

# Probability 1

## Chapter 05 : Continuous Random Variables - Part 1

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(based on the notes of Prof. Davide La Vecchia)

Spring Semester 2021

# Objectives

- Present the Exponential distribution
- Understand the consequences of the transformation of Random Variables.

## 1 Exponential distribution

## 2 Variable Transformation

- Transformation of Discrete Random Variables
- Transformation through the CDF
- Transformation of Continuous Random Variables through the PDF

## 1 Exponential distribution

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# Exponential distribution

## Definition

Let  $X$  be a continuous random variable, having the following characteristics:

- $X$  is defined on the **positive real numbers**  $(0; \infty)$  i.e.  $\mathbb{R}^+$ ;
- the PDF and CDF are

$$f_X(x) = \lambda \exp\{-\lambda x\}, \lambda > 0; \quad F_X(x) = 1 - \exp(-\lambda x);$$

then we say that  $X$  has an **exponential distribution**.

We write  $X \sim \text{Exp}(\lambda)$ .

## Remark (Expectation and Variance)

For  $X \sim \text{Exp}(\lambda)$  we have that:

$$E[X] = \frac{1}{\lambda} \quad \text{and} \quad \text{Var}(X) = \frac{1}{\lambda^2}. \quad (1)$$

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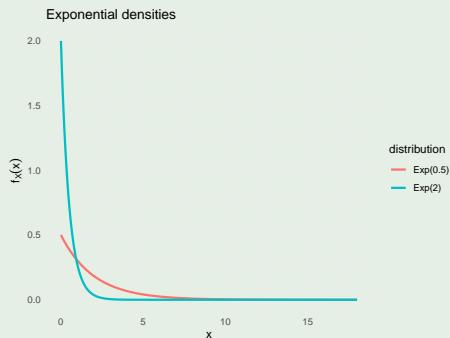
## Remark (Applications and Properties)

$X$  is typically applied to model the time until an event occurs, when events are always occurring at a rate  $\lambda > 0$ .

The sum of independent exponential random variables has a Gamma distribution (see the exercises).

# Exponential distribution

## Example (A graphical illustration)





## Example (Illustration of use (Spoiler of the Exercises!))

The lifetime  $X$  in years of a television follows an exponential law with density:

$$f(x) = \lambda e^{-\lambda x}, \quad x \geq 0.$$

1. Compute the cumulative distribution function  $F(x)$ .
2. Compute the  $\alpha$ -quantile  $q_\alpha = F^{-1}(\alpha)$ .
3. The expected life of your television is 8 years. What is the probability that the lifetime of your television is more than 8 years ? Evaluate the median.
4. Compute the variance of  $X$  for any  $\lambda$ .

# Exponential distribution

Unlike the Normal, the CDF of an exponential has a closed-form expression

## Example (CDF of Exponential)

Let  $X \sim \text{Exp}(\lambda)$ , with  $\lambda = 0.5$ . Thus

$$f_X(x) = \begin{cases} 0.5 \exp(-0.5x) & x > 0 \\ 0 & \text{otherwise} \end{cases}$$

Then, find the CDF.

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Then, find the CDF.

For  $x > 0$ , we have

$$\begin{aligned} F_X(x) &= \int_0^x f_X(u) du \\ &= 0.5 \left( -2 \exp(-0.5u) \right) \Big|_{u=0}^{u=x} \\ &= 0.5(-2 \exp(-0.5x) + 2 \exp(0)) \\ &= 1 - \exp(-0.5x) \end{aligned}$$

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so, finally,

$$F_X(x) = \begin{cases} 0 & x \leq 0 \\ 1 - \exp(-0.5x) & x > 0 \end{cases}$$

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# Variable Transformation

## Remark (Problem)

Consider a random variable  $X$  and suppose we are interested in  $Y = \psi(X)$ , where  $\psi$  is a one to one function

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## Definition

A **function**  $\psi(x)$  is **one to one** (1-to-1) if there are no two numbers,  $x_1, x_2$  in the domain of  $\psi$  such that  $\psi(x_1) = \psi(x_2)$  but  $x_1 \neq x_2$ .

