for finite sample points)

True statement or rule

any set is subset of S

subsets :- Facts $A \subset S$, $\emptyset \subset S$

$S \subset S \rightarrow$ proper subset

Event :- Any subset of sample space S is an event

Complement event :- Set of all elements of A that are not present in A.

Basic Principle of counting :-

Multiplication Principle.

Manish Gang's class

5/1/18

Set : collection of well defined objects

eg $S = \{ \text{set of all presidents} \}$ not well defined not set

$S = \{ \text{set of all students} \}$ " "

$S = \{ 1, 2, 3, 4, 5, 6 \}$ set

$S = \{ \text{set of } 1 \text{ feet tall students in INNIT} \}$ empty set

operations $\rightarrow A \cup B$, $A \cap B$, $A - B$ { In A but not B }

In set repetition does not matter. Two sets are unequal if element in one is not present in another.

$$\{ 2, 4, 3 \} = \{ 2, 4, 3, 2 \}$$

For dice $S = \{ 1, 2, 3, 4, 5, 6 \}$ $\rightarrow 2$ (universal set)

For coin $S = \{ \text{head, tail} \}$

Countable set :- A set which is either finite or empty or countably infinite is called countable set otherwise called uncountable set
 $S_n : n \in \mathbb{N}$

Countable Union $\rightarrow \bigcup_{n=1}^{\infty} S_n = S_1 \cup S_2 \cup S_3 \dots \infty = [x : x \in S_n \text{ for some } n]$

Countable Intersection $\bigcap_{n=1}^{\infty} S_n = S_1 \cap S_2 \cap S_3 \dots \infty = [x : x \in S_n \text{ for some } n]$

if $\alpha \in (0, 1)$

$\bigcup_{\alpha \in (0, 1)} S_\alpha = [x : x \in S_\alpha \text{ for some } \alpha]$ Uncountable Union

$\bigcap_{\alpha \in (0, 1)} S_\alpha = [x : x \in S_\alpha \text{ for some } \alpha]$ Uncountable Intersection

Countable Infinite Set :- An infinite set S is called countable infinite if there exist a mapping $f : \mathbb{N} \rightarrow S$ s.t. f is one-one & onto.

eg Set of all positive integers :-

Define $f : \mathbb{N} \rightarrow S$ $f(n) = 2n$

one-one : $f(n) = f(m) \Rightarrow n = m$ onto :
 $2n = 2m \Rightarrow n = m$
 $\boxed{n = m}$

$x \in S \exists n \in \mathbb{N}$
s.t. $x = 2n$
 $\boxed{f(n) = x}$

S is countable infinite set

eg For $S = \text{set of all integers}$:-

Define $f : \mathbb{N} \rightarrow I$

$$f(n) = \begin{cases} n/2 & \text{if } n \text{ is even} \\ -(n-1) & \text{if } n \text{ is odd} \end{cases}$$

for n is even and n is odd

4 cases (It is done before)

Use another Method :-

one-one if $m \neq n$

$$\frac{m}{2} \neq \frac{n}{2}$$

$$-\frac{(m-1)}{2} \neq -\frac{(n-1)}{2} \Rightarrow f(m) \neq f(n)$$

for onto: $m \in I$; if $m = 0$ $f(1) = 0$

$$\text{if } m < 0 \quad f(1-2m) = m$$

$$\text{if } m > 0 \quad f(2m) = m$$

It is onto (every codomain element should have pre image in domain)

Combinational Analysis :-

Basic Principle :- Suppose k experiments are to be formed. Suppose exp 1, has ' m ' possible outcomes and for each possible outcomes of exp 1, exp 2 has ' n ' possible outcomes. Then there are mn possible outcomes.

$$(1,1) \quad (1,2) \quad (1,3) \quad \dots \quad (1,n)$$

$$(2,1) \quad (2,2) \quad (2,3) \quad \dots \quad (2,n)$$

:

$$(m,1) \quad (m,2) \quad (m,3) \quad \dots \quad (m,n)$$

Generalisation :- k - experiment

$$\text{exp 1} \rightarrow n_1$$

$$\text{exp } k \rightarrow n_k$$

$$\text{exp 2} \rightarrow n_2$$

$$\text{Total } n_1, n_2, \dots, n_k$$

Ques 1 :- A small community 10 women each of whom have 3 children. If one women and one of her children are chosen as mother and child of the year. Then how many diff choices are possible

$$\underline{\text{Sol}}' \quad \text{Total } 10 \times 3 = 30$$

Q2 A college planning committee consists of 3 freshman, 4-sophomores, 5-Juniors, 2-Seniors. A subcommittee of 4 consisting 1 person from each class is to be chosen. How many diff sub-committees combination are possible.

$$\text{Total } = 3 \times 5 \times 2 \times 4 = 120$$

Q-3

How many diff 7 place license plates are possible if the first 3 places are to be occupied by letters & final 4 by numbers.

Ans

$$(26)^3 \times (10)^4$$

If repetition not allowed \(\rightarrow 26 \times 25 \times 24 \times 10 \times 9 \times 8 \times 7\)

Q-4

How many different ordered arrangement of a, b, c are possible?

abc, acb, bac, bca, cab, cba

These arrangement is called permutation.

For n objects $n!$ way to arrange

Q-5

Miss John has 10 books that she is going to put on bookshelf. Of these 4 are of Maths, 3-Chem, 2-History, 1-lang, she wants to arrange s.t. all books dealing same sub are on same place. How many arrangements are possible?

Sol

$$4 \rightarrow \text{Maths} \rightarrow 4!$$

$$3 - \text{Chem} \rightarrow 3!$$

$$2 - \text{History} \rightarrow 2!$$

$$1 - \text{language} \rightarrow 1!$$

Total \rightarrow

$$(4! \times 3! \times 2! \times 1!) \times$$

arranging 4 ways of bundles.

Q-6

A class in probability consists of 6 men and 4 women. An examination is given, and the students are ranked according to their performance. Assume that no two students obtain the same score.

i)

How many different ranking are possible if men are ranked themselves and women are also ranked themselves. How many diff ranking are possible?

ii)

(i) $(10!)$

[Total ranking]

(ii) $6! \times 4!$

[Multiplication due to basic principle]

Result :- Suppose 'n' objects are given. Out of these n_1 are alike, n_2 are alike ... n_r are same. How many different arrangements are possible?

$$\text{Ans} = \left[\frac{(n!)}{n_1! n_2! \dots n_r!} \right] \quad \begin{array}{l} \text{diff arrangements} \\ \text{when some items} \\ \text{are identical} \end{array}$$

Q PEPPER How Many diff. arrangements are possible?

3 - P

2 - E

1 - R

$$\rightarrow \frac{6!}{3! 2! 1!} = \frac{2}{1} \times 5 \times 6 = 60$$

Q suppose a chess tournaments has 10 competitors of which 4 are Russian, 3 are from US, 2 are from England, and 1 is from Brazil. If the tournament results lists just the nationalities of the player in the order in which they plays. How Many outcomes are possible?

→ Result is in terms of their countries

10!

All 4 Russian can be identified as $4! 3! 2! 1!$
identical because result will be like *winner is from Russia*.

COMBINATIONS :-

How Many diff groups of 3 could be selected from 5 objects A, B, C, D, E ?

$$\text{Total } \underline{\underline{5}} \underline{\underline{4}} \underline{\underline{3}} = 60 \quad \underline{\underline{5 \times 4 \times 3}}$$

Ideal arrangement

- In that expression each selection is counted 6 times. ABC, ACB, BAC, BCA, CAB, CBA so when ordering counts ${}^5C_3 = \frac{5!}{3!} = 10$ ways

nC_u are the ways

$$0! = 1$$

$${}^nC_u = 0$$

$${}^nC_0 = 1 = {}^nC_n$$

$$\text{if } u < 0, u > n$$

- Q From a group of 5 women and 7 men how many different committees consisting of 2 women and 3 Men can be formed?
- (ii) what if 2 of the Men are feuding and refused to serve on the committee together?

Sol :-

W	M	Total
5	7	(i) ${}^5C_2 \times {}^7C_3$
2	3	ans = 350

(ii) How many to select 3 men from 7 7C_3
 & feuding $\frac{7 \times 6 \times 5}{3 \times 2} = 35$

Two person which cannot be selected together ${}^2C_2 \times {}^5C_1 = 5$ ways To select 1 more from selecting both of those two person men
 By this those both person will be in same committee ; Total ways $35 - 5 = 30$ ways
 30 ways when both feuding men will not be together.

$$\begin{aligned} \text{Total ways} &= {}^5C_2 \times 30 \\ &= 10 \times 30 = 300. \end{aligned}$$

Identity :-

$$[n-1]C_{r-1} + [n-1]C_r = [n]C_r \quad 1 \leq r \leq n$$

L.H.S

$$\begin{aligned} & [n-1]C_{r-1} + [n-1]C_r = \frac{(n-1)!}{(r-1)! (n-r)!} + \frac{(n-1)!}{(r)! (n-r-1)!} \\ &= \frac{(n-1)!}{(r-1)! (n-r-1)!} \left[\frac{1}{n-r} + \frac{1}{r} \right] \\ &= \frac{(n-1)!}{(r-1)! (n-r-1)!} \left[\frac{n}{(n-r)r} \right] = \frac{n!}{r! (n-r)!} \\ &= \underline{\underline{n}C_r} \end{aligned}$$

Result :-

$$(x+y)^n = \sum_{R=0}^n nC_R x^R y^{n-R} \quad - \textcircled{1}$$

Proof by M.I \rightarrow for $n=1$

$$\begin{aligned} \text{R.H.S} \rightarrow \sum_{R=0}^1 nC_R x^R y^{1-R} &= {}^1C_0 x^0 y + {}^1C_1 x^1 y^0 \\ &= \underline{\underline{x+y}} \end{aligned}$$

Suppose $\textcircled{1}$ is true for $(n-1)$, we need to prove it for n .

$$(x+y)^n = (x+y) (x+y)^{n-1}$$

$$(x+y)^n = (x+y) \sum_{R=0}^{n-1} n-1 C_R x^R y^{n-1-R}$$

$$= \sum_{R=0}^{n-1} n-1 C_R x^{R+1} y^{n-1-R} + \sum_{R=0}^{n-1} n-1 C_R x^R y^{n-R}$$

Putting $i=R+1$ in the first expression and
 $R=i-1$ in the second expression.

$$\sum_{i=1}^n n-1 C_{i-1} x^i y^{n-i} + \sum_{i=0}^{n-1} x^i y^{n-i}$$

$$\begin{aligned} & \sum_{i=1}^{n-1} n-1 C_{i-1} x^i y^{n-i} + n-1 C_{n-1} x^n y^0 + n-1 C_0 x^0 y^n \\ & + \sum_{i=1}^{n-1} x^i y^{n-i} \end{aligned}$$

$$\begin{aligned}
 & \rightarrow \sum_{i=1}^{n+1} [{}^n C_{i-1} + {}^n C_i] x^i y^{n-i} + x^n + y^n \\
 & \rightarrow x^n + \sum_{i=1}^{n+1} {}^n C_i x^i y^{n-i} + y^n \\
 & = \boxed{\sum_{i=0}^n {}^n C_i x^i y^{n-i}} = (x+y)^n
 \end{aligned}$$

Q How many subsets are there of a set consisting of n elements?

Ans Total subsets $\sum_{i=0}^n {}^n C_i = 2^n$

$$\begin{aligned}
 & {}^n C_0 + {}^n C_1 + {}^n C_2 + \dots + {}^n C_n \\
 & = (1+1)^n = 2^n
 \end{aligned}$$

* There are $(n-1)!$ arrangements of the n objects around the table.

Q A set of n objects is to be divided into r distinct groups of size n_1, n_2, \dots, n_r respectively such that $\sum_{i=1}^r n_i = n$

How many diff. divisions are possible?

It is diff
from previous
one.

$$sol = \frac{(n)!}{(n_1!)(n_2!) \dots (n_r)!}$$

Order is
important

Q 10 children are to be divided into a team A and team B of 5 each. The team A will play in one league and team B in another. How many diff. divisions are possible?

Ans

Here teams
are separated

$$\frac{10!}{5! 5!} =$$

10 players are
divided into 2 gms
of 5 each.

$$= \frac{10!}{2! 5! 5!}$$

We are not separating the gms. clearly

Q 10 - policemen
patrolling
5 men

station gp
3 men

Reserve
23 men

All these 3 groups are separated properly.

Then $\frac{10!}{5! 3! 2!}$

when Order is not Relevant Then $\frac{10!}{5! 3! 2!} \quad 3!$

Probability

sample space \rightarrow The set of all possible outcomes in an experiment is called sample space.

S or Ω

Tossing a coin $S = \{\{H\}, \{T\}\}$

Tossing of two coins $S = \{\{HH\}, \{HT\}, \{TH\}, \{TT\}\}$

3) Rolling of two dices $= S = \{(i, j) ; i, j = 1, 2, 3, 4, 5, 6\}$

4) If the experimental measures (in hours) the life of a transistor.

$$S = \{n : 0 \leq n < \infty\}$$

Event :- A subset of a sample space is called event. 'E'.

ex 1) $E = \{\{H\}\}$

2) $E = \{\{HT\}, \{TH\}\}$

3) $E = \{(4,6), (5,5), (6,4)\}$

→ If the transistor does not last more than 4 hours.