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## Remark

*A sufficient condition for  $\psi(x)$  to be 1-to-1 is that it be monotonically increasing (or decreasing) in  $x$ .*



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## Remark

A sufficient condition for  $\psi(x)$  to be 1-to-1 is that it be monotonically increasing (or decreasing) in  $x$ . Note that the inverse of a 1-to-1 function  $y = \psi(x)$  is a

1-to-1 function  $\psi^{-1}(y)$  such that

$$\psi^{-1}(\psi(x)) = x \text{ and } \psi(\psi^{-1}(y)) = y.$$

To transform  $X$  to  $Y$ , we need to consider all the values  $x$  that  $X$  can take  
We first transform  $x$  into values  $y = \psi(x)$

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# Variable Transformation

## Transformation of Discrete Random Variables

- To transform a discrete random variable  $X$ , into the random variable  $Y = \psi(X)$ , we transfer the probabilities for **each**  $x$  to the values  $y = \psi(x)$ :

*Probability function for  $X$*

$X$	$P(\{X = x_i\}) = p_i$
$x_1$	$p_1$
$x_2$	$p_2$
$x_3$	$p_3$
$\vdots$	$\vdots$
$x_n$	$p_n$

$\Rightarrow$

*Probability function for  $X$*

$Y$	$P(\{X = x_i\}) = p_i$
$\psi(x_1)$	$p_1$
$\psi(x_2)$	$p_2$
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Probability function for $X$			Probability function for $X$	
$X$	$P(\{X = x_i\}) = p_i$		$Y$	$P(\{X = x_i\}) = p_i$
$x_1$	$p_1$	$\Rightarrow$	$\psi(x_1)$	$p_1$
$x_2$	$p_2$		$\psi(x_2)$	$p_2$
$x_3$	$p_3$		$\psi(x_3)$	$p_3$
$\vdots$	$\vdots$		$\vdots$	$\vdots$
$x_n$	$p_n$		$\psi(x_n)$	$p_n$

- Note that this is equivalent to applying the function  $\psi(\cdot)$  inside the probability statements:

$$\begin{aligned}P(\{X = x_i\}) &= P(\{\psi(X) = \psi(x_i)\}) \\&= P(\{Y = y_i\}) \\&= p_i\end{aligned}$$

# Variable Transformation

## Transformation of Discrete Random Variables

### Example (option pricing)

Let us imagine that we are tossing a balanced coin ( $p = 1/2$ ), and when we get a “Head” ( $H$ ) the stock price moves up of a factor  $u$ , but when we get a “Tail” ( $T$ ) the price moves down of a factor  $d$ . We denote the price at time  $t_1$  by  $S_1(H) = uS_0$  if the toss results in head ( $H$ ), and by  $S_1(T) = dS_0$  if it results in tail ( $T$ ). After the second toss, the price will be one of:

$$S_2(HH) = uS_1(H) = u^2S_0, \quad S_2(HT) = dS_1(H) = duS_0,$$

$$S_2(TH) = uS_1(T) = udS_0, \quad S_2(TT) = dS_1(T) = d^2S_0.$$

Indeed, after two tosses, there are four possible coin sequences,

$$\{HH, HT, TH, TT\}$$

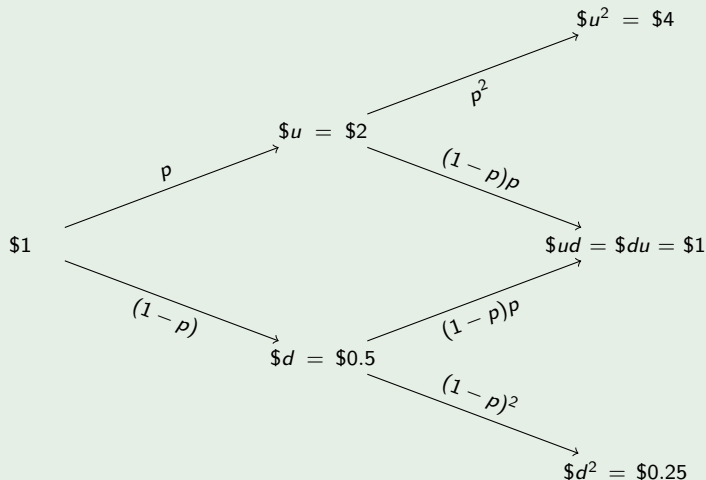
although not all of them result in different stock prices at time  $t_2$ .