Than event is

$$E = \{n ; 0 \leq n \leq 4\}$$

→ Let we have 2 events E and F.
Then an another event EUF → contain all outcomes which exist in either E or F.

$$E = \{ (H, H), (T, H) \}$$

$$F = \{ (H, T) \}$$

$$\underline{EUF} = \{ (H, H), (T, H), (H, T) \}$$

at least one head occurs.

ENF → contains all outcomes present in both E and F.

$$E = \{ (H, H), (T, H), (H, T) \} \xrightarrow{1} \text{at least one head}$$

$$F = \{ (T, H), (T, T), (H, T) \} \xrightarrow{2} \text{at least one tail}$$

$$ENF = \{ (H, T), (T, H) \}$$

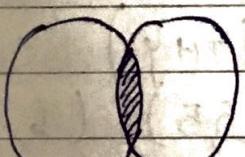
→ Intersection can also be written as EF or when there events are present ENFNL → EFL

* Mutually Exclusive Events :- Two events are called mutually exclusive if $ENF = \emptyset$

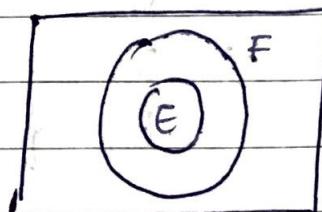
eg A and A^c

* ECF → all the outcomes of E and present in F.

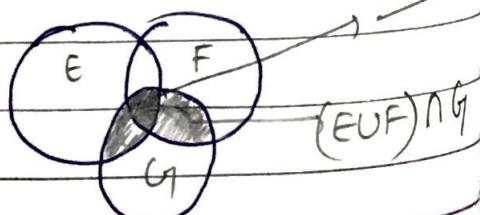
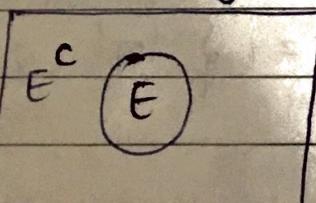
Venn diagram :-



ENF
S



ENFAG



(EUF)AG

some Results :-

$$\left(\bigcup_{i=1}^n E_i \right)^c = \bigcap_{i=1}^n E_i^c \quad \text{--- } ①$$

$$\left(\bigcap_{i=1}^n E_i \right)^c = \bigcup_{i=1}^n E_i^c \quad \text{--- } ②$$

De Morgan's Law

Proof of 1st $\rightarrow x \in \left(\bigcup_{i=1}^n E_i \right)^c$

$$\Rightarrow x \notin \bigcup_{i=1}^n E_i$$

$\Rightarrow x \notin E_i^c$ for any $i = 1, 2, 3 \dots n$.

$\Rightarrow x \in E_i^c$ for all $i = 1, 2, 3 \dots n$.

$$\Rightarrow x \in \bigcap_{i=1}^n E_i^c$$

$$\cdot \left(\bigcup_{i=1}^n E_i \right)^c \subseteq \bigcap_{i=1}^n E_i^c \quad \text{--- } ③$$

$$\rightarrow \text{let } x \in \bigcap_{i=1}^n E_i^c$$

$x \in E_i^c$, for all $i = 1, 2, \dots n$

$\Rightarrow x \notin E_i$ for any $i = 1, 2, \dots n$

$$x \notin \bigcup_{i=1}^n E_i$$

$$x \in \left(\bigcup_{i=1}^n E_i \right)^c$$

$$\bigcap_{i=1}^n E_i^c \subseteq \left(\bigcup_{i=1}^n E_i \right)^c \quad \text{--- } ④$$

$$\boxed{\left(\bigcup_{i=1}^n E_i \right)^c = \bigcap_{i=1}^n E_i^c}$$

OR we can do it \rightarrow

$$\left[\left(\bigcup_{i=1}^n E_i \right)^c \right]^c = \bigcap_{i=1}^n (E_i^c)^c$$

$$\left(\bigcup_{i=1}^n E_i^c \right)^c = \bigcap_{i=1}^n E_i \quad \xrightarrow{\text{Taking complement again}}$$

$$\left(\bigcap_{i=1}^n E_i \right)^c = \bigcup_{i=1}^n E_i^c$$

* Probability :- let S be a sample space & E be the event. we define a number $P(E)$ [called the probability of event $\{E\}$] if it satisfy the following axioms.

1) $0 \leq P(E) \leq 1$

2) $P(S) = 1$

3) For any sequence of Mutually exclusive events $E_1, E_2, \dots, E_n \dots$ $[E_i \cap E_j = \emptyset, i \neq j]$

$$P\left(\bigcup_{i=1}^{\infty} E_i\right) = \sum_{i=1}^{\infty} P(E_i)$$

ex $P(E^c) = 1 - P(E)$

$$E \cap E^c = \emptyset \quad (\text{mutually exclusive})$$

$$E \cup E^c = S$$

$$P(E \cup E^c) = P(S)$$

$$= 1$$

By Axiom 3 we have

$$P(E) + P(E^c) = 1$$

$$\boxed{P(E^c) = 1 - P(E)}$$

result $\boxed{P(\emptyset) = 0}$

we can write $E \Rightarrow E \cup \emptyset$

$$P(E) = P(E \cup \emptyset)$$

$$P(E) = P(E) + P(\emptyset)$$

$$\boxed{P(\emptyset) = 0}$$

→ if $E \subset F$ then $P(E) \leq P(F)$

since $E \subset F$

$$F = E \cup (E^c \cap F)$$

$$P(F) = P(E \cup (E^c \cap F))$$

$$P(F) = P(E) + P(E^c \cap F)$$

$$P(F) \geq P(E) \quad \text{from Axiom ①}$$

$$\rightarrow P(E \cup F) = P(E) + P(F) - P(E \cap F)$$

$$E \cup F = E \cup (E^c \cap F)$$

$$P(E \cup F) = P(E) + P(E^c \cap F) \quad \text{--- ①}$$

$$F = (E \cap F) \cup (E^c \cap F)$$

$$P(F) = P(E \cap F) + P(E^c \cap F) \quad \text{--- ②}$$

② in ①

$$P(E \cup F) = P(E) + P(F) - P(E \cap F)$$

→ suppose E_1, E_2, \dots, E_n are events

$$P(E_1 \cup E_2 \cup \dots \cup E_n) = \sum_{i=1}^n P(E_i) -$$

$$\sum_{i < i_2} P(E_i \cap E_{i_2}) \quad \dots \quad + (-1)^{r-1} \sum_{i < i_2 < i_r} P(E_i \cap E_{i_2} \dots \cap E_{i_r})$$

$$(E_i \cap E_{i_2} \dots \cap E_{i_r}) + (-1)^{n-r} P(E_1 \cap E_2 \dots \cap E_n)$$

$\sum_{i < i_2 < i_r} P(E_i \cap E_{i_2} \dots \cap E_{i_r})$ is taken over all the possible combination nCr of subset of size r from set n .

$$\begin{aligned} P(E \cup F \cup G) &= P(E) + P(F) + P(G) \\ &\quad - P(E \cap F) - P(E \cap G) - P(F \cap G) \\ &\quad + P(E \cap F \cap G) \end{aligned}$$

one at a time; then two; then three and so on

Normally we assume that o/p's are equally likely.