# Variable Transformation

## Transformation of Discrete Random Variables

### Example (continued)

Let us set  $S_0 = 1$ ,  $u = 2$  and  $d = 1/2$ : we represent the price evolution by a tree:



# Variable Transformation

## Transformation of Discrete Random Variables

### Example (continued)

Now consider an European option call with maturity  $t_2$  and strike price  $K = 0.5$ , whose random pay-off at  $t_2$  is  $C = \max(0; S_2 - 0.5)$ . Thus,

$$\begin{aligned} C(HH) &= \max(0; 4 - 0.5) = \$3.5 & C(HT) &= \max(0; 1 - 0.5) = \$0.5 \\ C(TH) &= \max(0; 1 - 0.5) = \$0.5 & C(TT) &= \max(0; 0.25 - 0.5) = \$0. \end{aligned}$$

Thus at maturity  $t_2$  we have

Probability function for $S_2$			Probability function for $C$	
$S_2$	$P(\{X = x_i\}) = p_i$	$\Rightarrow$	$C$	$P(\{C = c_i\}) = p_i$
$\$u^2$	$p^2$		$\$3.5$	$p^2$
$\$ud$	$2p(1-p)$		$\$0.5$	$2p(1-p)$
$\$d^2$	$(1-p)^2$		$\$0$	$(1-p)^2$

Since  $ud = du$  the corresponding values of  $S_2$  and  $C$  can be aggregated, without loss of info.

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# Variable Transformation

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## Transformation through the CDF

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- Let  $Y = \psi(X)$  with  $\psi(x)$  **1-to-1 and monotone increasing**. Then

$$\begin{aligned}F_Y(y) &= P(\{Y \leq y\}) \\&= P(\{\psi(X) \leq y\}) = P(\{X \leq \psi^{-1}(y)\}) \\&= F_X(\psi^{-1}(y))\end{aligned}$$

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### Example

Let  $Y = \psi(X) = \exp X$  where  $X \sim F_X$  on all values  $x \in \mathbb{R}$

$$\begin{aligned}F_Y(y) &= P(\{Y \leq y\}) \\&= P(\{\exp X \leq y\}) = P(\{X \leq \ln(y)\}) \\&= F_X(\ln(y)) \text{ only for } y > 0.\end{aligned}$$

# Variable Transformation

## Transformation through the CDF

- **Monotone decreasing functions** work in a similar way, but require **changing the sense of the inequality**.
- Let  $Y = \psi(X)$  with  $\psi(x)$  1-to-1 and **monotone decreasing**. Then

$$\begin{aligned}F_Y(y) &= P(\{Y \leq y\}) \\&= P(\{\psi(X) \leq y\}) = P(\{X \geq \psi^{-1}(y)\}) \\&= 1 - F_X(\psi^{-1}(y))\end{aligned}$$

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### Example

Example: let  $Y = \psi(X) = -\exp X$  where  $X \sim F_X$  on all values  $x \in \mathbb{R}$

$$\begin{aligned}F_Y(y) &= P(\{Y \leq y\}) = P(\{-\exp^X \leq y\}) \\&= P(\{\exp X \geq -y\}) = P(\{X \geq \ln(-y)\}) \\&= 1 - F_X(\ln(-y)) \text{ only for } y < 0.\end{aligned}$$

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# Variable Transformation

## Transformation of Continuous Random Variables through the PDF

- For continuous random variables, if  $\psi(x)$  1-to-1 and monotone **increasing**, we have

$$F_Y(y) = F_X(\psi^{-1}(y))$$

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- Notice this implies that the pdf of  $Y = \psi(X)$  must satisfy

$$\begin{aligned} f_Y(y) &= \frac{dF_Y(y)}{dy} = \frac{dF_X(\psi^{-1}(y))}{dy} \\ &= \frac{dF_X(x)}{dx} \times \frac{d\psi^{-1}(y)}{dy} && \text{(chain rule)} \\ &= f_X(x) \times \frac{d\psi^{-1}(y)}{dy} && \text{(derivative of CDF (of } X) \text{ is pdf)} \\ &= f_X(\psi^{-1}(y)) \times \frac{d\psi^{-1}(y)}{dy} && \text{(substitute } x = \psi^{-1}(y) \text{)} \end{aligned}$$



# Variable Transformation

## Transformation of Continuous Random Variables through the PDF

- What happens when  $\psi(x)$  1-to-1 and monotone **decreasing**? We have

$$F_Y(y) = 1 - F_X(\psi^{-1}(y))$$

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- but  $\frac{d\psi^{-1}(y)}{dy} < 0$  since here  $\psi(\cdot)$  is monotone decreasing, hence we can write

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- This expression (called Jacobian-formula) is valid for  $\psi(x)$  1-to-1 and monotone (whether increasing or decreasing)

# Variable Transformation

## Transformation of Continuous Random Variables through the PDF

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## Transformation of Continuous Random Variables through the PDF