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→ sample space having equally likely events :-

consider a finite sample space $S = \{1, 2, \dots, n\}$ having equally likely outcome.

$$P(\{1\}) = P(\{2\}) = \dots = P(\{n\})$$

$$P(\{1\} \cup \{2\} \cup \{3\} \dots \cup \{n\}) = P(\{1\}) + P(\{2\}) + P(\{3\}) + \dots + P(\{n\})$$

→ Because all these are mutually exclusive events

$$P(S) = 1 = n P(\{1\})$$

$P(\{1\}) = \frac{1}{n}$

→ Uniform Discrete Probability law.

E is an event in S.

$P(E) = \frac{\text{number of points in } E}{\text{number of points in } S}$
only for \leftarrow E = number of favourable outcomes to E
equally likely \rightarrow total no. of outcomes in the exp.

$$P(\{1, 2, 3\}) = \frac{3}{n}$$

equally likely

Q-1 If three balls are randomly drawn from a bag containing 6 white and 5 black ball what is the probability that one of the ball is white and the other 2 are black.

Solution:

$$\frac{6C_1 \times 5C_2}{11C_3} = \frac{60}{165}$$

E: Three balls are drawn one - white Two - Black

$$n(S) = 11C_3 = \frac{34 \times 10 \times 9}{3 \times 2} = 165$$

Q2

Two dice are rolled. Find the probability of getting an even no. on the 1st dice or a total of 8.

→ Two events

$$\frac{18}{36} + \frac{5}{36} - \frac{3}{36} = \frac{20}{36}$$

we want $P(E_1 \cup E_2)$ in this

$$P(E_1) + P(E_2) - P(E_1 \cap E_2)$$

Q Matching - Problem

- 1) Let us have n letters corresponding to which there exists n envelopes having different addresses. Considering various letters being put in various envelopes at Random. A Match is said to occur if a letter goes into the right envelope.
- 2) If in part (b) there are n persons with n different hats. A match is said to occur if in the process of selecting hats at Random. The i^{th} person rightly gets the i^{th} hat.
- 3) Suppose n male and female couples are at the party and that the males and females are randomly paired for a dance. A Match occurs if a couple happens to be paired together.

Q1

- 1) What is the probability that none of the letter occupies right envelope.
- 2) What is the probability that exactly r letters are kept in correct envelope.

Define E_i^o = a Match occurs at i^{th} place

= i^{th} letter goes to i^{th} envelope

$$P(E_i^o) = \frac{(n-1)!}{(n)} = \frac{1}{n} \quad \begin{matrix} \text{only one letter goes to} \\ \text{right place} \end{matrix}$$

$i=1, 2, 3, \dots, n$ ways

none of the letters
goes to Right place

$$P(\bar{E}_1 \cap \bar{E}_2 \cap \dots \cap \bar{E}_n) = 1 - P(E_1 \cup E_2 \cup E_3 \cup \dots \cup E_n)$$

$$\Rightarrow 1 - \left(\sum_{i=1}^n P(E_i) - \sum_{i_1 < i_2} P(E_{i_1} \cap E_{i_2}) + (-1)^{n-1} \sum_{i_1 < i_2 < \dots < i_r} P(E_{i_1} \cap E_{i_2} \cap \dots \cap E_{i_r}) \right)$$

2 letters are
going right
envelope

$$P(E_1 \cap E_2) = \frac{(n-2)!}{n!} = \frac{1}{n(n-1)}$$

$$P(E_1 \cap E_2 \cap E_3) = \frac{(n-3)!}{n!} = \frac{1}{n(n-1)(n-2)}$$

$$P(E_1 \cap E_2 \cap \dots \cap E_n) = \frac{1}{n(n-1)(n-2) \dots 1} = \frac{1}{n!}$$

Putting all values in eqn ①

$$\rightarrow P(\bar{E}_1 \cap \bar{E}_2 \cap \dots \cap \bar{E}_n) = 1 - \left(n \cdot \frac{1}{n!} - \frac{nC_2}{n!} \cdot \frac{1}{n(n-1)} + \frac{nC_3}{n!} \cdot \frac{1}{n(n-1)(n-2)} - \dots - \frac{(-1)^{n-1}}{n!} \right)$$

$$\Rightarrow 1 - \left[1 - \frac{1}{2!} + \frac{1}{3!} - \dots - \frac{(-1)^{n-1}}{n!} \right]$$

$$\Rightarrow \sum_{R=0}^n \frac{(-1)^R}{R!}$$

Solution $R=0$ & $R=1$ term will get cancelled

$$\text{If } n \text{ is large } 1 - \frac{1}{1!} + \frac{1}{2!} - \frac{1}{3!} - \dots = e^{-1} = 0.367$$

Probability that atleast one letter is kept to Right envelope

$$\Rightarrow 1 - P(\bar{E}_1 \cap \bar{E}_2 \cap \bar{E}_3 \cap \dots \cap \bar{E}_n)$$

$$\Rightarrow 1 - 0.367 = 0.633$$

Matching Problem :-

E_i = Match occurrence at i^{th} place

$$P(E_i) = \frac{1}{n}$$

$$P(E_1 \cap E_2 \cap \dots \cap E_n) = \frac{1}{n!} \quad \text{all letters go into right place}$$

$$P(\text{no match}) = \sum_{k=0}^n \frac{(-1)^k}{k!}$$

$$P(\text{at least one match}) = 1 - \sum_{k=0}^n \frac{(-1)^k}{k!}$$

Probability (exactly r matches occurs) :-

$P(r$ letters going to their envelope) =

$$\rightarrow \frac{(n-r)!}{n!} = \frac{1}{n(n-1)\dots(n-r+1)}$$

$P(\text{remaining } n-r \text{ does not go to their own envelopes}) = \sum_{k=0}^{n-r} \frac{(-1)^k}{k!}$

$P(\text{only } r \text{ letters going to their envelopes}) =$
 we are selecting $\begin{cases} 1 \\ n(n-1) \dots (n-r+1) \end{cases}$ from n
 only one set of r from n

$$P(\text{required Probability}) = \frac{n(r)}{n(n-1)\dots(n-r+1)} \sum_{k=0}^{n-r} \frac{(-1)^k}{k!}$$

${}^n C_r$ possibilities

$$= \frac{1}{r!} \sum_{k=0}^{n-r} \frac{(-1)^k}{k!}$$

Birthday Problems :-

If n persons are present in room what is the probability that no two persons have their Birthday at the same day of a year?

- * each person can have 365 possibilities to celebrate his/her Bday.

sample space $n(s) = (365)^n$

$$n(E) = (365)C_n n! \rightarrow \text{no. of events}$$

$$\rightarrow P(\text{no two person have same B-day}) = \frac{365 C_n (n)!}{(365)^n}$$

$$\rightarrow \frac{(365)(364)\dots(365-n-1)}{(365)^n}$$

$$\rightarrow P(\text{at least two person have their b-day at same day}) = 1 - P[\text{no two person have b-day at same day}]$$

\rightarrow If $n \geq 23$ value of $P[\text{no two person have b-day at same day}]$ is less than half then $P[\text{at least two person --}]$ is greater than half.

\rightarrow If $n = 23$ At least two person from all of them have same Bday $P > \frac{1}{2}$.

\rightarrow If $n = 1$	$P(\text{At least two})$	= zero
$n = 23$	$P(n)$	= 50.2%
$n = 50$		= 97.1%
$n = 70$		= 99.9%
$n = 100$		= 99.9999%

$\rightarrow P = 100\%$, $n = 366$ person

If 10 Married couples are seating at Random around a Round table. Find the prob. That no wife sit next to her husband.

$E_i^o = i^{th}$ couple sit next to each other
 $n(S) = (19)!$

$$\begin{aligned} P(\text{no couple}) &= 1 - P\left[\bigcup_{i=1}^n E_i^o\right] \\ &= 1 - \left[\sum_{i=1}^n P(E_i^o) - \sum_{i_1 < i_2} P(E_{i_1}^o \cap E_{i_2}^o) \right. \\ &\quad \left. + (-1)^{n-1} P(E_1^o \cap E_2^o \cap \dots \cap E_n^o) \right] \end{aligned}$$

Probability n couples sit next to each other.

$P[E_1^o \cap E_2^o \cap \dots \cap E_n^o]$

Total 20 candidates ; considering each of the couple as a single entity

Only n couples are sitting ~~two~~ next to each other

$$\frac{20 - 2n}{20 - n} + n$$

$(19-n)!$ ways That the n couples sit next to each other.

$2^n (19-n)!$

each person have two choices

example → 3-couples

6 people A, B, C, D, E, F

couples

$n=2$

sit together

[CD]

[EF]

$6 - 4 + 2 = 4$

$20 - 2n + n$

$P(E_1^o \cap E_2^o \cap \dots \cap E_n^o) = \frac{2^n (19-n)!}{(19)!}$

n couples sitting next to each other

$$\begin{aligned}
 P(\text{no couples sit next to each other}) &= 1 - \left(\frac{10C_1}{10!} \cdot 2 \cdot (18)! + \frac{10C_2}{10!} \cdot \frac{2^2 \cdot (17)!}{19!} \right) \\
 &= 0.3395
 \end{aligned}$$

Probability as a continuous set function:

- A sequence of events $\{E_n\}_{n \geq 1}$ is called an increasing sequence if $E_1 \subset E_2 \subset E_3 \subset \dots \subset E_n \dots$

- We define a new event denoted as $\lim_{n \rightarrow \infty} E_n$ defined as

$$\lim_{n \rightarrow \infty} P(E_n) = P\left(\bigcup_{n=1}^{\infty} E_n\right)$$

- ### # A sequence of events $\{E_n\}_{n \geq 1}$ is called decreasing if $E_1 \supset E_2 \supset E_3 \supset \dots \supset E_n \dots$

- We define a new event denoted as $\lim_{n \rightarrow \infty} E_n$ & defined as :-

$$\lim_{n \rightarrow \infty} P(E_n) = P\left(\bigcap_{n=1}^{\infty} E_n\right)$$

Theorem : If $\{E_n\}_{n \geq 1}$ is either increasing or decreasing sequence of events then

$$\lim_{n \rightarrow \infty} P(E_n) = P\left(\lim_{n \rightarrow \infty} E_n\right)$$

Proof: Suppose $\{E_n\}_{n \geq 1}$ is an increasing function / sequence. Define an event

$$F_n = E_n \cap \left(\bigcup_{i=1}^{n-1} E_i \right)^c$$

F_n contains those outcomes of E_n which are not in earlier in any of the events.

E_n contains outcomes which are in the previous events.

Clearly event F_n contains of those outcomes of E_n that are not in any of the earlier Event E_i . where $i < n$.

$$F_n = E_n \cap E_{n-1}^c \quad \text{--- (1)}$$

$$\underline{F_n \cap F_{n+1}} = (E_n \cap E_{n-1}^c) \cap (E_{n+1} \cap E_n^c)$$

F_n are mutually exclusive events

$$\rightarrow \bigcup_{i=1}^{\infty} F_n = \bigcup_{i=1}^{\infty} E_n \quad \text{--- (2)}$$

$$\rightarrow \bigcup_{i=1}^n F_n = \bigcup_{i=1}^n E_n \quad \text{--- (3)}$$

Taking RHS

$$P(\lim_{n \rightarrow \infty} E_n) = P\left(\bigcup_{n=1}^{\infty} E_n\right)$$

$$= P\left(\bigcup_{n=1}^{\infty} F_n\right) \rightarrow F_n \text{ are all mutually exclusive}$$

By Axiom 3 \rightarrow

$$\sum_{n=1}^{\infty} P(F_n)$$

$$\rightarrow \lim_{n \rightarrow \infty} \sum_{n=1}^i P(F_n)$$

$$\rightarrow \lim_{i \rightarrow \infty} P\left(\bigcup_{n=1}^i F_n\right)$$

$$\rightarrow \lim_{i \rightarrow \infty} P\left(\bigcup_{n=1}^i E_n\right) = \lim_{i \rightarrow \infty} P(E_i)$$

case 2nd :-

If $\{E_n\}_{n \geq 1}$ is decreasing sequence.