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- Recall that  $Y$  has a **lognormal distribution** when  $\ln(Y) = X$  has a Normal distribution
- $\Rightarrow$  if  $X \sim \mathcal{N}(\mu, \sigma^2)$ , then  $Y = \exp X \sim \text{lognormal}(\mu, \sigma^2)$

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  - Corresponding to  $\psi(x) = \exp x$  and  $\psi^{-1}(y) = \ln(y)$
- The *pdf* of  $X$  is

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2}(x - \mu)^2\right\}$$

for any  $-\infty < x < \infty$

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- Using  $\psi(x) = \exp x$  we know we'll have possible values for  $Y$  only on  $0 < y < \infty$

# Variable Transformation

## Transformation of Continuous Random Variables through the PDF

### Example (continued)

- We know that

$$f_Y(y) = f_X(\psi^{-1}(y)) \times \left| \frac{d\psi^{-1}(y)}{dy} \right|$$

# Variable Transformation

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- $\Rightarrow$  the *pdf* of  $Y$  is

$$f_Y(y) = \frac{1}{y\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{1}{2\sigma^2} (\ln(y) - \mu)^2 \right\}$$

for any  $0 < y < \infty$

# Variable Transformation

## Transformation of Continuous Random Variables through the PDF

### Example (continued)

- Both the Normal and the lognormal are characterized by only two parameters ( $\mu$  and  $\sigma$ ). The *median* of the lognormal distribution is  $\exp \mu$ , since

$$P(\{X \leq \mu\}) = 0.5,$$

and hence

$$\begin{aligned} 0.5 &= P(\{X \leq \mu\}) \\ &= P(\{\exp X \leq \exp \mu\}) \\ &= P(\{Y \leq \exp \mu\}). \end{aligned}$$

# Variable Transformation

## A Word of Warning

When  $X$  and  $Y$  are two random variables, **we should pay attention to their transformations.**

For instance, let us consider

$$X \sim \mathcal{N}(\mu, \sigma^2) \quad \text{and} \quad Y \sim \text{Exp}(\lambda).$$

Then, let's transform  $X$  and  $Y$

- in a linear way:  $Z = X + Y$ . We know that

$$E[Z] = E[X + Y] = E[X] + E[Y]$$

- in a nonlinear way  $W = X/Y$ . One can show that

$$E[W] = E\left[\frac{X}{Y}\right] \neq \frac{E[X]}{E[Y]}.$$

# Variable Transformation

## The big picture

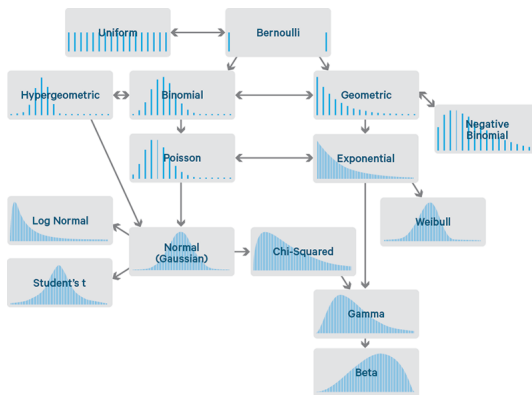
Despite exotic names, the common distributions relate to each other in intuitive and interesting ways. Several follow naturally from the Bernoulli distribution, for example.



# Variable Transformation

The big picture

- ▷ 'Common probability distributions: the data scientist's crib sheet'  
([goo.gl/NJRIXn](https://goo.gl/NJRIXn)):



- The Exponential distribution helps to model *duration* data.
- To compute the probabilities of transformed Discrete Random Variables we proceed on the Probability Function:

$$P(X = x) = P(\psi(X) = \psi(x)) = P(Y = \psi(x))$$

- This principle applies to the inequalities in the CDF (give or take the sense of the monotonicity)

$$F_X(x) = P(X \leq x) = P(\psi(X) \leq \psi(x)) = P(Y \leq \psi(x)) = F_Y(\psi(x))$$

$$F_X(x) = P(X \leq x) = P(\psi(X) \geq \psi(x)) = P(Y \geq \psi(x)) = 1 - F_Y(\psi(x))$$

- The density of a transformed variable can be found with the formula:

$$f_Y(y) = f_X(\psi^{-1}(y)) \times \left| \frac{d\psi^{-1}(y)}{dy} \right|$$

Thank You for your Attention!

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“See you” Next Week