$(E_n^c)_{n \geq 1}$ is an increasing sequence

$$P(\lim_{n \rightarrow \infty} E_n^c) = P\left(\bigcup_{i=1}^{\infty} E_i^c\right) = \lim_{n \rightarrow \infty} P(E_n^c)$$

$$= P\left(\bigcap_{i=1}^{\infty} E_i\right)^c = \lim_{n \rightarrow \infty} (1 - P(E_n))$$

$$= 1 - P\left(\bigcap_{i=1}^{\infty} E_i\right) = 1 - \lim_{n \rightarrow \infty} P(E_n)$$

$$\Rightarrow P\left(\bigcap_{i=1}^{\infty} E_i\right) = \lim_{n \rightarrow \infty} P(E_n)$$

$$\boxed{P(\lim_{n \rightarrow \infty} E_n) = \lim_{n \rightarrow \infty} P(E_n)}$$

conditional Probability :-

ex → Throwing two dice

(A) event → sum is even

$$\Omega = \{2, 3, \dots, 12\}$$

$$A = \{2, 4, 6, 8, 10, 12\}$$

$$P(A) = \frac{6}{11}$$

if is said that $\frac{\text{sum}}{\text{no}}$ is greater than 5.

event : (B) sum is greater than 5

$$\Omega = \{6, 7, \dots, 12\}$$

$$A = \{6, 8, 10, 12\}$$

$$\text{when } B \text{ has occurred } P(A) = \frac{4}{7}$$

event A depends upon event B.

conditional probability of A when B has occurred
 $P(A/B) = \text{Probability of A when B has occurred.}$

$$P(A/B) = \frac{P(A \cap B)}{P(B)}$$

Multiplication Rule :-

$$P(A \cap B) = P(A) \cdot P(B/A) = P(B) P(A|B)$$

Suppose E_1, E_2, \dots, E_n events.

$$P(E_1 \cap E_2 \cap \dots \cap E_n) = \frac{P(E_1) \cdot P(E_2)}{P(E_1)} \cdot \frac{P(E_3)}{P(E_1 \cap E_2)} \cdot \dots \cdot \frac{P(E_n)}{P(E_1 \cap E_2 \cap \dots \cap E_{n-1})}$$

$$\Rightarrow P(E_1) \cdot P(E_2/E_1) \cdot P(E_3/E_1 \cap E_2) \cdots P(E_n/(E_1 \cap E_2 \cap \dots \cap E_{n-1}))$$

Ques Two cards are drawn one after the another from a deck of 52 cards. Find the probability that both are spade cards if the 1st card is (i) Replaced or (ii) Not Replaced.

Sol : event A = first card is spade [2nd card is spade]
 event B =

$$P(A \cap B) = P(A) \cdot P(B/A)$$

$$= \frac{13}{52} \cdot \frac{13}{52} = \frac{1}{16}$$

(i) event A = event B

$$= \frac{13}{52} \times \frac{12}{51} = \frac{3}{51}$$

Q A die is thrown twice & the sum of the no. appearing is noted to be 8. what is the conditional prob. that no. 5 has appeared atleast once.

Sol:

A = the no. 5 appeared atleast once
 B = the sum is eight.

$$\Omega = 36$$

$$P(A/B) = \frac{P(A) \cdot P(B)}{P(A \cap B)}$$

$$= \frac{16}{36} \times \frac{2}{5}$$

$$A = \{(1,5), (2,5), (3,5), (4,5), (6,5), (5,1), (5,2), (5,3), (5,4), (5,5)\}$$

$$B = \{(2,6), (3,5), (4,4), (5,3), (6,2)\}$$

#

Total Probability Theorem :-

if E_1, E_2, \dots, E_n are events defined on S then $E_1, E_2, E_3, \dots, E_n$ are called exhaustive if

$$E_1 \cup E_2 \cup E_3 \cup \dots \cup E_n = S$$

Further;

$$E_i \cap E_j = \emptyset ; i \neq j$$

let B_1, B_2, \dots, B_n are mutually exclusive and exhaustive events on S with $P(B_n) \neq 0$ for $n = 1, 2, \dots, k$ let A be any random event on S

$$P(A) = \sum_{n=1}^k P(B_n) P(A|B_n)$$

$A \cap B_1, A \cap B_2, \dots, A \cap B_n$ are M.E.

$$A = (A \cap B_1) \cup (A \cap B_2) \cup \dots \cup (A \cap B_R)$$

$$P(A) = \sum_{i=1}^{\infty/R} P(A \cap B_i)$$

$$= \sum_{i=1}^{\infty/R} P(B_i) \cdot P(A|B_i)$$

Ques Box A contains 6 Red and 4 Blue balls
Box B contains 3 Red and 7 Blue.

Two balls are drawn at random from Box A
and placed in the Box B. Now a ball is drawn
at random from Box B. Find the prob. that
it is a Blue ball.

3 Possibilities

- Sol :- B1 i) Both are Red B2 ii) One Red - One Blue
B3 iii) Both are Blue

$$P(B_1) = \frac{6C_2}{10C_2} = \frac{1}{3} \quad P(B_2) = \frac{6C_1 \times 4C_1}{10C_2} = \frac{8}{15}$$

$$P(B_3) = \frac{4C_2}{10C_2} = \frac{2}{15}$$

Define an event A \rightarrow drawing a ^{Blue} ball from Box B
after the transfer.

- Both Red 5R, 7B \rightarrow $P(A|B_1) = \frac{7C_1}{12C_1} = \frac{7}{12}$

- One Red - One Blue 4R, 8B \rightarrow $P(A|B_2) = \frac{8}{12}$

- Both Blue 3R, 9B $= \frac{9}{12}$ $P(A|B_3)$

$$P(A) = \sum_{i=1}^{R/3} P(B_i) \cdot P(A|B_i)$$

$$\rightarrow \frac{1}{3} \times \frac{7}{12} + \frac{8}{15} \times \frac{8}{12} + \frac{2}{15} \times \frac{9}{12}$$

$$\Rightarrow 0.65$$

Baye's Theorem :-

Let B_1, B_2, \dots, B_R are mutually exclusive and exhaustive events on S with $P(B_i) \neq 0$ for $i = 1, 2, \dots, R$

Let A be any event on S with $P(A) \neq 0$

$$P\left(\frac{B_i}{A}\right) = P(B_i) \cdot P(A|B_i)$$

$$\sum_{i=1}^R P(B_i) \cdot P(A|B_i)$$

$$P(B_i \cap A) =$$

$$P(A)$$

Q A Bag contains 8 white balls and 4 black balls and Bag B \rightarrow 5 white, 6 black. One ball is drawn at random from Bag A and placed in B. Now a ball is drawn at random from Bag B. It is found that this ball is white. Find the Prob. That a black ball has been transferred from Bag A.

Sol : $B_1 =$ transferring a white ball to B] ?
 $B_2 =$ " " " " black ball to B .

$$P(B_1) = \frac{8}{12} = \frac{2}{3}$$

$$P(B_2) = \frac{4}{12} = \frac{1}{3}$$

$A =$ drawing a white ball from Box after transfer

$$P(A|B_1) = \frac{6}{12} = \frac{1}{2}$$

$$P(A|B_2) = \frac{5}{12}$$

$$P(B_2|A) = \frac{P(B_2) P(A|B_2)}{\sum_{i=1}^2 P(B_i) P(A|B_i)}$$

$$\Rightarrow \underline{\underline{5/17}}$$

Ques 3 cards are drawn at random from a deck without replacement. Find the probability that none of the 3 cards is a heart.

Sol $A_i = \{i^{\text{th}} \text{ card is not a heart}\}$

$$\begin{aligned} P(A_1 \cap A_2 \cap A_3) &= P(A_1) P\left(\frac{A_2}{A_1}\right) P\left(\frac{A_3}{A_1 \cap A_2}\right) \\ &= \left(\frac{39C_1}{52C_1}\right) \times \left(\frac{38C_1}{51C_1}\right) \times \left(\frac{37C_1}{50C_1}\right) \end{aligned}$$

$$\begin{aligned} \text{Probability Based} \\ \text{on Multiplication} \\ \text{Theorem} \end{aligned} = \left(\frac{39}{52}\right) \times \left(\frac{38}{51}\right) \times \left(\frac{37}{50}\right)$$

If Probability of A is not dependent on Probability of B Then $P(A|B) = P(A)$; A independent of B

$$P(A \cap B) = P(B) P(A|B) = P(B) P(A)$$

This Relation is also valid for $P(B) = 0$

= Two events A & B are Independent if

$$P(A \cap B) = P(A) P(B)$$

A/B events are disjoint (non empty)

$$A \cap B = \emptyset$$

$$P(A \cap B) = 0 \neq P(A) P(B)$$

$$P(A) \neq 0 \quad P(B) \neq 0$$

- if A and B are independent
- * A^c and B are also independent
- * A and B^c are also independent
- * A^c and B^c are also independent.

Prove A^c and B are independent.

$$\begin{aligned} B &= (A \cap B) \cup (A^c \cap B) \\ P(B) &= P(A \cap B) + P(A^c \cap B) \\ \Rightarrow P(A^c \cap B) &= P(B) - P(A) P(B) \\ \Rightarrow P(A^c \cap B) &= P(B) (1 - P(A)) \\ \Rightarrow P(A^c \cap B) &= P(A^c) P(B) \end{aligned}$$

 A^c and B are independent

$$\begin{aligned} \# \quad P(A^c \cap B^c) &= P(A \cup B)^c \\ &= 1 - P(A \cup B) \\ &= 1 - P(A) - P(B) + P(A \cap B) \\ &= 1 - P(A) - P(B) + P(A) P(B) \\ &= (1 - P(A)) (1 - P(B)) \\ \boxed{P(A^c \cap B^c)} &= P(A^c) P(B^c) \end{aligned}$$

example

$$\Omega = \{HH, HT, TH, TT\}$$

$$P(\omega) = \frac{1}{4} \rightarrow \text{prob. of each event}$$

$$A = \{HH, HT\} \quad P(A) = 1/2$$

$$B = \{HH, TH\} \quad P(B) = 1/2$$

$$C = \{HT, TT\} \quad P(C) = 1/2$$

$$A \cap B = \{HH\} \Rightarrow P(A \cap B) = 1/4$$

$$B \cap C = \{\emptyset\} \Rightarrow P(B \cap C) = 0$$

$$P(A) P(B) = \frac{1}{4} = P(A \cap B)$$

Independent of each other

$$P(B) P(C) = \frac{1}{4} \neq P(B \cap C)$$

Dependent on each other

Def: we say that events A_1, A_2, \dots, A_m are independent if for any $A_{i_1}, A_{i_2}, \dots, A_{i_m}$ from $\{A_1, A_2, \dots, A_m\}$ we have

$$P(A_{i_1} \cap A_{i_2} \cap \dots \cap A_{i_m}) = P(A_{i_1}) P(A_{i_2}) \dots P(A_{i_m})$$

for 3 events A, B, C

$$P(A \cap B) = P(A) P(B)$$

$$P(B \cap C) = P(B) P(C)$$

$$P(C \cap A) = P(C) P(A)$$

$$P(A \cap B \cap C) = P(A) P(B) P(C)$$

Pairwise

- ①

disjoint events

- ②

example: $\Omega = \{1, 2, 3, 4\}$

$$P(i) = \frac{1}{4} \quad i = 1, 2, 3, 4$$

$$A = \{1, 2\}$$

$$A \cap B = \{1\}$$

$$B \cap C = \{1\}$$

$$C \cap A = \{1\}$$

$$B = \{1, 3\}$$

$$\Rightarrow P(A \cap B) = 1/4$$

$$\Rightarrow P(B \cap C) = 1/4$$

$$C = \{1, 4\}$$

$$\Rightarrow P(C \cap A) = 1/4$$

$$P(A) = 1/2 = P(B) = P(C)$$

$$A \cap B \cap C = \{1\}$$

$$P(A \cap B \cap C) = 1/4$$

$$P(A) P(B) P(C) = \frac{1}{8}$$

It means these events are pairwise independent
But the eqn ② is not satisfied.

Q → consider & independent roll of a die & consider the following events.

$$A = \{ \text{1st roll is } 1, 2 \text{ or } 3 \}$$

$$B = \{ \text{1st roll is } 3, 4, 5 \}$$

$$C = \{ \text{sum of two roll is } 9 \}$$

$$A = \{ (1, i), (2, i), (3, i) \} \quad \begin{matrix} i = 1, 2, 3 \\ i = 6 \end{matrix}$$

$$B = \{ (3, i), (4, i), (5, i) \} \quad \begin{matrix} i = 1, 2, 3 \\ i = 6 \end{matrix}$$

$$C = \{ (5, 4), (4, 5), (6, 3), (3, 6) \}$$

$$A \cap B = \{ (3, i) \} \quad \begin{matrix} i = 1, 2, 3 \\ i = 6 \end{matrix}$$

$$B \cap C = \{ (5, 4), (4, 5), (3, 6) \}$$

$$C \cap A = \{ (3, 6) \}$$

$$A \cap B \cap C = \{ (3, 6) \}$$

$$P(A) = \frac{18}{36} = \frac{1}{2}$$

$$P(C) = \frac{4}{36} = \frac{1}{9}$$

$$P(B) = \frac{18}{36} = \frac{1}{2}$$

$$P(A \cap B \cap C) = \frac{1}{36}$$

$$P(A) P(B) P(C) = \frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{9} = \frac{1}{36}$$

In this eqⁿ ① { Pair conditions are not satisfied }

Q A problem in a question paper is given to 3 students in a class to be solved. The prob. of there solving the question is 0.5, 0.7, 0.8 resp. Find the prob that the problem will be solved.

sol

$$\begin{aligned}
 & 0.5 \times 0.3 \times 0.2 + 0.5 \times 0.7 \times 0.2 + 0.5 \times 0.3 \times 0.8 \\
 & + 0.5 \times 0.7 \times 0.2 + 0.5 \times 0.3 \times 0.8 + 0.5 \times 0.7 \times 0.8 \\
 & + 0.5 \times 0.7 \times 0.8
 \end{aligned}$$

→ P(Problem will be solved) = $1 - P(\bar{A} \cap \bar{B} \cap \bar{C})$

$$\begin{aligned}
 & = 1 - P(\bar{A}) P(\bar{B}) P(\bar{C}) \\
 & = 1 - (1 - P(A)) (1 - P(B)) (1 - P(C)) \\
 & = 1 - (0.5 \times 0.3 \times 0.2)
 \end{aligned}$$

#

Let I be an index set we say {A_i: i ∈ I} is independent if any finite sub collection is independent.

24/1/18
#

Random Variable :- Random variable is a mapping from the sample space → R.
• function from sample space to real.

ex-1

Tossing of 3 coin;

X = number of heads occurred

X = 0, 1, 2, 3

$$X = 0 = \{(T, T, T)\}$$

$$X = 1 = \{(HTT), (THT), (TTH)\} \xrightarrow{\text{Prob of this event}}$$

$$X = 2 = \{(HHT), (HTH), (THH)\}$$

$$X = 3 = \{HHH\}$$

- The value of the random variable is determined by the possible outcomes.

$$P(X=0) = \frac{1}{8}$$

$$P(X=1) = \frac{3}{8}$$

$$P(X=2) = \frac{3}{8}$$

$$P(X=3) = \frac{1}{8}$$

$$P\left(\bigcup_{i=0}^3 (x=i)\right) = \sum_{i=0}^3 P\{x=i\}$$

Q. 3 Balls are to be randomly selected without replacement from a bag containing 20 Balls no. 1 - 20. If we bet atleast one of the Ball that are drawn has a number as large as or greater than 17. What is the prob. that we win the bet?

Sol 1:- largest no. in the Ball is 17, 18, 19, 20

X = the largest Ball among selected.

$X = 3 - 20$ (because min no. of those 2 Balls can only be 1 or 2)

$$P\{X=i\} = \frac{i-1}{20} C_2$$

largest selected ball 3 is 17 $i = 3, 4, 5 - 20$

$$P\{X=17\} = \frac{16}{20} C_2 \quad \text{then the other two balls can be of Range 1-16}$$

$$P\{X=18\} = \frac{17}{20} C_2 \quad \text{Total Probability} = P(X=17) + P(X=18) +$$

$$P\{X=19\} = \frac{18}{20} C_2 \quad P(X=19) + P(X=20)$$

(To win the bet)

$$P\{X=20\} = \frac{19}{20} C_2$$

Q. 3 Balls are randomly chosen from a Bag containing 3 W, 3 Red, 5 Black Balls. Suppose that we win moneys \$1 for each white Ball selected and lose \$1 for each red Ball selected. Find the Probability that we win the money.

Sol For Black No loss (Nothing)

X = denote the total winning amount

$$X = 0, \pm 1, \pm 2, \pm 3$$

$$\begin{aligned}
 P\{X=1\} + P\{X=2\} + P\{X=3\} \\
 P\{X=1\} = \frac{3C_1}{11C_3} + \frac{5C_2}{3C_2} + \frac{3C_1}{3C_1} = P\{X=-1\} \\
 P\{X=2\} = \frac{3C_2}{11C_3} = P\{X=-2\} \\
 P\{X=3\} = \frac{3C_3}{11C_3} = P\{X=-3\}
 \end{aligned}$$

Taking summation of all values

$$\text{Probability} = \frac{1}{3}$$

Properties of Random variable :-

- i) If X and Y are random variables (μv) then $X+Y$, $X-Y$, XY , X/Y ($Y \neq 0$) are also μv . Moreover $ax+by$ is μv , $a, b \in R$.
- ii) If X is μv , $f: R \rightarrow R$ is function then $f(X)$ is also μv .
- iii) If X and Y are two μv and f is a function of two variables then $f(X, Y)$ are also μv .

• Discrete Random variable :- A random variable whose range is either finite or countably infinite is DRY. otherwise called continuous Random variable. eg measurement of height

• Probability Mass function :- corresponding to only discrete Random function. It is denoted by

$$\text{PMF} \rightarrow f_X(\cdot)$$

In probability :-

$$f_X(x) = P\{X=x\}$$

Definition :- A real valued function (f) defined by $f(x) = P\{X=x\}$ is called Probability Mass function (pmf of drv).

ex

Tossing of 2 coins :-

$$\Omega = \{(HH), (TT), (HT), (TH)\}$$

$X = \text{no. of heads.}$

$$f_X(x) = \begin{cases} 1/4 & \text{if } x=0 \\ 2/4 = 1/2 & \text{if } x=1 \\ 1/4 & \text{if } x=2 \\ & \text{otherwise} \end{cases}$$

your choice
want to
define it or not

This the Prob. Mass function of X .

• Properties of Probability Mass function :-

→ Let 'f' be a PMF of a random variable X then

1) $f(x) \geq 0$ for all $x \in R$

2) $S = \{x \in R : f(x) > 0\}$ is either finite or
countably infinite

3) $\sum_{x \in S} f(x) = 1$
 $\rightarrow x$ is taking all the values of X

1). $f_X(x)$ is denoting Prob. which is always greater than 0

2). It is talking about Range of the Random variable
is discrete \rightarrow and it is always finite or countably
infinite

3). All the Probabilities are mutually exclusive so
Total Prob will always be one.

→ If you have a $f^n : R \rightarrow R$ which satisfy all
the above four properties then there exist a
Random variable X whose PMF is given by
 $f_X(x)$. So there must exist a Random
variable whose PMF is $f_X(x)$.